Weekly Homework 4 570

Weiping Zhang Financial Engineering

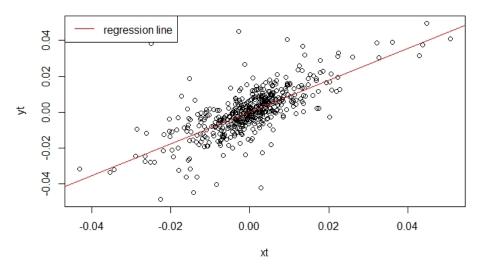
November 26,2018

Problem 1.

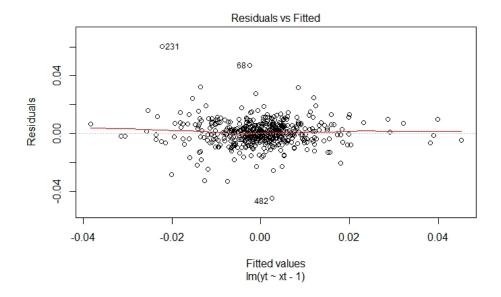
Solution.

Firstly, we estimate the co-intergrating realtion. We use y_t (the log return of Chevron) and x_t to do a linear regression. We first do the normal resgression y_t x_t , the result shows that the term of the intercept is not significant. Therefore, we fit the model agagin without the intercept term, the coefficient of x_t is significant. And the α here is 0.8882734, c here is 0.0004814.

scatter plot for xt and yt



Following the procedure, we the need to test for stationarity of the residual z_t useing ADF unit root test. The result gave that the p-value is smaller than 0.01, therefore, under the cofidence level equals 0.01, we can reject the null hypothesis, which is that the series is non-stationary. So we can see that the residual is stationary in this case.



Finally,we calculate Δ and follow the trading rules. When $|y_t - \alpha x_t - c + \Delta|$ is less than 0.001, we set the signal to 1. When $|y_t - \alpha x_t - c - \Delta|$ is less than 0.001, we set the signal to -1. Modify a new dataframe,we can get the result of long and short signals along with dates.

```
Date order act 8 2015-12-10 8 short 9 2015-12-11 9 long 13 2015-12-17 13 long 62 2016-03-01 62 short 167 2016-07-29 167 short 363 2017-05-10 363 short 368 2017-05-17 368 long
```

Based on the signals, we can calculate the pnl. We assume that we long 1 porfolio is that we short 1 shares B and long ratio*1 shares A.We short 1 porfolio is that we buy 1 shares B and short ratio*1 shares A. Based on this rule, the pnl here is 12.23

R code attached:

```
install.packages("tseries")
library(tseries)
#Problem 1
xom<-read.csv((file.choose()))</pre>
head(xom)
cvx<-read.csv((file.choose()))</pre>
head(cvx)
xom_close<-xom$Close
cvx_close<-cvx$Close
n1<-length(xom_close)</pre>
n2<-length(cvx_close)</pre>
xt<-log(xom_close[-1]/xom_close[-n1])</pre>
yt <-log(cvx_close[-1]/cvx_close[-n2])
plot(xt,yt,main="scatter plot for xt and yt")
fit<-lm(yt~xt)</pre>
summary(fit)
coef<-fit$coefficients</pre>
alpha <- as. vector (coef) [2]
alpha
const<-as.vector(coef)[1]</pre>
const
abline(fit,col="red")
legend("topleft", lty=c(1), col=c("red"),
        legend=c("regression line"))
zt<-fit$residuals
plot(fit, which=1)
#the non-linear trend is not obvious
adf.test(zt)
```

```
#p-value less than 0.01, rejected the null hypothesis. Under alpha=0.01, it's static
delta<-2*sd(zt)</pre>
date <-as.Date(xom$Date[-1])</pre>
yt_axt<-yt-alpha*xt
table <- as.data.frame(cbind(xt,yt,yt_axt))</pre>
table $Date = date
table<-table[,c('Date','xt','yt','yt_axt')]</pre>
table$signal=c(rep(0,nrow(table)))
table[abs((table$yt_axt+delta-const))<=0.001,]$signal<-1
table[abs((table$yt_axt-delta-const)) <= 0.001,] $signal <--1
order <-c(1:nrow(table))
table $ order = order
table$act=c(rep("act",nrow(table)))
table [table $signal == 1,] $act <- "long"
table[table$signal==-1,]$act<-"short"
trade \leftarrow subset(table[,c(-2,-3,-4,-5)],table = "act")
#calculate the pnl at the accuracy of 0.001
cal_xp<-xom_close[(trade$order+1)]</pre>
cal_cp<-cvx_close[(trade$order+1)]</pre>
#Assuming that we long 1 porfolio is that we short 1 shares B and long ratio shars
#we short 1 porfolio is that we buy 1 shares B and short ratio shares A
#calculate the pnl under the accuracy of 0.01
sig<-trade$act
sig
sig[sig=="short"]=1
sig[sig=="long"]=-1
sig<-as.numeric(sig)</pre>
sig_A<-sig*alpha
sig_A[length(sig_A)] = -2*alpha
sig_B < -sig_*(-1)
sig_B[length(sig_B)]=2
pnl_A<-sig_A%*%cal_xp
pnl_B<-sig_B%*%cal_cp
pnl_A
pnl_B
cal_xp
pnl<-pnl_A+pnl_B
pnl
```

Problem 2.

Solution.

The time series r_t model follows the model:

$$r_t = 0.01 + 0.2r_{t-2} + a_t$$

Take expectation on both sides:

$$Er_t = 0.01 + 0.2Er_{t-2} + a_t$$

$$\mu = 0.01 + 0.2\mu + 0$$

$$(1 - 0.2)\mu = 0.01(*)$$

$$\mu = 0.0125$$

Using (*), the model can be rewritten as:

$$r_t - \mu = 0.2 (r_{t-2} - \mu) + a_t \tag{1}$$

By repeated substitutions, the prior equation implies that:

$$r_t - \mu = a_t + 0.2r_{t-2} + 0.2^2r_{t-4} + \dots$$
$$= \sum_{i=0}^{\infty} 0.2^i a_{t-2i}$$

Therefore, we have $E[a_t(r_t - \mu)] = \sigma_a^2$, $E[a_t(r_{t-1} - \mu)] = 0$, $E[a_t(r_{t-2} - \mu)] = 0$. Multiplying (1) by $(r_{t-2} - \mu)$, $(r_t - \mu)$, $(r_{t-1} - \mu)$. Then take expectation on both sides, we have:

$$\gamma_2 = 0.2\gamma_0 \tag{2}$$

$$\gamma_0 = 0.2\gamma_2 + \sigma_a^2 \tag{3}$$

$$\gamma_1 = 0.2\gamma_1 \tag{4}$$

From (4), we know that $\gamma_1=0$. From (2) and (3), we can solve that

$$\gamma_0 = \frac{1}{48}$$

$$\gamma_2 = \frac{1}{240}$$

Dividing γ_1 and γ_2 by γ_0 , we then have the lag-1 and lag-2 correlations or rt:

$$\rho_1 = 0$$

$$\rho_2 = \frac{1}{5}$$

1-Step Ahead Forecast:

The point forecast of r_{t+1} given $F_t = r_t, r_t - 1, ...$ is the conditional expectation

$$\hat{r}_t(1) = E(r_{t+1}|\mathscr{F}_t)$$

= 0.01 + 0.2 r_{t-1}

Therefore, the 1-step forecast is 0.01+0.2*0.02=0.014 and the associated forecast error is:

$$e_t(1) = r_{t+1} - \hat{r}_t(1) = a_{t+1}$$

Consequently, the associated standard deviations of the 1-step forecast error is

$$\sqrt{Var[e_t(1)]} = \sqrt{Var(a_{t+1})} = \sigma_a = 0.14142$$

The point forecast of r_{t+2} given $F_t = r_t, rt-1, ...$ is the contional expectation

$$\hat{r}_t(2) = E\left(r_{t+2} \middle| \mathscr{F}_t\right)$$
$$= 0.01 + 0.2r_t$$

Therefore, the 2-step forecast is 0.01+0.2*(-0.01)=0.008 and the associated forecast error is:

$$e_t(2) = r_{t+2} - \hat{r}_t(2) = a_{t+2}$$

Consequently, the associated standard deviations of the 2-step forecast error is

$$\sqrt{Var[e_t(2)]} = \sqrt{Var(a_{t+2})} = \sigma_a = 0.14142$$