# CSUEB - STAT 6305 - Winter 2017 - Prof Yan Zhou

# Homework 5 - Henry Lankin, Gui Larangeira

February 14, 2017

#### HW 5: 16.5, 16.6

### 16.5

- **16.5** A study was designed to evaluate treatments for hypertension. The researchers were concerned that whether the patient smoked may impact the effectiveness of the treatments so they also recorded the number of cigarettes smoked daily by the patients. After one month on the treatment, the treating doctors assigned each patient an index based on blood pressure, cholesterol level, and amount of exercise, which reflected the patient's risk of cardiovascular disease (CVD). The index ranged from 0 to 100 with the higher values indicating a greater risk of CVD. The data are presented here with notation: RISK = risk index for CVD, NOCIG = number of cigarettes smoked daily, C = standard treatment, I = new treatment 1, II = new treatment 2.
  - a) Write a model for the above experiment. Make sure to identify all variables and parameters in your model.

ANCOVA design model:

$$y_{ij} = \beta_0 + \tau_i + \beta_1 x_{ij} + \varepsilon_{ij}$$
  
 $i = 1,2,3; j = 1,2,...,10$ 

Using an ANCOVA treatment structure with:

 $y_{ij}$  – the risk index given to each patient representing the observed response variable: 30 observations from 10 replications of 3 possible treatments

 $\beta_0$  – regression intercept of  $y_{ij}$  on  $x_{ij}$ 

 $\tau_i$  – the effect due to the treatment: 3 treatment levels

 $\beta_1$  – regression (slope) coefficient of  $y_{ij}$  on  $x_{ij}$ 

 $x_{ii}$  – the covariate effect due to the number of cigarettes smoked daily: 30 total values with  $j^{th}$  experimental unit receiving  $i^{th}$  level of treatment

 $\varepsilon_{ii}$  – random error associated with each treatment-cigarette value combination: 30 residual errors

Reduced Model I (ANOVA):

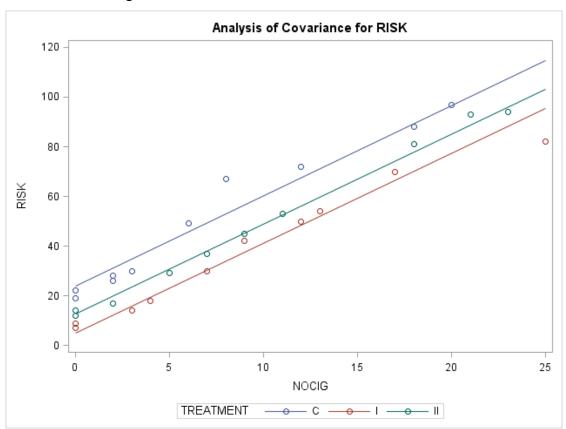
$$y_{ij} = \beta_0 + \tau_i + \varepsilon_{ij}$$

$$i = 1,2,3; j = 1,2,...,10$$

Reduced Model II (ANOVA):

$$y_{ij} = \beta_0 + \beta_1 x_{ij} + \varepsilon_{ij}$$
  
 $i = 1,2,3; j = 1,2,...,10$ 

b) Provide a scatterplot of the data with regression lines which would allow a visual assessment of whether there is a significant relationship between risk index of CVD and the number of cigarettes smoked.



c) From your scatterplot in (b), do the three lines appear to have similar slopes?

As we can see in the scatterplot above, the three lines have very similar slopes, implying the relationship between covariate and the observed response is linear.

## 16.6

#### 16.6 Refer to Exercise 16.5.

a) Test the hypothesis that the relationships between risk index and number of cigarettes have equal slopes for the three treatments at the  $\alpha = 0.05$  level.

Consider the GLM model:

$$E(y) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_1 x_2 + \beta_5 x_1 x_3$$

$$x_1 = covariate\ effect$$

$$x_2 = 1\ if\ treatment\ I, 0\ otherwis \square$$

$$x_3 = 1\ if\ treatment\ II, 0\ otherwis \square$$

Treatment I:

$$E(y) = \beta_0 + \beta_1 x_1 + \beta_2(1) + \beta_3(0) + \beta_4 x_1(1) + \beta_5 x_1(0)$$

$$= \beta_0 + \beta_1 x_1 + \beta_2 + \beta_4 x_1$$

$$= \beta_0 + \beta_2 + (\beta_1 + \beta_4) x_1$$

Treatment II:

$$E(y) = \beta_0 + \beta_1 x_1 + \beta_2(0) + \beta_3(1) + \beta_4 x_1(0) + \beta_5 x_1(1)$$

$$= \beta_0 + \beta_1 x_1 + \beta_3 + \beta_5 x_1$$

$$= \beta_0 + \beta_3 + (\beta_1 + \beta_5) x_1$$

Control:

$$E(y) = \beta_0 + \beta_1 x_1 + \beta_2(0) + \beta_3(0) + \beta_4 x_1(0) + \beta_5 x_1(0)$$
  
= \beta\_0 + \beta\_1 x\_1

Thus, our hypotheses are:

$$H_0$$
:  $\beta_4 = \beta_5 = 0$   
 $H_a$ : at least one  $\beta_i$  is not  $0$ 

Test statistic:

$$F_{1} \ model \ (no \ interaction/same \ slope) : y_{ij} = \beta_{0} + \tau_{i} + \beta_{1}x_{ij} + \varepsilon_{ij}$$
 
$$F_{2} \ model \ (interaction/different \ slopes) : y_{ij} = \beta_{0} + \tau_{i} + \beta_{i}x_{ij} + \varepsilon_{ij}$$
 
$$F = \frac{MS_{\beta_{4}\beta_{5}}}{MSE_{F_{2}}} = \frac{SS_{\beta_{4}\beta_{5}}/(t-1)}{SSE_{F_{2}}/(n-2t)} = \frac{\left(SSE_{F_{1}} - SSE_{F_{2}}\right)/df_{treatment}}{SSE_{F_{2}}/df_{interaction \ error}} = \frac{\left(SSE_{F_{1}} - SSE_{F_{2}}\right)/(t-1)}{SSE_{F_{2}}/(n-2t)}$$

No interaction/equal slopes model:

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	3	23230.96256	7743.65419	327.90	<.0001
Error	26	614.00411	23.61554		
Corrected Total	29	23844.96667			

Interaction model/unequal slopes model:

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	23358.04637	4671.60927	230.26	<.0001
Error	24	486.92029	20.28835		
Corrected Total	29	23844.96667			

$$F = \frac{\left(SSE_{F_1} - SSE_{F_2}\right)/(t-1)}{SSE_{F_2}/(n-2t)} = \frac{(614.00 - 486.92)/(3-1)}{486.92/(30-2\cdot3)} = 3.13$$

Rejection region:

At the  $\alpha = 0.05$  significance level, the rejection region is given by  $F_{2,24}(0.05) = 3.40$ . Conclusion:

Since F = 3.13 < 3.40, we fail to reject the null hypothesis and conclude that the relationships between the risk index and number of cigarettes has equal slopes for each level of treatment.

This can also be seen with the p-value of 0.0619 < 0.05 for the interaction term in the following table, confirming that we fail to reject the null hypothesis and the slopes are equal for each level of treatment.

Source	DF	Type III SS	Mean Square	F Value	Pr > F
NOCIG	1	22164.04588	22164.04588	1092.45	<.0001
TREATMENT	2	469.53955	234.76978	11.57	0.0003

<b>NOCIG*TREATMENT</b> 2 127.08382 63.54191 3.13 0.0619
---

b) Does there appear to be a difference in the mean risk index for the three treatments?

Hypothesis:

$$H_0$$
:  $au_1 = au_2 = au_3 = 0$   
 $H_a$ : at least one  $au_i$  does not equal  $0$ 

Test statistic:

$$\begin{split} Reduced\ model\ F_{3}(regression) \colon & y_{ij} = \beta_{0} + \beta_{1}x_{ij} + \varepsilon_{ij} \\ F_{1}(full\ model) \colon & y_{ij} = \beta_{0} + \tau_{i} + \beta_{1}x_{ij} + \varepsilon_{ij} \\ F = \frac{MST}{MSE_{F_{1}}} = \frac{SST/df_{treatment}}{SSE_{F_{1}}/df_{error}} = \frac{\left(SSE_{F_{3}} - SSE_{F_{1}}\right)/(t-1)}{SSE_{F_{1}}/(n-t-1)} \end{split}$$

Reduced model (regression):

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	21410.85065	21410.85065	246.29	<.0001
Error	28	2434.11602	86.93271		
Corrected Total	29	23844.96667			

Full model AOV table:

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	3	23230.96256	7743.65419	327.90	<.0001
Error	26	614.00411	23.61554		
Corrected Total	29	23844.96667			

$$F = \frac{\left(SSE_{F_3} - SSE_{F_1}\right)/(t-1)}{SSE_{F_1}/(n-t-1)} = \frac{(2434 - 614)/(3-1)}{614/(30-3-1)} = 38.5$$

Rejection region:

At the  $\alpha = 0.05$  significance level, the rejection region is given by  $F_{2,24}(0.05) = 3.40$ .

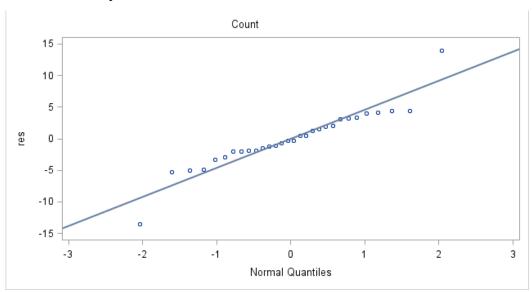
#### Conclusion:

Since F = 38.5 > 3.40, we reject the null hypothesis and conclude that there is not a significant difference in the mean risk index between the three treatments.

This can also be seen with the p-value < 0.0001 < 0.05 for the treatment term in the following table, confirming that we reject the null hypothesis and that there is not a significant difference in the mean risk index between the three treatments.

Source	DF	Type I SS	Mean Square	F Value	Pr > F
NOCIG	1	21410.85065	21410.85065	906.64	<.0001
TREATMENT	2	1820.11191	910.05595	38.54	<.0001

- c) Are the necessary conditions for conducting the tests of hypotheses in parts (a) and (b) satisfied with this data set?
  - 1. Normality of residuals:



Tests for Normality						
Test	Statistic p Valu		alue			
Shapiro-Wilk	W	0.916843	Pr < W	0.0222		

The QQ-plot above shows the residuals do not deviate significantly from the line, but the p-value of 0.02 < 0.05 implies that we should actually reject the null hypothesis of normality and conclude that the residuals may not follow a normal distribution.(???)

# 2. Equal variances:

Levene's Test for Homogeneity of risk_adj Variance ANOVA of Squared Deviations from Group Means						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
TREATMENT	2	3997.0	1998.5	0.91	0.4138	
Error	27	59183.8	2192.0			

The Levene test for homogeneity of variances shows the p-value of 0.4138, implying that we fail reject the null hypothesis of equal variances. Thus, we conclude that the residuals have a common equal variance.

- 3. Independence (test is skipped)
- 4. Each slope: The covariate and the response, risk index, have an equal slope for every level of treatment as shown in question 16.6(b).
- 5. Linearity of covariate: The scatterplots in question 16.6(a) show a straight line is sufficient to describe the relationship between the owe for each level of treatment.

#### SAS code:

```
data hypertension;
                         TREATMENT$
                                          RISK
                                                   NOCIG;
        input Patient
        cards;
                         0
                         2
3
        С
                 49
                         6
                 67
        С
                         8
5
        С
                 72
                         12
6
        С
                 19
        С
                 28
8
        С
                         20
                 97
9
        С
                 88
                         18
        С
10
                 30
11
        Ι
                         0
12
        Ι
                 9
                         0
                 14
1.3
        Ι
                         3
14
        Ι
                 18
15
        Ι
                 30
16
        Ι
                 42
17
        Ι
                         12
                 50
18
        I
                 54
                         13
                 70
19
```

```
20
                        25
              12
14
        ΙI
21
                        0
2.2
        ΙI
                        0
               17
23
       ΙI
24
        ΙI
               37
25
       TT
       ΙI
26
              45
53
                        9
27
        ΙI
                        11
28
       ΙI
29
        ΙI
               93
                        21
              94
30
       ΙI
                        23
run;
proc print data=hypertension;
run;
* full ANCOVA;
proc glm data=hypertension;
        class treatment;
        model risk = nocig treatment / solution;
output out=residuals r=res;
run;
quit;
proc print data = residuals;
run:
* test normality;
proc univariate normal plot data=residuals;
       var res;
run;
quit;
* reduced model 1 ANOVA;
proc glm data=hypertension;
        class treatment;
        model risk = treatment;
        output out=residuals_treat r=res;
run;
quit;
* test normality;
proc univariate normal plot data=residuals_treat;
        var res;
run:
quit;
* reduced model 2 regression;
proc glm data=hypertension;
       model risk = nocig;
        output out=residuals_reg r=res;
run:
quit;
* test normality;
proc univariate normal plot data=residuals_reg;
       var res;
run;
quit;
* interaction model to provide SSEinter for check for equal slopes;
proc glm data=hypertension;
        class treatment;
        model risk = nocig treatment nocig*treatment;
        output out=residuals_interact r=res;
run:
quit;
** results from full ANCOVA gives SSEnointeract and interaction model gives SSEinteract
       used to find F=MSbeta4beta5/MSEinteraction ~ (MSEnointeraction-MSEinteraction)/MSEinteraction
* test normality;
proc univariate normal plot data=residuals_interact;
        var res;
run;
quit;
```

```
* plot scatterplot to see regression relationship -- already plotted in full model output;
proc plot data=hypertension;
        plot risk*nocig=treatment;
run;
quit;
proc means data = hypertension;
       var nocig;
run;
* add new column with risk-mean;
data hypertension; set hypertension;
       risk_adj = risk - 3.62352665*(nocig - 8.5666667);
run;
* levene test;
proc glm data=hypertension;
       class treatment;
        model risk_adj = treatment;
means treatment / hovtest = levene;
run;
quit;
proc glm data=hypertension;
        class treatment;
        model risk = treatment nocig nocig*nocig;
run;
quit;
```