# Data Mining Assignment 6

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#### 1: Employee

```
> library(xtable)
> data = read.csv('performance.csv')
> attach(data)
(a)
> midX = median(X)
> group1 = subset(data, X >= midX)
> group2 = subset(data, X < midX)</pre>
> Y0count1 = sum(!group1$Y)
> Y1count1 = sum(group1$Y)
> Y0count2 = sum(!group2$Y)
> Y1count2 = sum(group2$Y)
> Q = c('aboveMid', 'belowMid')
> YOCount = c(YOcount1, YOcount2)
> Y1Count = c(Y1count1, Y1count2)
> tableOfCount = data.frame(Q, YOCount, Y1Count)
> YOProp = YOCount / c(dim(group1)[1], dim(group2)[1])
> Y1Prop = Y1Count / c(dim(group1)[1], dim(group2)[1])
> tableOfProp = data.frame(Q, YOProp, Y1Prop)
> # xtable(tableOfCount, caption = 'tableOfCount')
> # xtable(tableOfProp, caption = 'tableOfProp')
```

	Q	Y0Count	Y1Count
1	aboveMid	4	10
2	belowMid	9	4

Table 1: tableOfCount

	Q	Y0Prop	Y1Prop
1	aboveMid	0.29	0.71
2	belowMid	0.69	0.31

Table 2: tableOfProp

It's obvious from the table we can see that employees are tend to performe better when their emotion is more stable.

The proportion of emotion stability above median employees who can perform a task is almost twice as much as those who's emotion stability is below median.

#### (b)

```
> quan = quantile(X)
> group1 = subset(data, X < quan[2] & X >= quan[1])
> group2 = subset(data, X < quan[3] & X >= quan[2])
> group3 = subset(data, X < quan[4] & X >= quan[3])
```

```
> group4 = subset(data, X >= quan[4])
> Q = c("Q1", "Q2", "Q3", "Q4")
> YOCount = c(sum(!group1$Y), sum(!group2$Y), sum(!group3$Y), sum(!group4$Y))
> Y1Count = c(sum(group1\$Y), sum(group2\$Y), sum(group3\$Y), sum(group4\$Y))
> YOprop = YOCount / c(dim(group1)[1], dim(group2)[1], dim(group3)[1], dim(group4)[1])
> Y1prop = Y1Count / c(dim(group1)[1], dim(group2)[1], dim(group3)[1],dim(group4)[1])
> tableOfCount = data.frame(Q, YOCount, Y1Count)
> tableOfProp = data.frame(Q, YOprop, Y1prop)
> ys = c(6,1,3,3,3,4,1,6)
> cs = rep(c("Y0", "Y1"), 4)
> qs = rep(Q, 1, each = 2)
> temp = data.frame(ys, cs, qs)
> ov = xtabs(ys ~ qs + cs, data = temp)
> ov
    CS
   YO Y1
qs
 Q1 6 1
 Q2 3 3
 Q3 3
 Q4 1 6
> prop.test(ov)
        4-sample test for equality of proportions without continuity
        correction
data: ov
X-squared = 7.2586, df = 3, p-value = 0.0641
alternative hypothesis: two.sided
sample estimates:
             prop 2
   prop 1
                       prop 3
                                 prop 4
0.8571429 0.5000000 0.4285714 0.1428571
> # summary(ov)
> # xtable(tableOfCount, caption = 'tableOfCount')
> # xtable(tableOfProp, caption = 'tableOfProp')
```

	Q	Y0Count	Y1Count
1	Q1	6	1
2	Q2	3	3
3	Q3	3	4
4	Q4	1	6

Table 3: tableOfCount

p-value of Pearson X-squared test is 0.064, null hypothesis rejected, which means there is association between X and Y, which is in consistant with (a).

```
(c)
> fit = glm(Y ~ X, family = binomial, data)
> summary(fit)
glm(formula = Y ~ X, family = binomial, data = data)
Deviance Residuals:
              1Q
                   Median
                                3Q
                                        Max
-1.7845 -0.8350
                   0.5065
                            0.8371
                                     1.7145
Coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) -10.308925
                        4.376997 -2.355
                                            0.0185 *
```

```
X 0.018920 0.007877 2.402 0.0163 *

---
Signif. codes: 0 âĂŸ***âĂŹ 0.001 âĂŸ**âĂŹ 0.01 âĂŸ*âĂŹ 0.05 âĂŸ.âĂŹ 0.1 âĂŸ âĂŹ 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 37.393 on 26 degrees of freedom
Residual deviance: 29.242 on 25 degrees of freedom
AIC: 33.242

Number of Fisher Scoring iterations: 4
```

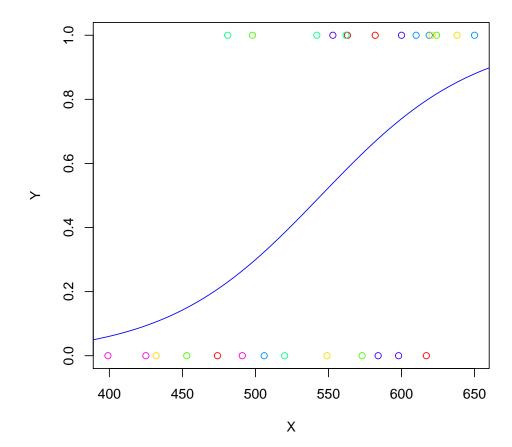
#### Assumtion:

Y are Benoulli random variables, independent, and  $P{Y=1}$  is related to X according to the specified logistic function.

As we can see from the summary,  $\beta_1$  is 0.0189. The intercept is -10.3.

(d)

```
> new.x = data.frame(X = seq(300, 1000))
> new.y <- predict(fit, newdata = new.x, se.fit=T, type="response")
> plot(X, Y, col = rainbow(7))
> lines(new.x$X, new.y$fit, col = "blue")
```



From the plot, I might say that the model fit is OK. But we can also see that there are lots of overlap from 450 to 600, so the model might not work well in this range.

(e)

#### > fit\$coefficients

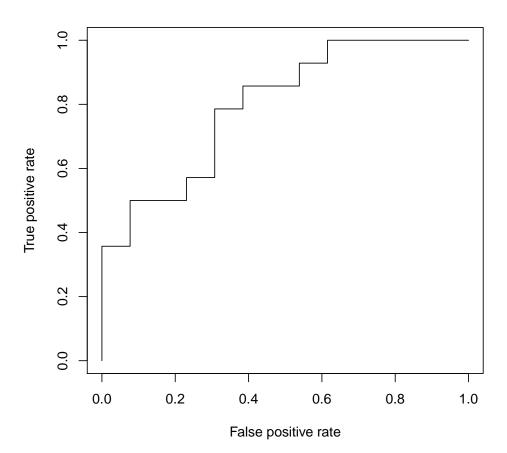
```
(Intercept) X -10.30892518 0.01891983
```

Since we get  $\beta_1 = 0.0189$  and its SE is only 0.0078 and its p-value is 0.0163, we are sure that there is association between X and Y. And the increase of log-odds of  $P{Y=1}$  with unit increase of X is 0.0189. This means that there is a positive association between X and Y. The bigger X, the more likely that Y is 1. This is in constant with (a) and (b) (f) > interval = confint(fit, level = c(0.9)) > interval 5 % 95 % (Intercept) -18.628406505 -3.90734449 0.007348506 0.03382477 > intervalX = c(interval[2, 1], interval[2, 2]) > e.beta1 = exp(intervalX) > exp(fit\$coefficients[2]) X 1.0191 > e.beta1 [1] 1.007376 1.034403 We get  $e^{\beta_1} = 1.0191$  and its 95% confident interval. This interval of e.beta1 means that: with 95% probability, unit increase in X will lead to odds of P{Y=1} increase by at minimum 1.007376 and at maximum 1.034403. **(g)** > predict(fit, data.frame(X = c(550)), type = "res") 0.5242263 (h) > pred.prob = predict(fit, data, type = "res") > pred.label = rep(0, 27) > pred.label[pred.prob > 0.5] = 1 > table(pred.label, Y) Y pred.label 0 1 0 8 3 1 5 11 We can see that there are 11 true positives, 8 true negatives, 5 false positives and 3 false negatives. (i) > library(ROCR) > pred <- prediction(pred.prob, labels = Y) > perf <- performance(pred, "tpr", "fpr")</pre>

[[1]] [1] 0.7967033

> plot(perf, colorize=F, main = "ROC curve")
> attributes(performance(pred, "auc"))\$y.values

## **ROC** curve



First, we can see that TPR increase with FPR. And they both reach 1.0 at upper-right corner. We can see that the AUC is 0.7967, which means that the model is on the right track, but not a very good job.

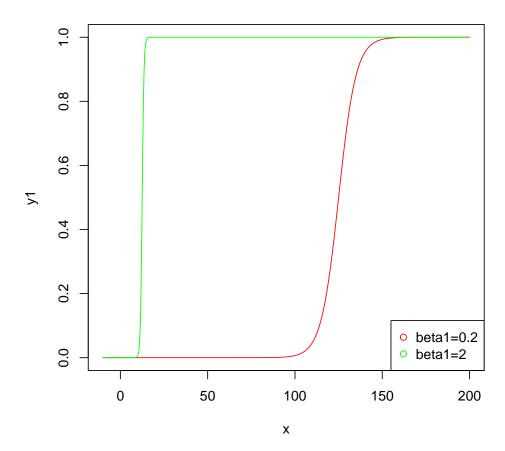
### 2: Experiment

```
> x = seq(from = -10, to = 200, length = 500)

(a)
> beta0 = -25
> beta1 = 0.2
> y1 = 1 / (1 + exp(-beta0 - beta1 * x))
> plot(y1 ~ x, col = 'red', type = 'l')
> beta11 = 2
> y2 = 1 / (1 + exp(-beta0 - beta11 * x))
> lines(y2 ~ x, col = 'green')
> title('sigmoid')
> legend('bottomright', legend = c('beta1=0.2', 'beta1=2'), col = c('red', 'green'), pch = 1:1)
```



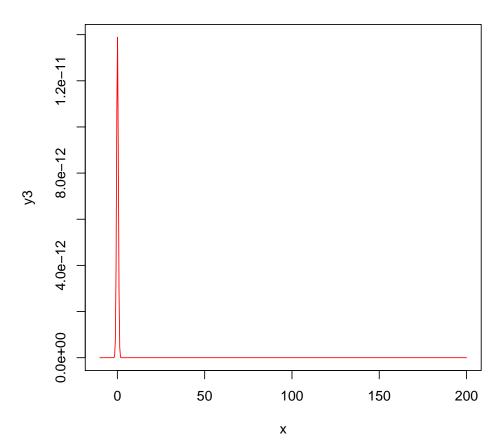
1 2



From the plot we can see that with a bigger  $\beta_1$ , the curve tends to be more steep and y reach 1 eailer. This is because  $\beta_1$  represents the change speed of log-odds of y = 1. With a bigger  $\beta_1$ , the log-odds becomes bigger. So y tends to reach 1 more quickly.

```
(b)
> beta2 = -2
> y3 = 1 / (1 + exp(-beta0 - beta1 * x - beta2 * I(x ^ 2)))
> plot(y3 ~ x, col = 'red', type = '1')
> title('sigmoid')
```





This plot is somewhat looks abnormal. But what it tells us is that  $X^2$  and Y has a strong negative association, and this association dominate. When x goes from -10 to 0,  $X^2$  becomes smaller so y goes up; When x passes 0,  $X^2$  becomes bigger, y goes down, until 0.