

# Day 1, Lecture 3

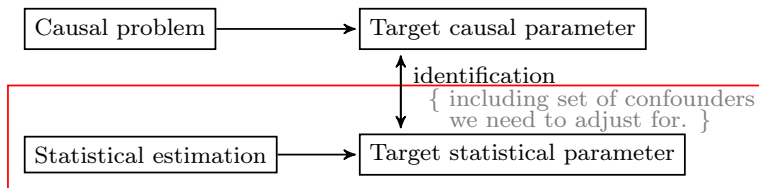
## Estimating the target

# Estimating the target

In this lecture, our goal is to:

1. Describe simple causal inference estimation methods for the average treatment effect, including the estimation equations estimator, an inverse probability weighted estimator and the g-formula based estimator.
2. Compare the assumptions and performance of “double robust” estimation methods to estimation methods based on inverse probability weighting and (g-formula based) standardization.

## Estimating the target



- ▶ one estimator is not more causal than another.
- ▶ different estimators are based on different nuisance parameters and have different statistical properties (bias/variance).

# G-formula versus IP-weighting

- G-formula
1. Estimate nuisance parameters  
 $f(a, x) = \mathbb{E}[Y \mid A = a, X = x]$  and the average over the marginal distribution  $\mu_X$  of  $X$
  2. Plug in to estimate the ATE:

$$\hat{\psi}_n^{\text{g-formula}} = \tilde{\Psi}(\hat{f}_n, \hat{\mu}_X) = \int_{\mathbb{R}^d} (\hat{f}_n(1, x) - \hat{f}_n(0, x)) d\hat{\mu}_X(x)$$

- IP-weighting
1. Estimate nuisance parameters  
 $\pi(a \mid x) = P(A = a \mid X = x)$  and the average over the distribution  $P$  of  $O$
  2. Plug in to estimate the ATE:

$$\hat{\psi}_n^{\text{ipw}} = \tilde{\Psi}_{\text{ipw}}(\hat{\pi}_n, \hat{P}_n) = \int_{\mathbb{R}^d} \sum_{a=0,1} \sum_{y=0,1} \left( \frac{ay}{\hat{\pi}_n(a \mid x)} - \frac{(1-a)y}{\hat{\pi}_n(a \mid x)} \right) d\hat{P}_n(x)$$

# Estimating equation (EE) estimator

EE estimator    1. Estimate nuisance parameters

$f(a, x) = \mathbb{E}[Y \mid A = a, X = x]$ ,  $\pi(a \mid x) = \mathbb{E}[A \mid X = x]$   
and the average over the distribution  $P$  of  $O$

2. Plug in to estimate the ATE:

$$\hat{\psi}_n^{\text{ee}} = \tilde{\Psi}_{\text{ee}}(\hat{f}_n, \hat{\pi}_n, \hat{P}_n) = \int_{\mathbb{R}^d} \sum_{a=0,1} \sum_{y=0,1} \left\{ \left( \frac{a}{\hat{\pi}_n(a \mid x)} - \frac{1-a}{\hat{\pi}_n(a \mid x)} \right) (y - \hat{f}_n(a, x)) \right. \\ \left. + \hat{f}_n(1, x) - \hat{f}_n(0, x) \right\} d\hat{P}_n(o)$$

## G-formula versus IP-weighting versus EE/TMLE

Estimation of the averages over  $\mu_X$  and  $P$  is straightforward using the empirical average over the observed data.

This yields:

G-formula estimator: 
$$\hat{\psi}_n^{\text{g-formula}} = \frac{1}{n} \sum_{i=1}^n \{ \hat{f}_n(1, X_i) - \hat{f}_n(0, X_i) \}$$

IP-weighted estimator: 
$$\hat{\psi}_n^{\text{ipw}} = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{A_i Y_i}{\hat{\pi}_n(A_i | X_i)} - \frac{(1 - A_i) Y_i}{\hat{\pi}_n(A_i | X_i)} \right\}$$

EE estimator: 
$$\hat{\psi}_n^{\text{ee}} = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{A_i}{\hat{\pi}_n(A_i | X_i)} - \frac{(1 - A_i)}{\hat{\pi}_n(A_i | X_i)} (Y_i - \hat{f}_n(A_i, X_i)) + \hat{f}_n(1, X_i) - \hat{f}_n(0, X_i) \right\}.$$

# G-formula versus IP-weighting versus EE/TMLE

G-formula estimator requires estimator  $\hat{f}_n$  for conditional expectation  $f$ .

- ▶ consistent if  $\hat{f}_n$  is consistent.

IP-weighted estimator requires estimator  $\hat{\pi}_n$  for the propensity score  $\pi$ .

- ▶ consistent if  $\hat{\pi}_n$  is consistent.

---

<sup>1</sup>which we get back to.

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**EE estimator** requires estimators  $\hat{f}_n$  and  $\hat{\pi}_n$  for conditional expectation  $f$  and propensity score  $\pi$ .

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**EE estimator** requires estimators  $\hat{f}_n$  and  $\hat{\pi}_n$  for conditional expectation  $f$  and propensity score  $\pi$ .

- ▶ consistent if either  $\hat{f}_n$  or  $\hat{\pi}_n$  is consistent (commonly known as "double robustness").

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# G-formula versus IP-weighting versus EE/TMLE

**G-formula estimator** requires estimator  $\hat{f}_n$  for conditional expectation  $f$ .

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- ▶ consistent if either  $\hat{f}_n$  or  $\hat{\pi}_n$  is consistent (commonly known as "double robustness").
- ▶ the EE estimator and the TMLE estimator share the same large-sample properties,<sup>1</sup> and particularly this property.

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<sup>1</sup>which we get back to.

# G-formula versus IP-weighting versus EE/TMLE

"Double robustness" —

## **SMALL EXERCISE:**

By the law of large numbers, the EE estimator converges in probability to:

$$\mathbb{E}_{P_0} \left[ \left( \frac{A}{\pi(A|X)} - \frac{1-A}{\pi(A|X)} \right) (Y - f(A, X)) + f(1, X) - f(0, X) \right] \quad (1)$$

where  $(f, \pi)$  denotes the limit of  $(\hat{f}_n, \hat{\pi}_n)$ . Compute (1) when:

1.  $f = f_0$  (i.e., the outcome regression is consistently estimated), and
2.  $\pi = \pi_0$  (i.e., the propensity score is consistently estimated).

## G-formula versus IP-weighting versus EE/TMLE

Can't we just construct a good g-formula estimator???

# G-formula versus IP-weighting versus EE/TMLE

Can't we just construct a good g-formula estimator???

- ▶ a logistic regression — great if correctly specified, but horrible if not.
- ▶ a random forest — properly tuned?

## A random forest — properly tuned?

Predictive performance of an estimator can be measured in terms of some distance<sup>2</sup> between:

- 1) the observed outcome:  $Y_i$
- 2) and the predicted conditional expectation:  $\hat{f}_n(A_i, X_i)$

---

<sup>2</sup>Measured in terms of a *loss function*.

## A random forest — properly tuned?

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- 1) the observed outcome:  $Y_i$
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One example of a loss function  $\mathcal{L}(f)(O)$  is the negative log-likelihood loss:

$$\mathcal{L}(\hat{f}_n)(Y_i, A_i, X_i) = -(Y_i \log(\hat{f}_n(A_i, X_i)) + (1 - Y_i) \log(1 - \hat{f}_n(A_i, X_i))).$$

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The estimator  $\hat{f}_n$  closest to the true  $f_0$  minimizes the risk:

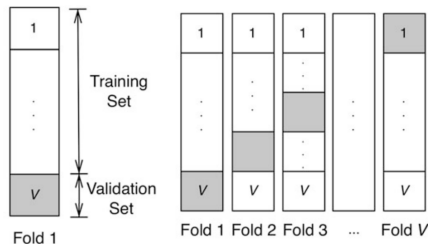
$$\mathbb{E}_{P_0}[\mathcal{L}(\hat{f}_n)(Y_i, A_i, X_i)].$$

---

<sup>2</sup>Measured in terms of a *loss function*.



# A random forest — properly tuned?



The risk can be estimated in a cross-validation scheme.<sup>a</sup>

I.e., for each sample split:

1. Each model is created and fitted on the training data:  $\hat{f}_n^{\text{train}}$ .
2. The quality of the model is checked on the validation data
  - ▶ Average of  $\mathcal{L}(\hat{f}_n^{\text{train}})(O_i)$  in the validation sample.

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<sup>a</sup>To measure performance on independent data.

# A random forest — properly tuned?

## Simulated example

- ▶  $X \sim \text{Unif}(-2, 2)$
- ▶  $X_1^{\text{noise}}, \dots, X_5^{\text{noise}} \sim N(0, 1)$
- ▶  $A \in \{0, 1\}$  with distribution given  $X$  given by:

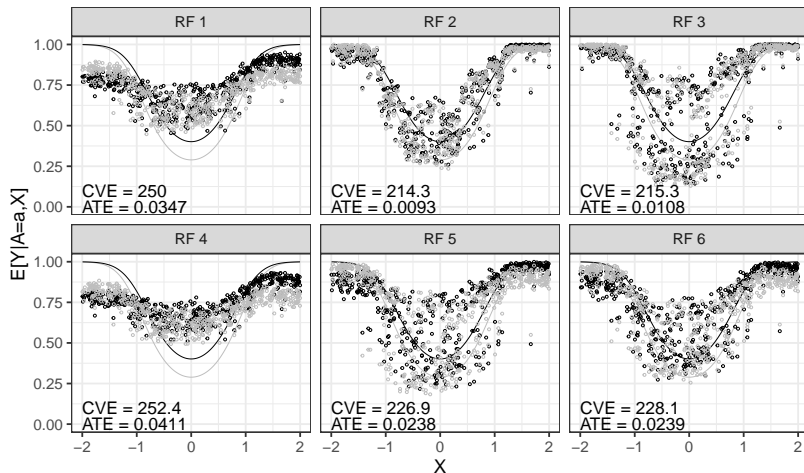
$$\text{logit } \mathbb{E}[A \mid X] = \gamma_0 + \gamma_X^\top X$$

- ▶  $Y \in \{0, 1\}$  with distribution given  $X$  and  $A$  given by:

$$\text{logit } \mathbb{E}[Y \mid A, X] = \beta_0 + \beta_A A + \beta_X^\top X^2$$

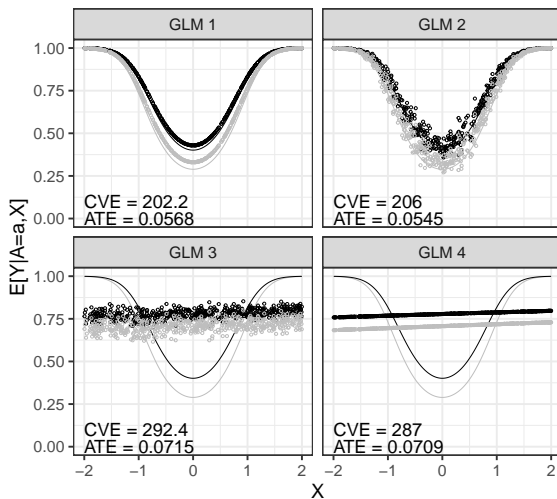
# A random forest — properly tuned?

RF fitted with different values of tuning parameters (nodesize, mtry):



# Different GLM models

GLM models fitted with different covariates and functional form of covariates:



## A random forest — properly tuned?

This is all about constructing a good estimator for the conditional expectation  $f$ .

This does not necessarily translate into a good estimator for the target  $\Psi(P)$ .

## A random forest — properly tuned?

This is all about constructing a good estimator for the conditional expectation  $f$ .

This does not necessarily translate into a good estimator for the target  $\Psi(P)$ .

TMLE is all about constructing a g-formula estimator which is a good estimator for *the target*.

# Simulating simple data structure in R

Fix randomness:

```
set.seed(5)
```

Fix a sample size:

```
n <- 500
```

Generate covariate  $X \in [-2, 2]$ :

```
X <- runif(n, -2, 2)
```

Generate binary treatment decision  $A$ :

```
A <- rbinom(n, 1, prob=plogis(-0.25 + 1.2*X))
```

(corresponding to logit  $\mathbb{E}[A | X] = \gamma_0 + \gamma_X X$ )

## Simulating simple data structure in R

Generate binary outcome  $Y$  according to

$$\text{logit } \mathbb{E}[Y \mid A, X] = \beta_0 + \beta_A A + \beta_X X^2$$



## Simulating simple data structure in R

Generate binary outcome  $Y$  according to

$$\text{logit } \mathbb{E}[Y \mid A, X] = \beta_0 + \beta_A A + \beta_X X^2$$

First generate counterfactuals:

```
Y1 <- rbinom(n, 1, prob=plogis(-0.9 + 1.9*X^2 + 0.5*1))  
Y0 <- rbinom(n, 1, prob=plogis(-0.9 + 1.9*X^2 + 0.5*0))
```

## Simulating simple data structure in R

Generate binary outcome  $Y$  according to

$$\text{logit } \mathbb{E}[Y \mid A, X] = \beta_0 + \beta_A A + \beta_X X^2$$

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```

We only observe the counterfactual outcome corresponding to the observed treatment level:

```
Y <- A*Y1 + (1-A)*Y0
```

# Simulating simple data structure in R

Observed data:

```
      X A Y
1: -1.1991422 0 0
2:  0.7408744 1 0
3:  1.6675031 1 1
4: -0.8624022 0 1
5: -1.5813995 0 1
---
496: -0.3978523 1 0
497: -1.5069379 0 1
498:  1.8340120 1 1
499:  0.6349484 1 1
500: -0.5214807 0 1
```

## Simulating simple data structure in R

Observed data:

	X	A	Y
1:	-1.1991422	0	0
2:	0.7408744	1	0
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---			
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Counterfactual data:

	X	Y1	Y0
1:	-1.1991422	0	1
2:	0.7408744	1	0
3:	1.6675031	1	1
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---			
496:	-0.3978523	0	1
497:	-1.5069379	0	1
498:	1.8340120	1	1
499:	0.6349484	0	0
500:	-0.5214807	0	0

## Simulating simple data structure in R

Simulating many observations of counterfactuals allows us to approximate the **true ATE**:

```
X <- runif(1e6, -2, 2)
Y1 <- rbinom(1e6, 1, prob=plogis(-0.9 + 1.9*X^2 + 0.5*1))
Y0 <- rbinom(1e6, 1, prob=plogis(-0.9 + 1.9*X^2 + 0.5*0))
```

The **true ATE** is then approximately:

```
(true.ate <- mean(Y1 - Y0))
```

```
[1] 0.070292
```

since  $ATE = \mathbb{E}_{P_0}[Y^1] - \mathbb{E}_{P_0}[Y^0]$ .

# Simulating simple data structure in R

Fit correctly specified parametric model:

```
fit.glm <- glm(Y~A+X.squared, data=dt[, X.squared:=X^2],  
              family=binomial)
```

Use model to estimate  $f(1, X)$  for all subjects:

```
dt[, pred.glm.A1:=predict(fit.glm, type="response", newdata=  
  copy(dt)[, A:=1])]
```

And similarly  $f(0, X)$  for all subjects:

```
dt[, pred.glm.A0:=predict(fit.glm, type="response", newdata=  
  copy(dt)[, A:=0])]
```

Then we can estimate the ATE by:

```
(fit.glm <- dt[, mean(pred.glm.A1-pred.glm.A0)])
```

```
[1] 0.04322891
```

## Simulating simple data structure in R

Using a random forest (no tuning):

```
library(randomForestSRC)
fit.rf <- rfsrc(Y~A+X, data=dt)
dt[, pred.rf.A1:=predict(fit.rf, type="response", newdata=
  copy(dt)[, A:=1])$predicted]
dt[, pred.rf.A0:=predict(fit.rf, type="response", newdata=
  copy(dt)[, A:=0])$predicted]
(fit.rf <- dt[, mean(pred.rf.A1-pred.rf.A0)])
```

```
[1] 0.07005063
```

# Simulating simple data structure in R

Using a misspecified parametric model:

```
fit.glm.mis <- glm(Y~A+X, data=dt, family=binomial)
dt[, pred.glm.mis.A1:=predict(fit.glm.mis, type="response",
                             newdata=copy(dt)[, A:=1])]
dt[, pred.glm.mis.A0:=predict(fit.glm.mis, type="response",
                             newdata=copy(dt)[, A:=0])]
(fit.glm.mis <- dt[, mean(pred.glm.mis.A1-pred.glm.mis.A0)])
```

[1] 0.09127889



# Simulating simple data structure in R

We can investigate the properties of different estimators —

- ▶ We know the true value of ATE:  $\psi_0 \approx 0.0702$
- ▶ We have generated the outcome  $Y$  according to

$$\text{logit } \mathbb{E}[Y \mid A, X] = \beta_0 + \beta_A A + \beta_X X^2$$

- ▶ We have generated the treatment  $A$  according to

$$\text{logit } \mathbb{E}[A \mid X] = \gamma_0 + \gamma_X X$$

If we repeat the experiment of drawing  $n$  observations we would every time end up with a different realization of the particular estimator.

## Different estimators

**G-formula estimator** Using an estimator  $\hat{f}_n$  for  $f(a, X) = \mathbb{E}[Y \mid A = a, X]$ , estimate the ATE by:

$$\hat{\psi}_n^{\text{g-formula}} = \frac{1}{n} \sum_{i=1}^n \{ \hat{f}_n(1, X_i) - \hat{f}_n(0, X_i) \}$$

**Inverse probability weighted estimator** Using an estimator  $\hat{\pi}_n$  for  $\pi(a \mid X) = P(A = a \mid X)$ , estimate the ATE by:

$$\hat{\psi}_n^{\text{ipw}} = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{A_i Y_i}{\hat{\pi}_n(A_i \mid X_i)} - \frac{(1 - A_i) Y_i}{\hat{\pi}_n(A_i \mid X_i)} \right\}$$

## Different estimators

**EE estimator** Using an estimator  $\hat{f}_n$  for  $f(a, X) = \mathbb{E}[Y \mid A = a, X]$  and an estimator  $\hat{\pi}_n$  for  $\pi(a \mid X) = P(A = a \mid X)$ , estimate the ATE by:

$$\hat{\psi}_n^{\text{ee}} = \frac{1}{n} \sum_{i=1}^n \left\{ \frac{A_i}{\hat{\pi}_n(A_i \mid X_i)} - \frac{(1 - A_i)}{\hat{\pi}_n(A_i \mid X_i)} (Y_i - \hat{f}_n(A_i, X_i)) + \hat{f}_n(1, X_i) - \hat{f}_n(0, X_i) \right\}$$

**TMLE estimator** Update the estimator  $\hat{f}_n \mapsto \hat{f}_n^*$  in a "targeted way" using the information from the estimator  $\hat{\pi}_n$ , then estimate the ATE by:

$$\hat{\psi}_n^{\text{tmle}} = \frac{1}{n} \sum_{i=1}^n \{ \hat{f}_n^*(1, X_i) - \hat{f}_n^*(0, X_i) \}$$

## Different estimators — `tmle` implementation

Today we will just (more or less blindly) use software to use TMLE.

## Different estimators — tmle implementation

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```
library(tmle)
```

```
tmle(Y, A, X,  
      gform,  
      Qform,  
      SL.library,  
      family="binomial",  
      cvQinit=FALSE,  
      ...  
)
```

- ▶  $Y \in \mathbb{R}$  or  $Y \in \{0, 1\}$
- ▶  $A \in \{0, 1\}$
- ▶  $X$  a vector, matrix or a data frame

## Different estimators — `tmle` implementation

- ▶ `gform`
  - ▶ optional regression formula for the propensity score  $\pi$
  - ▶ on the form  $A \sim X_1 + X_2$
  - ▶ (overrides call to `SuperLearner`)
- ▶ `Qform`
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- ▶ `cvQinit=FALSE`
  - ▶ default is `TRUE` which means cross-validated predicted values are estimated
- ▶ `gbound`
  - ▶ truncation of predicted probabilities of treatment



# Different estimators — `tmle` implementation

## On a sidenote — tomorrow

- ▶ `Q.SL.library`
  - ▶ optional vector of prediction algorithms to use for SuperLearner in initial estimation of  $f$
- ▶ `g.SL.library`
  - ▶ optional vector of prediction algorithms to use for SuperLearner in initial estimation of  $\pi$
- ▶ `Q.discreteSL`
  - ▶ if TRUE, a discrete super learner is used (rather than ensemble)
  - ▶ default is FALSE
- ▶ `g.discreteSL`
  - ▶ if TRUE, a discrete super learner is used (rather than ensemble)
  - ▶ default is FALSE

**Note:** The discrete super learner simply picks an algorithm from its library by minimizing the cross-validated empirical risk with respect a loss function.

## Different estimators — tmle implementation

What were the estimated IP weights?

```
summary(fit.tmle$g$g1W)
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.04751	0.19441	0.49405	0.49400	0.79710	0.94109

Note that weights close to 0 or to 1 would indicate positivity issues.

## Different estimators — tmle implementation

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```

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Note that weights close to 0 or to 1 would indicate positivity issues.

What truncation level was used?

```
fit.tmle$gbound
```

```
[1] 0.03598084 1.00000000
```

I.e., no weights were truncated.

# Practical 1: Explorations based on simulated data

As part of the exercise we will explore —

1. Comparing g-formula estimators for different estimators for  $f$ ; either different logistic regressions or different machine learning algorithms.
2. Properties of the g-formula estimator and the IP-weighted estimator, compared to the TMLE estimator.
3. Double robustness: Misspecification of the outcome regression ( $f$ ).

The exercise is described in detail in: **day1-practical1.pdf**.