# STAT 135 Lecture 20

# Henry Liev

# 7 August 2025

#### Method 0.1 (Permutation Test)

Data from cholesterol-lowering study

Control	Treatment
103	100
115	105
123	121
131	

 $H_0$ : Treatment does nothing

 $H_1$ : Treatment does something Mann Whitney:

- 1. Under  $H_0$ , all the data is interchangeable
- 2. Throw away data, replace with ranks
- 3. Define test statistic T, sum of the ranks
- 4. Calculate t
- 5. Compare to the rank sum distribution

#### Permutation Test:

- 1. Under  $H_0$ , all the data is interchangeable
- 2. Define test statistic  $\Delta = \bar{Y} \bar{X}$
- 3. Calculate  $\delta$
- 4. Use the distribution of  $\Delta$  across all equally likely assignments of data to treatment

We don't need to use all  $\binom{m+n}{n}$ 

- Pick B random assignments of data to treatment
- Calculate  $T_{(b)}$  for each assignment  $(\bar{Y} \bar{X})$
- Estimated *p*-value  $\hat{P} =$  fraction of B assignments for which t > T
- If P is true p-value:  $\mathbb{E}(\hat{p}) = p, SE(\hat{p}) = \sqrt{\frac{p(1-p)}{B}}$  Pick B large enough for desired accuracy
- End result: Totally nonparametric two sample test
- Obviously better than Mann-Whitney

# Method 0.2 (Bootstrap (nonparametric))

Due to Brad Efron (1979)  $X_1, \ldots, X_n \stackrel{iid}{\sim} F$ 

Interested in some parameter  $\theta(F)$ 

Have in mind an estimator  $\hat{\theta} = T(X_1, \dots, X_n)$ 

Want to get  $SE(\hat{\theta})$ 

$$\hat{\theta} = \bar{X}, SE(\bar{X}) = \frac{\sigma}{\sqrt{n}}$$

 $\hat{\theta} = f(\bar{X}), SE(f(\bar{X}))$  Delta Method

Big Idea of nonparametric (bootstrap)

Estimate the unknown F by  $\hat{F}$ 

 $\hat{F}$  is the empirical distribution which probability  $\frac{1}{n}$  on each data point, "pretend the sample is the population"

 $\hat{F}_n$  is a good approximation to F

Want  $SE_F(\hat{\theta})$ , Randomness comes from F

 $SE_{\hat{F}}(\theta^*)$ 

# Example 0.3

 $X_1, \ldots, X_n$  Assume data points are all distinct

Sampling from  $\hat{F}$ 

There are  $N = \binom{2n-1}{n}$  samples of size n drawn with replacement from  $X_1, \ldots, X_n$ 

1. In principle, we could draw all N possible samples of size n from the data

2. For b = 1, ..., N Get resampled data  $X_{(b)}, \theta_{(b)}^* = T(X_{(b)})$ 

$$SE_{\hat{F}} = \sqrt{\frac{1}{N} \sum_{b=1}^{N} (\hat{\theta}_{(b)}^* - \hat{\theta}_{(.)}^*)^2}, \hat{\theta}_{(.)}^* = \frac{1}{N} \sum_{i=1}^{N} \hat{\theta}_{(b)}^*$$

Gives us an estimate of  $SE_F(\theta)$  by replacing F with  $\hat{F}$ 

$$\hat{\theta} = T(x_1, \dots, x_n)$$

Based on randomness in F

$$\theta^* = T(Y_1, \dots, Y_n), Y_i \sim \hat{F}$$
 puts mass  $\frac{1}{n}$  on each  $X_i$ 

Choose a large number B (200, 1000)

For  $b = 1, \ldots, B$ 

Draw resamples  $X_{(b)}$  from  $X_1, \ldots, X_n$ 

Draw a random sample of size n with replacement from  $X_1, \ldots, X_n$ 

$$\widehat{SE} = \sqrt{\frac{1}{B-1} \sum_{b=1}^{B} \left[ \hat{\theta}_{(b)}^* - \hat{\theta}_{(.)}^* \right]^2} \Rightarrow \hat{\theta}_{(.)}^* = \frac{1}{B} \sum_{b=1}^{B} \hat{\theta}_{(b)}^*$$

Nonparametric method for estimating  $SE(\hat{\theta})$  with no assumptions and no theory,  $\hat{\theta} = \text{median}$ Are using "large n",  $\hat{F}_n \approx F$  asymptotic

#### Remark 0.4 (Bootstrap confidence intervals)

Many methods for using bootstrap to get CIs

	Seems to make sense	Automatically works well	Confusing
Basic Bootstrap	✓	×	A little
Percentile Interval	×	✓	No
Bootstrap-t	✓	?	Yes

# Method 0.5 (Basic bootstrap)

Basic bootstraps works on the same principle as parametric bootstrap: Approximating  $\hat{\theta} - \theta \Rightarrow$ 

$$P(a \le \hat{\theta} - \theta \le b) = 0.95 \rightarrow (\hat{\theta} - b, \hat{\theta} - a)$$

Generate  $\theta^*$  parametrically via  $f(x|\hat{\theta})$  (Parametric bootstrap)

- Basic bootstrap "makes" sense by same logic as parametric
- But doesn't work well becaues  $\theta^* \theta$  is not a good estimate of  $\hat{\theta} \theta$  in a nonparametric setting

#### Method 0.6 (Percentile Interval)

As before, generate many instance  $\theta_{(b)}^*$  and take the percentiles to generate confidence intervals for  $\theta$ 

#### Method 0.7 (Bootstrap-t)

Normal data setting

Know that  $Z = \frac{\bar{X} - \mu}{\text{SE}(\bar{X})} \sim T$   $P(t_{0.025} \leq \frac{\bar{X} - \mu}{\text{SE}(\bar{X})} \leq t_{0.975}) = 95\% \rightarrow (\bar{X} - t_{0.975} \text{SE}(\bar{X}), \bar{X} + t_{0.975}) \text{SE}(\bar{X})$ 

Build your own custom t-distribution from the data by bootstrapping  $Z(b) = \frac{\hat{\theta}_{(b)}^* - \hat{\theta}}{\operatorname{SE}^*(b)}$ 

For b = 1, ..., B generate resamples compute Z(b) to generate the confidence intervals above For each bootstrap need to conduct a bootstrap to find  $SE^*(b)$