W261 - Week 7
Distributed Supervised Machine Learning:
Part 2
Regularized Linear models

UC Berkeley - MIDS Summer 2023

Table of Contents

- Housekeeping
- Regularization
- Speeding up gradient descent
- Maximum likelihood Estimates (MLE) and probabilistic Loss functions [optional]

HW4 release and Due date

Week	Торіс	Deadlines
1	Intro to Machine Learning at Scale	
2	Parallel Computation Frameworks	HW 1 due Sunday midnight at the end of week 2
3	Map-Reduce Algorithm Design	
4	Intro to Spark/Map-Reduce with RDDs (part 1)	HW 2 due Sunday midnight at the end of this week
5	Intro to Spark/Map-Reduce with RDDs (part 2)	
6	Distributed Supervised ML (part 1)	HW 3 due Sunday midnight at the end of this week
7	Distributed Supervised ML (part 2)	
8	Big Data Systems and Pipelines	IW4 Due Weel
9	Graph Algorithms at Scale (part 1)	
10	Graph Algorithms at Scale (part 2)	

MidTerm Course Evaluation

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Take 10 minutes to complete now (due this week)

Please fill out the course evals over the next 24 course and for the program. Thank you!!

sible) here. These are important for the

https://course-

s.berkeley.edu/berkeley/



7: Distributed Supervised Machine Learning: Part II

Week 7

1h 0m Total Video Time

Course Content	Description
7.1 Weekly Introduction 7	Sequence 1m 6s View Sequence Outline
7.2 Loss Functions and General Framework for Gradient Descent	Interactive Video 16m 8s
7.3 Quiz 7-1	Sequence View Sequence Outline
7.4 Logistic Regression at Scale	Interactive Video 17m 23s
7.5 SVMs	Interactive Video 8m 34s
7.6 Quiz 7-2	Sequence View Sequence Outline
7.7 SGD-Based SVMs at Scale	Interactive Video 16m 22s

Regularized Linear Regression

3.5. LASSO Regression update via GD

Lasso Regression cost function

LASSO_MSE(
$$\theta$$
) = MSE(θ) + $\eta \sum_{i=1}^{n} |\theta_i|$

$$MSE(\theta) = \frac{1}{m} ||\mathbf{X} \cdot \theta - \mathbf{y}||_2^2 \text{ (i.e., the euclidean distance }^2)$$

$$= \frac{1}{m} [(\mathbf{X} \cdot \theta - \mathbf{y}) \cdot (\mathbf{X} \cdot \theta - \mathbf{y})]$$

Lasso Regression subgradient vector

$$\nabla_{\text{LASSO_MSE}}(\theta) = \nabla_{\text{MSE}}(\theta) + \eta \begin{pmatrix} 0 \\ \text{sign}(\theta_1) \\ \text{sign}(\theta_2) \\ \vdots \\ \text{sign}(\theta_n) \end{pmatrix} \quad \text{where } \text{sign}(\theta_i) = \begin{cases} -1 & \text{if } \theta_i < 0 \\ 0 & \text{if } \theta_i = 0 \\ +1 & \text{if } \theta_i > 0 \end{cases}$$

The gradient for Lasso Regression results in a vector from adding two vectors:

$$\nabla_{\text{LASSO_MSE}}(\theta) = \nabla_{MSE} + \eta * \nabla_{L1}$$

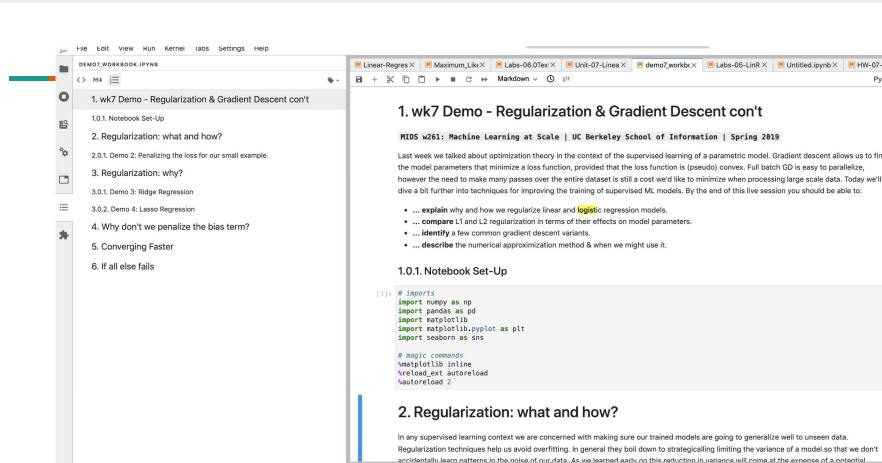
$$= \begin{bmatrix} \frac{\partial}{\partial \theta_0} \text{MSE}(\theta) \\ \frac{\partial}{\partial \theta_1} \text{MSE}(\theta) \\ \vdots \\ \frac{\partial}{\partial \theta_n} \text{MSE}(\theta) \end{bmatrix} + \eta * \begin{bmatrix} 0 \\ \text{sign}(\theta_1) \\ \vdots \\ \text{sign}(\theta_n) \end{bmatrix}$$

$$= \begin{bmatrix} \frac{2}{m} \mathbf{X_0}^T \cdot (\mathbf{X} \cdot \theta - \mathbf{y}) \\ \frac{2}{m} \mathbf{X_1}^T \cdot (\mathbf{X} \cdot \theta - \mathbf{y}) \\ \vdots \\ \frac{2}{m} \mathbf{X_n}^T \cdot (\mathbf{X} \cdot \theta - \mathbf{y}) \end{bmatrix} + \eta * \begin{bmatrix} 0 \\ \text{sign}(\theta_1) \\ \vdots \\ \text{sign}(\theta_n) \end{bmatrix}$$

$$= \frac{2}{m} \mathbf{X}^T \cdot (\mathbf{X} \cdot \theta - \mathbf{y}) + \eta * \nabla_{L1}$$

Gradient Descent step

$$\theta^{\text{(next step)}} = \theta - \eta \nabla_{\text{LASSO_MSE}}(\theta)$$



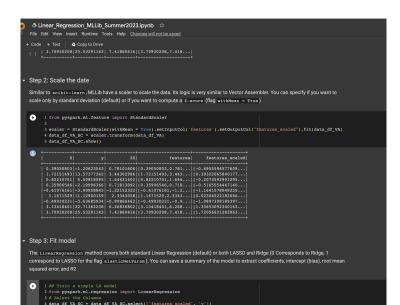
Important links for live session

The slides for module 7 are located here:

https://docs.google.com/presentation/d/1dbjllRxGnb7pcEx-HGT4igZxlidSvOV5jsnesSbIUKw/edit?usp=sharing

And a short supplemental Colab notebook on Dataframes, mllib/ml is located here:

https://colab.research.google.com/drive/1z-eHrD7Wu03DiX5hNros6lOinDgc3RkD?usp=sharing



Linear Classifiers and Maximum Likelihood

9.1. Deriving Maximum-Likelihood Estimates

Consider a training data set of the form $(X^{(1)}, y^{(1)}, \dots, (X^{(m)}, y^{(m)})$, where each $X^{(i)}$ is a an n-dimensional vector, and $y^{(i)}$ is the true label (aka ground truth label) of the corresponding example. Given a (learnt) logistic regression model one can calculate a probability for the ground truth class of each training example (as was defined above and presented here in context for convenience) as follows:

$$p(Y = y^{(i)} \mid x_1^{(i)}, \dots, x_n^{(i)})$$

Let's assume the training data are independently sampled from the same distribution; this is known as the *iid assumption*, which stands for "independent and identically distributed". Given an IID dataset, we can estimate the joint probability of the dataset as a function of the parameters of the learnt model as follows using the class probability associated with the ground truth class of each training example:

$$P(D|\theta) = \prod_{i=1}^{m} p(Y = y^{(i)} \mid x_1^{(i)}, \dots, x_n^{(i)})$$

We refer to this joint probability distribution, $P(D|\theta)$, as the **(conditional) likelihood** $L(\theta)$. Here for convenience we use θ to refer to the vector (for the binary classifier and a matrix for an n-ary classifier) of model coefficients, and bias term for all parameters $i \in 1, \ldots, n$. For a multinomial classifier, θ will be an $n \times K$ matrix for features $i \in 1, \ldots, n$ and classes $j \in [1, \ldots, K]$. The likelihood is a function of the parameter values, and the training examples. We usually work with the \log likelihood (because of numerical underflow issues), which is given by:

$$LL(\theta) \triangleq \log P(D|\theta) = \sum_{i=1}^{m} \log P(y^{(i)}|x^{(i)})$$

Here \triangleq denotes "is defined by". This decomposes into a sum of terms, one per example. Generally, we wish to maximize this quantity, i.e., for each example, we would desire for $p(Y = y^{(i)} \mid x_1^{(i)}, \dots, x_n^{(i)})$ to be maximum (or as close to 1 as possible for the ground truth class). This then leads to Maximum Likelihood Estimation (MLE) objective function that is given by:

$$\hat{\theta}_{MLE} = \arg\max_{\theta; D} \sum_{i=1}^{m} log P(y^{(i)} | x^{(i)})$$

$$\hat{\theta}_{MLE} = \arg\max_{\theta; D} \sum_{i=1}^{m} log P(y^{(i)} | x^{(i)})$$

Since most optimization algorithms and software implementations are designed to minimize objective functions, the MLE objective function is flipped to a minimization problem, commonly known as the (conditional) negative log likelihood or NLL as follows:

$$NLL(\theta; D) \triangleq -\log p(D|\theta) = -\sum_{i=1}^{m} log P(y^{(i)}|x^{(i)})$$

Minimizing this will also give the maximum likelihood estimate (MLE) of the model parameters.

$$\hat{\theta}_{MLE} = \arg\min_{\theta; D} - \sum_{i=1}^{m} log P(y^{(i)} | x^{(i)})$$
 (1)

In Equation (1), Vector θ , in the case of Naive Bayes, could refer to the model parameters $Pr(C_j)$ and $Pr(x_{ij|C=\nu_j})$ for $j \in [1, \dots, K]$ and $i \in 1, \dots, n$. In the case of logistic regression, θ would correspond to the the vector of coefficients and bias term for a binary classification proble,.

As an aside, if the model is unconditional (unsupervised, as in, say, the case of the EM clustering algorithm), the MLE becomes:

$$\hat{\theta}_{MLE} = \arg\min_{\theta; D} - \sum_{i=1}^{m} log P(x^{(i)}, \theta)$$
 (2)

In summary, maximum likelihood estimation (MLE) is a method of estimating the parameters of a model (of probability distributions in the case of Naive Bayes machine learnt model) by maximizing a likelihood function, so that under the assumed statistical model the observed data is most probable (i.e., the joint probability of the data is highest given by the likelihood function). The model, a point in the parameter space, that maximizes the likelihood function is called the maximum likelihood estimate. The logic of maximum likelihood is both intuitive and flexible, and as such the method has become a dominant means of statistical inference (think machine learning).

Most machine learning algorithms can be reframed probabilistically to leverage the maximum likelihood estimation framework. If the likelihood function is differentiable, the derivative test for determining minima (or maxima) can be applied. In some cases, the

Recall that in linear regression, the goal is to make predictions of the form

$$\hat{y} = \mathbf{x}^T \mathbf{w} + b,$$

where $\hat{y}, b \in \mathbb{R}$ and $w, x \in \mathbb{R}^n$. In contrast, in the case of classification, the interest is in asking the question "what is the probability that example x belongs to the positive class?" A regular linear model is a poor choice here because it can output values greater than 1 or less than 0. To coerce reasonable answers from our model (in the [0,1] interval), the linear regression model can be modified slightly, by running the linear function through a sigmoid activation function σ :

$$\hat{y} = \sigma(\mathbf{x}^T \mathbf{w} + b).$$

The sigmoid function σ , sometimes called a squashing function or a *logistic* function – thus the name logistic regression – maps a real-valued input to the range 0 to 1. Indeed, the logistic function $\sigma(z)$ is a good choice since it has the form of a probability, i.e. $\sigma(-z) = 1 - \sigma(z)$ and $\sigma(z) \in (0, 1)$ as $z \to \pm \infty$. Here z can be viewed as the perpendicular distance from a data point to a class hyperplane in the binary case. If target labels are set $y \in (0, 1)$ this will result in the following:

$$\mathbb{P}(y = 1|z) = \sigma(z) = \frac{1}{1 + e^{-z}}$$

$$\mathbb{P}(y = 0|z) = 1 - \sigma(z) = \frac{1}{1 + e^{z}}$$

which can be written more compactly as $\mathbb{P}(y|z) = \sigma(z)^y (1 - \sigma(z))^{1-y}$. The logistic function can be visualized as follows:

$$\prod_{i=1}^{n} \left(\sigma(z^{(i)}) \right)$$

 $= \prod^{n} P(y^{(i)} \mid x^{(i)}; \mathbf{w})$

 $L(\mathbf{w}) = P(\mathbf{y} \mid \mathbf{x}; \mathbf{w})$

$$=\prod_{i=1}^n \left(\sigma(z^{(i)})\right)^{y^{(i)}} \left(1-\sigma(z^{(i)})\right)^{1-y^{(i)}}$$
 In practice, it is easier to maximize the (natural) log of this equation, which is

called the log-likelihood function:

$$\log L(\mathbf{w})$$

$$l(\mathbf{w}) = \log L(\mathbf{w})$$

$$\Sigma \Gamma (i)$$
 , $\Gamma (i)$, $\Gamma (i)$, $\Gamma (i)$

$$\sum_{i=1}^{n} \left(\frac{1}{\sigma(z^{(i)})} \right) + \left(1 - \frac{1}{\sigma(z^{(i)})} \right) \log \left(1 - \frac{1}{\sigma(z^{(i)})} \right)$$

$$[u^{(i)} \log (\sigma(z^{(i)})) + (1 - u^{(i)}) \log (1 - \sigma(z^{(i)}))$$

$$= \sum [y^{(i)} \log (\sigma(z^{(i)})) + (1 - y^{(i)}) \log (1 - \sigma(z^{(i)}))]$$

Negative Log-Likelihood Loss

In practice, it is even more convenient to <u>minimize negative log-likelihood</u> instead of <u>maximizing log-likelihood</u>:

$$\mathcal{L}(\mathbf{w}) = -l(\mathbf{w})$$

$$= -\sum_{i=1}^{n} \left[y^{(i)} \log \left(\sigma(z^{(i)}) \right) + \left(1 - y^{(i)} \right) \log \left(1 - \sigma(z^{(i)}) \right) \right]$$

(in code, we also usually add a 1/n scaling factor for further convenience, where n is the number of training examples or number of examples in a minibatch)

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Since now linear classification is focussed on outputing class probabilities, one natural objective is to choose the weights (or parameters θ) that give the actual labels in the training data highest probability. For n samples $\{x_i, y_i\}$ the objective would boil down to maximizing the joint probability of the all the training data labels given their inputs:

$$\underset{\theta}{\operatorname{argmax}} \mathbb{P}_{\theta} (y_1, \ldots, y_n | \boldsymbol{x}_1, \ldots \boldsymbol{x}_n)$$

Because each example is independent of the others, and each label depends only on the features of the corresponding examples, the above can be simplified/decomposed and rewritten as follows:

$$\underset{\theta}{\operatorname{argmax}} \prod_{i}^{n} \mathbb{P}_{\theta}(y_{i} | \boldsymbol{x}_{i}) = \underset{\theta}{\operatorname{max}} \mathbb{P}_{\theta}(y_{1} | \boldsymbol{x}_{1}) \mathbb{P}_{\theta}(y_{2} | \boldsymbol{x}_{2}) \cdots \mathbb{P}_{\theta}(y_{n} | \boldsymbol{x}_{n})$$

where y_i corresponds to the class of the example, either positive or negative, \hat{p}_1 or \hat{p}_0 .

This function is a product over the training data. Since these predicted probabilies are less than or equal to one the product of many of such small numbers will get extremely small, so small that the numerical precision of modern day computers will run into numerical precision problems. To avoid such complications it is a lot easier to work with a loss function that, instead of working with the product of small numbers, works as a sum. This is accomplished by taking the log of the above equation. This results in the following:

$$\underset{\theta}{\operatorname{argmax}} \log \left(\prod_{i}^{n} \mathbb{P}(y_{i} | \mathbf{x}_{i}) \right) = \sum_{i}^{m} \log \left(\mathbb{P}(y_{i} | \mathbf{x}_{i}) \right) = \log \left(\mathbb{P}(y_{1} | \mathbf{x}_{1}) \right) + \dots + \log \left(\mathbb{P}(y_{n} | \mathbf{x}_{n}) \right)$$

The log transformation is monotonic and therefore does not change the dynamics/location of the solution to a maximization problem.

4.2. Learning a binomial logistic regression model via gradient descent using BXE (Binomial cross entropy)

The objective function for the learning a binomial logistic regression model (log loss) can be stated as follows:

$$\underset{\theta}{\operatorname{argmin}} \left[\operatorname{BXE} \right] = \underset{\theta}{\operatorname{argmin}} \left[-\frac{1}{m} \sum_{i=1}^{m} \left[y^{(i)} \log \left(\hat{p}^{(i)} \right) + (1 - y^{(i)}) \log \left(1 - \hat{p}^{(i)} \right) \right] \right]$$

• where probability of class 1:
$$|\hat{p}| = \mathbb{P}(y=1|X) = \sigma(X \cdot w + b) = \hat{p} = \sigma(t) = \frac{1}{1 + \exp(-t)}$$

The corresponding gradient function of partial derivatives is as follows (after a little bit of math):

$$\nabla_{\text{CXE}}(\theta) = \begin{pmatrix} \frac{\partial}{\partial \theta_0} \text{BXE}(\theta) \\ \frac{\partial}{\partial \theta_1} \text{BXE}(\theta) \\ \vdots \\ \frac{\partial}{\partial \theta_n} \text{BXE}(\theta) \end{pmatrix}$$
$$= \frac{1}{m} \mathbf{X}^T \cdot (\hat{p} - \mathbf{y})$$

For completeness learning a binomial logistic regression model via gradient descent would use the following step iteratively:

$$\theta^{\text{(next step)}} = \theta - \eta \nabla_{\text{BXE}}(\theta)$$

3.6. The gradient for Ridge Logistic Regression results is a vector from adding two vectors:** ¶

Binary Cross Entropy (BXE) function (log loss)

$$BXE(\theta) = -\frac{1}{m} \sum_{i=1}^{m} \left[y^{(i)} log \left(\hat{p}^{(i)} \right) + (1 - y^{(i)}) log \left(1 - \hat{p}^{(i)} \right) \right]$$
$$\frac{\partial}{\partial \theta_{j}} BXE(\theta) = \frac{1}{m} \sum_{i=1}^{m} x_{j}^{(i)} \left(\sigma(\mathbf{x}^{(i)} \theta) - y^{(i)} \right)$$
$$\nabla_{BXE}(\theta) = \frac{1}{m} \mathbf{X}^{T} \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y})$$

For more details on deriving the Gradient Equation for logistic regression see:

- Dan Jurafsky's Chapter on Logistic Regression is located here
- Or locally here.

$$\nabla_{\text{RIDGE_BXE}}(\theta) = \nabla_{BXE} + \eta * \nabla_{L2}$$

$$= \begin{bmatrix} \frac{\partial}{\partial \theta_0} \text{BXE}(\theta) \\ \frac{\partial}{\partial \theta_1} \text{BXE}(\theta) \\ \vdots \\ \frac{\partial}{\partial \theta_n} \text{BXE}(\theta) \end{bmatrix} + \eta * \begin{bmatrix} 0 \\ \theta_1 \\ \vdots \\ \theta_n \end{bmatrix}$$

$$= \begin{bmatrix} \frac{1}{m} \mathbf{X_0}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \\ \frac{1}{m} \mathbf{X_1}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \\ \vdots \\ \frac{1}{m} \mathbf{X_n}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \end{bmatrix} + \eta * \begin{bmatrix} 0 \\ \theta_1 \\ \vdots \\ \theta_n \end{bmatrix} \quad \text{where } \mathbf{X_i} \text{ denotes column } i \text{ of } \mathbf{X}$$

$$= \frac{1}{m} \mathbf{X}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) + \eta * \nabla_{L2}$$

Ridge Logistic Regression Gradient Descent step

$$\theta^{\text{(next step)}} = \theta - \eta \nabla_{\text{RIDGE_BXE}}(\theta)$$

$$\begin{split} BXE(\theta) &= -\frac{1}{m} \sum_{i=1}^{m} \left[y^{(i)} log \left(\hat{p}^{(i)} \right) + (1 - y^{(i)}) log \left(1 - \hat{p}^{(i)} \right) \right] \\ &\frac{\partial}{\partial \theta_{j}} \text{BXE}(\theta) = \frac{1}{m} \sum_{i=1}^{m} x_{j}^{(i)} \left(\sigma(\mathbf{x^{(i)}} \, \theta) - y^{(i)} \right) \\ &\nabla_{\text{BXE}}(\theta) = \frac{1}{m} \mathbf{X}^{T} \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \end{split}$$

3.7. The gradient for Lasso Logistic Regression results i a vector from adding two vectors:**

- Dan Jurafsky's Chapter on Logistic Regression is located here
- Or locally here.

$$\nabla_{\text{LASSO_BXE}}(\theta) = \nabla_{BXE} + \eta * \nabla_{L1}$$

$$= \begin{bmatrix} \frac{\partial}{\partial \theta_0} \text{BXE}(\theta) \\ \frac{\partial}{\partial \theta_1} \text{BXE}(\theta) \\ \vdots \\ \frac{\partial}{\partial \theta_n} \text{BXE}(\theta) \end{bmatrix} + \eta * \begin{bmatrix} 0 \\ \text{sign}(\theta_1) \\ \vdots \\ \text{sign}(\theta_n) \end{bmatrix}$$

$$= \begin{bmatrix} \frac{1}{m} \mathbf{X_0}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \\ \frac{1}{m} \mathbf{X_1}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \\ \vdots \\ \frac{1}{m} \mathbf{X_n}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) \end{bmatrix} + \eta * \begin{bmatrix} 0 \\ \text{sign}(\theta_1) \\ \vdots \\ \text{sign}(\theta_n) \end{bmatrix}$$

$$= \frac{2}{m} \mathbf{X}^T \cdot (\sigma(\mathbf{X} \cdot \theta) - \mathbf{y}) + \eta * \nabla_{L1}$$

where \mathbf{X}_i denotes column i of \mathbf{X}

$$\theta^{(\text{next step})} = \theta - \eta \nabla_{\text{LASSO_BXE}}(\theta)$$