

Oct 25

Conditional Prob

$$\text{recall: } P(A|B) = \frac{P(A \cap B)}{P(B)}$$

and

$$\text{Discrete Joint: } P_{X|Y}(x|y) = \frac{P_{XY}(x,y)}{P_Y(y)} \quad P_Y(y) \neq 0$$

and

$$\text{Cont Joint: } f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)} \leftarrow \begin{matrix} \text{Joint} \\ \text{marginal} \end{matrix}$$

* If X, Y are independent R.V.:

$$f_{X|Y}(x|y) = f_X(x) \quad (y \text{ gives you no new info})$$

Suppose $p(x,y)$ is a joint pmf of X, Y ,

$$p(0,0) = .4 \quad p(1,0) = .1 \quad X = \{0, 1\}$$

$$p(0,1) = .2 \quad p(1,1) = .3 \quad Y = \{0, 1\}$$

Calculate the conditional prob mass function of X given $Y=1$.

$$P_{X|Y}(x|y=1) = \underline{P_{XY}(x,1)}$$

$$P_Y(1)$$

sum X

$$P_Y(1) = p(0,1) + p(1,1) = .2 + .3 = .5$$

pmf: $P_{X|Y}(0|1) = \frac{p(0,1)}{P_Y(1)} = \frac{.2}{.5} = \frac{2/10}{5/10} = \frac{2}{5}$

$$P_{X|Y}(1|1) = \frac{p(1,1)}{P_Y(1)} = \frac{.3}{.5} = \frac{3}{5}$$

ex $f_{XY}(x,y) = \begin{cases} \frac{12}{5}x(2-x-y) & 0 < x < 1, 0 < y < 1 \\ 0 & \text{Otherwise} \end{cases}$

Compute the conditional density (pdf)
of X given $Y=y$ where $0 < y < 1$.

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} \leftarrow f_{XY}(x,y)$$

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} f(x,y) dx = \int_0^1 \frac{12}{5}x(2-x-y) dx \\ &= \frac{12}{5} \int_0^1 (2x - x^2 - xy) dx \end{aligned}$$

$$\begin{aligned}
 &= \frac{12}{5} \left(\frac{2x^2}{a} - \frac{x^3}{3} - \frac{x^2}{a} y \right) \Big|_{x=0}^{x=1} \\
 &= \frac{12}{5} \left(1 - \frac{1}{3} - \frac{y}{2} \right) \\
 &= \frac{12}{5} \left(\frac{2}{3} - \frac{y}{2} \right)
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow f_{X|Y}(x|y) &= \frac{f(x,y)}{f_Y(y)} = \frac{\frac{12}{5}x(2-x-y)}{\frac{12}{5}\left(\frac{2}{3}-\frac{y}{2}\right)} \\
 &= \frac{6x(2-x-y)}{4-3y}
 \end{aligned}$$

Sums of R.V.

X, Y be continuous, independent R.V.,
the pdf of $X+Y$ is given by
the convolution

$$f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$

$$\text{(derivation: } P(X+Y \leq a) = \iint_{\{X+Y \leq a\}} f_{XY}(x,y) dx dy \text{)}$$

Ind

$$f_X \cdot f_Y$$

Ex Suppose we choose 2 #'s at random from $[0, \infty)$ w/ an exp density w/ parameter λ . What is the pdf of sum?

$$f_X(x) = \lambda e^{-\lambda x} \quad x \geq 0$$

$$f_Y(y) = \lambda e^{-\lambda y} \quad y \geq 0$$

$$\underbrace{f_{X+Y}(a)}_{Z} = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$

$$= \int_0^{\infty} \lambda e^{-\lambda(a-y)} \lambda e^{-\lambda y} dy$$

$$= \lambda^2 \int_0^{\infty} e^{-\lambda a} e^{\lambda y} e^{-\lambda y} dy$$

$$e^{\lambda y - \lambda y} = e^0 = 1$$

$$= \lambda^2 e^{-\lambda a} \int_0^{\infty} dy$$

λ, a
constants

$$= \lambda^2 e^{-\lambda a} \int_0^a dy$$

$$= \lambda^2 e^{-\lambda a} y \Big|_0^a = \lambda^2 e^{-\lambda a} \cdot a$$

$0 \leq a - y < \infty$

$$f_{X+Y}(x, y) = \lambda^2 a e^{-\lambda a}$$

Sum of 2
exp R.V. $\lambda = \lambda_1 + \lambda_2$

Sums of Normal R.V.

if $X_i \quad i=1, \dots, n$ are independent normal R.V.

w/ respective μ_i, σ_i^2 , then

$\sum_{i=1}^n X_i$ is normally distributed with

$$\text{mean} = \sum_{i=1}^n \mu_i \quad \& \quad \text{var} = \sum_{i=1}^n \sigma_i^2$$

$\overrightarrow{\text{ex}}$ distribution

$$X \sim \mathcal{N}(\mu_x, \sigma_x^2)$$

then

$$X+Y \sim \mathcal{N}(\mu_x + \mu_y, \sigma_x^2 + \sigma_y^2)$$

$$Y \sim \mathcal{N}(\mu_y, \sigma_y^2)$$

$$X-Y \sim \mathcal{N}(\mu_x - \mu_y, \sigma_x^2 + \sigma_y^2)$$

ex Class 1 score on midterm $C_1 \sim N(95, 25)$
 Class 2 score on " $C_2 \sim N(65, 36)$

a) What is the prob that the combined average grades will be less than 90?

$$C_1 + C_2 \sim N(95 + 65, 25 + 36) = N(160, 61)$$

$$\begin{aligned} P\left(\frac{C_1 + C_2}{2} < 90\right) &= P(C_1 + C_2 < 180) \\ &= P\left(\frac{C_1 + C_2 - 160}{\sqrt{61}} < \frac{180 - 160}{\sqrt{61}}\right) \\ &= P(Z \leq 2.56) = \underline{\Phi}(2.56) \\ &= \underline{.9961} \end{aligned}$$

b) What is the prob that someone's grade from C_2 will have scored higher than

someone in C_1 ? $C_2 - C_1 \sim N(65 - 95, 61) = N(-30, 61)$

$$\begin{aligned} P(C_2 > C_1) &= P(C_2 - C_1 > 0) \\ &= P\left(\frac{C_2 - C_1 - (-30)}{\sqrt{61}}, \frac{0 - (-30)}{\sqrt{61}}\right) \end{aligned}$$

$$P\left(Z > \frac{30}{\sqrt{61}}\right) = P(Z > 3.84)$$

$$= 1 - P(Z \leq 3.84) = 1 - \Phi(3.84)$$

$$= 0.00006$$

Discrete (Sums)

$$P_{X+Y}(a) = \sum_y P_X(x) P_Y(a-x)$$

ex: 2 Poisson λ_1 & λ_2

$$\hookrightarrow \text{Sum: } P(X+Y=k) = \frac{e^{-(\lambda_1+\lambda_2)} (\lambda_1+\lambda_2)^k}{k!}$$

ex: 2 Binomial $X \sim \text{Bin}(n, p)$

$Y \sim \text{Bin}(m, p)$

$(X+Y) \sim \text{Bin}(n+m, p)$

$$P_{X+Y}(X+Y=k) = \binom{n+m}{k} p^k (1-p)^{(n+m)-k}$$

Ch. 7 (Ross) Expectations for Joint R.V.

Previously : $E[x]$ } get marginal pdf/pmf first
 $E[Y]$ } compute as usual

$$\text{Now: } E[g(x, Y)] = \sum_x \sum_y g(x, y) p(x, y) \\ = \int \int g(x, y) f(x, y) dx dy$$

$$\text{if: } g(x, Y) = X + Y : E[X + Y] = \int \int (x + y) f(x, y) dx dy \\ = \int \int x f(x, y) dx dy \\ + \int \int y f(x, y) dx dy$$

$$E[X + Y] = \sum \sum (x + y) p(x, y) \\ = \sum \sum x p(x, y) \\ + \sum \sum y p(x, y)$$

$$\Rightarrow E[X + Y] = E[X] + E[Y]$$

* careful $E[XY] \neq E[X]E[Y]$ in all cases

ex the pmf for X & Y :

		Y
		0 1
X	0	0.2 0.7
	1	0.0 0.1

Find $E[XY] = \sum_{\substack{x \\ i}} \sum_{\substack{y \\ j}} x_i y_j P(x_i, y_j)$

$$\begin{aligned} &= 0 \cdot 0 \cdot p(0,0) + 0 \cdot 1 \cdot p(0,1) \\ &\quad + 1 \cdot 0 \cdot p(1,0) + 1 \cdot 1 \cdot p(1,1) \\ &= p(1,1) = \textcircled{0.1} \end{aligned}$$

ex $f(x,y) = \begin{cases} 10xy^2 & 0 < x < y & 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$

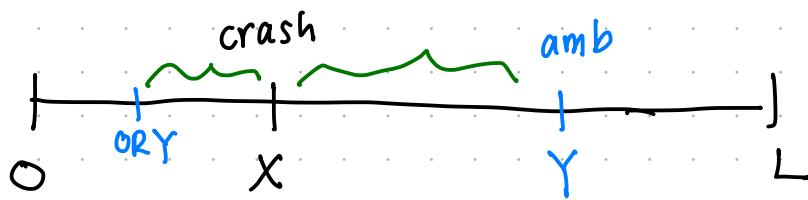
Find $E[XY] = \int_0^1 \int_0^y xy \cdot 10xy^2 dx dy$

$$= \int_0^1 10y^3 \int_0^y x^2 dx dy$$

$$= \int_0^1 10y^3 \cdot \frac{x^3}{3} \Big|_0^y dy$$

$$= \int_0^1 \frac{10}{3} y^6 dy = \frac{10}{3} \frac{y^7}{7} \Big|_0^1 = \frac{10}{3} \cdot \frac{1}{7} = \textcircled{\frac{10}{21}}$$

Ex An accident occurs at a point X that is uniformly distributed on a road of length L . At the time of accident, ambulance is at a location Y that is also uniformly distributed. X & Y indep find the expected distance btwn ambulance & accident.



Want $E[|X-Y|]$

$$|X-Y| = \begin{cases} X-Y \\ -(X-Y) = Y-X \end{cases}$$

X, Y uni $(0, L)$, indep

$$f_X(x) = \begin{cases} \frac{1}{L} & 0 < x < L \\ 0 & \text{Other} \end{cases}$$



$$f_Y(y) = \begin{cases} \frac{1}{L} & 0 < y < L \\ 0 & \text{Other} \end{cases}$$

$$f_{XY}(x,y) = f_X \cdot f_Y = \begin{cases} \frac{1}{L^2} & 0 < x < L, 0 < y < L \\ 0 & \text{Otherwise} \end{cases}$$

$$E[|X-Y|] = \int_0^L \int_0^L |X-Y| \frac{1}{L^2} dy dx = \frac{1}{L^2} \int_0^L \int_0^L |X-Y| dy dx$$

inner integral : $\int_0^L |X-Y| dy = \int_0^X (X-y) dy + \int_X^L (y-X) dy$

$$= \left(XY - \frac{y^2}{2} \right) \Big|_0^X + \left(\frac{y^2}{2} - XY \right) \Big|_X^L$$

$$= X^2 - \frac{X^2}{2} - 0 + \left(\frac{L^2}{2} - LX \right) - \left(\frac{X^2}{2} - X^2 \right)$$

$$= \frac{L^2}{2} - LX + X^2$$

$$= \frac{1}{L^2} \int_0^L \left(\frac{L^2}{2} - LX + X^2 \right) dx = \frac{1}{L^2} \left(\frac{L^2}{2} \cdot X - \frac{LX^2}{2} + \frac{X^3}{3} \right) \Big|_{x=0}^{x=L}$$

$$= \frac{L}{3} = E[|X-Y|]$$

Covariance → tells us information btwn 2 R.V.

$$\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

* if $\text{Cov}(X, Y) > 0$ then X & Y tend to move together

(when X is high, then Y tends to be high)

$\text{Cov}(X, Y) < 0$ then X & Y tend to move opposite

$\text{Cov}(X, Y) = 0$ then X & Y do not consistently move together

→ If X, Y are indep $\rightarrow E[XY] = E[X] \cdot E[Y]$

$$\hookrightarrow \text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 0$$

\hookrightarrow if X, Y are indep, then $\text{Cov}(X, Y) = 0$

* However, if $\text{Cov}(X, Y) = 0$, it doesn't automatically give indep

ex $P(X=0) = P(X=1) = P(X=-1) = \frac{1}{3}$

$$Y = \begin{cases} 0 & \text{if } X \neq 0 \\ 1 & \text{if } X = 0 \end{cases}$$

} dependent
 Y depends on X

$\text{Cov}(X, Y)$?

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

$$E[XY] = \sum_x \sum_y xy p(x,y)$$

$$= 0 \cdot 0 \cdot p(0,0) + 0 \cdot 1 \cdot p(0,1) + 1 \cdot 0 \cdot p(1,0)$$

$$+ 1 \cdot 1 \underbrace{p(1,1)}_0 + -1 \cdot 0 \cdot p(-1,0) + -1 \cdot 1 \underbrace{p(-1,1)}_0$$

X	Y
0	0
0	1
1	0
1	1
-1	0
-1	1

by design of Y

$$= 0$$

$$E[X] = 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3} - 1 \cdot \frac{1}{3} = \frac{1}{3} - \frac{1}{3} = 0$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 0$$

Properties of Covariance X, Y, Z R.V.

a, c constants

b

$$\textcircled{1} \quad \text{Cov}(X, X) = \text{Var}(X)$$

$$\textcircled{2} \quad \text{Cov}(X, Y) = \text{Cov}(Y, X)$$

$$\textcircled{3} \quad \text{Cov}(aX, Y) = a \text{Cov}(X, Y)$$

$$\textcircled{4} \quad \text{Cov}(X+c, Y) = \text{Cov}(X, Y) \quad \text{Cov}(X, Y+c) = \text{Cov}(X, Y)$$

$$\textcircled{5} \quad \text{Cov}(X+Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z)$$

$$\textcircled{6} \quad \text{Var}(aX+bY) = a^2 \text{Var}X + 2ab\text{Cov}(X, Y) + b^2 \text{Var}Y$$

* if X, Y are indep $\rightarrow \text{Cov}(X, Y) = 0$

$$\text{Var}(aX+bY) = a^2 \text{Var}X + b^2 \text{Var}Y$$

Correlation Coefficient (of X & Y R.V.)

↓ greek

$$\rho(x,y) = \frac{\text{Cov}(X,Y)}{\sqrt{\text{Var}(X) \text{Var}(Y)}} \quad -1 \leq \rho(x,y) \leq 1$$

$\rho(x,y) > 0$ "positive correlation"

$\rho(x,y) < 0$ "negative" " "

$= 0$ "uncorrelated"

ex $f(x,y) = \begin{cases} x+y & 0 < x < 1 \quad 0 < y < 1 \\ 0 & \text{otherwise} \end{cases}$

Find

a) $\text{Cov}(X,Y)$

b) $\text{Var}(X)$ & $\text{Var}(Y)$

c) $\rho(x,y)$

$$\Rightarrow E[XY] = \iint_0^1 xy \cdot (x+y) dx dy = \iint_0^1 x^2 y + xy^2 dx dy$$

$$= \left. \int_0^1 \left(\frac{x^3}{3} y + \frac{x^2}{2} y^2 \right) \right|_{x=0}^{x=1} dy = \int_0^1 \left(\frac{y}{3} + \frac{y^2}{2} \right) dy$$

$$= \left(\frac{y^2}{6} + \frac{y^3}{6} \right) \Big|_0^1 = \frac{1}{6} + \frac{1}{6} = \frac{1}{3}$$

$$f_X(x) = \int_0^1 (x+y) dy = \left(xy + \frac{y^2}{2} \right) \Big|_0^1 = x + \frac{1}{2}$$

$$f_Y(y) = \int_0^1 (x+y) dx = y + \frac{1}{2}$$

$$E[X] = \int_0^1 x \left(x + \frac{1}{2} \right) dx = \int_0^1 \left(x^2 + \frac{x}{2} \right) dx = \left(\frac{x^3}{3} + \frac{x^2}{4} \right) \Big|_0^1$$

$$= \frac{1}{3} + \frac{1}{4}$$

$$= \frac{7}{12}$$

$$E[Y] = \int_0^1 y \left(y + \frac{1}{2} \right) dy = \frac{7}{12}$$

$$\Rightarrow \text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

$$= \frac{1}{3} - \frac{7}{12} \cdot \frac{7}{12} = -\frac{1}{144}$$

b) $\text{Var}(X), \text{Var}(Y)$:

$$E[X^2] = \int_0^1 x^2 \left(x + \frac{1}{2} \right) dx = \int_0^1 \left(x^3 + \frac{1}{2} x^2 \right) dx = \frac{5}{12}$$

$$E[Y^2] = \int_0^1 y^2(y + \frac{1}{2}) dy = \frac{5}{12}$$

$$\Rightarrow \text{Var}(X) = E[X^2] - (E[X])^2 \\ = \frac{5}{12} - \left(\frac{7}{12}\right)^2 = \frac{11}{144} = \text{Var}(Y)$$

$$c) \rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X) \cdot \text{Var}(Y)}} = \frac{-\frac{1}{144}}{\sqrt{\frac{11}{144} \cdot \frac{1}{144}}} = \frac{-1}{11}$$

$$\approx -0.0909\dots$$

Cauchy-Schwarz inequality

$$(E[XY])^2 \leq E[X^2] \cdot E[Y^2]$$