第3章:参数估计(续)

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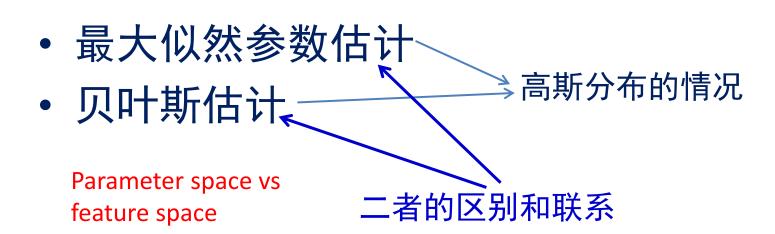
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上次课主要内容回顾

- 离散变量贝叶斯决策
- 复合模式分类





提纲

- 第3章
 - -特征维数问题
 - 期望最大法
 - 隐马尔可夫模型



特征维数问题

- 统计模式分类
 - 特征空间划分
 - 贝叶斯决策:最小风险规则,MAP
- 增加特征有什么好处
 - 判别性: 类别间有差异的特征有助于分类
- 带来什么问题
 - 计算
 - 存储
 - 泛化性能,Overfitting



分类错误率与特征的关系

- 二类高斯分布
 - $-p(\mathbf{x}|\omega_i)^{\sim}N(\mu_i,\Sigma), j=1,2$,等协方差矩阵
 - Bayes error rate

$$P(e) = \frac{1}{\sqrt{2\pi}} \int_{r/2}^{\infty} e^{-u^2/2} du \qquad r^2 = (\mu_1 - \mu_2)^t \Sigma^{-1} (\mu_1 - \mu_2)$$

$$P(error) = P(\mathbf{x} \in \mathcal{R}_2, \omega_1) + P(\mathbf{x} \in \mathcal{R}_1, \omega_2)$$

$$= P(\mathbf{x} \in \mathcal{R}_2 | \omega_1) P(\omega_1) + P(\mathbf{x} \in \mathcal{R}_1 | \omega_2) P(\omega_2)$$

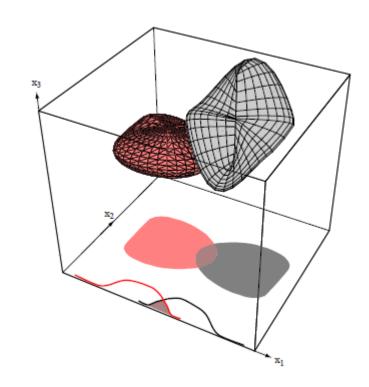
$$= \int_{\mathcal{R}_2} p(\mathbf{x} | \omega_1) P(\omega_1) d\mathbf{x} + \int_{\mathcal{R}_1} p(\mathbf{x} | \omega_2) P(\omega_2) d\mathbf{x}.$$

- Conditionally independent case $\Sigma = diag(\sigma_1^2, ..., \sigma_d^2)$
 - 每一维二类均值之间距离反映区分度,决定错误率
 - 特征增加有助于减小错误率(r²增大)

$$r^2 = \sum_{i=1}^d \left(\frac{\mu_{i1} - \mu_{i2}}{\sigma_i}\right)^2$$



- 特征维数决定可分性的例子
 - 3D空间完全可分
 - 2D和1D投影空间有重叠



然而,增加特征也可能导致分类性能更差,因为有模型估计误差(wrong model)



计算复杂度

- 最大似然估计
 - 高斯分布, d维特征, n个样本
 - 参数估计的复杂度, 主要由Σ决定

$$g(\mathbf{x}) = -\frac{1}{2} (\mathbf{x} - \hat{\boldsymbol{\mu}})^t \widehat{\boldsymbol{\Sigma}}^{-1} (\mathbf{x} - \hat{\boldsymbol{\mu}}) - \underbrace{\frac{O(1)}{d} \ln 2\pi}_{O(n)} - \underbrace{\frac{O(d^2n)}{1} \ln |\widehat{\boldsymbol{\Sigma}}|}_{O(n)} + \underbrace{\ln P(\omega)}_{O(n)}$$

$$\widehat{\boldsymbol{\Sigma}} = \frac{1}{n} \sum_{k=1}^{n} (\mathbf{x}_k - \mathbf{m}_n) (\mathbf{x}_k - \mathbf{m}_n)^t$$

• 参数存储复杂度

$$c(d+d(d+1)/2)$$

• 分类复杂度?



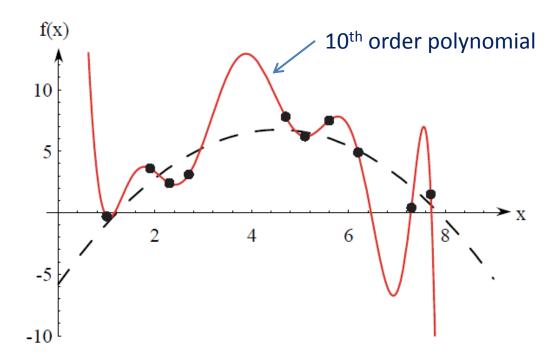
过拟合(Overfitting)

- Overfitting
 - 特征维数高、训练样本少导致模型参数估计不准确
 - 比如协方差矩阵需要样本数在d以上
- 克服办法
 - 特征降维: 特征提取(变换)、特征选择
 - 参数共享/平滑
 - 共享协方差矩阵Σ₀
 - Shrinkage (a.k.a. Regularized Discriminant Analysis)

$$\Sigma_i(\alpha) = \frac{(1-\alpha)n_i\Sigma_i + \alpha n\Sigma}{(1-\alpha)n_i + \alpha n}$$
$$\Sigma(\beta) = (1-\beta)\Sigma + \beta I$$



• 过拟合的例子



$$f(x) = ax^2 + bx + c + \epsilon$$
 where $p(\epsilon) \sim N(0, \sigma^2)$

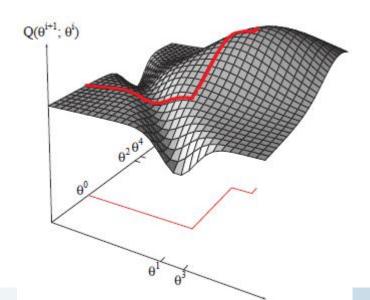


期望-最大法(EM)

- 数据缺失情况下的参数估计
 - Good features, missing/bad features $\mathbf{x}_k = \{\mathbf{x}_{kg}, \mathbf{x}_{kb}\}$

$$\mathcal{D} = \{\mathbf{x}_1, ..., \mathbf{x}_n\} = \mathcal{D}_g \cup \mathcal{D}_b$$

- 已知参数θ′情况下估计新参数θ
 - 对缺失数据求期望(marginalize) $\max Q(\theta; \; \theta^i) = \mathcal{E}_{\mathcal{D}_b}[\ln p(\mathcal{D}_q, \mathcal{D}_b; \; \theta) | \mathcal{D}_q; \; \theta^i]$





Expectation-Maximization (EM)

Algorithm 1 (Expectation-Maximization)

```
1 begin initialize \theta^{0}, T, i = 0
2 do i \leftarrow i + 1
3 E step: compute Q(\theta; \theta^{i})
5 M step: \theta^{i+1} \leftarrow \arg \max_{\theta} Q(\theta; \theta^{i})
6 until Q(\theta^{i+1}; \theta^{i}) - Q(\theta^{i}; \theta^{i-1}) \leq T
7 return \hat{\theta} \leftarrow \theta^{i+1}
8 end
```

The EM algorithm guarantees that the log-likelihood of good data increases monotonically.



Example: EM for a 2D Gaussian

$$\mathcal{D} = \{\mathbf{x}_{1}, \mathbf{x}_{2}, \mathbf{x}_{3}, \mathbf{x}_{4}\} = \left\{\binom{0}{2}, \binom{1}{0}, \binom{2}{2}, \binom{*}{4}\right\}$$
 parameters $\theta = \begin{pmatrix} \mu_{1} \\ \mu_{2} \\ \sigma_{1}^{2} \\ \sigma_{2}^{2} \end{pmatrix}$ initially $\theta^{0} = \begin{pmatrix} 0 \\ 0 \\ 1 \\ 1 \end{pmatrix}$
$$Q(\theta; \, \theta^{0}) = \mathcal{E}_{x_{41}}[\ln p(\mathbf{x}_{g}, \mathbf{x}_{b}; \, \theta | \theta^{0}; \, \mathcal{D}_{g})]$$

$$= \int_{-\infty}^{\infty} \left[\sum_{k=1}^{3} \ln p(\mathbf{x}_{k} | \theta) + \ln p(\mathbf{x}_{4} | \theta)\right] p(x_{41} | \theta^{0}; \, x_{42} = 4) \, dx_{41}$$

$$= \sum_{k=1}^{3} [\ln p(\mathbf{x}_{k} | \theta)] + \int_{-\infty}^{\infty} \ln p\left(\binom{x_{41}}{4} \middle| \theta\right) \frac{p\left(\binom{x_{41}}{4} \middle| \theta^{0}\right)}{\left(\int_{-\infty}^{\infty} p\left(\binom{x_{41}}{4} \middle| \theta^{0}\right) dx_{41}^{\prime}\right)} dx_{41}$$

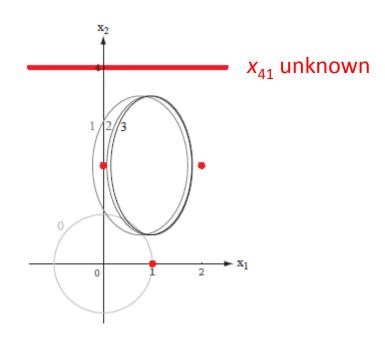
$$= \sum_{k=1}^{3} [\ln p(\mathbf{x}_{k} | \theta)] + \frac{1}{K} \int_{-\infty}^{\infty} \ln p\left(\binom{x_{41}}{4} \middle| \theta\right) \frac{1}{2\pi \left|\binom{1}{0} \right|} \exp\left[-\frac{1}{2}(x_{41}^{2} + 4^{2})\right] dx_{41}$$

$$= \sum_{k=1}^{3} [\ln p(\mathbf{x}_{k} | \theta)] - \frac{1 + \mu_{1}^{2}}{2\sigma_{1}^{2}} - \frac{(4 - \mu_{2})^{2}}{2\sigma_{2}^{2}} - \ln (2\pi\sigma_{1}\sigma_{2}).$$

$$\max \sum_{k=1}^{3} [\ln p(\mathbf{x}_{k}|\boldsymbol{\theta})] - \frac{1 + \mu_{1}^{2}}{2\sigma_{1}^{2}} - \frac{(4 - \mu_{2})^{2}}{2\sigma_{2}^{2}} - \ln (2\pi\sigma_{1}\sigma_{2})$$

$$\boldsymbol{\theta}^{1} = \begin{pmatrix} 0.75 \\ 2.0 \\ 0.938 \\ 2.0 \end{pmatrix}$$

After 3 iterations
$$\mu = \begin{pmatrix} 1.0 \\ 2.0 \end{pmatrix}$$
, and $\Sigma = \begin{pmatrix} 0.667 & 0 \\ 0 & 2.0 \end{pmatrix}$



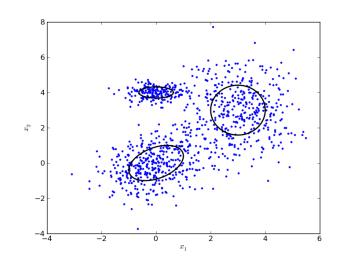
- EM for Gaussian mixture
 - 参数型概率密度函数,可以表示复杂的分布

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k p(\mathbf{x} \mid \theta_k)$$

subject to
$$\sum_{k=1}^{K} \pi_k = 1$$

Gaussian component

$$p(\mathbf{x} \mid \theta_k) = \mathcal{N}(\mathbf{x} \mid \mu_k, \Sigma_k)$$



- 参数估计: Maximum Likelihood (ML)

$$\max LL = \log \prod_{n=1}^{N} p(\mathbf{x}_n) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k p(\mathbf{x}_n \mid \theta_k)$$

$$\nabla_{\pi_k} LL = 0$$
, $\nabla_{\mu_k} LL = 0$, $\nabla_{\Sigma_k} LL = 0$

不能解析求解



EM Algorithm for Gaussian mixture

Incomplete data X, complete data {X,Z}

$$z_{nk} \in \{0,1\}, \quad k = 1, ..., K$$

Expectation of complete data log-likelihood

$$Q(\Theta, \Theta^{old}) = \sum_{\mathbf{Z}} \log p(\mathbf{X}, \mathbf{Z} \mid \Theta) p(\mathbf{Z} \mid \mathbf{X}, \Theta^{old})$$

- 1. Choose an initial set of parameters for Θ^{old}
- 2. E-step: Evaluate p(Z|X, Θ^{old})
- 3. M-step: Update parameters

$$\Theta^{new} = \arg\max_{\Theta} Q(\Theta, \Theta^{old})$$

4. If convergence condition is not satisfied

$$\Theta^{old} \leftarrow \Theta^{new}$$

5. Return to E-step

EM Algorithm for Gaussian mixture

E-step
$$p(\mathbf{X}, \mathbf{Z} \mid \Theta) = \prod_{n=1}^{N} \prod_{k=1}^{K} \pi_{k}^{z_{nk}} \mathcal{N}(\mathbf{x}_{n} \mid \mu_{k}, \Sigma_{k})^{z_{nk}}$$

$$Q(\Theta, \Theta^{old}) = \mathbf{E}_{\mathbf{Z}}[\log p(\mathbf{X}, \mathbf{Z} \mid \Theta)] = \sum_{n=1}^{N} \sum_{k=1}^{K} \gamma(z_{nk}) \{\log \pi_{k} + \log \mathcal{N}(\mathbf{x} \mid \mu_{k}, \Sigma_{k})\}$$

$$\gamma(z_{nk}) = P(z_{nk} = 1 \mid \mathbf{x}_{n}) = \frac{\pi_{k} N(\mathbf{x}_{n} \mid \mu_{k}, \Sigma_{k})}{\sum_{j=1}^{K} \pi_{j} N(\mathbf{x}_{n} \mid \mu_{j}, \Sigma_{j})}$$

M-step
$$\nabla_{\pi_k} Q = 0$$
, $\nabla_{\mu_k} Q = 0$, $\nabla_{\Sigma_k} Q = 0$

$$\mu_k^{new} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$$

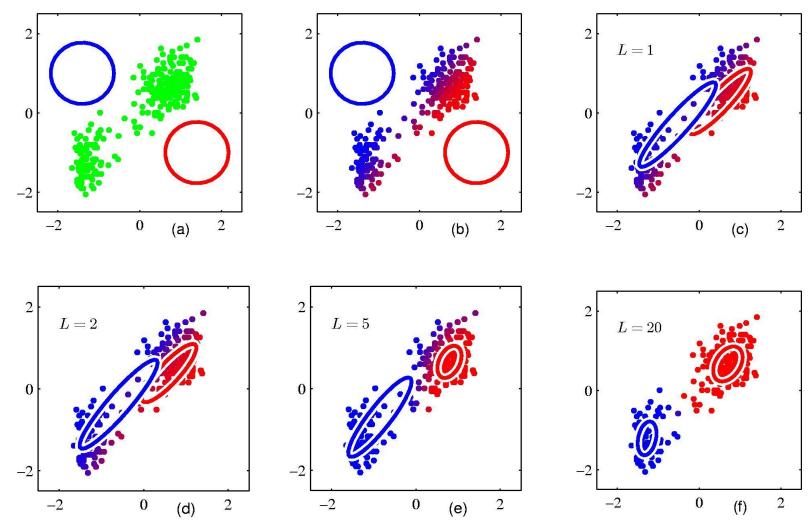
$$\sum_k^{new} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \mu_k^{new}) (\mathbf{x}_n - \mu_k^{new})^T$$

$$\pi_k^{new} = \frac{N_k}{N}$$

$$N_k = \sum_{n=1}^N \gamma(z_{nk})$$



An example, from (C.M. Bishop, *Pattern Recognition and Machine Learning*, 2006. Figure 9.8)



Break



隐马尔可夫模型

- Sequential (Temporal) Pattern
 - Variable length
 - Distortion



- Ambiguous boundary between primitives (symbols)
- Bayesian Classification
 - Sequence of patterns (observations) $\mathbf{O} = O_1 O_2 \cdots O_T$
 - Sequence class (states) $\mathbf{q} = q_1 q_2 \cdots q_T$
 - Posterior probability $P(\mathbf{q} \mid \mathbf{O}) = \frac{p(\mathbf{O} \mid \mathbf{q})P(\mathbf{q})}{p(\mathbf{O})}$
- Hidden Markov Model (HMM)
 - Model $p(\mathbf{O}|\mathbf{q}), p(\mathbf{O},\mathbf{q})$



Markov Chain

Sequence of States (classes)

$$P(q_1 q_2 \cdots q_T) = P(q_1) P(q_2 | q_1) P(q_3 | q_1 q_2) \cdots P(q_T | q_1 \cdots q_{T-1})$$

$$q_t \in \{S_1, \dots, S_N\}$$

First-Order Markov

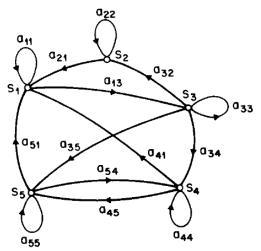
$$P(q_{t} = S_{j} | q_{t-1} = S_{i}, q_{t-2} = S_{k}, \dots) = P(q_{t} = S_{j} | q_{t-1} = S_{i})$$

$$P(q_{1}q_{2} \dots q_{T}) = P(q_{1})P(q_{2} | q_{1})P(q_{3} | q_{2}) \dots P(q_{T} | q_{T-1})$$

State transition probabilities

$$a_{ij} = P(q_t = S_j | q_{t-1} = S_i), \quad 1 \le i, j \le N$$

$$\sum_{i=1}^{N} a_{ij} = 1$$





State duration (self-transition)

$$O = \{S_i, S_i, S_i, \dots, S_i, S_j \neq S_i\}$$
1 2 3 d d+1
$$P(O \mid \text{Model}, q_1 = S_i) = (a_{ii})^{d-1} (1 - a_{ii}) = p_i(d)$$

Expected duration of specific state

$$\overline{d}_{i} = \sum_{d=1}^{\infty} d p_{i}(d)$$

$$= \sum_{d=1}^{\infty} d (a_{ii})^{d-1} (1 - a_{ii}) = \frac{1}{1 - a_{ii}}$$

Example: Transition of Weather

State 1: rain or (snow)
State 2: cloudy
A =
$$\{a_{ij}\}$$
 = $\begin{bmatrix} 0.4 & 0.3 & 0.3 \\ 0.2 & 0.6 & 0.2 \\ 0.1 & 0.1 & 0.8 \end{bmatrix}$

Expected number of days for sunny and cloudy?



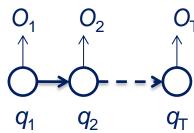
Hidden Markov Model (HMM)

- Markov Chain: States are Observable
- Hidden States: An Example
 - Imagine you are in a un-windowed room, cannot see the weather outside. Instead, you can guess the weather from the temperature and humidity in room
 - Observations: temperature, humidity
 - Hidden states: weather
 - Hidden Markov Model (HMM): Doubly embedded stochastic process

$$P(O_1, O_2, ..., O_T)$$

 $P(q_1, q_2, ..., q_T)$

Infer states from observations



$$q_i \in \{S_1, S_2, \dots, S_N\}$$



Elements of an HMM

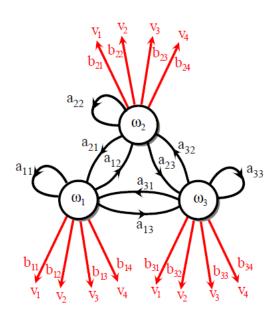
$$\lambda = (A, B, \pi)$$

- N: number of states in the model, $S=\{S_1, S_2, ..., S_N\}$
- M: number of observation symbols, $V=\{v_1, v_2, ..., v_M\}$
- State transition probability distribution $A = \{a_{ij}\}$

$$a_{ij} = P(q_{t+1} = S_j | q_t = S_i), \quad 1 \le i, j \le N$$

- Observation symbol (emission) probability distribution $B=\{b_j(k)\}$ $b_j(k)=P(v_k \text{ at } t \mid q_t=S_j), \ 1 \le j \le N, \ 1 \le k \le M$
- Initial state distribution $\pi = {\pi_i}$

$$\pi_i = P(q_1 = S_i), 1 \le i \le N$$





Three Basic Problems of HMM

– Problem 1 (Evaluation):

How to efficiently compute the probability of observation sequence $P(O|\lambda)$

– Problem 2 (Decoding):

How to choose the best state sequence responding to an observation sequence

– Problem 3 (Training):

How to estimate the model parameters



Evaluation Problem

- Given model $\lambda = (\mathbf{A}, \mathbf{B}, \pi)$ and observation sequence $O = O_1 O_2 ... O_T$, compute $P(O \mid \lambda)$
 - Direct computation

$$P(O \mid \lambda) = \sum_{all \ Q} P(O \mid Q, \lambda) P(Q \mid \lambda)$$

$$= \sum_{q_1, q_2, \dots, q_T} \pi_{q_1} b_{q_1}(O_1) a_{q_1 q_2} b_{q_2}(O_2) \cdots a_{q_{T-1} q_T} b_{q_T}(O_T)$$

Conditional independence
$$P(O \mid Q, \lambda) = \prod_{t=1}^{T} P(O_t \mid q_t, \lambda)$$

$$= b_{q_1}(O_1)b_{q_2}(O_2)\cdots b_{q_T}(O_T)$$

$$P(Q \mid \lambda) = \pi_{q_1}a_{q_1q_2}a_{q_2q_3}\cdots a_{q_{T-1}q_T}$$

$$= b_{q_1}(O_1)b_{q_2}(O_2)\cdots b_{q_T}(O_T)$$
Markov chain of states

– Complexity: O(2TN^T)!



Evaluation: Forward Procedure

Define forward variable

$$\alpha_t(i) = P(O_1 O_2 \cdots O_t, q_t = S_i \mid \lambda)$$

- Initialization $\alpha_1(i) = \pi_i b_i(O_1), 1 \le i \le N$
- Induction

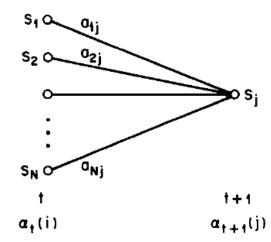
$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_{t}(i) a_{ij}\right] b_{j}(O_{t+1}),$$

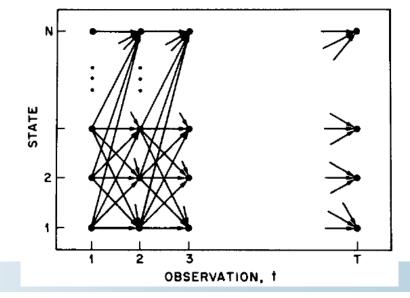
$$1 \le t \le T - 1, \qquad 1 \le j \le N.$$

Termination

$$P(O \mid \lambda) = \sum_{i=1}^{N} \alpha_{T}(i)$$

Complexity: O(TN²)







Evaluation: Backward Procedure

Define backward variable

$$\beta_t(i) = P(O_{t+1}, \dots, O_T \mid q_t = S_i, \lambda)$$

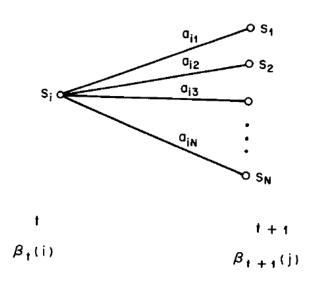
Initialization

$$\beta_T(i) = 1, \quad 1 \le i \le N$$

Induction

$$\beta_{t}(i) = \sum_{j=1}^{N} a_{ij} b_{j}(O_{t+1}) \beta_{t+1}(j),$$

$$1 \le t \le T - 1, \quad 1 \le i \le N$$



Termination

$$P(O \mid \lambda) = \sum_{i=1}^{N} \pi_i b_i(O_1) \beta_1(i) = \sum_{i=1}^{N} \alpha_1(i) \beta_1(i)$$

– Complexity?



Decoding Problem

- This is Pattern Recognition
- Optimal Sequence of States

$$\max_{q_1q_2\cdots q_T} P(q_1q_2\cdots q_T \mid O, \lambda) = \max_{q_1q_2\cdots q_T} P(q_1q_2\cdots q_T, O \mid \lambda)$$

- Viterbi Algorithm (DP)
 - Define variable

$$\delta_t(i) = \max_{q_1, q_2, \dots, q_{t-1}} P(q_1 q_2 \dots q_t = i, O_1 O_2 \dots O_t / \lambda)$$

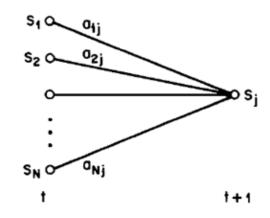
- DP

$$\delta_{t+1}(j) = \left[\max_{i} \delta_{t}(i) a_{ij} \right] \cdot b_{j}(O_{t+1})$$

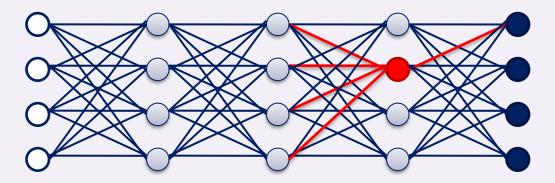
Initialization

$$\delta_1(i) = \pi_i b_i(O_1), \quad 1 \le i \le N$$

$$\psi_1(i) = 0.$$



- Appendix: Dynamic Programming (DP) Principle (Bellman Principle of Optimality)
 - The best path through a particular, intermediate place is the best way from start to it, followed by the best way from it to the goal.
 - Implication: from multiple ways reaching an intermediate place, only retain the best one
 - Often used in sequence matching and HMMs



$$\delta_t(i) = \max_{q_1, q_2, \dots, q_{t-1}} P(q_1 q_2 \cdots q_t = i, O_1 O_2 \cdots O_t \mid \lambda)$$

Viterbi Algorithm (Cont.)

- Recursion

$$\delta_{t}(j) = \left[\max_{1 \le i \le N} \delta_{t-1}(i) a_{ij}\right] b_{j}(O_{t}),$$

$$\psi_{t}(j) = \arg\max_{1 \le i \le N} \delta_{t-1}(i) a_{ij},$$

$$2 \le t \le T, \quad 1 \le j \le N$$

Termination

$$P^* = \max_{1 \le i \le N} \delta_T(i)$$

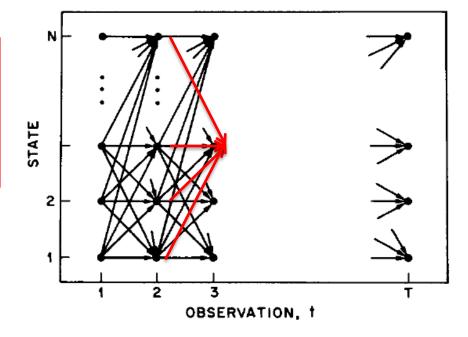
$$q_T^* = \arg\max_{1 \le i \le N} \delta_T(i)$$

Backtracking

$$q_t^* = \psi_{t+1}(q_{t+1}^*),$$

 $1 \le t \le T - 1$

- Complexity: $O(TN^2)$



Training Problem

- $\max P(O | \lambda)$ Maximum Likelihood (ML)
- Baum-Welch Algorithm (EM)

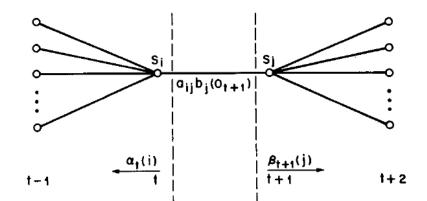
$$\max_{\overline{\lambda}} Q(\lambda, \overline{\lambda}) = \sum_{\mathcal{Q}} \log P(Q, O \mid \overline{\lambda}) P(Q, O \mid \lambda)$$
- Define variable
$$\xi_t(i, j) = P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda)$$

$$\xi_t(i,j) = P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda)$$

$$\xi_{t}(i,j) = \frac{\alpha_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}{P(O|\lambda)} P(O,q_{t} = S_{i}, q_{t+1} = S_{j}|\lambda)$$

$$= \frac{\alpha_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}{\sum_{i=1}^{N}\sum_{j=1}^{N}\alpha_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}$$

$$= \frac{\alpha_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}{\sum_{i=1}^{N}\sum_{j=1}^{N}\alpha_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}$$



Define probability

$$\gamma_{t}(i) = P(q_{t} = S_{i} \mid O, \lambda) = \frac{\alpha_{t}(i)\beta_{t}(i)}{P(O \mid \lambda)} = \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{i=1}^{N} \alpha_{t}(i)\beta_{t}(i)} = \sum_{j=1}^{N} \xi_{t}(i, j)$$

Baum-Welch Algorithm (Cont.)

Reestimation formulas

 $\bar{\pi}_i$ = expected frequency (number of times) in state S_i at time (t=1)= $\gamma_1(i)$

$$\overline{a}_{ij} = \frac{\text{expected number of transitions from state } S_i \text{ to state } S_j}{\text{expected number of transitions from state } S_i}$$

$$= \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \gamma_t(i)} = \frac{\sum_{t=1}^{T-1} \alpha_t(i) a_{ij} b_j(O_{t+1}) \beta_{t+1}(j)}{\sum_{t=1}^{T-1} \alpha_t(i) \beta_t(i)}$$

$$\overline{b}_{j}(k) = \frac{\text{expected number of times in state } j \text{ and observing symbol } v_{k}}{\text{expected number of times in state } j}$$

$$= \frac{\sum_{t=1, \text{ s.t. } O_t = v_k}^{T-1} \gamma_t(j)}{\sum_{t=1}^{T-1} \gamma_t(j)} = \frac{\sum_{t=1, \text{ s.t. } O_t = v_k}^{T} \alpha_t(j) \beta_t(j)}{\sum_{t=1}^{T} \alpha_t(j) \beta_t(j)}$$



Continuous Density HMM

- Handling Continuous Observations
 - Continuous features: vector \boldsymbol{O}_{T}
 - Discretization: vector quantization (VQ)
 - Each vector replaced with its closest codevector, which is viewed as a symbol
 - Small codebook: distortion
 - Large codebook: large data required in emission probability estimation
 - Continuous emission density: Gaussian mixture (GM)

$$b_{j}(O) = \sum_{m=1}^{M} c_{jm} \mathcal{N}(O; \mu_{jm}, U_{jm}), \quad 1 \leq j \leq N$$

Parameter Estimation of Continuous HMM (omitted)

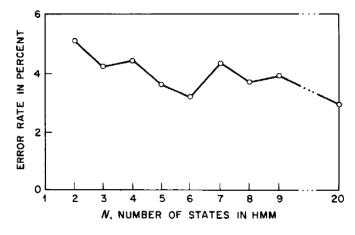


Application to Speech Recognition

- Isolated Word Recognition
 - Given HMM λ^{v} for each word in the vocabulary
 - Input observation sequence O, Bayes decision (assuming equal prior probabilities)

$$v^* = \arg \max_{1 \le v \le V} P(\lambda^v \mid O) = \arg \max_{1 \le v \le V} P(O \mid \lambda^v)$$

- Acoustic features (O_t) (details omitted)
- Vector quantization (discrete observation symbols)
- Choice of model parameters
 - Number of states: empirical, correspond roughly to the number of sounds (phonemes) in a word, maybe equal for all word models
 - Number of components in GM



Extensions of HMM

- Hybrid HMM/Neural
 - HMM: parametric $b_j(O_t)=p(O_t|q_t=S_j)$, conditional independence
 - Neural: discriminative emission probability $p(q_t=S_i|O_t)$
 - Neural network outputs approximate posterior probabilities
 - Replace $p(\mathbf{x}_t|q_t)$ with $p(q_t|\mathbf{x}_t)/p(q_t)$

$$\frac{p(\mathbf{x}_t \mid q_t)}{p(\mathbf{x}_t)} = \frac{p(q_t \mid \mathbf{x}_t)}{P(q_t)}$$

ANN may input multiple frames to learn the correlation

Latest: deep neural networks



讨论

- 特征维数与过拟合
 - 克服过拟合的方法?
- 期望最大法(EM)
 - 对数似然度对缺失数据的期望
 - EM for Gaussian mixture
- 隐马尔可夫模型(HMM)
 - Three basic problems
 - Viterbi Algorithm
 - Extensions



下次课内容

- 第4章 非参数法
 - 密度估计
 - Parzen窗方法
 - K近邻估计
 - 最近邻规则
 - 距离度量
 - Reduced Coulomb Energy Network
 - Approximation by Series Expansion

