Generalized Distand Association

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Given two random vectors X and Y, we are interested in testing their probablistic association given n pairs of independent and identically distributed random samples $\{(X_i, Y_i)\}_{i=1}^n$. Peng et al. (2015) proposed Mean Distance Association (MeDiA), a set of probablistic association statistics as functions of observation distances. The theoretical foundation of MeDiA relies upon the result below:

Theorem 1. (from Peng et al. (2015)) Denote the distance between two independent random samples from (X,Y) as d_{XY} , and the distance between two dependent random samples from (\hat{X},\hat{Y}) as $d_{\hat{X}\hat{Y}}$. Then we have

$$E(d_{XY}) \ge E(d_{\hat{X}\hat{Y}})$$

In this paper, we would like to expand the theory above to general functions on the observation graph. The generalized mean distance would encompass a number of existing methods, like mutual information. Besides, the generalized mean distance naturally leads to the construction of several other probabilistic association statistics.

Theorem 2. (univariate g-transformation) Using the same notation, denote a monotonically increasing continuously differentiable function $g(\cdot)$. Denote the g-transformed distance as

$$\begin{array}{lcl} \tilde{d}_{XY} & = & g(d_{XY}) \\ \tilde{d}_{\hat{X}\hat{Y}} & = & g(d_{\hat{X}\hat{Y}}) \end{array}$$

Then we have:

$$E(\tilde{d}_{XY}) \ge E(\tilde{d}_{\hat{X}\hat{Y}})$$

and the average of the transformed distances \tilde{d} follow asymptotic normal distribution.

Proof. The proof follows directly from delta method. More specifically, for given $d_{\hat{X}\hat{Y}}$ and d_{XY} , there exists d'_{XY} , such that:

$$\tilde{d}_{\hat{X}\hat{Y}} = g(d_{\hat{X}\hat{Y}})
= g[(d_{\hat{X}\hat{Y}} - d_{XY}) + d_{XY}]
= g(d_{XY}) + g'(d'_{XY})(d_{\hat{X}\hat{Y}} - d_{XY})
= \tilde{d}_{XY} + g'(d'_{XY})(d_{\hat{X}\hat{Y}} - d_{XY})$$

Following Theorem 1, taking expectation on both sides, and realizing that $g'(\cdot) \geq 0$, we conclude the proof.

Following Theorem 2, we can see that distance based mutual information statistic $MI = \sum \log(d_{ij})$ actually falls into the generalized mean distance family.

However, it would be helpful to realize that Theorem 2 has not yet encompass functions on the observation graph that give different weights depending on the value. We would make this up with the results below:

Theorem 3. (Multivariate f-transformation) Using the same notation as above, n-variate function f is monotonically increasing on every dimension of input. Define

$$\bar{d}_{XY} = f(d_{i1}^{XY}, \dots, d_{in}^{XY})$$

$$\bar{d}_{\hat{X}\hat{Y}} = f(d_{i1}^{\hat{X}\hat{Y}}, \dots, d_{in}^{\hat{X}\hat{Y}})$$

Then we have:

$$E(\bar{d}_{XY}) \ge E(\bar{d}_{\hat{X}\hat{Y}})$$

Proof. When f is monotonically increasing and continuously differentiable, the proof to Theorem 3 is straight forward and similar to the proof to Theorem 2 usin delta method.

In addition, when f is monotonically increasing but not continuously differentiable, there exists a sequence of monotonically increasing and continuously differentiable functions $\{f_i(\cdot)\}_{i=1}^{+\infty}$, such that

$$\lim_{i \to \infty} ||f_i - f||_{d_{XY}} \to 0 \tag{1}$$

Define

$$\bar{d}_{XY}^{i} = f_{i}(d_{i1}^{XY}, \dots, d_{in}^{XY})
\bar{d}_{\hat{X}\hat{Y}}^{i} = f_{i}(d_{i1}^{\hat{X}\hat{Y}}, \dots, d_{in}^{\hat{X}\hat{Y}})$$

Then for each i, we have:

$$E(\bar{d}_{XY}^i) \ge E(\bar{d}_{\hat{X}\hat{Y}}^i)$$

Summing the results above, we have

$$E(\bar{d}_{XY}) \ge E(\bar{d}_{\hat{X}\hat{Y}}) \tag{2}$$

Theorem 3 shows that k-nearest neighbour edge sum as defined in Mira score, and k-nearest neighbour log edge sum as defined in Mutual Information, also falls into this category and can be used to identify random vector associations.

References

Peng, Hesen, Ma, Junjie, Bai, Yun, Lu, Jianwei, & Yu, Tianwei. 2015. Media: Mean distance association and its applications in nonlinear gene set analysis. *Plos one*, **10**(4), e0124620.