Day 4, Lecture 2

Longitudinal TMLE (LTMLE)

## Longitudinal TMLE (LTMLE)

#### In this lecture, our goal is to:

- Describe the implementation of the targeting algorithm for estimation of parameters defined under dynamic treatment interventions.
- Explain the application of TMLE to longitudinal data structures to evaluate effects of static and dynamic effects in presence and without presence of right-censoring and with and without use of super learning.

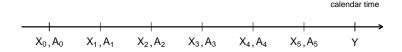
#### Overview

- 1. Targeting algorithm
  - Practical 2 (Task 1)
- 2. ltmle software package
  - Practical 2 (Task 2ff)
- 3. A note on right-censoring

#### Longitudinal data structure

#### Longitudinal data structure:

- $O = (X_0, A_0, X_1, A_1, \dots, X_K, A_K, Y) \in (\mathbb{R}^d \times \{0, 1\})^K \times \{0, 1\}$
- Covariates  $X = (X_0, X_1, \dots, X_K)$  change over time
- ► Treatment decisions  $A = (A_0, A_1, ..., A_K)$  are updated over time
- Covariates and treatment decisions interact in complex ways



NB: Still keeping right-censoring (and competing risks) out of the picture.

# Longitudinal targeting

(For the representation in terms of iterated conditional expectations)

#### Recall —

TMLE is a two-step procedure:

- Step 1 Construct initial estimator  $\hat{P}_n$  for P.
- Step 2 Update the estimator  $\hat{P}_n \mapsto \hat{P}_n^*$  such that  $\hat{P}_n^*$  solves the efficient influence curve equation.

```
Step 1 = "initial estimation step"
```

Step 2 = "targeting step"

#### Recall — for the ATE:

TMLE is a two-step procedure:

- Step 1 Construct initial estimators  $\hat{f}_n$ ,  $\hat{\pi}_n$  for f,  $\pi$ .
- Step 2 Update the estimator  $\hat{f}_n \mapsto \hat{f}_n^*$  for f such that  $\hat{f}_n^*$  for the fixed  $\hat{\pi}_n$  solves the efficient influence curve equation.

```
Step 1 = "initial estimation step"
```

Step 2 = "targeting step"

The relevant part of P needed to evaluate our target parameter:

$$\mathbb{E}_{P}[Y^{A_{0}=a_{0}^{*},A_{1}=a_{1}^{*},...,A_{K}=a_{K}^{*}}]=\tilde{\Psi}(\bar{Q}),$$

with 
$$\bar{Q} = (\bar{Q}_k)_{1 \le k \le K+1}$$
.

Starting backwards from the last time-point:

$$\bar{Q}_{K+1}(\bar{x}_K, \bar{a}_K) = \mathbb{E}_P[Y \mid \bar{X}_K = \bar{x}_K, \bar{A}_K = \bar{a}_K]$$

and iteratively for k = K, K - 1, ..., 1,

$$\bar{Q}_k(\bar{x}_{k-1},\bar{a}_{k-1}) = \mathbb{E}_P\big[\bar{Q}_{k+1}(\bar{X}_k,a_k^*,\bar{A}_{k-1}) \mid \bar{X}_{k-1} = \bar{x}_{k-1},\bar{A}_{k-1} = \bar{a}_{k-1}\big]$$

so that

$$\mathbb{E}_{P}\big[\,Y^{A_{0}=a_{0}^{*},A_{1}=a_{1}^{*},\ldots,A_{K}=a_{K}^{*}}\,\big]=\mathbb{E}_{P}\big[\,\bar{Q}_{1}\big(X_{0},a_{0}^{*}\big)\big]=\tilde{\Psi}\big(\,\bar{Q}\big).$$

We need the efficient influence function:

- ▶ Tells us what we need to estimate (to construct TMLE)
- Guides the construction of the targeting step

Construction of the targeting step for a given target parameter  $\Psi: \mathcal{M} \to \mathbb{R}$  with efficient influence function  $\phi^*(P)$  involves:

- (i) A parametric submodel  $\{ar{Q}_{arepsilon}: arepsilon \in \mathbb{R}\} \subset \mathcal{M}$
- (ii) A loss function  $(O, \bar{Q}) \mapsto \mathcal{L}(\bar{Q})(O)$

such that: (1) 
$$\bar{Q}_{\varepsilon=0} = \bar{Q}$$
, and, (2)  $\frac{d}{d\varepsilon}\Big|_{\varepsilon=0} \mathcal{L}(\bar{Q}_{\varepsilon})(O) = \phi^*(P)(O)$ 

The efficient influence function is given by:

$$\begin{split} \phi^*(P)(O) &= \tilde{\phi}^*(\bar{Q},\pi)(O) \\ &= \sum_{k=1}^{K+1} \bigg( \prod_{l=0}^{k-1} \frac{1\{A_l = a_l^*\}}{\pi_{A_l}(a_l^* \mid \bar{x}_l, \bar{a}_{l-1}^*)} \bigg) \Big\{ \bar{Q}_{k+1}(\bar{X}_k, a_k^*, \bar{A}_{k-1}) - \bar{Q}_k(\bar{X}_{k-1}, \bar{A}_{k-1}) \Big\} \\ &\quad + \bar{Q}_1(X_0, a_0^*) - \Psi(P) \end{split}$$
 (with  $\bar{Q}_{K+2} \coloneqq Y$ )

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 (with  $\bar{Q}_{K+2} \coloneqq Y$ )

- ▶ Need initial estimators for:  $\pi = (\pi_{A_k})_{0 \le k \le K}$ ,  $\bar{Q} = (\bar{Q}_k)_{1 \le k \le K+1}$
- Submodel and loss function for each  $\bar{Q}_k$  in turn to solve the k-specific part of the efficient influence curve equation,

$$\tilde{\phi}_k^*(\bar{Q},\pi)(O) = \left(\prod_{l=0}^{k-1} \frac{1\{A_l = a_l^*\}}{\pi_{A_l}(a_l^* \mid \bar{X}_l, \bar{a}_{l-1}^*)}\right) \left\{\bar{Q}_{k+1}(\bar{X}_k, a_k^*, \bar{A}_{k-1}) - \bar{Q}_k(\bar{X}_{k-1}, \bar{A}_{k-1})\right\}$$

We construct a submodel  $\bar{Q}_{k,\varepsilon}$  through a given  $\bar{Q}_k$  and a loss function  $(O,\bar{Q}_k)\mapsto \mathcal{L}_{\bar{Q}_{k+1}}(\bar{Q}_k)(O)$  such that

(1) 
$$|\bar{Q}_{k,\varepsilon=0}| = |\bar{Q}_k|$$
 and, (2)  $|\frac{d}{d\varepsilon}|_{\varepsilon=0} \mathscr{L}_{\bar{Q}_{k+1}}(\bar{Q}_{k,\varepsilon})(O) = \phi_k^*(\bar{Q},\pi)(O)$ 

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We define:

$$\mathcal{L}_{\bar{Q}_{k+1}}(\bar{Q}_k) = -(\bar{Q}_{k+1}\log(\bar{Q}_k) + (1 - \bar{Q}_{k+1})\log(1 - \bar{Q}_k))$$

and,

$$\bar{Q}_{k,\varepsilon}(O) = \mathrm{expit} \big( \mathrm{logit} \big( \bar{Q}_k \big( \bar{X}_{k-1}, \bar{A}_{k-1} \big) \big) + \varepsilon H_k(\pi)(O) \big)$$

with the "clever covariate":  $H_k(\pi)(O) \coloneqq \prod_{l=0}^{k-1} \frac{1\{A_l = a_l^*\}}{\pi_{A_l}(a_l^* \mid \bar{x}_l, \bar{a}_{l-1}^*)}$ 

Another valid choice of loss function and submodel would be:

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and,

$$\begin{split} \bar{Q}_{k,\varepsilon}(O) &= \mathrm{expit} \big( \mathrm{logit} \big( \bar{Q}_k \big( \bar{X}_{k-1}, \bar{A}_{k-1} \big) \big) + \varepsilon \big) \\ \text{using the "clever covariate"} \ H_k(\pi)(O) &\coloneqq \prod_{l=0}^{k-1} \frac{1\{A_l = a_l^*\}}{\pi_{A_l} \big( a_l^* \mid \bar{X}_l, \bar{a}_l^* \mid z_l \big)} \end{split}$$

as a weight.

The targeting step becomes a targeting algorithm that proceeds iteratively along the sequence of iterated conditional expectations, *starting from the last time-point*.

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We then solve:

$$\frac{1}{n} \sum_{i=1}^{n} H_{K+1}(\hat{\pi})(O_i) \left\{ \hat{\bar{Q}}_{K+2}^* - \hat{\bar{Q}}_{K+1}^* (\bar{X}_{K,i}, \bar{A}_{K,i}) \right\} = 0$$

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We then solve:

$$\frac{1}{n}\sum_{i=1}^{n}H_{k}(\hat{\pi})(O_{i})\left\{\hat{\bar{Q}}_{k+1}^{*}(\bar{X}_{k,i},a_{k}^{*},\bar{A}_{k-1,i})-\hat{\bar{Q}}_{k}^{*}(\bar{X}_{k-1,i},\bar{A}_{k-1,i})\right\}=0$$

The targeting step becomes a targeting algorithm that proceeds iteratively along the sequence of iterated conditional expectations, starting from the last time-point.

This procedure gives a sequence of updated estimators  $\hat{\bar{Q}}^* = (\hat{\bar{Q}}_{K+1}^*, \hat{\bar{Q}}_K^*, \dots, \hat{\bar{Q}}_1^*)$  that solves the efficient influence curve equation:

$$\begin{split} \frac{1}{n} \sum_{i=1}^{n} \sum_{k=1}^{K+1} \left( \prod_{l=0}^{k-1} \frac{1\{A_{l,i} = a_{l}^{*}\}}{\hat{\pi}_{A_{l}}(a_{l}^{*} \mid \bar{X}_{l,i}, \bar{a}_{l-1}^{*})} \right) \left\{ \hat{\bar{Q}}_{k+1}^{*}(\bar{X}_{k,i}, a_{k}^{*}, \bar{A}_{k-1,i}) \right. \\ & \left. - \hat{\bar{Q}}_{k}^{*}(\bar{X}_{k-1}, \bar{A}_{k-1}) \right\} + \hat{\bar{Q}}_{1}^{*}(X_{0}, a_{0}^{*}) - \Psi(\hat{\bar{Q}}^{*}) \end{split}$$

where

$$\tilde{\Psi}(\hat{\bar{Q}}^*) = \frac{1}{n} \sum_{i=1}^n \hat{\bar{Q}}_1^*(X_{0,i}, a_0^*).$$

# Practical 2: Implementing the targeting step (Task 1)

Read the first page of  $day3\_practical2.pdf$  and then go through the steps of Task 1.

- Specifying static and dynamic interventions
- Initial estimation and super learning
- LTMLE and ltmle for right-censored data structures

```
See also: https:
//cran.r-project.org/web/packages/ltmle/ltmle.pdf
```

```
install.packages(ltmle)
library(ltmle)
```

Some useful arguments to know:

```
ltmle(data,
          Anodes, Lnodes, Cnodes, Ynodes,
          abar, rule,
          Qform, gform,
          SL.library,
          gbounds,
          ...)
```

data: data frame in wide format

NB the order of columns correspond to the order of variables.

Anodes: names of treatment variables

Lnodes: names of covariates

Cnodes: names of censoring variables

Ynodes: names of outcome variables

NB All variables except baseline covariates must be specified in Anodes, Lnodes, Cnodes or Ynodes.

For the simulated dataset from the practicals:

```
X0.1 X0.2 X0.3 A0 X1.1 X1.2 A1 Y

1: 1.2915853 -0.29726781 0 0 1 0 0 1

2: 0.8407983 -0.48032616 0 0 0 1 0 0

3: 1.8633391 -0.50152323 0 1 1 1 0 1

4: -1.6855569 -1.57948481 0 1 0 0 1 1

5: -1.7854517 0.04438914 0 0 1 0 1

6: 0.2999129 -1.41470763 0 1 0 1 0 0
```

```
Anodes = paste0("A",0:1)
Lnodes = c(paste0("X0.", 1:3), paste0("X1.", 1:2))
Ynodes = "Y"
```

The ordering of columns dictates the temporal ordering of variables:

$$\begin{pmatrix} X_{0,1} \\ X_{0,2} \\ X_{0,3} \end{pmatrix} \rightarrow A_0 \rightarrow \begin{pmatrix} X_{1,1} \\ X_{1,2} \end{pmatrix} \rightarrow A \rightarrow Y$$

### Specifying interventions:

- abar: a binary vector of treatment assignments of length = length(Anodes) or a list of two elements to contrast treatment regimes
  - to specify static treatment regimes
- rule: a function that can be applied to each row of the data to return a binary vector of treatment assignments of length = length(Anodes)
  - to specify dynamic treatment regimes

### For example:

[output next slide]

```
Treatment Estimate:
   Parameter Estimate: 0.3029
   Estimated Std Err: 0.086738
              p-value: 0.00047919
   95% Conf Interval: (0.1329, 0.4729)
Control Estimate:
   Parameter Estimate: 0.36688
   Estimated Std Err: 0.015419
             p-value: <2e-16
   95% Conf Interval: (0.33666, 0.3971)
Additive Treatment Effect:
   Parameter Estimate: -0.063978
   Estimated Std Err: 0.088095
              p-value: 0.46769
   95% Conf Interval: (-0.23664, 0.10868)
```

Here note that Treatment Estimate and Control Estimate were fitted completely separately, and that they could had been obtained with separate calls:

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### Treatment Estimate:

### Control Estimate:

Parameter Estimate: 0.3029 Parameter Estimate: 0.36688
Estimated Std Err: 0.086738 Estimated Std Err: 0.015419

p-value: 0.00047919 p-value: <2e-16

95% Conf Interval: (0.1329, 0.4729) 95% Conf Interval: (0.33666, 0.3971)

► For each parameter we get the standard error, a the p-value of the null hypothesis that that quantity equals zero, and confidence intervals

ainfluence curve based if the argument variance.method="ic" is specified, otherwise a more robust variance estimator based on TMLE is used.

### Effect of a dynamic regime:

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Static regimes can also be specified as a dynamic regime:

- Specifying static and dynamic interventions
- Initial estimation and super learning
- LTMLE and ltmle for right-censored data structures

- Qform: character vector of regression formulas for the outcome regressions
  - Qform indicates what variables are included in each outcome regression
  - default is NULL which means that all variables from previous time-points are included
  - (does not mean that GLM is used)
- gform: character vector of regression formulas for the propensity scores
  - gform indicates what variables are included in each propensity score regression
  - default is NULL which means that all variables from previous time-points are included
  - (does not mean that GLM is used)

#### Default:

```
Qform not specified, using defaults:
formula for X1.1:
Q.kplus1 \sim X0.1 + X0.2 + X0.3 + A0
formula for Y:
Q.kplus1 \sim X0.1 + X0.2 + X0.3 + A0 + X1.1 + X1.2 + A1
gform not specified, using defaults:
formula for AO:
A0 \sim X0.1 + X0.2 + X0.3
formula for A1:
A1 \sim X0.1 + X0.2 + X0.3 + A0 + X1.1 + X1.2
speedglm failed, using glm instead. If you see a lot of this message, an
```

Extracting the first (last time-point) regression fit:

$$\bar{Q}_2(\bar{X}_1,\bar{A}_1) = \mathbb{E}_P[Y \mid \bar{X}_1,\bar{A}_1]$$

```
fit.ltmle$fit$Q[[1]]$Y
```

```
Estimate Std. Error t value
                                                  Pr(>|t|)
(Intercept) -0.91411304 0.10514428 -8.6938923 7.179800e-18
X0.1
             0.03225723 0.04248034 0.7593449 4.477361e-01
X0.2
            -0.05878768 0.04851366 -1.2117759 2.257419e-01
X0.3
             0.03744604 0.12002695 0.3119803 7.550882e-01
AΩ
            -0.29864339 0.10343002 -2.8873958 3.926343e-03
X1.1
             1.16277977 0.10398452 11.1822394 3.378290e-28
X1.2
             0.08077220 0.10204795 0.7915123 4.287394e-01
A 1
            -0.70956699 0.26311107 -2.6968345 7.059175e-03
```

Extracting the next/last (first time-point) regression fit:

$$\bar{Q}_1(X_0, A_0) = \mathbb{E}_P[\bar{Q}_2(\bar{X}_1, \bar{A}_1) \mid X_0, A_0]$$

```
fit.ltmle$fit$Q[[1]]$X1.1
```

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-1.45814424	0.01937218	-75.270032	0.0000000000000000000000000000000000000
X0.1	0.04941305	0.01106743	4.464726	0.00000846490464004
X0.2	-0.04858049	0.01266330	-3.836320	0.00012877312827154
X0.3	0.11706316	0.03047067	3.841830	0.00012593634196146
AO	0.17059618	0.02525066	6.756106	0.0000000001853852

(If you ask me, the naming convention is quite strange).

Extracting the propensity score regression fit:

$$\pi_{A_1}(1 \mid \bar{X}_1, A_0) = \mathbb{E}_P[A_1 \mid \bar{X}_1, A_0]$$

```
fit.ltmle$fit$g[[1]]$A1
```

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) -4.16188796 0.33230949 -12.5241321 1.074919e-34
X0.1 0.15421078 0.08538514 1.8060610 7.105961e-02
X0.2 0.09551047 0.10273564 0.9296722 3.526534e-01
X0.3 0.36500100 0.23542597 1.5503854 1.212078e-01
A0 4.84411485 0.34081015 14.2135287 9.610307e-44
X1.1 -5.64663180 0.51985544 -10.8619269 9.664175e-27
X1.2 -5.21307450 0.42763985 -12.1903383 5.081408e-33
```

Extracting the propensity score regression fit:

$$\pi_{A_{\mathbf{0}}}(1\mid X_0) = \mathbb{E}_P[A_0\mid X_0]$$

### fit.ltmle\$fit\$g[[1]]\$A0

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.03114875 0.05043403 -0.6176137 0.5369005
X0.1 0.01069894 0.03928253 0.2723587 0.7853744
X0.2 0.01483267 0.04490924 0.3302810 0.7412223
X0.3 -0.04204218 0.11017986 -0.3815777 0.7028152
```

- ▶ There are many opportunities for misspecification!!
- Double robustness helps a bit for consistency
- Inference still requires getting all of it right

 SL.library: list with entries named Q and g specifying super learner libraries to pass to SuperLearner for the outcome regressions and the propensities scores

 SL.library: list with entries named Q and g specifying super learner libraries to pass to SuperLearner for the outcome regressions and the propensities scores

You can see available models for the super learner here: https://cran.r-project.org/web/packages/SuperLearner/ vignettes/Guide-to-SuperLearner.html (Section 4)

- NB Some algorithms are really very slow
- NB Some algorithms may not converge (gives error messages)
- NB Think about what you know about each particular algorithm and do not just blindly include a ton of heavy algorithms

You can also write your own algorithms.

For example, one could specify:

#### And then call ltmle:

We can extract the super learner weights applied to each algorithm from the ltmle object:

```
fit.ltmle.sl$fit$Q[[1]]
```

### \$X1.1

	Risk	Coef
SL.glm_All	0.01616588	0.97450398
SL.mean_All	0.01708647	0.02549602
${\tt SL.glm.interaction\_All}$	0.01624704	0.00000000
SL.glmnet_All	NA	0.00000000
SL.gam_All	0.01617732	0.00000000

#### \$Y

	Risk	Coef
SL.glm_All	0.2091238	0.74283811
SL.mean_All	0.2342586	0.02679133
${\tt SL.glm.interaction\_All}$	0.2115791	0.0000000
SL.glmnet_All	0.2094033	0.00000000
SL.gam_All	0.2091607	0.23037056

We can extract the super learner weights applied to each algorithm from the ltmle object:

```
fit.ltmle.sl$fit$g[[1]]
```

#### \$AO

```
Risk Coef
SL.glm_All 0.2508915 0
SL.mean_All 0.2502013 1
SL.glmnet_All 0.2502013 0
SL.gam_All 0.2512009 0
```

#### \$A1

	Risk	Coef
SL.glm_All	0.03332285	0.000000
SL.mean_All	0.08031019	0.000000
SL.glmnet_All	0.03330272	0.822445
SL.gam_All	0.03337456	0.177555

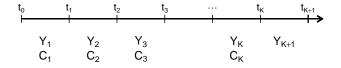
- ▶ Guidelines on what algorithms to use??¹
- Trade-off between computation time and performance??
- Studying performance on simulated data can be useful.
  - (but this is not trivial either).

<sup>&</sup>lt;sup>1</sup>Phillips, R. V., van der Laan, M. J., Lee, H., & Gruber, S. (2023). Practical considerations for specifying a super learner. International Journal of Epidemiology.

- Specifying static and dynamic interventions
- Initial estimation and super learning
- LTMLE and ltmle for right-censored data structures

Outcome process  $Y_k$  can jump from 0 to 1 at any time-point  $t_k$ .

Censoring process  $C_k$  can jump from 0 to 1 at any time-point  $t_k$ .



- ▶ Time to event:  $\tilde{T} = k' \wedge K + 1$  where  $k' = \min(k : Y_k = 1 \text{ or } C_k = 1)$
- Event indicator:  $\Delta = 1\{Y_{\tilde{T}} = 1\}$

One way that we can represent the survival data over a grid of K = 7 time-points (long format), e.g., for three different individuals:

```
id k Y C
                1: 5 1 0 0
  id k Y C
                2: 5 2 0 0
1: 2 1 0 0
                3: 5 3 0 0
                                    id k Y C
2: 2 2 0 0
              4: 5 4 0 0
                                 1: 7 1 0 0
3: 2 3 0 0
              5: 5500
                                 2: 7 2 0 1
4: 2410
                6: 5600
                7: 5 7 0 0
                8: 5800
```

One way that we can represent the survival data over a grid of K = 7 time-points (long format), e.g., for three different individuals:

```
id k Y C
  id k Y C
               2: 5 2 0 0
  2 1 0 0
               3: 5300
                                  id k Y C
2: 2200
             4: 5400
                               1: 7 1 0 0
3: 2300
             5: 5500
                               2: 7 2 0 1
4: 2410
               6:
                   5 6 0 0
               7:
                   5 7 0 0
               8:
                   5800
```

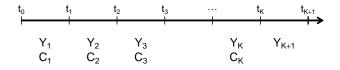
Same data may also be presented as (wide format):

One way that we can represent the survival data over a grid of K = 7 time-points (long format), e.g., for three different individuals:

```
id k Y C
               2: 5 2 0 0
  id k Y C
  2 1 0 0
               3: 5300
                                  id k Y C
2:
   2 2 0 0
              4: 5400
                               1: 7 1 0 0
3: 2300
             5: 5500
                               2: 7 2 0 1
4: 2410
               6:
                   5600
               7:
                   5 7 0 0
               8:
                   5800
```

Same data may also be presented as (wide format):

(The wide format is needed for ltmle).



### Example:

- ▶ Baseline covariate vector:  $X_0 = (X_{0,1}, X_{0,2}, X_{0,3})$
- ▶ Baseline treatment decision:  $A_0 \in \{0, 1\}$
- ▶ Time to event:  $\tilde{T} = k' \wedge K + 1$  where  $k' = \min(k : Y_k = 1 \text{ or } C_k = 1)$
- Event indicator:  $\Delta = 1\{Y_{\tilde{T}} = 1\}$

### Generally:

$$O = (X_0, A_0, Y_1, C_1, X_1, A_1, \dots, Y_K, C_K, X_K, A_K, Y = X_{K+1}).$$

- Covariates  $X = (X_0, X_1, \dots, X_K)$  change over time.
- ▶ Treatment decisions  $A = (A_0, A_1, ..., A_K)$  are updated over time.
- Censoring status  $C = (C_1, C_2, ..., C_K)$  change over time.
- ▶ Outcome (death) status  $Y = (Y_1, Y_2, ..., Y_K)$  change over time.

[After death/censoring, all variables are deterministically set to their last observed values].

Inferring on the uncensored event time is like imposing a simple static intervention on all right-censoring nodes to impose 'no censoring'.

Target causal parameter:

$$\Psi(P) = \mathbb{E}\big[Y_K^{A_0 = a_0^*, C_1 = 0, A_1 = a_k^*, \dots, C_K = 0, A_K = a_K^*}\big] - \mathbb{E}\big[Y_K^{A_0 = 0, C_1 = 0, A_0 = 0, \dots, C_K = 0, A_K = 0}\big]$$

I.e., the effect of the treatment had there been no loss to follow-up.

= the absolute risk by time  $t_{k^*}$  if everyone had received treatment  $(a_0^*, a_1^*, \dots, a_K^*)$  contrasted to the absolute risk if everyone had not been treated.

The ordering of

$$O = (X_0, A_0, Y_1, C_1, X_1, A_1, \dots, Y_K, C_K, X_K, A_K, Y = X_{K+1}),$$

implies a temporal ordering:

$$X_0 \to A_0 \to \cdots \to Y_k \to C_k \to X_k \to A_k \to \cdots \to Y_{K+1}.$$

- ▶ at each time-point  $t_k$ , a patient is only at risk of dying if they did not yet die and they were not yet right-censored  $(Y_{k-1} = 0, C_{k-1} = 0)$ .
- ▶ at each time-point  $t_k$ , a patient is only at risk of being right-censoring if they did not yet die at this time and they were not yet right-censored ( $Y_k = 0$ ,  $C_{k-1} = 0$ ).

Factorization of the density p of  $P \in \mathcal{M}$ :

$$\begin{split} p(o) &= \mu_{X_0}(x_0) \pi_{A_0}(a \mid x_0) \\ &\prod_{k=1}^K \left( \mu_{Y_k}(y_k \mid \bar{y}_{k-1}, \bar{c}_{k-1}, \bar{x}_{k-1}, \bar{a}_{k-1}) \pi_{C_k}(c_k \mid \bar{y}_k, \bar{c}_{k-1}, \bar{x}_k, \bar{a}_{k-1}) \right. \\ & \times \mu_{X_k}(x_k \mid \bar{y}_k, \bar{c}_k, \bar{x}_{k-1}, \bar{a}_{k-1}) \pi_{A_k}(a_k \mid \bar{y}_k, \bar{c}_k, \bar{x}_k, \bar{a}_{k-1}) \right) \\ & \times \mu_{Y_{K+1}}(y_{K+1} \mid \bar{y}_K, \bar{c}_K, \bar{x}_K, \bar{a}_K) \end{split}$$

Without going into too many details, note for example that:

- $\mu_{Y_k}(1 \mid \bar{0}_{k-1}, \bar{0}_{k-1}, \bar{x}_{k-1}, \bar{a}_{k-1})$  is the risk of outcome (dying) for a subject who did not yet die nor were right-censored plus had covariate and treatment history equal to  $\bar{x}_{k-1}, \bar{a}_{k-1}$ .
- $\pi_{C_k}(1 | \bar{0}_k, \bar{0}_{k-1}, \bar{x}_{k-1}, \bar{a}_{k-1})$  is the probability of being right-censored for a subject who did not yet die nor were right-censored plus had covariate and treatment history equal to  $\bar{x}_{k-1}, \bar{a}_{k-1}$ .

Identification of the target parameter as without censoring:

$$\begin{split} \bar{Q}_{K}(\bar{x}_{K-1},\bar{a}_{K-1},\bar{y}_{K-1}) &= \mathbb{E}_{P}\big[\,Q_{K+1}(\bar{X}_{K},a_{K}^{*},\bar{A}_{K-1},\bar{Y}_{K-1}) \mid \bar{Y}_{K-1} = \bar{y}_{K-1},\bar{C}_{K-1} = 0,\bar{Z}_{K-1},\bar{$$

 $\bar{Q}_{K+1}(\bar{x}_K, \bar{a}_K, \bar{v}_K) = \mathbb{E}_P[Y_{K+1} \mid \bar{Y}_K = \bar{v}_K, \bar{C}_K = 0, \bar{X}_K = \bar{x}_K, \bar{A}_K = \bar{a}_K]$ 

+ note that 
$$\bar{Q}_k(\bar{x}_{k-1}, \bar{a}_{k-1}, \bar{y}_{k-1}) = 1$$
 if  $y_{k'-1} = 1$  for  $k' \le k$ .

 $\bar{Q}_1(x_0, a_0) = \mathbb{E}_P[Q_2(\bar{X}_1, a_1^*) \mid X_0 = x_0, A_0 = a_0]$ 

The efficient influence function is given by:

$$\begin{split} \tilde{\phi}^*(\bar{Q},\pi)(O) \\ &= \sum_{k=1}^{K+1} 1\{Y_k = 0\} \left( \prod_{l=0}^{k-1} \frac{1\{A_l = a_l^*\} 1\{C_l = 0\}}{\pi_{A_l}(a_l^* \mid \bar{0}_l, \bar{N}_l, \bar{a}_{l-1}^*) \pi_{C_l}(0 \mid \bar{0}_{l-1}, \bar{0}_{l-1}, \bar{x}_{l-1}, \bar{a}_{l-1}^*)} \right) \\ &\times \left\{ \bar{Q}_{k+1}(\bar{X}_k, a_k^*, \bar{A}_{k-1}, 0) - \bar{Q}_k(\bar{X}_{k-1}, \bar{A}_{k-1}, 0) \right\} + \bar{Q}_0(X_0) - \Psi(P). \end{split}$$

- Right-censoring nodes are specified in the Cnodes argument.
- The formatting of Cnodes is a bit peculiar it should be a factor variable with the values 0 and 1 and the labels "uncensored" and "censored".
- Note that we further specify survivalOutcome=TRUE, so that Ynodes are treated as indicators of a terminating event.

- Qform: character vector of regression formulas for the outcome regressions
  - Qform indicates what variables are included in each outcome regression
  - default is NULL which means that all variables from previous time-points are included
  - (does not mean that GLM is used)
- gform: character vector of regression formulas for the propensity scores and the hazards of censoring
  - gform indicates what variables are included in each propensity score regression and the hazards of censoring
  - default is NULL which means that all variables from previous time-points are included
  - (does not mean that GLM is used)

For the simulated example:

```
XO.2 XO.3 AO Y1 C1 Y2 C2 Y3 C3 Y4 C4 Y5 C5 Y6 C6 Y7 C7 Y8
 iд
1: 1 0.408 -0.196
              0 1 0 1
                     0 1
                         0 1
                            0 1
2:
  2 -1.220 0.595 1 1 0 0 0 0
                           0 0 0 0 0
0 1 0 0 0 0 0 0 0 0 0 0 0 0 0
4: 4 0.604 0.041
5: 5 -0.532 -1.251
              0 0 0 0 0 0 0 0 0 0 0 0 0 0
6: 6 1.955 0.133
```

```
for (k in 1:7)
    sim.data[, (paste0("C", k)):=BinaryToCensoring(is.censored=get(
    paste0("C", k)))]
```

#### Treatment Estimate:

Parameter Estimate: 0.40391
Estimated Std Err: 0.031247
p-value: <2e-16
95% Conf Interval: (0.34267, 0.46515)

#### Control Estimate:

Parameter Estimate: 0.49545 Estimated Std Err: 0.029052 p-value: <2e-16 95% Conf Interval: (0.43851, 0.55239)

#### Additive Treatment Effect:

Parameter Estimate: -0.091539
Estimated Std Err: 0.041653
p-value: 0.027973
95% Conf Interval: (-0.17318, -0.0099008)

#### Relative Risk:

Parameter Estimate: 0.81524
Est Std Err log(RR): 0.095263
p-value: 0.032009

```
ltmle.fit$fit$g[[1]]
```

\$AO

Risk Coef SL.glm\_All 0.2502361 0.4568067 SL.mean\_All 0.2501919 0.5431933

\$C1

Risk Coef SL.glm\_All 0.1533290 0.98494501 SL.mean\_All 0.1667804 0.01505499

\$C2

Risk Coef SL.glm\_All 0.1365783 0.956276 SL.mean\_All 0.1428211 0.043724

\$C3

Risk Coef SL.glm\_All 0.1373765 0.92097765 SL.mean\_All 0.1440681 0.07902235

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# Practical 2: Application of 1tmle (Task 2ff)

In this part of the practical, we consider application of ltmle. There are three different 'topics' as follows:

- 1. Static and dynamic interventions (continuing Task 1)
- 2. Super learning
- 3. Right-censored data structures

Proceed from Task 2 of day3\_practical2.pdf.