# STAT 135 Lecture 11

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## Lemma 0.1 (Neyman Pearson Lemma)

Fill cart with items that are the best deal  $\frac{P(X|H_1)}{P(X|H_0)}$ ,  $P(X|H_1)$  is the value and  $P(X|H_0)$  is the price, until you run out of budget. Maximizes the power  $(1 - \beta)$  for your  $\alpha$ .

## Example 0.2 (Duality of Confidence Intervals and Hypothesis Tests)

Let  $X_1, \ldots, X_n \stackrel{iid}{\sim} \mathcal{N}(\mu, 1)$ 

Testing:  $H_0: \mu = \mu_0$ 

 $H_1: \mu \neq \mu_0$ 

**Solution.** Get a level 0.05 test by reject when  $|\bar{X} - \mu_0| > \frac{1.96}{\sqrt{n}}$ Accept when  $\frac{-1.96}{\sqrt{n}} \le \mu_0 - \bar{X} \le \frac{1.96}{\sqrt{n}} \to \bar{X} - \frac{1.96}{\sqrt{n}} \le \mu_0 \le \bar{X} + \frac{1.96}{\sqrt{n}}$  $\{\mu : \mu \in 95CI\} = \{\mu : H_0 : \mu = \mu_0 \text{ not rejected}\}$ 

Set of  $\mu$  in CI is the same as set of  $\mu$  which would not be rejected testing From 95% CI  $\Rightarrow$  Hypothesis test, reject outside the interval

From Hypothesis test ⇒ CI (Those null hypothesis wouldn't be rejected) ⇒ Can work back and forth between CI and Hypothesis Testing

### Remark 0.3 (Comparing two samples)

 $X_1, \dots, X_n \overset{iid}{\sim} f, \qquad Y_1, \dots, Y_m \overset{iid}{\sim} g$  Applications:

- Treatment vs control
- A/B Testing

#### Example 0.4

$$X_1, \ldots, X_n \stackrel{iid}{\sim} \mathcal{N}(\mu_X, \sigma^2)$$

$$Y_1, \ldots, Y_m \stackrel{iid}{\sim} \mathcal{N}(\mu_Y, \sigma^2)$$

Example 0.4
$$X_1, \dots, X_n \overset{iid}{\sim} \mathcal{N}(\mu_X, \sigma^2)$$

$$Y_1, \dots, Y_m \overset{iid}{\sim} \mathcal{N}(\mu_Y, \sigma^2)$$

$$\bar{X} - \bar{Y} \sim \mathcal{N}(\mu_Y - \mu_X, \sigma^2(\frac{1}{n} + \frac{1}{m}))$$

$$\frac{\bar{X} - \bar{Y} - (\mu_X - \mu_X)}{\sigma \sqrt{\frac{1}{n} + \frac{1}{m}}} \sim N(0, 1)$$

$$\frac{X-Y-(\mu_x-\mu_x)}{\sigma\sqrt{\frac{1}{n}+\frac{1}{m}}}\sim N(0,1)$$

$$\bar{X} - \bar{Y} \pm 1.96 SE(\bar{X} - \bar{Y}) = \bar{X} - \bar{Y} \pm 1.96 \sigma \sqrt{\frac{1}{n} + \frac{1}{m}}$$

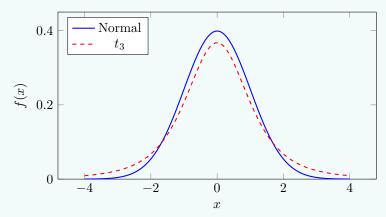
Reject  $\mu_x = \mu_y$  when  $\mu_X - \mu_Y = 0$ 

# Method 0.5 (One sample t-test)

 $X_1,\dots,X_n \overset{iid}{\sim} \mathcal{N}(\mu,\sigma^2)$  with  $\sigma^2$  known  $\bar{X}\pm 1.96\frac{\sigma}{\sqrt{n}}$  Exact 95%

If we estimate  $\sigma^2$  by  $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$ 

Then  $\frac{\sqrt{n}(\bar{X}-\mu)}{s} \sim t_{n-1}$ 



If we want a  $(1 - \alpha)$ % Confidence interval for  $\mu$ .

 $\bar{X} \pm t_{n-1}(\frac{\alpha}{2})\frac{s}{\sqrt{n}}, \qquad t_{n-1}(\frac{\alpha}{2}) > 1.96$ 

# Method 0.6 (Back to two sample problem)

 $X_1, \dots, X_n \overset{iid}{\sim} \mathcal{N}(\mu_X, \sigma^2)$   $Y_1, \dots, Y_m \overset{iid}{\sim} \mathcal{N}(\mu_Y, \sigma^2)$   $\sigma^2$  unknown

$$\frac{(\bar{X} - \bar{Y}) - (\mu_X - \mu_Y)}{S_p \sqrt{\frac{1}{n} + \frac{1}{m}}} \sim t_{n+m-2}$$

$$S_p^2 = \frac{(n-1)S_X^2 + (m-1)S_Y^2}{n+m-2}$$
 "Pooled Estimate"

$$S_X^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$S_Y^2 = \frac{1}{m-1} \sum_{i=1}^m (y_m - \bar{y})^2$$

$$E(S_p^2) = \frac{(n-1)E(S_X^2) + (m-1)E(S_Y^2)}{n+m-2} = \frac{(n-1+m-1)\sigma^2}{n+m-2} = \sigma^2$$

# Method 0.7 (Inference for two sample problem)

Exact  $(1 - \alpha)$  CI for  $\mu_X - \mu_Y$ 

$$\bar{X} - \bar{Y} \pm t_{n+m-2}(\frac{\alpha}{2})S_p\sqrt{\frac{1}{n} + \frac{1}{m}}$$

 $H_0: \mu_X = \mu_Y$ 

 $H_1: \mu_X \neq \mu_Y$ 

Level  $\alpha$  test: Reject for large  $\frac{|\bar{X} - \bar{Y}|}{S_p \sqrt{\frac{1}{n} + \frac{1}{m}}}$ 

## Definition 0.8

 $H_0: \mu_X = \mu_Y$ 

One-Sided Hypothesis:

 $H_1: \mu_X > \mu_Y$  $H_1: \mu_X < \mu_Y$ 

Two-Sided Hypothesis

 $H_1: \mu_X \neq \mu_Y$ 

Use two-sided test unless you know direction a priori

## Takeaways:

- Newyman Pearson Lemma Intuition
- Duality of CI/Hypothesis Tests
- Two Sample Problem
- Solved it for normally distributed data with common unknown variance