Day 2, Lecture 2

A tiny overview

"Classical" causal inference estimators:

- consistency of g-formula estimators rely on correct specification of the outcome regression $f(a,x) = \mathbb{E}[Y \mid A = a, X = x]$
- consistency of ipw estimators rely on correct specification of the propensity score π(a | x) = P(A = a | X = x)
- both types of estimators work poorly with machine learning (cross-validation does not help)

TMLE (and other estimators based on the efficient influence curve):

- built-in bias correction
- double robustness in consistency
- inference based on the efficient influence curve
- even when incorporating machine learning (under conditions!!)

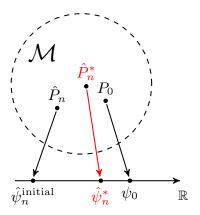
TMLE is a two-step procedure:

- Step 1 Construct initial estimator \hat{P}_n for P.
- Step 2 Update the estimator $\hat{P}_n \mapsto \hat{P}_n^*$ such that \hat{P}_n^* solves the efficient influence curve equation, i.e.,

$$\mathbb{P}_n \phi^* (\hat{P}_n^*) = \frac{1}{n} \sum_{i=1}^n \phi^* (\hat{P}_n^*) (O_i) \approx 0.$$

Step 1 = "initial estimation step"

Step 2 = "targeting step"



- in contrast to the estimating equation (EE) estimator (and the IPW estimator), TMLE is a substitution estimator.
- ► TMLE *may* show better small-sample performance in settings with unstable weights.

$$\Psi(\hat{P}_{n}) - \Psi(P_{0}) = \mathbb{P}_{n}\phi^{*}(P_{0}) + o_{P}(n^{-1/2}) + R(\hat{P}_{n}, P_{0}) - \mathbb{P}_{n}\phi^{*}(\hat{P}_{n})$$

- ▶ The role of the targeting step (Step 2):
 - Gain double robustness in consistency.
 - Easier to achieve asymptotic lineariy (amounts to getting rid of second-order remainder).
- ▶ The role of the initial estimation step (Step 1):
 - ▶ This should be done well enough to get rid of the second-order remainder.
 - ► The second-order remainder tells us if/how machine learning estimators can be incorporated while still achieving asymptotic linearity.

Specifically for the ATE:

$$\begin{split} \tilde{\Psi}(\hat{f}_n^*) - \tilde{\Psi}(f_0) &= \mathbb{P}_n \tilde{\phi}^*(f_0, \pi_0) + o_P(n^{-1/2}) \\ &+ \tilde{R}(\hat{f}_n^*, \hat{\pi}_n, f_0, \pi_0) - \underbrace{\mathbb{P}_n \tilde{\phi}^*(\hat{f}_n^*, \hat{\pi}_n)}_{=0, \text{ by targeting.}} \end{split}$$

When $\mathbb{P}_n \tilde{\phi}^*(\hat{f}_n^*, \hat{\pi}_n) = 0$, recall that:

$$|\tilde{R}(\hat{f}_{n}^{*},\hat{\pi}_{n},f_{0},\pi_{0})| \leq \sum_{a=0,1} \delta^{-1} \|\pi_{0}(a|\cdot) - \hat{\pi}_{n}(a|\cdot)\|_{\mu_{0}} \|f_{0}(a|\cdot) - \hat{f}_{n}^{*}(a|\cdot)\|_{\mu_{0}}$$

What this tells us:

• Asymptotic linearity when π_0 and f_0 are estimated at rate at least $n^{-1/4}$.

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Asymptotic linearity when π_0 and f_0 are estimated at rate at least $n^{-1/4}$.

We often see this structure of the second-order remainder, but note that it has to be verified on a case-by-case basis.

How can we perform estimation of π_0 and f_0 such as to achieve rate at least $n^{-1/4}$?

- Correctly specified parametric models
 - although consistency is guaranteed, inference cannot be based on the efficient influence curve when one is misspecified!
- There are no results on this being the case for generic implementations of, for example, random forests.
- Lasso, highly adaptive lasso (HAL), ...
- ► Loss-based "super learning"
 - oracle property: the super learner achieves the rate of convergence of the best estimator in its library.

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 - oracle property: the super learner achieves the rate of convergence of the best estimator in its library.
 - this is about minimizing expected loss; tuning is still important.

$$f(A,X) = \mathbb{E}_P[Y \mid A,X]$$

A loss function $\mathcal{L}(f)(O)$ measuring the distance between an estimator f and the observed outcome Y, e.g., the negative log-likelihood:

$$\mathscr{L}(\hat{f}_n)(Y_i, A_i, X_i) = -(Y_i \log(\hat{f}_n(A_i, X_i)) + (1 - Y_i) \log(1 - \hat{f}_n(A_i, X_i))).$$

▶ The estimator \hat{f}_n closest to the true f_0 minimizes the risk:

$$\mathbb{E}_{P_0}[\mathscr{L}(\hat{f}_n)(Y_i,A_i,X_i)].$$

Loss-based super learning: Minimizing the cross-validated empirical risk with respect to the loss function \(\mathcal{L} \) over the statistical model.

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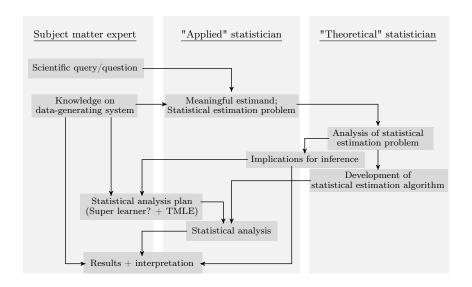
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Again, this does not usually by simple plug-in yield a good estimator for the particular feature of interest (the target parameter). But, combined with targeting, it is used to get rid of the second-order remainder.



- 1. Scientific question \Rightarrow causal parameter
- 2. Causal parameter ⇒ statistical parameter
- Statistical estimation problem = statistical parameter + statistical model
 - Efficient influence function
 - Second-order remainder
- 4. Identify relevant components that need targeting

 - ▶ Targeting algorithm
- Construct strong initial learners!!
 - ▶ *a priori* specified
- 6. Inference based on the efficient influence function

... from a theoretical perspective, for any new type of problem.

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... from a more applied perspective.

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Comment — substitution estimation

The g-formula estimator and the TMLE are substitution estimators:

$$\begin{split} \hat{\psi}_{n}^{\text{g-formula}} &= \tilde{\Psi}(\hat{f}_{n}, \hat{\mu}_{X}) = \int_{\mathbb{R}^{d}} \left(\hat{f}_{n}(1, x) - \hat{f}_{n}(0, x)\right) d\hat{\mu}_{X}(x) \\ \hat{\psi}_{n}^{\text{tmle}} &= \tilde{\Psi}(\hat{f}_{n}^{*}, \hat{\mu}_{X}) = \frac{1}{n} \sum_{i=1}^{n} \left\{\hat{f}_{n}^{*}(1, X_{i}) - \hat{f}_{n}^{*}(0, X_{i})\right\} \end{split}$$

The IPW estimator and the EE estimator are <u>not</u> substitution estimators:

$$\hat{\psi}_{n}^{\text{ipw}} = \frac{1}{n} \sum_{i=1}^{n} \left\{ \frac{A_{i} Y_{i}}{\hat{\pi}_{n}(A_{i} \mid X_{i})} - \frac{(1 - A_{i}) Y_{i}}{\hat{\pi}_{n}(A_{i} \mid X_{i})} \right\}
\hat{\psi}_{n}^{\text{ee}} = \hat{\psi}_{n}^{\text{ee}} = \frac{1}{n} \sum_{i=1}^{n} \left\{ \frac{A_{i}}{\hat{\pi}_{n}(A_{i} \mid X_{i})} - \frac{(1 - A_{i})}{\hat{\pi}_{n}(A_{i} \mid X_{i})} (Y_{i} - \hat{f}_{n}(A_{i}, X_{i})) + \hat{f}_{n}(1, X_{i}) - \hat{f}_{n}(0, X_{i}) \right\}$$

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The g-formula estimator and the TMLE are both based on the parametrization of the ATE as follows:

$$\tilde{\Psi}(f,\mu_X) = \mathbb{E}_P[f(1,X) - f(0,X)] = \int_{\mathbb{R}^d} (f(1,x) - f(0,x))\mu_X(x)$$
 (1)

corresponding to an average outcome under the post-interventional distribution.

- This will always have a statistical interpretation as a covariate-adjusted difference of treatment-specific risks (but may have no causal validity).
- ▶ Estimators based on plugging estimators for *f* (respecting the model constraints for *f*) into (1) will always respect the global constraints of the observed data model.

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