Day 2, Lecture 1

The foundation of TMLE: The constructive proof of TMLE

Day 2: 8 - 9

Introduction to TMLE. The constructive proof of TMLE.

The decomposition and the role of Step 1 & Step 2.

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The targeting step. Solving the efficient influence curve equation.

▷ Implementation of the targeting step.

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Lunch.

Day 2: 13 - 15

Step 2: Super learning. (Thomas).

- 1. Data is a random variable O with a probability P_0
- 2. P_0 belongs to a statistical model $\mathcal M$
- 3. Our target is a parameter $\Psi: \mathcal{M} \to \mathbb{R}$
- 4. Construct estimator \hat{P}_n for (relevant part of) P_0 and estimate the target parameter by $\hat{\psi}_n = \Psi(\hat{P}_n)$
- 5. Quantify uncertainty for the estimator $\hat{\psi}_n = \Psi(\hat{P}_n)$

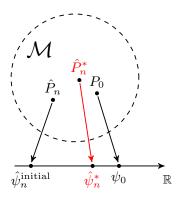
Estimation paradigm

- 1. Minimal amount of parametric assumptions on P_0 (goal)
- 2. Go after optimal/efficient estimation of $\psi_0 = \Psi(P_0)$

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Estimation paradigm

- 1. Minimal amount of parametric assumptions on P_0 (goal)
- 2. Go after optimal/efficient estimation of $\psi_0 = \Psi(P_0) \Rightarrow$ based the efficient influence function (canonical gradient)



Tools from semiparametric efficiency theory and empirical process theory tell us how to construct an optimal estimator for a given target parameter $\Psi:\mathcal{M}\to\mathbb{R}$

- asymptotic linearity/normality
- asymptotic efficiency

Recap notation:¹

 \triangleright For a function $h: \mathcal{O} \to \mathbb{R}$ and distribution P

$$Ph = \mathbb{E}_{P}[h(O)] = \int hdP = \int_{\mathcal{O}} h(o)dP(o)$$

where $\mathcal{O} = \mathbb{R}^d \times \{0,1\} \times \{0,1\}$ is the sample space of $\mathcal{O} = (X,A,Y)$.

 \triangleright For the empirical measure \mathbb{P}_n of the sample O_1, \ldots, O_n :

$$\mathbb{P}_n h = \int h d\mathbb{P}_n = \frac{1}{n} \sum_{i=1}^n h(O_i);$$

note: the right-hand-side is really just the empirical average.

 $\triangleright X_n = o_P(1)$ means that $X_n \stackrel{P}{\to} 0$; $X_n = o_P(n^{-1/2})$ means that $n^{1/2}X_n \stackrel{P}{\to} 0$.

¹van der Vaart, A. W. (2000). Asymptotic statistics (Vol. 3). Cambridge university press.

Asymptotic linearity

An estimator $\hat{\psi}_n$ is asymptotically linear if,

$$\sqrt{n}(\hat{\psi}_n - \Psi(P_0)) = \sqrt{n} \mathbb{P}_n \phi(P_0) + o_P(1) \tag{*}$$

with $P_0\phi(P_0)=0$. We say that $\hat{\psi}_n$ has influence function $O\mapsto\phi(P_0)(O)$.

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Use CLT on $(*) \Rightarrow$ asymptotic normality:

$$\hat{\psi}_n \stackrel{as}{\sim} N(\Psi(P_0), \sigma_0^2/n),$$

where $\sigma_0^2 = P_0 \phi(P_0)^2$.

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Efficient influence function

There exists $\phi^*(P_0)$ such that $P_0\phi(P_0)^2 \ge P_0\phi^*(P_0)^2$.

 $\phi^*(P_0)$ is called the efficient influence function.

 \Rightarrow estimator with the smallest possible asymptotic variance

EXAMPLE: Average treatment effect (ATE)

Observed data
$$O = (X, A, Y) \in \mathbb{R}^d \times \{0, 1\} \times \{0, 1\} = \mathcal{O}$$

- * $X \in \mathbb{R}^d$ are covariates
- * $A \in \{0,1\}$ is a binary exposure variable (treatment decision)
- * $Y \in \{0,1\}$ is a binary outcome variable

 $O \sim P_0$ where P_0 assumed to belong to nonparametric model \mathcal{M} .

We are interested in estimating the ATE:

$$\Psi(P) = \mathbb{E}_P \big[\mathbb{E}_P \big[Y \mid A = 1, X \big] - \mathbb{E}_P \big[Y \mid A = 0, X \big] \big].$$

For the ATE, as we have seen, we can also write the target parameter $\Psi:\mathcal{M}\to\mathbb{R}$ as

$$\Psi(P) = \tilde{\Psi}(f, \mu_X) = \int_{\mathbb{R}^d} (f(1, x) - f(0, x)) d\mu_X(x) \qquad (*)$$

where

$$f(a,x) = \mathbb{E}[Y \mid A = a, X = x]$$

and μ_X is the marginal distribution of X.

I.e.,
$$\hat{\psi}_n = \tilde{\Psi}(\hat{f}_n, \hat{\mu}_n)$$
.

An estimator $\hat{\psi}_n$ is asymptotically linear if,

$$\sqrt{n}(\hat{\psi}_n - \Psi(P_0)) = \sqrt{n} \mathbb{P}_n \phi^*(P_0) + o_P(1). \quad (*)$$

For $P \in \mathcal{M}$, define: $R(P, P_0) := \Psi(P) - \Psi(P_0) + P_0 \phi^*(P)$.

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$$\Psi(\hat{P}_{n}) - \Psi(P_{0}) = -P_{0}\phi^{*}(\hat{P}_{n}) + R(\hat{P}_{n}, P_{0})$$

$$= -P_{0}\phi^{*}(\hat{P}_{n}) + R(\hat{P}_{n}, P_{0})$$

$$+ \mathbb{P}_{n}\phi^{*}(\hat{P}_{n}) - \mathbb{P}_{n}\phi^{*}(\hat{P}_{n})$$

$$+ (\mathbb{P}_{n} - P_{0})\phi^{*}(P_{0}) - (\mathbb{P}_{n} - P_{0})\phi^{*}(P_{0})$$

- (1)
- (2)
- (3)

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$$= (\mathbb{P}_{n} - P_{0})\phi^{*}(P_{0})$$

$$+ (\mathbb{P}_{n} - P_{0})(\phi^{*}(\hat{P}_{n}) - \phi^{*}(P_{0}))$$

$$+ R(\hat{P}_{n}, P_{0})$$

$$- \mathbb{P}_{n}\phi^{*}(\hat{P}_{n})$$
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For $P \in \mathcal{M}$, define: $R(P, P_0) := \Psi(P) - \Psi(P_0) + P_0 \phi^*(P)$.

We decompose as follows:

$$\Psi(\hat{P}_{n}) - \Psi(P_{0}) = -P_{0}\phi^{*}(\hat{P}_{n}) + R(\hat{P}_{n}, P_{0})$$

$$= -P_{0}\phi^{*}(\hat{P}_{n}) + R(\hat{P}_{n}, P_{0})$$

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$$+ R(\hat{P}_{n}, P_{0})$$

$$- \mathbb{P}_{n}\phi^{*}(\hat{P}_{n})$$
(2)

i.e., want (1)–(3) to be $o_P(n^{-1/2})$.

An estimator
$$\hat{\psi}_n$$
 is asymptotically linear if,
$$\sqrt{n} (\hat{\psi}_n - \Psi(P_0)) = \sqrt{n} \, \mathbb{P}_n \phi^*(P_0) + o_P(1). \quad (*)$$

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(1)
(2)

- ▶ (1) handled by empirical process theory if Donsker.²
- (2) depends on the target parameter.
- ▶ (3) is called the efficient influence curve equation.

²Lemma 19.24 of van der Vaart, A. W. (2000): Asymptotic statistics yields then that $(\mathbb{P}_n - P_0)(\phi^*(\hat{P}_n) - \phi^*(P_0)) = o_P(n^{-1/2})$.

That is it.

That is it.

Conditions (asymptotic linearity and efficiency)

- (C1) Solve the efficient influence curve equation: $\mathbb{P}_n \phi^*(\hat{P}_n) = o_P(n^{-1/2})$
- (C2) Remainder $R(\hat{P}_n, P_0) = o_P(n^{-1/2})$
- (C3) Donsker class conditions for $\{\phi^*(P): P \in \mathcal{M}\}$

Then: $\Psi(\hat{P}_n) \stackrel{as}{\sim} N(\Psi(P_0), P_0\phi^*(P_0)^2/n)$

Side note: Usually, we will assume the Donsker class condition (C3)

- this is a way of nonparametrically characterizing the complexity of nuisance parameters.
- classes of functions that are Donsker: Indicator functions, bounded monotone functions, Lipschitz parametric functions, smooth functions, . . .

Donsker classes also include traditional parametric functions.

We will not discuss this further. For a nice intro see Sections 4.2 and 4.3 of Kennedy, E. H. (2016): Semiparametric theory and empirical processes in causal inference.

$$\Psi(\hat{P}_{n}) - \Psi(P_{0}) = \mathbb{P}_{n}\phi^{*}(P_{0}) + o_{P}(n^{-1/2})$$
$$+ R(\hat{P}_{n}, P_{0})$$
$$- \mathbb{P}_{n}\phi^{*}(\hat{P}_{n})$$

For a given target parameter $\Psi : \mathcal{M} \to \mathbb{R}$, we need to

- 1. Know the efficient influence curve, so that we can solve the efficient influence curve equation.
- 2. Analyze the remainder $R(P, P_0) \coloneqq \Psi(P) \Psi(P_0) + P_0 \phi^*(P)$.

EXAMPLE: Average treatment effect (ATE)

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1. The efficient influence function:

$$\phi^{*}(P)(O) = \tilde{\phi}^{*}(f,\pi)(O)$$

$$= \left(\frac{A}{\pi(A \mid X)} - \frac{1 - A}{\pi(A \mid X)}\right) (Y - f(A,X)) + f(1,X) - f(0,X) - \Psi(P)$$

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2. The remainder:

$$\begin{split} R(P,P_0) &= \tilde{R}(f,\pi,f_0,\pi_0) \\ &= \int_{\mathbb{R}^d} \sum_{a=0,1} (2a-1) \frac{\pi_0(a\mid x) - \pi(a\mid x)}{\pi(a\mid x)} \big(f_0(a,x) - f(a,x) \big) d\mu_{0,X}(x) \end{split}$$

$$f(A,X) = \mathbb{E}_{P}[Y \mid A,X], \ \pi(A \mid X) = P(A = a \mid X) \\ f_{0}(A,X) = \mathbb{E}_{P_{0}}[Y \mid A,X], \ \pi_{0}(A \mid X) = P(A = a \mid X)$$

$$R(P, P_0) := \Psi(P) - \Psi(P_0) + P_0 \phi^*(P).$$

Deriving the remainder for the ATE:

$$R(P, P_{0}) = \mathbb{E}_{P}[f(1, X) - f(0, X)] - \mathbb{E}_{P_{0}}[f_{0}(1, X) - f_{0}(0, X)]$$

$$+ \mathbb{E}_{P_{0}}\left[\left(\frac{A}{\pi(A \mid X)} - \frac{1 - A}{\pi(A \mid X)}\right)(Y - f(A, X))\right]$$

$$+ \mathbb{E}_{P_{0}}[f(1, X) - f(0, X)] - \Psi(P)$$

$$\stackrel{*}{=} \int_{\mathbb{R}^{d}} \sum_{a=0,1} (2a - 1)\left(\frac{\pi_{0}(a \mid x)}{\pi(a \mid x)} - 1\right)(f_{0}(a, x) - f(a, x))d\mu_{0, X}(x)$$

$$= \int_{\mathbb{R}^{d}} \sum_{a=0,1} (2a - 1)\frac{\pi_{0}(a \mid x) - \pi(a \mid x)}{\pi(a \mid x)}(f_{0}(a, x) - f(a, x))d\mu_{0, X}(x)$$

the equality marked by * is detailed on the next slide.

We used that:

$$\begin{split} \mathbb{E}_{P_{\mathbf{0}}} & \left[\left(\frac{A}{\pi(A \mid X)} - \frac{1 - A}{\pi(A \mid X)} \right) (Y - f(A, X)) \right] \\ & = \mathbb{E}_{P_{\mathbf{0}}} \left[\frac{2A - 1}{\pi(A \mid X)} (Y - f(A, X)) \right] \\ & = \mathbb{E}_{P_{\mathbf{0}}} \left[\mathbb{E}_{P_{\mathbf{0}}} \left[\frac{2A - 1}{\pi(A \mid X)} (Y - f(A, X)) \mid A, X \right] \right] \\ & = \mathbb{E}_{P_{\mathbf{0}}} \left[\frac{2A - 1}{\pi(A \mid X)} (f_{\mathbf{0}}(A, X) - f(A, X)) \right] \\ & = \int_{\mathbb{R}^{d}} \sum_{a = 0, 1} \frac{2a - 1}{\pi(a \mid x)} (f_{\mathbf{0}}(a, x) - f(a, x)) \pi_{\mathbf{0}}(a \mid x) d\mu_{\mathbf{0}, X}(x) \\ & = \int_{\mathbb{R}^{d}} \sum_{a = 0, 1} (2a - 1) \frac{\pi_{\mathbf{0}}(a \mid x)}{\pi(a \mid x)} (f_{\mathbf{0}}(a, x) - f(a, x)) d\mu_{\mathbf{0}, X}(x) \end{split}$$

The remainder determines the asymptotic bias.

For the ATE, the remainder has a really nice structure!

$$R(P, P_0) = \tilde{R}(f, \pi, f_0, \pi_0)$$

$$= \int_{\mathbb{R}^d} \sum_{a=0,1} (2a-1) \frac{\pi_0(a \mid x) - \pi(a \mid x)}{\pi(a \mid x)} (f_0(a, x) - f(a, x)) d\mu_{0,X}(x)$$

A "double robust" structure, which has some important implications.

$$|R(P, P_0)| = |\tilde{R}(f, \pi, f_0, \pi_0)|$$

$$\leq \sum_{a=0}^{\infty} \int_{\mathbb{R}^d} \frac{|\pi_0(a \mid x) - \pi(a \mid x)|}{\pi(a \mid x)} |f_0(a, x) - f(a, x)| d\mu_{0, X}(x)$$

$$\begin{split} |R(P,P_{0})| &= |\tilde{R}(f,\pi,f_{0},\pi_{0})| \\ &\leq \sum_{a=0,1} \int_{\mathbb{R}^{d}} \frac{|\pi_{0}(a\mid x) - \pi(a\mid x)|}{\pi(a\mid x)} |f_{0}(a,x) - f(a,x)| d\mu_{0,X}(x) \\ &\stackrel{*}{\leq} \sum_{a=0,1} \frac{1}{\pi(a\mid x)} \sqrt{\int_{\mathbb{R}^{d}} \left\{ \pi_{0}(a\mid x) - \pi(a\mid x) \right\}^{2} d\mu_{0,X}(x)} \\ &\times \sqrt{\int_{\mathbb{R}^{d}} \left\{ f_{0}(a,x) - f(a,x) \right\}^{2} d\mu_{0,X}(x)} \end{split}$$

$$\begin{split} |R(P,P_{0})| &= |\tilde{R}(f,\pi,f_{0},\pi_{0})| \\ &\leq \sum_{a=0,1} \int_{\mathbb{R}^{d}} \frac{|\pi_{0}(a\mid x) - \pi(a\mid x)|}{\pi(a\mid x)} |f_{0}(a,x) - f(a,x)| d\mu_{0,X}(x) \\ &\stackrel{*}{\leq} \sum_{a=0,1} \frac{1}{\pi(a\mid x)} \sqrt{\int_{\mathbb{R}^{d}} \left\{ \pi_{0}(a\mid x) - \pi(a\mid x) \right\}^{2} d\mu_{0,X}(x)} \\ &\times \sqrt{\int_{\mathbb{R}^{d}} \left\{ f_{0}(a,x) - f(a,x) \right\}^{2} d\mu_{0,X}(x)} \end{split}$$

Thus, if $\pi(a \mid X) > \delta > 0$ a.s., then:

$$|\tilde{R}(\hat{f}_{n}^{*}, \hat{\pi}_{n}, f_{0}, \pi_{0})| \leq \sum_{a=0,1} \delta^{-1} \|\pi_{0}(a \mid \cdot) - \hat{\pi}_{n}(a \mid \cdot)\|_{\mu_{0}} \|f_{0}(a \mid \cdot) - \hat{f}_{n}(a \mid \cdot)\|_{\mu_{0}}$$

What does this imply for estimation?

Double robustness in consistency

$$\begin{split} |\tilde{R}(\hat{f}_{n}^{*},\hat{\pi}_{n},f_{0},\pi_{0})| &\leq \sum_{a=0,1} \delta^{-1} \underbrace{\|\pi_{0}(a\mid\cdot) - \hat{\pi}_{n}(a\mid\cdot)\|_{\mu_{\mathbf{0}}}}_{o_{P}(1),\text{ or }} \underbrace{\|f_{0}(a\mid\cdot) - \hat{f}_{n}^{*}(a\mid\cdot)\|_{\mu_{\mathbf{0}}}}_{o_{P}(1)} \end{split}$$
 then $\tilde{\Psi}(\hat{f}_{n}^{*}) - \tilde{\Psi}(f_{0}) = o_{P}(1).$

Asymptotic linearity (easier to establish due to double robust structure)

$$\begin{split} |\tilde{R}(\hat{f}_n^*,\hat{\pi}_n,f_0,\pi_0)| &\leq \sum_{a=0,1} \delta^{-1} \underbrace{\|\pi_0(a\,|\,\cdot) - \hat{\pi}_n(a\,|\,\cdot)\|_{\mu_0}}_{=o_P(n^{-1/4})} \underbrace{\|f_0(a\,|\,\cdot) - \hat{f}_n^*(a\,|\,\cdot)\|_{\mu_0}}_{=o_P(n^{-1/4})} \end{split}$$
 i.e.,
$$\tilde{R}(\hat{f}_n^*,\hat{\pi}_n,f_0,\pi_0) = o_P(n^{-1/2}).$$

I.e., bias is converging at fast enough rate for reliable confidence intervals.

Side note: Showing the double robustness in consistency . . .

Say we have estimators $(\hat{f}_n^*, \hat{\pi}_n)$;

- converging to (f, π)
- solving the efficient influence curve equation.

Per definition, $\tilde{R}(\hat{f}_n^*, \hat{\pi}_n, f_0, \pi_0) = \tilde{\Psi}(\hat{f}_n^*) - \tilde{\Psi}(f_0) + P_0 \tilde{\phi}^*(\hat{f}_n^*, \hat{\pi}_n)$.

i.e.,
$$\tilde{\Psi}(\hat{f}_n^*) - \tilde{\Psi}(f_0) = -P_0 \tilde{\phi}^*(\hat{f}_n^*, \hat{\pi}_n) + \tilde{R}(\hat{f}_n^*, \hat{\pi}_n, f_0, \pi_0)$$
$$= (\mathbb{P}_n - P_0) \phi^*(\hat{f}_n^*, \hat{\pi}_n) + \tilde{R}(\hat{f}_n^*, \hat{\pi}_n, f_0, \pi_0)$$

The first term is an empirical process term which equals $(\mathbb{P}_n - P_0)\tilde{\phi}^*(f,\pi)$ plus an $o_P(n^{-1/2})$ -term.

This then gives

$$\tilde{\Psi}(\hat{f}_{n}^{*}) - \tilde{\Psi}(f_{0}) = \underbrace{(\mathbb{P}_{n} - P_{0})\tilde{\phi}^{*}(f, \pi)}_{\text{LLN applies}} + \tilde{R}(\hat{f}_{n}^{*}, \hat{\pi}_{n}, f_{0}, \pi_{0}) + o_{P}(n^{-1/2})$$

which yields that
$$\tilde{\Psi}(\hat{f}_n^*) - \tilde{\Psi}(f_0) = o_P(1)$$
 if $\tilde{R}(\hat{f}_n^*, \hat{\pi}_n, f_0, \pi_0) = o_P(1)$.

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$$\phi^*(P)(O) = \tilde{\phi}^*(f,\pi)(O) = \left(\frac{A}{\pi(A|X)} - \frac{1-A}{\pi(A|X)}\right) (Y - f(A,X)) + f(1,X) - f(0,X) - \Psi(P)$$

2. The remainder:

$$R(P, P_0) = \tilde{R}(f, \pi, f_0, \pi_0)$$

$$= \int_{\mathbb{R}^d} \sum_{a=0,1} (2a-1) \frac{\pi_0(a \mid x) - \pi(a \mid x)}{\pi(a \mid x)} (f_0(a, x) - f(a, x)) d\mu_{0,X}(x)$$

Deriving these is done once for a given target parameter $\Psi: \mathcal{M} \to \mathbb{R}$.

Conditions (asymptotic linearity and efficiency)

- (C1) Solve the efficient influence curve equation: $\mathbb{P}_n\phi^*(\hat{P}_n)=o_P(n^{-1/2})$
- (C2) Remainder $R(\hat{P}_n, P_0) = o_P(n^{-1/2})$
- (C3) Donsker class conditions for $\{\phi^*(P): P \in \mathcal{M}\}$

Then: $\Psi(\hat{P}_n) \stackrel{as}{\sim} \mathcal{N}(\Psi(P_0), P_0 \phi^*(P_0)^2/n)$

Conditions (asymptotic linearity and efficiency)

- (C1) Solve the efficient influence curve equation: $\mathbb{P}_n\phi^*(\hat{P}_n)=o_P(n^{-1/2})$
- (C2) Remainder $R(\hat{P}_n, P_0) = o_P(n^{-1/2})$
- (C3) Donsker class conditions for $\{\phi^*(P): P \in \mathcal{M}\}$

Then: $\Psi(\hat{P}_n) \stackrel{as}{\sim} N(\Psi(P_0), P_0\phi^*(P_0)^2/n)$

TMLE is a two-step procedure:

- Step 1 Construct initial estimator \hat{P}_n for P such that $R(\hat{P}_n, P_0) = o_P(n^{-1/2})$.
- Step 2 Update the estimator $\hat{P}_n \mapsto \hat{P}_n^*$ such that \hat{P}_n^* solves the efficient influence curve equation.

Introduction to TMLE: Summary so far

- ▶ The role of the targeting step (Step 2):
 - Gaining double robustness in consistency.
 - Easier to get rid of second-order remainder.
- ▶ The role of the initial estimation step (Step 1):
 - This should be done well enough to get rid of the second-order remainder.

Introduction to TMLE: Overview of today

Day 2: 8 - 9

Introduction to TMLE. The constructive proof of TMLE.

▶ The decomposition and the role of Step 1 & Step 2.

Day 2: 9 - 12

The targeting step. Solving the efficient influence curve equation.

Lunch.

Day 2: 13 - 15

Step 2: Super learning. (Thomas).