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http://idaejin.github.io/bcam-courses/



1/32

GitHub: idaeiin

Outline

Introduction

Case study: mammography experience study



Introduction

- ▶ In the previous lecture we have looked at the logistic regression case when the dependent variable is **binary**, i.e. it has *only two categories*.
- Now we consider the case where we have nominal variables with more than two categories.
- We can extend the logistic regression model to do this. We call this extended method multinomial logistic regression.
- ► The basic principle of multinomial logistic regression is similar to that for (binomial) logistic regression, in that it is based on the *probability of membership of each category of the dependent variable*.
- That is why it can be viewed as a classification method that generalizes logistic regression to multi-class problems.
- ▶ It is also known as: Multiclass Log Reg, Multinomial Logit, or Maximum Entropy Classifier or Multinomial Choice Model.



3/32

Introduction

- ightharpoonup We assume that the categories of the outcome variable Y, are coded 0,1, or 2. In the three outcome category we need two logit functions.
- ▶ We have to decide which outcome categories to compare.
- ▶ The obvious extension is to use Y=0 as the **referent** or **baseline outcome** and form two logits comparing Y=1 and Y=2 to it.



4/32

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D.-J. Lee (BCAM) Intro to GLM's with R

Introduction

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- ▶ We have to decide which outcome categories to compare.
- ▶ The obvious extension is to use Y=0 as the **referent** or **baseline outcome** and form two logits comparing Y=1 and Y=2 to it.
- ► We denote the two logit functions as:

$$g_1(X) = \log \left(\frac{\Pr(Y = 1 | X)}{\Pr(Y = 0 | X)} \right) = \beta_{10} + \beta_{11}x_1 + \beta_{12}x_2 + \ldots + \beta_{1k}x_k = \mathbf{X}\boldsymbol{\beta}_1$$

$$g_2(X) = \log \left(\frac{\Pr(Y = 2 | X)}{\Pr(Y = 0 | X)} \right) = \beta_{10} + \beta_{11}x_1 + \beta_{12}x_2 + \ldots + \beta_{1k}x_k = \mathbf{X}\boldsymbol{\beta}_2$$



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Introduction

► It follows that the conditional probabilities of each outcome category given the covariate vector are

$$\Pr(Y = 0|X) = \frac{1}{1 + \exp(g_1(X) + g_2(X))}$$



5/32

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$$\begin{array}{lcl} \Pr(Y=0|X) & = & \frac{1}{1+\exp(g_1(X)+g_2(X))} \\ \Pr(Y=1|X) & = & \frac{e^{g_1(X)}}{1+\exp(g_1(X)+g_2(X))} \end{array}$$



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► A general expression for *K* categories is:

$$\Pr(Y = k|X) = \frac{\exp(g_k(X))}{\sum_{k=0}^K \exp(g_k(X))},$$

where $\Pr(Y=k|X)$ is the probability of belonging to group k, and $\beta_0=0$ and $g_0(X)=0$.

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Case study

- ▶ Data from a study undertaken to assess factors associated with women's knowledge, attitude and behaviour towards the benefits of mammography (Zapka et al. 1991).
- ▶ **Source:** Hosmer and Lemeshow (2000). *Applied Logistic Regression, 2nd edition.* John Wiley Sons Inc.,New York. Section 8.1.2, page 264.

See Multinom.R script

```
> mamexp=read.table("data/mammexp.txt",header=TRUE)
```

> names(mamexp)

```
[1] "me" "symp" "pb" "hist" "bse" "dect"
```



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Mammography Experience Study

▶ A data frame with 412 observations on the following 6 variables.

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 develop symptoms' A factor with levels 1=Strongly Agree, 2=Agree,
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Different alternatives in R

► There are different libraries in R with functions to fit a multinomial regression model.



Different alternatives in R

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```
> library(VGAM)
```

- > library(nnet)
- > library(mlogit)
- ▶ vglm() in VGAM: estimation is based on likelihood-inference
- ▶ multinom() in nnet: estimation is based on neural networks
- mlogit() in mlogit: uses a specific format, so data frame has to be reshaped



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Case study

▶ We begin by considering a model containing a single dichotomous covariate coded 0 or 1.



9/32

Case study

- We begin by considering a model containing a single dichotomous covariate coded 0 or 1.
- ▶ In this case, in the binary outcome model the estimated slope coefficient is identical to the $log\ OR$ obtained from the 2×2 cross-clasifying table for the outcome and the covariate.
- ► The cross-classification of Mammography Experience (me) by Family History of Breast Cander (hist) is

- ▶ me=0 is the reference outcome value
- $ightharpoonup \widehat{\mathsf{OR}}_1 = \frac{19/14}{85/220} = 3{,}51 \ \text{(i.e. } \exp(eta_1) \text{)}$
- $ightharpoonup \widehat{\mathsf{OR}}_2 = \frac{11/14}{63/220} = 2{,}74 \text{ (i.e. } \mathsf{exp}(\beta_2)\text{)}$



9/32

Fitting in R

vglm is a very large class of models that includes generalized linear models (GLM's) as a special case.

```
> vglm1 <- vglm(me ~ hist, family=multinomial(refLevel=1), data=mamexp)
```



10/32

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```
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```

► Usually we will use multinom() in library(nnet) because it includes more options (e.g. anova()) and it is very similar to glm() function. But in general, the results are quite similar.

```
Value Std.Error t value p value
Intercept.1 -0.9509794 0.1277115 -7.446312 9.598590e-14
Intercept.2 -1.2504803 0.1428926 -8.751189 2.111168e-18
hist1.1 1.2565310 0.3746633 3.353761 7.972127e-04
hist1.2 1.0093843 0.4275097 2.361079 1.822183e-02
```



multinom()

```
> mii1t.i1
Call:
multinom(formula = me ~ hist, data = mamexp)
Coefficients:
  (Intercept) hist1
 -0.9509794 1.256531
2 -1.2504803 1.009384
Residual Deviance: 792.3399
AIC: 800.3399
> coefficients(multi1) # We can extract the coefficients of the model
```



11/32

multinom()

- ► The output is similar to logistic regression
 - By default the reference level is the 1st level.
 - Now we have two groups of parameters, one for each logit (for me=0 as baseline):

```
\begin{array}{lll} \mbox{logit(me=1|hist)} & = & -0.95 + 1.25 \times \mbox{hist} \\ \mbox{logit(me=2|hist)} & = & -1.25 + 1.01 \times \mbox{hist} \end{array}
```

ightharpoonup p-values are not provided (but can be easily obtained) \Rightarrow we would need LRT

```
> exp(coefficients(multi1))
(Intercept) hist1
1 0.3863624 3.513213
2 0.2863672 2.743911
```

Note that the OR are the exponentials of the β's and they identical to the values obtained in the 2 × 2 cross-table

multinom()

▶ 95 % Confidence intervals for the OR are obtained as:

```
> exp(confint(multi1))
, , 1

2.5 % 97.5 %
(Intercept) 0.3008061 0.4962529
hist1 1.6857397 7.3218101
, , 2

2.5 % 97.5 %
(Intercept) 0.2164178 0.3789254
hist1 1.1870612 6.3425948
```

► We interpret each estimated **OR** and its corresponding CI's as if it came from a logistic regression model (in few seconds).



13/32

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multinom()

► In order to test if the predictor is significant, we use the anova() function, and the Likelihood Ratio Test (LRT).



14/32

multinom()

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► Hence the varible hist is significant



Interpretation of the model

► The interpretation of hist ("Family history of breast cancer") on the frequency of a mammography screening me is as follows

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Interpretation of the model

- ► The interpretation of hist ("Family history of breast cancer") on the frequency of a mammography screening me is as follows
 - ► The odds among women with a family history of breast cancer (HIST=Yes) having a mammogram Within a Year is 3,51 times greater than the odds among women without family history (HIST=No).

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 - st In other words, women with a history of breast cancer are 2.7 times as likely to have had a mammogram over one year ago than women without a family history of breast cancer.
- ► Thus, hist is a significant factor in use of mammography screening.
- ► If we include instead of HIST, the variable DECT (women opinion towards a mammography)?

polychotomous covariate

▶ We now consider the variable dect that has three levels 0=Not likely, 1=Somewhat likely and 2=Very likely



polychotomous covariate

▶ We now consider the variable dect that has three levels 0=Not likely, 1=Somewhat likely and 2=Very likely

Cross classification table of me by dect:

> xtabs(~me+dect,data=mamexp)

me=0 is the reference outcome value and dect=0 as the reference covariate values

$$ightharpoonup \widehat{OR}_1(1,0) = \frac{12 \times 13}{77 \times 1} = 2.03$$

$$ightharpoonup \widehat{OR}_1(2,1) = \frac{91 \times 13}{144 \times 1} = 8,22$$

$$\widehat{\mathsf{OR}}_2(1,0) = \frac{16 \times 13}{77 \times 4} = 0.68$$

$$ightharpoonup \widehat{OR}_2(2,1) = \frac{54 \times 13}{144 \times 4} = 1,22$$



16/32

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822

805.1980

```
> multi2 <- multinom(me~dect,data=mamexp)
> anova(multi0,multi2)

Likelihood ratio tests of Multinomial Models

Response: me
    Model Resid. df Resid. Dev Test Df LR stat. Pr(Chi)
```

dect 818 778.4011 1 vs 2 4 26.79694 2.184912e-05

▶ Thus, woman's opinion on the ability of a mammogram to detect a new case of breast cancer is significantly associated to her decision to have a mammogram.



17/32

```
> coef(multi2)
  (Intercept)     dect1     dect2
1     -2.565201  0.7062589  2.1062453
2     -1.178617 -0.3926042  0.1977986
> exp(coef(multi2))
     (Intercept)     dect1     dect2
1     0.07690374  2.0263961  8.217330
2     0.30770391  0.6752959  1.218717
```

```
> exp(confint(multi2))
, , 1
                2.5 % 97.5 %
(Intercept) 0.01005804 0.5880059
dect.1
           0.24245780 16.9360656
dect2
           1.05675131 63.8982038
, , 2
               2.5 % 97.5 %
(Intercept) 0.1003340 0.9436648
dect1
           0.1947759 2.3412783
dect2
           0.3807324 3.9010890
```

18/32

Examining the **OR** and the **CI**'s we see that the association is strongest when comparing the women who have had a mammogram with the last year me=1, to those who never had one, and comparing dect=2 to dect=0, i.e. $OR_1(2,1) = 8,22$.

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18/32

Examining the **OR** and the **CI**'s we see that the association is strongest when comparing the women who have had a mammogram with the last year me=1, to those who never had one, and comparing dect=2 to dect=0, i.e. $OR_1(2,1) = 8,22$.

Interpretation: The odds of having a mammogram within the last year among the women who feel that a mammogram is very likely to detect a new case of breast cancer is 8,22 times larger than the odds among women who feel that it is no likely.

- ► If the predictor is continuous, it will be modelled as linear in the logit, and will have a single parameter for each logit function
- ► This coefficient, when exponentiated, gives the estimated odds ratio (OR) for a change of one unit in the continuous variable.



19/32

Variables selection

- ▶ The steps to determine which covariates should be included in the model are similar to those in the logistic regression case.
- ▶ Let us consider the rest of variables in the study:

univariate/multivariate fits

```
> multi3 <- multinom(me~symp,data=mamexp); anova(multi0,multi3)</pre>
> multi4 <- multinom(me~bse,data=mamexp); anova(multi0,multi4)
```

- > multi5 <- multinom(me~pb,data=mamexp); anova(multi0,multi5)</pre>
- > multi6 <- multinom(me~symp+pb+bse+hist+dect,data=mamexp)</pre>



20/32

Variables selection

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- ▶ Let us consider the rest of variables in the study:

univariate/multivariate fits

```
> multi3 <- multinom(me~symp,data=mamexp); anova(multi0,multi3)
> multi4 <- multinom(me~bse,data=mamexp); anova(multi0,multi4)
> multi5 <- multinom(me~pb,data=mamexp); anova(multi0,multi5)
> multi6 <- multinom(me~symp+pb+bse+hist+dect,data=mamexp)</pre>
```

vglm

```
> vglm6 <- vglm(me ~ symp+pb+bse+hist+dect,
+ family=multinomial(refLevel=1), data=mamexp)</pre>
```



▶ vglm function computes Wald statistics and p-values, summary(vglm6) gives:

```
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept):1 -2.99875 1.53905 -1.948 0.05136 .
(Intercept):2 -0.98609 1.11184 -0.887 0.37513
            0.11004 0.92273 0.119 0.90508
symp2:1
       -0.29008 0.64406 -0.450 0.65243
symp2:2
symp3:1 1.92471 0.77757 2.475 0.01331 *
symp3:2 0.81731 0.53979 1.514 0.12999
symp4:1 2.45699 0.77530 3.169 0.00153 **
symp4:2 1.13224 0.54767 2.067 0.03870 *
pb:1
         -0.21944 0.07551 -2.906 0.00366 **
pb:2
          -0.14821 0.07637 -1.941 0.05230 .
hist1:1 1.36624 0.43752 3.123 0.00179 **
hist1:2 1.06544 0.45940
                             2.319 0.02038 *
bse1:1
         1.29167 0.52988 2.438 0.01478 *
           1.05214 0.51499
                             2.043 0.04105 *
bse1:2
dect1:1
            0.01702 1.16169 0.015 0.98831
dect1:2
      -0.92439 0.71375 -1.295 0.19528
dect2:1 0.90414 1.12661 0.803 0.42225
dect2:2
      -0.69053
                    0.68712 -1.005 0.31491
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

► It suggests that with the possible exception of dect, each of the variables contribute to the model

► Let us look at the coefficients for the variable symp

- symp2:1 and symp2:2 estimates the log odds for Agree VS Strongly Agree (the reference value). This suggests that both categories are similar since the Wald statistics is significant.
- ► Then is symp significant?



22/32

► We can use a simpler model by coding symp into two levels: 0=Agree/Strongly agree and 1=Disagree or Strongly disagree

```
> mamexp$symp01<-mamexp$symp # create a new variable symp01
> require(car)
> mamexp$symp01<-recode(mamexp$symp,"1=0;2=0;3=1;4=1")</pre>
```

► Fit the model with symp01 a look at the significance of the coefficients

```
> vglm6a <- vglm(me ~ symp01+pb+bse+hist+dect,
+ family=multinomial(refLevel=1), data=mamexp)</pre>
```



23/32

Variables selection

▶ Let us compare multi6 with a model that excludes dect, and apply a LRT:

The LRT indicates that dect is not significant and should be removed from the model.



24/32

Variables selection

▶ Let us compare multi6 with a model that excludes dect, and apply a LRT:

The LRT indicates that dect is not significant and should be removed from the model.

But can we try something such as re-group the categories of dect?

► Let us explore collapsing dect into two categories: 0=NotLikely/SomewhatLikely and 1=VeryLikely



 Let us explore collapsing dect into two categories: 0=NotLikely/SomewhatLikely and 1=VeryLikely

```
> table(dect)
dect
    0     1     2
    18 105 289
> mamexp$dect01 <- mamexp$dect
> mamexp$dect01 <- recode(mamexp$dect01, "0=0; 1=0; 2=1")</pre>
```



► Let us explore collapsing dect into two categories: 0=NotLikely/SomewhatLikely and 1=VeryLikely

```
> table(dect)
dect
    0     1     2
18 105 289
> mamexp$dect01 <- mamexp$dect
> mamexp$dect01 <- recode(mamexp$dect01,"0=0; 1=0; 2=1")</pre>
```

 Fit the model with symp01 and dect01 a look at the significance of the coefficients

```
> vglm7a <- vglm(me ~ symp01+pb+bse+hist+dect01,
+ family=multinomial(refLevel=1), data=mamexp)</pre>
```

Can we use dect01 in our model?



Variables selection

▶ Now, we consider the variable pb (continuous and discrete) which is the sum of five scaled responses, each on a four point scale. A low value is indicative of a woman with strong agreement with the benefits of mammography.

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 $^{^1\}mathrm{We}$ cannot use the function gam as we did before, as the multinomial family is not available in mgcv

Variables selection

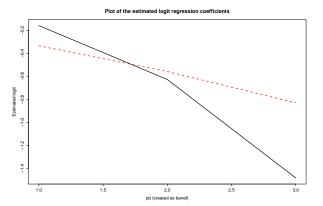
- Now, we consider the variable pb (continuous and discrete) which is the sum of five scaled responses, each on a four point scale. A low value is indicative of a woman with strong agreement with the benefits of mammography.
- ▶ Now, we want to know if the relationship of pb and the logit is linear or not¹
- ▶ Let us recode the variable pb, range of pb is [5,17], into

```
> mamexp$benef <- mamexp$pb
> mamexp$benef[mamexp$benef<=5] = 0
> mamexp$benef[mamexp$benef>5 & mamexp$benef <=7] = 1
> mamexp$benef[mamexp$benef>7 & mamexp$benef <=9] = 2
> mamexp$benef[mamexp$benef>9] = 3
> mamexp$benef <- factor(mamexp$benef)</pre>
```

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 $^{^1\}mbox{We}$ cannot use the function gam as we did before, as the multinomial family is not available in \mbox{mgcv}

- ▶ benef is a factor variable created to evaluate the scale of the effect of pb
- ► We plot the estimated logits with the design variable benef
- ► See Multinom.R script



► The plot shows evidence of linearity in the logits in pb, and hence we can leave pb as linear continuous predictor.

Variables selection

- As in the case of logistic regression, the next step is to assess the need to include interaction terms in the model.
- ▶ In this case, none of the possible interactions are significant.
- ► Our preferred model would be vglm7a or equivalently with multinom()

2 0.16139719 2.895563 3.087680 2.601307 0.8569992 1.120929



28/32

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```
> multi9 <- multinom(me~hist+symp01+bse+pb+dect01,data=mamexp)
> exp(coef(multi9))
(Intercept) hist1 symp011 bse1 pb dect011
1 0.07252958 3.705677 8.123425 3.445302 0.7792099 2.423432
2 0.16139719 2.895563 3.087680 2.601307 0.8569992 1.120929
```

► How do you interpret the OR?



interpretation

```
> exp(coef(multi9))
```

```
(Intercept) hist1 symp011 bse1 pb dect011
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29/32

interpretation

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```

- ► The estimated OR's show that generally, it increases from "1=Within a year" to "2=Over a year" wrt to the baseline category "0=Never", except on the variable pb.
- ► The coefficients of the variable pb are negative (and hence OR<1). It suggests that larger values indicate less belief in the benefit of mammography screening.
- ► Hence, the OR for pb reflect that "less belief" is significantly associated with less frequent use.

interpretation

► The function effect in library(effects) construct terms effects for a multinomial model

```
> library(effects)
> plot(effect("symp01",multi9))
> plot(effect("hist",multi9))
> plot(effect("bse",multi9))
> plot(effect("dect01",multi9))
> plot(effect("pb",multi9),style="stacked")
> plot(effect("pb",multi9),style="lines")
```



30/32

interpretation

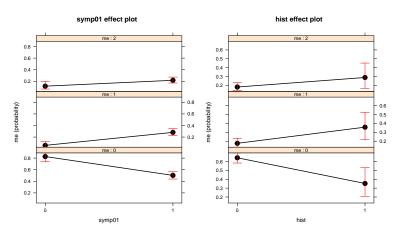
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```



30/32

Effects plots

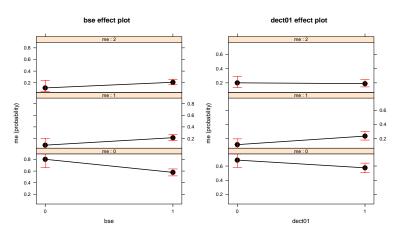




31/32

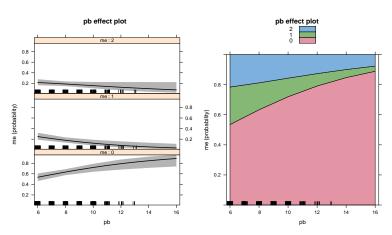
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Effects plots





Effects plots





Conclusions

- ► The real challenge when fitting a multinomial logistic regression model is the fact that there are multiple odds ratios for each model covariate.
- ▶ This complicates the interpretation of the model.
- ► However, using a multinomial outcome provide more complete description of the process being studied.
- ► For instance, in the mammography experience study, if we had combined the outcome into a binary response (e.g.: ever VS never) we would have missed the gradation in odds ratios.
- ► In practice, one should not pool the outcome categories unless the estimated coefficients in the logits are not significantly different from each other.
- ► In summary, fitting and interpreting the results from a multinomial logistic regression model follows the same paradigm as in the binary logistic model.

