

Project 2, MA8001

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September 2019

Exercise a)

We start by simulating a markov chain with binary outcome, $x_i \in \{0, 1\}$, $i = 1, \dots, n$ and stationary conditional probabilities $p(x_{i+1} = l | x_i = k) = P(k, l)$.

Data $\mathbf{y} = (y_1, \dots, y_n)$ are conditionally independent, given the variable of interest. We define a Gaussian likelihood, $p(y_i | x_i) = N(x_i, \tau^2)$, $i = 1, \dots, n$.

To generate data, we simulate a markov chain with $n = 250$, $p = P(0, 0) = P(1, 1) = 0.9$, $p(x_1 = 1) = 0.5$. Then, for each x_i , we generate data $y_i = N(x_i, \tau^2)$, $i = 1, \dots, n$, with $\tau = 0.4$.

The generated data are plotted in figure 1. We see that the chain stays in the same value for several iterations before transitioning to another state. The data for \mathbf{y} is also fairly mixed, and it is difficult to determine \mathbf{x} based purely on the plot of \mathbf{y} .

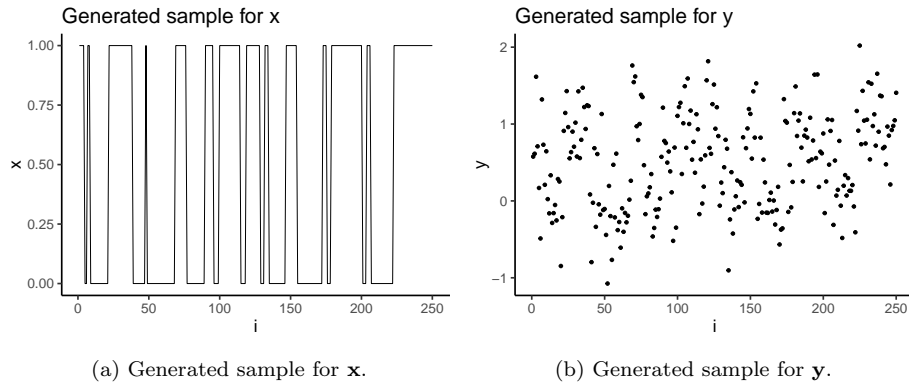


Figure 1: Generated data for this exercise. \mathbf{y} are now considered known, or observed. \mathbf{x} are "hidden" and to be estimated.

Exercise b)

We now do forward recursion to evaluate the marginal likelihood of the data, $p(\mathbf{y})$.

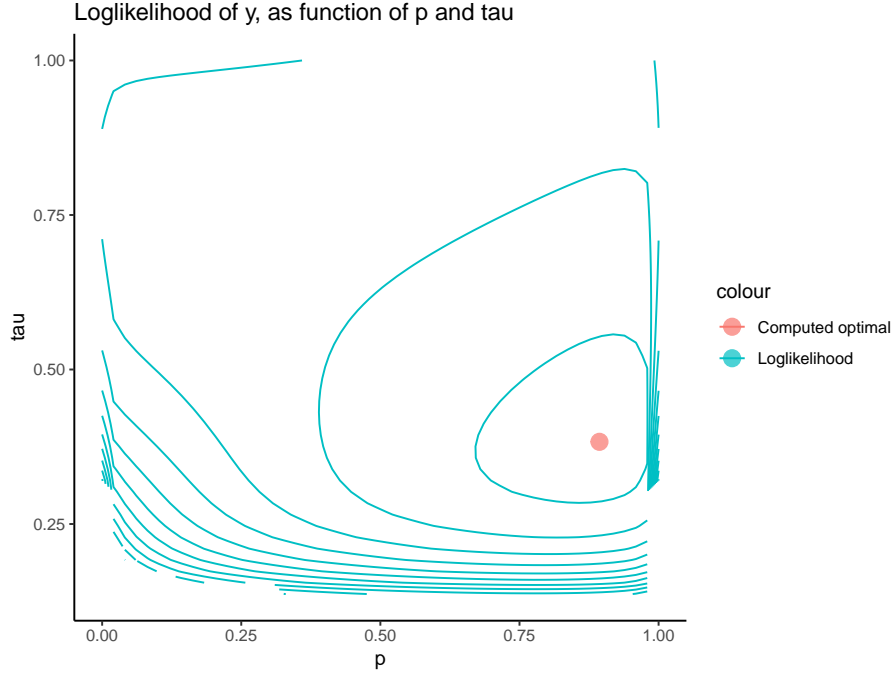


Figure 2: Contour plot of the loglikelihood $p(\mathbf{y})$, with \mathbf{y} from a). The indicated point is the optimal value for p and τ .

We first initialize the first iteration by setting

$$C_1 = \left(\sum_{i=0}^1 p(y_1|x_1 = i) \times p(x_1 = i) \right)^{-1}$$

$$p(x_1 = i|y_1) = C_1 \times p(y_1|x_1 = i) \times p(x_1 = i), i \in \{0, 1\}.$$

Then, we do forward recursion for $t = 2, \dots, n$ by the formula

$$C_t = \left(\sum_{i=0}^1 \sum_{j=0}^1 p(y_t|x_t = j) \times P(i, j) \times p(x_{t-1} = i|\mathbf{y}_{1:(t-1)}) \right)^{-1},$$

$$p(x_{t-1} = i, x_t = j|\mathbf{y}_{1:t}) = C_t \times p(y_t|x_t = j) \times P(i, j) \times p(x_{t-1} = i|\mathbf{y}_{1:(t-1)}),$$

$$p(x_t = i|\mathbf{y}_{1:t}) = \sum_{i=0}^1 p(x_{t-1} = i, x_t = j|\mathbf{y}_{1:t}).$$

Then, finally, we arrive at the expression $p(\mathbf{y}) = \prod_{t=1}^n 1/C_t$.

We then compute the marginal likelihood model for a grid of values for p and τ , given the data generated in a). Then, we plot the log of the marginal likelihood in figure 2. The loglikelihood is then optimized using the R function `optim`. The resulting values are $p = 0.9001172$, $\tau = 0.4188034$, which are almost equal to the "true" values set in a).

Exercise c)

We then use the backward recursion to find the marginal probabilities $p(x_i = 1|y_1, \dots, y_n)$, for all $i = 1, \dots, n$. We first note that $p(x_n = i|\mathbf{y})$ is known from the forward recursion, and this is used as initialization for the backward recursion. Then, we can compute the backward probabilities by the following equations for $t = n, \dots, 2$

$$p(x_{t-1} = i, x_t = j|\mathbf{y}) = \frac{p(x_{t-1} = i, x_t = j|\mathbf{y}_{1:t})}{p(x_t = i|\mathbf{y}_{1:t})} \times p(x_t = i|\mathbf{y}),$$

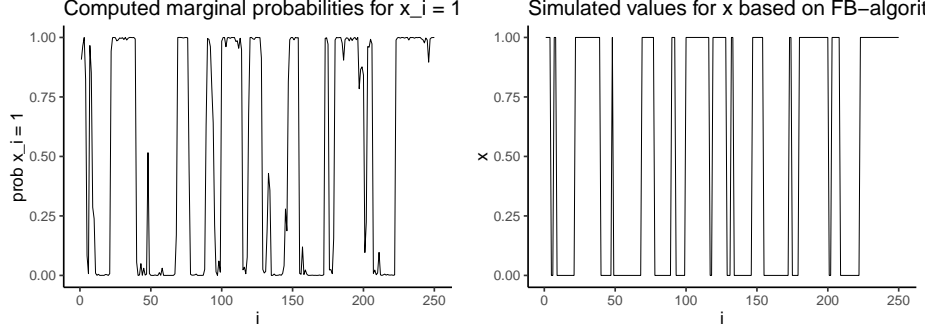
$$p(x_{t-1} = i|\mathbf{y}) = \sum_{j=0}^1 p(x_{t-1} = i, x_t = j|\mathbf{y}),$$

$$p(x_t = j|x_{t-1} = i, \mathbf{y}) = \frac{p(x_{t-1} = i, x_t = j|\mathbf{y})}{p(x_{t-1} = i|\mathbf{y})}.$$

Then, we note that the full posterior model can be written by the backward probabilities as

$$p(\mathbf{X} = \mathbf{x}|\mathbf{y}) = p(x_1|\mathbf{y}) \times \prod_{t=2}^n p(x_t|x_{t-1}, \mathbf{y}).$$

The probabilities for $p(x_i = 1|\mathbf{y})$ are plotted in figure 3a.



(a) Marginal probabilities for $p(x_i = 1|\mathbf{y})$. (b) Generated sample of \mathbf{x} based on $p(\mathbf{x}|\mathbf{y})$.

Figure 3: Marginal probability and one generated sample based on the backward recursion algorithm.

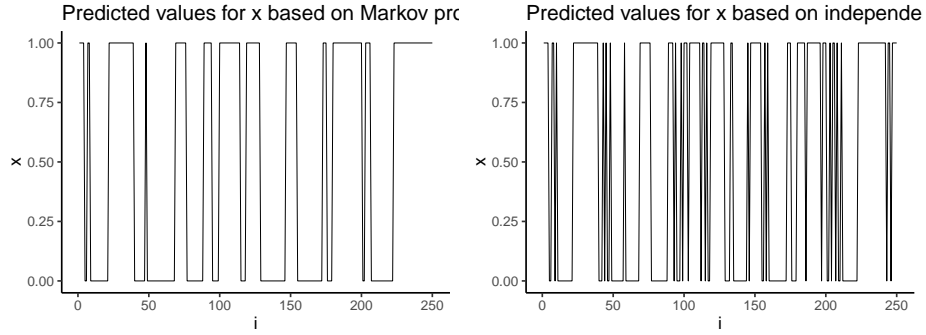
Having computed the probabilities by the backward recursion algorithm, we are also able to simulate values for \mathbf{x} . First, we sample x_1^s from $p(x_1|\mathbf{y})$, and then we can sample $x_i^s \sim p(x_i|x_{i-1}^s, \mathbf{y})$ for $i = 2, \dots, n$. The resulting sample \mathbf{x}^s is plotted in figure 3b.

Exercise d)

We now want to compare results between predicted values for \mathbf{x} with and without the markov assumption. If we assume markov dependence, $x_i = 1$

if $p(x_i = 1|\mathbf{y}) > 0.5$, and $x_i = 0$ if $p(x_i = 1|\mathbf{y}) < 0.5 \Rightarrow p(x_i = 0|\mathbf{y}) > 0.5$. The resulting prediction is plotted in figure 4a.

If we assume no markov dependence, then x_i will be predicted based solely on the individual observations y_i . This means that x_i will receive the value that is most likely based on given y_i . Because y_i has a Gaussian likelihood, it is symmetric. Hence, assuming no markov dependence, we predict $x_i = 0$ if $y_i < 0.5$ and $x_i = 1$ if $y_i > 0.5$. The resulting prediction for \mathbf{x} is plotted in figure 4b.



(a) Predicted values of \mathbf{x} based on Markov as- (b) Predicted values of \mathbf{x} based on no depen-
sumption. dence in Markov chain.

Figure 4: Comparison of predicted values for \mathbf{x} using the Markov chain property and simply using the likelihood model at each individual location.

Comparing figure 4 to figure 1a, we see that the markov assumption provides a better prediction than independence assumption. The independence assumption has too frequent shifts in the values of \mathbf{x} , and evidently does not capture that the underlying \mathbf{x} follows a markov chain. A comparison between the probabilities of $p(x_i = 1|\mathbf{y})$ for markov and independence assumption is also included in figure 5 for reference.

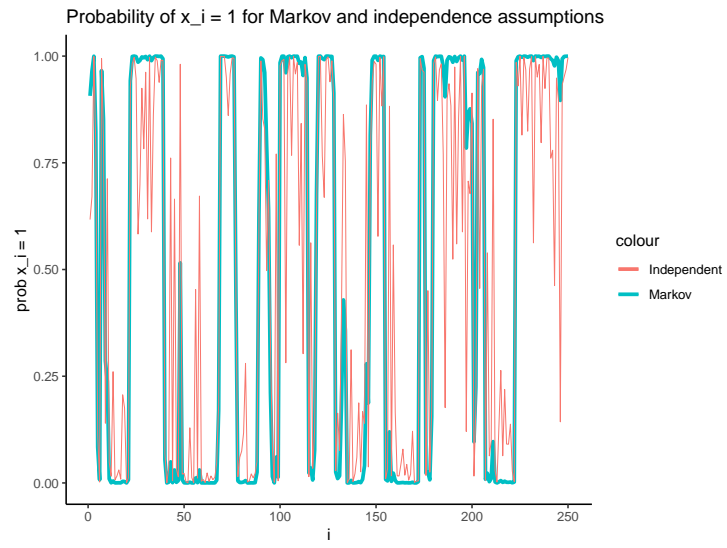


Figure 5: Comparison of probability for $x_i = 1$ for Markov chain model and no Markov dependence model.