#### Lecture 2

June 30, 2025

▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$ 

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ▶ A sequence of random variables  $Y_1, Y_2,...$

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ► A sequence of random variables *Y*<sub>1</sub>, *Y*<sub>2</sub>, . . .
- ▶ The pair  $\{(Y_k, \mathcal{F}_k)\}$  is a martingale if

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ▶ A sequence of random variables  $Y_1, Y_2,...$
- ▶ The pair  $\{(Y_k, \mathcal{F}_k)\}$  is a martingale if
  - ▶  $\{Y_k\}_{k\geq 1}$  is adapted  $\{\mathcal{F}_k\}_{k\geq 1}$ , i.e.,  $Y_k \in \mathcal{F}_k$  for all  $k\geq 1$ .

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ► A sequence of random variables  $Y_1, Y_2, ...$
- ▶ The pair  $\{(Y_k, \mathcal{F}_k)\}$  is a martingale if
  - ▶  $\{Y_k\}_{k\geq 1}$  is adapted  $\{\mathcal{F}_k\}_{k\geq 1}$ , i.e.,  $Y_k \in \mathcal{F}_k$  for all  $k\geq 1$ .
  - $ightharpoonup \mathbb{E}(|Y_k|) < \infty.$

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ▶ A sequence of random variables  $Y_1, Y_2,...$
- ▶ The pair  $\{(Y_k, \mathcal{F}_k)\}$  is a martingale if
  - ▶  $\{Y_k\}_{k\geq 1}$  is adapted  $\{\mathcal{F}_k\}_{k\geq 1}$ , i.e.,  $Y_k \in \mathcal{F}_k$  for all  $k\geq 1$ .
  - $ightharpoonup \mathbb{E}(|Y_k|) < \infty.$
  - $\mathbb{E}(Y_{k+1}|\mathcal{F}_k) = Y_k \text{ for all } k \geq 1.$

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ▶ A sequence of random variables  $Y_1, Y_2, ...$
- The pair {(Y<sub>k</sub>, F<sub>k</sub>)} is a martingale if
  {Y<sub>k</sub>}<sub>k>1</sub> is adapted {F<sub>k</sub>}<sub>k>1</sub>, i.e., Y<sub>k</sub> ∈ F<sub>k</sub> for all k ≥ 1.
  - $\mathbb{E}\left(|Y_k|\right) < \infty.$
  - $\mathbb{E}(Y_{k+1}|\mathcal{F}_k) = Y_k \text{ for all } k > 1.$
- ▶ Often  $\mathcal{F}_k = \sigma(X_1, \dots, X_k)$ , in which case we say  $\{Y_k\}$  is martingale with respect to  $\{X_k\}$ .

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ▶ A sequence of random variables  $Y_1, Y_2, ...$
- The pair {(Y<sub>k</sub>, F<sub>k</sub>)} is a martingale if
  Y<sub>k</sub>}<sub>k>1</sub> is adapted {F<sub>k</sub>}<sub>k>1</sub>, i.e., Y<sub>k</sub> ∈ F<sub>k</sub> for all k ≥ 1.
  - $\mathbb{E}(|Y_k|) < \infty.$
  - $\mathbb{E}(Y_{k+1}|\mathcal{F}_k) = Y_k \text{ for all } k \geq 1.$
- ▶ Often  $\mathcal{F}_k = \sigma(X_1, \dots, X_k)$ , in which case we say  $\{Y_k\}$  is martingale with respect to  $\{X_k\}$ . The key condition in this case is

$$\mathbb{E}\left(Y_{k+1}|X_1,\ldots,X_k\right)=Y_k$$

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ► A sequence of random variables  $Y_1, Y_2, ...$
- ▶ The pair  $\{(Y_k, \mathcal{F}_k)\}$  is a martingale if
  - ▶  $\{Y_k\}_{k\geq 1}$  is adapted  $\{\mathcal{F}_k\}_{k\geq 1}$ , i.e.,  $Y_k \in \mathcal{F}_k$  for all  $k\geq 1$ .
  - $ightharpoonup \mathbb{E}(|Y_k|) < \infty.$
  - ▶  $\mathbb{E}(Y_{k+1}|\mathcal{F}_k) = Y_k$  for all  $k \ge 1$ .
- ▶ Often  $\mathcal{F}_k = \sigma(X_1, \dots, X_k)$ , in which case we say  $\{Y_k\}$  is martingale with respect to  $\{X_k\}$ . The key condition in this case is

$$\mathbb{E}\left(Y_{k+1}|X_1,\ldots,X_k\right)=Y_k$$

One of the most general dependence structures in probability.

- ▶ A filteration  $\{\mathcal{F}_k\}_{k\geq 1}$  is a nested sequence of  $\sigma$ -fields  $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots$
- ► A sequence of random variables  $Y_1, Y_2, ...$
- ▶ The pair  $\{(Y_k, \mathcal{F}_k)\}$  is a martingale if
  - ▶  $\{Y_k\}_{k>1}$  is adapted  $\{\mathcal{F}_k\}_{k>1}$ , i.e.,  $Y_k \in \mathcal{F}_k$  for all  $k \ge 1$ .
  - $ightharpoonup \mathbb{E}(|Y_k|) < \infty.$
  - ▶  $\mathbb{E}(Y_{k+1}|\mathcal{F}_k) = Y_k$  for all  $k \ge 1$ .
- ▶ Often  $\mathcal{F}_k = \sigma(X_1, \dots, X_k)$ , in which case we say  $\{Y_k\}$  is martingale with respect to  $\{X_k\}$ . The key condition in this case is

$$\mathbb{E}\left(Y_{k+1}|X_1,\ldots,X_k\right) = Y_k$$

- One of the most general dependence structures in probability.
- ► Allow relaxing independence assumptions in classical limit theorems.

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ .

$$\mathbb{E}\left(S_{k+1}|X_1,\ldots,X_k\right) = \mathbb{E}\left(\sum_{i=1}^{k+1}X_i|X_1,\ldots,X_k\right)$$

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right) = \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1})$$

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_i$$

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_i = S_k$$

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ . Simply notice that

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_i = S_k$$

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ . Simply notice that

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_i = S_k$$

$$\mathbb{E}\left(L_{k+1} \mid X_1, \dots, X_k\right) = \mathbb{E}\left(\prod_{j=1}^{k+1} X_j \mid X_1, \dots, X_k\right)$$

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ . Simply notice that

$$\mathbb{E}(S_{k+1}|X_{1},...,X_{k}) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_{i} | X_{1},...,X_{k}\right)$$

$$= \mathbb{E}\left(\sum_{i=1}^{k} X_{i} + X_{k+1} | X_{1},...,X_{k}\right)$$

$$= \sum_{i=1}^{k} X_{i} + \mathbb{E}(X_{k+1}|X_{1},...,X_{k})$$

$$= \sum_{i=1}^{k} X_{i} + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_{i} = S_{k}$$

$$\mathbb{E}(L_{k+1} | X_1, \dots, X_k) = \mathbb{E}\left(\prod_{j=1}^{k+1} X_j | X_1, \dots, X_k\right)$$
$$= \mathbb{E}(X_{k+1} | X_1, \dots, X_k) \cdot \prod_{j=1}^k X_j$$

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ . Simply notice that

$$\mathbb{E}(S_{k+1}|X_{1},...,X_{k}) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_{i} | X_{1},...,X_{k}\right)$$

$$= \mathbb{E}\left(\sum_{i=1}^{k} X_{i} + X_{k+1} | X_{1},...,X_{k}\right)$$

$$= \sum_{i=1}^{k} X_{i} + \mathbb{E}(X_{k+1} | X_{1},...,X_{k})$$

$$= \sum_{i=1}^{k} X_{i} + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_{i} = S_{k}$$

$$\mathbb{E}(L_{k+1} | X_1, \dots, X_k) = \mathbb{E}\left(\prod_{j=1}^{k+1} X_j | X_1, \dots, X_k\right)$$

$$= \mathbb{E}(X_{k+1} | X_1, \dots, X_k) \cdot \prod_{j=1}^k X_j$$

$$= \mathbb{E}(X_{k+1}) \cdot \prod_{j=1}^k X_j$$

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ . Simply notice that

$$\mathbb{E}(S_{k+1}|X_{1},...,X_{k}) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_{i} | X_{1},...,X_{k}\right)$$

$$= \mathbb{E}\left(\sum_{i=1}^{k} X_{i} + X_{k+1} | X_{1},...,X_{k}\right)$$

$$= \sum_{i=1}^{k} X_{i} + \mathbb{E}(X_{k+1} | X_{1},...,X_{k})$$

$$= \sum_{i=1}^{k} X_{i} + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_{i} = S_{k}$$

$$\mathbb{E}(L_{k+1} | X_1, \dots, X_k) = \mathbb{E}\left(\prod_{j=1}^{k+1} X_j | X_1, \dots, X_k\right)$$

$$= \mathbb{E}(X_{k+1} | X_1, \dots, X_k) \cdot \prod_{j=1}^k X_j$$

$$= \mathbb{E}(X_{k+1}) \cdot \prod_{j=1}^k X_j$$

$$= 1 \cdot \prod_{j=1}^k X_j.$$

Partial sums  $S_k = \sum_{i=1}^k X_i$  of an iid zero-mean sequence  $\{X_k\}$ . Simply notice that

$$\mathbb{E}(S_{k+1}|X_1,...,X_k) = \mathbb{E}\left(\sum_{i=1}^{k+1} X_i | X_1,...,X_k\right)$$

$$= \mathbb{E}(\sum_{i=1}^{k} X_i + X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1} | X_1,...,X_k)$$

$$= \sum_{i=1}^{k} X_i + \mathbb{E}(X_{k+1})$$

$$= \sum_{i=1}^{k} X_i = S_k$$

$$\mathbb{E}(L_{k+1} | X_1, \dots, X_k) = \mathbb{E}\left(\prod_{j=1}^{k+1} X_j | X_1, \dots, X_k\right)$$

$$= \mathbb{E}(X_{k+1} | X_1, \dots, X_k) \cdot \prod_{j=1}^k X_j$$

$$= \mathbb{E}(X_{k+1}) \cdot \prod_{j=1}^k X_j$$

$$= 1 \cdot \prod_{j=1}^k X_j.$$

$$= \prod_{j=1}^k X_j.$$

$$L_k = \prod_{i=1}^k \frac{f_1(X_i)}{f_0(X_i)}.$$

$$L_k = \prod_{i=1}^k \frac{f_1(X_i)}{f_0(X_i)}.$$

This is a martingale since

$$\mathbb{E}\left(\frac{f_1(X_i)}{f_1(X)}\right) = \int \frac{f_1(x)}{f_1(x)} f_0(x) dx = \int f_1(x) dx = 1$$

$$\mathbb{E}\left(\frac{f_1(X_i)}{f_0(X_i)}\right) = \int \frac{f_1(x)}{f_0(x)} f_0(x) dx = \int f_1(x) dx = 1.$$

$$L_k = \prod_{i=1}^k \frac{f_1(X_i)}{f_0(X_i)}.$$

This is a martingale since

$$\mathbb{E}\left(\frac{f_1(X_i)}{f_0(X_i)}\right) = \int \frac{f_1(x)}{f_0(x)} f_0(x) dx = \int f_1(x) dx = 1.$$

▶ Doob's martingale: For any integrable Z (i.e.  $\mathbb{E}(|Z|) < \infty$  ) the following is always a martingale:

$$Y_k := \mathbb{E}\left(Z|X_1, \dots X_k
ight).$$

$$L_k = \prod_{i=1}^k \frac{f_1(X_i)}{f_0(X_i)}.$$

This is a martingale since

$$\mathbb{E}\left(\frac{f_1(X_i)}{f_0(X_i)}\right) = \int \frac{f_1(x)}{f_0(x)} f_0(x) dx = \int f_1(x) dx = 1.$$

▶ Doob's martingale: For any integrable Z (i.e.  $\mathbb{E}(|Z|) < \infty$ ) the following is always a martingale:

$$Y_k := \mathbb{E}(Z|X_1,\ldots X_k).$$

Notice that

$$\mathbb{E}(Y_{k+1}|X_1,\ldots,X_k) = \mathbb{E}(\mathbb{E}(Z|X_1,\ldots X_{k+1})|X_1,\ldots X_k)$$

$$L_k = \prod_{i=1}^k \frac{f_1(X_i)}{f_0(X_i)}.$$

This is a martingale since

$$\mathbb{E}\left(\frac{f_1(X_i)}{f_0(X_i)}\right) = \int \frac{f_1(x)}{f_0(x)} f_0(x) dx = \int f_1(x) dx = 1.$$

▶ Doob's martingale: For any integrable Z (i.e.  $\mathbb{E}(|Z|) < \infty$ ) the following is always a martingale:

$$Y_k := \mathbb{E}(Z|X_1,\ldots X_k).$$

Notice that

$$\mathbb{E}(\underline{Y_{k+1}}|X_1,\ldots,X_k) = \mathbb{E}(\underline{\mathbb{E}}(\underline{Z}|X_1,\ldots,X_{k+1})|X_1,\ldots,X_k)$$
  
=  $\mathbb{E}(Z|X_1,\ldots,X_k)$ 

$$L_k = \prod_{j=1}^k \frac{f_1(X_j)}{f_0(X_j)}.$$

This is a martingale since

$$\mathbb{E}\left(\frac{f_1(X_i)}{f_0(X_i)}\right) = \int \frac{f_1(x)}{f_0(x)} f_0(x) dx = \int f_1(x) dx = 1.$$

▶ Doob's martingale: For any integrable Z (i.e.  $\mathbb{E}(|Z|) < \infty$ ) the following is always a martingale:

$$Y_k := \mathbb{E}(Z|X_1,\ldots X_k).$$

Notice that

$$\mathbb{E}(Y_{k+1}|X_1,\ldots,X_k) = \mathbb{E}(\mathbb{E}(Z|X_1,\ldots X_{k+1})|X_1,\ldots X_k)$$

$$= \mathbb{E}(Z|X_1,\ldots X_k)$$

$$= Y_k.$$

▶ Theorem 7 (Azuma–Hoeffding). Let  $X = (X_1, ..., X_n)^{\top}$  be a random vector and let Z = f(X). Consider the Doob's martingale

$$Y_i := \mathbb{E}_i(Z) := \mathbb{E}(Z|X_1,\ldots,X_i)$$

and let  $\Delta_i = Y_i - Y_{i-1}$ . Assume that

$$\mathbb{E}_{i-1}\left(e^{\lambda\Delta_i}\right) \leq e^{\sigma_i^2\lambda^2/2}, \quad \forall \lambda \in \mathbb{R}$$
 (1)

almost surely for all i = 1, ..., n. Then,  $Z - \mathbb{E}(Z)$  is sub-Gaussian with parameter  $\sigma = \sqrt{\sum_{i=1}^{n} \sigma_i^2}$ .

▶ Theorem 7 (Azuma–Hoeffding). Let  $X = (X_1, ..., X_n)^{\top}$  be a random vector and let Z = f(X). Consider the Doob's martingale

$$Y_i := \mathbb{E}_i(Z) := \mathbb{E}(Z|X_1,\ldots,X_i)$$

and let  $\Delta_i = Y_i - Y_{i-1}$ . Assume that

$$\mathbb{E}_{i-1}\left(e^{\lambda\Delta_i}\right) \leq e^{\sigma_i^2\lambda^2/2}, \quad \forall \lambda \in \mathbb{R}$$
 (1)

almost surely for all  $i=1,\ldots,n$ . Then,  $Z-\mathbb{E}(Z)$  is sub-Gaussian with parameter  $\sigma=\sqrt{\sum_{i=1}^n\sigma_i^2}$ .

In particular, we have the tail bound

$$\mathbb{P}\left(|Z - \mathbb{E}(Z)| \geq t\right) \leq 2 \exp\left(-rac{t^2}{2\sigma^2}\right).$$

▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .

- ▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .
- Noting that  $\mathbb{E}_0(Z) = \mathbb{E}(Z)$  and  $Y_n = \mathbb{E}_n(Z) = \mathbb{E}(f(X)|X_1,\ldots,X_n) = f(X) = Z,$

$$S_n = \sum_{i=1}^n \Delta_i = \sum_{i=1}^n (Y_i - Y_{i-1}) = Y_n - Y_0 = Z - \mathbb{E}(Z).$$

- ▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .
- Noting that  $\mathbb{E}_0(Z) = \mathbb{E}(Z)$  and  $Y_n = \mathbb{E}_n(Z) = \mathbb{E}(f(X)|X_1, \dots, X_n) = f(X) = Z$ ,

$$S_n = \sum_{i=1}^n \Delta_i = \sum_{i=1}^n (Y_i - Y_{i-1}) = Y_n - Y_0 = Z - \mathbb{E}(Z).$$

By properties of conditional expectation, and assumption (1)

$$\mathbb{E}_{n-1}(e^{\lambda S_n}) = e^{\lambda S_{n-1}} \mathbb{E}_{n-1} \left( e^{\lambda \Delta_n} \right)$$

- ▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .
- Noting that  $\mathbb{E}_0(Z) = \mathbb{E}(Z)$  and  $Y_n = \mathbb{E}_n(Z) = \mathbb{E}(f(X)|X_1,\ldots,X_n) = f(X) = Z$ ,

$$S_n = \sum_{i=1}^n \Delta_i = \sum_{i=1}^n (Y_i - Y_{i-1}) = Y_n - Y_0 = Z - \mathbb{E}(Z).$$

By properties of conditional expectation, and assumption (1)

$$\mathbb{E}_{n-1}(e^{\lambda S_n}) = e^{\lambda S_{n-1}} \mathbb{E}_{n-1}\left(e^{\lambda \Delta_n}\right) \leq e^{\lambda S_{n-1}} e^{\sigma_n^2 \lambda^2/2}.$$

#### **Proof**

- ▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .
- Noting that  $\mathbb{E}_0(Z) = \mathbb{E}(Z)$  and  $Y_n = \mathbb{E}_n(Z) = \mathbb{E}(f(X)|X_1,\ldots,X_n) = f(X) = Z,$

$$S_n = \sum_{i=1}^n \Delta_i = \sum_{i=1}^n (Y_i - Y_{i-1}) = Y_n - Y_0 = Z - \mathbb{E}(Z).$$

By properties of conditional expectation, and assumption (1)

$$\mathbb{E}_{n-1}(e^{\lambda S_n}) = e^{\lambda S_{n-1}} \mathbb{E}_{n-1}\left(e^{\lambda \Delta_n}\right) \leq e^{\lambda S_{n-1}} e^{\sigma_n^2 \lambda^2/2}.$$

▶ Taking  $\mathbb{E}_{n-2}$  of both sides:

$$\mathbb{E}_{n-2}\left(e^{\lambda S_n}\right) \leq$$

#### **Proof**

- ▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .
- Noting that  $\mathbb{E}_0(Z) = \mathbb{E}(Z)$  and  $Y_n = \mathbb{E}_n(Z) = \mathbb{E}(f(X)|X_1,\ldots,X_n) = f(X) = Z$ ,

$$S_n = \sum_{i=1}^n \Delta_i = \sum_{i=1}^n (Y_i - Y_{i-1}) = Y_n - Y_0 = Z - \mathbb{E}(Z).$$

By properties of conditional expectation, and assumption (1)

$$\mathbb{E}_{n-1}(e^{\lambda S_n}) \,=\, e^{\lambda S_{n-1}} \mathbb{E}_{n-1}\left(e^{\lambda \Delta_n}\right) \,\leq\, e^{\lambda S_{n-1}} e^{\sigma_n^2 \lambda^2/2}.$$

▶ Taking  $\mathbb{E}_{n-2}$  of both sides:

$$\mathbb{E}_{n-2}\left(e^{\lambda S_n}\right) \leq e^{\sigma_n^2 \lambda^2/2} \mathbb{E}_{n-2}(e^{\lambda S_{n-1}})$$

•

#### **Proof**

- ▶ Let  $S_j = \sum_{i=1}^j \Delta_i$  which is only a function of  $X_i$  for  $i \leq j$ .
- Noting that  $\mathbb{E}_0(Z) = \mathbb{E}(Z)$  and  $Y_n = \mathbb{E}_n(Z) = \mathbb{E}(f(X)|X_1,\ldots,X_n) = f(X) = Z$ ,

$$S_n = \sum_{i=1}^n \Delta_i = \sum_{i=1}^n (Y_i - Y_{i-1}) = Y_n - Y_0 = Z - \mathbb{E}(Z).$$

By properties of conditional expectation, and assumption (1)

$$\mathbb{E}_{n-1}(e^{\lambda S_n}) = e^{\lambda S_{n-1}} \mathbb{E}_{n-1}\left(e^{\lambda \Delta_n}\right) \leq e^{\lambda S_{n-1}} e^{\sigma_n^2 \lambda^2/2}.$$

▶ Taking  $\mathbb{E}_{n-2}$  of both sides:

$$\mathbb{E}_{n-2}\left(e^{\lambda S_n}\right) \, \leq \, e^{\sigma_n^2 \lambda^2/2} \mathbb{E}_{n-2}(e^{\lambda S_{n-1}}) \, \leq \, e^{\lambda S_{n-2}} e^{(\sigma_n^2 + \sigma_{n-1}^2)\lambda^2/2}.$$

▶ Taking  $\mathbb{E}_{n-2}$  of both sides:

$$\mathbb{E}_{n-2}\left(e^{\lambda S_n}\right) \, \leq \, e^{\sigma_n^2 \lambda^2/2} \mathbb{E}_{n-2}(e^{\lambda S_{n-1}}) \, \leq \, e^{\lambda S_{n-2}} e^{(\sigma_n^2 + \sigma_{n-1}^2)\lambda^2/2}$$

▶ Taking  $\mathbb{E}_{n-2}$  of both sides:

$$\mathbb{E}_{n-2}\left(e^{\lambda S_n}\right) \leq e^{\sigma_n^2 \lambda^2/2} \mathbb{E}_{n-2}(e^{\lambda S_{n-1}}) \leq e^{\lambda S_{n-2}} e^{(\sigma_n^2 + \sigma_{n-1}^2)\lambda^2/2}$$

Repeating the process, we get

$$\mathbb{E}_0\left(oldsymbol{e}^{\lambda\mathcal{S}_n}
ight) \, \leq \, \exp\left(rac{\lambda}{2}\sum_{i=1}^n\sigma_i^2
ight) \, = \, \exp\left(rac{\lambda^2\sigma^2}{2}
ight).$$

▶ Taking  $\mathbb{E}_{n-2}$  of both sides:

$$\mathbb{E}_{n-2}\left(e^{\lambda S_n}\right) \leq e^{\sigma_n^2 \lambda^2/2} \mathbb{E}_{n-2}(e^{\lambda S_{n-1}}) \leq e^{\lambda S_{n-2}} e^{(\sigma_n^2 + \sigma_{n-1}^2)\lambda^2/2}$$

► Repeating the process, we get

$$\mathbb{E}_0\left(e^{\lambda S_n}\right) \, \leq \, \exp\left(\frac{\lambda}{2} \sum_{i=1}^n \sigma_i^2\right) \, = \, \exp\left(\frac{\lambda^2 \sigma^2}{2}\right).$$

▶ Or

$$\mathbb{E}\left(e^{\lambda(Z-\mathbb{E}(Z))}\right) \, \leq \, \exp\left(\frac{\lambda}{2} \sum_{i=1}^n \sigma_i^2\right) \, = \, \exp\left(\frac{\lambda^2 \sigma^2}{2}\right).$$

and the claim follows.

# Bounded difference inequality

Conditional sub-G. assump. holds under bounded difference property:

```
|f(x_1,\ldots,x_{i-1},x_i,x_{i+1},\ldots,x_n)-f(x_1,\ldots,x_{i-1},x_i',x_{i+1},\ldots,x_n)| \le L_i (2)
```

for all  $x_1, \ldots, x_n, x_i' \in \mathcal{X}$  and some constants  $(L_1, \ldots, L_n)$ .

# Bounded difference inequality

Conditional sub-G. assump. holds under bounded difference property:

$$|f(x_1,\ldots,x_{i-1}, \mathbf{x}_i, x_{i+1},\ldots,x_n) - f(x_1,\ldots,x_{i-1}, \mathbf{x}_i', x_{i+1},\ldots,x_n)|$$

$$\leq L_i$$
(2)
for all  $x_1,\ldots,x_n, x_i' \in \mathcal{X}$  and some constants  $(L_1,\ldots,L_n)$ .

▶ Theorem 8 (Bounded difference). Assume that  $X = (X_1, ..., X_n)^{\top}$  has independent coordinates, and assume that  $f : \mathcal{X}^n \to \mathbb{R}$  satisfies the bounded difference property (2). Then

$$\mathbb{P}\left(|f(X) - \mathbb{E}(f(X))| \geq t\right) \leq 2 \exp\left(-\frac{2t^2}{\sum_{i=1}^n L_i^2}\right), \ t \geq 0.$$

# Bounded difference inequality

Conditional sub-G. assump. holds under bounded difference property:

$$|f(x_1,\ldots,x_{i-1}, \mathbf{x}_i, x_{i+1},\ldots,x_n) - f(x_1,\ldots,x_{i-1}, \mathbf{x}_i', x_{i+1},\ldots,x_n)|$$

$$\leq L_i$$
(2)
for all  $x_1,\ldots,x_n, x_i' \in \mathcal{X}$  and some constants  $(L_1,\ldots,L_n)$ .

▶ Theorem 8 (Bounded difference). Assume that  $X = (X_1, \dots, X_n)^{\top}$  has independent coordinates, and assume that  $f : \mathcal{X}^n \to \mathbb{R}$  satisfies the bounded difference property (2). Then

$$\mathbb{P}\left(|f(X) - \mathbb{E}(f(X))| \geq t\right) \leq 2 \exp\left(-\frac{2t^2}{\sum_{i=1}^n L_i^2}\right), \ t \geq 0.$$

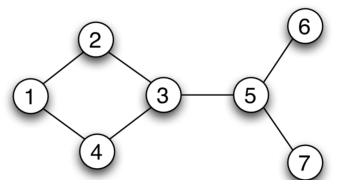
**Example.**  $f(X) = \sum_{i=1}^{n} X_i, X_i \in [a_i, b_i].$ 

## Clique number of Erdős-Rényi

Let *G* be an undirected graph on *n* nodes.

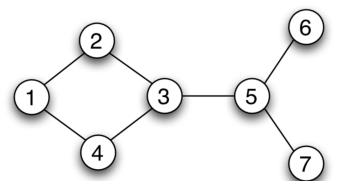
## Clique number of Erdős-Rényi

Let *G* be an undirected graph on *n* nodes.



## Clique number of Erdős-Rényi

Let *G* be an undirected graph on *n* nodes.

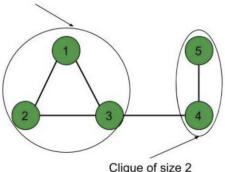


► The edge set is  $E(G) := \{\{1,2\},\{1,4\},\{2,3\},\{3,4\},\{3,5\},\{5,6\},\{5,7\}\}.$ 

► A clique in *G* is a complete (induced) sub-graph.

▶ A clique in *G* is a complete (induced) sub-graph.

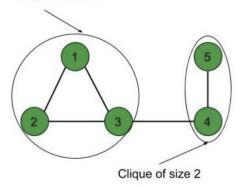
Clique of size 3



Clique of size 2

► A clique in *G* is a complete (induced) sub-graph.

Clique of size 3



▶ Clique number of G—denoted as  $\omega(G)$ —is the size of the largest clique(s).

► For two graphs *G* and *G'* that differ in at most 1 edge,

$$|\omega(G) - \omega(G')| \leq 1.$$

► For two graphs *G* and *G'* that differ in at most 1 edge,

$$|\omega(G) - \omega(G')| \leq 1.$$

▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.

► For two graphs G and G' that differ in at most 1 edge,

$$|\omega(G) - \omega(G')| \leq 1.$$

- ▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.
- ▶ Let G be an Erdős-Rényi random graph: Edges are independently drawn with probability p.

$$|\omega(G) - \omega(G')| \leq 1.$$

- ▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.
- Let G be an Erdős-Rényi random graph: Edges are independently drawn with probability p. Then, with  $m = \binom{n}{2}$

$$\mathbb{P}(|\omega(G) - \mathbb{E}(\omega(G))| \ge \frac{\delta}{\delta}) \le 2 \exp(-2\delta^2 / \sum_{i=1}^{m} 1)$$

$$|\omega(G) - \omega(G')| \leq 1.$$

- ▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.
- Let G be an Erdős-Rényi random graph: Edges are independently drawn with probability p. Then, with  $m = \binom{n}{2}$

$$\mathbb{P}(|\omega(G) - \mathbb{E}(\omega(G))| \ge \frac{\delta}{\delta}) \le 2 \exp(-2\delta^2 / \sum_{i=1}^{m} 1)$$

$$= 2 \exp(-2\delta^2 / m).$$

$$|\omega(G) - \omega(G')| \leq 1.$$

- ▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.
- Let G be an Erdős-Rényi random graph: Edges are independently drawn with probability p. Then, with  $m = \binom{n}{2}$

$$\begin{array}{lcl} \mathbb{P}\left(|\omega(G) - \mathbb{E}(\omega(G))| \geq \frac{\delta}{\delta}\right) & \leq & 2\exp\left(-2\delta^2/\sum_{i=1}^m 1\right) \\ & = & 2\exp\left(-2\frac{\delta^2}{m}\right). \end{array}$$

or setting  $\bar{\omega}(G) = \omega(G)/m$ ,

$$\mathbb{P}\left(|\bar{\omega}(G) - \mathbb{E}(\bar{\omega}(G))| \geq \delta\right) = \mathbb{P}\left(|\omega(G) - \mathbb{E}(\omega(G))| \geq m\delta\right)$$

$$|\omega(G) - \omega(G')| \leq 1.$$

- ▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.
- Let G be an Erdős-Rényi random graph: Edges are independently drawn with probability p. Then, with  $m = \binom{n}{2}$

$$\begin{array}{lcl} \mathbb{P}\left(|\omega(G) - \mathbb{E}(\omega(G))| \geq \frac{\delta}{\delta}\right) & \leq & 2\exp\left(-2\delta^2/\sum_{i=1}^m 1\right) \\ & = & 2\exp\left(-2\frac{\delta^2}{m}\right). \end{array}$$

or setting  $\bar{\omega}(G) = \omega(G)/m$ ,

$$\mathbb{P}(|\bar{\omega}(G) - \mathbb{E}(\bar{\omega}(G))| \ge \delta) = \mathbb{P}(|\omega(G) - \mathbb{E}(\omega(G))| \ge m\delta) \le 2 \exp(-2(m\delta)^2/m)$$

$$|\omega(G) - \omega(G')| \leq 1.$$

- ▶ Thus  $E(G) \rightarrow \omega(G)$  has bounded difference property with L = 1.
- Let G be an Erdős-Rényi random graph: Edges are independently drawn with probability p. Then, with  $m = \binom{n}{2}$

$$\begin{array}{lcl} \mathbb{P}\left(|\omega(G) - \mathbb{E}(\omega(G))| \geq \frac{\delta}{\delta}\right) & \leq & 2\exp\left(-2\delta^2/\sum_{i=1}^m 1\right) \\ & = & 2\exp\left(-2\frac{\delta^2}{m}\right). \end{array}$$

or setting  $\bar{\omega}(G) = \omega(G)/m$ ,

$$\mathbb{P}(|\bar{\omega}(G) - \mathbb{E}(\bar{\omega}(G))| \ge \delta) = \mathbb{P}(|\omega(G) - \mathbb{E}(\omega(G))| \ge m\delta) \\
\le 2\exp(-2(m\delta)^2/m) \\
\le 2\exp(-2m\delta^2)$$

▶ A function  $f : \mathbb{R}^n \to \mathbb{R}$  is *L*-Lipschitz w.r.t.  $\| \cdot \|_2$  if

$$|f(x)-f(y)| \leq L||x-y||_2, \quad \forall x, y \in \mathbb{R}^n.$$

▶ A function  $f : \mathbb{R}^n \to \mathbb{R}$  is *L*-Lipschitz w.r.t.  $\| \cdot \|_2$  if

$$|f(x)-f(y)| \leq L||x-y||_2, \quad \forall x,y \in \mathbb{R}^n.$$

For instance, if *f* is differentiable with bounded derivative.

▶ A function  $f : \mathbb{R}^n \to \mathbb{R}$  is *L*-Lipschitz w.r.t.  $\| \cdot \|_2$  if

$$|f(x)-f(y)| \leq L||x-y||_2, \quad \forall x,y \in \mathbb{R}^n.$$

- For instance, if *f* is differentiable with bounded derivative.
- ▶ The function  $f: \mathbb{R} \to \mathbb{R}$ ,  $f(x) = e^{-x^2}$  is Lipschitz.

▶ A function  $f : \mathbb{R}^n \to \mathbb{R}$  is *L*-Lipschitz w.r.t.  $\|\cdot\|_2$  if

$$|f(x)-f(y)| \leq L||x-y||_2, \quad \forall x,y \in \mathbb{R}^n.$$

- For instance, if *f* is differentiable with bounded derivative.
- ▶ The function  $f: \mathbb{R} \to \mathbb{R}$ ,  $f(x) = e^{-x^2}$  is Lipschitz.
- ▶ The function  $f:[0,1] \to \mathbb{R}$ ,  $f(x) = \sqrt{x}$  is not Lipschitz.

▶ Theorem 9 (Gaussian concentration). Let  $X \sim N(0, I_n)$  be a standard Gaussian vector and assume that  $f: \mathbb{R}^n \to \mathbb{R}$  is L-Lipschitz w.r.t. the Euclidean norm. Then,

$$\mathbb{P}(f(X) - \mathbb{E}(f(X)) \ge t) \le \exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

▶ Theorem 9 (Gaussian concentration). Let  $X \sim N(0, I_n)$  be a standard Gaussian vector and assume that  $f : \mathbb{R}^n \to \mathbb{R}$  is L-Lipschitz w.r.t. the Euclidean norm. Then,

$$\mathbb{P}\left(f(X) - \mathbb{E}(f(X)) \ge t\right) \le \exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

Notice that it follows that

$$\mathbb{P}\left(|f(X) - \mathbb{E}(f(X))| \geq t\right) \leq 2\exp\left(-\frac{t^2}{2L^2}\right), \qquad t \geq 0.$$

**Theorem 9 (Gaussian concentration).** Let  $X \sim N(0, I_n)$  be a standard Gaussian vector and assume that  $f: \mathbb{R}^n \to \mathbb{R}$  is L-Lipschitz w.r.t. the Euclidean norm. Then,

$$\mathbb{P}\left(f(X) - \mathbb{E}(f(X)) \ge t\right) \le \exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

Notice that it follows that

$$\mathbb{P}\left(|f(X) - \mathbb{E}(f(X))| \ge t\right) \le 2\exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

▶ In other words, f(X) is sub-Gaussian with parameter L.

**Theorem 9 (Gaussian concentration).** Let  $X \sim N(0, I_n)$  be a standard Gaussian vector and assume that  $f : \mathbb{R}^n \to \mathbb{R}$  is L-Lipschitz w.r.t. the Euclidean norm. Then,

$$\mathbb{P}\left(f(X) - \mathbb{E}(f(X)) \ge t\right) \le \exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

Notice that it follows that

$$\mathbb{P}\left(|f(X) - \mathbb{E}(f(X))| \ge t\right) \le 2\exp\left(-rac{t^2}{2L^2}
ight), \qquad t \ge 0.$$

- ▶ In other words, f(X) is sub-Gaussian with parameter L.
- Deep result, no easy proof!

▶ Theorem 9 (Gaussian concentration). Let  $X \sim N(0, I_n)$  be a standard Gaussian vector and assume that  $f : \mathbb{R}^n \to \mathbb{R}$  is L-Lipschitz w.r.t. the Euclidean norm. Then,

$$\mathbb{P}(f(X) - \mathbb{E}(f(X)) \ge t) \le \exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

Notice that it follows that

$$\mathbb{P}\left(|f(X) - \mathbb{E}(f(X))| \ge t\right) \le 2\exp\left(-\frac{t^2}{2L^2}\right), \qquad t \ge 0$$

- ▶ In other words, f(X) is sub-Gaussian with parameter L.
- Deep result, no easy proof!
- Has far-reaching consequences.

▶ Consider a matrix  $X \in \mathbb{R}^{n \times d}$  where n > d.

- ▶ Consider a matrix  $X \in \mathbb{R}^{n \times d}$  where n > d.
- Let  $\sigma_1(X) \ge \sigma_2(X) \ge ... \ge \sigma_k(X)$  be (ordered) singular values of X.

- ▶ Consider a matrix  $X \in \mathbb{R}^{n \times d}$  where n > d.
- Let  $\sigma_1(X) \ge \sigma_2(X) \ge ... \ge \sigma_k(X)$  be (ordered) singular values of X.
- ▶ By Weyl's theorem, for any  $X, Y \in \mathbb{R}^{n \times d}$ :

$$|\sigma_k(X) - \sigma_k(Y)| \le ||X - Y||_{op} \le ||X - Y||_F.$$

Note that this is a generalization of order-statistics inequality.)

- ▶ Consider a matrix  $X \in \mathbb{R}^{n \times d}$  where n > d.
- Let  $\sigma_1(X) \ge \sigma_2(X) \ge ... \ge \sigma_k(X)$  be (ordered) singular values of X.
- ▶ By Weyl's theorem, for any  $X, Y \in \mathbb{R}^{n \times d}$ :

$$|\sigma_k(X) - \sigma_k(Y)| \le ||X - Y||_{op} \le ||X - Y||_F.$$

Note that this is a generalization of order-statistics inequality.)

▶ Thus,  $X \to \sigma_1(X)$  is 1-Lipschitz.

# Example: Singular values

- ▶ Consider a matrix  $X \in \mathbb{R}^{n \times d}$  where n > d.
- Let  $\sigma_1(X) \ge \sigma_2(X) \ge ... \ge \sigma_k(X)$  be (ordered) singular values of X.
- ▶ By Weyl's theorem, for any  $X, Y \in \mathbb{R}^{n \times d}$ :

$$|\sigma_k(X) - \sigma_k(Y)| \leq ||X - Y||_{op} \leq ||X - Y||_F.$$

Note that this is a generalization of order-statistics inequality.)

- ▶ Thus,  $X \rightarrow \sigma_1(X)$  is 1-Lipschitz.
- ▶ **Proposition 6.** Let  $X \in \mathbb{R}^{n \times d}$  be a random matrix with iid N(0,1) entries. Then,

$$\mathbb{P}\left(|\sigma_k(X) - \mathbb{E}\left(\sigma_k(X)\right)| \geq \delta\right) \leq 2e^{-\delta^2/2}, \quad \delta \geq 0.$$

# Example: Singular values

- ▶ Consider a matrix  $X \in \mathbb{R}^{n \times d}$  where n > d.
- Let  $\sigma_1(X) \ge \sigma_2(X) \ge ... \ge \sigma_k(X)$  be (ordered) singular values of X.
- ▶ By Weyl's theorem, for any  $X, Y \in \mathbb{R}^{n \times d}$ :

$$|\sigma_k(X) - \sigma_k(Y)| \le ||X - Y||_{op} \le ||X - Y||_F.$$

Note that this is a generalization of order-statistics inequality.)

- ▶ Thus,  $X \rightarrow \sigma_1(X)$  is 1-Lipschitz.
- ▶ **Proposition 6.** Let  $X \in \mathbb{R}^{n \times d}$  be a random matrix with iid N(0,1) entries. Then,

$$\mathbb{P}(|\sigma_k(X) - \mathbb{E}(\sigma_k(X))| \geq \delta) \leq 2e^{-\delta^2/2}, \quad \delta \geq 0.$$

▶ It remains to characterize  $\mathbb{E}(\sigma_k(X))$ .

▶ When d > n there is no hope of estimating  $\theta^*$ ,

- ▶ When d > n there is no hope of estimating  $\theta^*$ ,
- unless we impose a low dimensional assumption on  $\theta^*$ .

- ▶ When d > n there is no hope of estimating  $\theta^*$ ,
- unless we impose a low dimensional assumption on  $\theta^*$ .
- ▶ Support of  $\theta^*$  (recall  $[d] = \{1, ..., d\}$ )

$$\operatorname{supp}(\theta^*) := \mathcal{S}(\theta^*) = \{ j \in [d] : \theta_j^* \neq 0 \}.$$

- ▶ When d > n there is no hope of estimating  $\theta^*$ ,
- unless we impose a low dimensional assumption on  $\theta^*$ .
- Support of  $\theta^*$  (recall  $[d] = \{1, \dots, d\}$ )

$$\operatorname{supp}(\theta^*) := \mathcal{S}(\theta^*) = \{j \in [d] \, : \, \theta_j^* \neq 0\}.$$

▶ Hard sparsity assumption:  $s = |S(\theta^*)| << d$ .

- ▶ When d > n there is no hope of estimating  $\theta^*$ ,
- unless we impose a low dimensional assumption on  $\theta^*$ .
- Support of  $\theta^*$  (recall  $[d] = \{1, \ldots, d\}$ )

$$\operatorname{supp}(\theta^*) := \mathcal{S}(\theta^*) = \{j \in [d] \, : \, \theta_j^* \neq 0\}.$$

- ▶ Hard sparsity assumption:  $s = |S(\theta^*)| << d$ .
- ▶ Weak sparsity via  $\ell_q$  balls for  $q \in [0, 1]$ :

$$heta^* \in \mathbb{B}_q(R_q) := \left\{ heta \in \mathbb{R}^d \, : \, \sum_{j=1}^d | heta_j|^q \leq R_q 
ight\}.$$

- ▶ When d > n there is no hope of estimating  $\theta^*$ ,
- unless we impose a low dimensional assumption on  $\theta^*$ .
- Support of  $\theta^*$  (recall  $[d] = \{1, \dots, d\}$ )

$$\operatorname{supp}(\theta^*) := \mathcal{S}(\theta^*) = \{j \in [d] : \theta_j^* \neq 0\}.$$

- ▶ Hard sparsity assumption:  $s = |S(\theta^*)| << d$ .
- ▶ Weak sparsity via  $\ell_q$  balls for  $q \in [0, 1]$ :

$$heta^* \in \mathbb{B}_q(R_q) := \left\{ heta \in \mathbb{R}^d \, : \, \sum_{j=1}^d | heta_j|^q \leq R_q 
ight\}.$$

▶ q = 1 gives the  $\ell_1$  ball.

- ▶ When d > n there is no hope of estimating  $\theta^*$ ,
- unless we impose a low dimensional assumption on  $\theta^*$ .
- Support of  $\theta^*$  (recall  $[d] = \{1, \ldots, d\}$ )

$$\operatorname{supp}(\theta^*) := S(\theta^*) = \{j \in [d] : \theta_j^* \neq 0\}.$$

- ▶ Hard sparsity assumption:  $s = |S(\theta^*)| << d$ .
- ▶ Weak sparsity via  $\ell_q$  balls for  $q \in [0, 1]$ :

$$heta^* \in \mathbb{B}_q(R_q) := \left\{ heta \in \mathbb{R}^d \, : \, \sum_{j=1}^d | heta_j|^q \leq R_q 
ight\}.$$

- ightharpoonup q = 1 gives the  $\ell_1$  ball.
- ▶ q = 0 gives the  $\ell_0$  ball, same as hard sparsity:

$$\|\theta^*\|_0 := |S(\theta^*)|.$$

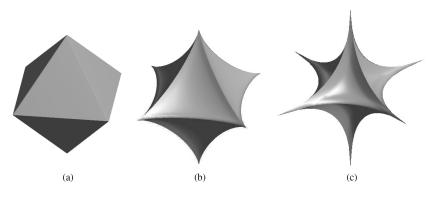


Figure 7.1 Illustrations of the  $\ell_q$ -"balls" for different choices of the parameter  $q \in (0, 1]$ . (a) For q = 1, the set  $\mathbb{B}_1(R_q)$  corresponds to the usual  $\ell_1$ -ball shown here. (b) For q = 0.75, the ball is a non-convex set obtained by collapsing the faces of the  $\ell_1$ -ball towards the origin. (c) For q = 0.5, the set becomes more "spiky", and it collapses into the hard sparsity constraint as  $q \to 0^+$ . As shown in Exercise 7.2(a), for all  $q \in (0, 1]$ , the set  $\mathbb{B}_q(1)$  is star-shaped around the origin.

# Compressed sensing or the equation $y = X\theta^*$

▶ When can we solve the equation  $y = X\theta^*$ ?

#### The classical answer

The classical theory of linear algebra, which we learn as undergraduates, is as follows:

If there are at least as many equations as unknowns  $(n \ge d)$ , and X has full rank, then the problem is determined or overdetermined, and one can easily solve  $y = X\theta$  uniquely (e.g. by gaussian elimination).

#### The classical answer

The classical theory of linear algebra, which we learn as undergraduates, is as follows:

- If there are at least as many equations as unknowns  $(n \ge d)$ , and X has full rank, then the problem is determined or overdetermined, and one can easily solve  $y = X\theta$  uniquely (e.g. by gaussian elimination).
- If there are fewer equations than unknowns (n < d), then the problem is underdetermined even when X has full rank. Knowledge of  $y = X\theta$  restricts  $\theta$  to an (affine) subspace of  $\mathbb{R}^d$ , but does not determine  $\theta$  completely.

# Sparse recovery

It is thus of interest to obtain a good estimator for underdetermined problems such as  $X\theta^* = y$  in the case in which  $\theta^*$  is expected to be "spiky" - that is, concentrated in only a few of its coordinates.

# Sparse recovery

- lt is thus of interest to obtain a good estimator for underdetermined problems such as  $X\theta^* = y$  in the case in which  $\theta^*$  is expected to be "spiky" that is, concentrated in only a few of its coordinates.
- ▶ A model case occurs when  $\theta^*$  is known to be *s*-sparse for some  $1 \le s \le d$ , which means that at most *s* of the coefficients of  $\theta^*$  can be non-zero.

## Sparse recovery

Sparsity is a simple but effective model for many real-life signals. For instance, an image may be many megapixels in size, but when viewed in the right basis (e.g. a wavelet basis), many of the coefficients may be negligible, and so the image may be compressible into a file of much smaller size without seriously affecting the image quality. In other words, many images are effectively sparse in the wavelet basis.

# Sparsity helps!

Intuitively, if a signal  $\theta^* \in \mathbb{R}^d$  is s-sparse, then it should only have s degrees of freedom rather than d. In principle, one should now only need s measurements or so to reconstruct  $\theta^*$ , rather than d.

signals are sparse in a known basis;

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

#### Such situations can arise in:

Imaging.

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

- Imaging.
- Sensor networks.

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

- Imaging.
- Sensor networks.
- MRI.

- signals are sparse in a known basis;
- measurements (or computation at the sensor end) are expensive; but computations at the receiver end are cheap.

- Imaging.
- Sensor networks.
- ► MRI.
- Astronomy.

▶ Consider the noiseless case  $y = X\theta^*$ .

- ▶ Consider the noiseless case  $y = X\theta^*$ .
- ▶ We assume that  $\|\theta^*\|_0$  is small.

- ▶ Consider the noiseless case  $y = X\theta^*$ .
- ▶ We assume that  $\|\theta^*\|_0$  is small.
- ▶ Ideal program to solve:

$$\min_{\theta} \|\theta\|_0$$
 subject to  $y = X\theta$ 

- ▶ Consider the noiseless case  $y = X\theta^*$ .
- ▶ We assume that  $\|\theta^*\|_0$  is small.
- Ideal program to solve:

$$\min_{\theta} \|\theta\|_0$$
 subject to  $y = X\theta$ 

 $ightharpoonup \|\cdot\|_0$  is highly nonconvex, relax to  $\|\cdot\|_1$ :

$$\min_{\theta} \|\theta\|_1$$
 subject to  $y = X\theta$ 

- ▶ Consider the noiseless case  $y = X\theta^*$ .
- ▶ We assume that  $\|\theta^*\|_0$  is small.
- Ideal program to solve:

$$\min_{\theta} \|\theta\|_0$$
 subject to  $y = X\theta$ 

 $ightharpoonup \|\cdot\|_0$  is highly nonconvex, relax to  $\|\cdot\|_1$ :

$$\min_{\theta} \|\theta\|_1$$
 subject to  $y = X\theta$ 

► This is called basis pursuit (regression).

- ► Consider the noiseless case  $y = X\theta^*$ .
- ▶ We assume that  $\|\theta^*\|_0$  is small.
- Ideal program to solve:

$$\min_{\theta} \|\theta\|_0$$
 subject to  $y = X\theta$ 

 $ightharpoonup \|\cdot\|_0$  is highly nonconvex, relax to  $\|\cdot\|_1$ :

$$\min_{\theta} \|\theta\|_1$$
 subject to  $y = X\theta$ 

- This is called basis pursuit (regression).
- The resulting problem is convex.

► In fact, can be written as a linear program. Notice that the problem is equivalent to

$$\min_{s} \sum_{j=1}^{d} s_{j}$$
 subject to  $y = X\theta$ ,  $|\theta_{j}| \leq s_{j}$ .

▶ In fact, can be written as a linear program. Notice that the problem is equivalent to

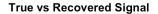
$$\min_{s} \sum_{j=1}^{d} s_{j}$$
 subject to  $y = X\theta$ ,  $|\theta_{j}| \leq s_{j}$ .

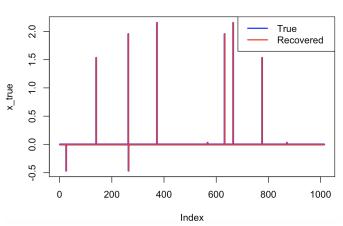
Global solutions can be obtained very efficiently.

## Basis pursuit example

```
library(CVXR)
3
   set.seed(123)
   n <- 1014
                    # Signal dimension
   m <- 64 # Number of measurements
   A <- matrix(rnorm(m * n), m, n) # Measurement matrix
   x_true <- rep(0, n)
   x_true[sample(1:n, 10)] <- rnorm(5) # Sparse true signal
   plot(x_true)
   b <- A %*% x true
  x <- Variable(n)
.8 objective <- Minimize(norm1(x))</pre>
 constraints <- list(A %*% x == b)
   problem <- Problem(objective, constraints)</pre>
22 result <- solve(problem)
23 x_est <- result$aetValue(x)</pre>
24
25
26 # Plot results
27 plot(x_true, type = "h", lwd = 3, col = "blue", ylim = range(c(x_true, x_est)),
        main = "True vs Recovered Signal")
29 lines(x_est, type = "h", col = "red", lwd = 2)
30 legend("topright", legend = c("True", "Recovered"), col = c("blue", "red"), lwd = 2)
31
32 mean((x_est - x_true)^2)
```

# Basis pursuit example





# Terence Tao and Emmanuel Candes have been celebrated for understanding of compressed sensing



Define

$$\mathbb{C}(S) := \{ \Delta \in \mathbb{R}^d : \|\Delta_{S^c}\|_1 \leq \|\Delta_S\|_1 \}.$$

Define

$$\mathbb{C}(S) := \{ \Delta \in \mathbb{R}^d : \|\Delta_{S^c}\|_1 \leq \|\Delta_S\|_1 \}.$$

### Theorem

The following two are equivalent:

Define

$$\mathbb{C}(\mathcal{S}) := \{ \Delta \in \mathbb{R}^d : \|\Delta_{\mathcal{S}^c}\|_1 \leq \|\Delta_{\mathcal{S}}\|_1 \}.$$

### **Theorem**

The following two are equivalent:

► For any  $\theta^* \in \mathbb{R}^d$  with support  $\subset S$ , the basis pursuit program applied to the data  $(X, y = X\theta^*)$  has unique solution  $\hat{\theta} = \theta^*$ .

Define

$$\mathbb{C}(\mathcal{S}) := \{ \Delta \in \mathbb{R}^d : \|\Delta_{\mathcal{S}^c}\|_1 \leq \|\Delta_{\mathcal{S}}\|_1 \}.$$

#### **Theorem**

The following two are equivalent:

- For any  $\theta^* \in \mathbb{R}^d$  with support  $\subset S$ , the basis pursuit program applied to the data  $(X, y = X\theta^*)$  has unique solution  $\hat{\theta} = \theta^*$ .
- The restricted null space (RNS) property holds, i.e.,

$$\mathbb{C}(S) \cap \ker(X) = \{0\}.$$

Recall that  $ker(X) = \{\theta \in \mathbb{R}^d : X\theta = 0\}.$ 

► Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :

$$\mathbb{T}(\theta^*) = \{ \Delta \in \mathbb{R}^d : \|\theta^* + t\Delta\|_1 \le \|\theta^*\|_1 \text{ for some } t > 0 \}.$$

Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :

$$\mathbb{T}(\theta^*) = \{\Delta \in \mathbb{R}^d \ : \ \|\theta^* + t\Delta\|_1 \le \|\theta^*\|_1 \ \text{ for some } \ t > 0\}.$$

i.e., the set of descent directions for the norm  $\ell_1$  at the point  $\theta^*$ .

Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :

```
\mathbb{T}(\theta^*) = \{\Delta \in \mathbb{R}^d \,:\, \|\theta^* + t\Delta\|_1 \leq \|\theta^*\|_1 \ \text{ for some } \ t>0\}.
```

- i.e., the set of descent directions for the norm  $\ell_1$  at the point  $\theta^*.$
- ► Feasible set is  $\theta^* + \ker(X)$ , i.e.,  $\ker(X)$  is the set of feasible directions  $\Delta = \theta \theta^*$ .

- Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :
  - $\mathbb{T}(\theta^*) = \{\Delta \in \mathbb{R}^d \ : \ \|\theta^* + t\Delta\|_1 \le \|\theta^*\|_1 \ \text{ for some } \ t > 0\}.$
  - i.e., the set of descent directions for the norm  $\ell_1$  at the point  $\theta^*$ .
- ► Feasible set is  $\theta^* + \ker(X)$ , i.e.,  $\ker(X)$  is the set of feasible directions  $\Delta = \theta \theta^*$ .
- ▶ Hence, there is a minimizer other than  $\theta^*$  if and only if

$$\mathbb{T}(\theta^*) \cap \ker(X) \neq \{0\}.$$

► Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :

$$\mathbb{T}(\theta^*) = \{\Delta \in \mathbb{R}^d \ : \ \|\theta^* + t\Delta\|_1 \le \|\theta^*\|_1 \ \text{ for some } \ t > 0\}.$$

- i.e., the set of descent directions for the norm  $\ell_1$  at the point  $\theta^*$ .
- ► Feasible set is  $\theta^* + \ker(X)$ , i.e.,  $\ker(X)$  is the set of feasible directions  $\Delta = \theta \theta^*$ .
- ▶ Hence, there is a minimizer other than  $\theta^*$  if and only if

$$\mathbb{T}(\theta^*) \cap \ker(X) \neq \{0\}.$$

And so the minizer is  $\theta^*$  if and only if  $\mathbb{T}(\theta^*) \cap \ker(X) = \{0\}$ .

Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :

$$\mathbb{T}(\theta^*) = \{\Delta \in \mathbb{R}^d \,:\, \|\theta^* + t\Delta\|_1 \leq \|\theta^*\|_1 \ \text{ for some } \ t>0\}.$$

- i.e., the set of descent directions for the norm  $\ell_1$  at the point  $\theta^*$ .
- ► Feasible set is  $\theta^* + \ker(X)$ , i.e.,  $\ker(X)$  is the set of feasible directions  $\Delta = \theta \theta^*$ .
- ▶ Hence, there is a minimizer other than  $\theta^*$  if and only if

$$\mathbb{T}(\theta^*) \cap \ker(X) \neq \{0\}.$$

And so the minizer is  $\theta^*$  if and only if  $\mathbb{T}(\theta^*) \cap \ker(X) = \{0\}$ .

▶ It is enough to show that

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

Consider the tangent cone to the  $\ell_1$  ball (of radius  $\|\theta^*\|_1$ ) at  $\theta^*$ :

$$\mathbb{T}(\theta^*) = \{\Delta \in \mathbb{R}^d \,:\, \|\theta^* + t\Delta\|_1 \leq \|\theta^*\|_1 \ \text{ for some } \ t>0\}.$$

i.e., the set of descent directions for the norm  $\ell_1$  at the point  $\theta^*$ .

- ► Feasible set is  $\theta^* + \ker(X)$ , i.e.,  $\ker(X)$  is the set of feasible directions  $\Delta = \theta \theta^*$ .
- ▶ Hence, there is a minimizer other than  $\theta^*$  if and only if

$$\mathbb{T}(\theta^*) \cap \ker(X) \neq \{0\}.$$

And so the minizer is  $\theta^*$  if and only if  $\mathbb{T}(\theta^*) \cap \ker(X) = \{0\}$ .

It is enough to show that

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

since then

$$\mathbb{C}(S) \cap \ker(X) = \bigcup_{\theta \in \mathbb{R}^d : \operatorname{supp}(\theta) \subset S} (\mathbb{T}(\theta) \cap \ker(X)).$$

$$\mathbb{C}(\mathcal{S}) = \bigcup_{\theta \in \mathbb{R}^d : \operatorname{supp}(\theta) \subset \mathcal{S}} \mathbb{T}(\theta)$$

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

Let  $\mathbb{T}_1(\theta^*)$  be the subset of  $\mathbb{T}(\theta^*)$  with t = 1. We can work with  $\mathbb{T}_1(\theta^*)$  instead of  $\mathbb{T}(\theta^*)$ .

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

- Let  $\mathbb{T}_1(\theta^*)$  be the subset of  $\mathbb{T}(\theta^*)$  with t = 1. We can work with  $\mathbb{T}_1(\theta^*)$  instead of  $\mathbb{T}(\theta^*)$ .
- Notice that

$$\|\Delta + \theta^*\|_1 = \|\Delta_{S^c} + \theta^*_{S^c}\|_1 + \|\Delta_{S} + \theta^*_{S}\|_1$$

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \operatorname{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

- Let  $\mathbb{T}_1(\theta^*)$  be the subset of  $\mathbb{T}(\theta^*)$  with t = 1. We can work with  $\mathbb{T}_1(\theta^*)$  instead of  $\mathbb{T}(\theta^*)$ .
- Notice that

$$\|\Delta + \theta^*\|_1 = \|\Delta_{S^c} + \theta^*_{S^c}\|_1 + \|\Delta_S + \theta^*_S\|_1 = \|\Delta_{S^c}\|_1 + \|\Delta_S + \theta^*_S\|_1.$$

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

- Let  $\mathbb{T}_1(\theta^*)$  be the subset of  $\mathbb{T}(\theta^*)$  with t = 1. We can work with  $\mathbb{T}_1(\theta^*)$  instead of  $\mathbb{T}(\theta^*)$ .
- Notice that

$$\|\Delta + \theta^*\|_1 \ = \ \|\Delta_{S^c} + \theta^*_{S^c}\|_1 + \|\Delta_S + \theta^*_S\|_1 \ = \ \|\Delta_{S^c}\|_1 + \|\Delta_S + \theta^*_S\|_1.$$

Hence,  $\Delta \in \mathbb{T}_1(\theta^*)$  if and only if

$$\|\Delta_{S^c}\|_1 + \|\Delta_S + \theta_S^*\|_1 = \|\Delta + \theta^*\|_1$$

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

- Let  $\mathbb{T}_1(\theta^*)$  be the subset of  $\mathbb{T}(\theta^*)$  with t = 1. We can work with  $\mathbb{T}_1(\theta^*)$  instead of  $\mathbb{T}(\theta^*)$ .
- Notice that

$$\|\Delta + \theta^*\|_1 = \|\Delta_{S^c} + \theta^*_{S^c}\|_1 + \|\Delta_{S} + \theta^*_{S}\|_1 = \|\Delta_{S^c}\|_1 + \|\Delta_{S} + \theta^*_{S}\|_1.$$

Hence,  $\Delta \in \mathbb{T}_1(\theta^*)$  if and only if

$$\|\Delta_{S^c}\|_1 + \|\Delta_S + \theta_S^*\|_1 = \|\Delta + \theta^*\|_1 \le \|\theta^*\|_1 = \|\theta_S^*\|_1$$

$$\mathbb{C}(\mathcal{S}) = igcup_{ heta \in \mathbb{R}^d \colon \mathrm{supp}( heta) \subset \mathcal{S}} \mathbb{T}( heta)$$

- Let  $\mathbb{T}_1(\theta^*)$  be the subset of  $\mathbb{T}(\theta^*)$  with t = 1. We can work with  $\mathbb{T}_1(\theta^*)$  instead of  $\mathbb{T}(\theta^*)$ .
- Notice that

$$\|\Delta + \theta^*\|_1 \ = \ \|\Delta_{S^c} + \theta^*_{S^c}\|_1 + \|\Delta_S + \theta^*_S\|_1 \ = \ \|\Delta_{S^c}\|_1 + \|\Delta_S + \theta^*_S\|_1.$$

Hence,  $\Delta \in \mathbb{T}_1(\theta^*)$  if and only if

$$\|\Delta_{\mathcal{S}^c}\|_1 \, + \, \|\Delta_{\mathcal{S}} + \theta_{\mathcal{S}}^*\|_1 = \|\Delta + \theta^*\|_1 \leq \|\theta^*\|_1 = \|\theta_{\mathcal{S}}^*\|_1$$

if and only if

$$\|\Delta_{S^c}\|_1 \le \|\theta_S^*\|_1 - \|\Delta_S + \theta_S^*\|_1.$$

▶ We have shown that  $\Delta \in \mathbb{T}_1(\theta^*)$  if and only if  $\|\Delta_{S^c}\|_1 \leq \|\theta_S^*\|_1 - \|\Delta_S + \theta_S^*\|_1$ .

- ▶ We have shown that  $\Delta \in \mathbb{T}_1(\theta^*)$  if and only if  $\|\Delta_{S^c}\|_1 \leq \|\theta_S^*\|_1 - \|\Delta_S + \theta_S^*\|_1$ .

 $\|\Delta_{S^c}\|_1 \le \sup_{\theta^* \in \mathbb{R}^d} \{ \|\theta_S^*\|_1 - \|\Delta_S + \theta_S^*\|_1 \}$ 

▶ We have that 
$$\Delta \in \mathbb{T}_1(\theta^*)$$
 for some  $\theta^*$  with supp $(\theta^*) \subset S$  iff

- ▶ We have shown that  $\Delta \in \mathbb{T}_1(\theta^*)$  if and only if  $\|\Delta_{S^c}\|_1 \leq \|\theta_S^*\|_1 - \|\Delta_S + \theta_S^*\|_1$ .

▶ We have that 
$$\Delta \in \mathbb{T}_1(\theta^*)$$
 for some  $\theta^*$  with  $\mathrm{supp}(\theta^*) \subset S$  iff

$$\begin{aligned} \|\Delta_{\mathcal{S}^c}\|_1 &\leq \sup_{\theta^* \in \mathbb{R}^d} \{\|\theta_{\mathcal{S}}^*\|_1 - \|\Delta_{\mathcal{S}} + \theta_{\mathcal{S}}^*\|_1\} \\ &= \|\Delta_{\mathcal{S}}\|_1 \end{aligned}$$

by the triangle inequality and setting  $\theta_S^* = -\Delta_S$ .