Actividad 1: Simulación Estocástica

Curso: Temas Selectos I: O25 LAT4032 1
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Funciones para graficar

```
[1]: import numpy as np
     import matplotlib.pyplot as plt
     import seaborn as sns
     sns.set(style='whitegrid')
[2]: def histograma(muestras, montecarlo, bins=50):
         fig, ax = plt.subplots(figsize=(10, 6))
         n, _, _ = ax.hist(muestras, bins=bins, alpha=0.7)
         ax.axvline(montecarlo, color='r', linestyle='--',
                    linewidth=2, zorder=3, label=f'Monte Carlo {montecarlo:.4f}')
         xmin, xmax = ax.get_xlim()
         if montecarlo < xmin or montecarlo > xmax:
             ax.set_xlim(min(xmin, montecarlo), max(xmax, montecarlo))
         ax.set_title('Muestras')
         ax.set_xlabel('Valor')
         ax.set_ylabel('Frecuencia')
         ax.legend(facecolor='white', edgecolor='none')
         plt.tight_layout()
         plt.show()
[3]: def tlc(muestras, valor_verdadero=None):
         k = len(muestras)
         medias = np.cumsum(muestras) / np.arange(1, k + 1)
         plt.figure(figsize=(10, 6))
         plt.plot(medias)
         plt.xlabel('Número de simulaciones')
         plt.ylabel('Media acumulada')
         plt.title(f'Convergencia de la media Monte Carlo')
         plt.axhline(y=medias[-1], color='r', linestyle='--', label=f'Media final_
      \hookrightarrow {medias[-1]:.4f}')
         if valor_verdadero is not None:
             plt.axhline(y=valor_verdadero, color='g', linestyle='--', label=f'Valor_
      ⇔verdadero {valor_verdadero:.4f}')
         plt.legend(facecolor='white', edgecolor='none')
         plt.gca().yaxis.set_ticks_position('right')
         plt.gca().yaxis.set_label_position('right')
         plt.show()
```

```
Si x_0 = 5 y x_n = 2x_{n-1} \mod 150. Encontrar x_1, \dots, x_{10}.
                                               x_n=ax_{n-1} \bmod m
     $$ 10 = 2.5 \mod 150
     20 = 10 \cdot 5 \bmod 150 \setminus
     40 = 20.5 \mod 150 \setminus
     80 = 40.5 \mod 150
     10 = 80.5 \mod 150
     $$
[4]: pseudoaleatorios = []
      x0 = 5
      a = 2
      m = 150
      for i in range(10):
           xn = (a * x0) % m
           x0 = xn
           pseudoaleatorios.append(xn)
      pseudoaleatorios
```

 $\int_0^1 \exp(e^x) \, dx$

Sea

$$\theta = \int_0^1 \exp(e^x) \, dx.$$

Reescritura como valor esperado con $U \sim \text{Unif}(0, 1)$:

$$\theta = \mathbb{E}\big[\exp(e^U)\big] \,.$$

Estimador Monte Carlo con $u_1, \ldots, u_K \stackrel{\text{iid}}{\sim} \text{Unif}(0, 1)$:

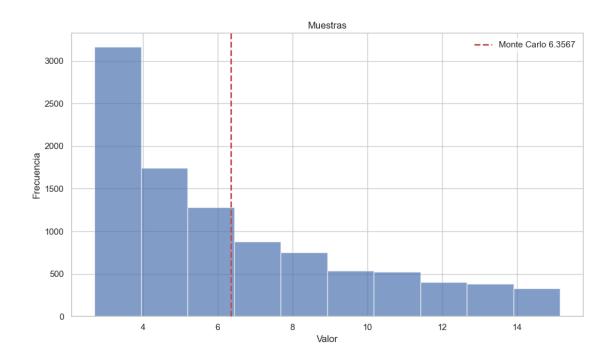
$$\widehat{\theta}_K = \frac{1}{K} \sum_{i=1}^K \exp(e^{u_i}).$$

- [5]: def h(u):
 return np.exp(np.exp(u))

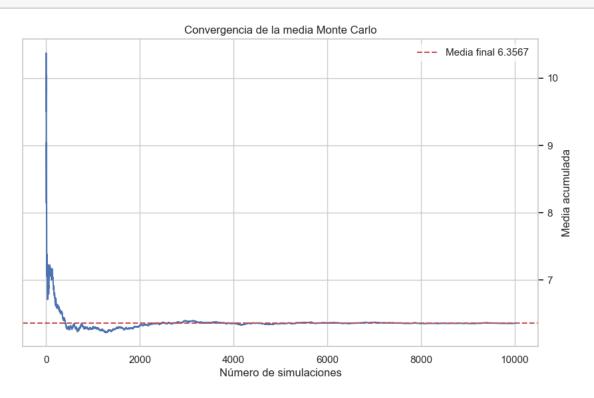
 k = 10000

 u = np.random.random(k)

 muestras = h(u)
 montecarlo = muestras.mean()
 montecarlo
- [5]: np.float64(6.35668829491894)
- [6]: histograma(muestras, montecarlo, bins=10)



[7]: tlc(muestras)



 $\int_{-2}^{2} e^{x+x^2} \, dx$

Sea

$$\theta = \int_{-2}^{2} e^{x+x^2} dx.$$

Cambio de variable a ([0,1]):

$$u = \frac{x - (-2)}{2 - (-2)} = \frac{x + 2}{4}, \qquad x = -2 + 4u, \qquad dx = 4 du.$$

Entonces

$$\theta = \int_0^1 4 \exp[(-2 + 4u) + (-2 + 4u)^2] du.$$

Forma de valor esperado con $U \sim \text{Unif}(0, 1)$:

$$\theta = \mathbb{E}[g(U)], \qquad g(u) = 4 \exp[(-2 + 4u) + (-2 + 4u)^2].$$

Estimador Monte Carlo:

$$\widehat{\theta}_K = \frac{1}{K} \sum_{i=1}^K g(u_i), \qquad u_i \stackrel{iid}{\sim} \text{Unif}(0,1).$$

[8]: def h(u):

return (b-a)*np.exp(a+(b-a)*u + (a+(b-a)*u)**2)

k = 10000

a = -2

b = 2

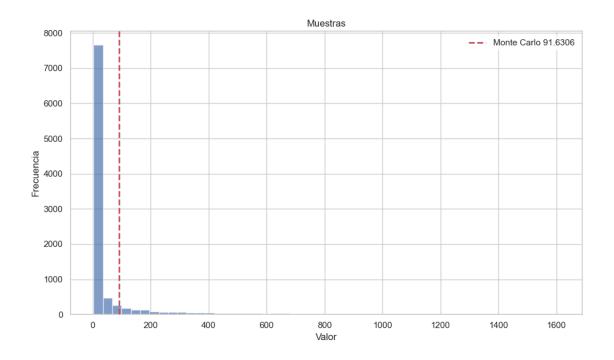
u = np.random.random(k)

muestras = h(u)

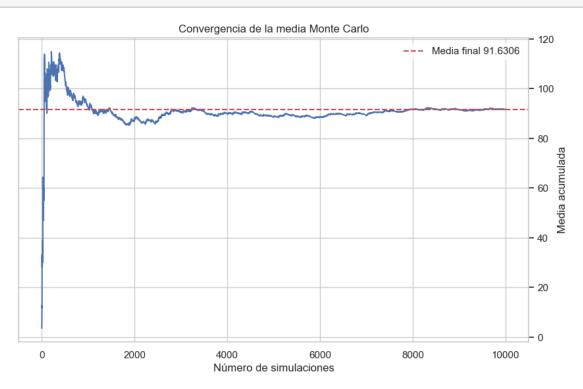
montecarlo = muestras.mean()

montecarlo

- [8]: np.float64(91.6305599447782)
- [9]: histograma(muestras, montecarlo)



[10]: tlc(muestras)



$$\int_0^\infty \frac{x}{(1+x^2)^2} \, dx$$

Estimación Monte Carlo

Sea:

$$\theta = \int_0^\infty \frac{x}{(1+x^2)^2} \, dx.$$

Cambio:

$$y = \frac{1}{x+1}$$
, $dy = -\frac{dx}{(x+1)^2} = -y^2 dx$.

Entonces:

$$\theta = \int_0^1 h(y) \, dy, \qquad h(y) = \frac{g\left(\frac{1}{y} - 1\right)}{y^2}, \quad g(x) = \frac{x}{(1 + x^2)^2}.$$

Forma de esperanza con $U \sim \text{Unif}(0, 1)$:

$$\theta = \mathbb{E}[h(U)].$$

Estimador Monte Carlo:

$$\widehat{\theta}_K = \frac{1}{K} \sum_{i=1}^K h(u_i), \quad u_i \stackrel{\text{iid}}{\sim} \text{Unif}(0,1).$$

[11]: def g(x): return x / (1 + x**2)**2

def h(u):
 return g(1/u-1)/u**2

 $k = 10_{-000}$

u = np.random.random(k)

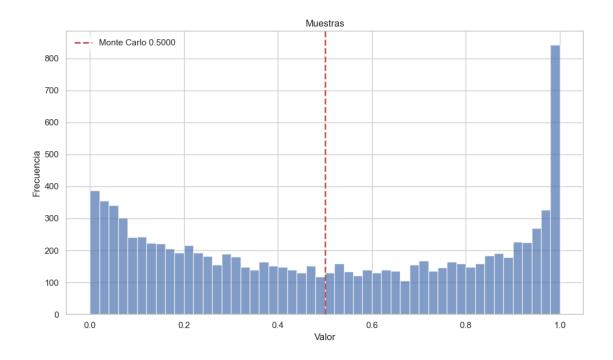
muestras = h(u)

montecarlo = muestras.mean()

montecarlo

[11]: np.float64(0.5000219155164649)

[12]: histograma(muestras, montecarlo)



Calculo analítico

Sea

$$\theta = \int_0^\infty \frac{x}{(1+x^2)^2} \, dx.$$

Integral impropia:

$$\theta = \lim_{b \to \infty} \int_0^b \frac{x}{(1+x^2)^2} \, dx.$$

Sustitución $u = 1 + x^2 \Rightarrow du = 2x dx$: cuando $x = 0 \Rightarrow u = 1$, cuando $x = b \Rightarrow u = 1 + b^2$. Entonces

$$\int_0^b \frac{x}{(1+x^2)^2} \, dx = \frac{1}{2} \int_1^{1+b^2} u^{-2} \, du.$$

Primitiva:

$$\int u^{-2} \, du = -u^{-1} + C.$$

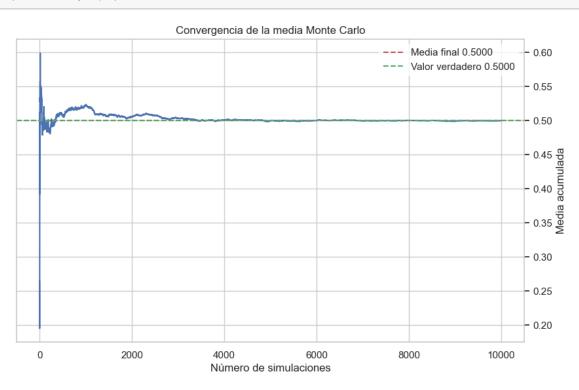
Evaluación:

$$\frac{1}{2} \left[-u^{-1} \right]_1^{1+b^2} = \frac{1}{2} \left(-\frac{1}{1+b^2} + 1 \right).$$

Límite:

$$\theta = \lim_{b \to \infty} \frac{1}{2} \left(1 - \frac{1}{1 + b^2} \right) = \frac{1}{2}.$$

[13]: tlc(muestras, 1/2)



 $\int_0^1 \int_0^1 e^{(x+y)^2} \, dy \, dx$

Sea:

$$\theta = \int_0^1 \int_0^1 e^{(x+y)^2} \, dy \, dx.$$

Entonces:

$$\theta = \int_0^1 \int_0^1 g(x_1, x_2) \, dx_1 \, dx_2, \qquad g(x_1, x_2) = e^{(x_1 + x_2)^2}.$$

Sabemos que:

$$\theta = \mathbb{E}[g(U_1, U_2)], \quad U_1, U_2 \stackrel{iid}{\sim} \text{Unif}(0, 1).$$

Estimador Monte Carlo:

$$\widehat{\theta}_k = \frac{1}{k} \sum_{i=1}^k g(u_{i1}, u_{i2}) = \frac{1}{k} \sum_{i=1}^k \exp((u_{i1} + u_{i2})^2), \quad (u_{i1}, u_{i2}) \stackrel{iid}{\sim} \text{Unif}(0, 1).$$

[14]: def h(u1, u2): return np.exp((u1 + u2)**2)

k = 10000

u1 = np.random.random(k)

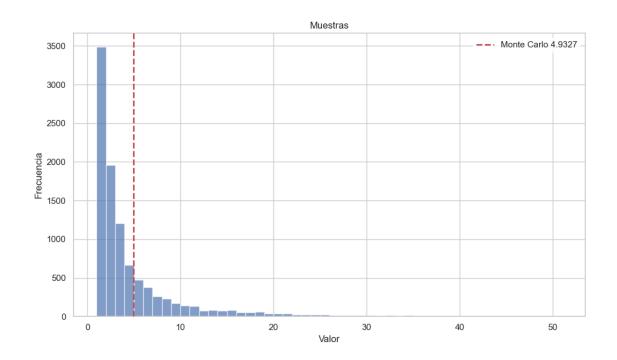
u2 = np.random.random(k)

muestras = h(u1, u2)

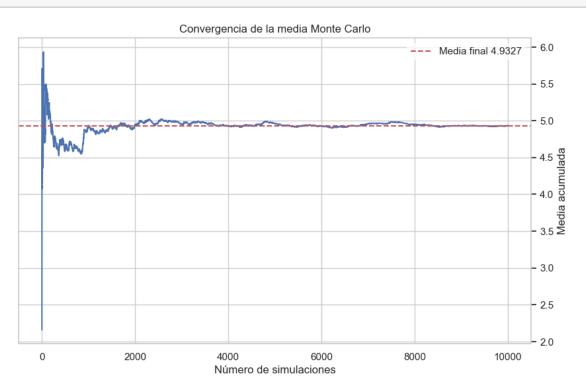
montecarlo = muestras.mean()

montecarlo

- [14]: np.float64(4.932739519320463)
- [15]: histograma(muestras, montecarlo)



[16]: tlc(muestras)



Usar simulación para aproximar $Cov(U, e^U)$, donde $U \sim \mathcal{U}(0, 1)$. Comparar con la respuesta exacta. Por definición,

$$\mathrm{Cov}(X,Y) = \mathbb{E}\big[(X-\mathbb{E}X)(Y-\mathbb{E}Y)\big].$$

Expansión lineal:

$$\mathrm{Cov}(X,Y) = \mathbb{E}[XY] - \mathbb{E}[X] \; \mathbb{E}[Y].$$

Aplicando a X = U y $Y = e^U$:

$$Cov(U, e^U) = \mathbb{E}[Ue^U] - \mathbb{E}[U] \mathbb{E}[e^U].$$

Estimación Monte Carlo

Sea $u_1, \ldots, u_K \stackrel{iid}{\sim} \mathrm{Unif}(0,1)$. Entonces

$$\widehat{\mu}_U = \frac{1}{K} \sum_{i=1}^K u_i, \qquad \widehat{\mu}_e = \frac{1}{K} \sum_{i=1}^K e^{u_i}, \qquad \widehat{m} = \frac{1}{K} \sum_{i=1}^K u_i e^{u_i}.$$

Entonces:

$$\widehat{\mathrm{Cov}}^{(MC)} = \widehat{m} - \widehat{\mu}_U \, \widehat{\mu}_e$$

Cálculo analítico

 $\mathbb{E}[U]$

$$\int_0^1 u \, du = \left[\frac{u^2}{2} \right]_0^1 = \frac{1^2}{2} - \frac{0^2}{2} = \frac{1}{2}.$$

 $\mathbb{E}[e^U]$

$$\int_0^1 e^u \, du = [e^u]_0^1 = e^1 - e^0 = e - 1.$$

 $\mathbb{E}[Ue^U]$ (por partes)

$$\int_0^1 x e^x dx, \qquad \begin{cases} u = x & \Rightarrow du = dx, \\ dv = e^x dx & \Rightarrow v = e^x. \end{cases}$$

$$\int_a^b u \, dv = \left[u \, v \right]_a^b - \int_a^b v \, du.$$

$$\left[x \, e^x \right]_0^1 - \int_0^1 e^x \, dx = \left(1 \cdot e - 0 \cdot 1 \right) - \left[e^x \right]_0^1 = e - (e - 1) = 1.$$

Covarianza

$$Cov(U, e^U) = \mathbb{E}[Ue^U] - \mathbb{E}[U]\mathbb{E}[e^U] = 1 - \frac{1}{2}(e - 1) = \frac{3 - e}{2} = 0.140859086$$

```
[17]: def valor_esperado_1(u):
    return u * np.exp(u)

def valor_esperado_2(u):
    return u

def valor_esperado_3(u):
    return np.exp(u)

k = 1000000

u = np.random.random(k)

montecarlo = valor_esperado_1(u).mean() - valor_esperado_2(u).mean() *_
    -valor_esperado_3(u).mean()
```

- [18]: montecarlo
- [18]: np.float64(0.14075041201920013)
- [19]: 1 1/2*(np.e -1)

[19]: 0.14085908577047745

Para variables aleatorias uniformes U_1, U_2, \ldots definir

contador += 1

return lista_contadores

suma += np.random.random()
lista_contadores.append(contador)

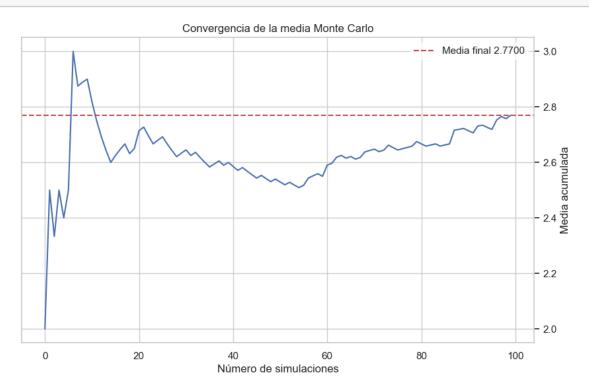
$$N = \min\left\{n: \sum_{i=1}^n U_i > 1\right\}.$$

Estimar $\mathbb{E}[N]$ por simulación con: a) 100 valores, b) 1000 valores, c) 10000 valores, d) Discutir el valor esperado.

```
PSEUDOCÓDIGO - MINIMO_N(k)
      total_contadores ← 0
      PARA i \leftarrow 1 HASTA k HACER:
           suma \leftarrow 0
           contador \leftarrow 0
           MIENTRAS suma < 1 HACER:
                \texttt{contador} \leftarrow \texttt{contador} + 1
                suma \leftarrow UNIFORME(0,1)
           total\_contadores \leftarrow total\_contadores + contador
      RETORNAR total_contadores / k
[20]: def minimo_N(k):
            lista_contadores = []
            for _ in range(k):
                 suma = 0
                 contador = 0
                 while suma < 1:</pre>
```

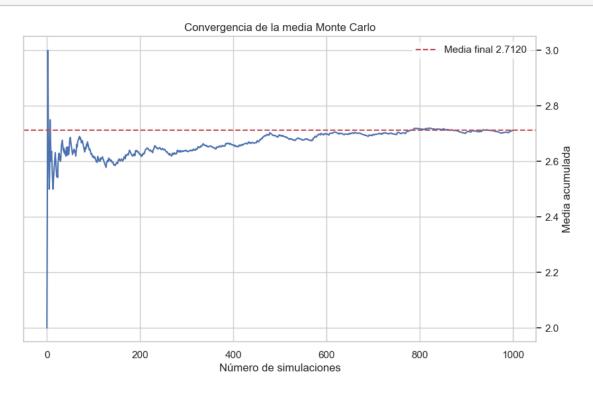
100 valores

[21]: tlc(minimo_N(100))



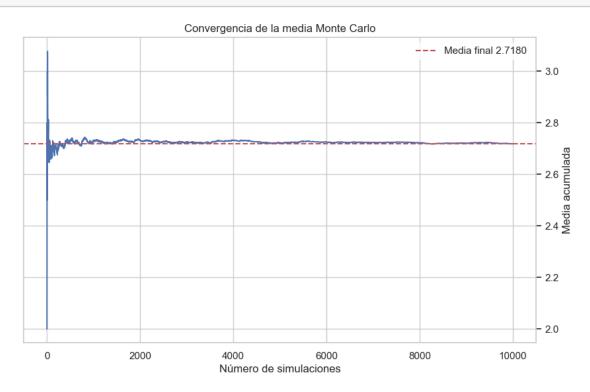
1000 valores

[22]: tlc(minimo_N(1000))



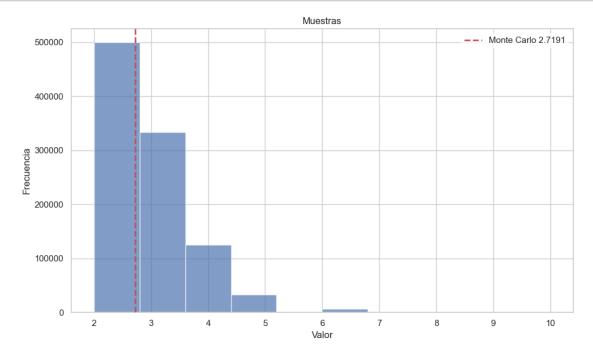
10000 valores

[23]: tlc(minimo_N(10000))



Discución de *n*

[24]: muestras = minimo_N(1000000)
montecarlo = np.mean(muestras)
histograma(muestras, montecarlo, bins=10)



[25]: np.e

[25]: 2.718281828459045

```
Si x_0 = 3 y x_n = (5x_{n-1} + 7) \mod 200. Encontrar x_1, \dots, x_{10}.
                                                 x_n = ax_{n-1} \mod m
      $$ 22 = (5 \cdot 3 + 7) \mod 200
      117 = (22 \cdot 3 + 7) \mod 200
      192 = (117 \cdot 3 + 7) \mod 200
      167 = (192 \cdot 3 + 7) \mod 200
      42 = (167 \cdot 3 + 7) \mod 200
      17 = (42 \cdot 3 + 7) \mod 200
      92 = (17 \cdot 3 + 7) \mod 200
      67 = (92 \cdot 3 + 7) \mod 200
      142 = (67 \cdot 3 + 7) \mod 200
      117 = (142 \cdot 3 + 7) \mod 200
      $$
[26]: pseudoaleatorios = []
       x0 = 3
       a = 5
       m = 200
       c = 7
       for i in range(10):
            xn = (a * x0 + c) % m
            x0 = xn
            pseudoaleatorios.append(xn)
       pseudoaleatorios
[26]: [22, 117, 192, 167, 42, 17, 92, 67, 142, 117]
```

$$\int_0^1 (1-x^2)^{3/2} \, dx.$$

[27]: def h(u):
 return (1-u**2)**(3/2)

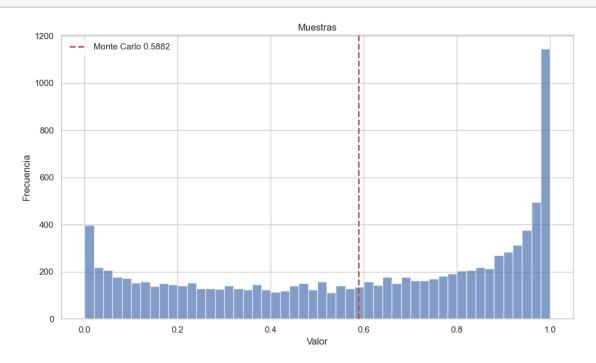
k = 10000

u = np.random.random(k)

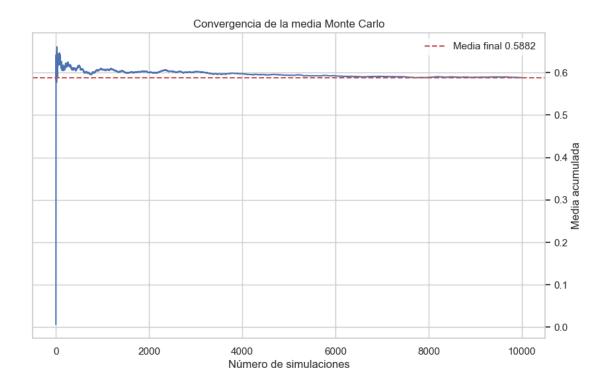
muestras = h(u)
montecarlo = muestras.mean()
montecarlo

[27]: np.float64(0.5882433134915226)

[28]: histograma(muestras, montecarlo)



[29]: tlc(muestras)



$$\int_0^\infty e^{-x}\,dx.$$

Integrando es la densidad $\mathrm{Exp}(1)$, por lo que el valor exacto es 1.

Estimación Monte Carlo

Sea:

$$\theta = \int_0^\infty e^{-x} \, dx.$$

Cambio:

$$y = \frac{1}{x+1}$$
, $dy = -\frac{dx}{(x+1)^2} = -y^2 dx$.

Entonces:

$$\theta = \int_0^1 h(y) \, dy, \qquad h(y) = \frac{g\left(\frac{1}{y} - 1\right)}{y^2}, \quad g(x) = e^{-x}.$$

Forma de esperanza con $U \sim \text{Unif}(0, 1)$:

$$\theta = \mathbb{E}[h(U)].$$

Estimador Monte Carlo:

$$\widehat{\theta}_K = \frac{1}{K} \sum_{i=1}^K h(u_i), \quad u_i \stackrel{\text{iid}}{\sim} \text{Unif}(0,1).$$

[30]: def g(x):

return np.exp(-x)

def h(u):

return $g(1/u-1)/u^{**2}$

 $k = 10_{-000}$

u = np.random.random(k)

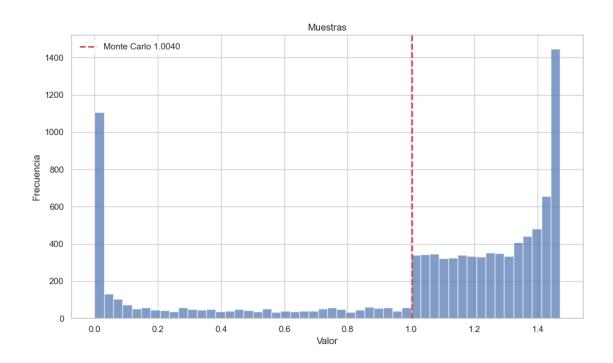
muestras = h(u)

montecarlo = muestras.mean()

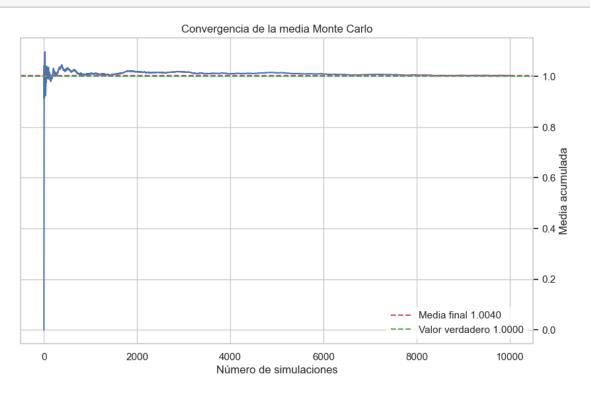
montecarlo

[30]: np.float64(1.003981922206577)

[31]: histograma(muestras, montecarlo)



[32]: tlc(muestras, 1)



$$\int_{-\infty}^{\infty} e^{-x^2} \, dx.$$

Integrando, el valor exacto es $\sqrt{\pi}$ (Lo vimos en análisis II).

Estimación Monte Carlo

Sea:

$$\theta = \int_{-\infty}^{\infty} e^{-x^2} dx.$$

Como tiene símetria par, g(x) = g(-x):

$$\theta = 2 \int_0^\infty e^{-x^2} \, dx.$$

Cambio:

$$y = \frac{1}{x+1}$$
, $dy = -\frac{dx}{(x+1)^2} = -y^2 dx$.

Entonces:

$$\theta = 2 \int_0^1 h(y) \, dy, \qquad h(y) = \frac{g\left(\frac{1}{y} - 1\right)}{y^2}, \quad g(x) = e^{-x^2}.$$

Forma de esperanza con $U \sim \text{Unif}(0, 1)$:

$$\theta = 2\mathbb{E}[\,h(U)\,].$$

Estimador Monte Carlo:

$$\widehat{\theta}_K = 2 \cdot \frac{1}{K} \sum_{i=1}^K h(u_i), \quad u_i \stackrel{\text{iid}}{\sim} \text{Unif}(0, 1).$$

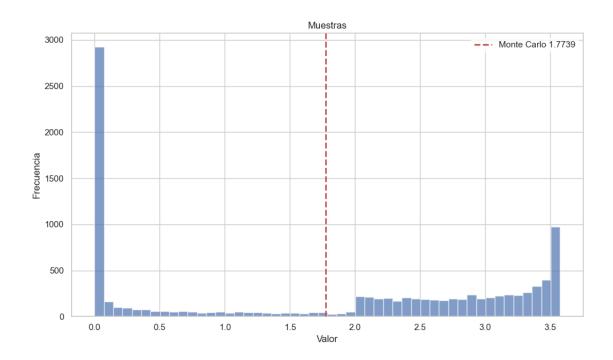
```
[33]: def g(x):
    return np.exp(-x**2)

def h(u):
    return 2* g(1/u-1)/u**2

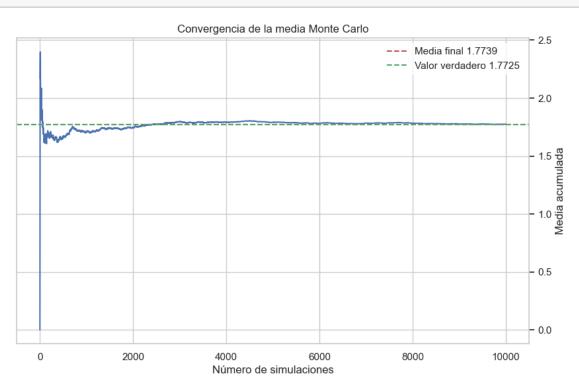
k = 10_000

u = np.random.random(k)
muestras = h(u)
montecarlo = muestras.mean()
montecarlo
```

- [33]: np.float64(1.7739489326465634)
- [34]: verdadero = np.sqrt(np.pi)
 verdadero
- [34]: np.float64(1.7724538509055159)
- [35]: histograma(muestras, montecarlo)



[36]: tlc(muestras, verdadero)



$$\int_0^\infty \int_0^x e^{-(x+y)} \, dy \, dx.$$

Solución analítica

$$\int_0^\infty \! \int_0^x e^{-(x+y)} \, dy \, dx = \int_0^\infty e^{-x} \! \left(\int_0^x e^{-y} \, dy \right) dx.$$

Integrando a $y \cos x$ fijo:

$$\int_0^x e^{-y} \, dy = \left[-e^{-y} \right]_0^x = 1 - e^{-x}.$$

Sustituyendo:

$$\int_0^\infty e^{-x} (1 - e^{-x}) \, dx = \int_0^\infty e^{-x} \, dx - \int_0^\infty e^{-2x} \, dx.$$

Evaluando ambas integrales impropias:

$$\int_0^\infty e^{-x} \, dx = \left[-e^{-x} \right]_0^\infty = 1, \qquad \int_0^\infty e^{-2x} \, dx = \left[-\tfrac{1}{2} e^{-2x} \right]_0^\infty = \tfrac{1}{2}.$$

Restando:

$$1 - \frac{1}{2} = \frac{1}{2}.$$

Estimación Monte Carlo

Sea:

$$\theta = \int_0^\infty \int_0^x e^{-(x+y)} \, dy \, dx.$$

Entonces:

$$\theta = \int_0^\infty \int_0^\infty \mathbf{1}\{x_2 \le x_1\} e^{-(x_1 + x_2)} dx_2 dx_1, \qquad g(x_1, x_2) = \mathbf{1}\{x_2 \le x_1\}.$$

Sabemos que:

$$\theta = \mathbb{E}[g(X_1, X_2)], \quad X_1, X_2 \stackrel{iid}{\sim} \text{Exp}(1).$$

Estimador Monte Carlo:

$$\widehat{\theta}_k = \frac{1}{k} \sum_{i=1}^k g(x_{i1}, x_{i2}) = \frac{1}{k} \sum_{i=1}^k \mathbf{1} \{ x_{i2} \le x_{i1} \}, \quad (x_{i1}, x_{i2}) \stackrel{iid}{\sim} \operatorname{Exp}(1).$$

(Para simular x_{ij} : $x_{ij} = -\ln(1 - u_{ij})$, $u_{ij} \stackrel{iid}{\sim} \text{Unif}(0, 1)$, con el método de la transformada inversa).

```
[37]: k = 1000
    u1 = np.random.random(k)
    u2 = np.random.random(k)

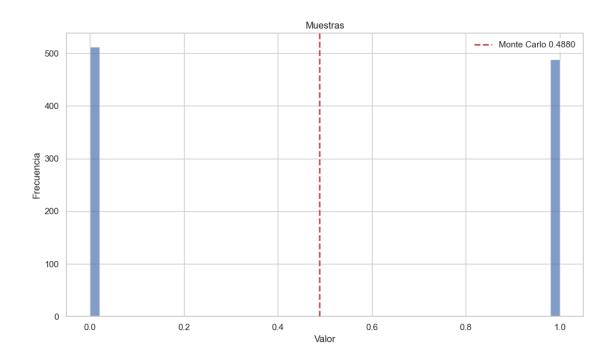
    e1 = -np.log(u1)
    e2 = -np.log(u2)

def g(e1, e2):
    if e1 < e2:
        return 1
    else:
        return 0

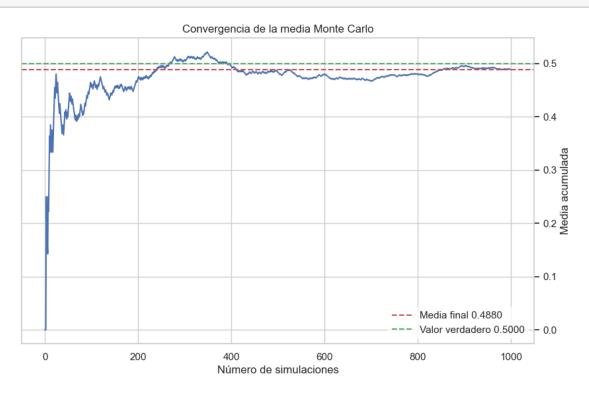
muestras = [g(e1[i], e2[i]) for i in range(k)]

montecarlo = np.mean(muestras)
montecarlo</pre>
```

- [37]: np.float64(0.488)
- [38]: histograma(muestras, montecarlo)



[39]: tlc(muestras, 1/2)



Sea $U \sim \mathcal{U}(0, 1)$. Aproximar por simulación:

- (a) Corr $(U, \sqrt{1-U^2})$,
- (b) Corr $(U^2, \sqrt{1-U^2})$.

Covarianza

$$\mathrm{Cov}(X,Y) = \mathbb{E} \big[(X - \mathbb{E} X) (Y - \mathbb{E} Y) \big].$$

Expansión lineal:

$$\mathrm{Cov}(X,Y) = \mathbb{E}[XY] - \mathbb{E}[X] \; \mathbb{E}[Y].$$

Aplicando a X = U y $Y = e^U$:

$$\mathrm{Cov}(U,\sqrt{1-U^2}) = \mathbb{E}\big[U\sqrt{1-U^2}\big] - \mathbb{E}[U] \; \mathbb{E}[\sqrt{1-U^2}].$$

Correlación

$$\widehat{\rho} = \frac{\widehat{\text{Cov}}}{\sqrt{s_U^2 s_Y^2}}.$$

Varianza

$$\mathrm{Var}(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2.$$

Estimación Monte Carlo

Sea $u_1, \ldots, u_K \stackrel{iid}{\sim} \mathrm{Unif}(0,1)$. Entonces

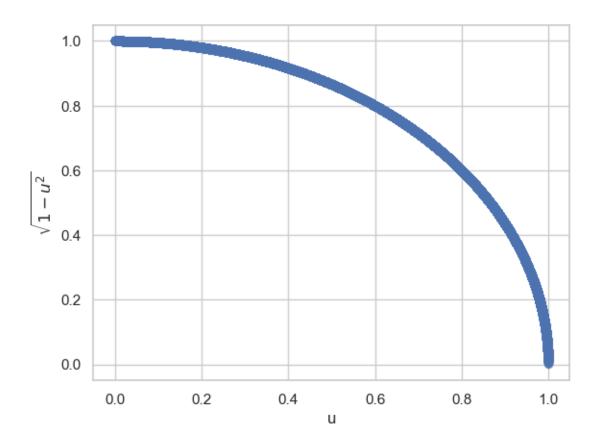
$$\widehat{\mu}_U = \frac{1}{K} \sum_{i=1}^K u_i, \qquad \widehat{\mu}_{\sqrt{1-U^2}} = \frac{1}{K} \sum_{i=1}^K \sqrt{1-U^2}, \qquad \widehat{m} = \frac{1}{K} \sum_{i=1}^K u_i \sqrt{1-U^2}.$$

Entonces:

$$\widehat{\mathrm{Cov}}^{(MC)} = \widehat{m} - \widehat{\mu}_U \, \widehat{\mu}_e$$

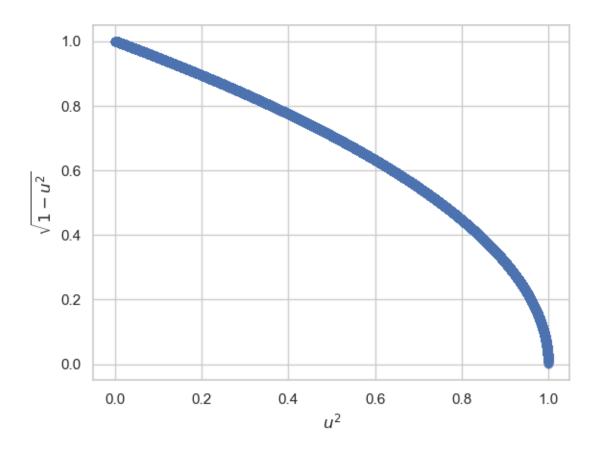
```
A
```

```
[40]: def valor_esperado_1(u):
          return u * np.sqrt(1-u**2)
      def valor_esperado_2(u):
          return u
      def valor_esperado_3(u):
          return np.sqrt(1-u**2)
     k = 1000000
      u = np.random.random(k)
      cov = valor_esperado_1(u).mean() - valor_esperado_2(u).mean() *_
       ⇔valor_esperado_3(u).mean()
      cov
[40]: np.float64(-0.0593370812703804)
[41]: def varianza(u):
          return (u**2).mean() - u.mean()**2
      corr = cov / np.sqrt(varianza(u) * varianza(np.sqrt(1-u**2)))
      corr
[41]: np.float64(-0.9213938620284924)
[42]: plt.scatter(u, np.sqrt(1-u^{**2}), alpha=0.1)
      plt.xlabel("u")
      plt.ylabel("$\sqrt{1-u^2}$")
      plt.show()
```



В

```
[43]: def valor_esperado_1(u):
          return u**2 * np.sqrt(1-u**2)
      def valor_esperado_2(u):
          return u**2
      def valor_esperado_3(u):
          return np.sqrt(1-u**2)
     k = 1000000
      u = np.random.random(k)
      cov = valor_esperado_1(u).mean() - valor_esperado_2(u).mean() *_
       ⇔valor_esperado_3(u).mean()
      cov
[43]: np.float64(-0.06548312826910846)
[44]: def varianza(u):
          return (u**2).mean() - u.mean()**2
      corr = cov / np.sqrt(varianza(u**2) * varianza(np.sqrt(1-u**2)))
      corr
[44]: np.float64(-0.9835397769788468)
[45]: plt.scatter(u**2, np.sqrt(1-u**2), alpha=0.1)
      plt.xlabel("$u^2$")
      plt.ylabel("$\sqrt{1-u^2}$")
      plt.show()
```



Sea $U_i \sim \mathcal{U}(0,1)$ i.i.d. Definir

$$N = \max \left\{ n : \prod_{i=1}^{n} U_i \ge e^{-3} \right\}, \quad \text{con } \prod_{i=0}^{0} U_i = 1.$$

- a) Estimar $\mathbb{E}[N]$ por simulación.
- b) Estimar $\mathbb{P}[N = i]$ para i = 0, 1, 2, 3, 4, 5, 6.

```
A
      ENTRADA: k
      SALIDA: E_hat
      acumulado \leftarrow 0
      PARA r \leftarrow 1 HASTA k HACER:
          S \leftarrow 0
          n \leftarrow 0
          MIENTRAS S 3 HACER:
               u \leftarrow UNIFORME(0,1)
               S \leftarrow S + (\log u)
               SI S 3 ENTONCES:
                    n \leftarrow n + 1
          acumulado \leftarrow acumulado + n
      E_hat \leftarrow acumulado / k
      RETORNAR E_hat
[46]: def estimar_E_N(k, seed=None):
           rng = np.random.default_rng(seed)
           total = 0
           for _ in range(k):
                S = 0.0
                n = 0
                while S <= 3.0:
                     S += -np.log(rng.random())
                    if S <= 3.0:
                         n += 1
                total += n
           return total / k
[47]: estimar_E_N(10000)
```

[47]: 3.0137

```
В
      ENTRADA: k
      SALIDA: p_hat[0..6]
     p_hat[0..6] \leftarrow 0
     PARA r \leftarrow 1 HASTA k HACER:
          S \leftarrow 0
          n \leftarrow 0
          MIENTRAS S 3 HACER:
              u \leftarrow UNIFORME(0,1)
              S \leftarrow S + (\log u)
              SI S 3 ENTONCES:
                   n \leftarrow n + 1
          SI 0 n 6 ENTONCES:
              p_hat[n] \leftarrow p_hat[n] + 1
      PARA i \leftarrow 0 HASTA 6 HACER:
          p_hat[i] \leftarrow p_hat[i] / k
      RETORNAR p_hat
[48]: def estimar_pmf_N(k, seed=None):
           rng = np.random.default_rng(seed)
           counts = np.zeros(7, dtype=int)
           for _ in range(k):
               S = 0.0
               n = 0
               while S <= 3.0:
                    S += -np.log(rng.random())
                    if S <= 3.0:
                        n += 1
               if 0 <= n <= 6:
                    counts[n] += 1
           return counts / k
[49]: for i in range(7):
           print(f"Estimación de P(N={i}): {estimar_pmf_N(10000)[i]:.4f}")
      Estimación de P(N=0): 0.0479
      Estimación de P(N=1): 0.1475
      Estimación de P(N=2): 0.2253
      Estimación de P(N=3): 0.2321
```

Estimación de P(N=4): 0.1740 Estimación de P(N=5): 0.0974 Estimación de P(N=6): 0.0493