

Bayes Rule.

$$P(A|B)P(B) = P(B|A)P(A)$$

or

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

A = Parameters ( $\theta$ )

B = Data ( $D$ )

$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)}$$

← LIKELIHOOD      ← PRIOR      ← PROBABILITY  
 ↓                    ↓                    ( like  
 POSTERIOR      P(D)      gaussian  
 PROBABILITY      ← EVIDENCE      with  
 ( Prob. after                mean  
 seeing data )                 $\theta$ )  
 { ALSO ACTS AS  
 REGULARIZER }

# MLE FOR LINEAR REGRESSION

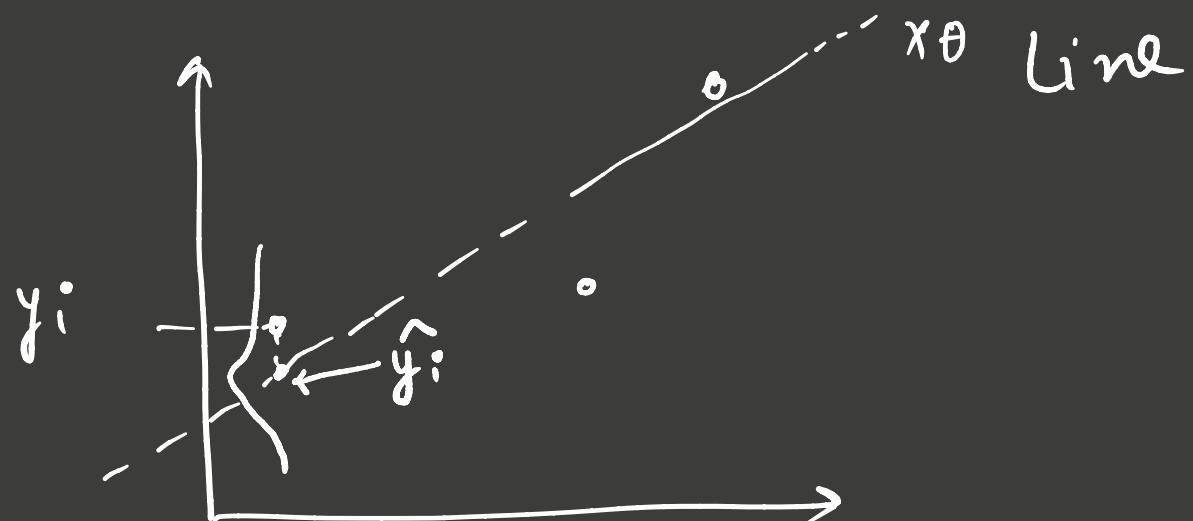
$$\hat{\theta}_{LS} = \underset{\theta}{\operatorname{argmin}} \epsilon^T \epsilon = \underset{\theta}{\operatorname{argmin}} (y - X\theta)^T (y - X\theta)$$

LEAST  
SQUARES  
SOL<sup>N</sup>

①  $\epsilon_i \sim N(0, \sigma^2)$  (0 MEAN, KNOWN VARIANCE)

②  $\epsilon_i$  INDEPENDENT ACROSS OBSERVATIONS

③  $y \sim N(X\theta, \sigma^2) \quad \{ X\theta + N(0, \sigma^2) \}$



Likelihood =  $P(\text{Data} | \text{Parameters})$   
 Data =  $\langle y_i, x_i \rangle$ ; Parameters =  $\theta, \sigma^2$

$$P(y | X, \theta, \sigma^2) = P(y_1 | x_1, \theta, \sigma^2) P(y_2 | x_2, \theta, \sigma^2) \dots$$

$$= \prod_{i=1}^n P(y_i | x_i, \theta, \sigma^2)$$

$$= \frac{n}{\prod_{i=1}^n} \left( \frac{1}{2\pi\sigma^2} \right)^{\frac{y_i - x_i \theta}{2}} e^{-\frac{(y_i - x_i \theta)^2}{2\sigma^2}}$$

$$\text{Log Likelihood} = \sum_{i=1}^n \text{Constant} * (y_i - x_i \theta)^2 = \text{Constant} * \sum_{i=1}^n (y_i - x_i \theta)^2$$

$$\hat{\theta}_{MLE} = \underset{\theta}{\operatorname{arg\max}} (y - x\theta)^T (y - x\theta) = \hat{\theta}_{LS}$$

UNDER NORMALLY DISTRIBUTED RESIDUALS

# MAP FOR LINEAR REGRESSION

- ①  $\epsilon \sim N(0, \sigma^2)$
- ②  $\epsilon$  is independent across observations
- ③  $y \sim N(x\theta, \sigma^2)$
- ④  $\theta$  has Gaussian prior i.e.  $\theta \sim N(0, \tau^2 I)$  Identity

$$P(\theta|D) \propto P(D|\theta) P(\theta) \Rightarrow \log P(\theta|D) \propto \log P(D|\theta) + \log P(\theta)$$

$$\hat{\theta}_{MAP} = \arg \max_{\theta} \{ \log P(D|\theta) + \log P(\theta) \}$$

$$= \arg \min_{\theta} (y - X\theta)^T (y - X\theta) + \lambda \theta^T \theta$$

MAP with GAUSSIAN PRIOR  $\Rightarrow$   
RIDGE REGRESSION

PRIOR  $\xrightarrow{\text{LEADS TO}}$  REGULARIZATION

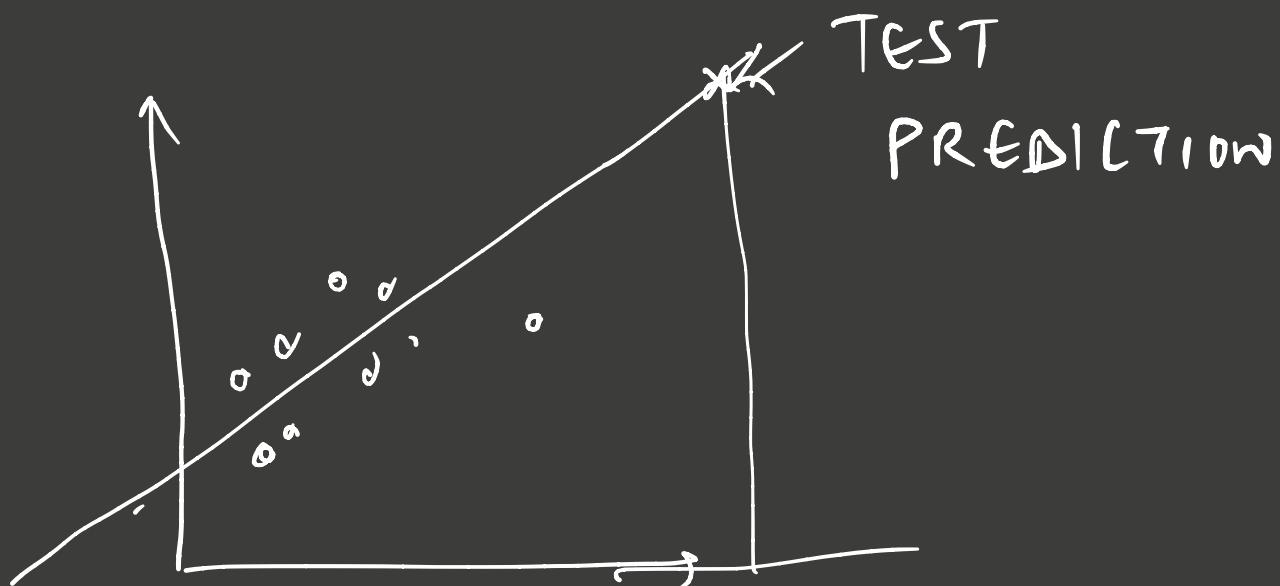
$\theta \sim N(0, \tau^2 I)$  : RIDGE REGRESSION

$\theta_i \sim \text{LAPLACE}(0, t)$  : LASSO

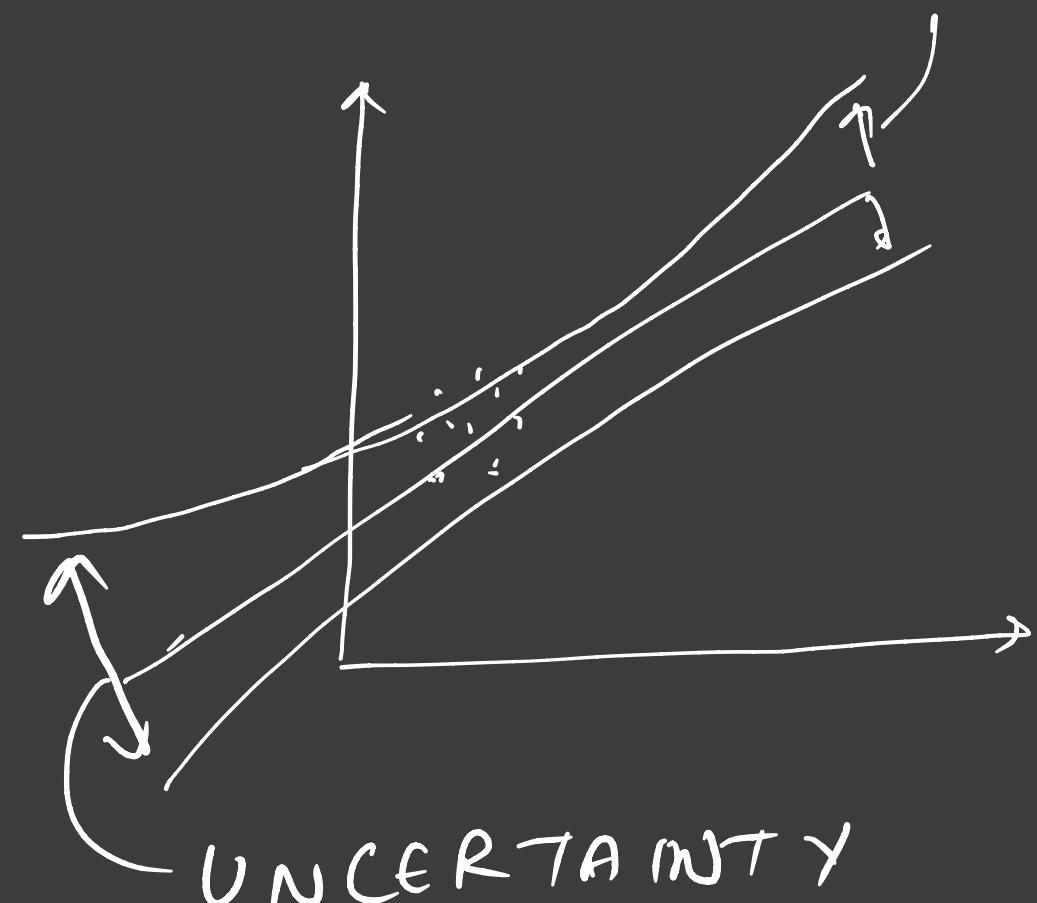
# BAYESIAN LINEAR REGRESSION

- \* POINT ESTIMATE  $\rightarrow$  NO MEASURE OF UNCERTAINTY

NON "FULLY"  
BAYESIAN  
(MAP; MLE)



"FULLY" BAYESIAN



Q : What is  $P(y^* | x^*, D)$

$$y_i \sim N(x_i \theta, \bar{a}^i) \xrightarrow{\text{Precision } (\alpha = 1/\sigma^2)} \text{UNIVARIATE}$$

where  $x_i \in \mathbb{R}^d$

$$\theta \sim N(\theta, b^{-1} I) \xrightarrow{\text{Precision}} \text{MULTIVARIATE}$$

$$a, b > 0$$

ASSUMPTION

$a, b$  are KNOWN

$$p(D|\theta) \propto e^{-\alpha_2 \frac{(y-x\theta)^T(y-x\theta)}{2}} \quad \{\text{likelihood}\}$$

$$\begin{aligned} p(\theta|D) &\propto p(D|\theta)p(\theta) \\ &\propto e^{-\alpha_2 \frac{(y-x\theta)^T(y-x\theta)}{2} - b_2 \frac{\theta^T \theta}{2}} \\ &\propto e^{-\alpha_2 \frac{(y-x\theta)^T(y-x\theta)}{2} - b_2 \theta^T \theta} \end{aligned}$$

$$\begin{aligned} \text{Let } Z &= \alpha (y-x\theta)^T (y-x\theta) + b \theta^T \theta \\ &= \alpha y^T y - 2\alpha \theta^T x^T y + \theta^T (\alpha x^T x + b I) \theta \end{aligned}$$

$$\text{Let } Z = (\theta - \mu)^T \Lambda (\theta - \mu)$$

$$\theta^T \Lambda \theta - 2 \theta^T \Lambda \mu + \text{constant} = \alpha y^T y - 2\alpha \theta^T x^T y + \theta^T (\alpha x^T x + b I) \theta$$

$$\Rightarrow \hat{\Lambda} = \alpha X^T X + b I$$

$$\text{and } \hat{\mu} = \alpha \hat{\Lambda}^{-1} X^T y$$

$$p(\theta | D) \sim N(\theta | \mu, \hat{\Lambda}^{-1})$$

NOW WE WANT

$p(y^* | x^*, D, \alpha, b)$  Predictive distribution  
 $\downarrow (x^*, y^*)$  new point

USE ALL  $\theta$

$$p(y^* | x^*, D, a, b) = \int p(y^* | \theta, x^*, a) p(\theta | D, a, b) d\theta$$

$$= \int N(y^* | \hat{x}^{\theta}, \hat{a}') N(\theta | \mu, \hat{\sigma}^2) d\theta$$

$$\propto \frac{-\alpha}{2} \left( y^* - \hat{x}^{\theta} \right)^2 - \frac{1}{2} (\theta - \mu)^T \Lambda (\theta - \mu) d\theta$$

$\uparrow$   
 $F$

$$\begin{aligned} F &= (\theta - \mu)^T \Lambda (\theta - \mu) \\ &= \theta^T (\alpha x^* x^{*\top} + \Lambda) \theta - 2 \theta^T (\alpha y^* a + \Lambda \mu) \\ &\quad + \alpha y^2 \end{aligned}$$

$$L = \alpha \mathbf{x}^* \mathbf{x}^{*\top} + \lambda$$

$$\mathbf{m} = L^{-1} (\alpha \mathbf{y}^* \mathbf{x}^* + \lambda \boldsymbol{\mu})$$

⋮  
⋮  
⋮  
⋮  
⋮

$$P(y^* | \mathbf{x}^*, D) = N(y^* | u, \frac{1}{\lambda})$$

$$u = \mu^T \mathbf{x}^*$$

$$\frac{1}{\lambda} = \frac{1}{\alpha} + \mathbf{x}^{*\top} \Lambda \mathbf{x}^*$$

GRAPHICAL  
MODEL  
FOR  
BAYESIAN  
LINEAR  
REGRESSION

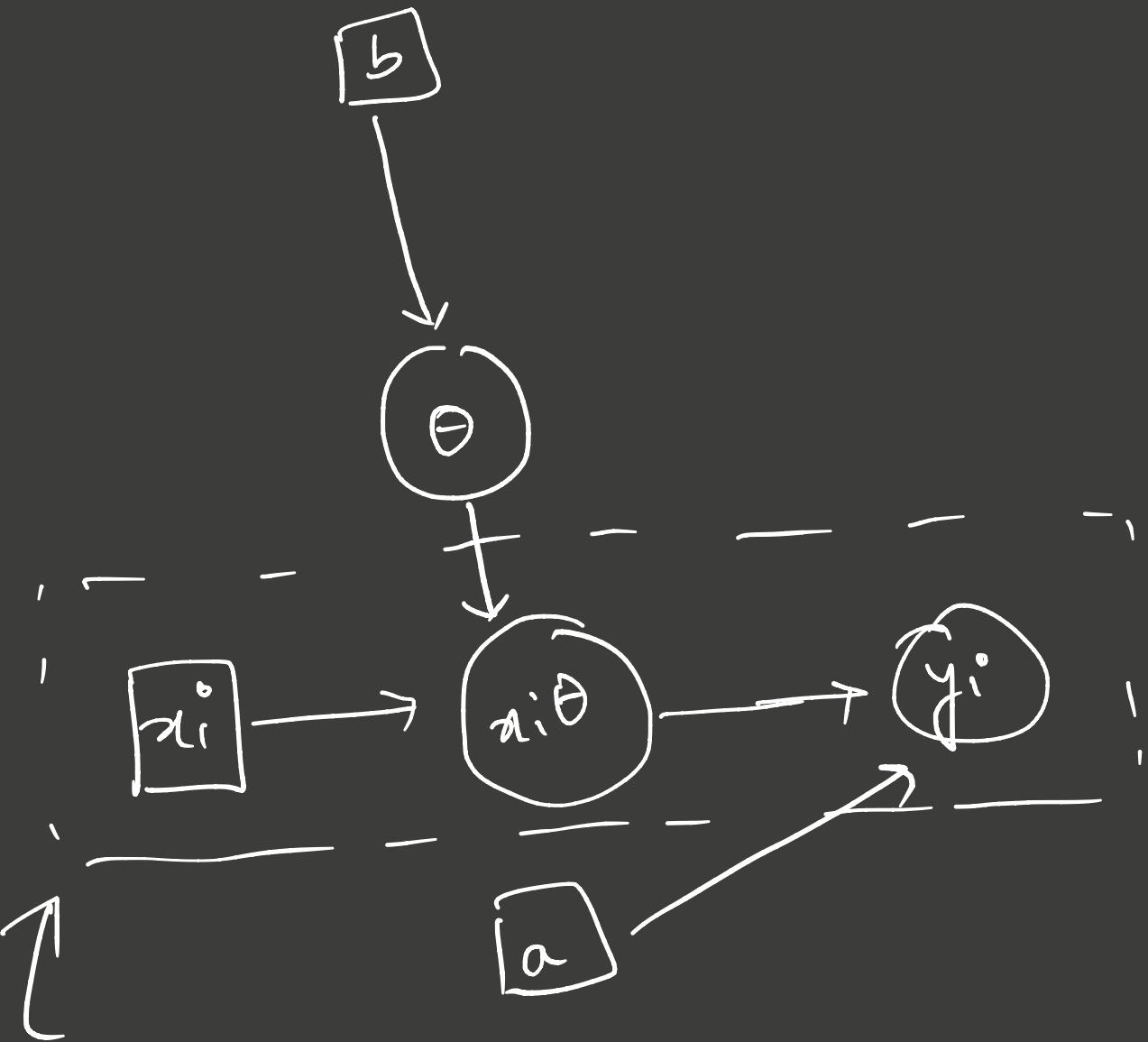


Plate  
model  
(repeated  
over time | i)