

Quiz 2.

Name: Yifu He

ID: 190003956.

Score.

Q1. For any sequence $\{X_n\}$, the convergence of S_n/n to 0 in probability would lead to that X_n/n in probability as well.

Proof: $S_n = X_1 + X_2 + \dots + X_n$
 $= \sum_{i=1}^n X_i.$

Use the information of $\frac{S_n}{n}$ converge to 0 in p.r.

That's to say, iff. $\forall \delta > 0,$

$$\lim_{n \rightarrow \infty} P\left[\left|\frac{S_n}{n} - 0\right| \leq \delta\right] = 1 \quad (1)$$

$$\frac{X_n}{n} = \frac{S_n - S_{n-1}}{n} < \left|\frac{S_n - S_{n-1}}{n}\right| < \left|\frac{S_n}{n}\right| + \left|\frac{S_{n-1}}{n}\right| \quad (2)$$

We need to prove, $\frac{X_n}{n}$ converge to 0 in p.r. too.

Then for $\forall \varepsilon > 0,$

we need to prove $\lim_{n \rightarrow \infty} P\left[\left|\frac{X_n}{n} - 0\right| \leq \varepsilon\right] = 1,$

Use the information of (1). for any ε , we have $\varepsilon = 2\delta$.

$$\text{thus, } \lim_{n \rightarrow \infty} P\left[\underbrace{\left|\frac{X_n}{n}\right|}_{\leq} < \underbrace{\left|\frac{S_n}{n}\right|}_{\leq} + \underbrace{\left|\frac{S_{n-1}}{n}\right|}_{\leq} < \underbrace{2\delta}_{\varepsilon}\right] = 1$$

proof finished.

Q2. Let $\{X_n, n \geq 1\}$ be a sequence of i.i.v. $S_n = X_1 + X_2 + \dots + X_n$, and for any t , define $N(t)$ (2)
 such that $S_{N(t)} \leq t \leq S_{N(t)+1}$. Show that:

$$E[N(t)] = \sum_{n=1}^{\infty} P[S_n \leq t]$$

Proof: $S_n = X_1 + X_2 + \dots + X_n$
 $= \sum_{i=1}^n X_i$
 then $S_{N(t)} = \sum_{i=1}^{N(t)} X_i$

the relationship between $P[N(t) \geq n]$, $P[S_n \leq t]$ is that

$$P[N(t) \geq n] = P[S_n \leq t]$$

$$\begin{aligned} E[N(t)] &= 1 \cdot P[N(t) = 1] + 2 \cdot P[N(t) = 2] + 3 \cdot P[N(t) = 3] + \dots + \infty \cdot P[N(t) = \infty] \\ &= P[N(t) \geq 1] + P[N(t) \geq 2] + \dots + P[N(t) \geq \infty] \\ &= \sum_{n=1}^{\infty} P[N(t) \geq n] \\ &= \sum_{n=1}^{\infty} P[S_n \leq t] \end{aligned}$$

Q3. Find the n th iterated convolution of an exponential distribution

Proof: the density function of exponential ~~function~~ distribution with λ is

$$f(x; \lambda) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

when $n=2$,

$$\begin{aligned} f_2(x; \lambda) &= \int_{-\infty}^{\infty} f_1(x-y) \cdot f_1(y) dy \\ &= \int_{-\infty}^{\infty} \lambda \cdot e^{-\lambda(x-y)} \lambda e^{-\lambda y} dy \\ &= \int_{-\infty}^{\infty} \lambda^2 e^{-\lambda x + \lambda y} \cdot e^{-\lambda y} dy \\ &= \int_{-\infty}^{\infty} \lambda^2 e^{-\lambda x} dy \\ &= \cancel{\lambda^2 e^{-\lambda x} \int_{-\infty}^{\infty} dy} = \lambda^2 e^{-\lambda x} \int_0^x dy = \lambda^2 \cdot x e^{-\lambda x} = \frac{\lambda^2}{2-1!} \cdot x^{2-1} \cdot e^{-\lambda x} \quad (1) \end{aligned}$$

when $n=k$, $k \geq 2$ Suppose $f_k(x; \lambda) = \frac{1}{(k-1)!} \lambda^k \cdot x^{k-1} \cdot e^{-\lambda x}$ (2)

when $n=k+1$,

$$\begin{aligned} f_{k+1}(x; \lambda) &= \int_0^x f_k(x-y) \cdot f_k(y) dy \quad (3) \\ &= \int_0^x \lambda e^{-\lambda(x-y)} \cdot \frac{1}{(k-1)!} \lambda^k \cdot y^{k-1} \cdot e^{-\lambda y} dy \\ &= \cancel{\lambda^{k+1} \int_0^x} \frac{\lambda^{k+1} \cdot e^{-\lambda x}}{(k-1)!} \int_0^x y^{k-1} e^{\lambda y - \lambda y} dy = \frac{\lambda^{k+1}}{(k-1)!} \cdot e^{-\lambda x} \cdot \int_0^x y^{k-1} dy \\ &= \frac{\lambda^{k+1}}{(k-1)!} \cdot e^{-\lambda x} \cdot \left. \frac{y^k}{k} \right|_0^x \\ &= \frac{\lambda^{k+1}}{k!} \cdot e^{-\lambda x} \cdot y^k \end{aligned}$$

according to (1) (2) (3). using Mathematical Induction

We conclude $f_n(x; \lambda) = \frac{1}{(n-1)!} \lambda^n \cdot x^{n-1} \cdot e^{-\lambda x}$