Modeling the cumulative incidence function of clustered competing risks data: a multinomial GLMM approach

master thesis defense



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- 1 Data
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Clustered competing risk data



Key terms:

- 1 Clustered: groups with a dependence structure (e.g. families);
- 2 Causes competing by something.

Something?

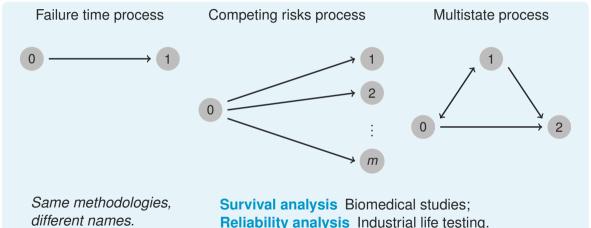
- Failure of an industrial or electronic component;
- Occurence or cure of a disease or some biological process;
- Progress of a patient clinic state.

Independent of the application, always the same framework

Cluster	ID	Cause 1	Cause 2	Censorship	Time	Feature
1	1	Yes	No	No	10	Α
1	2	No	No	Yes	8	Α
2	1	No	No	Yes	7	В
2	2	No	Yes	No	5	Α

Big picture: Failure time data/time-to-event outcomes





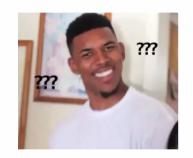


A comprehensive reference is Kalbfleisch and Prentice (2002)'s book.

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Modeling clustered competing risks data









What? Why? How?

Modeling failure time data



First of all, we have to choose which scale we model the survival experience.

1 Usually, is in the

hazard (failure rate) scale:
$$\lambda(t \mid \text{features}) = \lambda_0(t) \times c(\text{features}).$$
 (1)

We have a Equation 1 for each competing cause.

The cluster dependence is something actually not measured...

Not measured dependence \rightarrow random/latent effects \rightarrow Frailty models.

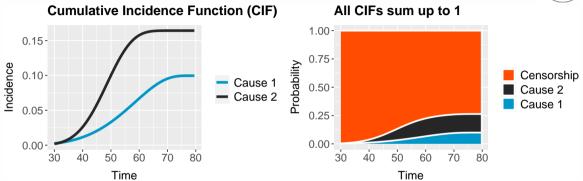
Frailty-based models for (multiple) survival experiences turn out in challengeable likelihood functions with inference routines mostly done via

 Elaborated and slow expectation—maximization (EM) algorithms; Inefficient Markov chain Monte Carlo (MCMC) schemes.

2 Not usually, the probability scale.

$\textbf{Probability scale} \rightarrow \textbf{Cause-specific CIF}$





i.e., $CIF = \mathbb{P}[\text{ failure time } \leq t, \text{ a given cause } | \text{ features \& latent effects }].$

Common applications: family studies.

↓ Keywords: within-family/cluster dependence; age at disease onset; populations.

Formally,



for a cause-specific of failure k, the cumulative incidence function (CIF) is defined as

$$F_k(t \mid \mathbf{x}) = \mathbb{P}[T \leqslant t, \ K = k \mid \mathbf{x}]$$

$$= \int_0^t f_k(z \mid \mathbf{x}) \, \mathrm{d}z \quad (f_k(t \mid \mathbf{x}) \text{ is the (sub)density for the time to a type } k \text{ failure})$$

$$= \int_0^t \underbrace{\lambda_k(z \mid \mathbf{x})}_{\text{cause-specific hazard function}} \underbrace{S(z \mid \mathbf{x})}_{\text{overall survival function}} dz, \quad t > 0, \quad k = 1, \dots, K.$$



Again, a comprehensive reference is Kalbfleisch and Prentice (2002)'s book.



Cederkvist et al. (2019)'s CIF specification



For two competing causes of failure, the cause-specific CIFs are specified in the following manner

$$F_k(t \mid \boldsymbol{x}, \ u_1, \ u_2, \ \eta_k) = \underbrace{\pi_k(\boldsymbol{x}, \ u_1, \ u_2)}_{\text{cluster-specific risk level}} \times \underbrace{\Phi[w_k g(t) - \boldsymbol{x} \gamma_k - \eta_k]}_{\text{cluster-specific failure time trajectory}}, \quad t > 0, \quad k = 1, \ 2, \quad (2)$$

with

1
$$\pi_k(\mathbf{x}, \mathbf{u}) = \exp\{\mathbf{x}\beta_k + u_k\} / \left(1 + \sum_{m=1}^{K-1} \exp\{\mathbf{x}\beta_m + u_m\}\right), \quad k = 1, 2, \quad K = 3;$$

 \bullet $\Phi(\cdot)$ is the cumulative distribution function of a standard Gaussian distribution;

In Cederkvist et al. (2019), this CIF specification is modeled under a *challengeable* pairwise composite likelihood approach (Lindsay 1988; Varin, Reid, and Firth 2011).

Our contribution: a full likelihood analysis



For two competing causes of failure, a subject i, in the cluster i, in time t, we have

For two competing causes of failure, a subject
$$i$$
, in the cluster j , in time t , we have
$$y_{ijt} \mid \underbrace{\{u_{1j}, u_{2j}, \eta_{1j}, \eta_{2j}\}}_{\text{latent effects}} \sim \text{Multinomial}(p_{1ijt}, p_{2ijt}, p_{3ijt})$$

$$\begin{bmatrix} u_1 \\ u_2 \\ \eta_1 \\ \eta_2 \end{bmatrix} \sim \text{Multivariate} \begin{pmatrix} \begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \sigma_{u_1}^2 & \text{cov}(u_1, u_2) & \text{cov}(u_1, \eta_1) & \text{cov}(u_1, \eta_2) \\ \sigma_{u_2}^2 & \text{cov}(u_2, \eta_1) & \text{cov}(u_2, \eta_2) \\ \sigma_{\eta_1}^2 & \sigma_{\eta_2}^2 & \sigma_{\eta_2}^2 \end{pmatrix}$$

$$\begin{bmatrix} \eta_1 \\ \eta_2 \end{bmatrix} \qquad \text{Normal} \qquad \begin{bmatrix} 0 \\ 0 \end{bmatrix}^{\gamma} \begin{bmatrix} \sigma_{\eta_1}^2 & \text{cov}(\eta_1, \eta_2) \\ \sigma_{\eta_2}^2 \end{bmatrix} \end{bmatrix}$$

$$p_{kijt} = \frac{\partial}{\partial t} F_k(t \mid \boldsymbol{x}, \boldsymbol{u}, \eta_k)$$

$$= \frac{\exp\{\boldsymbol{x}_{kij}\beta_k + u_{kj}\}}{1 + \sum_{m=1}^{K-1} \exp\{\boldsymbol{x}_{mij}\beta_m + u_{mj}\}}$$

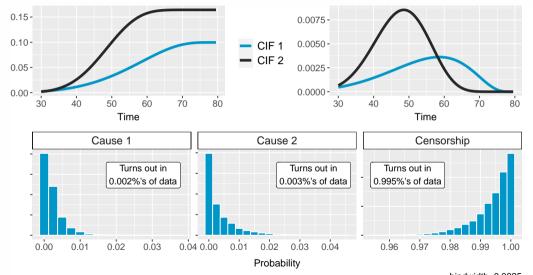
$$\times w_k \frac{\delta}{2\delta t - 2t^2} \Phi\left(w_k \operatorname{arctanh}\left(\frac{t - \delta/2}{\delta/2}\right) - \boldsymbol{x}_{kij}\gamma_k - \eta_{kj}\right), \quad k = 1, 2.$$

Simulating from the model



dCIF 1

dCIF 2



Marginal likelihood function for two competing causes



$$L(\theta; y) = \prod_{j=1}^{J} \int_{\mathfrak{R}^4} \pi(y_j \mid \boldsymbol{r}_j) \times \pi(\boldsymbol{r}_j) \, d\boldsymbol{r}_j$$

$$= \prod_{j=1}^{J} \int_{\mathfrak{R}^4} \left\{ \prod_{i=1}^{n_j} \prod_{t=1}^{n_{ij}} \left(\frac{(\sum_{k=1}^{K} y_{kijt})!}{y_{1ijt}! y_{2ijt}! y_{3ijt}!} \prod_{k=1}^{K} \rho_{kijt}^{y_{kijt}} \right) \right\} \times$$
fixed effect component
$$(2\pi)^{-2} |\Sigma|^{-1/2} \exp\left\{ -\frac{1}{2} \boldsymbol{r}_j^{\mathsf{T}} \Sigma^{-1} \boldsymbol{r}_j \right\} d\boldsymbol{r}_j$$

$$= \prod_{j=1}^{J} \int_{\mathfrak{R}^4} \left\{ \underbrace{\prod_{i=1}^{n_j} \prod_{t=1}^{n_{ij}} \prod_{k=1}^{K} \rho_{kijt}^{y_{kijt}}}_{\text{fixed effect}} \right\} \underbrace{(2\pi)^{-2} |\boldsymbol{\Sigma}|^{-1/2} \exp\left\{-\frac{1}{2} \boldsymbol{r}_j^\top \boldsymbol{\Sigma}^{-1} \boldsymbol{r}_j\right\}}_{\text{latent effect component}} \mathrm{d}\boldsymbol{r}_j,$$

latent effect component

with p_{kijt} from Equation 3 and where $\theta = [\beta \ \gamma \ \mathbf{w} \ \sigma^2 \ \rho]^{\top}$ is the parameters vector.

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TMB: Automatic Differentiation and Laplace Approximation





Kristensen et al. (2016).

An R (R Core Team 2021) package for the quickly implementation of complex random effect models through simple C++ templates.

Workflow

- Write your objective function in a .cpp through a #include <TMB.hpp>;
- 2 Compile and load it in R via TMB::compile() and base::dyn.load(TMB::dynlib());
- 3 Compute your objective function derivatives with obj <- TMB::MakeADFun();</p>
- 4 Perform the model fitting, opt <- base::nlminb(obj\$par, obj\$fn, obj\$gr);</pre>
- 5 Compute the parameters standard deviations, TMB::sdreport(obj).



For details about TMB, AD, and Laplace approximation: Laureano (2021).

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Simulation study model designs



Risk model

Latent effects only on the risk level i.e.,

$$\Sigma = \begin{bmatrix} \sigma_{u_1}^2 & \mathsf{COV}_{u_1, u_2} \\ & \sigma_{u_2}^2 \end{bmatrix}.$$

Block-diag model

Latent effects on the risk and time levels without cross-correlations i.e.,

$$\Sigma = egin{bmatrix} \sigma_{u_1}^2 & \operatorname{cov}_{u_1,u_2} & 0 & 0 \ & \sigma_{u_2}^2 & 0 & 0 \ & & \sigma_{\eta_1}^2 & \operatorname{cov}_{\eta_1,\eta_2} \ & & & \sigma_{\eta_2}^2 \end{bmatrix}.$$

Time model

Latent effects only on the failure time trajectory level i.e.,

$$\Sigma = egin{bmatrix} \sigma_{\eta_1}^2 & \mathsf{cov}_{\eta_1,\eta_2} \ & \sigma_{\eta_2}^2 \end{bmatrix}.$$

Complete model

A complete latent effects structure i.e..

$$\Sigma = \begin{bmatrix} \sigma_{u_1}^2 & \text{cov}_{u_1, u_2} & 0 & 0 \\ & \sigma_{u_2}^2 & 0 & 0 \\ & & \sigma_{\eta_1}^2 & \text{cov}_{\eta_1, \eta_2} \\ & & & \sigma_{\eta_2}^2 \end{bmatrix}. \qquad \Sigma = \begin{bmatrix} \sigma_{u_1}^2 & \text{cov}_{u_1, u_2} & \text{cov}_{u_1, \eta_1} & \text{cov}_{u_1, \eta_2} \\ & \sigma_{u_2}^2 & \text{cov}_{u_2, \eta_1} & \text{cov}_{u_2, \eta_2} \\ & & & \sigma_{\eta_1}^2 & \text{cov}_{\eta_1, \eta_2} \\ & & & & \sigma_{\eta_2}^2 \end{bmatrix}.$$

Simulation study setup

Four latent effects structures:



1 Risk model;

2 Time model;

Block-diag model;

4 Complete model.

Two CIF configurations:

Low max incidence \approx 0.15;

High max incidence \approx 0.60.

For each of those $4 \times 2 = 8$ scenarios, we vary the sample and cluster sizes:

5000 data points

- 2500 clusters of **size 2**;
- 1000 clusters of **size 5**;
- 500 clusters of size 10.

30000 data points

- 15000 clusters of size 2;
- 6000 clusters of size 5;
- 3000 clusters of size 10.

60000 data points

- 30000 clusters of **size 2**;
- 12000 clusters of **size 5**;
- 6000 clusters of **size 10**.

Totalizing, $\mathbf{8} \times \mathbf{3} \times \mathbf{3} = \mathbf{72}$ scenarios.

For each scenario, we simulate 500 samples, totalizing $72 \times 500 = 36000$ model fittings.



First of all, the **time**.

 The non-complete models (2D Laplace aprox.) are kind of fast, taking always less than 5 min.

In the most expensive scenarios (30K 4D Laplaces).

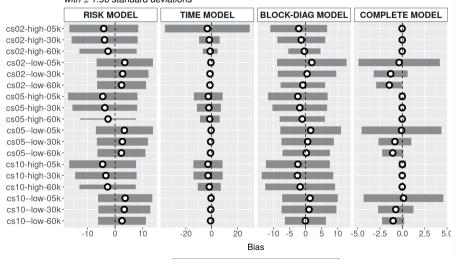
- the complete model takes 30 min.
 In a full R implementation with 10K 4D Laplaces, it took 30hrs. TMB is fast.
- We also did a Bayesian analysis via Stan/NUTS-HMC (Stan Development Team 2020).
 - 1 week of parallelized processing for a 2500 size 2 clusters scenario with tuned NUTS.
 This just reinforces the MCMC impracticability for some complex models.

Model identifiability.

The non-complete models fail to learn the data.
 They appear to be not structured enough to capture the data characteristics.

Parameter: β₁

with + 1.96 standard deviations



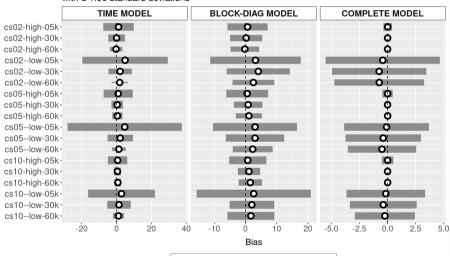
400 450 500

Fittinas



Parameter: $log(\sigma_4^2)$

with ± 1.96 standard deviations



350

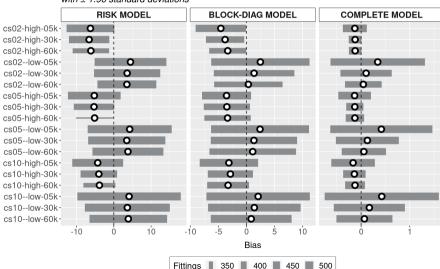
400 450 500

Fittings



Parameter: $z(\rho_{12})$

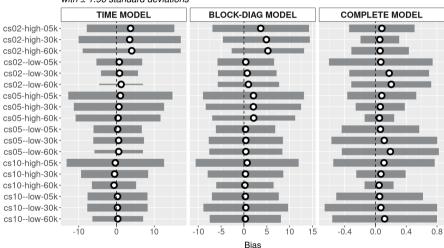
with ± 1.96 standard deviations





Parameter: $z(\rho_{34})$

with ± 1.96 standard deviations



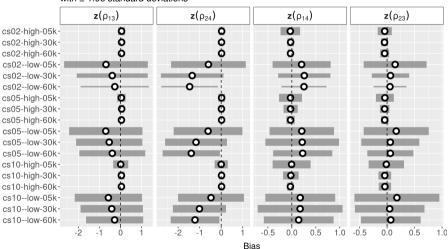
Fittings

350 400 450 500



Complete model's cross-correlations

with + 1.96 standard deviations



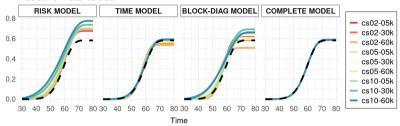


Simulation study results: High CIF scenario



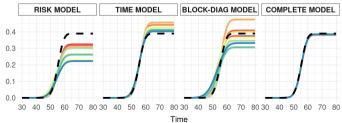


True curve in dashed black



CIF of failure cause 2

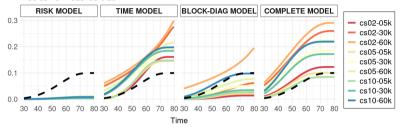
True curve in dashed black



Simulation study results: Low CIF scenario

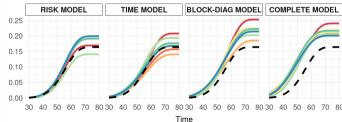
CIF of failure cause 1

True curve in dashed black



CIF of failure cause 2

True curve in dashed black



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Take-home message



- The complete model works. It's not magnificent, but it works.
- 1 It works better in the high CIF scenarios;
- 2 As expected, as the sample size increases the results get better;
- 3 Given the low data representativity, the model needs a considerable amount of data to perform well;
- 4 In standard multinomial GLMMs, as bigger the clusters better the results. In our CIF-based formulation, this characteristic is not so clear.

What else can we do?

- 1 Instead of a conditional approach (latent effects model), we can try a marginal approach e.g., an McGLM (Bonat and Jørgensen 2016);
- We can also try a copula (Embrechts 2009), on maybe two fronts:1) for a full specification; 2) to accommodate the within-cluster dependence.



For more read Laureano (2021) master thesis.

Thanks for watching and have a great day



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Joint work with

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http://leg.ufpr.br/~wagner

Paulo Justiniano Ribeiro Jr. http://leg.ufpr.br/~paulojus

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References

Bonat, W. H., and B. Jørgensen. 2016. "Multivariate Covariance Generalized Linear Models." *Journal of the Royal Statistical Society, Series C (Applied Statistics)* 65 (5): 649–75.

Cederkvist, L., K. K. Holst, K. K. Andersen, and T. H. Scheike. 2019. "Modeling the Cumulative Incidence Function of Multivariate Competing Risks Data Allowing for Within-Cluster Dependence of Risk and Timing." *Biostatistics* 20 (2): 199–217.

Embrechts, P. 2009. "Copulas: A Personal View." The Journal of Risk and Insurance 76 (3): 639–50.

Kalbfleisch, J. D., and R. L. Prentice. 2002. *The Statistical Analysis of Failure Time Data*. Second Edition. Hoboken, New Jersey: John Wiley & Sons, Inc.

Kristensen, K., A. Nielsen, C. W. Berg, H. J. Skaug, and B. M. Bell. 2016. "TMB: Automatic Differentiation and Laplace Approximation." *Journal of Statistical Software* 70 (5): 1–21.

Laureano, H. A. 2021. "Modeling the Cumulative Incidence Function of Clustered Competing Risks Data: A Multinomial Glmm Approach." Master's thesis, Federal University of Paraná (UFPR).

Lindsay, B. G. 1988. "Composite Likelihood Methods." Comtemporary Mathematics 80 (1): 221–39.

R Core Team. 2021. R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing. Vienna, Austria.

Stan Development Team. 2020. "RStan: The R Interface to Stan." https://mc-stan.org/.

Varin, C., N. Reid, and D. Firth. 2011. "An Overview of Composite Likelihood Methods." Statistica Sinica 21 (1): 5-42.