GMMs by Mean Field Variational Inference

Jiajun He

In Bayesian framework, we want to find a set of parameters

$$\theta = \{\pi, z_1, z_2, ..., z_N, \mu_1, \Sigma_1, \mu_2, \Sigma_2, ..., \mu_K, \Sigma_K\}$$

, which maximizes the posterior probability $P(\boldsymbol{\theta}|\boldsymbol{X})$. But the form of the posterior distribution is relatively complicated in GMMs, so we use $q(\boldsymbol{\theta})$ to approximate it.

According to Bayes Rule, the probability of observing $X = \{x_1, x_2, ..., x_N\}$ is

$$P(\boldsymbol{X}) = \frac{P(\boldsymbol{\theta}, \boldsymbol{X})}{P(\boldsymbol{\theta}|\boldsymbol{X})} = \frac{P(\boldsymbol{\theta}, \boldsymbol{X})}{q(\boldsymbol{\theta})} \frac{q(\boldsymbol{\theta})}{P(\boldsymbol{\theta}|\boldsymbol{X})}$$

Take logarithm and expectation on both sides, we get

$$\int q(\boldsymbol{\theta}) \log P(\boldsymbol{X}) d\boldsymbol{\theta} = \int q(\boldsymbol{\theta}) \log \frac{P(\boldsymbol{\theta}, \boldsymbol{X})}{q(\boldsymbol{\theta})} d\boldsymbol{\theta} + \int q(\boldsymbol{\theta}) \log \frac{q(\boldsymbol{\theta})}{P(\boldsymbol{\theta}|\boldsymbol{X})} d\boldsymbol{\theta}$$

i.e.,

$$\log P(\boldsymbol{X}) = \int q(\boldsymbol{\theta}) \log \frac{P(\boldsymbol{\theta}, \boldsymbol{X})}{q(\boldsymbol{\theta})} d\boldsymbol{\theta} + D_{KL}[q(\boldsymbol{\theta})||p(\boldsymbol{\theta}|\boldsymbol{X})]$$

The tern on the left-hand side is a constant, so minimizing KL-divergence between $q(\theta)$ and real posterior distribution is equivalent to maximizing the first term on the right-hand side, which is named as ELBO.

Under mean field assumption, i.e., $q(\theta) = \prod_{\theta_i \in \theta} q(\theta_i)$, each variable can be optimized successively:

Specifically, for each $\theta_i \in \boldsymbol{\theta}$,

$$\begin{split} ELBO &= \int q(\boldsymbol{\theta}) \log(\frac{P(\boldsymbol{\theta}, \boldsymbol{X})}{q(\boldsymbol{\theta})}) \mathrm{d}\boldsymbol{\theta} \\ &= \int q(\boldsymbol{\theta}) \log P(\boldsymbol{\theta}, \boldsymbol{X}) \mathrm{d}\boldsymbol{\theta} - \int q(\boldsymbol{\theta}) \log q(\boldsymbol{\theta}) \mathrm{d}\boldsymbol{\theta} \\ &= \mathbb{E}_{\theta_i} [\mathbb{E}_{\boldsymbol{\theta} - \{\theta_i\}} [\log P(\boldsymbol{\theta}, \boldsymbol{X})]] - \mathbb{E}_{\boldsymbol{\theta}} [\log q(\theta_i)] - \sum_{j \neq i} \mathbb{E}_{\boldsymbol{\theta}} [\log q(\theta_j)] \\ &= \int q(\theta_i) \mathbb{E}_{\boldsymbol{\theta} - \{\theta_i\}} [\log P(\boldsymbol{\theta}, \boldsymbol{X})] \mathrm{d}\theta_i - \int q(\theta_i) \log q(\theta_i) \mathrm{d}\theta_i - \sum_{j \neq i} \int q(\theta_j) \log q(\theta_j) \mathrm{d}\theta_j \end{split}$$

The last term is a constant for θ_i , i.e.,

$$ELBO = \int q(\theta_i) \mathbb{E}_{\boldsymbol{\theta} - \{\theta_i\}} [\log P(\boldsymbol{\theta}, \boldsymbol{X})] d\theta_i - \int q(\theta_i) \log q(\theta_i) d\theta_i + const$$

As the expectation term is positive, we can view it as an log-probability $\log \tilde{P}$, then

$$ELBO = -D_{KL}[q(\theta_i)||\tilde{P}] + const$$

To maximum ELBO is equivalent to maximum the KL-divergence between each $q(\theta_i)$ and corresponding \tilde{P} .

Here, our task is just to find an MAP assignment, so to furthermore simplify the calculation, I set all q to be one-point distribution, in which situation, ELBO is maximized when

$$q(\theta_i) = \mathbb{I}\Big\{\theta_i = \text{mode}\{\tilde{P}\}\Big\}$$

Hereinafter, to simplify the notation, I will use $\theta_i = a$ to represent $q(\theta_i) = \mathbb{I}\{\theta_i = a\}$.

In our Gaussian Mixture Model, in each iteration step, I successively optimize π , μ_1 , Σ_1 , μ_2 , Σ_2 , ..., μ_K , Σ_K , z_1 , z_2 , ..., z_N as follows:

Optimize π :

All terms without π in it can be viewed as constants, and integrate out variables that do not occur in log term, we can get

$$\log q(\boldsymbol{\pi}) = \int q(z_1)q(z_2)...q(z_N) \log P(\boldsymbol{\pi})P(z_1|\boldsymbol{\pi})P(z_2|\boldsymbol{\pi})...P(z_N|\boldsymbol{\pi}) dz_1 dz_2...dz_N + const$$
$$= \log P(\boldsymbol{\pi}) + \log P(z_1, z_2, ..., z_N|\boldsymbol{\pi}) + const$$

The last step is correct as $q(z_i)$ is one-point distribution according to my assumption.

Considering that $\pi \sim \text{Dir}(\alpha)$ and $P(z_1, z_2, ..., z_N | \pi) \sim \text{Multi}(\pi)$, the right-hand side is just the posterior Dirichlet distribution. Therefore, $\pi = \text{mode} \left\{ \text{Dir}(\alpha_1 + N_1, \alpha_2 + N_2 ..., \alpha_K + N_K) \right\}$ where N_k is the number of observations in cluster k.

Optimize μ_i, Σ_i :

When optimizing μ_i , Σ_i , given that the entanglement of them does not bring us any inconvenience because we have a well-formed distribution over μ_i , Σ_i , I do not entangle them into two q. Similar to π , we can cancel out irrelevant variables and only get

$$\log q(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i) = \log P(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i) + \sum_{z_j = i} \log P(\boldsymbol{x}_j | \boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i) + const$$

The right-hand side is just the posterior Normal-Inverse-Wisart distribution. Thus, $\boldsymbol{\mu}_{j}, \boldsymbol{\Sigma}_{j} = \text{mode}\Big\{\mathcal{N}\mathcal{I}\mathcal{W}(\frac{\lambda\boldsymbol{\mu}_{0}+N_{j}\bar{\boldsymbol{x}}}{\lambda+N_{j}}, \lambda+N_{j}, \nu+N_{j}, \boldsymbol{\Psi}+\mathbf{C}+\frac{\lambda N_{j}}{\lambda+N_{j}}(\bar{\boldsymbol{x}}-\boldsymbol{\mu}_{0})(\bar{\boldsymbol{x}}-\boldsymbol{\mu}_{0})^{T})\Big\}$, where $\mathbf{C}=\sum_{i}(\boldsymbol{x}_{i}-\bar{\boldsymbol{x}})(\boldsymbol{x}_{i}-\bar{\boldsymbol{x}})^{T}$ and $\bar{\boldsymbol{x}}$ is the mean of samples in class j;

Optimize $z_1, z_2, ..., z_N$:

Via similar calculation, it is easy to find that each z_i can be optimized as $z_i = \max_{i=1}^k \{ \pi_j P(x_i | \mu_j, \Sigma_j) \}$.

Predict z for new x:

After finding the MAP assignments for all parameters (i.e., one-point posterior distribution), we can predict class labels for new input \boldsymbol{x} .