

Master Research Report: Multilevel Multivariate Imputation by Chained Equations through Bayesian Additive Regression Trees

*Methodology and Statistics for the Behavioural, Biomedical and Social
Sciences*

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Word count:

Candidate Journal:

FETC Case Number:

Supervisors:

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Computational Statistics & Data Analysis
23-1778

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1 Introduction

1.1 Introducing missing data, multiple imputation & multilevel data structure

Incomplete data is a common challenge in many fields of research. A common approach for dealing with incomplete data is to remove all missing values from the data. However, this could possibly lead to biased results if the data is not Missing Completely At Random (MCAR) [van Buuren, 2018, Kang, 2013, Enders, 2017, Austin et al., 2021]. MCAR is one of the missing data mechanisms described by Rubin [Rubin, 1976]. Where MCAR means the cause of the missing data are unrelated to the data, Missing At Random (MAR) that it is related to observed data and Missing Not At Random (MNAR) that it is related to unobserved data [van Buuren, 2018, Rubin, 1976]. Furthermore, other approaches to dealing with incomplete data include: pairwise deletion, mean imputation and regression imputation, which also yield biased results [van Buuren, 2018].

Multiple imputation (MI) is considered a valid method for dealing with incomplete data [Mistler and Enders, 2017, van Buuren, 2018, Enders, 2017, Burgette and Reiter, 2010, Austin et al., 2021, Audigier et al., 2018, Van Buuren, 2007, Grund et al., 2021, Hughes et al., 2014]. MI imputes each missing value more than once, thereby considering necessary variation associated with the missingness problem. The multiply imputed data sets are analyzed, and the corresponding inferences are pooled according to Rubin’s rules [van Buuren, 2018, Austin et al., 2021, Rubin, 1987]. Generally, multiple imputation operates under two frameworks: joint modeling (JM) and fully conditional specification (FCS) [Mistler and Enders, 2017, van Buuren, 2018, Enders et al., 2018a, Enders et al., 2018b, Hughes et al., 2014]. JM employs a multivariate data distribution and regresses incomplete variables on complete variables to impute missing values. FCS, or chained equations, iteratively imputes one variable with missing values at a time through conditional univariate distributions regressing an incomplete variable complete and previously imputed variables [Mistler and Enders, 2017, van Buuren, 2018, Enders et al., 2018a, Enders et al., 2018b, Hughes et al., 2014].

JM and FCS are extended to a multilevel or hierarchical context. Multilevel data is hierarchically structured, where, for example, students are nested within schools, or patients are nested within hospitals [Hox et al., 2017]. JM is extended by defining a multivariate linear mixed model. FCS is extended by defining a series of univariate linear mixed models [Mistler and Enders, 2017].

1.2 Literature review (difficulty of imputing multilevel data)

Two ad-hoc strategies for dealing with multilevel missing data are: ignoring the multilevel structure and fixed effect imputation: adding group dummy variables representing the group effects [Lüdtke et al., 2017, Enders et al., 2016]. However, these strategies produce biased estimates of variance components and multilevel regression coefficients [Lüdtke et al., 2017]. Currently, the implementation of JM and FCS in a multilevel context are appropriate in a two-level random intercept context with normally distributed data. However, they differ beyond that: JM is more capable of handling within- and between- cluster relationships, random intercepts and incomplete categorical variables, while FCS is better suited for random slopes and restricted to normally distributed variables [Enders et al., 2016]. Also, they differ in their handling of missing level-2 data. Overall, FCS is believed to be more flexible than JM [Audigier et al., 2018] and, thus, may be better suited for multilevel data.

1.3 Relevance of research

Currently, the specifications of the imputation models in a multilevel context are quite complex [van Buuren, 2018]: they should at least be as general as the analysis model [Grund et al., 2018] and preferably all-encompassing. However, the complexity of the multilevel analysis model is built step-wise with non-linearities [Hox et al., 2017] and a very complex model might not converge [van Buuren, 2018]. The package MICE [Buuren and Groothuis-Oudshoorn, 2011] implements the following methods in the FCS framework: *21.bin*, *21.lmer*, *21.pan*, *21.continuous*, *21.jomo*, *21.glm.norm*, *21.norm*, *21.2stage.norm*, *21.pmm*, and *21.2stage.pmm*.

Bayesian Additive Regression Trees (BART) is a sum-of-trees model proposed by Chipman et al. [Chipman et al., 2010]. Regression trees are its building blocks [Chipman et al., 2010, Hill et al., 2020, James et al., 2021]. Regression trees model non-linearities well and automatically through recursive binary partitioning of the predictor space [Hill et al., 2020]. Recursive binary partitioning doesn’t assume a specific data form; it divides the predictor space to maximize variance explanation by automatically identifying best fitting splits [Hastie, 2017, James et al., 2021, Salditt et al., 2023]. BART models can be described as:

$$y_i = f(\mathbf{x}_i) + \epsilon_i, \quad (1.1)$$

$$y_i = g(\mathbf{x}_i, T_1, M_1) + g(\mathbf{x}_i, T_2, M_2) + \dots + g(\mathbf{x}_i, T_m, M_m) + \epsilon_i, \quad (1.2)$$

where y_i is the outcome variable for person i , $f(\mathbf{x}_i)$ is the sum-of-trees many regression trees, and ϵ_i is the error term; $\epsilon \sim \mathcal{N}(0, \sigma^2)$. \mathbf{x} are the predictors included in the model, T_2 is the tree and M_2 is the collection of leaf parameters within each tree [Chipman et al., 2010, Hill et al., 2020, James et al., 2021]. Next to the sum-of-trees model, BART also includes a regularization prior that constrains the size and fit of each tree so that each contributes only a small part of the overall fit to prevent overfitting [Chipman et al., 2010, Hill et al., 2020, James et al., 2021]. The Bayesian backfitting Markov Chain Monte Carlo (MCMC) algorithm is used to obtain estimates from BART. It updates each tree, conditional on the remaining trees, their associated parameters and σ , by fitting a new tree to the partial residuals, r_i , perturbing the tree from the previous iteration. The partial residuals, r_i , are defined as:

$$r_i = y_i - \sum_{k' < k} \hat{f}_{k'}^b(x_i) - \sum_{k' > k} \hat{f}_{k'}^{b-1}(x_i), \quad (2)$$

where $\hat{f}_{k'}^b(x_i)$ is the prediction of the k' th tree in the b th iteration for person i .

In a single-level context, the use of tree-based models like regression trees, random forests or BARTs simplified imputation models and performed better than parametric methods: the estimates showed better confidence interval coverage of the population parameters, lower variance and lower bias [Burgette and Reiter, 2010, Xu et al., 2016]. Also in a multilevel prediction context, BART provides better estimates with a lower Mean Squared Error (MSE) and lower relative bias compared to the standard multilevel models [Wagner et al., 2020, Chen, 2020]. However, their use in multiple imputation in a multilevel context is yet to be implemented, even though their performance in a single-level context seems promising [Burgette and Reiter, 2010, Xu et al., 2016].

1.4 Research question

Can multivariate imputation by chained equations through a multilevel bayesian additive regression trees model improve the bias, variance and coverage of the estimates in a multilevel context compared to current practices?

1.5 Hypotheses

Given the success of non-parametric models in single-level multiple imputation, I anticipate that employing multilevel BART models in a multilevel missing data context will reduce bias, accurately model variance, and improve estimate coverage compared to classical multilevel imputation through 21.pmm in MICE.

2 Method

3 Results

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