

# Master Thesis: Multilevel Multivariate Imputation by Chained Equations through Bayesian Additive Regression Trees

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**Heleen** Brüggem



**Utrecht  
University**

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**Supervisors:**

T. Volker MSc.

Dr. G. Vink

H. Oberman MSc.

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Utrecht University

Utrecht University

Utrecht University

# Contents

<b>1</b>	<b>Introduction</b>	<b>2</b>
<b>2</b>	<b>Method</b>	<b>3</b>
2.1	Theoretical background . . . . .	3
2.1.1	Bayesian Additive Regression Trees (BART) . . . . .	3
2.1.2	Multilevel-BART (M-BART) . . . . .	5
2.1.3	stan4bart . . . . .	5
2.2	Simulation study . . . . .	6
2.2.1	Data generating mechanism . . . . .	6
2.2.2	Simulation design . . . . .	7
2.2.3	Missing data generation . . . . .	8
2.2.4	Evaluation . . . . .	9
<b>3</b>	<b>Results</b>	<b>9</b>
3.1	Bias . . . . .	9
3.2	Coverage . . . . .	12
3.3	Confidence interval width . . . . .	13
<b>4</b>	<b>Discussion</b>	<b>13</b>
<b>5</b>	<b>Conclusion</b>	<b>13</b>
<b>6</b>	<b>Appendix</b>	<b>14</b>
	<b>References</b>	<b>19</b>

# 1 Introduction

Incomplete data is a common challenge in many fields of research. Frequently used ad hoc strategies to deal with missing data, such as complete case analysis or mean imputation, often lead to erroneous inferences in realistic situations. Missingness can follow a multivariate mechanism that may depend on observed data or even unobserved data, leading to biased estimates and inaccurate variance estimates when using one of these ad hoc strategies (Austin et al., 2021; Enders, 2017; Kang, 2013; Little and Rubin, 2002; van Buuren, 2018). Multiple imputation (MI; Rubin, 1987) is proven to be an effective method for dealing with multivariate incomplete data supported by a considerable amount of methodological research (Audigier et al., 2018; Austin et al., 2021; Burgette and Reiter, 2010; Enders, 2017; Grund et al., 2021; Hughes et al., 2014; Little and Rubin, 2002; Mistler and Enders, 2017a; Van Buuren, 2007; van Buuren, 2018).

MI separates the missing data problem from the analysis problem (Audigier et al., 2018; Austin et al., 2021; Bartlett et al., 2015; Burgette and Reiter, 2010; Carpenter and Kenward, 2013; Enders, 2017; Grund et al., 2021; Hughes et al., 2014; Little and Rubin, 2002; Mistler and Enders, 2017a; Van Buuren, 2007; van Buuren, 2018). A statistical model specifying the variables used for imputation, i.e. the imputation model, is defined for every variable with missing values. Each missing value in the dataset is imputed  $m$  times by drawing values from their posterior predictive distribution conditional on the observed data and parameters from the imputation model. By repeatedly drawing values from the posterior predictive distributions – in other words, the distribution of plausible replacement values – the necessary variation associated with the missingness problem is considered. After imputation, each of the imputed datasets are analyzed according to the model of interest, i.e. the substantive analysis model. Then, their  $m$  corresponding model parameters are pooled together according to Rubin’s rules (Rubin, 1987). One central requirement for MI is the concept of congeniality; the imputation model should be at least as general as the analysis model and preferably all-encompassing (Bartlett et al., 2015; Enders et al., 2018a; Grund et al., 2016, 2018b; Little and Rubin, 2002; Meng, 1994). If not, the imputation model will not be compatible with the analysis model and the pooled estimates of the latter may be biased.

When MI is applied in a multilevel data context, concerns regarding the concept of congeniality become more pronounced (Audigier et al., 2018; Dong and Mitani, 2023; Enders et al., 2020, 2018a,b, 2016; Grund et al., 2016, 2018a,b, 2021; Lüdtke et al., 2017; Mistler and Enders, 2017a; Quartagno and Carpenter, 2022; Resche-Rigon and White, 2018; Taljaard et al., 2008; van Buuren, 2018). Multilevel data is hierarchically structured, where, for example, students are nested within classes within schools or patients within hospitals (Hox and Roberts, 2011; Hox et al., 2017). When analyzing multilevel data, this hierarchical structure should be taken into consideration. Ignoring it will underestimate the intra-class correlation (ICC) and standard errors, as conventional statistical analyses assume independence of observations (Hox and Roberts, 2011; Lüdtke et al., 2017; Taljaard et al., 2008; van Buuren, 2018). The ICC can be interpreted as the proportion of the total variance at level-2 (Gulliford et al., 2005; Hox and Roberts, 2011; Shieh, 2012). Accounting for this structure, can be done using multilevel models (MLMs; Hox and Roberts, 2011; Hox et al., 2017; Lüdtke et al., 2017). MLMs can contain variables relating to the individual level – level-1 variables – or to the grouping structure – level-2 variables or potentially higher order structures. For example, imagine a case where students are nested within classes. Here, the academic performance of a student is a level-1 variable, whereas the teacher’s experience is a level-2 variable. Additionally, MLMs allow you to specify random intercepts, indicating that some classes have students that significantly perform better or worse academically on average; random slopes, indicating that the relationship between the performance of students and the outcome variable differs between classes; and cross-level interactions, indicating that the effect of performance of students can differ with the teacher’s experience (Hox and Roberts, 2011; Hox et al., 2017). Typically, the complexity of the multilevel analysis model is built step-wise with non-linearities, meaning the analysis model is not determined beforehand: predictors, random intercepts, random slopes, and cross-level interactions are added in a stepwise manner to the model (Hox and Roberts, 2011; Hox et al., 2017). Thus, ensuring congeniality for the imputation model can be complex, since the final analysis model is not pre-determined. Furthermore, including the hierarchical structure along with cross-level interactions or other complicated non-linearities in imputation models is quite challenging (Burgette and Reiter, 2010; Hox and Roberts, 2011; van Buuren, 2018), also because very complex models might not converge (van Buuren, 2018).

A popular and flexible implementation of MI in a multilevel context, is fully conditional specification (FCS), otherwise known as chained equations (Audigier et al., 2018; Burgette and Reiter, 2010; Grund et al., 2018a; Van Buuren, 2007). FCS employs univariate linear mixed models to account for the hierarchical structure of multilevel models (Enders et al., 2018a; Mistler and Enders, 2017a; Resche-Rigon and White,

2018) and iteratively imputes each incomplete variable conditional on observed and previously imputed variables (Enders et al., 2018a,b, 2016; Grund et al., 2018a; Hughes et al., 2014; Mistler and Enders, 2017a; van Buuren, 2018). Furthermore, it can impute non-linearities, such as cross-level interactions, by using ‘passive imputation’ or defining a separate imputation model for the non-linearities (Grund et al., 2018b; van Buuren, 2018). However, including these non-linearities in FCS is still very complicated (Grund et al., 2018b, 2021; van Buuren, 2018). FCS can also handle random intercepts and slopes, yet, once again, correctly specifying an imputation model accounting for these random effects can be challenging (Grund et al., 2018b, 2021; van Buuren, 2018).

Non-parametric, tree-based models might alleviate these complexities when defining imputation models. They do not assume a specific data distribution. So, they implicitly model non-linear relationships and can simultaneously handle continuous and categorical variables (Breiman et al., 1984; Burgette and Reiter, 2010; Chipman et al., 2010; Hill et al., 2020; James et al., 2021; Lin and Luo, 2019; Salditt et al., 2023). Studies showed that the use of tree-based, non-parametric models like regression trees, random forests, or Bayesian Additive Regression Trees (BART) in imputation of single-level data simplified the imputation process (Burgette and Reiter, 2010; Silva and Gutman, 2022; Waljee et al., 2013; Xu et al., 2016). They showed better model parameter estimates than parametric methods. Specifically, the imputations showed better confidence interval coverage of the parameters, lower variance and lower bias, especially in non-linear and interactive contexts (Burgette and Reiter, 2010; Silva and Gutman, 2022; Xu et al., 2016). Waljee et al. (2013) also found lower missclassification error rate for the predicted class as well as lower imputation error when imputing with a random forest algorithm compared to multivariate imputation by chained equations (*mice*) using linear, logistic, and polytomous logistic regression imputation models, K-nearest neighbors (KNN) and mean imputation.

In prediction, multilevel-BART models (M-BART) have predominantly been implemented with random intercepts only (Chen, 2020; Tan et al., 2016; Wagner et al., 2020; Wundervald et al., 2022). Wagner et al. (2020) have found that this random intercept M-BART model provided better predictions with a lower mean squared error (MSE) compared to a parametric MLM, Tan et al. (2016) found higher area under the curve (AUC) values compared to a single-level BART model and linear logistic random intercept model, and Chen (2020) found better predictions and better coverage of the parameter estimates compared to parametric models and a single-level BART model. Other researchers modeled the random intercept as an extra split on each terminal node and found a lower MSE compared to a standard BART model and parametric MLMs (Wundervald et al., 2022). Dorie et al. (2022) developed a multilevel BART model that included random intercepts, random slopes and cross-level interactions by modeling these random parts with a Stan (Lee et al., 2017) model and the fixed parts with a BART model. Their results showed that their algorithm `stan4bart` showed better coverage of the population values and lower root mean squared error (RMSE) compared to BART models with varying intercept, BART models ignoring the multilevel structure, bayesian causal forests, and parametric MLMs.

Despite these promising findings, M-BART models have yet to be implemented in a multilevel multiple imputation context. Thus, my research question will be: *Can multivariate imputation by chained equations through a multilevel bayesian additive regression trees model improve the bias, variance, and coverage of the multilevel model parameter estimates compared to current practices?* Given the success of non-parametric models in single-level MI, I anticipate that employing M-BART models in a multilevel missing data context will reduce bias, accurately model variance, and improve estimate coverage compared to conventional implementations of multilevel MI, single-level MI, and complete case analysis in the R-package *mice* (Buuren and Groothuis-Oudshoorn, 2011).

## 2 Method

### 2.1 Theoretical background

#### 2.1.1 Bayesian Additive Regression Trees (BART)

BART is a sum-of-trees model proposed by Chipman et al. (2010) with regression trees as its building blocks (Chipman et al., 2010; Hill et al., 2020; James et al., 2021). Regression trees recursively split the data into binary subgroups based on the predictors included in the model. At each step down the tree, these splits are based on the predictor that minimizes the variability within the subgroups from all predictors. Observations are then assigned to a certain subgroup according to these splits. This is continued until a certain stopping criterion is reached; for example, we desire a minimal number of observations within a subgroup (Breiman et al., 1984; Hastie, 2017; James et al., 2021; Salditt et al.,

2023). Recursive binary partitioning of the predictor space doesn't assume a specific data form. This making regression trees, and as a consequence, BART, non-parametric models (Breiman et al., 1984; Hastie, 2017; James et al., 2021; Salditt et al., 2023) and allows regression trees to model non-linearities and other complicated relationships well and automatically (Burgette and Reiter, 2010; Hill et al., 2020). Chipman et al. (2010) define the BART model as:

$$f(\mathbf{x}) = \sum_{k=1}^K g(\mathbf{x}; T_k, M_k), \quad (1)$$

where  $f(\mathbf{x})$  is the overall fit of the model: the sum of  $K$  regression trees,  $\mathbf{x}$  are the predictor variables,  $T_k$  is the  $k^{\text{th}}$  tree and  $M_k$  is the collection of leaf parameters within the  $k^{\text{th}}$  tree, i.e. the collection of predictions for its terminal nodes (Chipman et al., 2006, 1998, 2010; Hill et al., 2020; James et al., 2021). The data are assumed to arise from a model with additive normally distributed errors:  $Y = \sum_{k=1}^K g(\mathbf{x}; T_k, M_k) + \epsilon$ ,  $\epsilon \sim \mathcal{N}(0, \sigma^2)$ . Next to the sum-of-trees model, BART also includes a regularization prior that constrains the size and fit of each tree so that each contributes only a small part of the variation in the outcome variables to prevent overfitting. The prior is imposed over all parameters of the sum-of-trees model, specifically,  $(T_1, M_1), \dots, (T_K, M_K)$  and  $\sigma$ . However, the specification of the regularization prior is simplified by a series of independence assumptions:

$$\begin{aligned} p((T_1, M_1), \dots, (T_K, M_K), \sigma) &= \left[ \prod_k p(T_k, M_k) \right] p(\sigma), \\ &= \left[ \prod_k p(M_k | T_k) p(T_k) \right] p(\sigma), \\ p(M_k | T_k) &= \prod_j p(\mu_{jk} | T_k), \end{aligned} \quad (2)$$

where  $\mu_{jk} \in M_k$ . These assumptions state that the trees  $(T_k)$ , leaf parameters  $(\mu_j | T_k)$ , and the standard deviation  $(\sigma)$  are independent of each other. Thus, priors only need to be specified for those parameters (Chipman et al., 2006, 1998, 2010; Hill et al., 2020). Chipman et al. (1998) define an independent prior for each tree. The probability that a node at depth  $d$  splits is defined as:

$$\alpha(1 + d)^{-\beta}, \alpha \in (0, 1), \beta \in [0, \infty), \quad (3)$$

where the default specification put forth by Chipman et al. (2006, 2010) is  $\alpha = .95$  and  $\beta = 2$ . This specification sets the probability of a tree with 1, 2, 3, 4, and 5 nodes at .05, .55, .28, .09, and .03 respectively. Thus, smaller trees are favoured. Chipman et al. (2006, 2010) also provide a default specification for the prior for the leaf parameters. They propose to rescale the response value to the interval  $[-.5, .5]$ . Then, the leaf parameter prior is defined as:

$$\mu_{jk} \sim \mathcal{N}(0, \sigma_\mu^2), \text{ with } \sigma_\mu^2 = \frac{.5}{t\sqrt{K}}, \quad (4)$$

where  $t$  is a preselected number and  $K$  is the number of trees. This prior shrinks the tree parameters  $\mu_{jk}$  towards 0, decreasing the effect of the individual tree components. If  $t$  or  $K$  increase, more shrinkage is applied. Chipman et al. (2006, 2010) found good results with and recommend using  $t = 2$  – or values between 1 and 3 – as a default choice. Furthermore, Chipman et al. (2006, 2010) propose the conjugate inverse chi-square distribution as the prior for the residual standard deviation  $\sigma^2 \sim \nu\lambda/\chi_\nu^2$ . They represent the degrees of freedom,  $\lambda$ , as the probability that the BART residual standard deviation,  $\sigma$ , is less than the estimated residual standard deviation from a linear regression model,  $\hat{\sigma}_{\text{OLS}}$ . Their default specification of the hyperparameters is  $\nu = 3$  and  $\Pr(\sigma < \hat{\sigma}_{\text{OLS}}) = .9$  (Chipman et al., 2006, 1998, 2010; Hill et al., 2020).

BARTs are estimated using the Bayesian back-fitting Markov Chain Monte Carlo (MCMC) algorithm (Chipman et al., 2006, 1998, 2010; Hill et al., 2020; James et al., 2021). Each tree is initialized with a single root node with the mean response value divided by the number of trees ( $\hat{f}_k^1(x) = \frac{1}{nK} \sum_{i=1}^n y_i$ , with sample size  $n$ ). Then, each pair  $(T_k, M_k)$  is updated considering the remaining trees, their associated parameters, and the residual standard deviation  $(\sigma)$  by sampling from the following conditional distribution:

$$(T_k, M_k) | T_{k'}, M_{k'}, \sigma, y. \quad (5)$$

However, this conditional distribution only depends on  $(T_{k'}, M_{k'}, y)$  through the partial residuals:

$$r_i = y_i - \sum_{k' < k} \hat{f}_{k'}^b(x_i) - \sum_{k' > k} \hat{f}_{k'}^{b-1}(x_i), \text{ with } i = 1, \dots, n, \quad (6)$$

where  $\hat{f}_k^b(x_i)$  is the prediction of the  $k^{\text{th}}$  tree in the  $b^{\text{th}}$  iteration for person  $i$  and sample size  $n$ . Thus, updating each pair  $(T_k, M_k)$  simplifies to proposing a new tree fit to the partial residuals,  $r_i$ , treating them as the data, by perturbing the tree from the previous iteration. Perturbations entail either *growing*, *pruning*, or *changing* a tree. *Growing* means adding additional splits, *pruning* removes splits, and *changing* changes decision rules. The algorithm stops after the specified number of iterations (Chipman et al., 2006, 1998, 2010; Hill et al., 2020; James et al., 2021).

### 2.1.2 Multilevel-BART (M-BART)

Tan et al. (2016); Wagner et al. (2020) and Dorie et al. (2024) define an M-BART model including a random intercept. The BART model (1) is extended to include a random intercept by:

$$m(\mathbf{x}) = \sum_{k=1}^K g(\mathbf{x}; T_k, M_k) + \alpha_j, \quad (7)$$

where, now,  $f(\mathbf{x})$  is the overall fit of the model incorporating random intercept  $\alpha_j$  for cluster  $j$  and. So, the data are now assumed to arise from the following model:

$$Y_{ij} = \sum_{k=1}^K g(\mathbf{x}; T_k, M_k) + \alpha_j + \epsilon_{ij}, \quad \epsilon_{ij} \sim \mathcal{N}(0, \sigma^2), \quad \alpha_j \sim \mathcal{N}(0, \tau^2), \quad (8)$$

where  $\alpha_j \perp \epsilon_{ij}$ . Now the joint prior distribution (2) becomes:

$$\begin{aligned} p((T_1, M_1), \dots, (T_K, M_K), \sigma) &= \left[ \prod_k p(T_k, M_k) \right] p(\sigma) p(\tau), \\ &= \left[ \prod_k p(M_k | T_k) p(T_k) \right] p(\sigma) p(\tau), \\ p(M_k | T_k) &= \prod_j p(\mu_{jk} | T_k). \end{aligned} \quad (9)$$

A Metropolis within Gibbs procedure is used to draw values from the posterior. First, the Gibbs sample for  $\sigma$ ,  $\tau$ , and  $\alpha_j$  are obtained from their respective posterior distributions. Then, we obtain  $\tilde{Y}_{ij} = Y_{ij} - \alpha_j$  and view  $\tilde{Y}_{ij} | \mathbf{X}_j$  as a BART model. So,  $\tilde{Y}$  is now used as the outcome variable in the BART algorithm described in the previous section, 2.1.1. (Tan et al., 2016; Wagner et al., 2020). Dorie et al. (2024) implemented this algorithm within the R-package `dbarts` with the function `rbart.vi()`. Where, the default prior for the random intercept is  $\tau \sim \text{Cauchy}(0, 2.5)$ : a Cauchy distribution with a scale parameter 2.5 times the original scale.

### 2.1.3 stan4bart

Dorie et al. (2022) developed a multilevel BART model that included random intercepts, random slopes, and cross-level interactions. They extend a Bayesian linear, mixed model with a BART model (1). The resulting model is:

$$h(\mathbf{x}) = \mathbf{x}^\beta \boldsymbol{\beta} + f(\mathbf{x}; T_K, M_K) + \boldsymbol{\lambda} \mathbf{w}, \quad (10)$$

where  $\mathbf{x}^\beta$  is a vector of 1 – for the intercept – and the linear predictors;  $\boldsymbol{\beta}$  is a vector of linear, parametric coefficients;  $\boldsymbol{\lambda}$  is a vector of all parametric random slopes and intercepts;  $\mathbf{w}$  is a vector of the coefficients for the random slopes and intercepts; and  $f(\mathbf{x}; T_K, M_K)$  is a non-parametric, sum-of-trees BART model (Dorie et al., 2022). So, the data are assumed to arise from the following model:

$$Y_{ij} = \mathbf{x}^\beta \boldsymbol{\beta} + f(\mathbf{x}; T_K, M_K) + \boldsymbol{\lambda} \mathbf{w} + \epsilon_{ij}, \quad \epsilon_{ij} \sim \mathcal{N}(0, \sigma^2), \quad \boldsymbol{\lambda} \sim \mathcal{N}(0, \boldsymbol{\Sigma}_\lambda), \quad (11)$$

where  $\Sigma_\lambda$  is the variance-covariance matrix for the random intercept and slopes. The model is implemented as a Gibbs sampler: a Hamiltonian Monte Carlo, no-U-turn sampler with a diagonal Euclidean adaptation matrix is used to jointly sample the linear, parametric components given the non-parametric components. The non-parametric components are sampled using the BART algorithm described in section 2.1.1 (Dorie et al., 2022). To accomplish this, a parametric Stan model (Lee et al., 2017) fits equation 10 with  $f(\mathbf{x}; T_K, M_K)$  as a generic linear offset. Dorie et al. (2022) combine a custom mutable Stan sampler object with a BART sampler with a fixed variance and offset term: First, the Stan sampler collects the current draws of the BART model into  $\text{vec}_i f(\mathbf{x}_i; T_K, M_K)$  and uses this to draw  $\beta, \lambda, \sigma, \Sigma_\lambda | \mathbf{Y}, \text{vec}_i f(\mathbf{x}_i; T_K, M_K)$ . Then,  $\sigma$  and  $\text{vec}_i [\mathbf{x}_i^\beta \beta + \lambda \mathbf{w}_i]$  are passed to BART, which produces  $M_k, T_k | \mathbf{Y}, \text{vec}_i [\mathbf{x}_i^\beta \beta + \lambda \mathbf{w}_i], \sigma, M_{k'}, T_{k'}$ . Then, the cycle is completed by passing  $\text{vec}_i f(\mathbf{x}_i; T_K, M_K)$  back to Stan. The process is continued for the set amount of posterior samples intended for inference. This algorithm is implemented in the R-package `stan4bart` (Dorie, 2023).

## 2.2 Simulation study

### 2.2.1 Data generating mechanism

We assembled a simulation study to evaluate the performance of multilevel BART models in a multilevel imputation context. The population data-generating mechanism is based on the following MLM:

$$y_{ij} = \beta_{0j} + \sum_{k=1}^7 \beta_{kj} X_{kij} + \epsilon_{ij}, \quad X_{kij} \sim \mathcal{MVN}(0, \Sigma_x), \quad (12a)$$

$$\beta_{0j} = \gamma_{00} + \sum_{p=1}^2 \gamma_{0p} Z_{pj} + v_{0j}, \quad (12b)$$

$$\beta_{kj} = \gamma_{k0} + \sum_{p=1}^2 \gamma_{kp} Z_{pj} + v_{kj}, \quad Z_{pj} \sim \mathcal{MVN}(0, \Sigma_z), \quad (12c)$$

where  $y_{ij}$  is a continuous level-1 outcome variable for person  $i$  in group  $j$  and  $X_{kij}$  are 7 continuous level-1 variables and  $Z_{pj}$  are 2 continuous level-2 variables. The predictors are multivariate normally distributed with means of 0 and variance-covariance matrix  $\Sigma_x$  and  $\Sigma_z$ , respectively:

$$\Sigma_x = \begin{pmatrix} 6.25 & & & & & & \\ 2.25 & 9 & & & & & \\ 1.5 & 1.8 & 4 & & & & \\ 2.25 & 3.06 & 2.04 & 11.56 & & & \\ 1.5 & 1.8 & 1.2 & 2.04 & 4 & & \\ 1.125 & 1.35 & 0.9 & 1.53 & .9 & 2.25 & \\ 3.3 & 3.96 & 2.64 & 4.488 & 2.64 & 1.98 & 19.36 \end{pmatrix}, \quad (13a)$$

$$\Sigma_z = \begin{pmatrix} 1 & \\ .48 & 2.56 \end{pmatrix}. \quad (13b)$$

The covariances between the variables are calculated such that the correlation between the variables is .3, aligned with Cohen's (1990) medium effect size benchmark. The residuals are normally distributed as,

$$\epsilon_{ij} \sim \mathcal{N}(0, 25). \quad (14)$$

The random intercept  $\beta_{0j}$  is determined by the overall intercept  $\gamma_{00}$ , the 2 group-level effects  $\gamma_{0p} Z_{pj}$  and the group-level random residuals  $v_{0j}$ . The overall intercept  $\gamma_{00}$  is set to 10 and the group-level effects  $\gamma_{01}$  and  $\gamma_{02}$  to .5. The 7 regression coefficients  $\beta_{kj}$  for the continuous variables  $X_{kij}$  depend on the intercepts  $\gamma_{k0}$ , the cross-level interactions  $\gamma_{kp} Z_{pj}$ , and the random slopes  $v_{kj}$ . The 7 intercepts, or within-group

effect sizes,  $\gamma_{k0}$  are set to .5, the cross-level interactions  $\gamma_{11}$ ,  $\gamma_{21}$ , and  $\gamma_{32}$  are set to .35.

$$\gamma_{00} = 10, \quad \gamma_{0p} = \begin{pmatrix} .5 \\ .5 \\ .5 \\ .5 \\ .5 \\ .5 \end{pmatrix}, \quad \gamma_{k0} = \begin{pmatrix} .5 \\ .5 \\ .5 \\ .5 \\ .5 \\ .5 \end{pmatrix}, \quad \gamma_{kp} = \begin{pmatrix} .35 & 0 \\ .35 & 0 \\ 0 & .35 \\ 0 & 0 \\ 0 & 0 \\ 0 & 0 \end{pmatrix}. \quad (15)$$

The random slopes are multivariate normally distributed with a mean of 0 and a variance-covariance matrix  $\mathbf{T}$  shown in equation 16. Again, the covariances are calculated to yield a correlation of .3.

$$v_j \sim \mathcal{MVN}(0, \mathbf{T}), \quad \mathbf{T} = \begin{pmatrix} t_{00} & & & & & & & \\ .3 & 1 & & & & & & \\ .3 & .3 & 1 & & & & & \\ .3 & .3 & .3 & 1 & & & & \\ 0 & 0 & 0 & 0 & 0 & & & \\ 0 & 0 & 0 & 0 & 0 & 0 & & \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & \\ 0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \end{pmatrix}. \quad (16)$$

The variance of  $v_{0j}$ , the group-level random residuals  $t_{00}$ , are scaled such that the specified ICC values as in table 1 was obtained. The following formula is used to calculate  $v_{0j}$  following the variance decomposition from Rights and Sterba (2019):

$$\text{ICC} = \frac{\gamma^{b'} \phi^b \gamma^b + \tau_{00}}{\gamma^{w'} \phi^w \gamma^w + \gamma^{b'} \phi^b \gamma^b + \text{tr}(\mathbf{T}\mathbf{\Sigma}) + \tau_{00} + \sigma^2}, \quad (17)$$

where  $\gamma^b$  and  $\gamma^w$  are the level-1 and level-2 fixed effects;  $\phi^b$  is the variance-covariance matrix of a vector with 1, for the intercept, and all level-2 predictors;  $\phi^w$  is the variance-covariance matrices of all cluster-mean-centered level-1 predictors;  $\tau_{00}$  is the variance of the random intercept;  $\mathbf{T}$  is the variance-covariance matrix of the random intercept and slopes;  $\mathbf{\Sigma}$  is the variance-covariance matrix of a vector containing 1, for the intercept, and the level-1 variables; and  $\sigma^2$  is the residual variance. The value for  $\tau_{00}$  is calculated using the function `uniroot()` in R (R Core Team, 2023).

## 2.2.2 Simulation design

Table 1 shows the design factors considered in the simulation study. These factors are either grounded in prior research or deemed realistic in real-world applications (Enders et al., 2020, 2018b; Grund et al., 2018b; Gulliford et al., 1999; Hox et al., 2017; Murray and Blitstein, 2003). According to Kreft and de Leeuw (2007), 30 groups is the smallest acceptable number in multilevel research and 50 groups is frequent in organizational research (Maas and Hox, 2005). Group sizes of 15 are typical in educational research (Lüdtke et al., 2017) and group sizes of 50 are often used in simulation studies (Akkaya Hocagil and Yucel, 2023; Enders et al., 2020, 2018a,b; Grund et al., 2018b; Maas and Hox, 2005). The ICC was chosen to be .5, which is often used as an upper limit in methodological research (Enders et al., 2020, 2018a,b; Grund et al., 2018b; Mistler and Enders, 2017b; Salditt et al., 2023). Oberman and Vink (2023) recommend including both Missing Completely At Random (MCAR) and Missing At Random (MAR) missingness mechanisms in simulation studies. They pose that the statistical properties of the imputation method are not deemed sound if it cannot yield valid inferences under MCAR. Furthermore, they pose that including observed-data-dependend missingness – for example, MAR – is of utmost importance in evaluating the imputation method’s performance. The amount of missingness in data sets is varied between 0% and 50%. 0% missingness is included as an additional benchmark and 50% missingness is often used in simulation studies as a high amount of missingness (Grund et al., 2016; Lüdtke et al., 2017; Schouten and Vink, 2021). For each combination of design factors, 100 datasets are simulated. 5 different imputation methods are compared:

1. conventional single-level imputation with PMM (predictive mean matching),
2. conventional multilevel imputation with PMM,



3. single-level BART imputation,
4. multilevel BART imputation accounting for random intercepts (Chen, 2020; Tan et al., 2016; Wagner et al., 2020),
5. multilevel BART imputation accounting for random effects and cross-level interactions (Dorie et al., 2022).

The first and second methods are implemented with the R-packages `mice` and `miceadds` (Robitzsch et al., 2024). The conventional single-level imputation is implemented with the imputation method `pmm` and the conventional multilevel imputation is implemented with the `2l.pmm` method for level-1 variables and `2lonly.mean` for level-2 variables.

The third, single-level BART, fourth, random intercept BART and fifth method, multilevel BART methods are implemented by writing new method-functions in R (R Core Team, 2023) for the package `mice`. The functions `bart` and `rbart_vi` from the `dbarts` package were used for the single-level and random intercept BART imputation methods (Dorie et al., 2024). The function `stan4bart` from the package `stan4bart` was used for the multilevel BART imputation method accounting for random effects and cross-level interactions (Dorie, 2023). The functions were written such that they can be used as imputation methods in the `mice` package. All three functions are implemented as follows: for every variable to be imputed, a respective BART model is fitted based on the predictor matrix. Then, the fitted values – the posterior means – are extracted for the observed and missing values. Imputations for the missing values are then obtain using predictive mean matching by matching the predicted values for the observed cases to the predicted values for the missing cases. The code for these functions can be found in the appendix – listing 1, 2, and 3.

For all imputation methods, the incomplete data sets are imputed 5 times with 10 iterations each. Then, each of the 5 imputed datasets are then analyzed using the R-package `lme4` (Bates et al., 2015) with an MLM reflecting the population generating mechanism:  $y = 1 + x_1 + x_2 + x_3 + x_4 + x_5 + x_6 + x_7 + z_1 + z_2 + x_1 * z_1 + x_2 * z_1 + x_3 * z_2 + (1 + x_1 + x_2 + x_3 \mid \text{group})$ . The estimates from the 5 imputed datasets are pooled together using the R-package `mice` (Buuren and Groothuis-Oudshoorn, 2011). These pooled estimates are compared on the bias, coverage, and the width of the 95% confidence intervals.

As an additional benchmark, the imputation methods will also be compared to analyses using listwise deletion, i.e. complete case analysis, and using the true data without missing values.

### 2.2.3 Missing data generation

Missing values in the variables are introduced by multivariate amputation using the function `ampute()` (Schouten et al., 2018) from package `mice`. As can be seen in table 1, the missing data mechanism is either Missing Completely At Random (MCAR) or Missing At Random (MAR). The missing data mechanism is said to be MCAR when the cause of the missing data is unrelated to the data and MAR when the missing data is related to the observed data (Rubin, 1976). The amount of missingness is either 0% or 50%, which is defined as the percentage of cases that have at least one missing value.

For both MCAR and MAR, all possible patterns with 1 to 5 missing values out of the 10 variables ( $x_1, x_2, x_3, x_4, x_5, x_6, x_7, z_1, z_2$ , and  $y$ ) per case are generated. They have the same relative frequency of occurrence in the data sets. So, 50% of the cases had 1 to 5 missing values.

For the MAR mechanism, the weighted sum of scores on the observed variables is used to predict the probability of missingness for a case. The weights of the variables  $x_4$  and  $z_1$  are set to 2 and 1.5 respectively when they remain observed in a specific pattern, while the weights of the other variables that remain observed in a specific pattern are set to 1. The type of missingness is set to ‘RIGHT’ meaning that cases with a higher weighted sum of scores have a higher probability of becoming incomplete. So, this means that cases with higher values on  $x_4$  and  $z_1$  are more likely to become incomplete.

In summary, either no missing values are introduced (0%), or up to 5 missing values are introduced in 50% of the cases. When data is MAR, the probability of a value being missing depends on the observed values of all other variables, with variables  $x_4$  and  $z_1$  having a greater influence on this probability.

**Table 1:** Simulation design

Parameter	Values
Number of clusters (J)	30, 50
Within-cluster sample size ( $n_j$ )	15, 50
Intraclass Correlation (ICC)	.5
Missing data mechanism	MCAR, MAR
Amount of missingness	0%, 50%

### 2.2.4 Evaluation

The estimates from the analysis models are evaluated in terms of absolute bias, coverage of 95% confidence intervals, with their respective Monte Carlo SE (MCSE), and the width of the 95% confidence intervals (Morris et al., 2019; Oberman and Vink, 2023):

$$\text{Bias} = \frac{1}{n_{\text{sim}}} \sum_{t=1}^{n_{\text{sim}}} (\hat{\theta}_t - \theta), \quad \text{MCSE}_{\text{Bias}} = \sqrt{\frac{\sum_{t=1}^{n_{\text{sim}}} (\hat{\theta}_t - \bar{\theta})^2}{n_{\text{sim}}(n_{\text{sim}} - 1)}}, \quad (18)$$

$$\text{Coverage} = \frac{1}{n_{\text{sim}}} \sum_{t=1}^{n_{\text{sim}}} 1(\hat{\theta}_{\text{low},i} \leq \theta \leq \hat{\theta}_{\text{upp},i}), \quad \text{MCSE}_{\text{Coverage}} = \sqrt{\frac{\text{Coverage}(1 - \text{Coverage})}{n_{\text{sim}}}}, \quad (19)$$

$$\text{CIW} = \frac{1}{n_{\text{sim}}} \sum_{t=1}^{n_{\text{sim}}} (\hat{\theta}_{\text{upp},i} - \hat{\theta}_{\text{low},i}), \quad (20)$$

where  $\hat{\theta}_t$  is the estimated parameter in simulation  $t$ ,  $\theta$  is the true value,  $\bar{\theta}$  is the mean of  $\hat{\theta}_t$ , and  $n_{\text{sim}}$  is the number of simulated datasets. The lower and upper bounds of the 95% confidence intervals are denoted as  $\hat{\theta}_{\text{low},i}$  and  $\hat{\theta}_{\text{upp},i}$  respectively. The coverage is the proportion of the 95% confidence intervals that contain the true value.

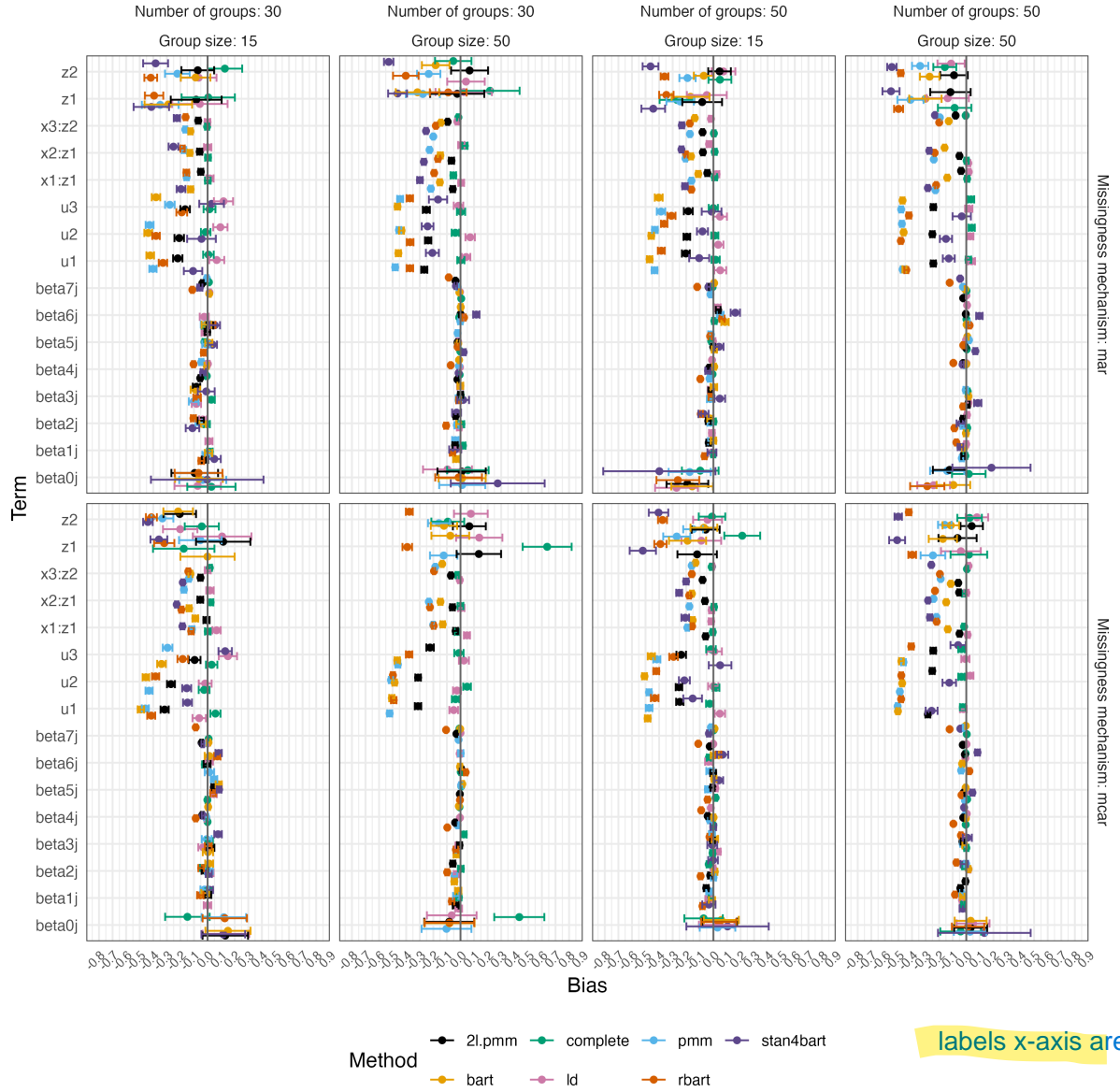
Enders et al. (2018a); Morris et al. (2019); Oberman and Vink (2023); van Buuren (2018) suggest that a coverage of 95% is acceptable. Poor coverage, i.e. below 95%, indicates biased estimates or too narrow intervals. While, coverage above 95% indicates that efficiency could still be gained. The width of the confidence intervals is a measure of the statistical precision of the estimates: a smaller width indicates a more precise estimate (Oberman and Vink, 2023; van Buuren, 2018).

## 3 Results

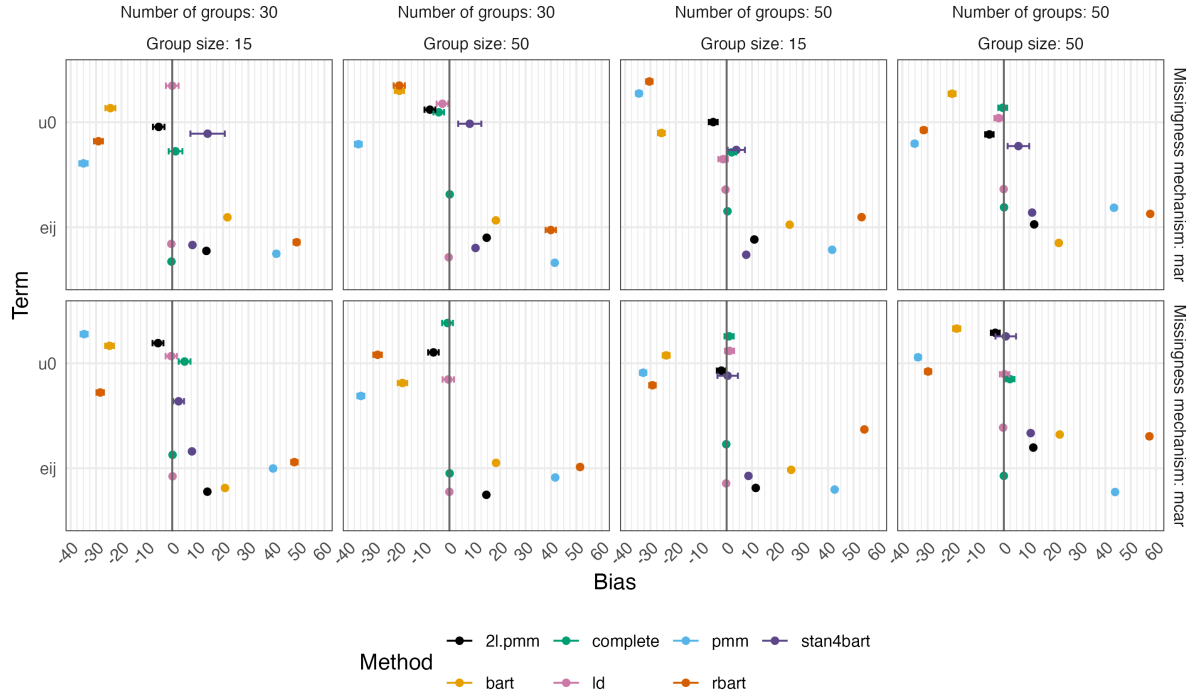
### 3.1 Bias

Figure 1 shows the absolute bias of the estimates of the linear mixed model – except the residual variance,  $\epsilon_{ij}$ , and the intercept variance,  $u_0$  for interpretability – for all imputation methods in consideration. The absolute bias of the residual variance and the intercept variance are shown in figure 2.

First, the estimates of the fixed effects – the overall intercept,  $\beta_{0j}$ ; level-1 effects  $\beta_{1j} : \beta_{7j}$ ; level-2 effects  $z_1$  and  $z_2$ ; and the cross-level interactions  $x_3 * z_2, x_2 * z_1, x_1 * z_1$  – will be considered in terms of absolute bias.

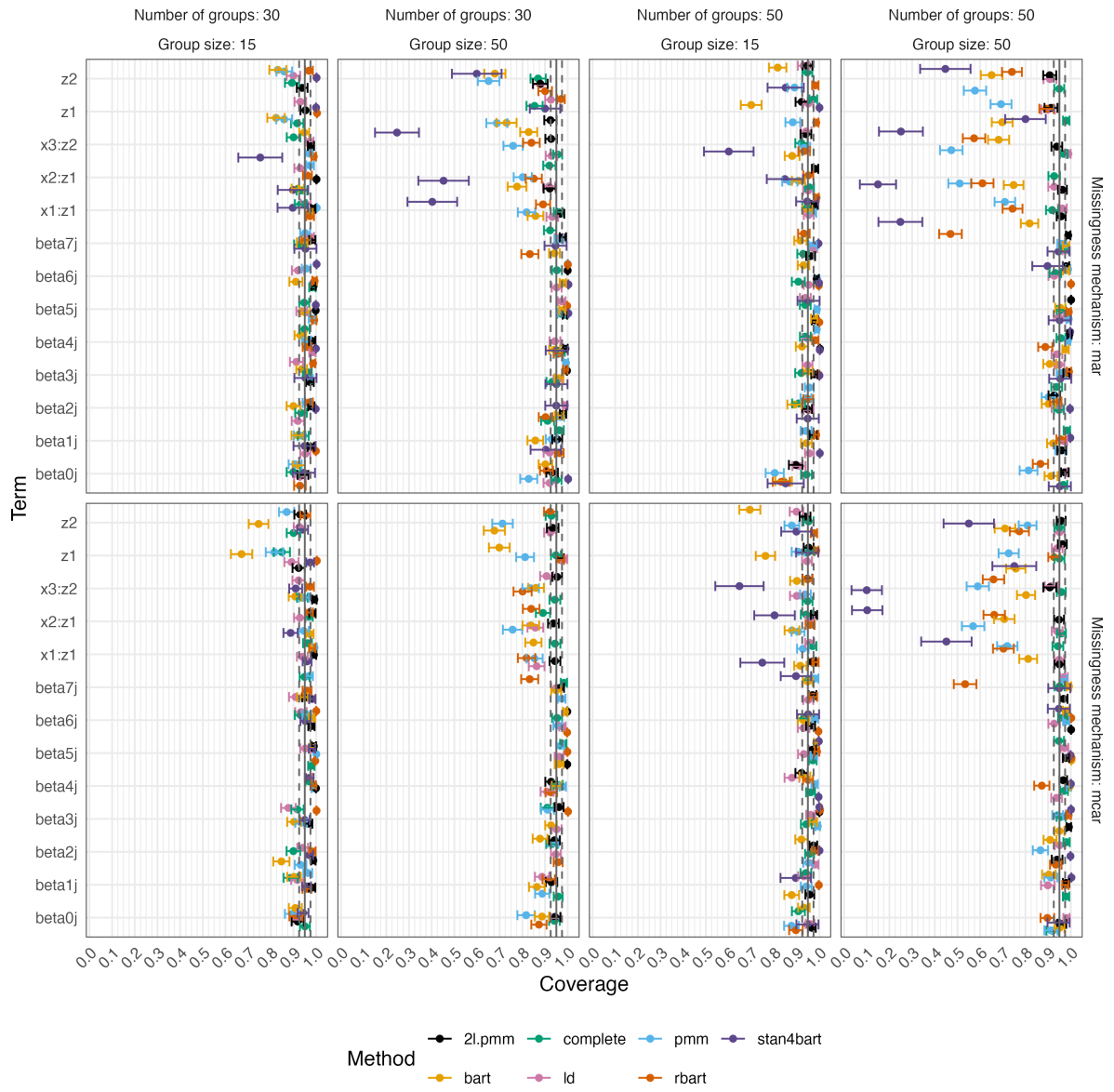


**Figure 1:** Absolute bias of the estimates of the linear mixed model for all simulated data sets over 100 simulations with  $ICC = .5$  excluding the residual variance ( $\epsilon_{ij}$ ) and intercept variance ( $u_0$ ).



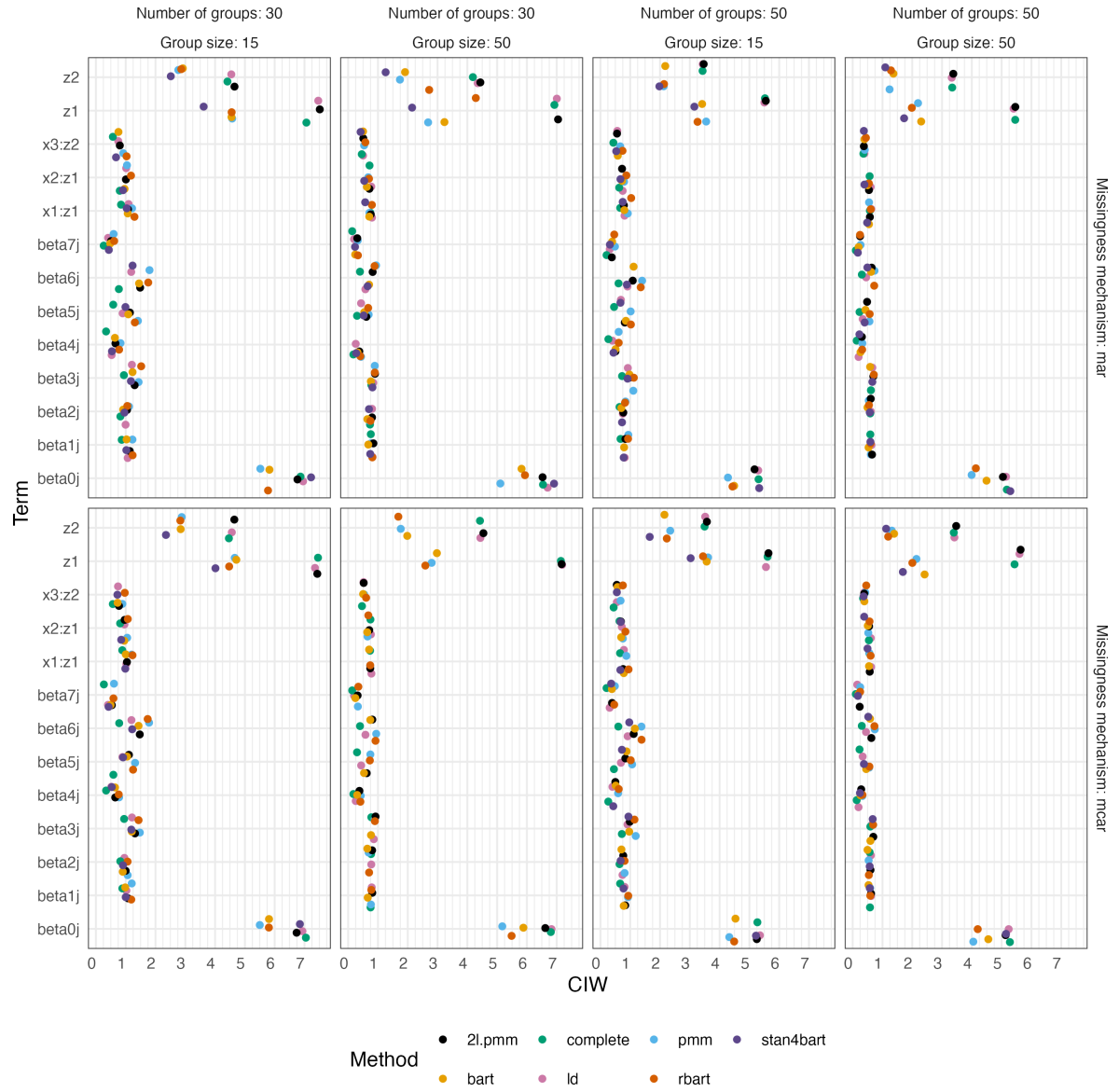
**Figure 2:** Bias of the  $e_{ij}$  and  $u_0$

### 3.2 Coverage



**Figure 3:** Coverage of the estimates

### 3.3 Confidence interval width



**Figure 4:** Confidence interval width of the estimates

## 4 Discussion

## 5 Conclusion

## 6 Appendix

**Listing 1:** Imputation function for single-level BART

```
1  mice.impute.bart <- function(y, ry, x, wy = NULL, use.matcher = FALSE, donors = 5L,  
... ) {  
2    install.on.demand("dbarts", ...)  
3    if (is.null(wy)) {  
4      wy <- !ry  
5    }  
6  
7    # Parameter estimates  
8    fit <- dbarts::bart(x, y, keeptrees = TRUE, verbose = FALSE)  
9  
10   yhatobs <- fitted(fit, type = "ev", sample = "train")[ry]  
11   yhatmis <- fitted(fit, type = "ev", sample = "train")[wy]  
12  
13   # Find donors  
14   if (use.matcher) {  
15     idx <- matcher(yhatobs, yhatmis, k = donors)  
16   } else {  
17     idx <- matchindex(yhatobs, yhatmis, donors)  
18   }  
19  
20   return(y[ry][idx])  
21 }
```

**Listing 2:** Imputation function for random intercept BART

```
1  mice.impute.2l.rbart <- function(y, ry, x, wy = NULL, type, use.matcher = FALSE,  
donors = 5L, ...) {  
2    install.on.demand("dbarts", ...)  
3    if (is.null(wy)) {  
4      wy <- !ry  
5    }  
6  
7    clust <- names(type[type == -2])  
8    effects <- names(type[type != -2])  
9    X <- x[, effects, drop = FALSE]  
10  
11    model <- paste0(  
12      "y ~ ", paste0(colnames(X), collapse = " + ")  
13    )  
14  
15    fit <- dbarts::rbart_vi(formula = formula(model), group.by = clust, data = data.  
frame(y, x), verbose = FALSE, n.threads = 1, n.samples = 500L, n.burn = 500L, ...)  
16  
17    yhatobs <- fitted(fit, type = "ev", sample = "train")[ry]  
18    yhatmis <- fitted(fit, type = "ev", sample = "train")[wy]  
19  
20    # Find donors  
21    if (use.matcher) {  
22      idx <- matcher(yhatobs, yhatmis, k = donors)  
23    } else {  
24      idx <- matchindex(yhatobs, yhatmis, donors)  
25    }  
26  
27    return(y[ry][idx])  
28 }
```

**Listing 3:** Imputation function for multilevel BART with random effects and cross-level interactions

```
1  mice.impute.2l.bart <- function(y, ry, x, wy = NULL, type, intercept = TRUE, use.  
matcher = FALSE, donors = 5L, ...) {  
2    install.on.demand("stan4bart", ...)  
3    if (is.null(wy)) {  
4      wy <- !ry  
5    }  
6  
7    if (intercept) {  
8      x <- cbind(1, as.matrix(x))  
9      type <- c(2, type)
```

```

10     names(type)[1] <- colnames(x)[1] <- "(Intercept)"
11 }
12
13 clust <- names(type[type == -2])
14 rande <- names(type[type == 2])
15 fixe <- names(type[type > 0])
16
17 lev <- unique(x[, clust])
18
19 X <- x[, fixe, drop = FALSE]
20 Z <- x[, rande, drop = FALSE]
21 xobs <- x[ry, , drop = FALSE]
22 yobs <- y[ry]
23 Xobs <- X[ry, , drop = FALSE]
24 Zobs <- Z[ry, , drop = FALSE]
25
26 # create formula
27 fr <- ifelse(length(rande) > 1,
28             paste0("+ (1 +", paste(rande[-1L], collapse = "+")),
29             " + (1 "
30             )
31 randmodel <- paste0(
32     "y ~ bart(", paste0(fixe[-1L], collapse = " + "), ")",
33     fr, "| ", clust, ")"
34 )
35 fit <- eval(parse(text = paste("stan4bart::stan4bart(", randmodel,
36     ", data = data.frame(y, x),
37     verbose = -1,
38     bart_args = list(k = 2.0, n.samples = 500L, n.burn = 500L, n.thin = 1L, n.
39     threads = 1)")),
40     collapse = ""
41 )))
42
43 yhatobs <- fitted(fit, type = "ev", sample = "train")[ry]
44 yhatmis <- fitted(fit, type = "ev", sample = "train")[wy]
45
46 # Find donors
47 if (use.matcher) {
48     idx <- matcher(yhatobs, yhatmis, k = donors)
49 } else {
50     idx <- matchindex(yhatobs, yhatmis, donors)
51 }
52
53 return(y[ry][idx])
54 }

```



## References

- Akkaya Hocagil, T. and Yucel, R. M. (2023). A computationally efficient sequential regression imputation algorithm for multilevel data. *Journal of Applied Statistics*, pages 1–21.
- Audigier, V., White, I. R., Jolani, S., Debray, T. P. A., Quartagno, M., Carpenter, J., Van Buuren, S., and Resche-Rigon, M. (2018). Multiple Imputation for Multilevel Data with Continuous and Binary Variables. *Statistical Science*, 33(2).
- Austin, P. C., White, I. R., Lee, D. S., and Van Buuren, S. (2021). Missing Data in Clinical Research: A Tutorial on Multiple Imputation. *Canadian Journal of Cardiology*, 37(9):1322–1331.
- Bartlett, J. W., Seaman, S. R., White, I. R., Carpenter, J. R., and for the Alzheimer’s Disease Neuroimaging Initiative\* (2015). Multiple imputation of covariates by fully conditional specification: Accommodating the substantive model. *Statistical Methods in Medical Research*, 24(4):462–487.
- Bates, D., Mächler, M., Bolker, B., and Walker, S. (2015). Fitting Linear Mixed-Effects Models Using **lme4**. *Journal of Statistical Software*, 67(1).
- Breiman, L., Friedman, J. H., Olshen, R. A., and Stone, C. J. (1984). *Classification And Regression Trees*. Routledge, 1 edition.
- Burgette, L. F. and Reiter, J. P. (2010). Multiple Imputation for Missing Data via Sequential Regression Trees. *American Journal of Epidemiology*, 172(9):1070–1076.
- Buuren, S. V. and Groothuis-Oudshoorn, K. (2011). **Mice** : Multivariate Imputation by Chained Equations in R. *Journal of Statistical Software*, 45(3).
- Carpenter, J. R. and Kenward, M. G. (2013). *Multiple Imputation and Its Application*. Wiley, 1 edition.
- Chen, S. (2020). *A New Multilevel Bayesian Nonparametric Algorithm and Its Application in Causal Inference*. PhD thesis, Texas A&M University.
- Chipman, H., George, E., and McCulloch, R. (2006). Bayesian Ensemble Learning. In *Advances in Neural Information Processing Systems*, volume 19. MIT Press.
- Chipman, H. A., George, E. I., and McCulloch, R. E. (1998). Bayesian CART Model Search. *Journal of the American Statistical Association*, 93(443):935–948.
- Chipman, H. A., George, E. I., and McCulloch, R. E. (2010). BART: Bayesian additive regression trees. *The Annals of Applied Statistics*, 4(1).
- Cohen, J. (1990). Statistical power analysis for the behavioral sciences. *Computers, Environment and Urban Systems*, 14(1):71.
- Dong, M. and Mitani, A. (2023). Multiple imputation methods for missing multilevel ordinal outcomes. *BMC Medical Research Methodology*, 23(1):112.
- Dorie, V. (2023). *Stan4bart: Bayesian Additive Regression Trees with Stan-Sampled Parametric Extensions*.
- Dorie, V., Chipman, H., McCulloch, R., Dadgar, A., Team, R. C., Draheim U., G., Bosmans, M., Tournayre, C., Petch, M., Valle, R. d. L., Johnson G., S., Frigo, M., Zaitseff, J., Veldhuizen, T., Maisonobe, L., Pakin, S., and Daniel G., R. (2024). Dbarts: Discrete Bayesian Additive Regression Trees Sampler.
- Dorie, V., Perrett, G., Hill, J. L., and Goodrich, B. (2022). Stan and BART for Causal Inference: Estimating Heterogeneous Treatment Effects Using the Power of Stan and the Flexibility of Machine Learning. *Entropy*, 24(12):1782.
- Enders, C. K. (2017). Multiple imputation as a flexible tool for missing data handling in clinical research. *Behaviour Research and Therapy*, 98:4–18.
- Enders, C. K., Du, H., and Keller, B. T. (2020). A model-based imputation procedure for multilevel regression models with random coefficients, interaction effects, and nonlinear terms. *Psychological Methods*, 25(1):88–112.

- Enders, C. K., Hayes, T., and Du, H. (2018a). A Comparison of Multilevel Imputation Schemes for Random Coefficient Models: Fully Conditional Specification and Joint Model Imputation with Random Covariance Matrices. *Multivariate Behavioral Research*, 53(5):695–713.
- Enders, C. K., Keller, B. T., and Levy, R. (2018b). A fully conditional specification approach to multilevel imputation of categorical and continuous variables. *Psychological Methods*, 23(2):298–317.
- Enders, C. K., Mistler, S. A., and Keller, B. T. (2016). Multilevel multiple imputation: A review and evaluation of joint modeling and chained equations imputation. *Psychological Methods*, 21(2):222–240.
- Grund, S., Lüdtke, O., and Robitzsch, A. (2016). Multiple imputation of missing covariate values in multilevel models with random slopes: A cautionary note. *Behavior Research Methods*, 48(2):640–649.
- Grund, S., Lüdtke, O., and Robitzsch, A. (2018a). Multiple Imputation of Missing Data at Level 2: A Comparison of Fully Conditional and Joint Modeling in Multilevel Designs. *Journal of Educational and Behavioral Statistics*, 43(3):316–353.
- Grund, S., Lüdtke, O., and Robitzsch, A. (2018b). Multiple Imputation of Missing Data for Multilevel Models: Simulations and Recommendations. *Organizational Research Methods*, 21(1):111–149.
- Grund, S., Lüdtke, O., and Robitzsch, A. (2021). Multiple imputation of missing data in multilevel models with the R package mdmb: A flexible sequential modeling approach. *Behavior Research Methods*, 53(6):2631–2649.
- Gulliford, M., Adams, G., Ukoumunne, O., Latinovic, R., Chinn, S., and Campbell, M. (2005). Intraclass correlation coefficient and outcome prevalence are associated in clustered binary data. *Journal of Clinical Epidemiology*, 58(3):246–251.
- Gulliford, M. C., Ukoumunne, O. C., and Chinn, S. (1999). Components of Variance and Intraclass Correlations for the Design of Community-based Surveys and Intervention Studies: Data from the Health Survey for England 1994. *American Journal of Epidemiology*, 149(9):876–883.
- Hastie, T. J., editor (2017). *Statistical Models in S*. Routledge, 1st edition.
- Hill, J., Linero, A., and Murray, J. (2020). Bayesian Additive Regression Trees: A Review and Look Forward. *Annual Review of Statistics and Its Application*, 7(1):251–278.
- Hox, J. and Roberts, J. K., editors (2011). *Handbook of Advanced Multilevel Analysis*. Routledge, 0 edition.
- Hox, J. J., Moerbeek, M., and Van De Schoot, R. (2017). *Multilevel Analysis: Techniques and Applications*. Routledge, Third edition. | New York, NY : Routledge, 2017. |, 3 edition.
- Hughes, R. A., White, I. R., Seaman, S. R., Carpenter, J. R., Tilling, K., and Sterne, J. A. (2014). Joint modelling rationale for chained equations. *BMC Medical Research Methodology*, 14(1):28.
- James, G., Witten, D., Hastie, T., and Tibshirani, R. (2021). *An Introduction to Statistical Learning: With Applications in R*. Springer Texts in Statistics. Springer US, New York, NY.
- Kang, H. (2013). The prevention and handling of the missing data. *Korean Journal of Anesthesiology*, 64(5):402.
- Kreft, I. and de Leeuw, J. (2007). *Introducing Multilevel Modeling*. Introducing Statistical Methods. SAGE, Los Angeles, Calif., reprinted edition.
- Lee, D., Carpenter, B., Li, P., Morris, M., Betancourt, M., Maverickg, Brubaker, M., Trangucci, R., Inacio, M., Kucukelbir, A., Buildbot, S., Bgoodri, Seantalts, Arnold, J., Tran, D., Hoffinan, M., Margossian, C., Modrák, M., Adler, A., Sakrejda, K., Stukalov, A., Lawrence, M., Goedman, R. J., Van Horn, K. S., Vehtari, A., Gabry, J., Casallas, J. S., and Bales, B. (2017). Stan-dev/stan: V2.17.1. Zenodo.
- Lin, S. and Luo, W. (2019). A New Multilevel CART Algorithm for Multilevel Data with Binary Outcomes. *Multivariate Behavioral Research*, 54(4):578–592.
- Little, R. J. A. and Rubin, D. B. (2002). *Statistical Analysis with Missing Data*. Wiley Series in Probability and Statistics. Wiley, 1 edition.

- Lüdtke, O., Robitzsch, A., and Grund, S. (2017). Multiple imputation of missing data in multilevel designs: A comparison of different strategies. *Psychological Methods*, 22(1):141–165.
- Maas, C. J. M. and Hox, J. J. (2005). Sufficient Sample Sizes for Multilevel Modeling. *Methodology*, 1(3):86–92.
- Meng, X.-L. (1994). Multiple-imputation inferences with uncongenial sources of input. *Statistical science*, pages 538–558.
- Mistler, S. A. and Enders, C. K. (2017a). A Comparison of Joint Model and Fully Conditional Specification Imputation for Multilevel Missing Data. *Journal of Educational and Behavioral Statistics*, 42(4):432–466.
- Mistler, S. A. and Enders, C. K. (2017b). A Comparison of Joint Model and Fully Conditional Specification Imputation for Multilevel Missing Data. *Journal of Educational and Behavioral Statistics*, 42(4):432–466.
- Morris, T. P., White, I. R., and Crowther, M. J. (2019). Using simulation studies to evaluate statistical methods. *Statistics in Medicine*, 38(11):2074–2102.
- Murray, D. M. and Blitstein, J. L. (2003). Methods To Reduce The Impact Of Intraclass Correlation In Group-Randomized Trials. *Evaluation Review*, 27(1):79–103.
- Oberman, H. I. and Vink, G. (2023). Toward a standardized evaluation of imputation methodology. *Biometrical Journal*, page 2200107.
- Quartagno, M. and Carpenter, J. R. (2022). Substantive model compatible multilevel multiple imputation: A joint modeling approach. *Statistics in Medicine*, 41(25):5000–5015.
- R Core Team (2023). *R: A Language and Environment for Statistical Computing*. Vienna, Austria.
- Resche-Rigon, M. and White, I. R. (2018). Multiple imputation by chained equations for systematically and sporadically missing multilevel data. *Statistical Methods in Medical Research*, 27(6):1634–1649.
- Rights, J. D. and Sterba, S. K. (2019). Quantifying explained variance in multilevel models: An integrative framework for defining R-squared measures. *Psychological Methods*, 24(3):309–338.
- Robitzsch, A., Simon Grund, and Henke, T. (2024). Miceadds: Some Additional Multiple Imputation Functions, Especially for 'mice'.
- Rubin, D. B. (1976). Inference and missing data. *Biometrika*, 63(3):581–592.
- Rubin, D. B. (1987). *Multiple Imputation for Nonresponse in Surveys*. Wiley, New York.
- Salditt, M., Humberg, S., and Nestler, S. (2023). Gradient Tree Boosting for Hierarchical Data. *Multivariate Behavioral Research*, pages 1–27.
- Schouten, R. M., Lugtig, P., and Vink, G. (2018). Generating missing values for simulation purposes: A multivariate amputation procedure. *Journal of Statistical Computation and Simulation*, 88(15):2909–2930.
- Schouten, R. M. and Vink, G. (2021). The Dance of the Mechanisms: How Observed Information Influences the Validity of Missingness Assumptions. *Sociological Methods & Research*, 50(3):1243–1258.
- Shieh, G. (2012). A comparison of two indices for the intraclass correlation coefficient. *Behavior Research Methods*, 44(4):1212–1223.
- Silva, G. C. and Gutman, R. (2022). Multiple imputation procedures for estimating causal effects with multiple treatments with application to the comparison of healthcare providers. *Statistics in Medicine*, 41(1):208–226.
- Taljaard, M., Donner, A., and Klar, N. (2008). Imputation Strategies for Missing Continuous Outcomes in Cluster Randomized Trials. *Biometrical Journal*, 50(3):329–345.
- Tan, Y. V., Flannagan, C. A. C., and Elliott, M. R. (2016). Predicting human-driving behavior to help driverless vehicles drive: Random intercept Bayesian Additive Regression Trees.

- Van Buuren, S. (2007). Multiple imputation of discrete and continuous data by fully conditional specification. *Statistical Methods in Medical Research*, 16(3):219–242.
- van Buuren, S. (2018). *Flexible Imputation of Missing Data*. Chapman & Hall/CRC Interdisciplinary Statistics Series. CRC Press, Taylor & Francis Group, Boca Raton London New York, second edition.
- Wagner, J., West, B. T., Elliott, M. R., and Coffey, S. (2020). Comparing the Ability of Regression Modeling and Bayesian Additive Regression Trees to Predict Costs in a Responsive Survey Design Context. *Journal of Official Statistics*, 36(4):907–931.
- Waljee, A. K., Mukherjee, A., Singal, A. G., Zhang, Y., Warren, J., Balis, U., Marrero, J., Zhu, J., and Higgins, P. D. (2013). Comparison of imputation methods for missing laboratory data in medicine. *BMJ Open*, 3(8):e002847.
- Wundervald, B., Parnell, A., and Domijan, K. (2022). Hierarchical Embedded Bayesian Additive Regression Trees.
- Xu, D., Daniels, M. J., and Winterstein, A. G. (2016). Sequential BART for imputation of missing covariates. *Biostatistics*, 17(3):589–602.