OSM Boot Camp Econ Notes

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• ! Note. Recall

$$\hat{\theta}_{\mathrm{MLE}} = \theta : \max_{\theta} \log \mathcal{L} \left(\mathbf{x} \mid \theta \right)$$

$$\hat{\theta}_{GMM} = \theta : \min_{\theta} \|m(\mathbf{x} \mid \theta) - m(\mathbf{x})\|$$

$$= \arg \min_{\theta} e(\mathbf{x} \mid \theta)^{\mathsf{T}} \mathbf{W} e(\mathbf{x} \mid \theta).$$

where we can set

$$m_1(\mathbf{x} \mid \mu, \sigma) = \begin{pmatrix} \mathrm{E}\left[\mathbf{x} \mid \mu, \sigma\right] \\ \mathrm{Var}\left[\mathbf{x} \mid \mu, \sigma\right] \end{pmatrix}$$

and

$$m_2(\mathbf{x} \mid \mu, \sigma)$$

• Topic. OLS.

- Consider

$$y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \varepsilon_i,$$

where

$$E\left[\varepsilon_{i}\right]=0$$

and

$$E[x_{ji}\varepsilon_i] = 0.$$

- Let

$$\hat{\theta}_{\mathrm{GMM}} = \theta : \min_{\theta} \varepsilon^{\mathsf{T}} \varepsilon.$$

This is just OLS regression, in the form of GMM.

- Now consider

$$E[x_{1i}\varepsilon_i]=0$$

and

$$E[x_{2i}\varepsilon_i] = 0.$$

- Computing a moment:

$$\varepsilon_i = y_i - \beta_0 - \beta_1 x_{1i} - \beta_2 x_{2i}.$$

Each error is a function of the data.

- Let

$$m_1(\mathbf{x} \mid \beta_0, \beta_1, \beta_2) = \frac{1}{N} \sum_{i=1}^{N}$$

= $\frac{1}{N} \sum_{i} (y_i - \beta_0 - \beta_1 x_{1i} - \beta_2 x_{2i})$
= 0.

MLE says that choose a set of params so that that sum adds up to zero.

$$m_2(\mathbf{x} \mid \beta_0, \beta_1, \beta_2) = \frac{1}{N} \sum_{i=1}^{N} x_{1i} (y_i - \beta_0 - \beta_1 x_{1i} - \beta_2 x_{2i})$$

= 0.

This is another moment condition.

$$m_3(\cdot) = \cdots x_{2i} \cdots$$

and so on.

• Topic. Brock-Mirman model.

$$1 = \frac{\beta \mathbf{E} \left[r_{t+1} u' \left(c_{t+1} \right) \right]}{u' \left(c_{t} \right)}$$

$$\implies \beta \mathbf{E} \left[\frac{r_{t+1} c_{t}}{c_{t+1}} \right] - 1 = 0.$$

Middle term is

$$\beta \mathbf{E} \left[\frac{\alpha \mathbf{e}^{z_{t+1}} k_{t+1}^{\alpha - 1} c_t}{c_{t+1}} \right] - 1 = 0.$$

This is equation (9) in the notes.

- Topic.
 - Consider the difference between

$$\hat{\boldsymbol{\theta}}_{\mathrm{MLE}} = \arg\max \boldsymbol{\theta} \log \mathcal{L} \left(\mathbf{x} \mid \boldsymbol{\theta} \right)$$

and

$$\hat{\boldsymbol{\theta}}_{\mathrm{GMM}} = \arg\min_{\boldsymbol{\theta}} \left\| \underbrace{m\left(\mathbf{x} \mid \boldsymbol{\theta}\right)}_{\mathsf{model}} - \underbrace{m(\mathbf{x})}_{\mathsf{data}} \right\|$$

- Topic.
 - Let

$$S \coloneqq \# \text{ of sims } (s)$$

$$\tilde{\mathbf{x}} \coloneqq \{\tilde{\mathbf{x}}_1, \dots, \tilde{\mathbf{x}}_S\} \implies \tilde{\mathbf{x}}_s = \begin{pmatrix} y_{1s} x_{11s} x_{21s} \\ \vdots \\ y_{is} x_{1is} x_{2is} \\ \vdots \\ y_{NS} x_{1NS} x_{2NS} \end{pmatrix}.$$

The model moments are

$$m\left(\tilde{\mathbf{x}}\mid\boldsymbol{\theta}\right) = \frac{1}{S}\sum_{s=1}^{S}m\left(\tilde{\mathbf{x}}_{s}\mid\boldsymbol{\theta}\right).$$

Note that S is usually a large number, like 10,000.

- Topic.
 - For SMM, we have

$$\hat{\boldsymbol{\theta}}_{\mathrm{SMM}} = \arg\min_{\boldsymbol{\theta}} \|m\left(\tilde{\mathbf{x}} \mid \boldsymbol{\theta}\right) - m(\mathbf{x})\|.$$

In the L^2 norm way,

$$\arg\min_{\boldsymbol{\theta}} e\left(\tilde{\mathbf{x}} \mid \boldsymbol{\theta}\right)^{\mathsf{T}} \mathbf{W} e\left(\tilde{\mathbf{x}} \mid \boldsymbol{\theta}\right),$$

where

$$e\left(\tilde{\mathbf{x}} \mid \boldsymbol{\theta}\right) := m\left(\tilde{\mathbf{x}} \mid \boldsymbol{\theta}\right) - m(\mathbf{x}).$$

• Topic.

- Taking draws from the truncated normal distribution.
 - Let the PDF be $\phi(\mathbf{x} \mid \boldsymbol{\theta})$ and the CDF be $\Phi(\mathbf{x} \mid \boldsymbol{\theta})$. To simulate a general distribution, here are the steps:
 - 1. Draw N values $u_i \sim \text{Unif}(0,1)$.
 - 2. Use $\Phi\left(\mathbf{x}\mid\boldsymbol{\theta}\right)$ to convert u_{i} to x_{i} (the implied values from this PDF) $\implies x_{i}\sim\phi\left(\mathbf{x}\mid\boldsymbol{\theta}\right)$.
 - Note. The SMM problem will be a bonus problem.

Mon, 23 Jul. 2018

- Topic. Lucas Tree Model.
- ! Note. Review of probability.
 - Start with probability space $Z := \{z_1, z_2, \dots, z_n\}$. Assume Z stays constant over time. Each z_i is mutually exclusive and exactly one must occur.
 - Take the infinite cartesian product of this set, Z^{∞} . We are interested in an infinite horizon. Call this set Ω
 - For $\omega \in \Omega$, we have $\omega = (z^1, z^2, z^3, \dots, z^t, \dots)$. Call this a path.
 - Random variable $X(\omega): \Omega \to \mathbb{R}$.
- \Diamond **Example.** Simple random variables.
 - Suppose $z^1 = z_3$. Then

$$X(\omega) = \{a \mid z^1 = a\}$$
$$= z_3.$$

Alternatively, let

$$X'(\omega) = \left\{ \left. a \right| z^2 = a \right\}.$$

Consider two different paths:

$$(z_1, z_2z_7, z_{14}, z_2, z_1, z_{25}, \ldots)$$

and

$$(z_1, z_7, z_{14}, z_2, z_{27}, \ldots)$$
.

These paths are isomorphic to decision/probability tree paths.

- \triangle **Def.** A **filtration** is a sequence of σ -algebras or partitions.
- \Diamond **Example.** Consider $Z = \{1, 2\}$. Two possible paths are

$$(2,1,2,2,1,1,1,2,1,\ldots)$$

and

$$(1,2,2,2,1,1,2,1,\ldots)$$
.

- We can partition this set into two possible sets:

$$\{\omega|\,z_1=1\}$$

and

$$\{\omega|\,z_1=2\}\,.$$

If we combined these two sets with \emptyset and Ω , we get a σ -algebra! Call this **S.A. 1.**

- Now consider S.A. 2. What could it be? We have

$$\{\omega | z_1 = 1, z_2 = 1\}, \dots, \{\omega | z_1 = 2, z_2 = 2\};$$

in total we have four sets. Now include \emptyset and Ω . If we include all the unions and cross-unions, we would also get a σ -algebra. In other words, this would be the σ -algebra generated by these sets. Note that this resultant σ -algebra is **finer** than S.A. 1. In other words, S.A. 1 is **coarser** than S.A. 2.

- We make a sequence of successively finer σ -algebras, as we learn more information. This would be called a **filtration**.
- Let $P\{1\} =: \pi(1) = \frac{1}{3}$ and $\pi(2) = \frac{2}{3}$, and the outcomes IID.
- Topic. Consider the investor's problem,

$$\max_{\mathbf{c}} \left\{ \mathbf{E} \left[\sum_{t=0}^{\infty} \beta^{t} u\left(c_{t}\right) \right] \right\}.$$

Let's say we have

Asset Price Dividend Asset 1
$$p_t$$
 $d_t(\omega)$,

where the price and the dividend are both random variables. The only things that we can consume are with the dividends that are paid out by these assets.

- Define θ_t as the household's portfolio at time t. This might be

$$\boldsymbol{\theta}_t = (\theta_{1t}, \theta_{2t}, \dots, \theta_{nt}),$$

i.e., how many shares of each stock you have in your portfolio.

Households take their wealth at time t and choose to eat some of it and invest the rest of it. At time
 t,

$$c_t + p_{1t}\theta_{1t} + p_{2t}\theta_{2t} + \cdots + p_{nt}\theta_{nt}$$
.

This represents how much the agent consumes and distributes investment across different assets at time t. Note that

$$c_t + p_{1t}\theta_{1t} + p_{2t}\theta_{2t} + \dots + p_{nt}\theta_{nt} \le (p_{1t} + d_{1t})\theta_{1,t-1} + \dots + (p_{nt} + d_{nt})\theta_{n,t-1}.$$

In vector notation,

$$\underbrace{c_t}_{\text{not a vector}} + \mathbf{p}_t' \boldsymbol{\theta}_t \le (\mathbf{p}_t' + \mathbf{d}_t') \, \boldsymbol{\theta}_{t-1}.$$

Rewriting, we have

$$c_t = \mathbf{p}_t' \boldsymbol{\theta}_{t-1} + \mathbf{d}_t' \boldsymbol{\theta}_{t-1} + \mathbf{p}_t' \boldsymbol{\theta}_t$$

= $\mathbf{p}_t' (\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}_t) + \mathbf{d}_t' \boldsymbol{\theta}_{t-1}.$

So our problem is

$$\max_{\boldsymbol{\theta}} \left\{ E \left[\sum_{t=0}^{\infty} \beta^{t} u \left(\underbrace{\mathbf{p}'_{t} (\boldsymbol{\theta}_{t-1} - \boldsymbol{\theta}_{t})}_{\uparrow} + \mathbf{d}'_{t} \boldsymbol{\theta}_{t-1} \right) \right] \right\}.$$

$$change in portfolio wealth$$

– So what is this individual going to do? Take the FOCs of the above with respect to θ_t :

$$E\left[\beta^{t}\left(-u'\left(\underbrace{c_{t}}\right)\mathbf{p}_{t}+\beta u'\left(c_{t+1}\right)\left(\mathbf{p}_{t+1}+\mathbf{d}_{t+1}\right)\right)\right]=\mathbf{0}^{\mathsf{T}}.$$

$$\uparrow \mathbf{p}_{t}'\boldsymbol{\theta}_{t-1}+\mathbf{d}_{t}'\boldsymbol{\theta}_{t-1}+\mathbf{p}_{t}'\boldsymbol{\theta}_{t}$$

For row i:

$$E\left[\beta^{t}\left(-u'\left(c_{t}\right)p_{it}+\beta u'\left(c_{t+1}\right)\left(p_{i,t+1}+d_{i,t+1}\right)\right)\right]=0.$$

Divide the LHS and the RHS by p_{it} :

$$E\left[\beta^{t}\left(-u'\left(c_{t}\right)+\beta u'\left(c_{t+1}\right)\underbrace{\frac{p_{i,t+1}+d_{i,t+1}}{p_{it}}}\right)\right]=0.$$

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Stochastic Discount Factor.

•

$$\begin{split} \mathbf{E} \left[-\beta^{t} u'\left(c_{t}\right) \mathbf{p}_{t} + \beta^{t+1} u'\left(c_{t+1}\right) \left(\mathbf{p}_{t+1} + \mathbf{d}_{t+1}\right) \right] &= 0 \\ \mathbf{E} \left[-u'\left(c_{t}\right) \mathbf{p}_{t} + \beta u'\left(c_{t+1}\right) \left(\mathbf{p}_{t+1} + \mathbf{d}_{t+1}\right) \right] &= 0 \\ \Longrightarrow -u'\left(c_{t}\right) + \beta \mathbf{E} \left[u'\left(c_{t+1}\right) R_{it} \middle| \Omega_{t} \right] &= 0 \\ \Longrightarrow u'\left(c_{t}\right) &= \beta \mathbf{E} \left[u'\left(c_{t+1}\right) R_{it} \middle| \Omega_{t} \right]. \end{split}$$

$$1 = E \left[\underbrace{\beta \frac{u'(c_{t+1})}{u'(c_t)}}_{\text{stochastic discount factor}} R_{it} \quad \middle| \Omega_t \right].$$

Note that for small agents, R_{it} is exogenous to choices. Also,

$$1 = \operatorname{E} \left[m_t R_{it} | \Omega_t \right]$$

$$1 = \operatorname{Cov} \left[m_t, R_{it} | \Omega_t \right] + \operatorname{E} \left[m_t | \Omega_t \right] \operatorname{E} \left[R_{it} | \Omega_t \right].$$

For now, write $\operatorname{Cov}\left[m_t, R_{it} \middle| \Omega_t\right] =: \operatorname{Cov}\left[m_t, R_{it}\right]$ for brevity. We have

$$\underbrace{\operatorname{Cov}\left[m_{t},R_{it}\right]+\operatorname{E}\left[m_{t}\right]\operatorname{E}\left[R_{it}\right]=\operatorname{Cov}\left[m_{t},R_{jt}\right]+\operatorname{E}\left[m_{t}\right]\operatorname{E}\left[R_{jt}\right]}_{\text{risk-return tradeoff}}$$

(for $i \neq j$ in general).

• Recall

$$m_t = \beta \frac{u'\left(c_{t+1}\right)}{u'\left(c_t\right)}$$

and

$$R_{it} = \frac{\mathbf{p}_{i,t+1} + \mathbf{d}_{i,t+1}}{\mathbf{p}_{it}}.$$

! Note. Small investors do not have control over R; they take it as given, and E[R] is a belief based on information that the investor has. Portfolio choices θ are encapsulated in m. As an investor, we are manipulation the Cov[m,R] term. In equilibrium, we all do that.

consider the FOC

$$-u'(c_t) + \beta \mathbf{E}\left[u'(c_{t+1})\left(\frac{1}{\mathbf{p}_{it}}\right)\right] = 0.$$

Rearranging, we have

$$\mathbf{p}_{it} = \beta \mathbf{E} \left[\frac{u'\left(c_{t+1}\right)}{u'\left(c_{t}\right)} \right].$$

This is the expected discount factor.

• Gross return:

$$R_{it} = \underbrace{1 + r_{it}}_{(*)}$$

$$= \frac{1}{\mathbf{p}_{it}}$$

$$= \left[\frac{1}{\beta \mathrm{E}\left[\frac{u'(c_{t+1})}{u'(c_t)}\right]}\right]$$

This is the implied rate of return by this model. This can be used to estimate the parameters of the SDF through GMM. This is a **moment condition**. Note that GMM was developed by Lars Hansen specifically to estimate these particular moment conditions. Also note that (*) is something that we can easily observe. We can use all the expected marginal utilities information to estimate what people's β is.

- We would assume that

$$u(c_{t+1}) = \frac{c_{t+1}^{1-\gamma}}{1-\gamma}.$$

• If we take the FOC

$$g(\gamma, \beta, R) = \beta E \left[\frac{u'(c_{t+1})}{u'(c_t)} R_{it} \right] - 1$$
$$= 0$$

as a moment condition. We will do this on the homework. Every period, calculate

$$\mathbb{E}\left[\frac{u'\left(c_{t+1}\right)}{u'\left(c_{t}\right)}R_{it}-1\right].$$

Figure out what the average is, adjust the γ , and figure out what γ would make this condition equal to zero. (This is Questions 1 & 2.)

Kyle (1985) Model

- Agents:
 - Market Makers: The ones submitting limit orders. We want to understand this party. To these guys, V is a random variable, but they know the distribution of V. They don't know which traders are informed or uninformed. They observe Y = X + U (the sum of informed and uninformed demand).
 - Informed Traders: The ones who actually know what V is exactly. They also know the distribution of V that market makers know. Their optimal demand is X(V).
 - Noise Traders or Uninformed/Liquidity Traders: Individuals that are trading in ways that are uncorrelated with the future value of the asset. There is no information in this trade. No "market timing" effects involved. They know nothing. They could be bad traders, or they could just be laypeople selling stock in order to gain idiosyncratic personal financial liquidity. We assume $U \sim \mathcal{N}\left(0, \sigma_u^2\right)$. Assume that U, V are uncorrelated.
- Let future value of asset is V. By "future" for this model, we refer to a small time horizon, i.e., less than a week or so. Assume $V \sim \mathcal{N}(p_0, \Sigma_0)$. The market makers know this distribution.

Key Concept 1: Price Function.

 \triangle **Def.** In **equilibrium**, a market maker sets a **price function** P(Y), where Y = X + U. The informed traders also have a **demand** such that $\operatorname{E}\left[(V - P(Y))(\underbrace{-Y}) \middle| Y\right] = 0$ (competitive risk-neutral market makers), and where X(U) maximizes $\operatorname{E}\left[(V - P(Y))X(U)\middle| V\right]$. In general, this is a very hard system of equations to solve.

• To solve for the equilibrium, we use guess and check. Assume that the equilibrium is $P(Y) = \mu + \lambda Y$. Assume linear. Note that this is the place where you would expect to see linearity—risk-neutral and linear-preference assumptions. Take this as given, and move directly to

$$E[(V - P(Y))X(U)|V].$$

They want to

$$\max_{X} \operatorname{E}\left[\left(V-\mu-\lambda\left(X+U\right)\right)X|\,V\right].$$

Note that for informed traders, U is the only thing that is unknown. Now some simplification:

$$= \mathbb{E}\left[VX - \mu X - \lambda X^2 - \lambda UX \middle| V\right]$$

$$= VX - \mu X - \lambda X^2 - \lambda X \underbrace{\mathbb{E}\left[U\middle| V\right]}_{=0}$$

$$= VX - \mu X - \lambda X^2.$$

Calculating the FOC and solving for optimal X,

$$V - \mu - 2\lambda X = 0$$

so

$$X = \frac{V}{2\lambda} - \frac{\mu}{2\lambda}$$

$$= -\frac{\mu}{2\lambda} + \frac{1}{2\lambda}V.$$

Note that the second derivative is negative if $\lambda > 0$. This is an important consideration.