Poisson Processes and Jump Diffusion

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Exponential Random Variables

Let τ be a random variable with pdf

$$f_{\tau}(t) = egin{cases} \lambda e^{-\lambda t}, t \geq 0 \\ 0, t < 0 \end{cases}$$

We have $\mathbb{E}[\tau] = \frac{1}{\lambda}$ and cdf

$$F_{\tau}(t) = \mathbb{P}(\tau \leq t) = 1 - e^{-\lambda t}, t \geq 0$$

This is a memoryless random variable, or

$$\mathbb{P}(au > t) = \mathbb{P}(au > t + s | au > s)$$



Construct a sequence of i.i.d. exponential random variables $\{\tau_i\}$ with parameter λ . The "jump" times can then be defined as

$$S_n = \sum_{i=1}^n \tau_i$$

We have a resulting Poisson process that we call N(t) that counts the number of jumps before a given time

$$N(t) = \left\{ egin{array}{ll} 0, 0 \leq t < S_1 \ 1, S_1 \leq t < S_2 \ dots \ n, S_n \leq t < S_{n+1} \ dots \end{array}
ight.$$

Lemma 11.2.2 The Poisson process N(t) with intensity λ has the distribution

$$\mathbb{P}(N(t) = k) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}, k = 0, 1, \dots$$

[1]

Theorem 11.2.3: Let N(t) be a Poisson process with intensity $\lambda > 0$, and let $0 = t_0 < t_1 < \cdots < t_n$ be given. Then the increments

$$N(t_1) - N(t_0), N(t_2) - N(t_1), \dots, N(t_n) - N(t_{n-1})$$

are stationary and independent, and

$$\mathbb{P}(N(t_{j+1}) - N(t_j) = k) = \frac{\lambda^k (t_{j+1} - t_j)^k}{k!} e^{-\lambda(t_{j+1} - t_j)}, k = 0, 1, \dots$$

Poisson Increments

As a result of Theorem 11.2.3, we have the following results about our Poisson increments for $0 \le s < t$

$$\mathbb{P}(N(t) - N(s) = k) = \frac{\lambda^k (t - s)^k}{k!} e^{-\lambda(t - s)}, k = 0, 1, \dots$$

$$\mathbb{E}[N(t) - N(s)] = \lambda(t - s)$$

$$\mathbb{V}[N(t) - N(s)] = \lambda(t - s)$$

Martingale Property

Theorem 11.2.4: Let N(t) be a Poisson process with intensity λ . We define the compensated Poisson process

$$M(t) = N(t) - \lambda t$$

Then M(t) is a martingale.[1] Proof: Let $0 \le s < t$ be given.

$$\mathbb{E}[M(t)|\mathcal{F}(s)] = \mathbb{E}[M(t) - M(s)|\mathcal{F}(s)] + \mathbb{E}[M(s)|\mathcal{F}(s)]$$

$$= \mathbb{E}[N(t) - N(s) - \lambda(t-s)|\mathcal{F}(s)] + M(s)$$

$$= \mathbb{E}[N(t) - N(s)] - \lambda(t-s) + M(s)$$

$$= M(s)$$

Compound Poisson Process

Let N(t) be a Poisson process with intensity λ , and let $\{Y_i\}$ be a sequence of i.i.d. random variables with $\mathbb{E}[Y_i] = \beta$ that are independent of both each other and of the Poisson process. We define the **compound Poisson process** as

$$Q(t) = \sum_{i=1}^{N(t)} Y_i, t \ge 0$$

Using this definition and previous results we have

$$\mathbb{E}[Q(t)] = \beta \lambda t$$

Theorem 11.3.1: Let Q(t) be the compound Poisson process. Then the compensated Poisson process

$$Q(t) - \beta \lambda t$$

is a martingale.[1]

Theorem 11.3.2: Let Q(t) be a compound Poisson process and let $0 = t_0 < t_1 < \cdots < t_n$ be given. The increments

$$Q(t_1) - Q(t_0), Q(t_2) - Q(t_1), \ldots, Q(t_n) - Q(t_{n-1})$$

are independent and stationary. In particular, the distribution of $Q(t_j) - Q(t_{j-1})$ is the same as the distribution of $Q(t_j - t_{j-1})[1]$



Definition 11.4.1: Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space, and let $\mathcal{F}(t), t > 0$, be a filtration on this space. We say that a Brownian motion W(t) is a Brownian motion relative to this filtration if W(t) is $\mathcal{F}(t)$ -measurable for every t and for every u > t the increment W(u) - W(t) is independent of $\mathcal{F}(t)$. Similarly, we say that a Poisson process *N* is a Poisson process relative to this filtration if N(t) is $\mathcal{F}(t)$ -measurable for every tand for every u > t the increment N(u) - N(t) is independent of $\mathcal{F}(t)$. Finally, we say that a compound Poisson process Q is a compound Poisson process relative to this filtration if Q(t) is $\mathcal{F}(t)$ -measurable for every t and for every u > t the increment Q(u) - Q(t) is independent of $\mathcal{F}(t)$.[1]

We need to define the stochastic integral

$$\int_0^t \Phi(s) dX(s)$$

where X can have jumps. Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space on which $\mathcal{F}(t), t \geq 0$, is a filtration. All processes are adapted to this filtration from the previous definition and will be right-continuous and can be expressed as

$$X(t) = X(0) + I(t) + R(t) + J(t)$$

A pure jump process J(t) is an adapted, right-continuous process with J(0)=0 and $J(t)=\lim_{s\downarrow t}J(s)$

Definition 11.4.2: A process X(t) of the form in the previous slide, with Itô integral part I(t), Riemann integral part R(t), and pure jump part I(t) will be called a jump process. The continuous part of this process is $X^c(t) = X(0) + I(t) + R(t)$.[1]

Definition 11.4.3: Let X(t) be a jump process of the previous form and let $\Phi(s)$ be an adapted process. The stochastic integral of Φ with respect to X is defined to be

$$\int_0^t \Phi(s) dX(s) = \int_0^t \Phi(s) \Gamma(s) dW(s) + \int_0^t \Phi(s) \Theta(s) ds + \sum_{0 < s \le t} \Phi(s) \Delta J(s)$$

or in differential notation,

$$\Phi(t)dX(t) = \Phi(t)dI(t) + \Phi(t)dR(t) + \Phi(t)dJ(t)$$

= $\Phi(t)dX^{c}(t) + \Phi(t)dJ(t)$





Theorem 11.4.5: Assume that the jump process X(s) is a martingale, the integrand $\Phi(s)$ is left-continuous and adapted, and

$$\mathbb{E}\left[\int_0^t \mathsf{\Gamma}^2(s) \mathsf{\Phi}^2(s) ds\right] < \infty, \forall t \geq 0$$

Then the stochastic integral $\int_0^t \Phi(s) dX(s)$ is also a martingale.

Quadratic Variation

For a partition Π , we denote, as usual, the sample quadratic variation of a process as:

$$Q_{\Pi}(X) = \sum_{j=0}^{n-1} (X(t_{j+1}) - X(t_j))^2$$

And then to get the actual quadratic variation of the process X(t), we let $||\Pi|| \to 0$ to get [X,X](t). For the processes $X_1(t)$ and $X_2(t)$ we define the sample cross variation as

$$C_{\Pi}(X) = \sum_{j=0}^{n-1} (X_1(t_{j+1}) - X_1(t_j))(X_2(t_{j+1}) - X_2(t_j))$$



Theorem 11.4.7: Let $X_1(t) = X_1(0) + I_1(t) + R_1(t) + J_1(t)$ be a jump process, where $I_1(t) = \int_0^t \Gamma_1(s) dW(s)$, $R_1(t) = \int_0^t \Theta(s) ds$, and $J_1(t)$ is a right continuous pure jump process. Then $X_1^c(t) = X_1(0) + I_1(t) + R_1(t)$ and

$$\begin{aligned} [X_1, X_1](T) &= [X_1^c, X_1^c](T) + [J_1, J_1](T) \\ &= \int_0^T \Gamma_1^2(s) ds + \sum_{0 < s \le T} (\Delta J_1(s))^2 \end{aligned}$$

Define $X_2(t)$ similarly to $X_1(t)$.

$$\begin{aligned} [X_1, X_2](T) &= [X_1^c, X_2^c](T) + [J_1, J_2](T) \\ &= \int_0^T \Gamma_1(s) \Gamma_2(s) ds + \sum_{0 < s \le T} \Delta J_1(s) \Delta J_2(s) \end{aligned}$$



Ito Formula for One Jump Process

Theorem 11.5.1: Let X(t) be a jump process and f(x) a function for which f'(x) and f''(x) are defined and continuous. Then

$$f(X(t)) = f(0, X(0)) + \int_0^t f'(X(s))dX^c(s) + \frac{1}{2} \int_0^t f''(X(s))dX^c(s)dX^c(s) + \sum_{0 < s \le T} [f(X(s)) - f(X(s-))]$$

Corollary 11.5.3: Let W(t) be a Brownian motion and let N(t) be a Poisson process with intensity $\lambda > 0$, both defined on the same probability space $(\Omega, \mathcal{F}, \mathbb{P})$ and relative to same filtration $\mathcal{F}(t)$, $t \geq 0$. Then the processes W(t) and N(t) are independent.

Two Dimensional Ito Formula

Theorem 11.5.4: Let $X_1(t)$ and $X_2(t)$ be jump processes, and let $f(t, x_1, x_2)$ be a function whose first and second partial derivatives appearing in the following formula are defined and continuous. Then

$$\begin{split} f(t,X_1(t),X_2(t)) &= f(0,X_1(0),X_2(0)) + \int_0^t f_t(s,X_1(s),X_2(s))ds \\ &+ \int_0^t f_{X_1}(s,X_1(s),X_2(s))dX_1^c(s) + \int_0^t f_{X_2}(s,X_1(s),X_2(s))dX_2^c(s) \\ &+ \frac{1}{2} \int_0^t f_{X_1,X_1}(s,X_1(s),X_2(s))dX_1^c(s)dX_1^c(s) \\ &+ \int_0^t f_{X_1,X_2}(s,X_1(s),X_2(s))dX_1^c(s)dX_2^c(s) \\ &+ \frac{1}{2} \int_0^t f_{X_2,X_2}(s,X_1(s),X_2(s))dX_2^c(s)dX_2^c(s) \\ &+ \sum_{0 < s \le t} [f(s,X_1(s),X_2(s)) - f(s-,X_1(s-),X_2(s-))] \end{split}$$

Ito Product Rule for Jump Processes

Corollary 11.5.5: Let $X_1(t)$ and $X_2(t)$ be jump processes, Then

$$\begin{split} X_1(t)X_2(t) &= X_1(0)X_2(0) + \int_0^t X_2(s)dX_1^c(s) + \int_0^t X_1(s)dX_2^c(s) \\ &+ [X_1^c, X_2^c](t) + \sum_{0 < s \le t} [X_1(s)X_2(s) - X_1(s-)X_2(s-)] \\ &= X_1(0)X_2(0) + \int_0^t X_2(s-)dX_1(s) + \int_0^t X_1(s-)dX_2(s) \\ &+ [X_1, X_2](t) \end{split}$$

[1]

Corollary 11.5.6: Let X(t) be a jump process. The **Doleans-Dade exponential** of X is defined to be the process

$$Z^{X}(t) = e^{X^{c}(t) - \frac{1}{2}[X^{c}, X^{c}](t)} \prod_{0 < s \le t} (1 + \Delta X(s))$$

This process is the solution to the stochastic differential equation

$$dZ^X(t) = Z^X(t-)dX(t)$$

with initial condition $Z^X(0) = 1$, which in integral form is

$$Z^X(t) = 1 + \int_0^t Z^X(s-)dX(s)$$

[1]



Let $\tilde{\lambda}$ be a positive number and define:

$$Z(t) = e^{(\lambda - \tilde{\lambda})t} \left(\frac{\tilde{\lambda}}{\lambda}\right)^{N(t)}$$

Fix a time T > 0, and we will use Z(T) to change to measure $\tilde{\mathbb{P}}$ under which N(t) has intensity $\tilde{\lambda}$ instead of λ .

Lemma 11.6.1: The process Z(t) satisfies:

$$dZ(t) = rac{ ilde{\lambda} - \lambda}{\lambda} Z(t_{-}) dM(t)$$

In particular, Z(t) is a martingale under \mathbb{P} and $\mathbb{E}[Z(t)] = 1$ for all t.[1]

Theorem 11.6.2 (Change of Poisson Intensity): Under the probability measure $\tilde{\mathbb{P}}$, the process N(t), $0 \le t \le T$, is Poisson with intensity $\tilde{\lambda}$.[1]

Define:

$$Z_m(t) = e^{(\lambda_m - \tilde{\lambda_m})t} \left(rac{ ilde{\lambda_m}}{\lambda_m}
ight)^{N_m(t)}$$
 $Z(t) = \prod_{m=1}^M Z_m(t)$

Theorem 11.6.4: The process Z(t) is a martingale. In particular, $\mathbb{E}[Z(t)] = 1$ for all t.

Theorem 11.6.5: Under $\widetilde{\mathbb{P}}$, Q(t) is a compound Poisson process with intensity $\widetilde{\lambda} = \sum_{m=1}^{M} \widetilde{\lambda}_m$, and Y_1, Y_2, \ldots are independent, identically distributed random variables with

$$\widetilde{\mathbb{P}}\{Y_i=y_m\}=\widetilde{p}(y_m)=\frac{\widetilde{\lambda}_m}{\widetilde{\lambda}}$$

[1]

Lemma 11.6.6: For Z(t) given by

$$Z(t) = e^{(\lambda - \tilde{\lambda})t} \prod_{i=1}^{N(t)} \frac{\tilde{\lambda}\tilde{f}(Y_i)}{\lambda f(Y_i)}$$

this process is a martingale. In particular, $\mathbb{E}[Z(t)] = 1$ for all $t \ge 0$.

Define:

$$Z_1(t) = e^{-\int_0^t \theta(u)dW(u) - \frac{1}{2}\int_0^t \theta^2(u)du}$$

$$Z_2(t) = e^{(\lambda - \tilde{\lambda})t} \prod_{i=1}^{N(t)} \frac{\tilde{\lambda}\tilde{f}(Y_i)}{\lambda f(Y_i)}$$

$$Z(t) = Z_1(t)Z_2(t)$$

Lemma 11.6.8: The process Z(t) is a martingale. In particular, $\mathbb{E}[Z(t)] = 1$ for all $t \ge 0$.[1]

Theorem 11.6.9: Under the probability measure $\widetilde{\mathbb{P}}$, the process

$$\widetilde{W}(t) = W(t) + \int_0^t \theta(s) ds$$

is a Brownian motion, Q(t) is a compound Poisson process with intensity $\tilde{\lambda}$ and independent, identically distributed jump sizes having density $\tilde{f}(y)$, and the processes $\widetilde{W}(t)$ and Q(t) are independent.[1]

[1] S.E. Shreve. *Stochastic Calculus for Finance II: Continuous-Time Models*. Number v. 11.