

Least Squares Objective

$$\hat{y} = w x$$

linear regression

- Instead of “exact y_i ”, we evaluate the size of error in prediction
- Classic way is setting slope ‘ w ’ to minimize the sum of squared errors:

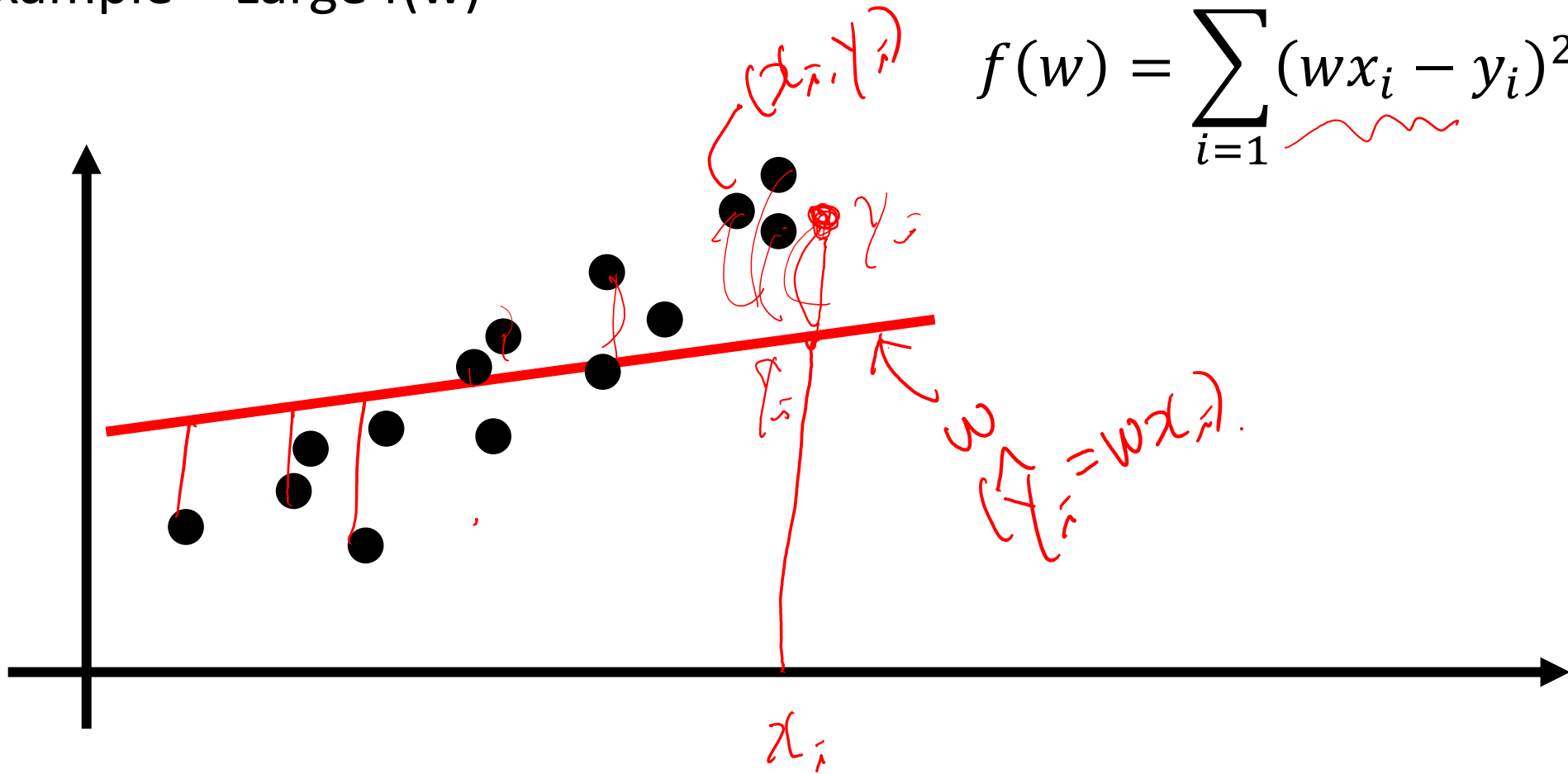
$$\arg \min_w f(w) = \sum_{i=1}^n (\hat{y}_i - y_i)^2$$

- A probabilistic interpretation is coming later in this course!
- But usually, it is done because it is easy to minimize.

Least Squares Objective

- Example – Large $f(w)$

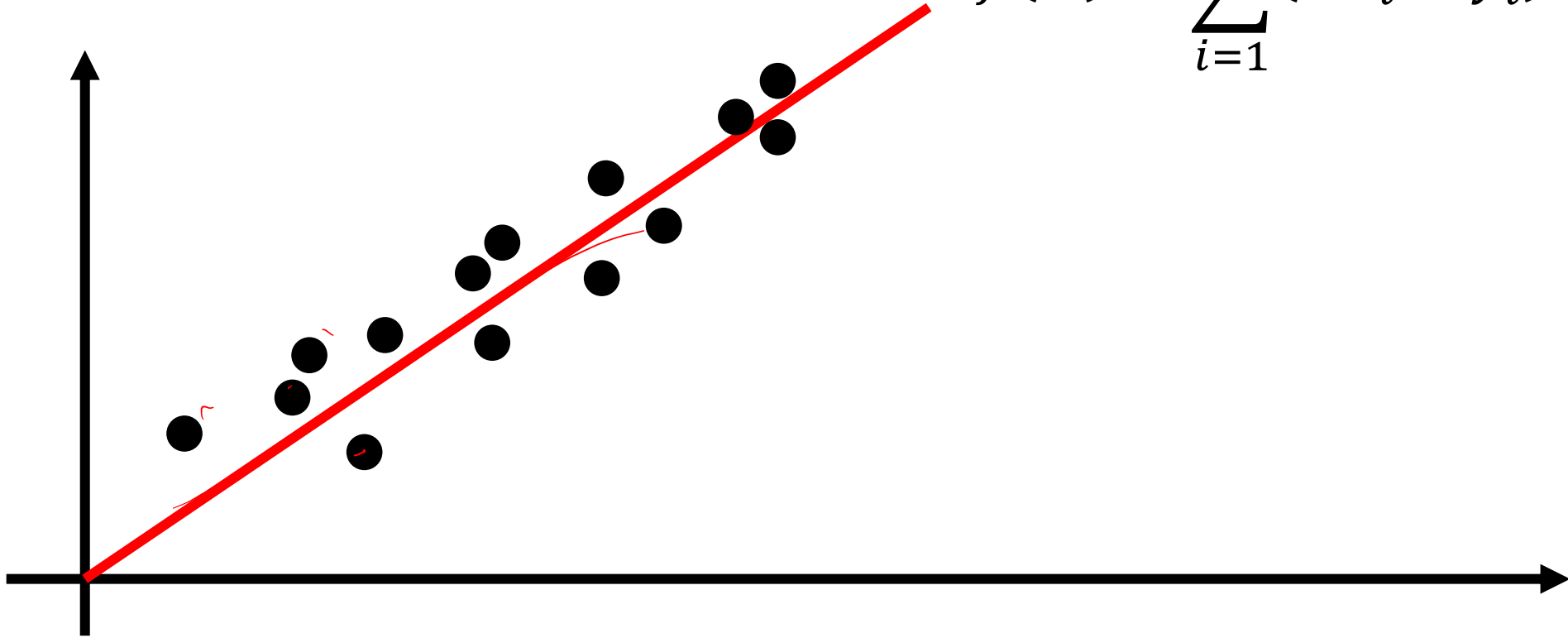
$$f(w) = \sum_{i=1}^n (wx_i - y_i)^2$$



Least Squares Objective

- Example – Small $f(w)$

$$f(w) = \sum_{i=1}^n (wx_i - y_i)^2$$



Finding Least Squares Solution

- Not change the solution!
 - Multiply 'f' by any positive constant
 - Add some constants to 'f'
- Finding 'w' that minimizes sum of squared errors:

$$\underline{f'(w)} = \underline{C_1} \sum_{i=1}^n (\underline{w}x_i - y_i)^2 + \underline{C_2}$$

$$\begin{aligned}\underline{f(w)} &= \frac{1}{2} \sum_{i=1}^n (wx_i - y_i)^2 = \frac{1}{2} \sum_{i=1}^n [w^2 x_i^2 - 2wx_i y_i + y_i^2] \\ &= \frac{w^2}{2} \sum_{i=1}^n x_i^2 - w \sum_{i=1}^n x_i y_i + \frac{1}{2} \sum_{i=1}^n y_i^2 = \frac{w^2}{2} \underline{a} - w \underline{b} + \underline{c}\end{aligned}$$

$$\underline{f'(w)} = wa - b = \underline{0}$$

$$w = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2}$$

Multiple Dimension Linear Function

- A simple way is with a d-dimensional linear model

- $\hat{y}_i = w_1 x_{i1} + w_2 x_{i2} + w_3 x_{i3} + \dots + w_d x_{id}$

- In words, our model is that the output is a weighted sum of the inputs

- We can re-write this in summation notation:

- $\hat{y}_i = \sum_{j=1}^d w_j x_{ij}$

- We can also re-write this in vector notation: (inner product)

- $\hat{y}_i = \mathbf{w}^T \mathbf{x}_i$

- In this course, a vector is a column vector

Least Squares in d-Dimensions

- The linear least squares model in d-dimensions minimizes:

$$f(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^n (\mathbf{w}^T \mathbf{x}_i - y_i)^2$$

- Least Squares Partial Derivatives for 1 sample

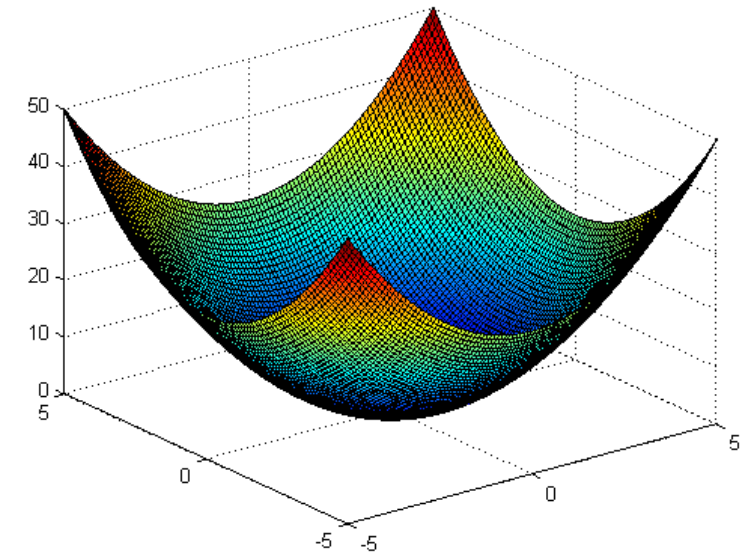
$$f(w_1, w_2, \dots, w_d) = \frac{1}{2} \left(\sum_{j=1}^d w_j x_{ij} \right)^2 - \left(\sum_{j=1}^d w_j x_{ij} \right) y_i + \frac{1}{2} y_i^2$$

$$\frac{\partial}{\partial w_k} f(w_1, w_2, \dots, w_d) = \left(\sum_{j=1}^d w_j x_{ij} \right) x_{ik} - y_i x_{ik} = (\mathbf{w}^T \mathbf{x}_i - y_i) x_{ik} = 0.$$

Least Squares in d-Dimensions

- Least Squares Partial Derivatives for all samples

$$\frac{\partial}{\partial w_k} f(w_1, w_2, \dots, w_d) = \sum_{i=1}^n (\mathbf{w}^T \mathbf{x}_i - y_i) x_{ik}$$



- Unfortunately, the partial derivative for w_j depends on all $\{w_1, w_2, \dots, w_d\}$
 - Thus, we can't just set equal to 0 and solve for w_j
 - **We need to find 'w' where the gradient vector equals the zero vector!**

$$\nabla f(w_1, w_2, \dots, w_d) = \left[\frac{\partial f}{\partial w_1}, \frac{\partial f}{\partial w_2}, \frac{\partial f}{\partial w_3}, \dots, \frac{\partial f}{\partial w_d} \right]^T = \mathbf{0}$$

Linear and Quadratic Gradients

$$A = A^T.$$

$$f(w) = aw^2 \Rightarrow \frac{\partial}{\partial w} f(w) = 2aw$$

$$g(w) = bw \Rightarrow \frac{\partial}{\partial w} g(w) = b$$

$$h(w) = c \Rightarrow \frac{\partial}{\partial w} h(w) = 0$$

$$f(\mathbf{w}) = \mathbf{w}^T \mathbf{A} \mathbf{w} \Rightarrow \nabla f(\mathbf{w}) = \mathbf{A} \mathbf{w}$$

If \mathbf{A} is symmetric

$$g(\mathbf{w}) = \mathbf{w}^T \mathbf{b} \Rightarrow \nabla g(\mathbf{w}) = \mathbf{b}$$

$$h(\mathbf{w}) = \mathbf{c} \Rightarrow \nabla h(\mathbf{w}) = \mathbf{0}$$

Linear and Quadratic Gradients

- We can re-write the d-dimensional quadratic:

$$\bullet f(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^n (\mathbf{w}^T \mathbf{x}_i - y_i)^2 = \frac{1}{2} \|\mathbf{X}\mathbf{w} - \mathbf{y}\|^2 = \frac{1}{2} \mathbf{w}^T \mathbf{X}^T \mathbf{X} \mathbf{w} - \mathbf{w}^T \mathbf{X}^T \mathbf{y} + \frac{1}{2} \mathbf{y}^T \mathbf{y}$$

$$\bullet f(\mathbf{w}) = \frac{1}{2} \mathbf{w}^T \mathbf{A} \mathbf{w} + \mathbf{w}^T \mathbf{b} + c$$

- Thus, the gradient is given by:

$$\bullet \nabla f(\mathbf{w}) = \mathbf{A} \mathbf{w} - \mathbf{b} = \mathbf{X}^T \mathbf{X} \mathbf{w} - \mathbf{X}^T \mathbf{y}$$

$= 0$

- Normal equations:

$$\bullet \mathbf{X}^T \mathbf{X} \mathbf{w} = \mathbf{X}^T \mathbf{y} \quad (\text{in the form of } \mathbf{A} \mathbf{x} = \mathbf{b})$$

$$\bullet \text{When } \mathbf{X}^T \mathbf{X} \text{ is invertible, } \mathbf{w} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

$$(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

↳ pseudo-inverse.

$$\mathbf{A} = \mathbf{X}^T \mathbf{X}, \quad \mathbf{A}^T = (\mathbf{X}^T \mathbf{X})^T = \mathbf{X}^T \mathbf{X}$$

$$\frac{1}{2} (\mathbf{X}\mathbf{w} - \mathbf{y})^T (\mathbf{X}\mathbf{w} - \mathbf{y})$$

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_n \end{bmatrix} \begin{bmatrix} w_1 \\ w_2 \\ \vdots \\ w_d \end{bmatrix} = \begin{bmatrix} \hat{y}_1 \\ \hat{y}_2 \\ \vdots \\ \hat{y}_n \end{bmatrix} - \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

Inverse Matrix & Pseudo-inverse Matrix

- For $m \times n$ system A ,
 - $Ax = b$
 - Let's find x that minimizes the energy of $\|Ax - b\|^2$
- The derivative of $\|Ax - b\|^2$ becomes
 - $Ax - b = 0$
 - $Ax = b$
 - But, we can't estimate the inverse of A because A is not square!

Handwritten red notes:

$$w = (X^T X)^{-1} X^T y$$

Dimensions: 100×100 (under $X^T X$)

$X \in \mathbb{R}^{n \times d}$

Inverse Matrix & Pseudo-inverse Matrix

- The derivative of $\|Ax - b\|^2$ becomes
 - $Ax - b = 0$
 - $Ax = b$
 - But, we can't estimate the inverse of A because A is not square!
- Then, let's make it square matrix
 - $A^T Ax = A^T b$
 - When the columns of A are linearly independent, $A^T A$ is invertible. Thus,
 - $x = (A^T A)^{-1} A^T b \equiv A^+ b$