



矩形分布：

$$T_1 \quad M_1 = \mathbb{E}X = np$$

$$p = \frac{n}{k} \quad M_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

$$p = \frac{\bar{x}}{n}$$

概率似然估计： $(x_1, x_2, \dots, x_n)$  由  $(X_1, X_2, \dots, X_n)$

- 求最大似然估计。

$$x_i = 1, 0$$

$$P(X=x) = p^x (1-p)^{1-x}$$

$$L(p) = \prod_{i=1}^n p^{x_i} (1-p)^{1-x_i}$$

$$= p^{\sum_{i=1}^n x_i} (1-p)^{n - \sum_{i=1}^n x_i}$$

$$\ln L(p) = (\sum_{i=1}^n x_i) \ln p + (n - \sum_{i=1}^n x_i) \ln (1-p)$$

$$\frac{d \ln L(p)}{dp} = \frac{\sum_{i=1}^n x_i}{p} + \frac{n - \sum_{i=1}^n x_i}{1-p} = 0$$

$$p = \frac{1}{n} \sum_{i=1}^n x_i$$

$$p = \frac{1}{n}$$

T<sub>2</sub> 矩形分布：

$$M_1 = \mathbb{E}X = \frac{\alpha+1}{\alpha+2}$$

$$1 - \frac{1}{\alpha+2} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\alpha = \frac{1}{1 - \frac{1}{n} \sum_{i=1}^n x_i} - 2$$

$$= \frac{2\bar{x} - 1}{1 - \bar{x}}$$

最大似然估计：

$(x_1, x_2, \dots, x_n) \rightarrow (X_1, X_2, \dots, X_n)$  - 每次出现

$$L(\alpha) = (\alpha+1)^n \left( \prod_{i=1}^n x_i \right)^\alpha$$

$$\ln L(\alpha) = n \ln(\alpha+1) + \alpha \ln \left( \prod_{i=1}^n x_i \right)$$

$$\frac{d \ln L(\alpha)}{d \alpha} = 0$$

$$\frac{n}{\alpha+1} = -\frac{1}{\ln \left( \prod_{i=1}^n x_i \right)}$$

$$\alpha = -n \ln \left( \prod_{i=1}^n x_i \right) - 1$$

$$\therefore \alpha = -n \sum_{i=1}^n \ln x_i - 1$$

$$\begin{aligned} T_3 & \int_0^{+\infty} \exp \left\{ -\frac{1}{2\sigma^2} (\ln x - \mu)^2 \right\} dx \\ & = \int_0^{+\infty} x \exp \left\{ -\frac{1}{2\sigma^2} (\ln x - \mu)^2 \right\} d \ln x \\ & = \int_{-\infty}^{+\infty} e^t \exp \left\{ -\frac{1}{2\sigma^2} (t - \mu)^2 \right\} dt \\ & = \int_{-\infty}^{+\infty} e^{-\frac{1}{2\sigma^2} t^2 + (\frac{\mu}{\sigma^2} + 1)t - \frac{\mu^2}{2\sigma^2}} dt \\ & = e^{\mu + \frac{\sigma^2}{2}} \int_{-\infty}^{+\infty} e^{-\frac{1}{2\sigma^2} t^2} dt \\ & = \sqrt{2\pi} e^{\mu + \frac{\sigma^2}{2}} \end{aligned}$$

$$\mu_1 = \sum_{i=1}^n x_i$$

$$\mu_2 = \sum_{i=1}^n x_i^2$$

$$\mu_1 = e^{\mu + \frac{\sigma^2}{2}}$$

$$\begin{aligned}
& \int_0^{+\infty} x^2 \exp \left\{ -\frac{1}{2\sigma^2} (\ln x - \mu)^2 \right\} d \ln x \\
&= \int_{-\infty}^{+\infty} e^{2t} \cdot e^{-\frac{1}{2\sigma^2}(t-\mu)^2} dt \\
&= \int_{-\infty}^{+\infty} e^{-\frac{1}{2\sigma^2}(t^2 - 2(\mu + \sigma^2)t + \mu^2)} dt \\
&= \int_{-\infty}^{+\infty} e^{-\frac{1}{2\sigma^2}(t - \mu - \sigma^2)^2} \cdot e^{2\sigma^2 + 2\mu} dt \\
&= \sqrt{2\pi\sigma} \cdot e^{2\sigma^2 + 2\mu} \\
\mu_2 &= e^{2\sigma^2 + 2\mu}
\end{aligned}$$

$$\sigma^2 = \ln \mu_2 - 2 \ln \mu_1$$

$$\mu = 2 \ln \mu_1 - \frac{1}{2} \ln \mu_2$$

$$\begin{aligned}
L(\mu, \sigma^2) &= (2\pi\sigma^2)^{-\frac{n}{2}} \cdot \frac{1}{\prod_{i=1}^n x_i} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^n (\ln x_i - \mu)^2 \right\} \\
\ln L &= -\frac{n}{2} \ln(2\pi\sigma^2) - \sum_{i=1}^n \ln x_i - \frac{1}{2\sigma^2} \sum_{i=1}^n (\ln x_i - \mu)^2 \\
\frac{\partial L}{\partial \mu} &= 0 \quad \frac{1}{\sigma^2} \sum_{i=1}^n (\mu - \ln x_i) = 0 \quad \mu = \frac{1}{n} \sum_{i=1}^n \ln x_i \\
&\quad \mu = \frac{1}{n} \sum_{i=1}^n \ln X_i \\
\frac{\partial L}{\partial \sigma^2} &= 0 \quad -\frac{n}{2\sigma^4} + \frac{1}{2\sigma^4} \sum_{i=1}^n (\ln x_i - \mu)^2 = 0 \\
&\quad \sigma^2 = \frac{1}{n} \sum_{i=1}^n (\ln x_i - \frac{1}{n} \sum_{i=1}^n \ln x_i)
\end{aligned}$$

T4 និច្ចបាត់ខាងក្រោម:

$$\begin{aligned}& \frac{1}{\theta} \int_m^{+\infty} x e^{-\frac{x-m}{\theta}} dx \\&= \theta \int_0^{+\infty} (m + \frac{m}{\theta}) e^{-m} e^{-\frac{x}{\theta}} dx \\&= \theta \left[ -m e^{-m} \Big|_0^{+\infty} + (1 - m/\theta) e^{-m} \Big|_0^{+\infty} \right] \\&= m + \theta\end{aligned}$$

$$\begin{aligned}& \frac{1}{\theta} \int_m^{+\infty} x^2 e^{-\frac{x-m}{\theta}} dx \\&= \theta \cdot e^{\frac{m}{\theta}} \int_{m/\theta}^{+\infty} m^2 e^{-m} dm \\&= \theta^2 \cdot e^{\frac{m}{\theta}} (-m^2 e^{-m} - 2me^{-m} - 2e^{-m}) \Big|_{m/\theta}^{+\infty} \\&= \theta^2 (m^2 + 2m\theta + 2\theta^2) \\&= m^2 + 2m\theta + 2\theta^2\end{aligned}$$

$$S^* = \sqrt{m^2 + 2m\theta + 2\theta^2 - (m + \theta)} = \theta$$

$$\theta = S^*$$

$$m = \bar{x} - S^*$$

概率密度函数： $(x_1, x_2 \dots x_n) \rightarrow (X_1, X_2 \dots X_n)$  - 问题

2. 由上式

$$L(\mu, \theta) = \frac{1}{\theta^n} \exp \left\{ -\left( \sum_{i=1}^n x_i - n\mu \right) / \theta \right\}$$

$$\frac{\partial L}{\partial \mu} = \frac{n}{\theta^{n+1}} \exp \left\{ -\left( \sum_{i=1}^n x_i - n\mu \right) / \theta \right\} \quad \text{单侧估计}.$$

$$2 \because \mu \leq \min \{x_1, x_2 \dots x_n\}$$

$$\therefore \mu = \min \{x_1, x_2 \dots x_n\} \quad \mu = \min \{X_1, X_2 \dots X_n\}$$

$$\frac{\partial L}{\partial \theta} = -\frac{n}{\theta^{n+1}} \exp \left\{ -\left( \sum_{i=1}^n x_i - n\mu \right) / \theta \right\} + \frac{\sum_{i=1}^n x_i - n\mu}{\theta^{n+2}} \exp \left\{ -\left( \sum_{i=1}^n x_i - n\mu \right) / \theta \right\}$$

$$\left\{ \begin{array}{l} \frac{\partial L}{\partial \theta} = 0 \\ n(\theta + \mu) = \sum_{i=1}^n x_i \end{array} \right.$$

$$\theta = \bar{X} - \min \{x_1, x_2 \dots x_n\}$$

5. 估计量：

$$\mu_1 = \theta$$

$$\theta = \bar{X}$$

最大似然量： $(x_1, x_2 \dots x_n) \rightarrow (X_1, X_2 \dots X_n)$  - 问题

$$\theta \text{ 为 } \left[ \max \{X_1, X_2 \dots X_n\} - \frac{1}{2}, \min \{X_1, X_2 \dots X_n\} \right]$$

两个端点值。

T<sub>b</sub> 簡便計:

$$M_1 = \frac{3}{2} \theta$$

$$\frac{3}{2} \theta = \bar{X}$$

$$\theta = \frac{2}{3} \bar{X}$$

第2大統計計:

$$\theta = [\min\{x_1, x_2, \dots, x_n\}, \frac{1}{2} \max\{x_1, x_2, \dots, x_n\}]$$

T<sub>7</sub> 簡便計

$$3 - 4\theta = 2$$

$$\theta = \frac{1}{4}$$

最大似然法.

$$L(\theta) = \theta^2 \cdot [2\theta(1-\theta)]^2 \cdot \theta^2 (1-2\theta)^4$$

$$\ln L(\theta) = \ln 4 + 6 \ln \theta + 2 \ln(1-\theta) + 8 \ln(1-2\theta)$$

$$\frac{\partial}{\partial \theta} \frac{1}{1-\theta} + \frac{4}{1-2\theta} = 0$$

$$6\theta^2 - 9\theta + 3 + 2\theta^2 - \theta + 4\theta^2 - 4\theta = 0$$

$$12\theta^2 - 14\theta + 3 = 0$$

$$\theta = \frac{14 - 2\sqrt{13}}{24}$$

$$\theta = \frac{7 - \sqrt{13}}{12}$$

$$T_8 \quad P(X \geq 0) = e^{-\lambda}$$

$(x_1, x_2, \dots, x_n)$  为  $(X_1, X_2, \dots, X_n)$  的观测值.

$$L(\lambda) = \prod_{i=1}^n \frac{\lambda^{x_i}}{x_i!} e^{-\lambda}$$

$$\ln L(\lambda) = -n\lambda + (\sum_{i=1}^n x_i) \cdot \ln \lambda - \sum_{i=1}^n \ln(x_i!)$$

$$\frac{d \ln L(\lambda)}{d \lambda} = 0 \quad -n + \frac{1}{\lambda} \sum_{i=1}^n x_i = 0$$

$$\lambda = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\lambda = \bar{X}$$

$$P(X \geq 0) = e^{-\bar{X}}$$

$$T_9, \textcircled{1} \quad E(\bar{\mu}_1) = \frac{1}{4}E X_1 + \frac{3}{4}E X_2 \\ = \frac{1}{4}\mu + \frac{3}{4}\mu \\ = \mu$$

$$M = D(\hat{\theta}(X_1, X_2))$$

$$= \sigma_1^2 + \sigma_2^2$$

$$= \frac{1}{16} + \frac{9}{16}$$

$$= \frac{5}{8}$$

$$\textcircled{2} \quad E(\bar{\mu}_2) = \frac{1}{2}E X_1 + \frac{1}{2}E X_2 \\ = \frac{1}{2}\mu + \frac{1}{2}\mu \\ = \mu.$$

$$M = D(\hat{\theta}(X_1, X_2))$$

$$= \sigma_1^2 + \sigma_2^2$$

$$= \frac{1}{4} + \frac{1}{4}$$

$$= \frac{1}{2}$$

$$\textcircled{1} > \textcircled{2}$$

$$\begin{aligned} T_{10} \sigma^2 &= C^2 \sigma_1^2 + (1-C)^2 \sigma_2^2 \\ &= (\sigma_1^2 + \sigma_2^2) C^2 - 2\sigma_1^2 C + \sigma_2^2 \end{aligned}$$

$$C = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2} \quad \text{if}$$

$$\begin{aligned} T_{11} \Sigma &\left( \sum_{i=1}^{n-1} (X_{i+1} - X_i)^2 \right) \\ &= 2 \left[ \sum_{i=1}^n (X_i - \bar{X})^2 + \sum_{i=2}^{n-1} (X_i - \bar{X})^2 + 2 \sum_{i=1}^{n-1} (X_i - \bar{X})(X_{i+1} - \bar{X}) \right] \\ &= 2(n-1) \sigma^2 \\ \therefore C &= \frac{1}{2(n-1)} \end{aligned}$$

$T_{12} (x_1, x_2, \dots, x_n)$  为  $(X_1, X_2, \dots, X_n)$  的似然函数。

$$L(\sigma^2) = (2\pi\sigma^2)^{-\frac{n}{2}} \cdot e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2}$$

$$\ln L(\sigma^2) = -\frac{n}{2}(\ln 2\pi + \ln \sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2$$

$$\frac{\partial L}{\partial \sigma^2} = 0 \quad -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (X_i - \mu)^2 = 0$$

$$\sigma^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$

$$\sigma^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$

$$\sigma^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \mu)^2$$

$$\sigma^2 > \sigma^2$$

$$T_{13} \quad U = \frac{\bar{X} - \mu}{\sigma / \sqrt{n}} \sim N(0, 1)$$

$$P\left(-u_{\alpha/2} < \frac{\bar{X} - \mu}{\sigma / \sqrt{n}} < u_{\alpha/2}\right) = 1 - \alpha$$

$$\text{区间} \quad \left( \bar{X} - u_{\alpha/2} \frac{\sigma}{\sqrt{n}}, \bar{X} + u_{\alpha/2} \frac{\sigma}{\sqrt{n}} \right)$$

$$\frac{2\sigma u_{\alpha/2}}{\sqrt{n}} < L$$

$$n > \frac{4\sigma^2 u_{\alpha/2}^2}{L^2}$$

$$T_{14} \quad \frac{\sqrt{n}(\bar{X} - \mu)}{S} \sim t(n-1)$$

$$\text{区间} \quad \left( \bar{X} - t_{\alpha/2}(n-1) \frac{s}{\sqrt{n}}, \bar{X} + t_{\alpha/2}(n-1) \frac{s}{\sqrt{n}} \right)$$

$$\bar{X} = 6 \quad S^2 = 0.33$$

$$\text{区间} \quad (5.558, 6.442)$$

$$T_{15} \quad \frac{(n-1)S^2}{s^2} \sim \chi^2(n-1)$$

$$\text{区间} \quad \text{区间} \quad \left( \frac{(n-1)S^2}{\chi^2_{\alpha/2}(n-1)}, \frac{(n-1)S^2}{\chi^2_{1-\alpha/2}(n-1)} \right)$$

$$(6.99 \times 10^{-4}, 6.63 \times 10^{-3})$$

$$T_{16} \text{ 问题} (\bar{X} - \bar{Y} - t_{\alpha/2}(n_1 + n_2 - 2) S_w \sqrt{\frac{1}{n_1} + \frac{1}{n_2}},$$

$$\bar{X} - \bar{Y} + t_{\alpha/2}(n_1 + n_2 - 2) S_w \sqrt{\frac{1}{n_1} + \frac{1}{n_2}})$$

$$S_w = \sqrt{\frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1+n_2-2}}$$

即为  $(-0.002, 0.006)$

$$T_{17} \quad \frac{\bar{X} - \mu}{S/\sqrt{n}} \approx t(n-1) \quad \mu = \mu$$

$$P(-t_{\alpha/2}(n-1) < \frac{\bar{X} - \mu}{S/\sqrt{n}} < t_{\alpha/2}(n-1))$$

$$\text{问题} (\bar{X} - t_{\alpha/2}(n-1) \cdot \frac{S}{\sqrt{n}}, \bar{X} + t_{\alpha/2}(n-1) \cdot \frac{S}{\sqrt{n}})$$

即为  $(0.203, 0.255)$

$$T_{18} \quad \frac{\bar{X} - \lambda}{S/\sqrt{n}} \approx N(0, 1)$$

$$(-u_{\alpha/2} < \frac{\bar{X} - \lambda}{S/\sqrt{n}} < u_{\alpha/2}) \quad S = \bar{X}$$

$$\text{问题} (\bar{X} - u_{\alpha/2} \sqrt{\frac{\bar{X}}{n}}, \bar{X} + u_{\alpha/2} \sqrt{\frac{\bar{X}}{n}})$$