

Introduction to penalized splines

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Outline

Introduction

Splines, regression splines and Smoothing splines

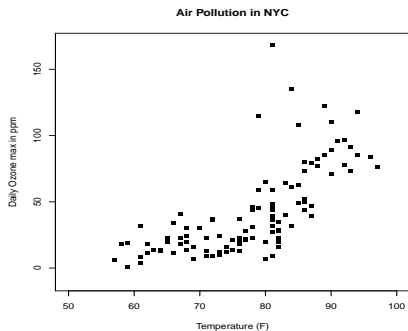
Penalized regression and semi-parametric models

Penalized splines in **R**

From linear regression to penalized regression

Basics

- Scatterplot of pairs (x_i, y_i) , $i = 1, \dots, n$



- **Assumption:** straight line fits data well

$$E[y_i|x_i] = \beta_0 + \beta_1 x_i$$

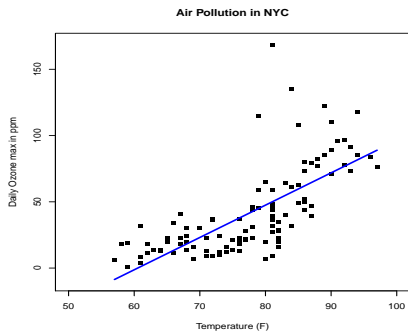
- Minimize least squares criteria

$$\min \|y - \hat{y}\|^2$$

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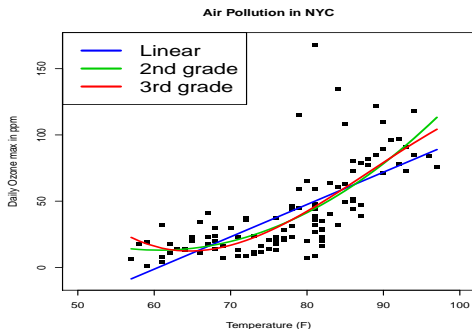
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From linear regression to penalized regression

Beyond linear regression

- ▶ Linear fit is **too simple** and not always OK
- ▶ **Alternative:** use higher degree powers of x , with $X = [1_n, x_i, x_i^2, \dots, x_i^p]$
- ▶ Columns of X are **basis functions** (polynomial regression)



- ▶ Same regression equations: $\hat{\beta} = (X'X)^{-1}X'y$

From linear regression to penalized regression

non-parametric regression or smoothing

- ▶ Instead of simple linear regression, we can fit the model:

$$E[y_i|x_i] = f(x)$$

where $f(x)$ is an arbitrary **smooth function**

- ▶ The model is “non-parametric” in the sense of no distributional assumptions for the parameters as in a linear model (**scatterplot smoothing**).
- ▶ Wide variety of methods (since 80's):
 - ▶ kernel smoothing, local linear methods, smoothing splines, etc ...
(**out-of-fashion!**)
 - ▶ **Penalized regression splines** are getting popular
 - ▶ combine a rich set of **basis functions**, with a **roughness penalty**

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 - ▶ Many knots/high degree: We run the risk of overfitting (variance).

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$$g_1(x) = 1, \quad g_2(x) = x, \dots, \quad g_{k+1}(x) = x^k, \\ g_{k+j+1}(x) = (x - t_j)_+^k, \quad j = 1, \dots, m.$$

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- ▶ The truncated power basis are:
 - ▶ Conceptually simple.
 - ▶ Simpler models are nested inside it, leading to straightforward tests of null hypotheses.
 - ▶ Computationally inefficient (singularity problems).

Splines (cont.)



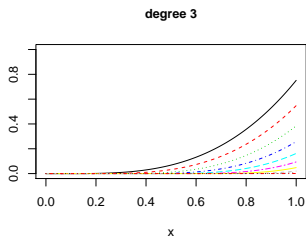
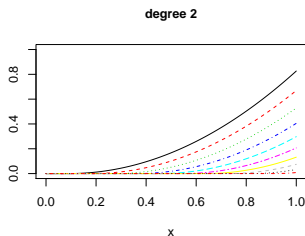
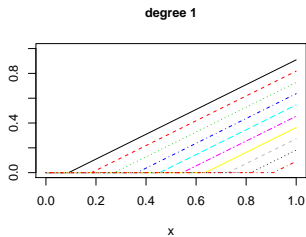
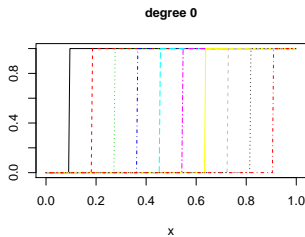
Splines (cont.)



Truncated power basis in R

```
> tpoly<-function(x,t,p){  
+ # p order truncated polynomials  
+ B=NULL  
+ for(i in 1:length(t)){  
+ B=cbind(B,(x-t[i])^p*(x>t[i]))}  
+ B  
+ }
```

Truncated power basis in R



B-splines (de Boor, 1978)

- ▶ a much better computational choice, both for speed and numerical accuracy, is the B-spline basis.
- ▶ **One linear B-spline**
 - ▶ Two pieces, each a straight line, everything else zero
 - ▶ Nicely connected at knots (t_1 to t_3) same value
 - ▶ Slope jumps at knots
- ▶ **One quadratic B-spline**
 - ▶ Three pieces, each a quadratic segment, rest zero
 - ▶ Nicely connected at knots (t_1 to t_4): same values and slopes
 - ▶ Shape similar to Gaussian curve.
- ▶ **One cubic B-spline**
 - ▶ Four pieces, each a cubic segment, rest zero
 - ▶ At knots (t_1 to t_5): same values, first and second derivatives
 - ▶ Shape more similar to Gaussian curve.

B-splines in R

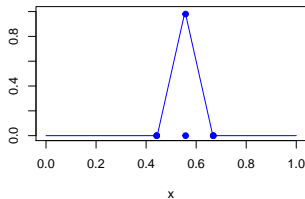
```

> library(splines)
> bspline <- function(x, xl, xr, ndx, bdeg){
+   dx <- (xr-xl)/ndx
+   knots <- seq(xl-bdeg*dx, xr+bdeg*dx, by=dx)
+   B <- spline.des(knots,x,bdeg+1,0*x)$design
+   B
+ }
>
> #   xl = left boundry of domain
> #   xr = right boundry of domain
> #   ndx = number of intervals for B-splines.
> #   bdeg = degree of B-spline

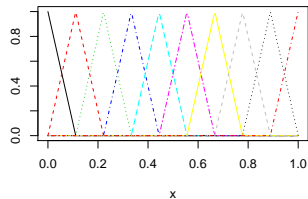
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B-splines basis in R

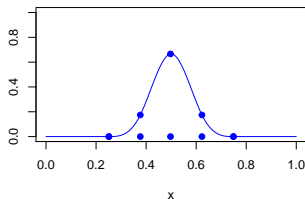
One linear B-spline



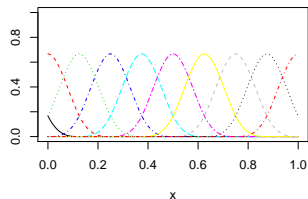
One linear B-spline basis



One cubic B-spline



One cubic B-spline basis



B-splines basis

- ▶ Basis matrix B
- ▶ Columns are B-splines

$$B = \begin{bmatrix} B_1(x_1) & B_2(x_1) & B_3(x_1) & \dots & B_m(x_1) \\ B_1(x_2) & B_2(x_2) & B_3(x_2) & \dots & B_m(x_2) \\ B_1(x_3) & B_2(x_3) & B_3(x_3) & \dots & B_m(x_3) \\ \vdots & \vdots & \vdots & \dots & \vdots \\ B_1(x_n) & B_2(x_n) & B_3(x_n) & \dots & B_m(x_n) \end{bmatrix}$$

- ▶ In each row only a few non-zero elements (degree plus one)
- ▶ Demo
 - > `library(gamlss.demo)`
 - > `demo.BSplines()`

Smoothing splines

- ▶ Regularized regression over the **natural spline basis**
- ▶ Minimize *penalized sum-of-squares*:

$$PSS(f, \lambda) = \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda \int_{x_1}^{x_n} f''(x)^2 dx$$

- ▶ First term balances the goodness-of-fit
- ▶ Second term penalizes the second derivative of the function (i.e. the curvature)
- ▶ λ is the so-called smoothing parameter that controls the balance between bias and variance.
- ▶ They put a knots on every data point x_1, \dots, x_n (solver the knots selection problem).
- ▶ $0 < \lambda < \infty$ if $\lambda = 0$ the fit interpolates the data, when $\lambda \rightarrow \infty$ the second derivative goes to zero, and the fit is linear.
- ▶ in R, the `smooth.spline()`.

Natural spline basis

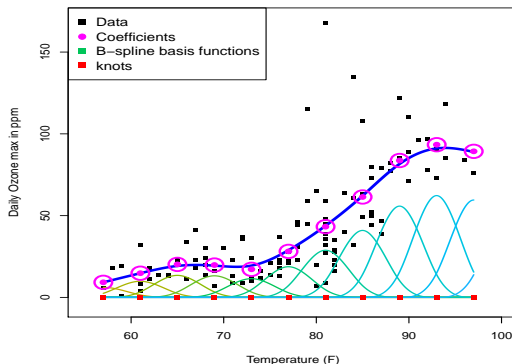
- ▶ One problem with regression splines is that the estimates tend to display erratic behavior, i.e., they have high variance, at the boundaries of the domain of x_1, \dots, x_n . This gets worse as the order k gets larger
- ▶ A way to remedy this problem is to force the piecewise polynomial function to have a lower degree to the left of the leftmost knot, and to the right of the rightmost knot.
- ▶ in R, the `ns()`.

- ▶ Smoothing splines often deliver similar fits to those from kernel regression.
- ▶ However, they are in a sense simpler.
- ▶ Both have a tuning parameter the bandwidth h for kernel regression, and the smoothing parameter λ for smoothing splines, which we would typically need to choose by cross-validation.
- ▶ But for smoothing splines, we don't require a choice of kernel.
- ▶ Smoothing splines are generally much more computationally efficient.

Penalized regression

P-splines (Eilers and Marx, 1996)

- ▶ The model: $f(x) = B\theta$, where B is regression basis and θ the new vector of coefficients which we penalize
- ▶ **Example:** Air pollution in NYC



Penalized regression

P-splines (Eilers and Marx, 1996)

- ▶ Estimation by **penalized least squares**, such that:

$$\min \|y - B\theta\|^2 \rightarrow \hat{\theta} = (B'B + P)^{-1} B'y$$

$$\hat{y} = B\hat{\theta}$$

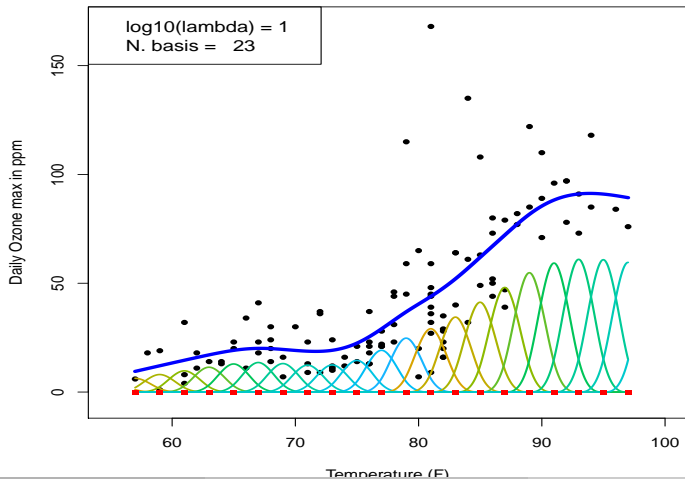
P is a **roughness penalty for smoothness** controlled by λ

- ▶ Choose the size of B and λ (and penalty order)
 - ▶ $20 < \text{knots} < 40$
 - ▶ $0 < \lambda < \infty$

Penalized regression

P-splines (Eilers and Marx, 1996)

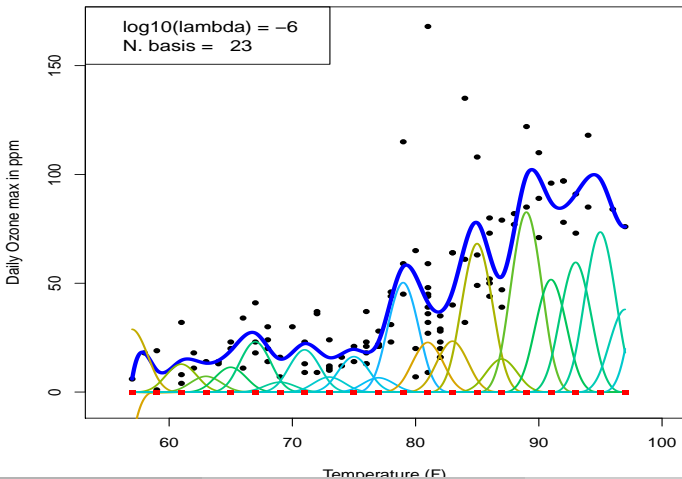
- Choose λ to tune the fit



Penalized regression

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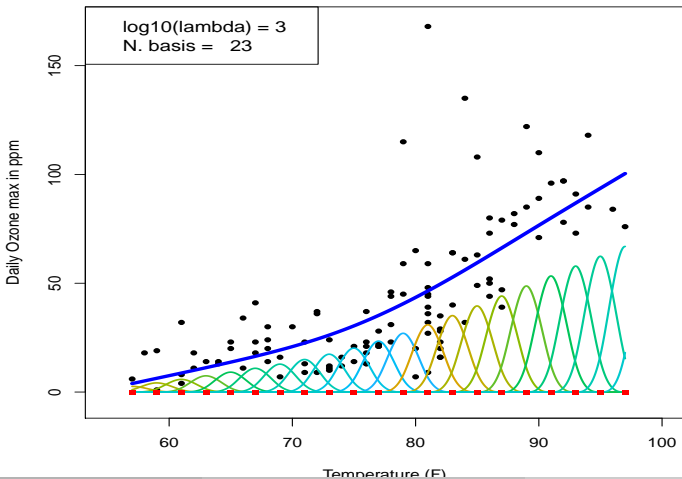
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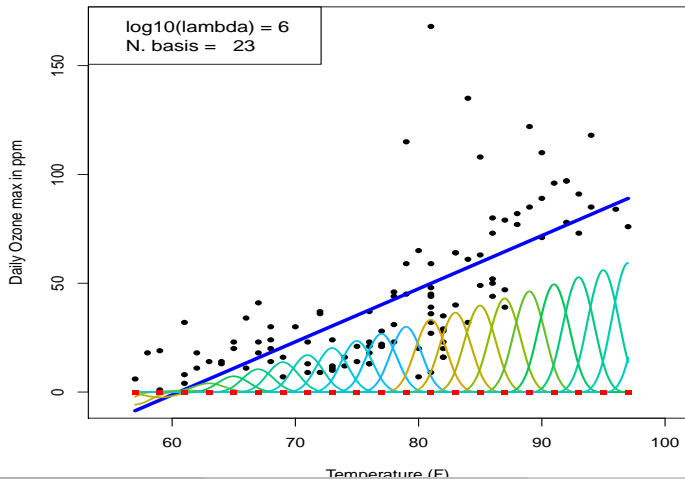
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Penalized regression

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P-splines

Knots and penalties

- ▶ Choose a moderate number of equidistant knots ($k \ll n$), e.g. $20 \leq k \leq 40$
- ▶ Add a penalty to the measure of fit, to tune smoothness
 - ▶ Discrete differences of coefficients

$$\|y - B\theta\|^2 + \lambda \sum_j (\Delta^d \theta_j)^2$$

where Δ^d is a difference operator of order d , i.e.

$$\Delta \theta_j = \theta_j - \theta_{j-1} \quad \text{(First order)}$$

$$\Delta^2 \theta_j = \theta_j - 2\theta_{j-1} + \theta_{j-2} \quad \text{(Second order)}$$

$$\vdots$$

P-splines

Penalties

- ▶ We are interested in differences on the coefficients
- ▶ $\Delta^d \theta = D_d \theta$ (*Difference matrix operator*)

$$D_1 = \begin{bmatrix} -1 & 1 & 0 & 0 & \cdots \\ 0 & -1 & 1 & 0 & \cdots \\ 0 & 0 & -1 & 1 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix} \quad \text{or} \quad D_2 = \begin{bmatrix} 1 & -2 & 1 & 0 & \cdots \\ 0 & 1 & -2 & 1 & \cdots \\ 0 & 0 & 1 & -2 & \cdots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

- ▶ The **penalty** becomes $P = \lambda D' D$
- ▶ Estimation is done by **penalized least squares**:

$$\min \|y - B\theta\|^2 + P \rightarrow \hat{\theta}_{\lambda^*} = (B' B + P)^{-1} B' y$$

- ▶ λ^* can be selected by criteria as **AIC**, **BIC**, or **GCV**

Selection of λ

- Optimal λ^* can be selected by criteria as **AIC**, **BIC**, or **GCV**

$$GCV = \sum_{i=1}^n \frac{(y_i - \hat{y}_i)^2}{n - \text{trace}(\mathbf{H})}; \quad \mathbf{H} = \mathbf{B}(\mathbf{B}'\mathbf{B} + \lambda\mathbf{D}'\mathbf{D})^{-1}\mathbf{B}'$$

$$AIC = 2\log \left(\sum_{i=1}^n (y_i - \hat{y}_i)^2 \right) - 2\log(n) + 2\log(\text{trace}(\mathbf{H}))$$

- Evaluate on a large grid of λ 's.

Demo P-splines

```
► library(gamlss.demo)  
   demoPsplines()
```

Smoothing and mixed models

a semi-parametric approach

- ▶ Reformulate $y = B\theta + \epsilon$ into:

$$y = X\beta + Z\alpha + \epsilon, \quad \begin{pmatrix} \alpha \\ \epsilon \end{pmatrix} \sim \mathcal{N} \left[\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} G & 0 \\ 0 & \sigma^2 I \end{pmatrix} \right]$$

where $G = \sigma_\alpha^2 I$, is the random effects covariance

λ estimation becomes the ratio $\sigma^2 / \sigma_\alpha^2$

- ▶ **Flexibility:**

- Easy incorporation of smoothing in **complex models**: hierarchical models, multi-level models, longitudinal data, correlated errors ...

- ▶ **Mixed models theory** for **estimation** and **inference**

- ▶ **Extension to non-gaussian data** (Poisson, Binomial, etc ...)

- ▶ Generalized Linear Models (GLM's) to GL(Mixed)M's

Bayesian P-splines

Lang and Brezger (2004)

- ▶ The bayesian analogue of P-splines replace differences with Gaussian random walks as priors on the regression coefficients θ_j
- ▶ First/second order θ_j , corresponds to a B-splines basis B :

$$\theta_j = \theta_{j-1} + v_j \quad \text{or} \quad \theta_j = 2\theta_{j-1} - \theta_{j-2} + v_j$$

where $v_j \sim N(0\tau^2)$

- ▶ The amount of smoothing is controlled by $\tau^2 = \sigma^2/\lambda$
- ▶ In general, we can rewrite P as:

$$\theta|\tau^2 \propto \exp\left(-\frac{1}{2\tau^2}\theta'P\theta\right)$$

the **rank** of P is $c-1$ for first order RW, and $c-2$ for 2nd order (improper prior)

- ▶ As previously

$$y|\theta \sim N(B\theta, I\sigma^2),$$

- ▶ Software BayesX, Inla
- ▶ Hierarchical models (mixed model representation): BuGS, JAGS

Multidimensional P-splines

- ▶ (Generalized) Additive Models (Hastie & Tibshirani, 1990)

$$\eta = f(\mathbf{x}_1) + f(\mathbf{x}_2) + \epsilon$$

- ▶ Smooth ANOVA models (Lee and Durban, 2011)

$$\eta = f(\mathbf{x}_1) + f(\mathbf{x}_2) + f(\mathbf{x}_1, \mathbf{x}_2) + \epsilon$$

- ▶ For higher dimensions one may incur in the **curse of dimensionality** (computational problems)
- ▶ Other approaches includes Thin plate splines (radial basis functions). Problems: knots selection and position in larger dimensions.
- ▶ Recommended approach use for interactions: **Tensor Products**

GAMs with P-splines

- ▶ Use B-splines $\eta = f(x_1) + f(x_2) = B_1\theta_1 + B_2\theta_2$
- ▶ Vectorize and combine

$$\eta = [B_1 : B_2] \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix} = \mathbf{B}\theta$$

- ▶ Difference penalty is blockdiagonal, i.e.

$$\mathbf{P} = \text{bdiag}(\mathbf{P}_1, \mathbf{P}_2) = \begin{bmatrix} \lambda_1 D_1' D_1 & \\ & \lambda_2 D_2' D_2 \end{bmatrix}$$

- ▶ In general $\lambda_1 \neq \lambda_2$ (Anisotropic smoothing)

Two-dimensional smoothing with P-splines

- ▶ Use B-splines $\eta = f(x_1) + f(x_2) = B_1\theta_1 + B_2\theta_2$
- ▶ Vectorize and combine

$$\eta = [B_1 : B_2] \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix} = B\theta$$

- ▶ Difference penalty is blockdiagonal, i.e.

$$P = \text{bdiag}(P_1, P_2) = \begin{bmatrix} \lambda_1 D_1' D_1 & \\ & \lambda_2 D_2' D_2 \end{bmatrix}$$

- ▶ In general $\lambda_1 \neq \lambda_2$ (Anisotropic smoothing)

2d P-splines

Bivariate smoothing

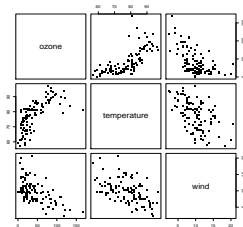
- ▶ Bivariate data, $(\mathbf{x}_1, \mathbf{x}_2, \mathbf{y})$, regression model $E[\mathbf{y}|\mathbf{x}_1, \mathbf{x}_2] = f(\mathbf{x}_1, \mathbf{x}_2)$
- ▶ B -splines on \mathbf{x}_1 and \mathbf{x}_2 domains, i.e. $B_k(\mathbf{x}_1)$ and $B_l(\mathbf{x}_2)$
- ▶ We aim to build a **surface** as a **sum of Tensor products**

$$f(\mathbf{x}_1, \mathbf{x}_2) = \sum_k \sum_l B_k(\mathbf{x}_1) B_l(\mathbf{x}_2) \theta_{kl}$$

- ▶ Now we have a matrix of coefficients $\mathbf{A} = [\theta_{kl}]$, to be penalized with λ_1 and λ_2

E.g.: Air quality data in NYC

- ▶ Covariates:
 - ▶ \mathbf{x}_1 = Daily max temperature (in F)
 - ▶ \mathbf{x}_2 = Wind speed (in mph)



2d P-splines

Bivariate smoothing

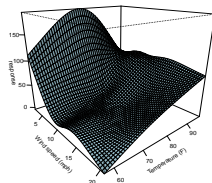
- ▶ Bivariate data, $(\mathbf{x}_1, \mathbf{x}_2, \mathbf{y})$, regression model $E[\mathbf{y}|\mathbf{x}_1, \mathbf{x}_2] = f(\mathbf{x}_1, \mathbf{x}_2)$
- ▶ B -splines on \mathbf{x}_1 and \mathbf{x}_2 domains, i.e. $B_k(\mathbf{x}_1)$ and $B_l(\mathbf{x}_2)$
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Multidimensional smoothing

Array structured data

- **Data** $\mathbf{Y} = y_{ij}$, $i = 1, \dots, n_1$ and $j = 1, \dots, n_2$
- **Array structure:** n_1 rows and n_2 columns

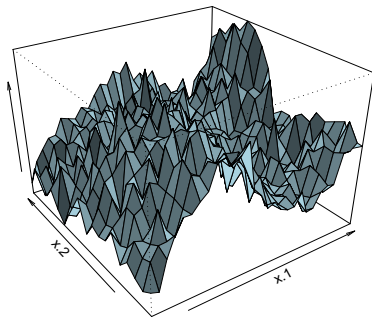
$$\mathbf{Y} = \begin{bmatrix} y_{11} & y_{12} & \cdots & y_{1n_2} \\ y_{21} & y_{22} & \cdots & y_{2n_2} \\ \vdots & \vdots & \ddots & \vdots \\ y_{n_1 1} & \cdots & \cdots & y_{n_1 n_2} \end{bmatrix}$$

- **Regressors:**

$$\mathbf{x}_1 = (x_{11}, \dots, x_{1n_1})'$$

$$\mathbf{x}_2 = (x_{21}, \dots, x_{2n_2})'$$

- **e.g.:** image data, mortality life tables, micro-arrays etc ...



Multidimensional smoothing

Array structured data

– Tensor Products of B -splines

► Marginal Basis:

- $B_1 = B_1(x_1)$, $n_1 \times c_1$
- $B_2 = B_2(x_2)$, $n_2 \times c_2$

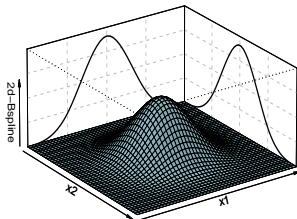
► $2d$ B -splines Basis:

- Kronecker Product (\otimes) of marginal basis:

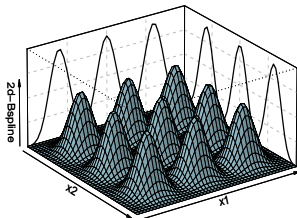
$$B = B_2 \otimes B_1, \quad n_1 n_2 \times c_1 c_2$$

- Computationally efficient methods
(GLAM, Currie, Durban and Eilers, 2006)

Tensor product of 2 cubic B -splines



B -spline basis of 3×3



References

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- [3] Lang, S. and Brezger, A. *Bayesian P-Splines* Journal of Computational and Graphical Statistics. Vol 13, 2004, pages 183-212
- [4] Lee, D.-J. and Durbán, M. (2011) *P-spline ANOVA-type interaction models for spatio-temporal smoothing* Statistical Modelling, Vol. 11, Issue 1, Pages 49-69.
- [5] Ruppert, D., Wand, M.P. and Carroll, R.J. (2009) *Semiparametric regression during 2003-2007* Electron. J. Statist. Volume 3 (2009), 1193-1256.
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Penalized splines in R

What's next?

<http://idaejin.github.io/bcam-courses/jjseb3/R-jjseb3.html>