

# STAT5040: Fourier Analysis of Time Series Ch8

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# The Spectrum

# Introduction

- We focus on the mathematical aspect in previous chapters, e.g.,
  - fitting sinusoids of known/unknown frequency, and
  - the fast Fourier transform (FFT) algorithms.
- We now discuss the statistical aspect with some real data sets.
  - How to do spectral analysis of time series?
  - How to deal with some practical issues such as change in scale?
  - How to interpret the results from spectral analysis?
- The data is available [here](#).
  - The link is not up-to-date in the book.
- My code is available [here](#).
  - Modern libraries like xts are used.

# Periodogram Analysis

# Definitions recall

- Discrete Fourier transform (Section 5.1):

$$d(f_j) = \frac{1}{n} \sum_{t=0}^{n-1} x_t \exp(-2\pi i f_j t), \quad j = 0, 1, \dots, n-1.$$

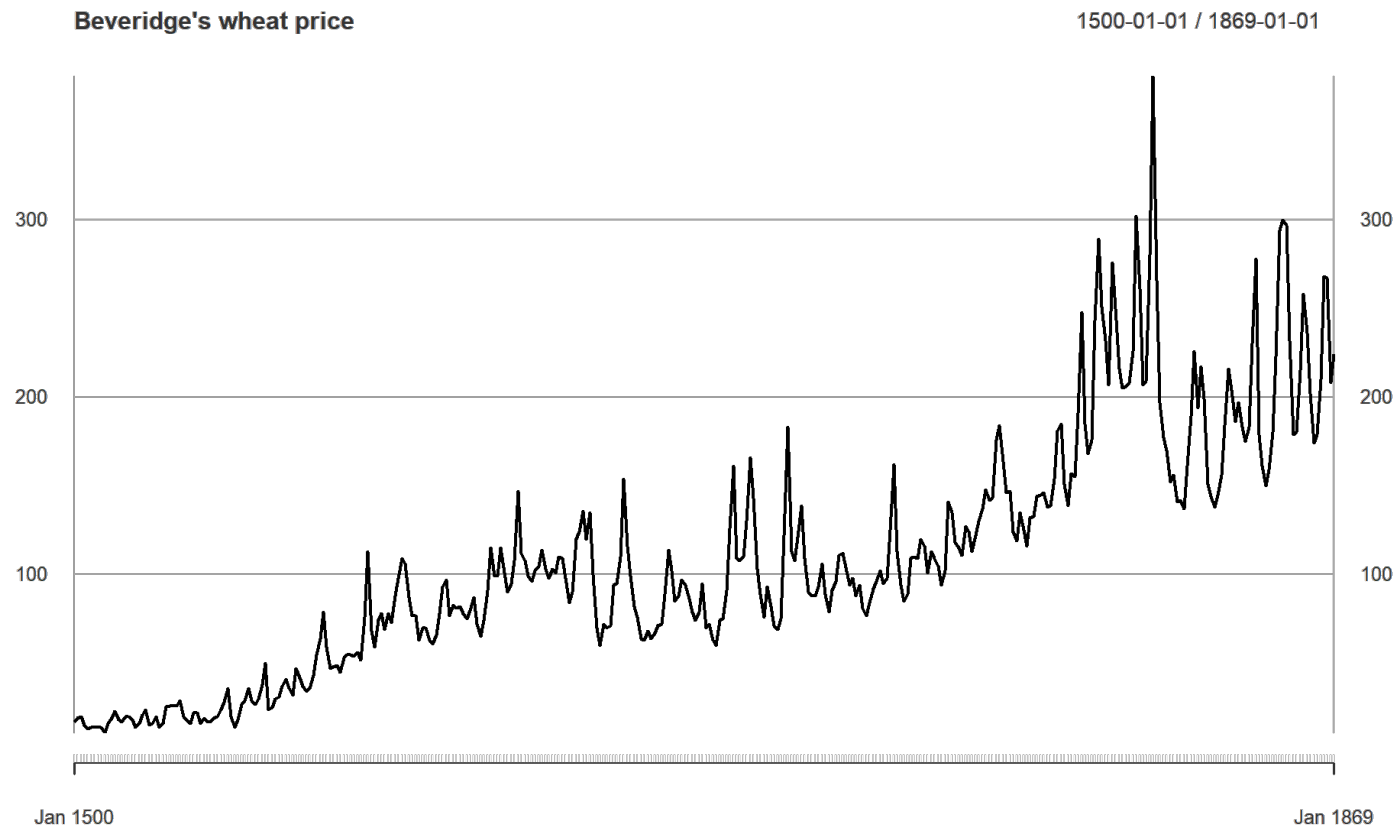
- Equivalent to fitting sinusoids (Section 4.1)
- Periodogram (Section 6.1):

$$I(f) = n|d(f)|^2.$$

- An estimate of the spectrum/spectral density

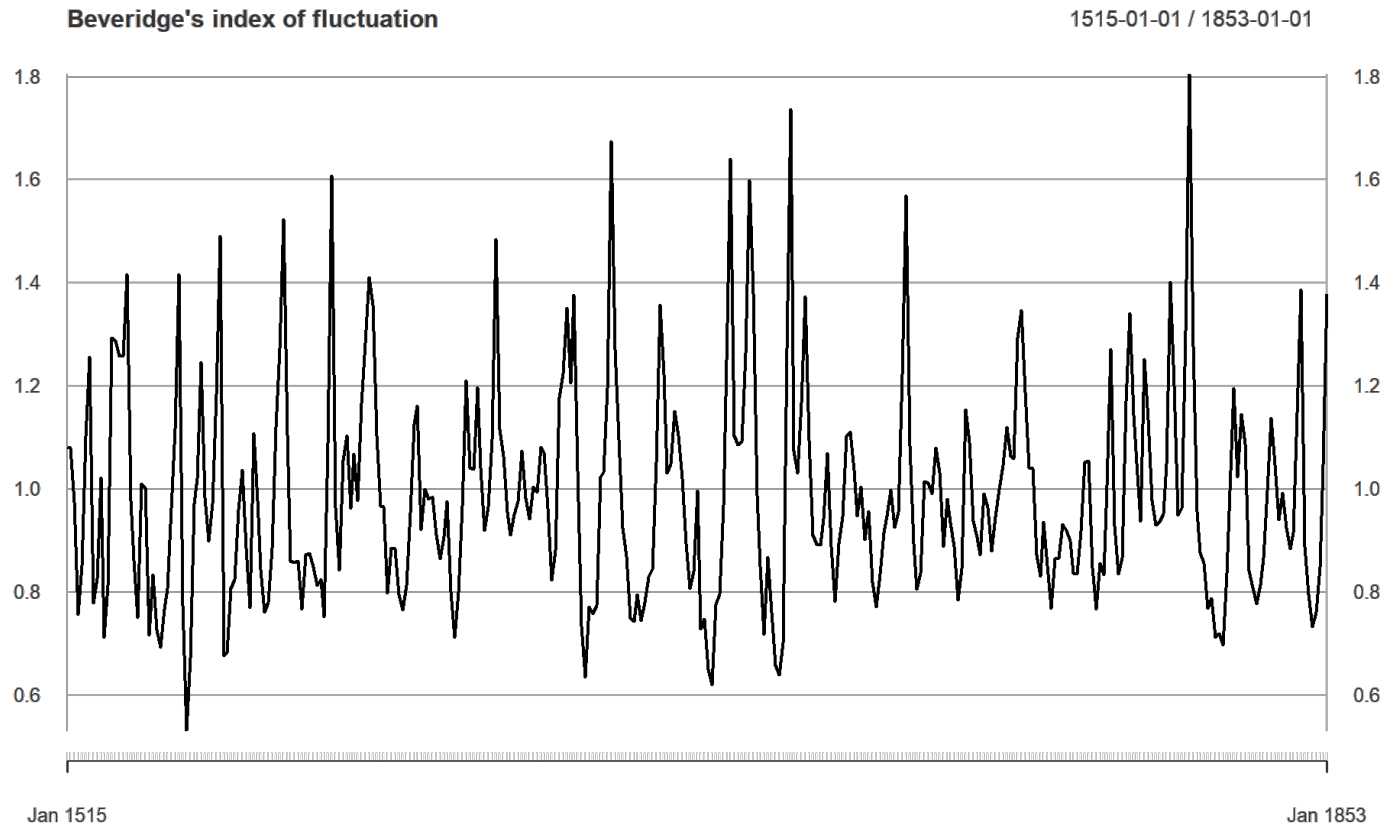
# Problem: change in scale

- No sinusoids can match oscillations that grow in amplitude
  - Possible cause: inflation?



# Beveridge: “index of fluctuation”

- Each value is divided by the average of 31 centered adjacent values
  - Idea: smoothing



# Motivation: unobserved components model

- Proposed by Harvey (1989)
- Bloomfield only considered the trend  $T_t$  and irregular component  $I_t$ 
  - Multiplicative model:  $x_t = T_t I_t$
  - Additive model:  $x_t = T_t + I_t$
- Interpretation for wheat price
  - $T_t$ : driven by long-run economic forces such as inflation
  - $I_t$ : caused by short-run effects such as changes in supply from year to year
  - $T_t$  is an unwanted complication in the analysis and approximated by 31-year moving average
  - $I_t$  is then estimated by the index of fluctuation
  - Closely related to filtering in Section 7.2



# Problem: spurious periodicities

- Beveridge (1921) gave periodogram ordinates based on index of fluctuation
  - It solved the change in scale problem in analyzing periodicity
  - However, Slutsky pointed out that operations involving linear filtering might lead to spurious periodicities in 1927
  - Intuitively, this means that transformation may distort the original periodicity
    - I try to replicate his estimates but it seems that additional cleanings are necessary (p.139)
    - As the original paper is not found online, I choose to discuss the idea only
- Under multiplicative model, we can correct for the transfer function to mitigate spurious periodicities
  - As also shown by Granger and Hughes (1971)
  - For index of fluctuation, the transfer function is  $D_{31}(f)$ , a Dirichlet kernel (Section 2.2)
  - After correction, the peak changes as compared with Beveridge (1921)

# (Part of) Beveridge's periodogram

Original Analysis		After Correction		
Period (years)	Periodogram Ordinate	Period (years)	Periodogram Ordinate	Previous Rank
15	47.28	15	50.48	1
11	40.93	11	46.51	2
20	32.44	36	36.92	7
17	29.35	13	36.41	5
13	27.81	12	25.99	9
24	26.48	17	24.59	4
36	26.27	35	36.92	10
16	20.14	20	22.39	3
12	20.11	16	18.90	8
35	17.52	24	18.51	6
18	17.26	34	13.50	15
25	14.95	18	13.23	11
6	12.29	7	12.12	16
8	12.05	6	11.53	13
34	11.04	8	11.31	14
7	10.43	25	10.81	12
30	7.86	30	7.38	17
23	7.54	23	5.15	18
22	7.50	22	5.06	19
21	6.33	10	5.06	21
10	5.39	31	5.04	23

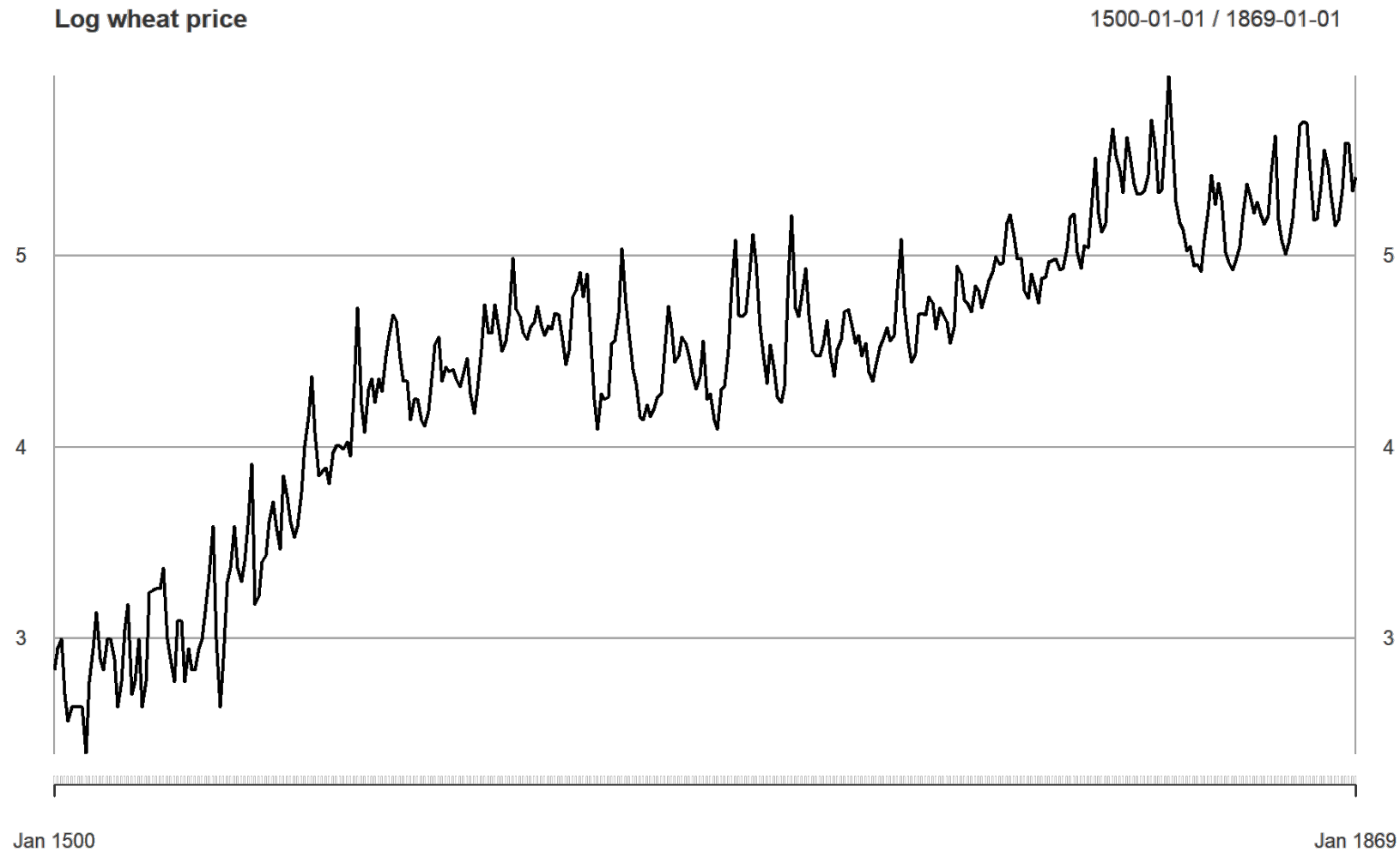
# Logarithmic transformation

- Alternatively, we can do logarithmic transformation under multiplicative model

$$\ln x_t = \ln T_t + \ln I_t$$

- Interpretation
  - $\ln T_t$ : the typical value of  $x_t$  in the neighborhood of  $t$
  - $I_t$ : a dimensionless quantity close to 1 so that  $\ln I_t$  fluctuates around 0
- Vs index of fluctuation
  - Nonsinusoidal behavior of a series introduce structure into its Fourier transform
  - However, they are not revealed by the periodogram (Section 6.5)
  - Logarithms can reduce this behavior such as spikiness of the peaks
    - Reasonable as we are interested in  $I_t$  but not the extreme behaviors

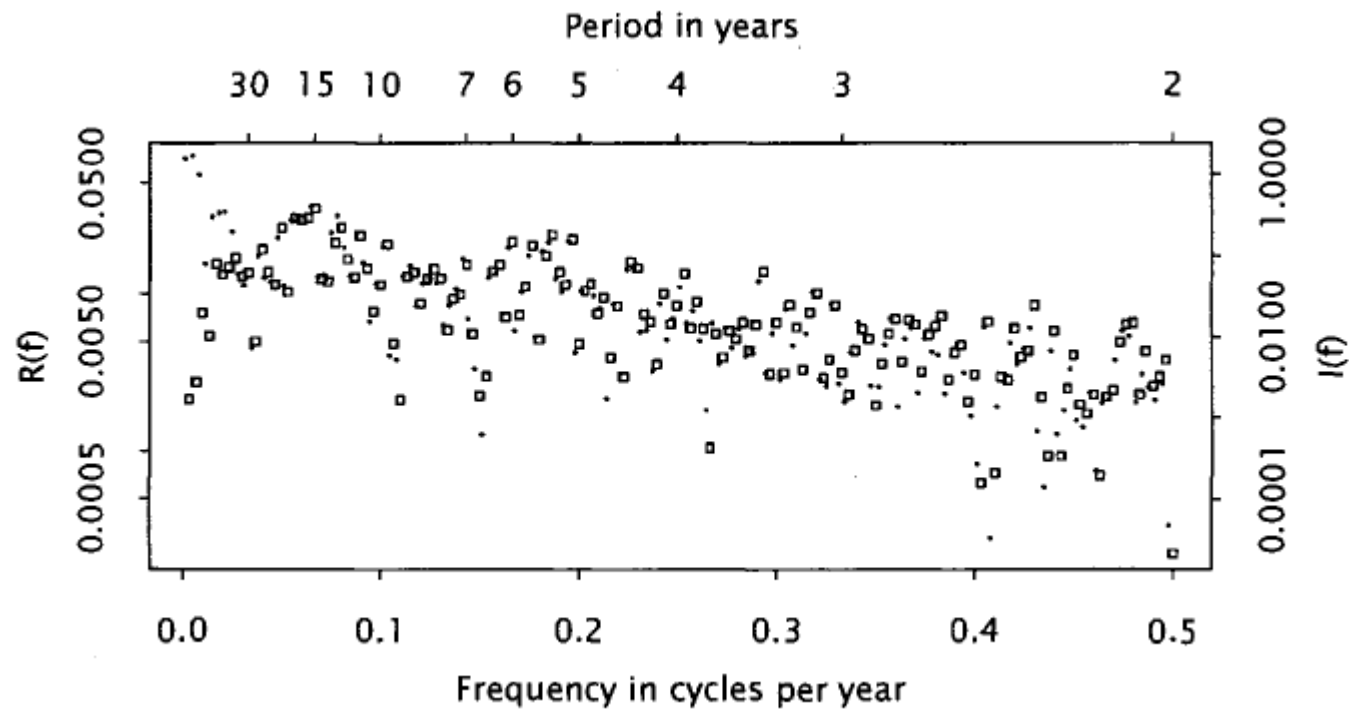
# Logarithmic transformation



# Data cleaning

- Beveridge (1921) argued that the early part of the series was unreliable due to fewer sources
- The later part was of a different nature due to economic changes in the 19th century
  - I have not looked into wheat market history but it sounds reasonable
  - We should also be aware of the underlying structure/data quality when we perform statistical analysis
- After cleaning, the periodogram of logarithms and index of fluctuation are similar

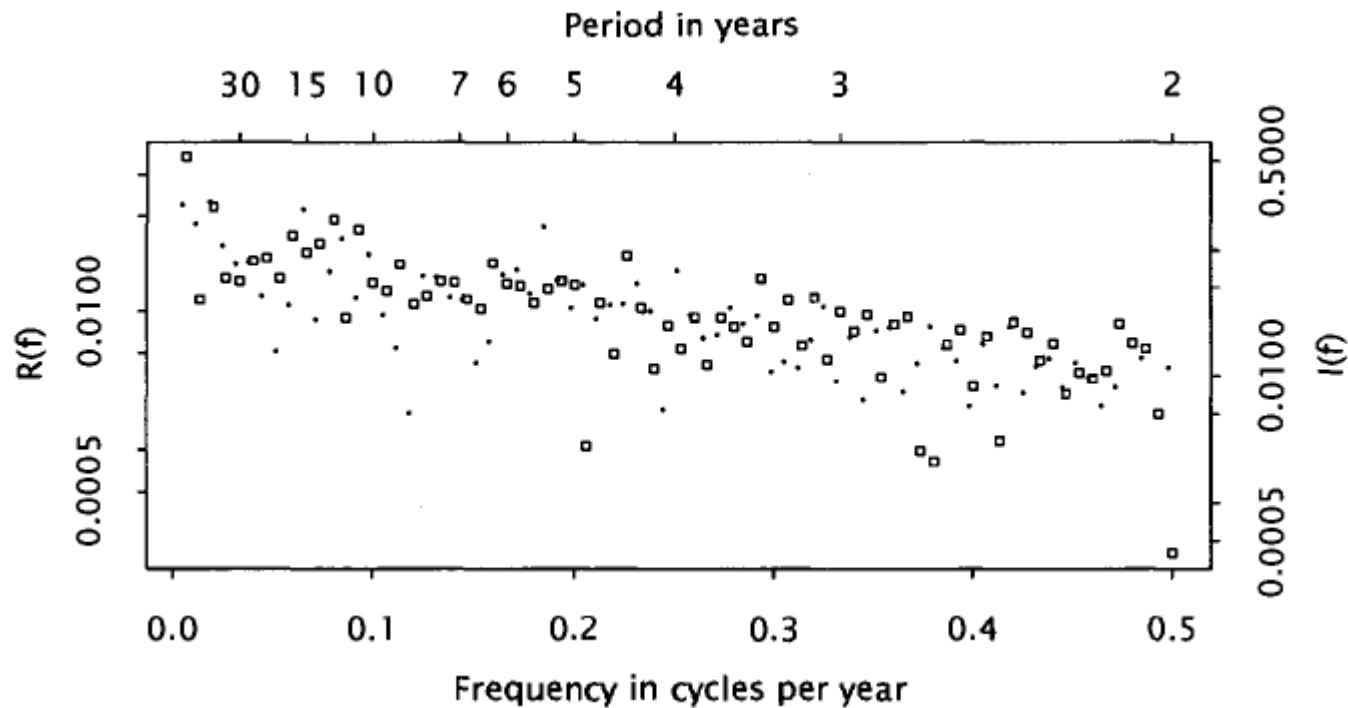
# Data cleaning



Periodogram of logarithms (dots) and index of fluctuation (squares)

# Analysis of segments: idea

- Idea: if there is periodicity in a series, it should hold for segments of the same series
- Beveridge (1921) also gave some terms from the periodogram of two halves of the series



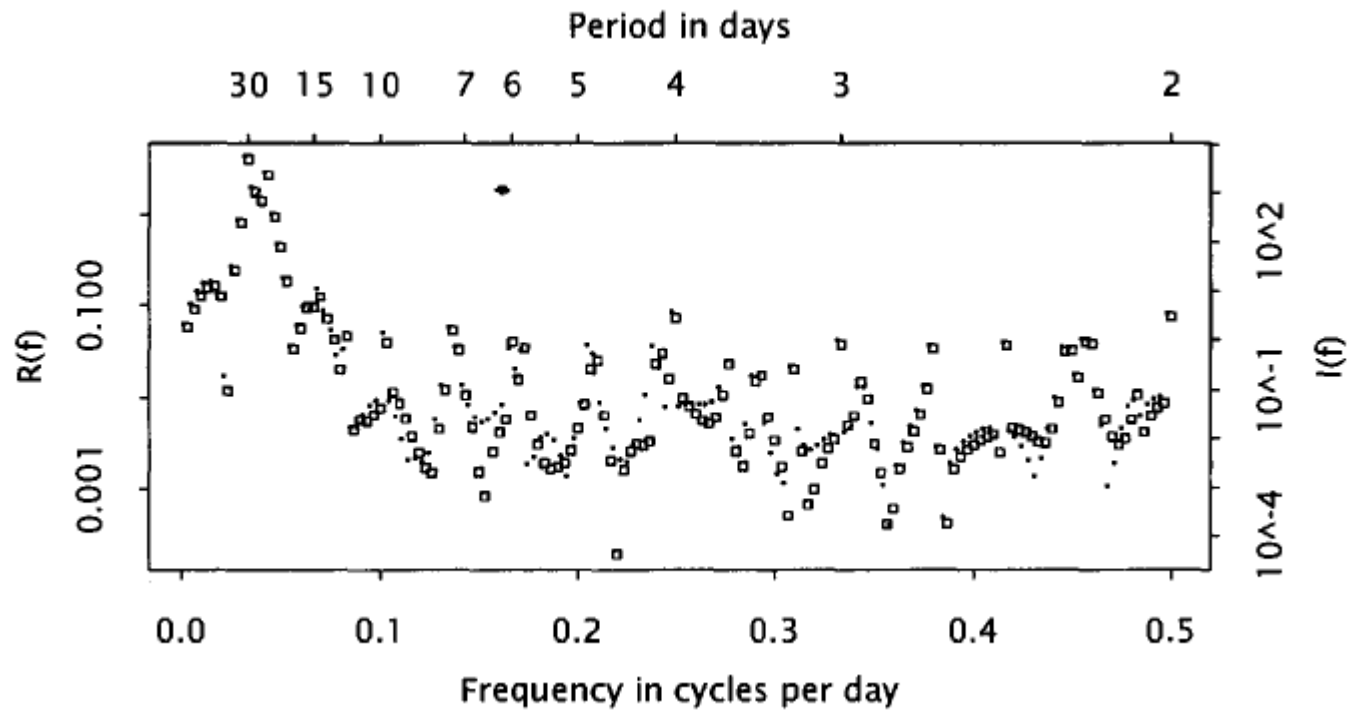
Periodogram of logarithms of two halves (1545-1694=squares, 1695-1844=dots)

# Analysis of segments: wheat price

- The two periodograms have the same general shape
  - Large at lowest frequencies and show a broad peak between 0.06 and 0.10 cycles per year
  - Then show a gentle decline over the rest of the periodogram with small fluctuations
- However, the fine structures are quite unrelated
  - A local peak in one is just as likely to be matched by a local trough as a local peak in the other
- These shows that the fine strucuture is not repeated from one segment to the next
  - But the broad features show a statistical regularity or consistency across segments
    - I think these terms can be confusing in the literature
- Thus the fine structure of these two periodograms is not characteristic of the series as a whole
  - Broad features do not appear to vary in this way and may be characteristic of the whole series
  - In contrast, the periodograms of the variable star data have similar fine structures



# Analysis of segments: variable stars



Periodogram of two halves of the variable star series

# Analysis of segments: conclusion

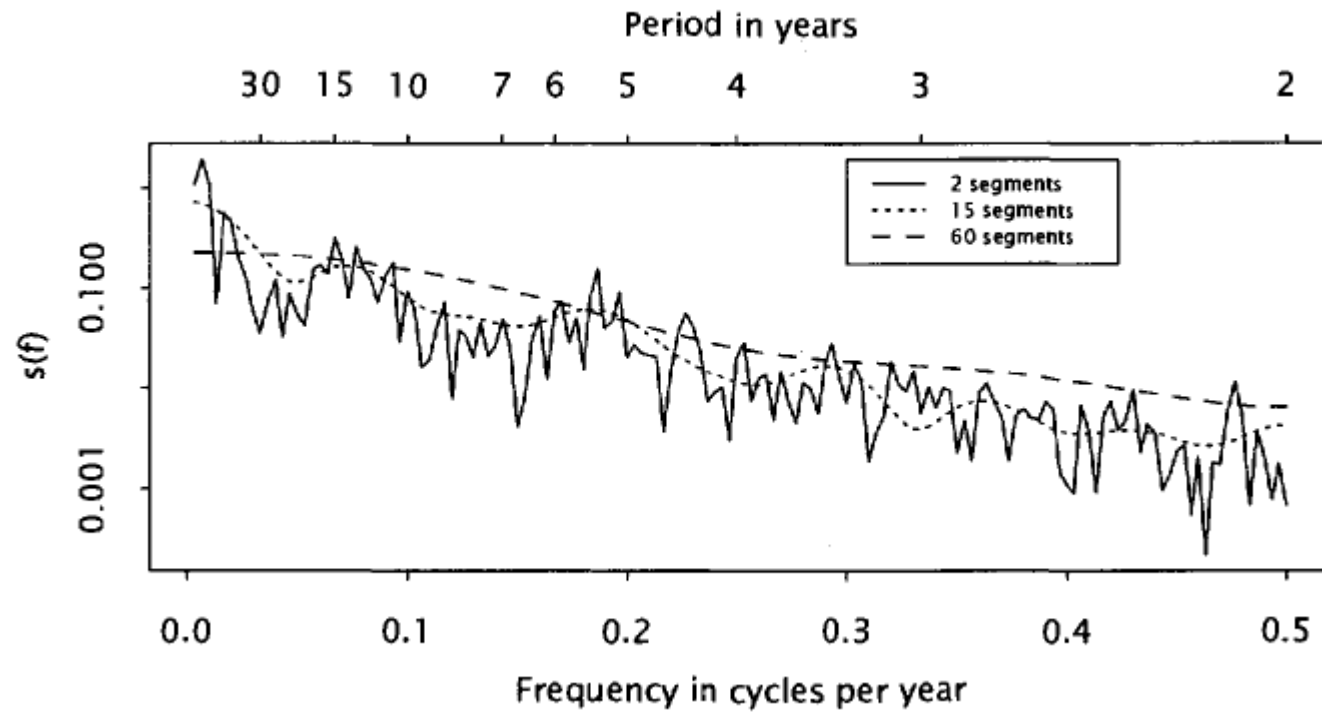
- Bloomfield is trying to motivate the concept of spectrum here
  - A raw periodogram may not be useful
    - For variable star, it is directly useful
    - For wheat price, the fine structure should be suppressed to focus on the broad behavior
  - This suggests a smoothed periodogram
    - So that periodograms of segments may be regarded as the same underlying smooth curve
    - Such smooth curve is called the spectrum or spectral density of the series
    - Spectrum exists for many time series models (Chapter 9)

# Spectrum Estimation

# Bartlett's method

- Suggested by Bartlett (1948)
  1. Split the series into  $k$  non-overlapping segments of similar length
  2. For each segment, compute its periodogram
  3. Average the result of the periodograms above for the  $k$  segments
- Welch's method: allow overlapping
  - It can reduce noise in exchange for reducing the frequency resolution
  - Often desirable in finite sample
- Average of logarithms or logarithms of averages can both be used
  - The latter is preferred since it puts more weight on larger values
  - Small values are the most susceptible to perturbations such as leakage from other frequencies

# Bartlett's method



Average segment periodograms for the logarithms of the wheat prices, 1545-1844

# Bartlett spectrum estimate

- To understand the effect of Bartlett's method, we try to find the "periodogram" that it is producing
- For simplicity, Bloomfield assumed  $x_t$  are deviations around 0
  - If they are not, we can do demean and arrive at a similar result
- Now note that

$$\begin{aligned} I(f) &= n|d(f)|^2 = n \cdot d(f) \cdot \overline{d(f)} \\ &= \frac{1}{n} \sum_t \sum_{t'} x_t x_{t'} \exp \{ -2\pi i f(t - t') \}. \end{aligned}$$

- As  $t - t' \in [-n + 1, n - 1] \cap \mathbb{Z}$ , we can rewrite the above as

$$\begin{aligned} I(f) &= \frac{1}{n} \sum_{r=-n+1}^{n-1} \sum_{t-t'=r} x_t x_{t'} \exp(-2\pi i f r) \\ &= \frac{1}{n} \sum_{r=-n+1}^{n-1} \exp(-2\pi i f r) \sum_{t-t'=r} x_t x_{t'}. \end{aligned}$$

# Bartlett spectrum estimate

- We can see that the periodogram is itself a Fourier series with  $n^{-1} \sum_{t-t'=r} x_t x_{t'}$  as the coefficients
- Some manipulation yields

$$I(f) = \sum_{|r| < n} c_r \exp(-2\pi i f r) \quad \text{with} \quad c_r = \begin{cases} n^{-1} \sum_{t=r}^{n-1} x_t x_{t-r}, & r \geq 0; \\ c_{-r}, & r < 0. \end{cases}$$

- If you are familiar with time series,  $c_r$  is the sample autocovariance of  $\{x_t\}$  at lag  $r$ 
  - Estimating the autocovariance structure is equivalent to estimating the spectrum
  - This also explains long run variance is the normalized spectrum at frequency 0
- In light of the symmetry of the autocovariances, we can also write

$$I(f) = c_0 + \sum_{r=1}^{n-1} c_r \cos(2\pi f r).$$

# Bartlett spectrum estimate

- Suppose the data are divided into  $k$  non-overlapping segments of length  $m = n/k$ 
  - $I_j(f)$ : the periodogram of the  $j$ -th segment
  - $c_{j,r}$ : the autocovariance of the  $j$ -th segment at lag  $r$
- The average of these periodogram is

$$\begin{aligned}\hat{s}(f) &= \frac{1}{k} \sum_{j=1}^k I_j(f) = \sum_{|r| < m} \left( \frac{1}{k} \sum_{j=1}^k c_{j,r} \right) \exp(-2\pi i f r) \\ &= \sum_{|r| < m} \left( 1 - \frac{|r|}{m} \right) \frac{1}{k(m - |r|)} \sum_{j=1}^k m c_{j,r} \exp(-2\pi i f r).\end{aligned}$$

- Now  $\sum_{j=1}^k m c_{j,r}$  is like  $n c_r$ , a sum of products of the form  $x_t x_{t+r}$ 
  - Except that the term is included only if  $x_t$  and  $x_{t+r}$  fall into the same segment
  - So  $\{k(m - |r|)\}^{-1} \sum_{j=1}^k m c_{j,r}$  is the average of these products
  - It makes sense to replace with  $n c_r / (n - |r|)$  (whole series version)



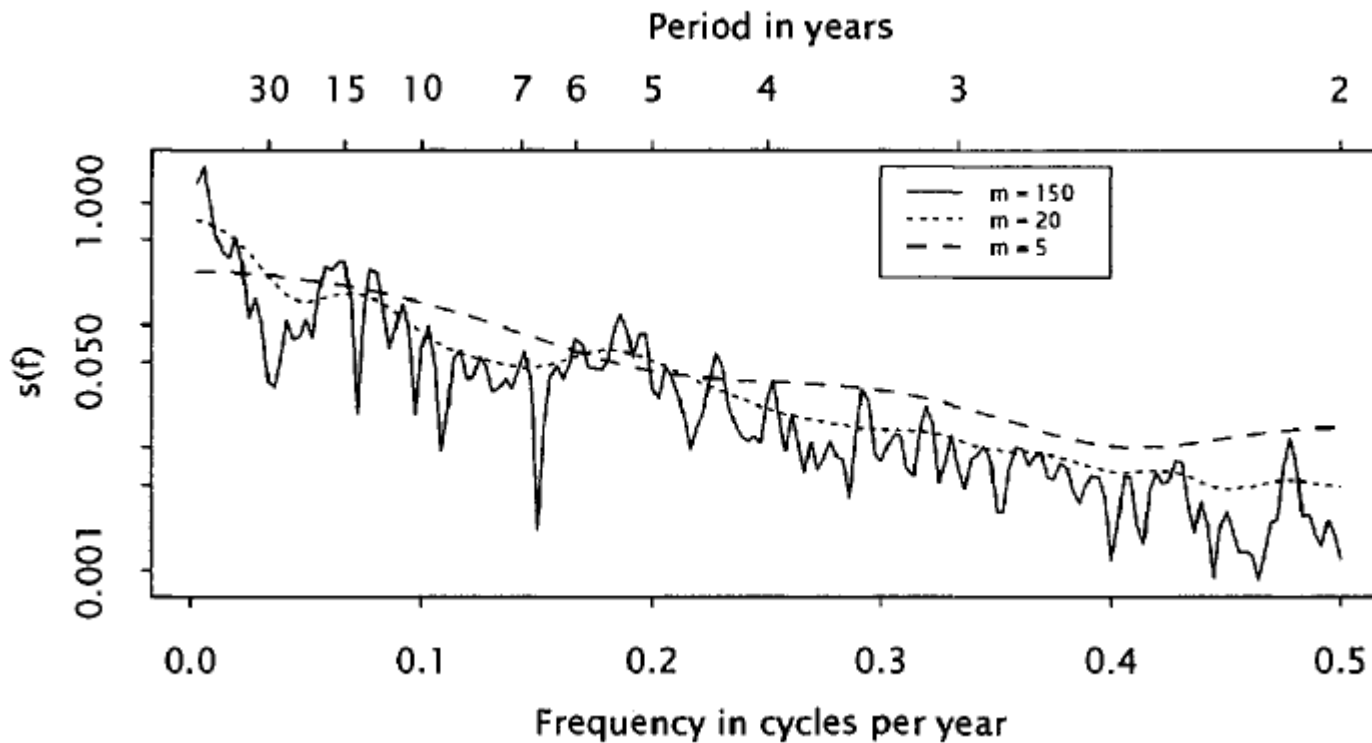
# Bartlett spectrum estimate

- Replacing  $\{k(m - |r|)\}^{-1} \sum_{j=1}^k m c_{j,r}$  with  $n c_r / (n - |r|)$ , the modified version is

$$\hat{s}_B(f) = \sum_{|r| < m} \frac{1 - |r|/m}{1 - |r|/n} c_r \exp(-2\pi i f r) = \sum_{|r| < m} w_r c_r \exp(-2\pi i f r),$$

- where  $w_r = (1 - |r|/m)/(1 - |r|/n)$ .
- $\hat{s}_B(f)$  is known as the Bartlett spectrum estimate
- $w_r$  is the weight at lag  $r$
- Vs periodogram
  - All terms with  $|r| \geq m$  will be omitted
  - The remaining terms are progressively reduced in magnitude by the weight  $w_r$
  - Both effects make  $\hat{s}_B(f)$  smoother than the periodogram
  - Varying the truncation point  $m$  provides control over the degree of smoothness
    - $m$  is also called the bandwidth
  - However,  $\hat{s}_B(f)$  is not guaranteed to be positive unlike  $I(f)$

# Bartlett spectrum estimate



Bartlett spectrum estimates for the logarithms of the wheat prices, 1545-1844

# Smoothing the periodogram

- The Bartlett spectrum estimate motivates smoothing the periodogram with other data windows  $w_r$ 
  - A significant property is that  $w_r$  decays from 1 when  $r = 0$  to 0 when  $r = \pm m$ 
    - This is common but not necessary, at least for long run variance estimation
    - See, e.g., Andrews (1991)
  - Anderson (1994) lists the most commonly used windows and their properties
  - Bartlett window is the only window that decays linearly and rarely used in practice
    - I think Bartlett window is quite popular, at least for long run variance estimation as well
    - It is also computationally efficient for spectrum estimation. See Xiao and Wu (2011)
- The smoothing strategy is attractive when the truncation point  $m$  is small
  - Only  $m$  autocovariances need to be computed
  - If  $m$  is large, the autocovariances can be computed efficiently using FFT and its inverse

# Computing the autocovariances

- From previous derivation (p.144), we know that the periodogram is itself a Fourier series
- This raises the possibility of using FFT to obtain autocovariance estimates
- When the periodogram is evaluated at a Fourier frequency  $f_j = j/n$ ,

$$I(f_j) = \sum_{r=0}^{n-1} (c_r + c_{n-r}) \exp(-2\pi i f_j r),$$

- provided that  $c_r$  is defined to be 0 for  $|r| \geq n$ .
- Therefore, we have

$$c_r + c_{n-r} = \frac{1}{n} \sum_{j=0}^{n-1} I(f_j) \exp(2\pi i f_j r),$$

- which may be computed using inverse FFT.
- However,  $c_r$  is not symmetric with  $c_{n-r}$
- This computation is only fine for small  $r$  as giving approximations of  $c_r$

# Computing the autocovariances

- The problem of  $c_r + c_{n-r}$  can be solved by using a finer grid  $f'_j = j/n'$  for some  $n' > n$
- We can pad the data with a block of  $n' - n$  zeros. Then

$$I(f'_j) = \sum_{r=0}^{n-1} (c_r + c_{n'-r}) \exp(-2\pi i f'_j r).$$

- Inversion now gives

$$c_r + c_{n'-r} = \frac{1}{n'} \sum_{j=0}^{n'-1} I(f'_j) \exp(2\pi i f'_j r),$$

- which yields  $c_0, c_1, \dots, c_{n'-n}$  exactly.
- If  $n' = 2n - 1$ , all autocovariance estimates are obtained.
  - Note that  $I(f'_j)$  only sum to  $n - 1$
  - Probably why the zeros may not contaminate the estimates
- FFT costs  $O(n \log n)$  times, which is faster than brute force  $O(n^2)$

# Representation of a spectrum estimate

- Suppose a spectrum estimate is given by  $\hat{s}(f) = \sum_{|r| < n} w_r c_r \exp(-2\pi i f r)$ ,
  - where a truncation point  $m$  is no longer assumed.
- By the integral inversion formula (p.40), we have  $c_r = \int_0^1 I(f) \exp(2\pi i f r) df$ .
  - Then  $\hat{s}(f) = \int_0^1 W_n(f - f') I(f') df'$ ,
  - where  $W_n(f) = \sum_{|r| < n} w_r \exp(-2\pi i f r)$ .
- Therefore, any  $\hat{s}(f)$  of the above form may be written as an integral average of  $I(f)$ 
  - The  $W_n(f)$  is called the spectral window associated with the spectral estimate
    - There is one-to-one relationship between  $W_n(f)$  and  $w_r$

# Spectral window

- The spectral window of the Bartlett estimate has no closed form

- The window is  $w_r = (1 - |r|/m)/(1 - |r|/n)$

- Modified Bartlett estimate

- The window is  $w_r = (1 - |r|/m)\mathbb{I}_{|r| < m}$

- The corresponding spectral window is

$$W_n(f) = \sum_{|r| < m} \left(1 - \frac{|r|}{m}\right) \exp(-2\pi i f r) = m D_m(f)^2,$$

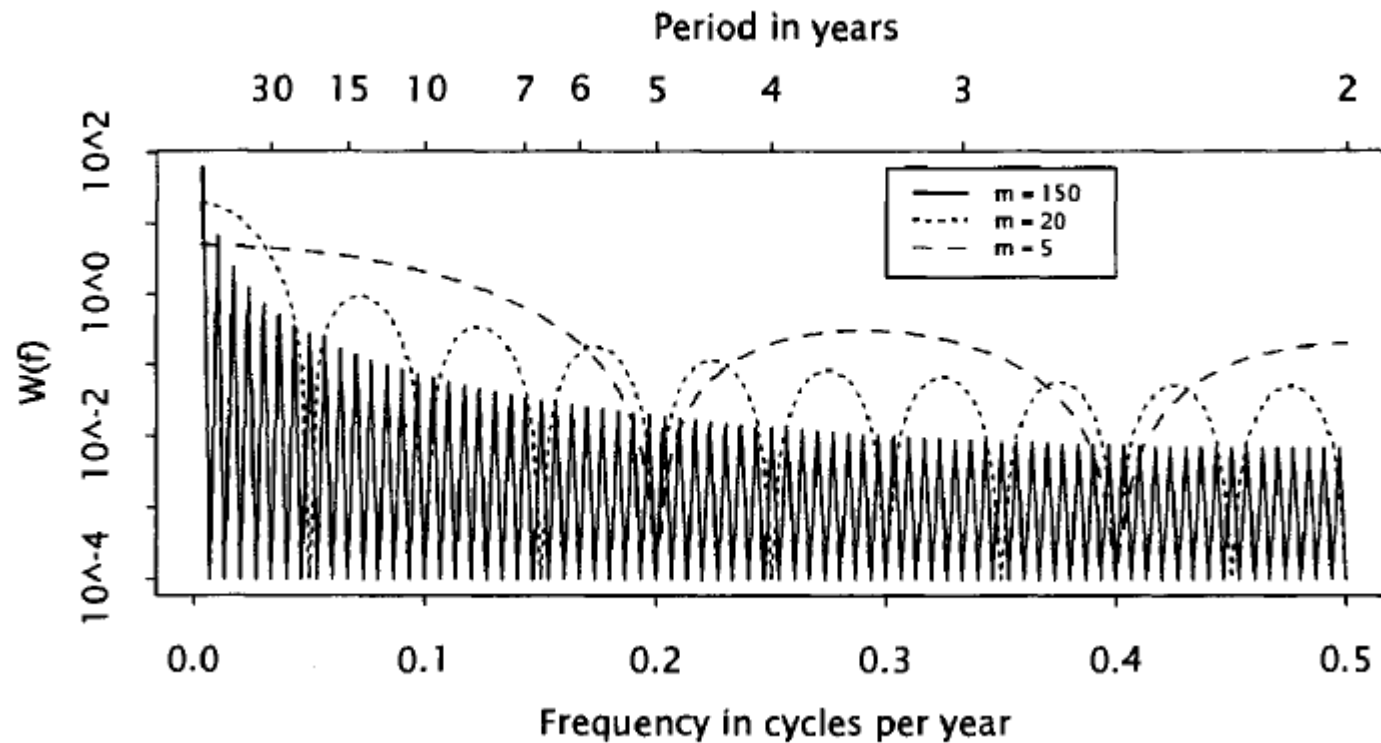
- where  $D_m(f)$  is the Dirichlet kernel (Section 2.2).

- If the spectral window and periodogram are both nonnegative,

- then the spectrum estimate is guaranteed to be nonnegative.

- Important in long run variance estimation

# Spectral window



Spectral windows for the modified Bartlett spectrum estimates



# Modified Bartlett spectrum estimate

- The central peak in the spectral window of the modified Bartlett estimate is of height  $m$
- The first zeros on either side are at  $f = \pm 1/m$  cycles per unit time
- However, a sizable proportion of the mass is contained in the sidelobes which decay slowly
- The periodogram values at some distance from  $f$  may also contribute substantially to the integral
- The estimated spectrum in one frequency may be swamped by leakage from another with high power
  - Even when these bands are not adjacent
  - Such leakage is different than the leakage in the periodogram itself
  - The major source is the sidelobes in the smoothing spectral window
- The sidelobes of the modified Bartlett window are larger and decay more slowly
  - As compared with Anderson (1994)
  - Thus it is rarely used
- Note that sidelobes are bound to exist for any spectrum estimates with truncation point  $m$

# Another representation of a spectrum estimate

- From the representation  $\hat{s}(f) = \int_0^1 W_n(f - f')I(f')df'$ , we have

$$\hat{s}(f) = \int_0^1 W(f')I(f - f')df',$$

- for suitable function  $W(f)$ .
- Furthermore, we can write  $w_r = \int_0^1 W(f) \exp(2\pi i f r) df$ .
  - Mathematically,  $w_r$  are the Fourier coefficients of  $W(f)$
  - $W_n(f)$  is a partial sum of the Fourier series for  $W(f)$ 
    - Recall  $W_n(f)$  is the spectral window
    - I understand this as constructing  $w_r$  reversely from the spectral window
- Daniell (1946) suggested  $W(f) = (2\delta)^{-1} \mathbb{I}_{|f| < \delta}$ .
  - The resulting estimate is the integral analog of simple moving average filter (Section 7.2)
  - It may be applied successively to build up more complex filter (p.157)
- I will skip the remaining of Section 8.5 which discuss the computation of discrete  $\hat{s}(f)$

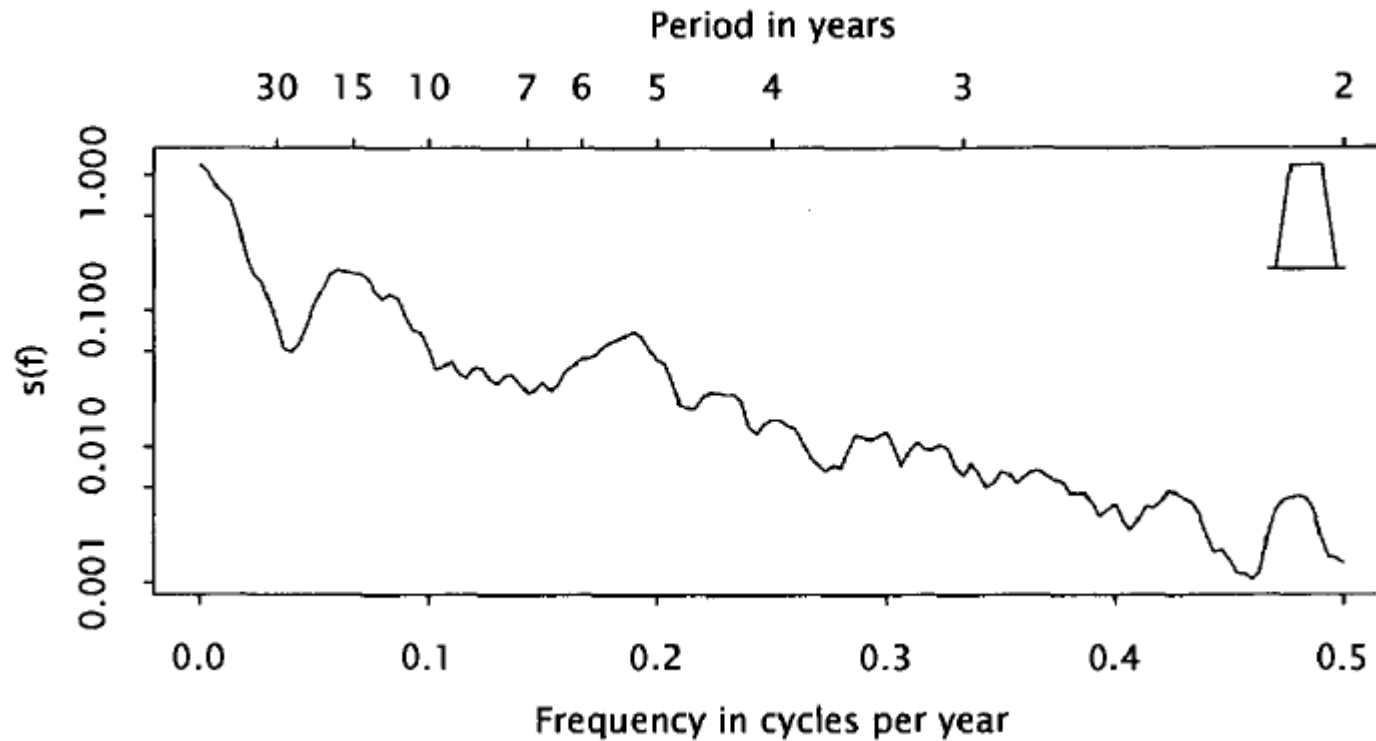
# Choice of a spectral window

- Four factors need to be considered when choosing a spectral window:
  - Resolution or bandwidth
  - Stability
  - Leakage
  - Smoothness
- Resolution: the ability of a spectrum estimate to represent fine structure in the frequency
  - Such as narrow peaks in the spectrum
  - A narrow peak in the periodogram is usually spread out into a broader peak in the spectrum
  - This peak is roughly an image of the spectral window and its width is the bandwidth
  - If the spectrum contains two close narrow peaks, they may overlap and form a single peak
  - In this case, the estimate has failed to resolve the peaks
- Stability: the extent to which estimates from different segments agree
  - In other words, it is the ability to remove irrelevant fine structure
  - Resolution and stability are conflicting requirements
    - Easy to see by definition
  - Section 9.5 gives a statistical treatment of the stability of spectrum estimates

# Choice of a spectral window

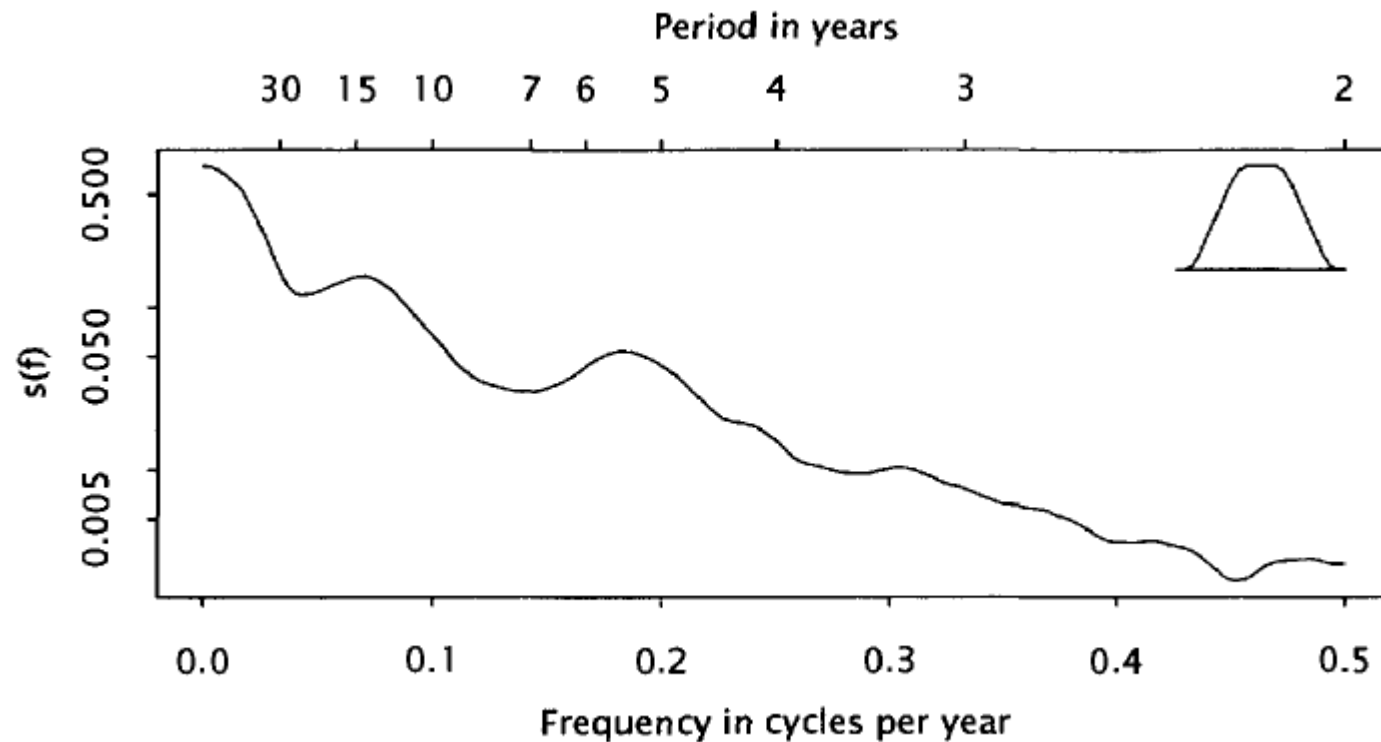
- Leakage
  - Caused by sidelobes in the spectral window
  - Always exists if there is a nontrivial truncation point  $m < n$
  - The computationally simpler discrete spectral averages (Section 8.5) can avoid leakage entirely
    - The part that I skip
  - First, use a data window to control leakage in the periodogram
  - Second, use a spectral window of a desired compact form
- Smoothness
  - Less tangible but important in visualization
  - The need for smoothness can introduce further conflict in choosing a window
  - Bloomfield gave an example of Daniell estimate here
  - Repeated smoothing is possible but yields a less stable estimate

# Example: wheat price



Smoothed periodogram of logarithms of wheat price index, with spectral window inset (Modified Daniell filter,  $m=6$ .)

# Example: wheat price

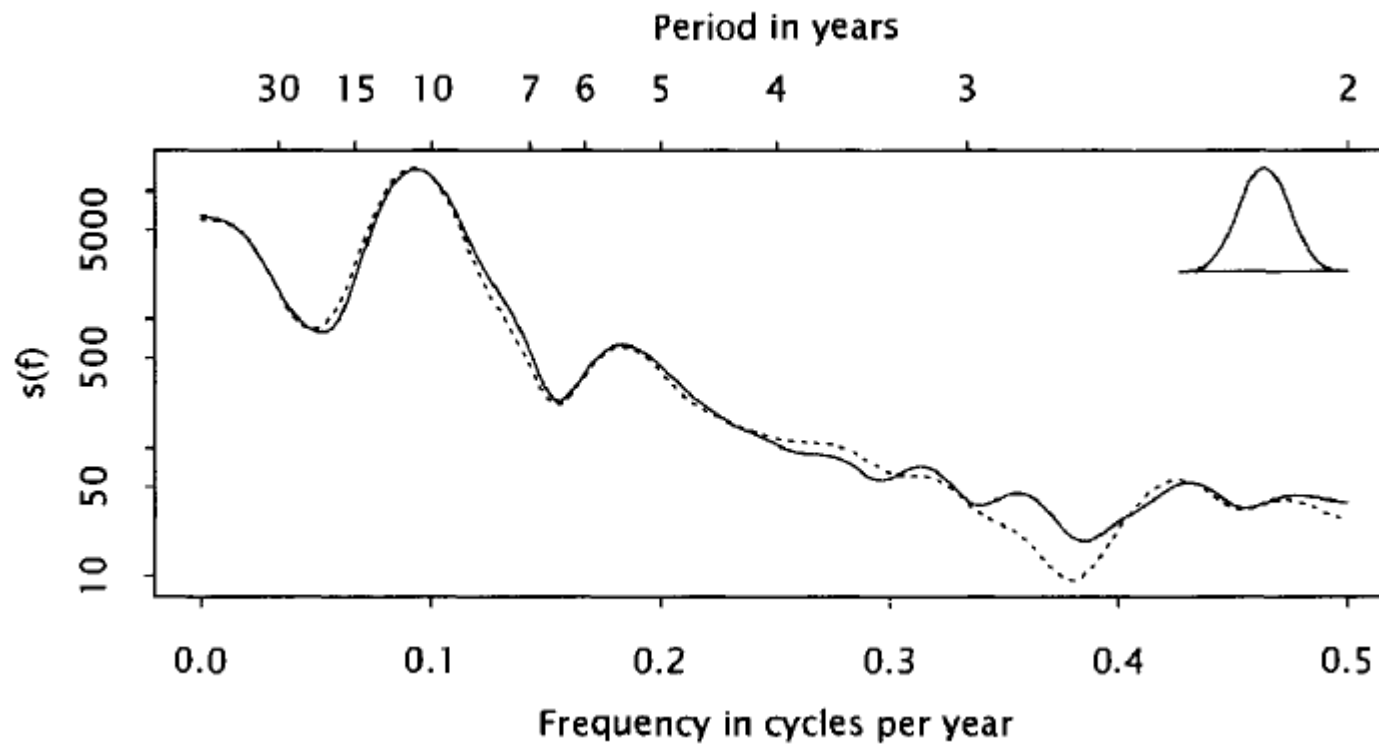


Smoothed periodogram of logarithms of wheat price index, with spectral window inset (Modified Daniell filter,  $m=6,12$ .)

# Example: sunspot

- For sunspot data, the spectral weights have the same span as the wheat price
  - Which means they cover the same number of periodogram ordinates
  - However, they are smoother and have a narrower peak
  - The spectral estimates correspondingly show more rounded but slightly larger fluctuations
  - This argument is made more precise in Chapter 9
- For the square root transformation, refer to Section 6.7 for the idea
- Further discussion of the choice of a spectral window available in Jenkins (1961) and Parzen (1961)
  - I try to find a more recent survey but have not yet found one

# Example: sunspot



Smoothed periodogram of yearly sunspot numbers (solid line) and their square roots (broken line), with spectral window inset (Modified Daniell filter,  $m=6,6,6$ .)



# Reroughing the spectrum

- Recall repeated smoothing may yield a less stable estimate
  - This seems to be the case for the wheat price spectrum
  - The trough between the two peaks at  $f = 0$  and  $f = 0.07$  is not as clear as it is before second smoothing
  - Such loss of resolution suggests the possibility of oversmoothing
  - Similar problem appears in the sunspot spectra
  - This motivates the idea of reroughing or twicing (Tukey, 1977)
- In the context of linear filters, we have rough = input - output
- Since spectra are nonnegative, we can instead define rough = input/output
- To be specific, we can define the rough in spectrum estimation as

$$r(f) = \frac{I(f)}{\hat{s}(f)}.$$

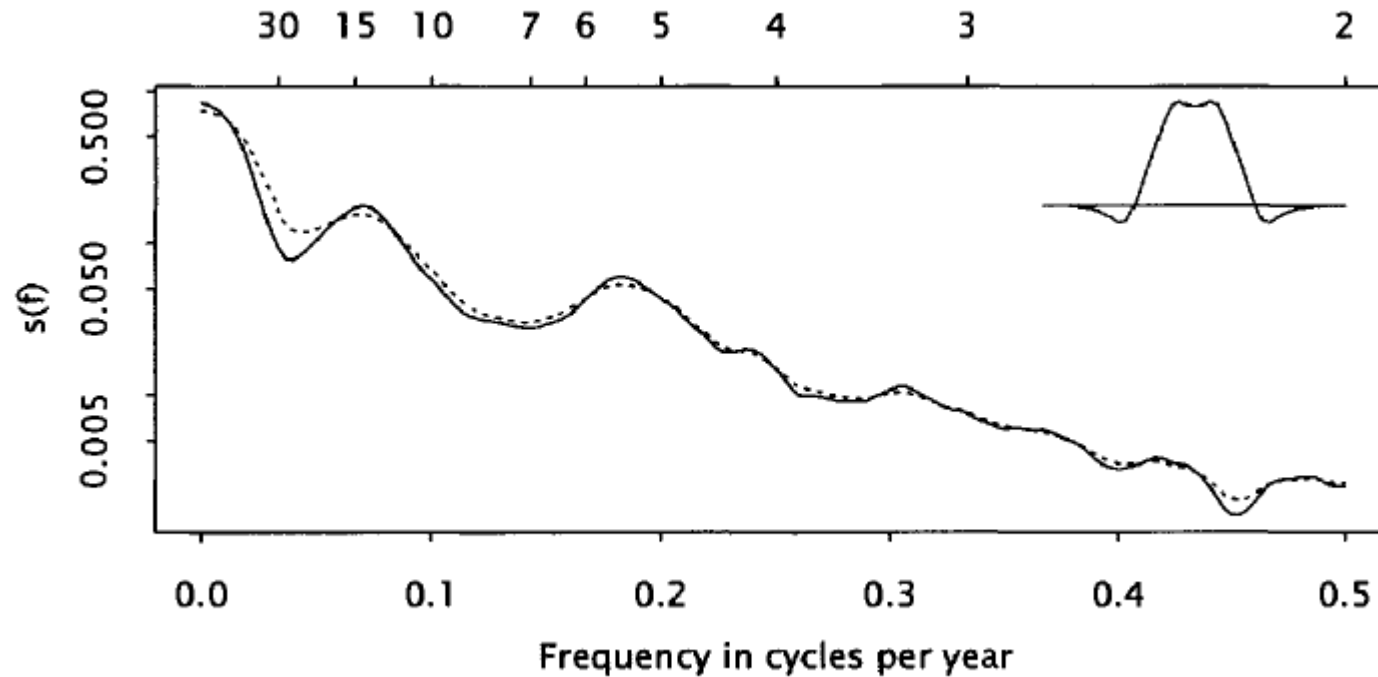
# Reroughing the spectrum

- If  $\hat{s}(f)$  suffers from oversmoothing,
  - there are narrow-band features in the periodogram that were not fully transferred to  $\hat{s}(f)$ .
  - They will appear partially in  $r(f)$ , which can be extracted with another round of smoothing:

$$\tilde{r}(f) = \sum_u g_u r(f - f_u).$$

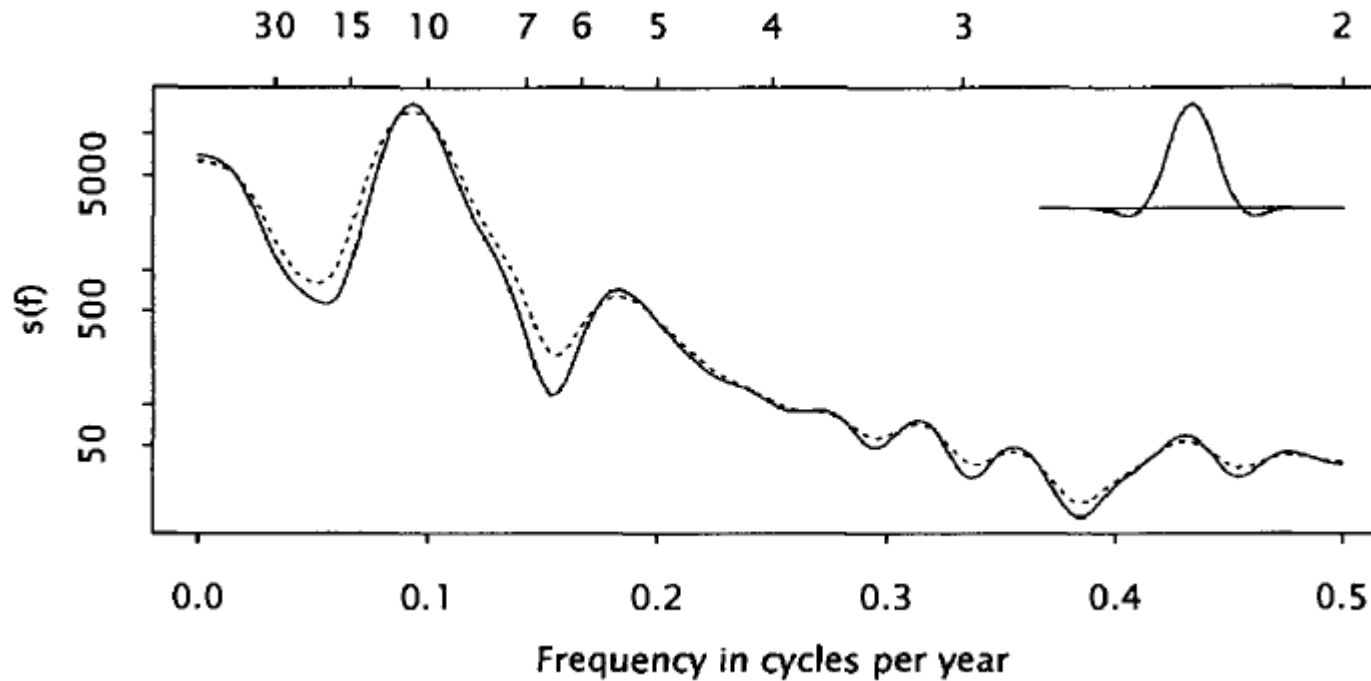
- The reroughed spectrum estimate is  $\hat{s}_r(f) = \tilde{r}(f)\hat{s}(f)$ . - If the same filter is used in the second round as in the first, the process is called twicing

# Example: wheat price



Twiced spectrum estimate of wheat price index (solid line) and original estimate (broken line) (Modified Daniell filter,  $m=6,12$ .)

# Example: sunspot



Twiced spectrum estimate of yearly sunspot numbers (solid line) and original estimate (broken line)  
(Modified Daniell filter,  $m=6,6,6$ .)

# Prewhitening

- Reroughing is closely related to prewhitening (Blackman and Tukey, 1959)
- Oversmoothing leads to leakage from frequency bands with high power to adjacent bands
  - From this perspective, oversmoothing is caused by a large dynamic range in the spectrum
- Prewhitening is a technique for reducing dynamic range prior to forming the periodogram
  - It reduces the leakage and allows the use of a more stable estimate with lower resolution
  - The simplest form of prewhitening is replacing the data by their first differences
    - I think differencing is also used as stationary transformation in practice
- Vs reroughing
  - Reroughing is an enhancement to spectrum estimation
  - Prewhitening is a form of preprocessing

# Concluding Remarks

# Comments

- Periodogram analysis
  - Some practical problems and their solutions are covered
  - The statistical regularity over different segments remains a concern
- Spectrum estimation
  - Smoothing the periodogram to focus on the broad behavior
  - From Fourier transform form to autocovariance form
  - Data windows and their corresponding spectral window
  - Choice of spectral window
  - Reroughing and prewhitening