

STAT 151A Lecture 13

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Remark 0.1 (Multivariate Normal Distribution (MVN))

Scalar (usual) normal: $y \sim \mathcal{N}(\mu, \sigma^2)$

Multivariate normal: $\vec{Y} \sim \mathcal{N}_n(\vec{\mu}, \Sigma)$, $\vec{\mu} \in \mathbb{R}^n$, $\Sigma \in \mathbb{R}^{n \times n}$, $\Sigma \succ 0$

Properties:

- If $\vec{Y} \sim \mathcal{N}_n(\vec{\mu}, \Sigma)$, linear functions of \vec{Y} are also MVN
- If $\vec{Y} \sim \mathcal{N}_n(\vec{\mu}, \Sigma)$, then $Y_j(\mu_j, \sigma_{jj}^2)$
- If \vec{Y} and \vec{Z} are both MVN and independent, $\vec{Y} + \vec{Z}$ is also MVN, $\vec{Y} + \vec{Z} \sim \mathcal{N}_n(\vec{\mu}_Y + \vec{\mu}_Z, \Sigma_{YY} + \Sigma_{ZZ} + \Sigma_{YY})$
- If $\vec{Z} \in \mathbb{R}^m$, $\vec{Y} \in \mathbb{R}^n$, $\begin{pmatrix} \vec{Z} \\ \vec{Y} \end{pmatrix} \sim MVN_{n+m}$ and $\text{Cov}(\vec{Z}, \vec{Y}) \Leftrightarrow \vec{Z} \perp \vec{Y}$

Counterexample: $X \sim \mathcal{N}(0, 1)$: X and X^2 , $\text{Cov}(X, X^2) = 0$

Method 0.2 (Normal Linear Model)

$\vec{Y}_{|X} \sim \mathcal{N}_n(\mathbf{X}\beta, \sigma^2 \mathbb{I}_n)$

$\vec{Y} = \mathbf{X}\beta + \vec{\varepsilon}$, $\vec{\varepsilon} \sim \mathcal{N}_n(\vec{0}, \sigma^2 \mathbb{I}_n)$ Stronger than Gauss-Markov model due to multivariate normal (No covariance)

Why? Back to $\hat{\beta}_j$

Test whether null hypothesis $H_0 : \beta_j = 0$

Use NLM: $\hat{\beta}_j$

$\hat{\beta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{Y}$

Distribution of $\hat{\beta}$ under NLM we get another Multivariate normal

$\hat{\beta} \sim \mathcal{N}_p((\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{X} \beta, (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \sigma^2 \mathbb{I}_n \mathbf{X} (\mathbf{X}^\top \mathbf{X})^{-1}) \Rightarrow \hat{\beta} \sim \mathcal{N}_p(\beta, \sigma^2 (\mathbf{X}^\top \mathbf{X})^{-1})$

Under H_0 :

$\hat{\beta}_j \stackrel{H_0}{\sim} \mathcal{N}(0, \sigma^2 (\mathbf{X}^\top \mathbf{X})_{jj})$

$\frac{\hat{\beta}_j}{\hat{\sigma} \sqrt{(\mathbf{X}^\top \mathbf{X})_{jj}}} \stackrel{H_0}{\sim} \mathcal{N}(0, 1)$

Since we have to estimate the standard deviation, we should instead use the t distribution

Example 0.3 (Derive the t -test for $\hat{\beta}_j$)

What is the t -distribution?

Take a $Z \sim \mathcal{N}(0, 1)$

Take a $W \sim \chi_k^2 \rightarrow$ sum of k independent standard normals, each squared

$$T = \frac{Z}{\sqrt{\frac{W}{k}}} \rightarrow T \sim t_k$$

Goal: Show that under Normal Linear Model and H_0 , $\frac{\hat{\beta}_j}{\hat{\sigma} \sqrt{(\mathbf{X}^\top \mathbf{X})_{jj}^{-1}}} \sim t_{n-p-1}$, $\hat{\sigma}$ is a function of \vec{e}

Steps:

- (a) $\hat{\beta}_j \sim \mathcal{N}(\mu = 0)$
 $\vec{e} \sim \mathcal{N}_n(\mu = 0)$
- (b) $\hat{\beta}$ and \vec{e} are uncorrelated \xrightarrow{NLM} independent
- (c) $\vec{e} \sim \mathcal{N}_n \rightarrow \frac{\vec{e}^\top \vec{e}}{\sigma^2} \sim \chi_{n-p-1}^2$
- (d) Put it all together

Proof.

- (a) $\hat{\beta} \sim \mathcal{N}_{p+1}(\beta, \sigma^2(\mathbf{X}^\top \mathbf{X})^{-1})$ under NLM

Under $H_0 \rightarrow \hat{\beta}_j \sim \mathcal{N}(0, \sigma^2(\mathbf{X}^\top \mathbf{X})_{jj}^{-1})$

$$\vec{e} = (\mathbb{I} - \mathbb{H})\vec{Y}$$

Under NLM:

$$\vec{e} \sim \mathcal{N}_n((\mathbb{I} - \mathbb{H})\mathbf{X}\beta, (\mathbb{I} - \mathbb{H})\sigma^2\mathbb{I}(\mathbb{I} - \mathbb{H})) = \mathcal{N}_n(\vec{0}, \sigma^2(\mathbb{I} - \mathbb{H}))$$

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