

# CSSS 512 HW 2

Chris Hess

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## Problem 1 - U.S. House of Representatives

### Part a

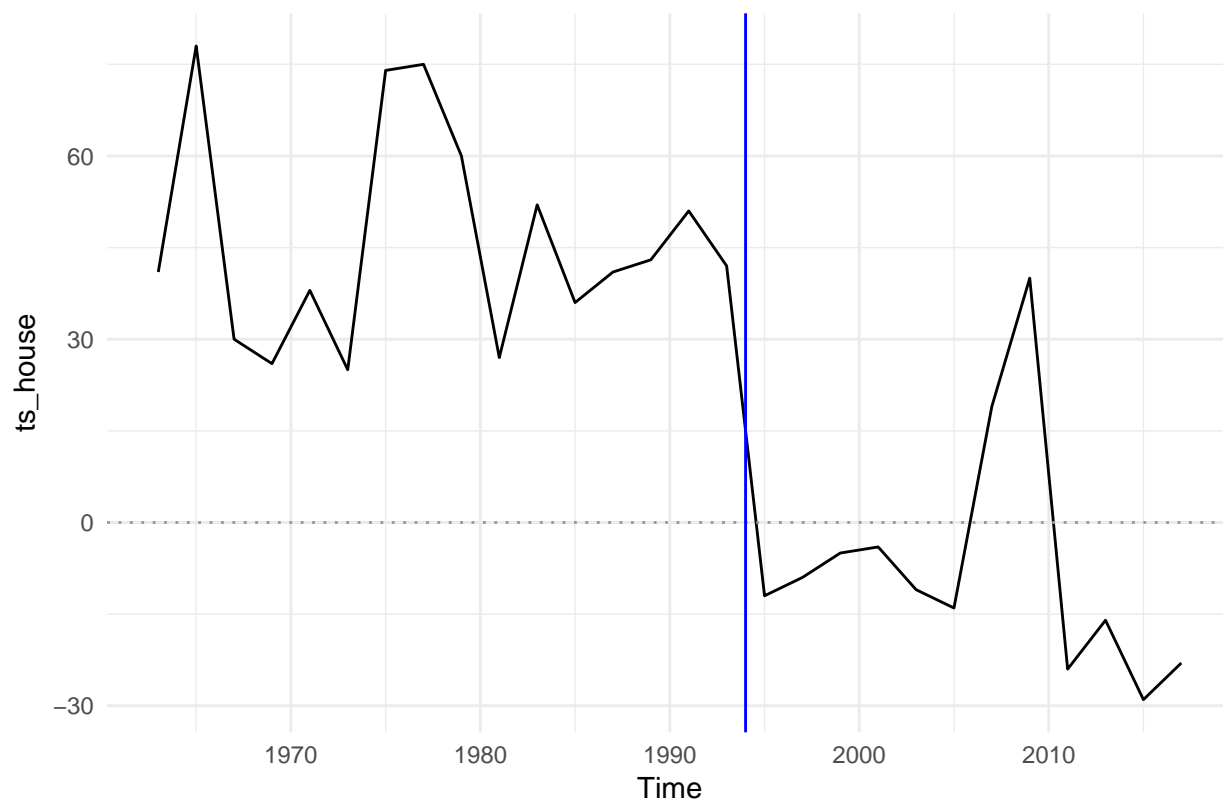
The time-series does show evidence of a substantial shock in 1994 where the average through the first portion of the series ( $\approx 46$  seat majority) does not approximate the short-run average post-1994 ( $\approx -7$ ). The ACF and PACF suggest that the time-series may be autoregressive order 1 before demeaning to the series, though there is not a strict pattern of geometric decline from L1. The Phillips-Perron and Augmented Dickey-Fuller tests disagree with each other, and this lack of conclusive evidence regarding stationarity likely is a function of the short series. Demeaning the series according to the 1994 structural break (i.e. pre and post) shows much stronger evidence in favor of stationarity, and also indicates that the initial evidence of autocorrelation was largely related to pre/post 1994 variation.

```
#plot ts, ACF, PACF
#run ADF and PP tests
#demean by pre/post 1994 periods
#plot ts, ACF, PACF
#what effect does 1994 structural break have on ts?

#pull ts, every other year frequency 1963-2017
ts_house <- ts(congress$DemHouseMaj,
               frequency = 1/2,
               start = 1963, end = 2017)

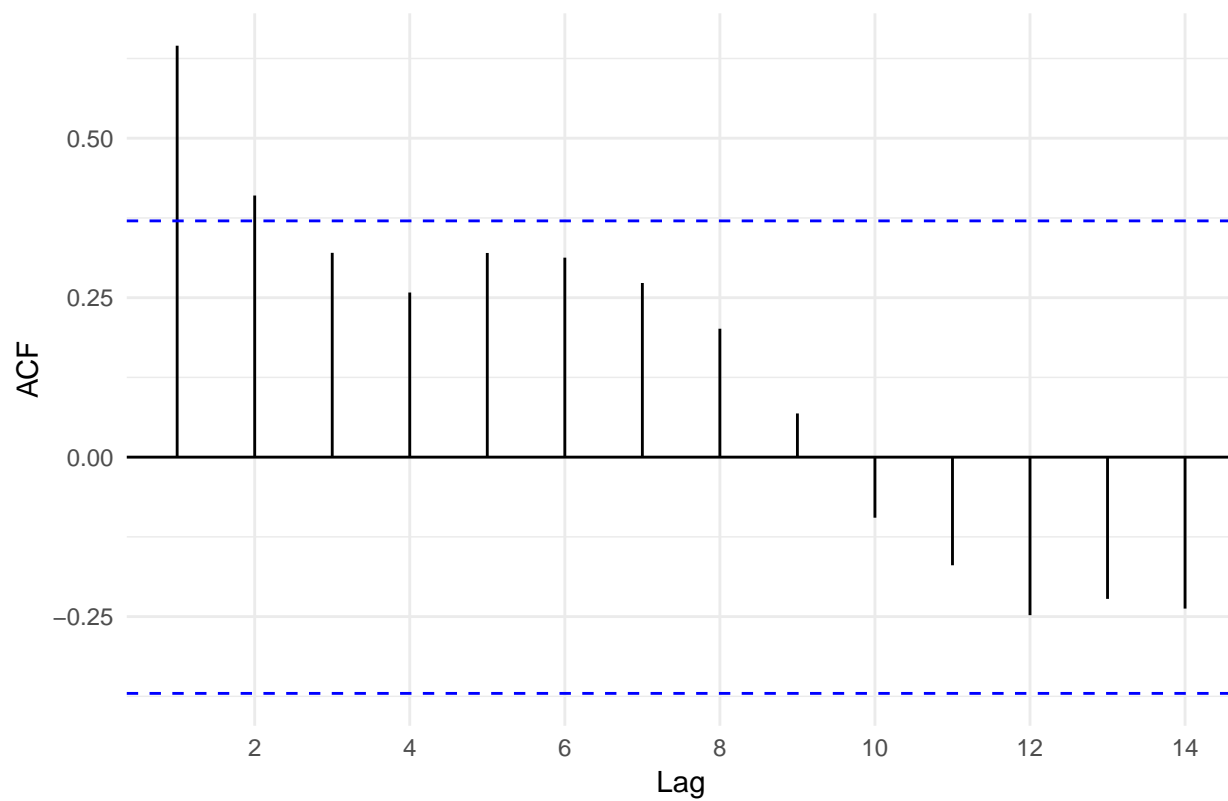
#plot as observed
autoplot(ts_house) +
  geom_hline(yintercept = 0, linetype = 3, color = "grey60") +
  geom_vline(xintercept = 1994, color = "Blue") +
  theme_minimal() +
  labs(title = "Democratic House Majority Time-series")
```

## Democratic House Majority Time-series



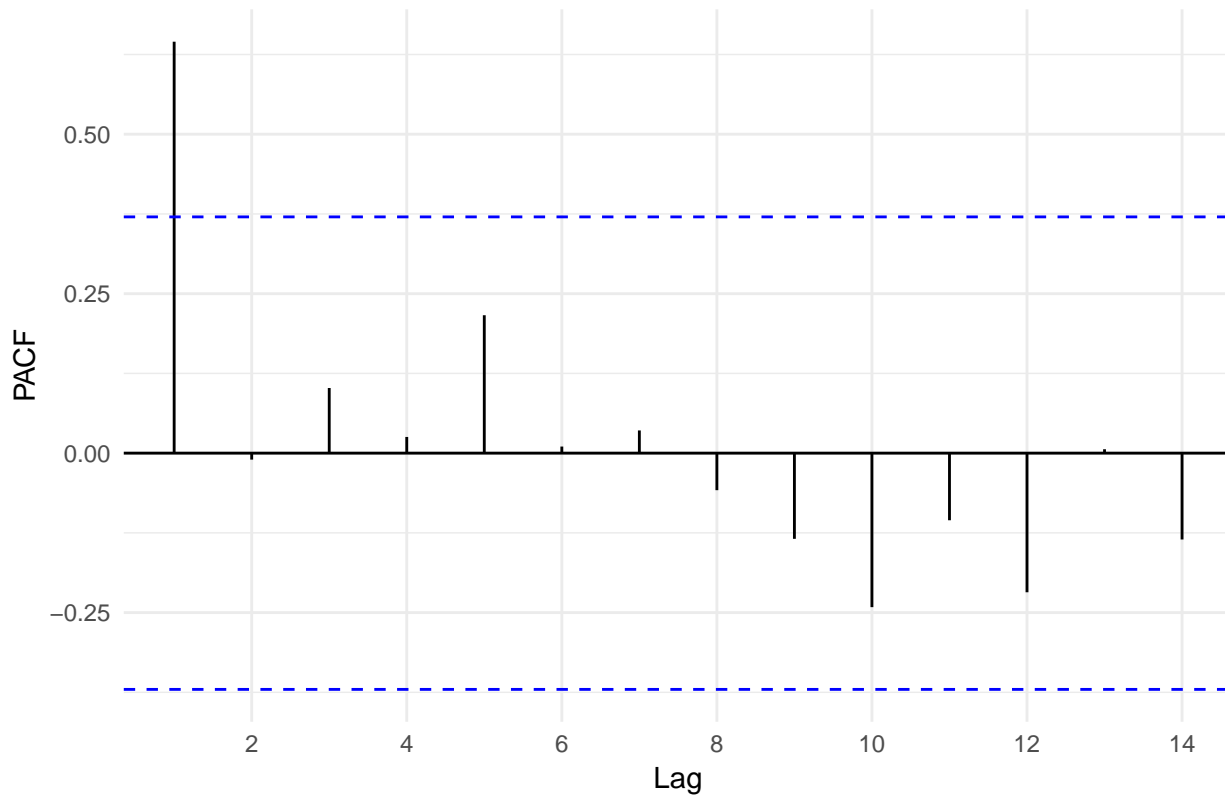
```
#observed ACF  
ggAcf(ts_house) + #substantial autocorrelation but not quite geometric decline  
  theme_minimal() +  
  labs(title = "Democratic House Majority ACF")
```

## Democratic House Majority ACF



```
#observed PACF
ggPacf(ts_house) + #AR(1), phi = .65 ish
  theme_minimal() +
  labs(title = "Democratic House Majority PACF")
```

## Democratic House Majority PACF



```
PP.test(ts_house)
```

```
##
## Phillips-Perron Unit Root Test
##
## data: ts_house
## Dickey-Fuller = -3.6974, Truncation lag parameter = 2, p-value =
## 0.04239
```

```
tseries::adf.test(ts_house)
```

```
##
## Augmented Dickey-Fuller Test
##
## data: ts_house
## Dickey-Fuller = -2.7382, Lag order = 3, p-value = 0.2901
## alternative hypothesis: stationary
```

```
#pre-1994 mean for ts
preMean <- congress %>%
  filter(StartYear < 1994) %>%
  summarize(mean = mean(DemHouseMaj)) %>%
  pull(mean)
```

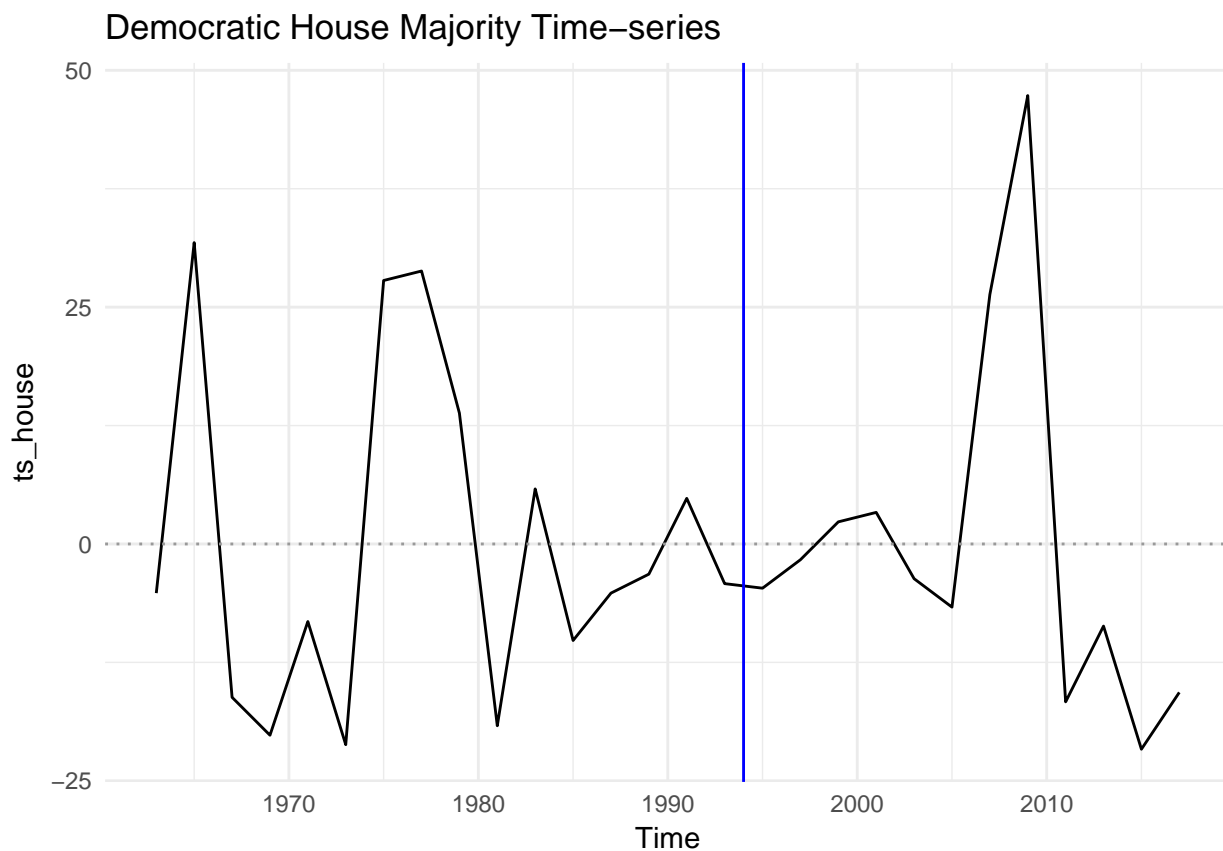
```
#post-1994 mean for ts
postMean <- congress %>%
  filter(StartYear >= 1994) %>%
  summarize(mean = mean(DemHouseMaj)) %>%
  pull(mean)
```

```

#demean based on pre/post 1994 (i.e. 1-16th obs vs 17-28th obs)
ts_house[1:16] <- ts_house[1:16] - preMean
ts_house[-1:-16] <- ts_house[-1:-16] - postMean

#demeaned ts - looks stationary now
autoplot(ts_house) +
  geom_hline(yintercept = 0, linetype = 3, color = "grey60") +
  geom_vline(xintercept = 1994, color = "Blue") +
  theme_minimal() +
  labs(title = "Democratic House Majority Time-series")

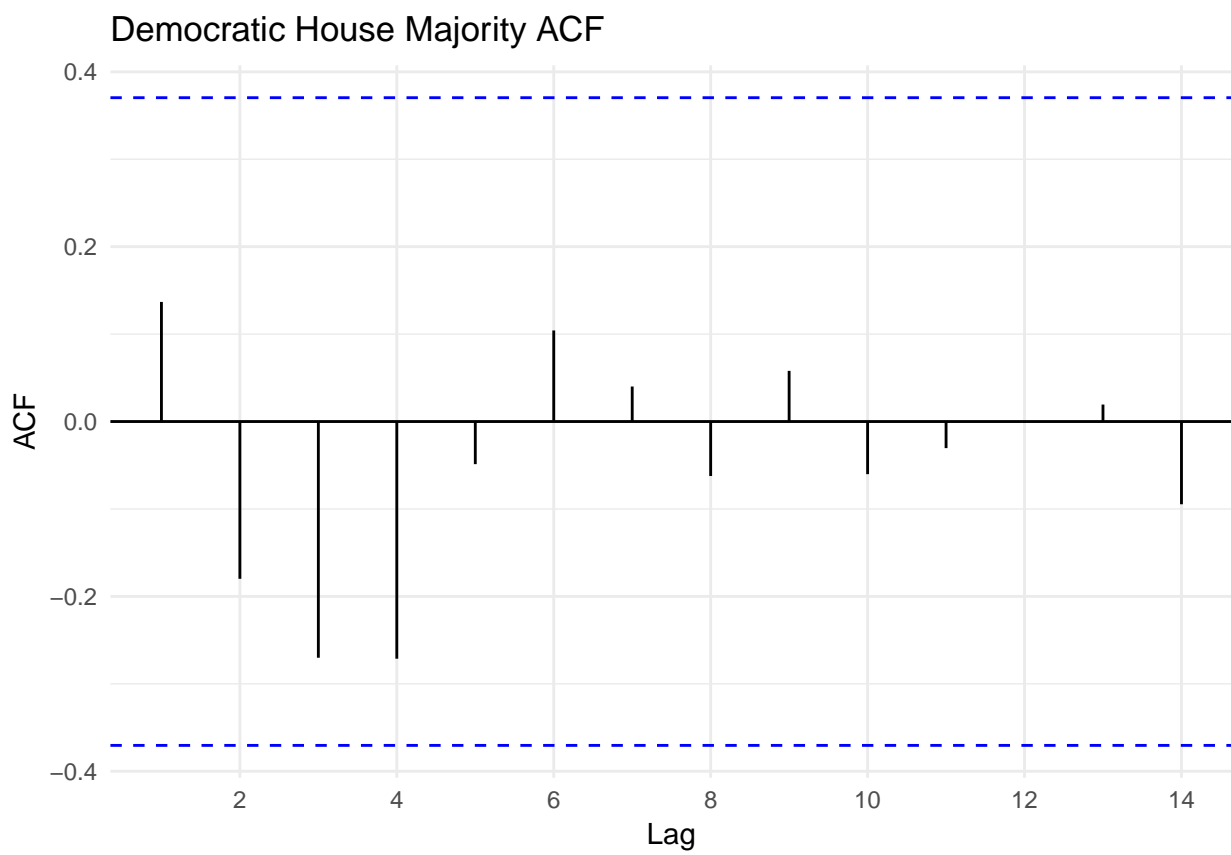
```



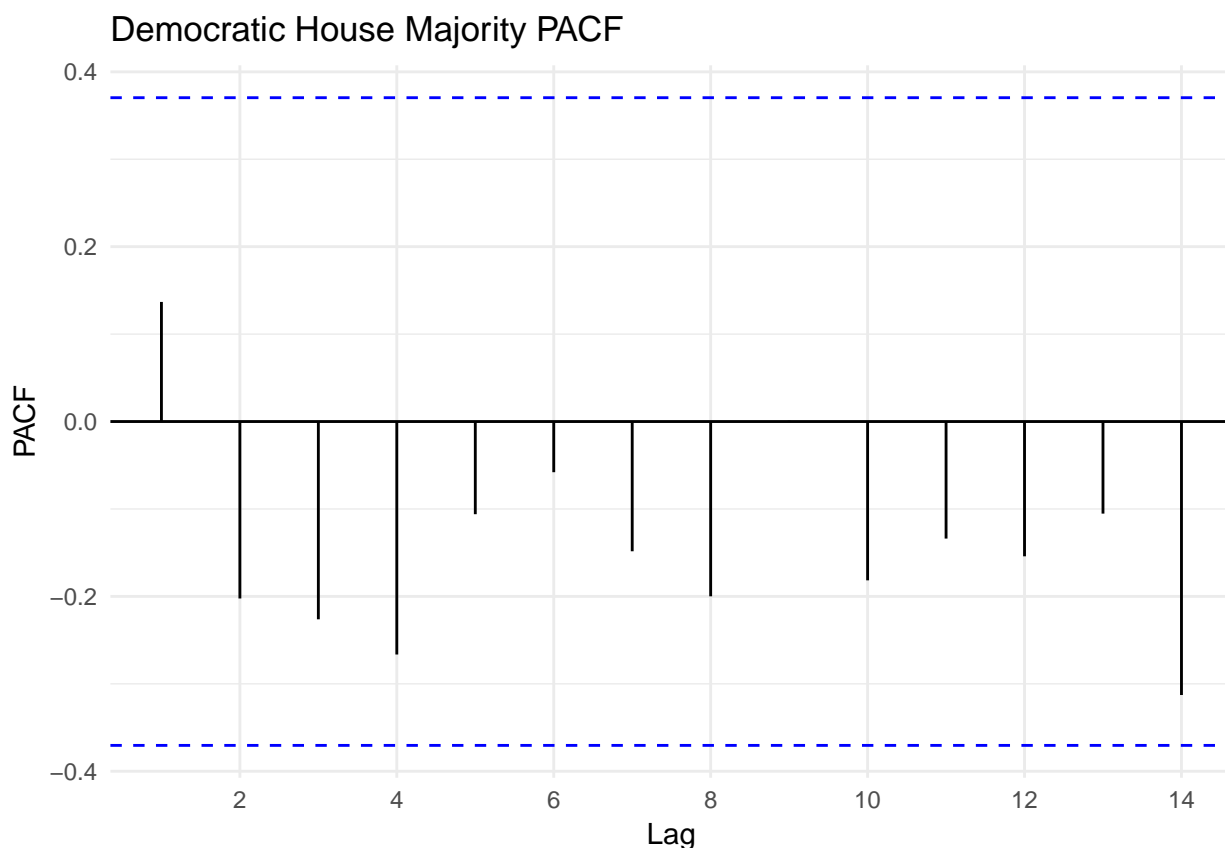
```

#demeaned ACF
ggAcf(ts_house) + #no autocorrelation
  theme_minimal() +
  labs(title = "Democratic House Majority ACF")

```



```
#demeaned PACF  
ggPacf(ts_house) + #no need for lags  
  theme_minimal() +  
  labs(title = "Democratic House Majority PACF")
```



#### Part b

The AR(0)/OLS model of the Democratic House majority time-series shows support for the coattails theory and for the re-alignment shock in 1994. The clearest relationships are that the Southern re-alignment switched the make-up of the House towards majority-Republican (a difference of nearly 48 seats for pre 1994 elections), and that the party of the newly-elected president (in presidential election years) typically sees a proportional boost in the House. The evidence for the partisan midterm resurgence is a bit uncertain but still suggests that midterms tend to associate with the party not holding the presidency taking back some seats. There doesn't seem to be much salience to unemployment for the Democratic House majority, net of the other covariates. The model tended to miss the true value of the time-series by about 12-14 seats depending on the specification.

```
#pull ts, every other year frequency 1963-2017
ts_house <- ts(congress$DemHouseMaj,
               frequency = 1/2,
               start = 1963, end = 2017)

#create df of covariates
covar <- congress %>% select(PartisanMidterm, PartisanUnem, Coattails, Pre1994)

#function to provide sum stats that can be passed to xtable()
armaFit <- function(ts, order = c(0, 0, 0), seasonal.order = c(1, 0, 0),
                    seasonal.period = NA, xdf = NULL){
  if(is.na(seasonal.period)){
    mod <- Arima(ts, order = order, xreg = xdf)
  } else{
    mod <- Arima(ts, order = order,
                  seasonal = list(order = seasonal.order,
                                  period = seasonal.period), xreg = xdf)
  }
}
```

```

lab.order <- paste0("(", order[1], ",", order[2], ",", order[3], ")")
aic <- round(mod[['aic']], 3)
rmse <- round(sqrt(mean((ts - mod$fitted)^2)), 3)
lenT <- length(ts)

phi1 <- round(mod$coef["ar1"], 3)
phi2 <- round(mod$coef["ar2"], 3)
psi1 <- round(mod$coef["ma1"], 3)
sar1 <- round(mod$coef["sar1"], 3)
phi1 <- ifelse(is.na(phi1), "", phi1)
phi2 <- ifelse(is.na(phi2), "", phi2)
psi1 <- ifelse(is.na(psi1), "", psi1)
sar1 <- ifelse(is.na(sar1), "", sar1)

coef <- round(mod[['coef']], 3)
coef <- round(coef[c((length(coef)-3),
                    (length(coef)-2),
                    (length(coef)-1),
                    length(coef))], 3)

se <- round(sqrt(diag(mod[['var.coef']])), 3)
se_phi1 <- round(se["ar1"], 3)
se_phi2 <- round(se["ar2"], 3)
se_psi1 <- round(se["ma1"], 3)
se_sar1 <- round(se["sar1"], 3)
se_phi1 <- ifelse(is.na(se_phi1), "", se_phi1)
se_phi2 <- ifelse(is.na(se_phi2), "", se_phi2)
se_psi1 <- ifelse(is.na(se_psi1), "", se_psi1)
se_sar1 <- ifelse(is.na(se_sar1), "", se_sar1)
se <- round(se[c((length(se)-3),
                (length(se)-2),
                (length(se)-1),
                length(se))], 3)

names(lab.order) <- "Order"
names(aic) <- "AIC"
names(rmse) <- "RMSE"
names(lenT) <- "$T$"
names(phi1) <- "$\\phi_1$"
names(phi2) <- "$\\phi_2$"
names(psi1) <- "$\\psi_1$"
names(sar1) <- "$\\phi_1$"

oneline <- c(lab.order, aic, rmse, lenT, phi1, phi2, psi1, sar1, coef)
twoline <- c("", "", "", "",
             ifelse(se_phi1 == "", "", paste0("(", se_phi1, ")")),
             ifelse(se_phi2 == "", "", paste0("(", se_phi2, ")")),
             ifelse(se_psi1 == "", "", paste0("(", se_psi1, ")")),
             ifelse(se_sar1 == "", "", paste0("(", se_sar1, ")")),
             paste0("(", se, ")"))
lines <- rbind(oneline, twoline)
rownames(lines) <- NULL

return(lines)
}

```



*#AR(0) with covariates*

```
ar0 <- armaFit(ts_house, order = c(0, 0, 0), xdf = covar)
```

```
print(xtable(ar0, digits = 3), booktabs = T, include.rownames = FALSE,
      sanitize.colnames.function=function(x){x})
```

Order	AIC	RMSE	$T$	$\phi_1$	$\phi_2$	$\psi_1$	$S\phi_1$	PartisanMidterm	PartisanUnem	Coattails	Pre1994
(0,0,0)	239.243	13.999	28					-7.27 (3.823)	-2.053 (1.733)	18.396 (5.331)	47.994 (5.703)

### Part c

Across the 5 specifications, there is fairly robust support for the 1994 structural break and the importance of presidential coattails. The partisan unemployment theory receives pretty weak support, since the coefficient is uncertain through 4 out of 5 specifications. The 5 models are largely indistinguishable in terms of AIC, and both the RMSE and SE of the regression are all relatively close across models too. The ARMA(1, 1) performs best in-sample, though this is somewhat expected given the additional model parameters. Accordingly, the improvement in fit for including both autoregressive and moving-average term is weaker than would be expected if the two parameters were both necessary (the AR(2) is similar in this respect). All models miss the true values of the time-series they were trained on by about 14 or 15 seats.

*#AR(1)*

```
ar1 <- armaFit(ts_house, order = c(1, 0, 0), xdf = covar)
```

*#AR(2)*

```
ar2 <- armaFit(ts_house, order = c(2, 0, 0), xdf = covar)
```

*#MA(1)*

```
ma1 <- armaFit(ts_house, order = c(0, 0, 1), xdf = covar)
```

*#ARMA(1,1)*

```
arma11 <- armaFit(ts_house, order = c(1, 0, 1), xdf = covar)
```

*#bind the sum stat rows together*

```
sums <- rbind(ar0, ar1, ar2, ma1, arma11)
```

```
print(xtable(sums, digits = 3), booktabs = T, include.rownames= FALSE,
      sanitize.colnames.function=function(x){x})
```

Order	AIC	RMSE	$T$	$\phi_1$	$\phi_2$	$\psi_1$	$S\phi_1$	PartisanMidterm	PartisanUnem	Coattails	Pre1994
(0,0,0)	239.243	13.999	28					-7.27 (3.823)	-2.053 (1.733)	18.396 (5.331)	47.994 (5.703)
(1,0,0)	240.221	13.732	28	0.234 (0.223)				-8.848 (3.86)	-2.42 (1.75)	15.364 (5.828)	46.656 (7.023)
(2,0,0)	239.662	13.034	28	0.448 (0.238)	-0.377 (0.215)			-10.726 (3.158)	-2.858 (1.744)	10.279 (5.862)	44.729 (6.038)
(0,0,1)	239.495	13.529	28				0.389 (0.3)	-9.866 (3.743)	-2.761 (1.841)	13.121 (6.315)	45.386 (7.58)
(1,0,1)	238.096	12.189	28	-0.382 (0.254)			1 (0.108)	-11.673 (3.204)	-3.846 (1.484)	14.522 (6.058)	42.887 (6.623)

### Part d

The best model given the in-sample and out-of-sample goodness of fit indicators is the AR(1). The in-sample indicators would suggest (somewhat weakly) that the ARMA(1,1) is the best model, however this more complicated model actually predicts worse out-of-sample

compared to the simpler AR(1) specification. Given that the two fits between AR(1) and ARMA(1,1) are not all that different, I'd opt for the more parsimonious model. The ARMA(1,1)'s psi of 1.0 is suspect too, though I'm less certain that this is a fatal flaw so much as a sign of odd fit.

```
#ar0
f_ar0 <- function(x, h){forecast(Arima(x, order=c(0,0,0)), h=h)}
e_ar0 <- tsCV(ts_house, f_ar0, h = 3,
             window = 20)
mae_ar0 <- round(apply(e_ar0, 2, function(x){mean(abs(ts_house - x), na.rm = T)}), 3)
avg_mae_ar0 <- round(mean(mae_ar0), 3)
names(avg_mae_ar0) <- "avgMAE"
e_ar0 <- c(ar0[1, 1], ar0[1, 2], ar0[1, 3], mae_ar0, avg_mae_ar0)

#ar1
f_ar1 <- function(x, h){forecast(Arima(x, order=c(1,0,0)), h=h)}
e_ar1 <- tsCV(ts_house, f_ar1, h = 3,
             window = 20)
mae_ar1 <- round(apply(e_ar1, 2, function(x){mean(abs(ts_house - x), na.rm = T)}), 3)
avg_mae_ar1 <- round(mean(mae_ar1), 3)
names(avg_mae_ar1) <- "avgMAE"
e_ar1 <- c(ar1[1, 1], ar1[1, 2], ar1[1, 3], mae_ar1, avg_mae_ar1)

#ar2
f_ar2 <- function(x, h){forecast(Arima(x, order=c(2,0,0)), h=h)}
e_ar2 <- tsCV(ts_house, f_ar2, h = 3,
             window = 20)
mae_ar2 <- round(apply(e_ar2, 2, function(x){mean(abs(ts_house - x), na.rm = T)}), 3)
avg_mae_ar2 <- round(mean(mae_ar2), 3)
names(avg_mae_ar2) <- "avgMAE"
e_ar2 <- c(ar2[1, 1], ar2[1, 2], ar2[1, 3], mae_ar2, avg_mae_ar2)

#ma1
f_ma1 <- function(x, h){forecast(Arima(x, order=c(0,0,1)), h=h)}
e_ma1 <- tsCV(ts_house, f_ma1, h = 3,
             window = 20)
mae_ma1 <- round(apply(e_ma1, 2, function(x){mean(abs(ts_house - x), na.rm = T)}), 3)
avg_mae_ma1 <- round(mean(mae_ma1), 3)
names(avg_mae_ma1) <- "avgMAE"
e_ma1 <- c(ma1[1, 1], ma1[1, 2], ma1[1, 3], mae_ma1, avg_mae_ma1)

#arma1,1
f_arma11 <- function(x, h){forecast(Arima(x, order=c(1,0,1)), h=h)}
e_arma11 <- tsCV(ts_house, f_arma11, h = 3,
               window = 20)
mae_arma11 <- round(apply(e_arma11, 2, function(x){mean(abs(ts_house - x), na.rm = T)}), 3)
avg_mae_arma11 <- round(mean(mae_arma11), 3)
names(avg_mae_arma11) <- "avgMAE"
e_arma11 <- c(arma11[1, 1], arma11[1, 2], arma11[1, 3], mae_arma11, avg_mae_arma11)

#compile model fit stats
forecastMAE <- rbind(e_ar0, e_ar1, e_ar2, e_ma1, e_arma11)

print(xtable(forecastMAE), booktabs = T, include.rownames=FALSE)
```

Order	AIC	RMSE	h=1	h=2	h=3	avgMAE
(0,0,0)	239.243	13.999	30.975	41.386	42.208	38.19
(1,0,0)	240.221	13.732	26.078	39.36	37.721	34.386
(2,0,0)	239.662	13.034	26.944	39.481	37.505	34.643
(0,0,1)	239.495	13.529	30.738	41.285	41.653	37.892
(1,0,1)	238.096	12.189	27.598	39.671	37.665	34.978

## Part e

The coattails effect is what the scenarios highlight more than anything. The unemployment rate is fairly inconsequential for the level of the forecasts, however, predicting that the Democrats lose the presidential election in 2020 would essentially wipe out the forecasted increase in Democratic House seats for  $h=2$  and  $h=3$ . The variation around the predictions is somewhat large though—e.g. f

```
#selecting AR(1) as final model
ar1 <- Arima(ts_house, order = c(1, 0, 0), xreg = covar)

#counterfactual 1 - unemployment stays at 4.6% for all three elections
cf1 <- data.frame(
  "Pre1994" = rep(0, 3), #all forecasts are post1994
  "ParisanUnemp" = rep(4.6, 3),
  "PartisanMidterm" = c(-1, 0, 1),
  "Coattails" = c(0, 1, 0)
)
cf1
```

```
##   Pre1994 ParisanUnemp PartisanMidterm Coattails
## 1      0          4.6             -1          0
## 2      0          4.6              0          1
## 3      0          4.6              1          0
```

```
#forecast forward three periods based on cf1 X's
pred_cf1 <- predict(ar1, newxreg = cf1)

##counterfactual 2 - unemployment falls to 3.6% for all three elections
cf2 <- data.frame(
  "Pre1994" = rep(0, 3), #all forecasts are post1994
  "ParisanUnemp" = rep(3.6, 3),
  "PartisanMidterm" = c(-1, 0, 1),
  "Coattails" = c(0, 1, 0)
)
cf2
```

```
##   Pre1994 ParisanUnemp PartisanMidterm Coattails
## 1      0          3.6             -1          0
## 2      0          3.6              0          1
## 3      0          3.6              1          0
```

```
#forecast forward three periods based on cf2 X's
pred_cf2 <- predict(ar1, newxreg = cf2)

##counterfactual 2 - unemployment rises to 5.6% for all three elections
cf3 <- data.frame(
  "Pre1994" = rep(0, 3), #all forecasts are post1994
  "ParisanUnemp" = rep(5.6, 3),
  "PartisanMidterm" = c(-1, 0, 1),
  "Coattails" = c(0, 1, 0)
)
```

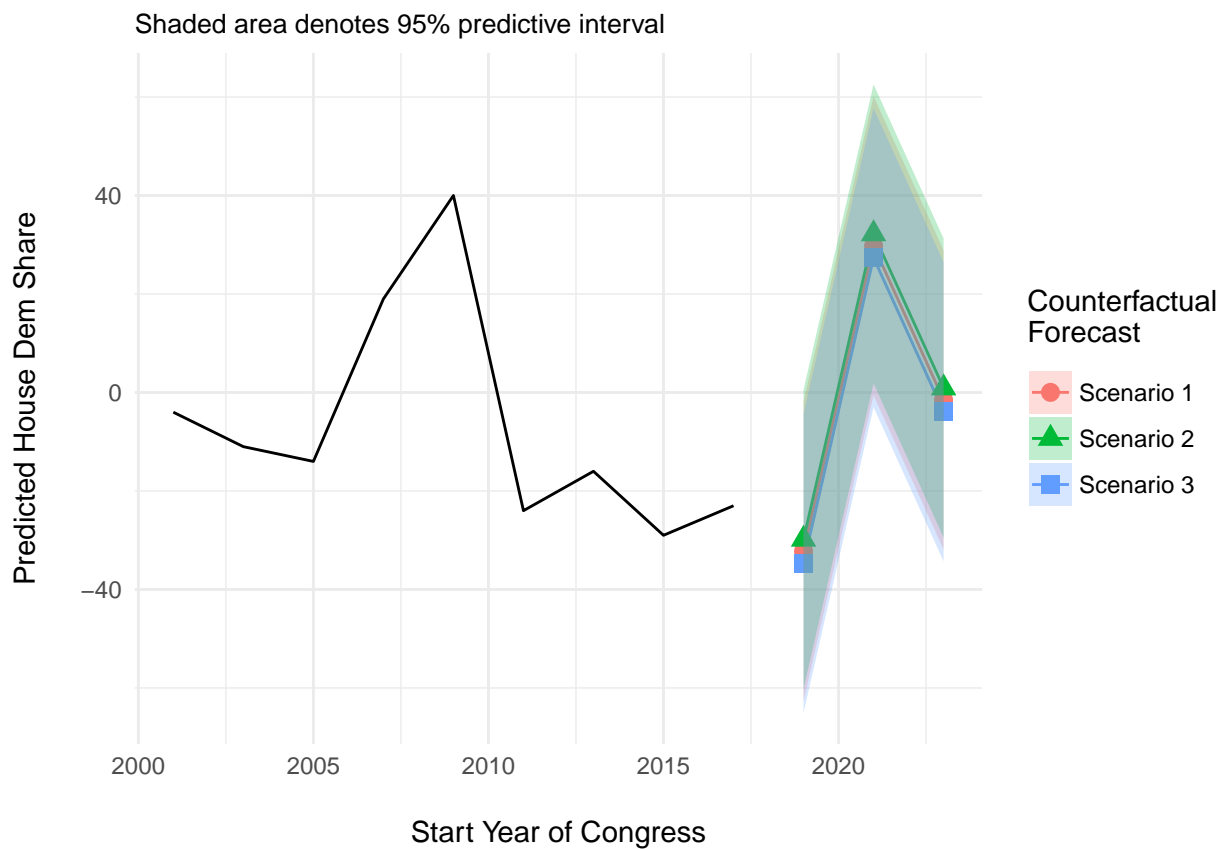
```
cf3

##   Pre1994 PartisanUnemp PartisanMidterm Coattails
## 1      0          5.6          -1          0
## 2      0          5.6           0          1
## 3      0          5.6           1          0

#forecast forward three periods based on cf2 X's
pred_cf3 <- predict(ar1, newxreg = cf3)

#construct tidy matrix for value, upper and lower
pred_vals <- data.frame(
  cf = c(rep("Scenario 1", 3), rep("Scenario 2", 3), rep("Scenario 3", 3)),
  time = rep(c(2019, 2021, 2023), 3),
  values = c(pred_cf1$pred, pred_cf2$pred, pred_cf3$pred),
  upper = c(pred_cf1$pred[1]+1.96*pred_cf1$se,
    pred_cf1$pred[2]+1.96*pred_cf1$se,
    pred_cf1$pred[3]+1.96*pred_cf1$se,
    pred_cf2$pred[1]+1.96*pred_cf2$se,
    pred_cf2$pred[2]+1.96*pred_cf2$se,
    pred_cf2$pred[3]+1.96*pred_cf2$se,
    pred_cf3$pred[1]+1.96*pred_cf3$se,
    pred_cf3$pred[2]+1.96*pred_cf3$se,
    pred_cf3$pred[3]+1.96*pred_cf3$se),
  lower = c(pred_cf1$pred[1]-1.96*pred_cf1$se,
    pred_cf1$pred[2]-1.96*pred_cf1$se,
    pred_cf1$pred[3]-1.96*pred_cf1$se,
    pred_cf2$pred[1]-1.96*pred_cf2$se,
    pred_cf2$pred[2]-1.96*pred_cf2$se,
    pred_cf2$pred[3]-1.96*pred_cf2$se,
    pred_cf3$pred[1]-1.96*pred_cf3$se,
    pred_cf3$pred[2]-1.96*pred_cf3$se,
    pred_cf3$pred[3]-1.96*pred_cf3$se))

ggplot() +
  geom_line(data=pred_vals, aes(x = time, y = values,
                                group = cf, color = cf)) +
  geom_point(data = pred_vals, aes(x = time, y = values,
                                   group = cf, shape = cf, color = cf),
             size = 3) +
  geom_ribbon(data = pred_vals, aes(x = time, ymin = lower, ymax = upper,
                                   group = cf, fill = cf),
            color = NA, alpha = .25) +
  geom_line(data = congress %>% filter(StartYear >= 2000), aes(x = StartYear, y = DemHouseMaj),
            color = "black") +
  scale_x_continuous() +
  xlab("\nStart Year of Congress") +
  ylab("Predicted House Dem Share\n") +
  labs(subtitle = "Shaded area denotes 95% predictive interval",
       fill = "Counterfactual\nForecast",
       color = "Counterfactual\nForecast",
       shape = "Counterfactual\nForecast") +
  theme_minimal()
```



## Problem 2 - U.S. Senate

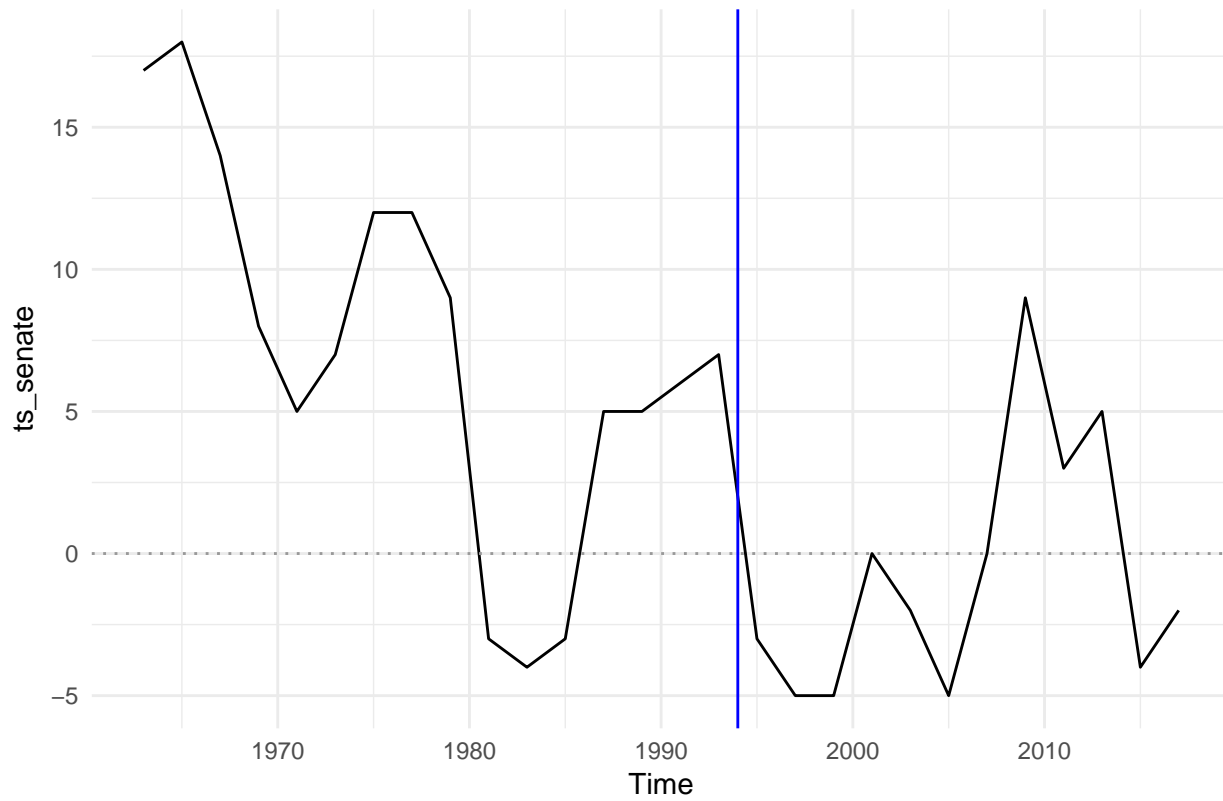
### Part a

```
#plot ts, ACF, PACF
#run ADF and PP tests
#demean by pre/post 1994 periods
#plot ts, ACF, PACF
#what effect does 1994 structural break have on ts?

#pull ts, every other year frequency 1963-2017
ts_senate <- ts(congress$DemSenateMaj,
  frequency = 1/2,
  start = 1963, end = 2017)

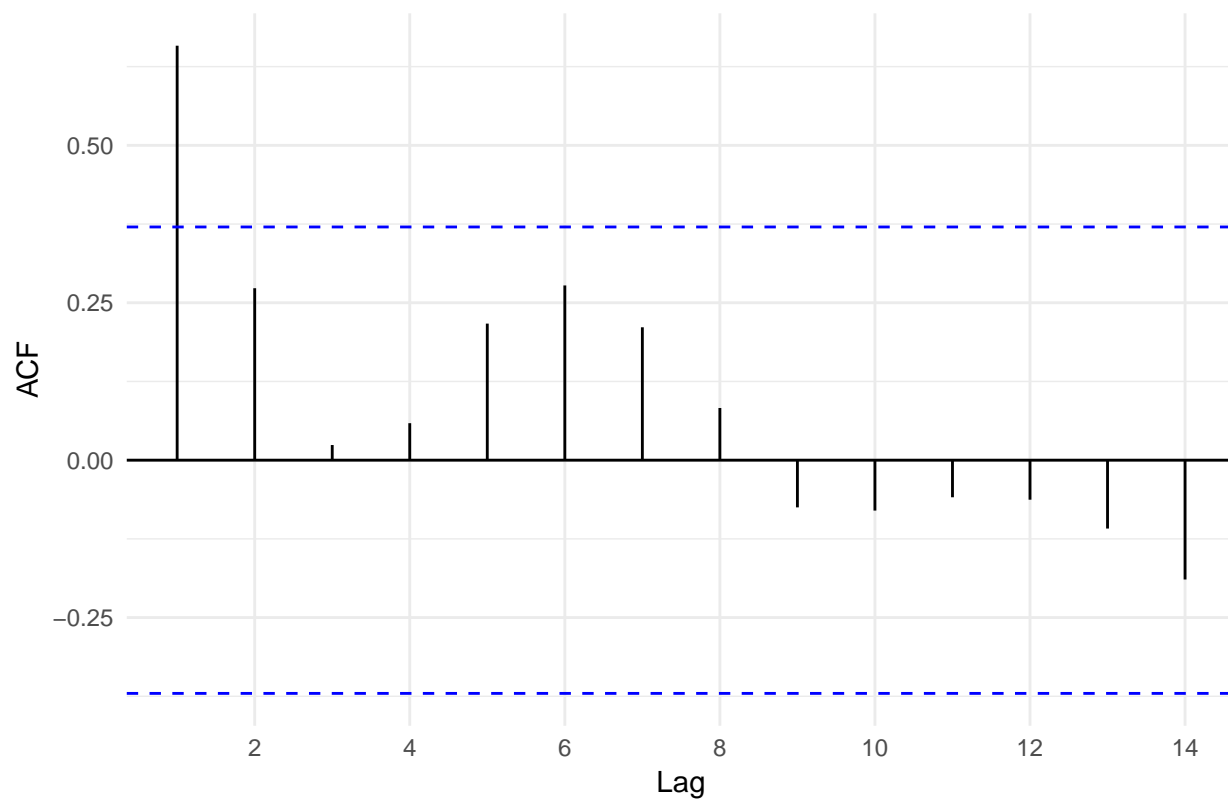
#plot as observed
autoplot(ts_senate) +
  geom_hline(yintercept = 0, linetype = 3, color = "grey60") +
  geom_vline(xintercept = 1994, color = "Blue") +
  theme_minimal() +
  labs(title = "Democratic Senate Majority Time-series")
```

## Democratic Senate Majority Time-series



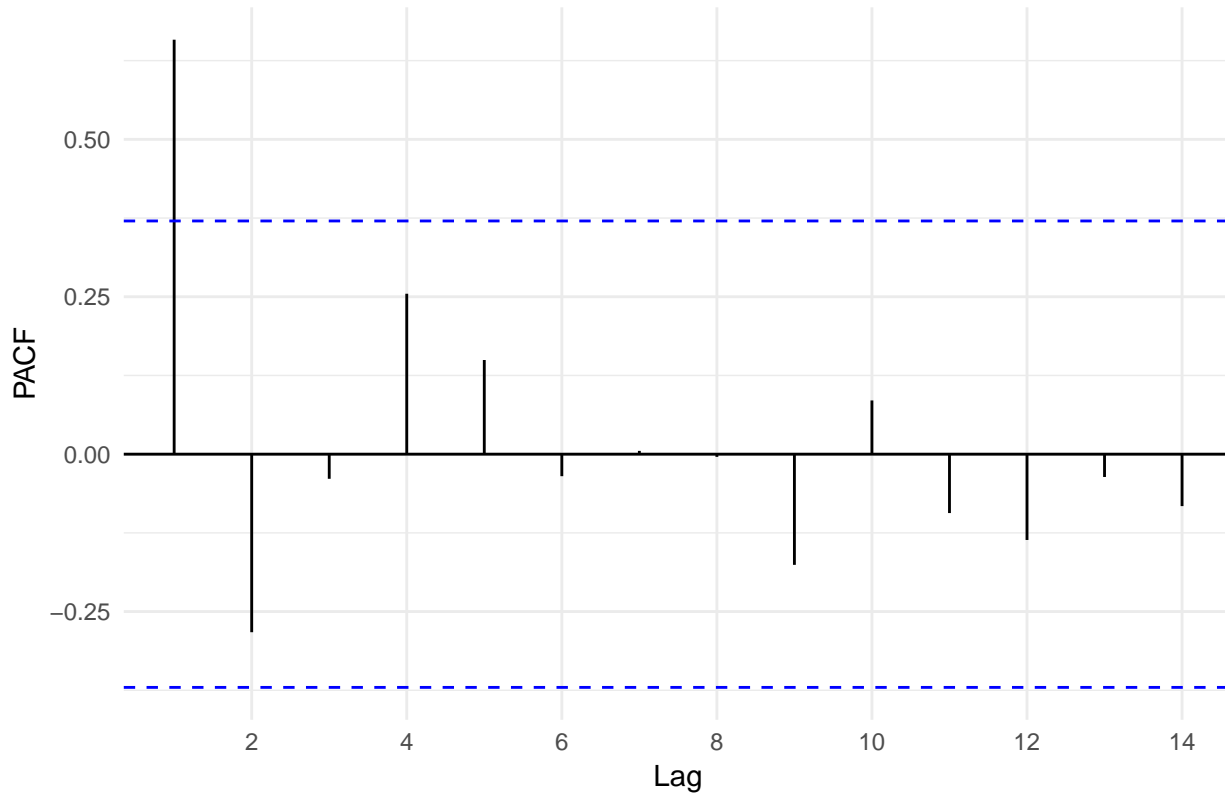
```
#observed ACF
ggAcf(ts_senate) + #could be AR(1)
  theme_minimal() +
  labs(title = "Democratic Senate Majority ACF")
```

## Democratic Senate Majority ACF



```
#observed PACF
ggPacf(ts_senate) + #AR(1), phi = .65 ish
  theme_minimal() +
  labs(title = "Democratic Senate Majority PACF")
```

## Democratic Senate Majority PACF



```
#pre-1994 mean for ts
preMean <- congress %>%
  filter(StartYear < 1994) %>%
  summarize(mean = mean(DemSenateMaj)) %>%
  pull(mean)

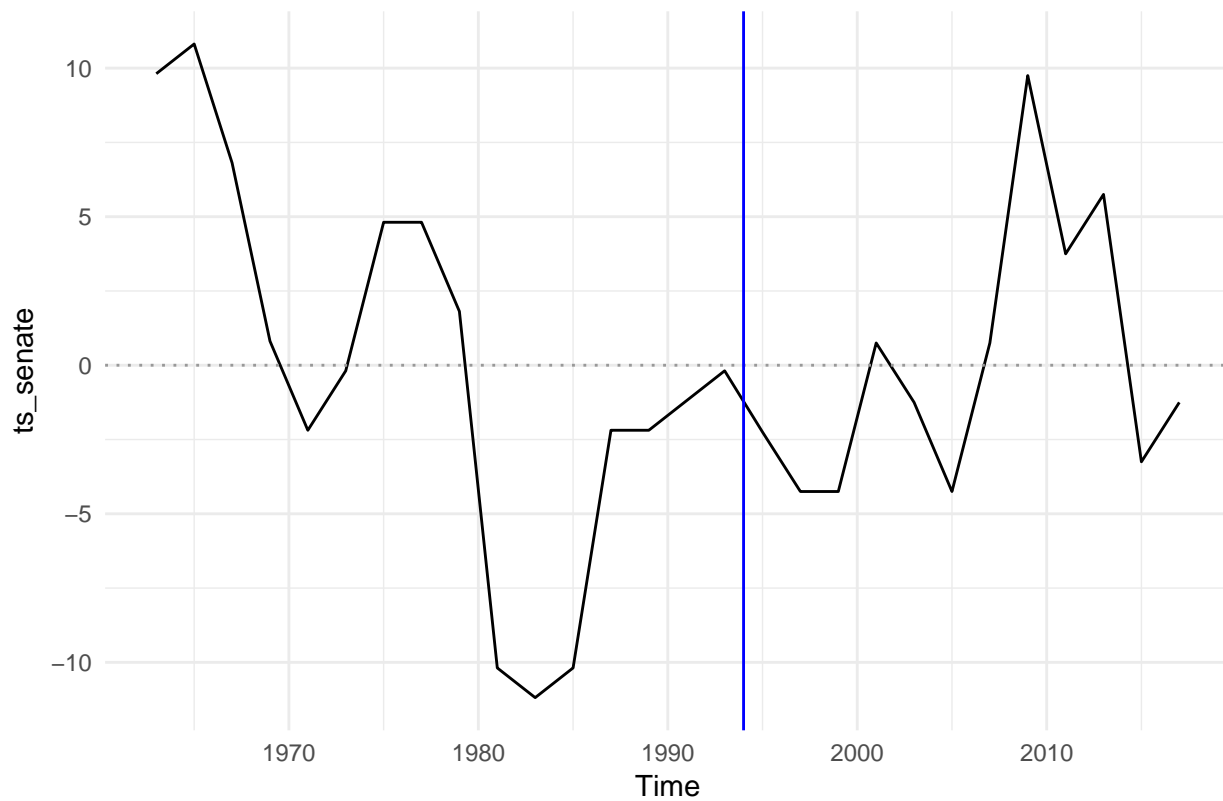
#post-1994 mean for ts
postMean <- congress %>%
  filter(StartYear >= 1994) %>%
  summarize(mean = mean(DemSenateMaj)) %>%
  pull(mean)

#demean based on pre/post 1994 (i.e. 1-16th obs vs 17-28th obs)
ts_senate[1:16] <- ts_senate[1:16] - preMean
ts_senate[-1:-16] <- ts_senate[-1:-16] - postMean

#demeaned ts - looks stationary now
autoplot(ts_senate) +
  geom_hline(yintercept = 0, linetype = 3, color = "grey60") +
  geom_vline(xintercept = 1994, color = "Blue") +
  theme_minimal() +
  labs(title = "Democratic Senate Majority Time-series")
```

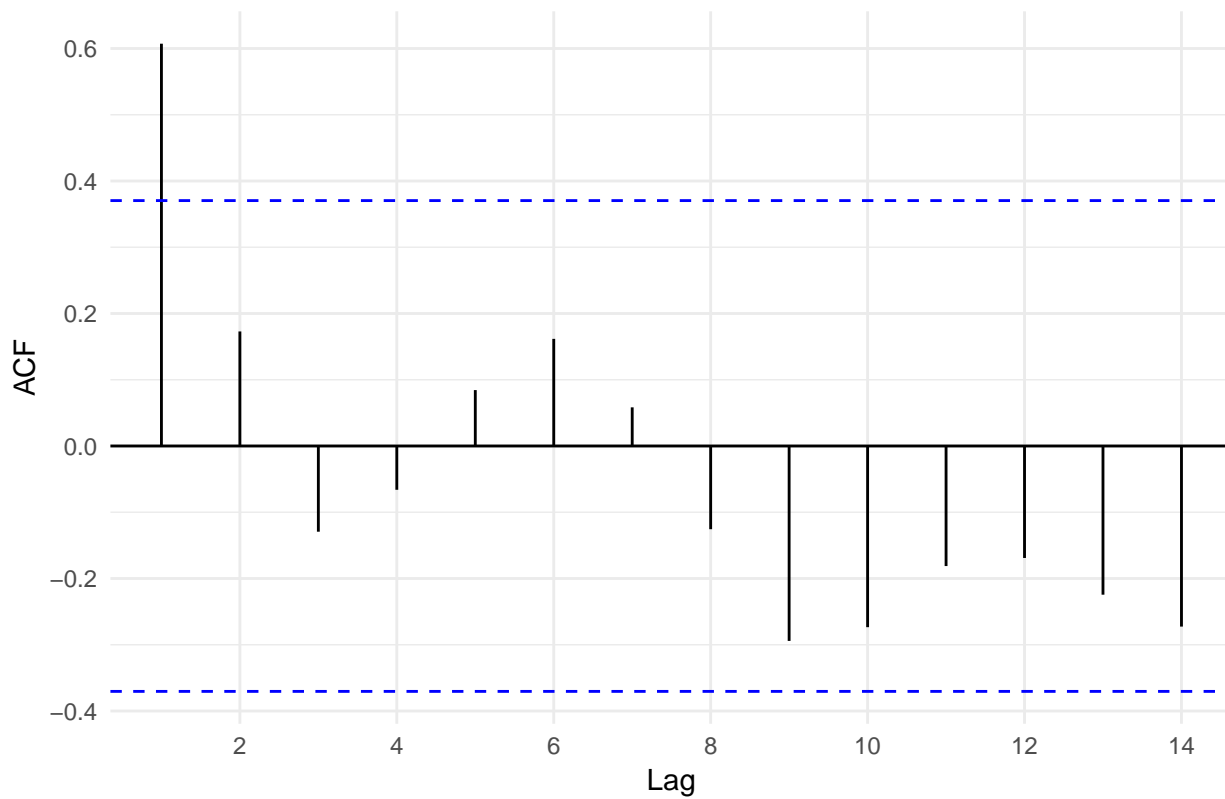


## Democratic Senate Majority Time-series



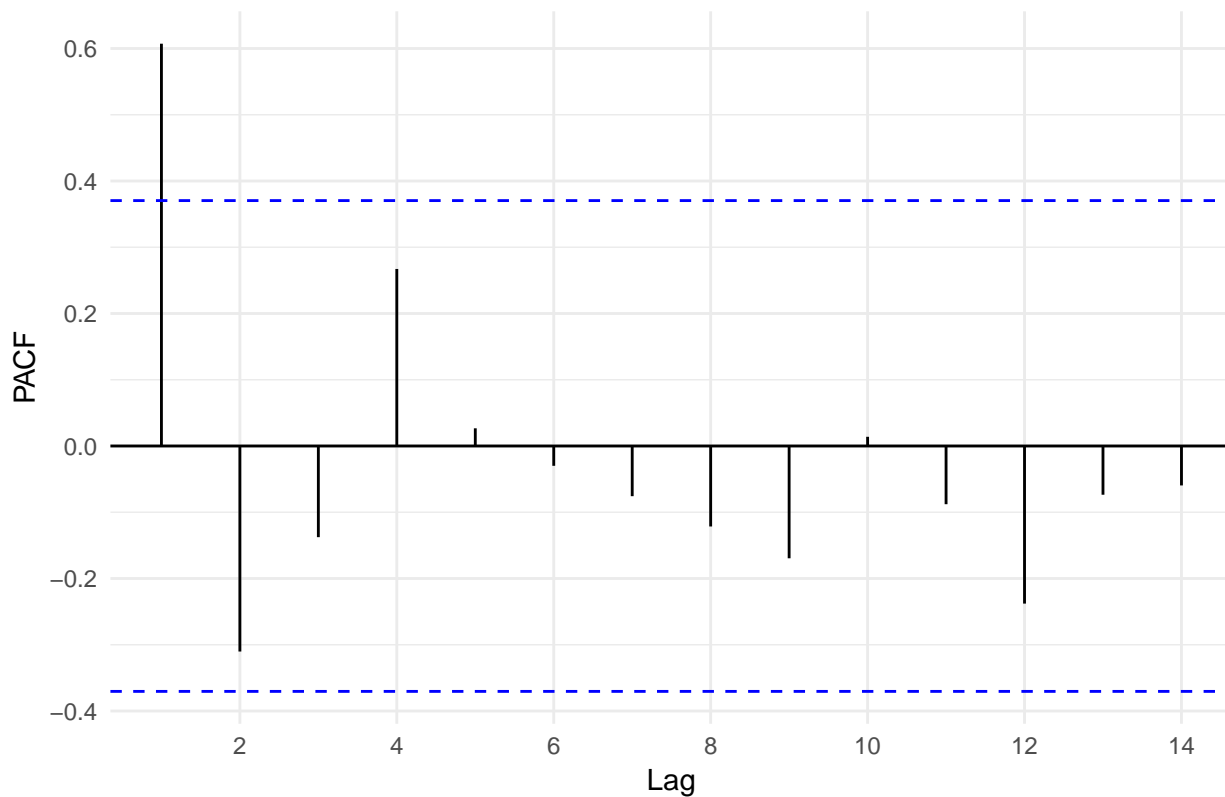
```
#demeaned ACF
ggAcf(ts_senate) + #still shows autocorrelation
  theme_minimal() +
  labs(title = "Democratic Senate Majority ACF")
```

## Democratic Senate Majority ACF



```
#demeaned PACF  
ggPacf(ts_senate) + #need an AR(1) with phi = .6 for this ts  
  theme_minimal() +  
  labs(title = "Democratic Senate Majority PACF")
```

## Democratic Senate Majority PACF



### Part b

```
#pull ts, every other year frequency 1963-2017
ts_senate <- ts(congress$DemSenateMaj,
               frequency = 1/2,
               start = 1963, end = 2017)

#create df of covariates
covar <- congress %>% select(PartisanMidterm, PartisanUnem, Coattails, Pre1994)

#AR(0) with covariates
ar0 <- armaFit(ts_senate, order = c(0, 0, 0), xdf = covar)

#AR(1)
ar1 <- armaFit(ts_senate, order = c(1, 0, 0), xdf = covar)
tab.ar1 <- xtable(ar1)

#AR(2)
ar2 <- armaFit(ts_senate, order = c(2, 0, 0), xdf = covar)
tab.ar2 <- xtable(ar2)

#MA(1)
ma1 <- armaFit(ts_senate, order = c(0, 0, 1), xdf = covar)
tab.ma1 <- xtable(ma1)

#ARMA(1,1)
arma11 <- armaFit(ts_senate, order = c(1, 0, 1), xdf = covar)
```

```

tab.arma11 <- xtable(arma11)

#bind the sum stat rows together
sums <- rbind(ar0, ar1, ar2, ma1, arma11)

print(xtable(sums), booktabs = T, include.rownames = FALSE,
      sanitize.colnames.function=function(x){x})

```

Order	AIC	RMSE	$T$	$\phi_1$	$\phi_2$	$\psi_1$	$S\phi_1$	PartisanMidterm	PartisanUnem	Coattails	Pre1994
(0,0,0)	183.044	5.132	28					1.792 (1.402)	0.358 (0.635)	3.403 (1.954)	8.503 (2.091)
(1,0,0)	172.3	4.039	28	0.696 (0.145)				-0.341 (0.945)	-0.268 (0.459)	1.892 (1.392)	8.151 (3.507)
(2,0,0)	168.047	3.563	28	1.087 (0.16)	-0.565 (0.174)			-0.806 (0.635)	-0.884 (0.367)	0.103 (1.091)	6.611 (2.581)
(0,0,1)	170.826	3.749	28			1 (0.122)		-0.464 (0.767)	-1.224 (0.339)	1.784 (1.467)	7.377 (2.204)
(1,0,1)	168.501	3.413	28	0.44 (0.187)		1 (0.148)		-0.72 (0.576)	-1.35 (0.24)	0.807 (1.134)	9.131 (2.171)

### Part c

```

ar1_sar1 <- armaFit(ts_senate,
                    order = c(1, 0, 0),
                    seasonal.order = c(1, 0, 0),
                    seasonal.period = 3,
                    xdf = covar)
ar1_sar1[1,1] <- "(1, 0, 0)(1, 0, 0)"
sums <- rbind(ar0, ar1, ar2, ma1, arma11, ar1_sar1)

```

Order	AIC	RMSE	$T$	$\phi_1$	$\phi_2$	$\psi_1$	$S\phi_1$	PartisanMidterm	PartisanUnem	Coattails	Pre 1994
(0,0,0)	183.044	5.132	28					1.792 (1.402)	0.358 (0.635)	3.403 (1.954)	8.503 (2.091)
(1,0,0)	172.3	4.039	28	0.696 (0.145)				-0.341 (0.945)	-0.268 (0.459)	1.892 (1.392)	8.151 (3.507)
(2,0,0)	168.047	3.563	28	1.087 (0.16)	-0.565 (0.174)			-0.806 (0.635)	-0.884 (0.367)	0.103 (1.091)	6.611 (2.581)
(0,0,1)	170.826	3.749	28			1 (0.122)		-0.464 (0.767)	-1.224 (0.339)	1.784 (1.467)	7.377 (2.204)
(1,0,1)	168.501	3.413	28	0.44 (0.187)		1 (0.148)		-0.72 (0.576)	-1.35 (0.24)	0.807 (1.134)	9.131 (2.171)
(1, 0, 0)(1, 0, 0)	169.047	3.622	28	0.818 (0.12)			-0.51 (0.185)	-0.494 (0.641)	-0.301 (0.388)	0.728 (1.029)	6.887 (3.045)