Day 2, Lecture 1

Targeted Minimum Loss-based Estimation (TMLE)

Overview of today

Before lunch (9-12):

- Targeted Minimum Loss-based Estimation (TMLE).
- The targeting step: updating/modifying initial nuisance parameter estimators.
- ▶ The ATE as a concrete example; the ATT as a different example.
- Valid inference still requires strong initial learners.
- * TMLE as a two-step procedure with involving an initial estimation step followed by a targeting step.
- * Implementation of the targeting step.
- * The link between the theoretical decomposition from yesterday, and TMLE as a practical estimation method.

After lunch (13 - 15): Super learning.

In this lecture, our goal is to:

- Explain the relationship between the two-step procedure of Targeted Minimum Loss-based Estimation (TMLE) and the decomposition guiding the construction of asymptotically linear estimators.
- Describe the process of constructing the targeting step in TMLE, highlighting the role of the efficient influence curve, and differentiate TMLE from other estimation methods solving the efficient influence curve.

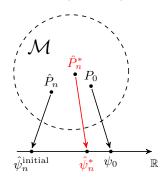
For a particular target parameter $\Psi: \mathcal{M}$ (and a particular statistical model \mathcal{M}),

we have seen that —

Conditions (asymptotic linearity and efficiency)

- (C1) Solve the efficient influence curve equation: $\mathbb{P}_n \phi^*(\hat{P}_n) = o_P(n^{-1/2})$
- (C2) Remainder $R(\hat{P}_n, P_0) = o_P(n^{-1/2})$
- (C3) Donsker class conditions for $\{\phi^*(P): P \in \mathcal{M}\}$

Then: $\Psi(\hat{P}_n) \stackrel{as}{\sim} N(\Psi(P_0), P_0\phi^*(P_0)^2/n)$



TMLE is a two-step procedure:

- Step 1 Construct initial estimator \hat{P}_n for P.
- Step 2 Update the estimator $\hat{P}_n \mapsto \hat{P}_n^*$ such that \hat{P}_n^* solves the efficient influence curve equation, i.e.,

$$\mathbb{P}_{n}\phi^{*}(\hat{P}_{n}^{*}) = \frac{1}{n}\sum_{i=1}^{n}\phi^{*}(\hat{P}_{n}^{*})(O_{i}) \approx 0.$$

Step 1 = "initial estimation step"

Step 2 = "targeting step"

$$\Psi(\hat{P}_{n}) - \Psi(P_{0}) = \mathbb{P}_{n}\phi^{*}(P_{0}) + o_{P}(n^{-1/2}) + R(\hat{P}_{n}, P_{0}) - \mathbb{P}_{n}\phi^{*}(\hat{P}_{n})$$

- ▶ The role of the targeting step (Step 2):
 - Gain double robustness in consistency.
 - Easier to achieve asymptotic lineariy (amounts to getting rid of second-order remainder).
- ▶ The role of the initial estimation step (Step 1):
 - This should be done well enough to get rid of the second-order remainder.

$$f(A,X) = \mathbb{E}_P[Y \mid A,X]$$

A loss function $\mathcal{L}(f)(O)$ measuring the distance between an estimator f and the observed outcome Y, e.g., the negative log-likelihood:

$$\mathcal{L}(\hat{f}_n)(Y_i, A_i, X_i) = -(Y_i \log(\hat{f}_n(A_i, X_i)) + (1 - Y_i) \log(1 - \hat{f}_n(A_i, X_i))).$$

▶ The estimator \hat{f}_n closest to the true f_0 minimizes the risk:

$$\mathbb{E}_{P_0}[\mathscr{L}(\hat{f}_n)(Y_i,A_i,X_i)].$$

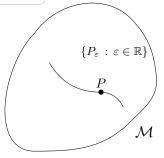
Loss-based super learning: Minimizing the cross-validated empirical risk with respect to the loss function & over the statistical model.

This is all about constructing a good estimator for the conditional expectation f;

does not necessarily yield a good estimator for the particular feature of interest, the target parameter.

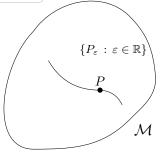
This is Step 1.

Step 2: We can minimize along a loss function in a certain way that results in a good estimator for the target.



Loss function $\mathcal{L}(f)(O)$ + clever choice of a parametric submodel $\{P_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$.

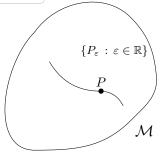
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Loss function $\mathcal{L}(f)(O)$ + clever choice of a parametric submodel $\{P_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$.

 \Rightarrow minimize the loss along the submodel, given the estimator \hat{f}_n from Step 1.

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Loss function $\mathcal{L}(f)(O)$ + clever choice of a parametric submodel $\{P_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$.

- \Rightarrow minimize the loss along the submodel, given the estimator \hat{f}_n from Step 1.
- \Rightarrow update \hat{f}_n along the path defined by P_{ε} : moving by $\hat{\varepsilon}_n$ that minimizes the loss.

Construction of the targeting step for a given target parameter $\Psi: \mathcal{M} \to \mathbb{R}$ with efficient influence function $\phi^*(P)$ requires:

- (i) A parametric submodel $\{P_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$
- (ii) A loss function $(O, P) \mapsto \mathcal{L}(P)(O)$

such that: (1)
$$P_{\varepsilon=0} = P$$
, and, (2) $\frac{d}{d\varepsilon}\Big|_{\varepsilon=0} \mathcal{L}(P_{\varepsilon})(O) = \phi^*(P)(O)$

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- ▶ Initial estimator \hat{P}_n^0
- ▶ Minimizer $\hat{\varepsilon}_{n,0}$ of $\varepsilon \mapsto \mathbb{P}_n \mathcal{L}(\hat{P}_{n,\varepsilon}^0)$
- Update: $\hat{P}_n^1 := \hat{P}_{\hat{\varepsilon}_{n,0}}^0$

Then:
$$\mathbb{P}_n \frac{d}{d\varepsilon}\Big|_{\varepsilon=\hat{\varepsilon}_{n,0}} \mathcal{L}(\hat{P}_{n,\varepsilon}^0)(O) = 0$$

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- Updated estimator \hat{P}_n^1
- ▶ Minimizer $\hat{\varepsilon}_{n,1}$ of $\varepsilon \mapsto \mathbb{P}_n \mathscr{L}(\hat{P}^1_{n,\varepsilon})$
- Update: $\hat{P}_n^2 := \hat{P}_{\hat{\varepsilon}_{n,1}}^1$

Then:
$$\mathbb{P}_n \frac{d}{d\varepsilon}\Big|_{\varepsilon=\hat{\varepsilon}_{n,1}} \mathcal{L}(\hat{P}_{n,\varepsilon}^1)(O) = 0$$

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- kth updated estimator \hat{P}_n^k
- ▶ Minimizer $\hat{\varepsilon}_{n,k}$ of $\varepsilon \mapsto \mathbb{P}_n \mathscr{L}(\hat{P}_{n,\varepsilon}^k)$
- Update: $\hat{P}_n^{k+1} := \hat{P}_{\hat{\varepsilon}_{n,k}}^k$

Then:
$$\mathbb{P}_n \frac{d}{d\varepsilon}\Big|_{\varepsilon=\hat{\varepsilon}_{n,k}} \mathcal{L}(\hat{P}_{n,\varepsilon}^k)(O) = 0$$

Construction of the targeting step for a given target parameter

$$\Psi: \mathcal{M} \to \mathbb{R}$$
 with efficient influence function $\phi^*(P)$ requires:

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- (ii) A loss function $(O, P) \mapsto \mathcal{L}(P)(O)$

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, so when $\hat{\varepsilon}_{n,k} \approx 0$: $\mathbb{P}_n \frac{d}{d\varepsilon} \Big|_{\varepsilon = 0} \mathcal{L}(\hat{P}_{n,\varepsilon}^k) = 0$.

Construction of the targeting step for a given target parameter $\Psi : \mathcal{M} \to \mathbb{R}$ with efficient influence function $\phi^*(P)$ requires:

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$$\{P_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$$

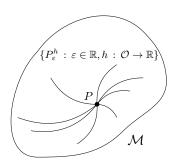
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What happens?



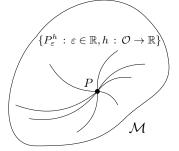
Parametric submodels $\{P_{\varepsilon} : \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$ are also what we use to:

define pathwise differentiability:¹

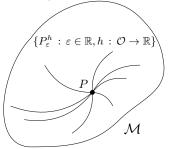
$$\frac{d}{d\varepsilon}\bigg|_{\varepsilon=0} \Psi(P_{\varepsilon}) = \int \phi(P)(o)b(o)dP(o), \tag{1}$$

derive a nonparametric lower bound on the variance.

¹(1) should hold across any smooth submodel $\{P_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$.

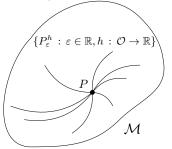


- ▶ Index submodel by its score function: $\{P_{\varepsilon}^h : \varepsilon \in \mathbb{R}, h : \mathcal{O} \to \mathbb{R}\}.$
 - i.e., $\frac{d}{d\varepsilon}\Big|_{\varepsilon=0}\log p_{\varepsilon}^h(o)=h(o)$.
- ▶ Easier to estimate Ψ in the smaller model $\{P_{\varepsilon}^h : \varepsilon \in \mathbb{R}\}$ than in \mathcal{M} .
- ▶ The supremum over Cramér-Rao bounds over all submodels $\{P_{\varepsilon}^h : \varepsilon \in \mathbb{R}\}$ for estimating $\varepsilon \mapsto \Psi(P_{\varepsilon}^h)$ at $\varepsilon = 0$ provides a lower bound on the variance for estimating Ψ in \mathcal{M} :



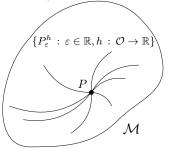
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$$\frac{(\frac{d}{d\varepsilon}\big|_{\varepsilon=0}\Psi(P_\varepsilon^h))^2}{Ph^2}$$



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$$\frac{\left(\frac{d}{d\varepsilon}\Big|_{\varepsilon=0}\Psi(P_{\varepsilon}^{h})\right)^{2}}{Ph^{2}} \stackrel{\text{PD}}{=} \frac{(P\phi^{*}(P)h)^{2}}{Ph^{2}} \tag{*}$$



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$$\frac{\left(\frac{d}{d\varepsilon}\Big|_{\varepsilon=0}\Psi(P_{\varepsilon}^{h})\right)^{2}}{Ph^{2}} \stackrel{\text{PD}}{=} \frac{(P\phi^{*}(P)h)^{2}}{Ph^{2}} \stackrel{\text{CS}}{\leq} P\{\phi^{*}(P)\}^{2} \tag{*}$$

The submodel which attains the supremum of the Cramér-Rao bounds over all parametric submodels is called the least favorable submodel;

 \triangleright It is the submodel for which the score is equal to the efficient influence function $\phi^*(P)$.

The TMLE step uses the least favorable submodel as a fluctuation model

- given a current estimator \hat{P}_n^k the updated estimator is found by fluctuating along the least favorable submodel;
- when no further fluctuation is needed ($\varepsilon \approx 0$), the estimator solves the efficient influence curve equation.

Conditions (asymptotic linearity and efficiency)

- (C1) Solve the efficient influence curve equation: $\mathbb{P}_n \phi^*(\hat{P}_n) = o_P(n^{-1/2})$
- (C2) Remainder $R(\hat{P}_n, P_0) = o_P(n^{-1/2})$
- (C3) Donsker class conditions for $\{\phi^*(P): P \in \mathcal{M}\}$

Then: $\Psi(\hat{P}_n) \stackrel{as}{\sim} N(\Psi(P_0), P_0\phi^*(P_0)^2/n)$

- ▶ The targeting step ensures that (C1) holds.
- Assume that (C2)^a and (C3) hold.

We can use the efficient influence function to compute an estimator for the standard error of the TMLE estimator:

$$\hat{\sigma}_n = \sqrt{\frac{\mathbb{P}_n\{\phi^*(\hat{P}_n)\}^2}{n}}$$

^aWith (C2) depending on how well we do for the initial estimation step.

EXAMPLE: Average treatment effect (ATE)

Observed data
$$O = (X, A, Y) \in \mathbb{R}^d \times \{0, 1\} \times \{0, 1\} = \mathcal{O}$$

- * $X \in \mathbb{R}^d$ are covariates
- * $A \in \{0,1\}$ is a binary exposure variable (treatment decision)
- * $Y \in \{0,1\}$ is a binary outcome variable

 $O \sim P_0$ where P_0 assumed to belong to nonparametric model \mathcal{M} .

We are interested in estimating the ATE:

$$\Psi(P) = \mathbb{E}_P \Big[\mathbb{E}_P \big[Y \mid A = 1, X \big] - \mathbb{E}_P \big[Y \mid A = 0, X \big] \Big].$$

EXAMPLE: Average treatment effect (ATE)

For the ATE, as we have seen, we can also write the target parameter $\Psi:\mathcal{M}\to\mathbb{R}$ as

$$\Psi(P) = \tilde{\Psi}(f, \mu_X) = \int_{\mathbb{R}^d} (f(1, x) - f(0, x)) d\mu_X(x) \qquad (*)$$

where

$$f(a,x) = \mathbb{E}[Y \mid A = a, X = x]$$

and μ_X is the marginal distribution of X.

I.e.,
$$\hat{\psi}_n = \tilde{\Psi}(\hat{f}_n, \hat{\mu}_n)$$
.

EXAMPLE: Average treatment effect (ATE)

- Step 1 Construct initial estimators \hat{f}_n , $\hat{\pi}_n$ for f, π .
- Step 2 Update the estimator $\hat{f}_n \mapsto \hat{f}_n^*$ for f such that \hat{f}_n^* for the fixed $\hat{\pi}_n$ solves the efficient influence curve equation.

For the ATE, Step 2 is simply just an additional logistic regression step.

EXAMPLE: Average treatment effect (ATE)

We need:

0. The efficient influence function:

$$\widetilde{\phi}^*(f,\pi)(O) = \left(\frac{A}{\pi(A\mid X)} - \frac{1-A}{\pi(A\mid X)}\right) (Y - f(A,X))$$

+
$$f(1,X) - f(0,X) - \tilde{\Psi}(f)$$

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Further, we need:

(i) A parametric submodel
$$\{f_{\varepsilon}: \varepsilon \in \mathbb{R}\} \subset \mathcal{M}$$

(ii) A loss function
$$(O, f) \mapsto \mathcal{L}(f)(O)$$

such that

(1)
$$f_{\varepsilon=0} = f$$
 (2) $\frac{d}{d\varepsilon}\Big|_{\varepsilon=0} \mathscr{L}(f_{\varepsilon})(O) = \tilde{\phi}_{f}^{*}(f,\pi)(O)$

(i) Log-likelihood loss function:

$$\operatorname{logit}(p) = \operatorname{expit}^{-1}(p) = \operatorname{log}\left(\frac{p}{1-p}\right)$$

$$\mathcal{L}(f)(O) = -(Y\log(f(A,X)) + (1-Y)\log(1-f(A,X)))$$

(ii) Logistic regression model:

$$f_\varepsilon(A,X) = \operatorname{expit} \bigl(\operatorname{logit}(f(A,X)) + \varepsilon H(A,X) \bigr)$$

with the so-called "clever covariate": $H(A, X) := \frac{2A-1}{\pi(A \mid X)}$.

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To show this, we verify that (i)–(ii) fulfill

(1)
$$f_{\varepsilon=0} = f$$
 (2) $\frac{d}{d\varepsilon}\Big|_{\varepsilon=0} \mathscr{L}(f_{\varepsilon})(O) = \tilde{\phi}_f^*(f,\pi)(O)$

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SMALL EXERCISE: To show this, we verify that (i)–(ii) fulfill

(1)
$$f_{\varepsilon=0} = f$$
 (2) $\frac{d}{d\varepsilon}\Big|_{\varepsilon=0} \mathscr{L}(f_{\varepsilon})(O) = \tilde{\phi}_f^*(f,\pi)(O)$

- Initial estimators \hat{f}_n , $\hat{\pi}_n$.
- ▶ Estimate clever covariate by:

$$\hat{H}_n(A,X) = \frac{2A-1}{\hat{\pi}_n(A\mid X)}.$$

▶ The minimizer $\hat{\varepsilon}_n$ of $\varepsilon \mapsto \mathbb{P}_n \mathcal{L}(\hat{f}_{n,\varepsilon})$ equals the maximum likelihood estimator for ε in the fixed-intercept logistic regression:

$$\operatorname{logit} \mathbb{E}[Y \mid A, X] = \operatorname{logit}(\hat{f}_n(A, X)) + \varepsilon \hat{H}_n(A, X)$$

• Update: $\hat{f}_n^* := \hat{f}_{n,\hat{\varepsilon}_n}$.

Then:
$$\mathbb{P}_n \frac{d}{d\varepsilon} \Big|_{\varepsilon = \hat{\varepsilon}_n} \mathcal{L}(\hat{f}_{n,\varepsilon}) = 0$$
, i.e., $\mathbb{P}_n \tilde{\phi}_f^*(\hat{f}_{n,\hat{\varepsilon}_n}, \hat{\pi}_n) = \mathbb{P}_n \tilde{\phi}_f^*(\hat{f}_n^*, \hat{\pi}_n) = 0$.

$$\tilde{\phi}^*(f,\pi)(O) = \underbrace{\left(\frac{A}{\pi(A\mid X)} - \frac{1-A}{\pi(A\mid X)}\right) \left(Y - f(A,X)\right)}_{=\tilde{\phi}^*_f(f,\pi)(O)} + \underbrace{f(1,X) - f(0,X) - \tilde{\Psi}(f)}_{=\tilde{\phi}^*_{\mu_X}(f)(O)}$$

Per construction we already have: $\mathbb{P}_n \phi_{\mu}^*(\hat{f}_n^*) = 0$,

since:
$$\tilde{\Psi}(\hat{f}_n^*) = \frac{1}{n} \sum_{i=1}^n (\hat{f}_n^*(1, X_i) - \hat{f}_n^*(0, X_i)) = \mathbb{P}_n(\hat{f}_n^*(1, \cdot) - \hat{f}_n^*(0, \cdot)).$$

The targeting step thus yields:

$$\mathbb{P}_n\tilde{\phi}^*(\hat{f}_n^*,\hat{\pi}_n) = \mathbb{P}_n\tilde{\phi}_f^*(\hat{f}_n^*,\hat{\pi}_n) + \mathbb{P}_n\phi_u^*(\hat{f}_n^*) = 0.$$

Doing the targeting in practice using the simulated dataset:

```
set.seed(5)
   n < -500
   X \leftarrow runif(n, -2, 2)
   A \leftarrow rbinom(n, 1, prob=plogis(-0.25 + 1.2*X))
   Y < - rbinom(n, 1, prob=plogis(-0.9 + 1.9*X^2 + 0.5*A))
   (sim.data <- data.table(id=1:n,X=X,A=A,Y=Y))</pre>
                  XAY
      id
     1 -1.1991422 0 1
  1:
  2: 2 0.7408744 1 1
  3: 3 1.6675031 1 1
  4: 4 -0.8624022 0 1
  5: 5 -1.5813995 0 1
496: 496 -0.3978523 1 0
497: 497 -1.5069379 0 1
498: 498 1.8340120 1 1
499: 499 0.6349484 1 0
500: 500 -0.5214807 0 1
```

Initial estimation:

```
#-- treatment distribution;
glm.A <- glm(A~X, data=sim.data, family=binomial)</pre>
pi.1 <- predict(glm.A, type="response")</pre>
#-- outcome distribution (misspecified);
glm.Y <- glm(Y~A+X, data=sim.data, family=binomial)</pre>
sim.data[, f:=predict(glm.Y, type="response")]
sim.data[, f.A1:=predict(glm.Y, type="response",
             newdata=copy(sim.data)[, A:=1])]
sim.data[, f.A0:=predict(glm.Y, type="response",
             newdata=copy(sim.data)[, A:=0])]
#-- initial estimate of the ATE:
fit.ate.initial <- sim.data[, mean(f.A1 - f.A0)]</pre>
```

Targeting step:

```
eps = -0.0157708436790858
```

Targeting step:

eps = -0.0157708436790858

```
#-- tmle update;
sim.data[, f.A1.tmle:=plogis(qlogis(f.A1) + eps/pi.1)]
sim.data[, f.A0.tmle:=plogis(qlogis(f.A0) - eps/(1-pi.1))]
```

i.e., f.A1.tmle is the estimate of $f(1,X) = \mathbb{E}[Y \mid A = 1,X]$, obtained via the submodel:

$$\hat{f}_n^*(1,X) = \hat{f}_{n,\hat{\varepsilon}_n}(1,X) = \operatorname{expit} \bigl(\operatorname{logit}\bigl(\hat{f}_n(1,X)\bigr) + \hat{\varepsilon}_n \hat{H}_n(1,X)\bigr),$$

and likewise with f.AO.tmle.

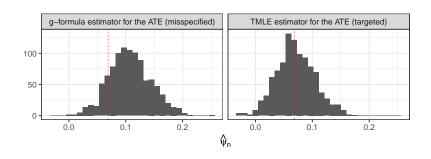
```
id
                 X A Y f.A1 f.A0 f.A1.tmle f.A0.tmle
  1:
      1 -1.1991422 0 1 0.7655621 0.6713853 0.7488795 0.6755825
  2:
      2 0.7408744 1 1 0.7396070 0.6399080 0.7349584 0.6504368
  3:
      3 1.6675031 1 1 0.7265721 0.6244167 0.7228545 0.6481588
 4:
      4 -0.8624022 0 1 0.7611886 0.6660214 0.7488197 0.6705960
  5:
     5 -1.5813995 0 1 0.7704590 0.6774205 0.7463439 0.6813231
 ___
496: 496 -0.3978523 1 0 0.7550638 0.6585507 0.7464799 0.6639337
497: 497 -1.5069379 0 1 0.7695108 0.6762494 0.7471142 0.6802008
498: 498 1.8340120 1 1 0.7241872 0.6216047 0.7205492 0.6495635
499: 499 0.6349484 1 0 0.7410712 0.6416611 0.7362345 0.6513868
500: 500 -0.5214807 0 1 0.7567041 0.6605467 0.7472996 0.6656728
```

```
id
                 X A Y f.A1 f.A0 f.A1.tmle f.A0.tmle
  1 •
      1 -1 1991422 0 1 0.7655621 0.6713853 0.7488795 0.6755825
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      3 1.6675031 1 1 0.7265721 0.6244167 0.7228545 0.6481588
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 5:
     5 -1.5813995 0 1 0.7704590 0.6774205 0.7463439 0.6813231
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500: 500 -0.5214807 0 1 0.7567041 0.6605467 0.7472996 0.6656728
```

```
fit.ate.tmle <- sim.data[, mean(f.A1.tmle - f.A0.tmle)]</pre>
```

```
initial ate est = 0.0975
tmle ate est = 0.0768
```

With 500 repeated simulations:



Practical 1: Implementing the targeting step

Practical Part 1 Implementing the targeting step.

Practical Part 2 Computing the variances of the ATE, the $\log RR$ and the $\log OR$.

Practical Part 3 Large-sample properties (simulation study).

The exercise is described in detail in: day2-practical1.pdf.

[More comments on the following slides].

We focused on the ATE as an example of a causal parameter.

But note that other simple causal parameters can be constructed from $\mathbb{E}_P[Y^1]$ and $\mathbb{E}_P[Y^0]$.

Like:

$$\Psi_{\mathsf{RR}}(P) = \frac{\mathbb{E}_{P}[Y^{1}]}{\mathbb{E}_{P}[Y^{0}]},$$

or,

$$\Psi_{\mathrm{OR}}(P) = \frac{\mathbb{E}_{P}[Y^{1}]/(1 - \mathbb{E}_{P}[Y^{1}])}{\mathbb{E}_{P}[Y^{0}]/(1 - \mathbb{E}_{P}[Y^{0}])},$$

For the targeting step, we can choose to target $\Psi_1(P) = \mathbb{E}_P[Y^1]$ and $\Psi_0(P) = \mathbb{E}_P[Y^0]$ separately.

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The efficient influence function for the treatment-specific mean $\Psi_a(P) = \mathbb{E}_P[Y^a]$:

$$\widetilde{\phi}_{a}^{*}(f,\pi)(O) = \underbrace{\frac{1\{A=a\}}{\pi(a|X)}}_{\text{clever covar.}} (Y - f(A,X)) + f(a,X) - \Psi_{a}(P)$$

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$$\widetilde{\phi}_{a}^{*}(f,\pi)(O) = \underbrace{\frac{1\{A=a\}}{\pi(a\mid X)}}_{\text{clever covar.}} (Y-f(A,X)) + f(a,X) - \Psi_{a}(P)$$

If we target $\Psi_1(P)$ and $\Psi_0(P)$ separately, we obtain two sets of updated estimators $\hat{f}_n \mapsto \hat{f}_{n,1}^*$ and $\hat{f}_n \mapsto \hat{f}_{n,0}^*$

- one to construct a targeted estimator $\hat{\psi}_{1,n}^*$ for $\Psi_1(P)$;
- ▶ and the other to construct a targeted estimator $\hat{\psi}_{0,n}^*$ for $\Psi_0(P)$.

We can then compute an estimate for the ATE as

$$\hat{\psi}_{n}^{*} = \hat{\psi}_{n,1}^{*} - \hat{\psi}_{n,0}^{*},$$

and we can estimate the variance of this estimator by

$$\mathbb{P}_{n}\{\tilde{\phi}_{1}^{*}(\hat{f}_{n,1}^{*},\hat{\pi}_{n})-\tilde{\phi}_{0}^{*}(\hat{f}_{n,0}^{*},\hat{\pi}_{n})\}^{2};$$

since efficient influence function for the ATE is

$$\tilde{\phi}^*(f,\pi) = \tilde{\phi}_1^*(f,\pi) - \tilde{\phi}_0^*(f,\pi).$$

Similarly we can construct estimators for the RR and the OR by simple plug-in:

$$\hat{\psi}_{\mathsf{RR},n}^* = \frac{\hat{\psi}_{1,n}^*}{\hat{\psi}_{0,n}^*},$$

and,

$$\hat{\psi}_{\mathsf{OR},n}^* = \frac{\hat{\psi}_{1,n}^*/(1-\hat{\psi}_{1,n}^*)}{\hat{\psi}_{0,n}^*/(1-\hat{\psi}_{0,n}^*)}.$$

Similarly we can construct estimators for the RR and the OR by simple plug-in:

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We can use the delta method to derive the efficient influence functions of $\Psi_{RR}(P)$ and $\Psi_{OR}(P)$.

Let $\phi^*(P)$ be the efficient influence function for a parameter $\Psi(P)$. Say that interest is in $h(\Psi(P))$ for a function h.

The delta method yields that:

If the first derivative $h'(\psi) = \frac{d}{d\psi}h(\psi)$ of h exists and is non-zero, then the efficient influence function of $h(\Psi(P))$ is:

$$\phi_h^*(P) = h'(\Psi(P))\phi^*(P).$$

So, once we have TMLE (targeted) estimators for $\Psi_1(P) = \mathbb{E}[Y^1]$ and $\Psi_0(P) = \mathbb{E}[Y^0]$:

- We can construct estimators for the ATE, the RR and the OR.
- \blacktriangleright We can compute the variance of the ATE estimator, the log RR estimator and the log OR estimator.

Practical 1: Implementing the targeting step

Practical Part 1 Implementing the targeting step.

Practical Part 2 Computing the variances of the ATE, the $\log RR$ and the $\log OR$.

Practical Part 3 Large-sample properties (simulation study).

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