

# DEMOCRACY AND ETHNIC FAVORITISM: EVIDENCE FROM AFRICA<sup>\*</sup>

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## ABSTRACT

What is the effect of democracy on ethnic favoritism? I estimate the relationship between co-ethnicity and five outcomes of public good provision – education, infant health, wealth, access to clean drinking water and access to electricity – using data from twenty-one African countries. Following previous research, I use variation in co-ethnicity across ethnic groups and over time. I first estimate the relationship between these five outcomes and co-ethnicity in the full sample and then split the analysis between anocracies and democracies. I find mixed evidence of ethnic favoritism in the full sample. I find more evidence of co-ethnic targeting when comparing within democracies only and within anocracies only.

*Keywords:* Ethnic favoritism; democracy; dictatorships; institutions.

*JEL Classifications:* D3, O1, O2.

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## I. INTRODUCTION

Ethnic favoritism has been widespread in Africa, and it negatively affects growth. One example of such action was from Zimbabwe, where the White minority ruled the country after independence. During colonization and following independence in the short-lived Rhodesia, Whites enjoyed special privileges. They were also granted free agricultural farms. Policies that favored a specific ethnic group created an economic gap between White and Black Zimbabweans. Even though Whites accounted for 0.3% of the population, they owned 70% of the farming land ([Batha, 2000](#)). Ethnic diversity could lead to lower growth as a result of conflicts and rent-seeking by politically dominant ethnic groups. [La Porta et al. \(1999\)](#), [Alesina et al. \(1999\)](#), [Alesina et al. \(2003\)](#) and [Alesina and Zhuravskaya \(2011\)](#) show that ethnic fractionalization and ethnic segregation affects the quality of governance and creates conflicts over the provisions of public goods. Ethnic diversity also affect the productivity of labor ([Lazear, 1999](#); [Habyarimana et al., 2007](#); [Hjort, 2014](#)). [Francois et al. \(2015\)](#) show that African leaders share the benefits of heading an ethnically diverse country among ministers from different ethnic groups to maximize their time in power. Instances like these could explain the inequality among the ethnic lines and why Africa lags behind economically and in public goods provisions ([Easterly and Levine, 1997](#); [Alesina and Ferrara, 2005](#); [Miguel and Gugerty, 2005](#); [Glennerster et al., 2013](#)). [Golden and Min \(2013\)](#), in a review of ethnic favoritism empirical studies, show that the majority of papers study ethnic favoritism in one democratic country or focus on one outcome. Understanding the effect of political systems on the persistence of ethnic favoritism is essential in understanding Africa's economic development. I aim to answer the following question. Do democratic institutions affect the prevalence of ethnic favoritism?

This paper makes two contributions to the literature. First, it investigates the existence of ethnic favoritism on a few outcomes—schooling, infant survival, wealth, electrification and access to clean drinking water—that are essential for economic development. Second, it provides empirical evidence from a sample of many countries on the existence of differential

co-ethnic effects between countries that are more or less democratic.

What causes some countries to distribute resources evenly among their citizens while others take a different path that benefits a specific ethnic group? What stops a leader from favoring their ethnic kins? The literature of ethnic favoritism is rife with research on the existence of the phenomena (Friedman, 2018; Franck and Rainer, 2012; Hodler and Raschky, 2014b; Kramon and Posner, 2013). Friedman (2018) found that Kenyans that are of the same ethnicity as the leader disproportionately received antiretroviral drugs. Franck and Rainer (2012) studied the effect of ethnic favoritism in health and education in multiple African countries. They found that the leader's ethnic group is more likely to complete primary school and have better health outcomes. Kramon and Posner (2013) found that ethnic favoritism in education exists in Africa. Kasara (2007) finds that African leaders tax their ethnic group more because they can exert more control over them. Suggesting that ethnic groups do not benefit from having ethnic kin in power. Kudamatsu (2009) exploits an exogenous change in the ethnicity of the president of Guinea to study ethnic favoritism on infant mortality. They find that the ethnic group of the president did not enjoy an advantage in the reduction of infant mortality.<sup>1</sup> Distributive policies that favor one group over another are not exclusive to developing countries. Goss (1972) documented how the members of the principal military committees in the United States received unrepresentative constituency benefits. Ferejohn (1974) shows that influence in congress and pork barrels over rivers and harbors are distributed based on political consideration and not improving the public good.

Therefore, all of these papers looked at the existence of ethnic favoritism in several countries. However, they did not attempt to find a relationship between ethnic favoritism and democracy, which I aim to do in this paper.

Others tried to investigate the presence of ethnic favoritism during periods of democracy in a specific country (Burgess et al., 2015; Dionne and Horowitz, 2016; Harris and Posner, 2019; Kramon and Posner, 2016). Harris and Posner (2019) found that members of the

1. Kudamatsu (2012), they studied the effect of democratization on infant mortality. They find that democratization reduced infant mortality by 1.2 percentage points.

Kenyan parliament do not disproportionately favor their constituents. [Burgess et al. \(2015\)](#) also tried to find the effect of democracy on ethnic favoritism in Kenya using data on road building. They found that ethnic favoritism exists in Kenya, but it disappears during periods of democracy. [Amodio and Chiovelli \(2016\)](#) investigated ethnic favoritism toward the Zulu nation in South Africa when it was a new democracy. They found that when the Zulu-based Inkatha Freedom Party (IFP) is in the majority, Zulu individuals experience better labor market outcomes. [Dionne and Horowitz \(2016\)](#) also tried to investigate the existence of ethnic favoritism in agricultural subsidies in Malawi. They found that the incumbent majority party, at the time, Democratic Progressive Party headed by president Bingu wa Mutharika did not systematically target the ethnic kins of Mutharika and their co-partisans. The analysis of this group of papers was limited to one or two countries. Therefore, they did not have enough observations of transitions from democracy to dictatorship and *vice versa*. For example, in [Burgess et al. \(2015\)](#), Kenya changed its political system twice. Thus, the analysis was technically restricted to two observations over time. I try to circumvent this problem in this paper by taking into consideration 21 countries that transitioned into and out of democracy twenty times.

Moreover, using a continuous democracy index, Polity IV, I can observe changes in democratic institutions and run a heterogeneous effect analysis. For example, in 1991, Mali was transitioning from dictatorship to democracy. Using a dummy variable for democracy, I would observe that Mali became a democracy in 1992. However, by using Polity IV, I found that in 1990, Mali's *Polity* score was -7, which increased to 4 in 1992 then 5 in 1997.

Additionally, the literature considered the effects of institutions on economic growth. This relates to the topic I propose because a democratic system is contained in the group of all institutions. [Alsan \(2015\)](#) looked for a link between the prevalence of the Tse Tse fly in Africa and how it affected institution building. She found that the fly impeded a society's ability to settle down and build institutions. Therefore, it put these countries on a lower growth trajectory. [Acemoglu et al. \(2001\)](#) tried to measure institutions using the number of

European Colonial fatalities as an instrumental variable. The authors found that countries with a more hospitable environment toward European settlers had better institutions than those that were not. They also found that the European mortality rate was negatively correlated with economic growth. In other words, countries with a high European mortality rate had a worse economic performance than those that did not. My paper would add to this literature by studying the effect of democratic institutions on ethnic favoritism that holds Africa back from growing faster.

Ethnic favoritism has played a role in political economy theories. [Fearon \(1999\)](#) and [Caselli and Coleman \(2006\)](#) introduced theories in which ethnicity is a tool of exclusion and to enforce coalition membership. [Francois et al. \(2015\)](#) introduced a model of power-sharing in Africa. They show that African countries, mainly autocracies, appoint a cabinet in a manner that represents the ethnic diversity of the population. More precisely, they find that large ethnic groups are slightly underrepresented, minorities are overrepresented and the leader's ethnic group has a slight premium. The leader, in return, distributes the benefits of being in power among the ministers to stay in power. [Padró i Miquel \(2007\)](#) show that patronage, taxation and spending change with changes in the ethnic group in power, which contributes to bad governance, ethnic bias and wasteful policies in Africa.

The literature lacks an empirical framework to explain the differences in ethnic bias among countries. Moreover, there is an even more significant gap investigating the existence of a link between democratic institutions and favoritism across nations. In this paper, I examine ethnic favoritism in twenty-one African countries and estimate its persistence during periods of democracy and dictatorship.

Using data from the Demographics and Health Survey (DHS), I estimate the effect of ethnic favoritism and democracy on a few outcomes of interest such as education, wealth, infant survival, electrification and access to clean drinking water. Using a rich data set, I was able to control for ethnic group, time and age fixed effects.

There are several tools that a leader could use to benefit the educational, health, wealth

and infrastructure outcomes of their ethnic kins. Leaders could hire more teachers, improve the physical health of schools (Glewwe and Kremer, 2006) or waive fees and pay people to go to school. A leader could also improve access to immunization and antiretroviral drugs and increase the transfers to the health care sector (Friedman, 2018; Jones et al., 2003). A leader could also improve the status of infrastructure, water and electricity, by favoring regions that are populated by their own ethnic group (Hodler and Raschky, 2014b).

Another way that a leader could improve the fortune of their ethnic group is by offering government jobs to them. By offering more public sector jobs to members of their group, the leader is decreasing unemployment among their group and increasing their wealth and income. An increase in income could in return affect the health and education outcomes of the ethnic group. Families, in this situation, are more likely to afford health care and to send their children to schools.

Figure II and table II summarize the effect of ethnic favoritism on education, wealth, infant survival, electrification and access to clean drinking water. People that spent all of their primary school when a leader was co-ethnic were 3.8 percentage points less likely to finish primary school, though the result was statistically insignificant. A child whose mom shared the same ethnicity as the leader of the country she lived in two years before birth is 1.3 percentage points more likely to survive the first 12 months of birth. Co-ethnics were as likely as their non-co-ethnic peers to be in the same wealth quintile and have electricity. Finally, co-ethnics have better access to clean drinking water than non-co-ethnics.

Figure III and table III summarize the effect of co-ethnicity on the outcomes of interest once I allow for the heterogeneous effect of democracy. The interaction between democratic bins and co-ethnicity is statistically significant when studying their effect on electrification and access to clean drinking water.<sup>2</sup> Co-ethnicity and more democracy do not eliminate ethnic favoritism when it comes to primary school completion, infant survival and wealth. Co-ethnics in democracies are as likely to finish primary schooling, survive the first year

2. Democratic bins include a bin for democracies with a polity score between 10 and 5, anocracies with a polity score between 5 and -5, and dictatorships with a polity score less than -5.

of birth and belong to the same wealth quintile as their non-co-ethnic peers. Co-ethnics in anocracies are 3 percentage points more likely to finish primary schooling but are as likely to survive the first year of birth and belong to the same wealth quintile as their non-co-ethnic peers. However, Co-ethnics in democracies are 15.4 percentage points more likely to have access to electricity. Co-ethnics in anocracies are 6.7 percentage points more likely to have access to electricity. Finally, co-ethnics in both democracies and anocracies are more likely to have better access to clean drinking water.

The rest of this paper will be divided into four sections. I will discuss the theoretical framework in Section II, section III will be on the data and the construction of the co-ethnic leader variable. Sections IV, V, and VI will respectively layout an empirical model, discuss the results and then conclude.

## II. THEORETICAL FRAMEWORK

### II.A. *Ethnic Favoritism Models*

The literature on ethnic favoritism provides us with a few models that explain the existence of the phenomena (Cox and McCubbins, 1986; Dixit and Londregan, 1996; Lindbeck and Weibull, 1987). I will discuss these models and offer implication that arises from them.

Cox and McCubbins (1986) present a model where they assume that a leader's utility is directly affected by the well-being of their ethnic kins. The strong assumption they make will steer a leader to favor their ethnic group regardless of the group's political behavior. In other words, a president of a country will always be allocated funds to favor their ethnic group even if they do not vote for them. This is referred to as the "ethnic altruism" model.

Dixit and Londregan (1996) assume that a leader's only objective is to maximize time in power. In this model, a leader will transfer funds to several groups, not only their own, to increase public support. They also assume that ethnic kins get utility from the mere fact that a co-ethnic person is holding high office. This is referred to as the "psychic benefit" (Chandra, 2007). The psychic benefit model provides an important implication: members

of an ethnic group will give unequivocal support to a co-ethnic leader even if they do not get anything in return. Under this scenario, a leader will have no incentives to favor their ethnic group, therefore, ethnic favoritism will not exist. This model is inconsistent with the empirical evidence that shows the existence of ethnic favoritism.

Finally, [Lindbeck and Weibull \(1987\)](#) established a model wherein a leader is an office-seeker. The office-seeking behavior requires political support. They drop the “psychic benefit” assumption. In this model, an ethnic group will only support a co-ethnic leader if she/he provides them with material benefits, like schools and roads. A leader will support their ethnic groups for two reasons. First, it could be the case that the support of the co-ethnic group is cheaper than other groups. The lower cost of allocating funds comes from the fact that the leader understands the needs of their ethnic group and is better at allocating resources to meet these needs. Second, if a leader is risk-averse, then they will trust the promises that their ethnic group will support them more than the other ethnicities. Therefore, it is less risky for them to buy the support of their group in exchange for benefits. This is called the *quid pro quo* model.

There are several tools that a leader could use to benefit the educational, health, wealth and infrastructure outcomes of their ethnic kins. Leaders could hire more teachers, improve the physical health of schools ([Glewwe and Kremer, 2006](#)) or waive fees and pay people to go to school. A leader could also improve access to immunization and antiretroviral drugs and increase the transfers to the health care sector ([Friedman, 2018](#); [Jones et al., 2003](#)). A leader could also improve the status of infrastructure, water and electricity, by favoring regions that are populated by their own ethnic group ([Hodler and Raschky, 2014b](#)).

Another way that a leader could improve the fortune of their ethnic group is by offering government jobs to them. By offering more public sector jobs to members of their group, the leader is decreasing unemployment among their group and increasing their wealth and income. An increase in income could in return affect the health and education outcomes of the ethnic group. Families now could afford health care and send their children to schools.



The theory on ethnic favoritism yields different results under different assumptions. These conclusions are in contrast with the empirical evidence that shows that ethnic favoritism exists. Therefore, testing for ethnic favoritism using different data sets, countries and analyses remains an important contribution to the literature of ethnic favoritism.

### III. DATA

The data I am using is from twenty-one African countries spanning the period from 1947 to 2000. I present the summary statistics of the sample in table I.

#### III.A. *Demographics and Health Survey*

The Demographics and Health Survey (DHS) is a nationally representative survey conducted in over 90 countries. It aims to improve the understanding of global development. The United States Agency for International Development, known as USAID, funds the DHS project. In their surveys, they collect a myriad of significant variables in measuring a country's development. Moreover, the female and male surveys collect detailed information, like schooling, access to water, and ethnicity, that are important in estimating ethnic favoritism. Since ethnicity is more salient in African countries, I will limit my data to the continent (Murphree, 1988; Posner, 2004). Out of the African countries surveyed by the DHS, I excluded the ones where ethnic information was not collected. For example, in Burundi, the DHS only asked a question about the nationality of the interviewee.

By merging this data set with other data that I am using, I will create a few databases consisting of between 822,347 and 2,651,752 observations from twenty-one countries. Moreover, the data include cohorts that were born between 1947 to 2000. I introduce the summary statistics in table I. The average age was 29.01, and females in the sample had an average of 3.02 children. On average, a household consisted of 7.34 members. Access to electricity and wealth index are measures of well-being.<sup>3</sup> 35% of the sample belong to the

3. The wealth index is a value of standardized scores and factor coefficient scores of wealth indicators.

**TABLE I**  
SUMMARY STATISTICS

Characteristic	N = 1,387,301
Age	29 (10) [10, 98]
Total children ever born	3.02 (3.99) [0.00, 99.00]
Total number of household members	7.3 (5.0) [1.0, 80.0]
Has electricity	0.31
Currently working	0.55
Wealth Quintile	
1	0.19
2	0.18
3	0.19
4	0.20
5	0.24
Total years of education	5.0 (4.7) [0.0, 26.0]
Completed primary school	0.65
Urban	0.36
Literacy	0.58

<sup>1</sup> makecell[l]Mean *SD*) [*Range*];%

poorest wealth quintiles, 18% to the middle class. Only 58% of the people surveyed had access to electricity. The mean total years of education was about 5 years.

### ***III.B. Democracy Data***

I use Polity IV ([Marshall et al., 2019](#)) as an indicator for democratic institutions and supplemented it with the Democracy-Dictatorship Index ([Cheibub et al., 2010](#)). Polity IV is an annual measure that scores each country, since its independence, from -10 to 10— 10 being the most democratic and -10 the most autocratic. The scoring is based on five criteria:

1. Competitiveness of the executive.
2. The openness of the executive.

3. Regulation of political participation.
4. The competitiveness of political participation
5. Constraints on the executive branch.

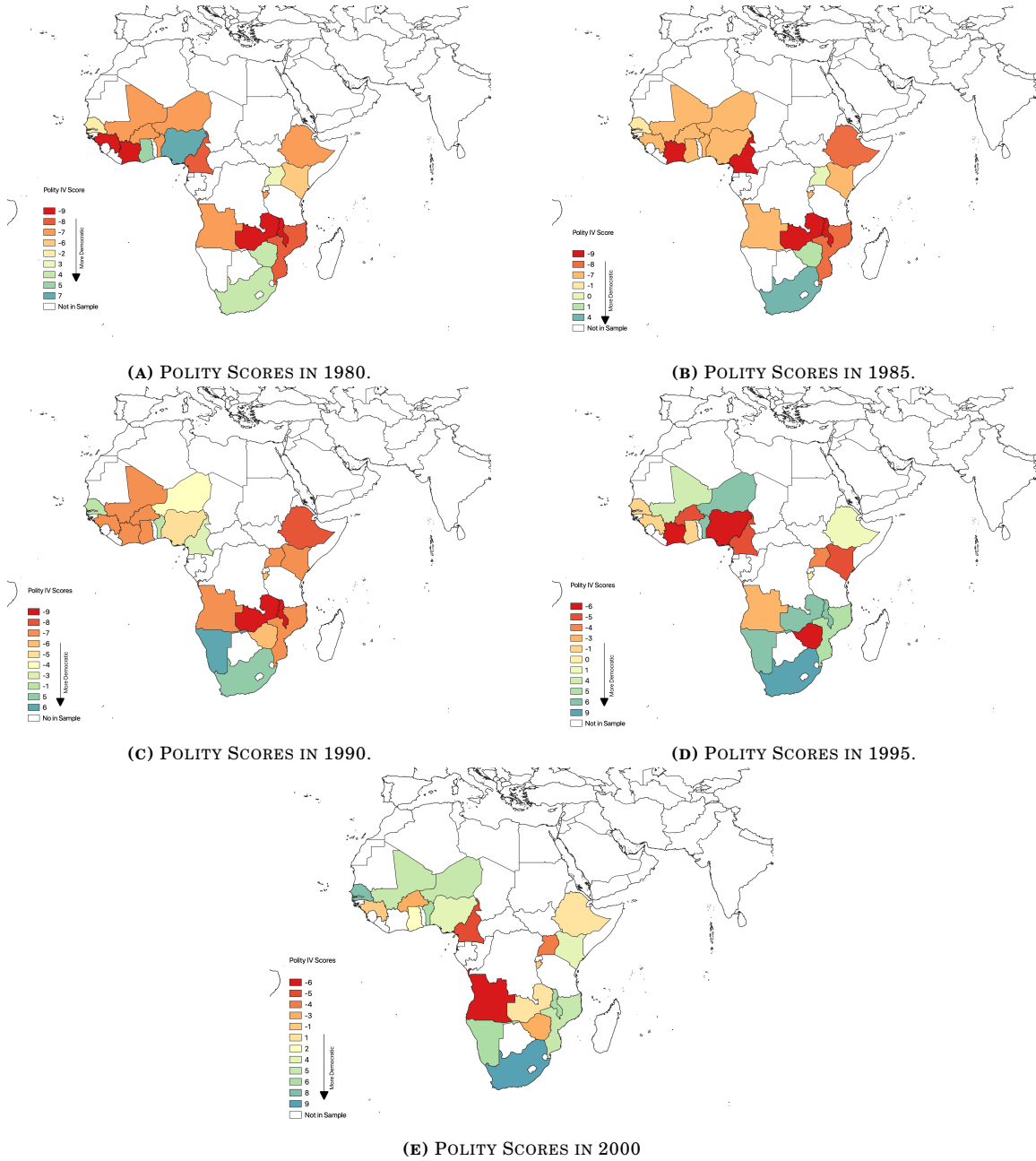
The first two conditions are concerned with barriers to entry, or lack of, to become the leader of the executive branch (i.e., the president or prime minister). The easier it is for a person to enter the election, the more democratic the country. The third and fourth benchmarks focus on political life in a country. The less constrained a person's freedoms are, the more democratic the state. The final requirement relates to the amount of power a country's leader could yield. On the one hand, in the United States, the president's powers are checked by the legislative and judicial branches of government. On the other hand, President Mobutu Sese Seko of the Democratic Republic of Congo ruled with an iron fist from 1965 till 1996. His powers were unchecked, and his decisions were undoubted.

Moreover, I use the Democracy-Dictatorship Index to complement the data set and not as a primary source of data on democracy. I will also use the Democracy-Dictatorship to check the robustness of my results (see these results in table V and figure IV). The Democracy-Dictatorship Index (Cheibub et al., 2010) has a list of the holders of countries' top two offices— for example, president and vice president— with their titles and source of power.<sup>4</sup> For instance, Idi Amin of Uganda became president in 1971 through a coup d'etat. While Nelson Mandela, of South Africa, was the first South African to be elected in a free election in 1994. The average polity score in the data set was -4.75 (see table I).

Furthermore, I present the evolution of democracy from 1980 to 2000 in figure I. Notice how Ghana, with a score of 5, and Nigeria, with a score of 7, were the most democratic countries in 1980. During the 1980s, the apartheid state's repeal in South Africa had just begun, which explains the score of 4. By the year 2000, five countries— Benin, Malawi, Namibia, Senegal, and South Africa— were democracies.

4. By source of power, I mean if a leader was elected or got into power through a coup.

**FIGURE I**  
**DEMOCRACY IN SAMPLE COUNTRIES FROM 1980 TO 2000.**



*Note.* In this figure, I present the changes in Polity IV scores in the countries I am studying from 1980 to 2000. Notice Africa's democratization.

### ***III.C. Leaders, their ruling period and ethnicity***

To construct a variable for co-ethnicity, I had to know the ethnic group of every leader in my sample. Consequently, I rely on Archigos, a data set that records the name of every leader

of a country since its independence (Goemans et al., 2006). I also use the data from Fearon et al. (2007) to construct a data set on the ethnicity of leaders around the world from 1945 to 1999. I supplement Fearon et al. (2007) the data to include more years, and consequently, more leaders.

## IV. THE EMPIRICAL MODEL

### IV.A. *Ethnic Favoritism in Africa.*

In this section, I will test for the existence of ethnic favoritism in Africa using the Difference-in-Differences (DD) method. Using the DHS data, I can investigate the existence of ethnic favoritism using a few outcomes of interest like education, wealth, electrification, access to drinking water and infant survival.

To estimate the average ethnic favoritism, I will run the following regression:

$$Y_{iceta} = \beta_0 + \beta_1 coethnic_{iceta} + \delta_t + \gamma_{ce} + \eta_a + \varepsilon_{iceta} \quad (1)$$

Where  $Y_{iceta}$  is the outcome of interest of person  $i$  in the country  $c$ , of an ethnic group  $e$  that was born at time  $t$ <sup>5</sup> and with age at the time of survey  $a$ . While  $coethnic_{iceta}$  is a variable that indicates the co-ethnicity between person  $i$  and the leader of a country at a given time. Under the specification I am using, co-ethnicity is a dummy variable that is equal to one if a person  $i$  is co-ethnic with the leader of a country at a given time. I also used other specifications for co-ethnicity in which the co-ethnicity variable takes values between 0 and 1 to indicate the portion of time spent under a co-ethnic leader. The variables  $\delta_t, \gamma_{ce}$  and  $\eta_a$  are, respectively, vectors of time, country-specific ethnic group and age at the time of survey fixed effects.

Coefficient  $\beta_1$  from equation 1 estimates the differences in outcomes between co-ethnics

5. In the cases when the outcomes are education, electrification, wealth and access to water  $t$  is the time of the survey. In the case when the outcome is infant survival,  $t$  is the year of birth.

and non-co-ethnics. If  $\beta_1 > 0$ , then the ethnic kins of a president enjoy an advantage in outcome  $Y$ . If  $\beta_1 < 0$ , the people that do not share the same ethnicity as their leader enjoy an advantage in outcome  $Y$ . Finally, when  $\beta_1 = 0$  there is no evidence that co-ethnics enjoy an advantage or disadvantage when compared to non-co-ethnics.

For  $\beta_1$  to be interpreted as ethnic favoritism, the differences in outcomes should be driven solely by a disproportional allocation of resources from a leader to their ethnic kins. The specification in equation 1 helps disentangle and weed out some of the threats to the identification. When using difference-in-differences, I am comparing the differences in outcomes of the different groups during periods when the president was from their groups to period when the president was not from their group. Thus, the difference-in-differences specification helps separate the effect of having a co-ethnic as a president from other factors, like positive shocks to agriculture. Moreover, including survey year and ethnic group, fixed effects can produce an estimate that separates the effect of having a co-ethnic leader in power from group-specific and time-invariant characteristics.

Furthermore, as pointed out by [Kramon and Posner \(2013\)](#), the problem of stock *versus* flow threaten identification. This problem is specifically severe when studying the effect of co-ethnicity on access to electricity and clean drinking water. Electrification and access to clean water can persist for a long time (stock). The fact that we can observe the different groups at a different point in time—using the different DHS waves— I can compare if having a co-ethnic president did affect access to the two services. As for education and infant survival, we know when an infant was born—a health care service was provided— and we know when a person received a year of primary education. Therefore, we can pin down the effect of co-ethnicity and make a stronger inference.

#### ***IV.B. Ethnic Favoritism and Democracy in Africa.***

The data I obtained allow me to construct a rich empirical model to investigate the existence of a differential and heterogeneous co-ethnic effect between countries that are more or

less democratic. My specification also allows me to identify the effect of democracy within countries. I describe the regressions in equations 2 and 3. Since I aim to study the effect of ethnic favoritism, I had to use the ethnicity variables provided by the DHS and the ethnic group of a leader from Fearon et al. (2007) to create the co-ethnicity dummy variable. This variable is equal to one if person  $i$  and the leader of the country they live in are from the same ethnic group.

The regression equations I will use are given by:

$$Y_{iceta} = \beta_0 + \beta_1 coethnic_{iceta} + \beta_2 DemocracyIndex_{ct} + \beta_3 coethnic_{iceta} \cdot DemocracyIndex_{ct} + \delta_t + \gamma_{ce} + \eta_a + \varepsilon_{iceta} \quad (2)$$

$$Y_{iceta} = \beta_0 + \beta_1 coethnic_{iceta} + \beta_2 I_{(5 \leq PolityIV \leq 10)} + \beta_3 I_{(-5 \leq PolityIV \leq 5)} + \beta_4 I_{(5 \leq PolityIV \leq 10)} \cdot coethnic_{iceta} + \beta_5 I_{(-5 \leq PolityIV \leq 5)} \cdot coethnic_{iceta} + \delta_t + \gamma_{ce} + \eta_a + \varepsilon_{iceta} \quad (3)$$

Where  $Y_{iceta}$  is the outcome of interest for person  $i$  in the country  $c$ , ethnic group  $e$ , at period  $t$ <sup>6</sup> and with age at the time of the survey  $a$ .  $coethnic_{iceta}$  is the dummy variable for co-ethnicity and is equal to one if the observation  $i$  is the same ethnicity as the leader at time  $t$  in the country  $c$ .  $DemocracyIndex_{ct}$  is the index for democracy. Variables  $I_{(5 \leq PolityIV \leq 10)}$ , and  $I_{(-5 \leq PolityIV \leq 5)}$  are indicator variables of democratic bins that respectively represent democracies, and anocracies.  $I_{(-10 \leq PolityIV \leq -5)}$  is the omitted dictatorship group. The variables  $\delta_t, \gamma_e$  and  $\eta_a$  are, respectively, vectors of time, country-specific ethnic group and age at the time of survey fixed effects.

We are interested in variables  $\beta_3$  from equation 2 and  $\beta_5$ , and  $\beta_6$  from 3.  $\beta_3$  and  $\beta_5$ , and  $\beta_6$  estimate the differences and relationships of ethnic favoritism in democratic/anocratic and autocratic countries on the outcomes of interest.

6. In the cases when the outcomes are education, electrification, wealth and access to water  $t$  is the time of the survey. In the case when the outcome is infant survival,  $t$  is the year of birth.

## V. RESULTS AND DISCUSSION

In this section, I will introduce the results of estimating my empirical models. First, I investigate the existence of ethnic favoritism on schooling, electrification, access to clean drinking water, wealth and infant survival. I find evidence of ethnic favoritism on infant survival and access to clean drinking water (equation 1). Second, I include a measure of democracy and investigate the effect of democracy on ethnic favoritism on the outcomes of interest. The interaction between democracy and co-ethnicity is statistically significant when studying their effect on electrification and access to clean drinking water (equation 3).

I find evidence of ethnic favoritism on infant survival and access to clean drinking water. When I introduce a measure of democracy, I find a statistically significant effect of the interaction between democracy and co-ethnicity on electrification and access to clean drinking water. A child whose mother is a co-ethnic of the leader in power two years before their birth is 1.3 percentage points more likely to survive their first year after birth. A person who was of the same ethnicity as the leader two years prior to the survey is more likely to have access to clean drinking water than a non-co-ethnic person. Putting the ethnic favoritism on infant survival in perspective, 87.5% of non-co-ethnic children in the sample survive the first year of life. Consequently, ethnic kins would, on average, have a survival rate of 88.8%. However, when comparing co-ethnic and non-co-ethnics in democracies or anocracies and autocracies, I find evidence of ethnic favoritism on electrification and access to clean drinking water in anocracies and democracies as well as education in anocracies.

### ***V.A. Existence of Ethnic Favoritism***

I look into the existence of ethnic favoritism on primary school completion, electrification, access to clean drinking water, wealth and infant survival.<sup>7</sup> I report the results in table II

7. The primary school completion is a dummy variable that is equal to 1 if a person completed primary school and zero otherwise. The co-ethnic variable for education as an outcome is equal to the total years between the ages of seven and thirteen that a person lived under the rule of a co-ethnic leader divided by seven. For electrification and access to water, the co-ethnic variable is equal to one of the leaders two years before the interview date was a co-ethnic. For wealth, the co-ethnic variable is equal to the total years a



and figure II. These results are the estimates of equation 1.

**TABLE II**  
ETHNIC FAOVIRTISM RESULTS.

	(1)	(2)	(3)	(4)	(5)
	Schooling	Infant Survival	Wealth	Electrification	Water
Coethnic	-0.038 (0.024)	0.013*** (0.005)	0.067 (0.043)	0.043 (0.044)	0.210** (0.091)
Observations	1,146,278	2,516,752	684,762	822,347	822,347

Standard errors in parentheses

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

<sup>1</sup> In this table, I am reporting the estimates of equation 1. I present the results of the effect of ethnic favoritism on primary school completions in column 1, infant survival in column 2, wealth quintile in column 3, electrification in column 4 and access to clean drinking water in column 4. Primary school completion is a dummy variable that is equal to one if a person completed primary school and zero otherwise. Infant survival is a dummy variable that is equal to one if an infant survived the first 12 months of life. Electrification is a dummy variable that is equal to one if a household has electricity. Finally, access to clean drinking water is an ordinal variable that has values from 1, worst water source, to 4.

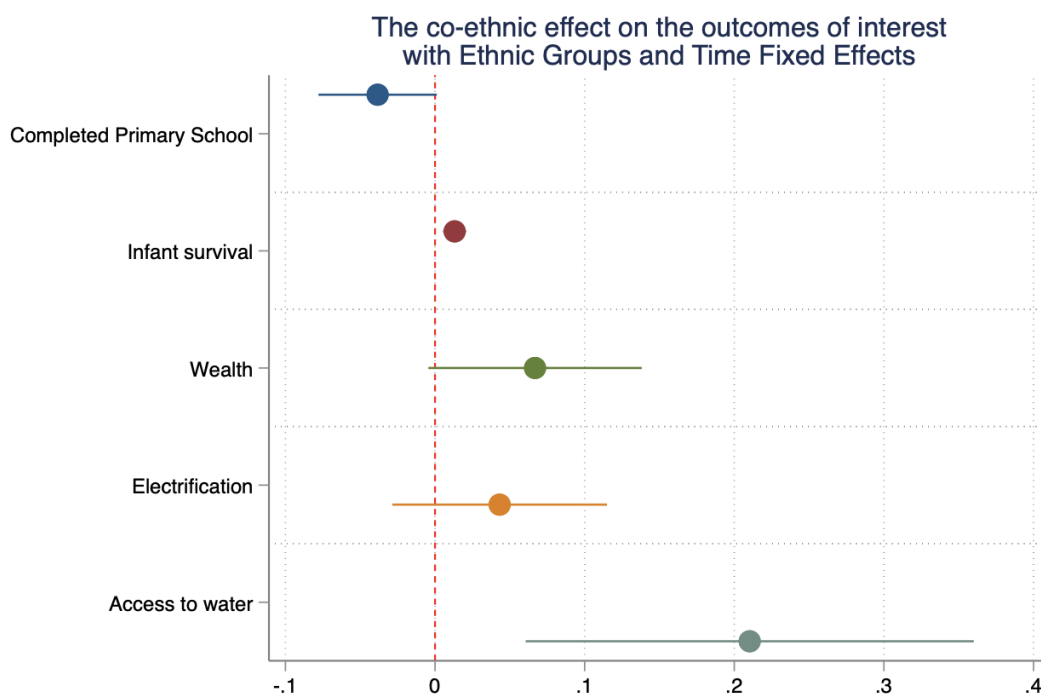
<sup>2</sup> Standard errors are clustered on ethnic groups. All results include ethnic group, time and age fixed effects.

In table II, I estimate the average effect of having a co-ethnic leader on primary school completion (column 1), infant survival (column 2), wealth quintile (column 3), electrification (column 4) and access to clean drinking water (column 5). A respondent that spent all of her primary school under the rule of a co-ethnic leader was 3.8 percentage points less likely to finish primary school, but the result was statistically insignificant (table II column 1). As for children's health (table II column 2), a child whose mother was of the same ethnic group as the president two years before birth was more likely to survive the first 12 months of their life. A co-ethnic child is 1.3 percentage points more likely to survive the first 12 months after birth.<sup>8</sup> A person that lived under the rule of a co-ethnic leader in the four years before the survey is more likely to be in a higher wealth quintile (table II column 3). Finally, a respondent that had a co-ethnic leader two years before the survey

person spent under the rule of a co-ethnic leader in the four years before the interview was conducted divided by four. For infant mortality, the co-ethnic variable is a dummy that is equal to one of the presidents who was of the same ethnicity as the mother the year the child was born. The infant mortality variable was constructed following Kudamatsu (2009, 2012).

8. The magnitude of the effect of ethnic favoritism on infant survival is larger than an intervention to reduce infant mortality (Jones et al., 2003).

was conducted is more than 4.3 percentage points more likely to have electricity and has better access to clean drinking water (table II columns 4 and 5). To the exception of the effect of co-ethnicity on infant survival and access to clean drinking water, none of the other results were statistically significant. The point estimate of the effect of ethnic favoritism on primary school completion is large but insignificant. The confidence interval, however, includes mostly negative values. Ranging from  $-8.55$  to around  $0.88$  percentage points. To put this into perspective,  $65\%$  of the sample completed primary school. Meaning that, on average,  $67.1\text{-}80.9\%$  of co-ethnics finished primary school while  $75.56\text{-}80.08\%$  of non-co-ethnics finished primary school. The point estimates of co-ethnicity on infant survival suggest the existence of ethnic favoritism. The effect of co-ethnicity on infant survival is  $1.3$  percentage points.



**FIGURE II**

EFFECT OF CO-ETHNICITY ON OUTCOMES OF INTEREST WITH ETHNIC GROUP AND TIME FIXED EFFECTS.

*Note.* In this figure, I present the estimates of the co-ethnicity variable on primary school completion, infant survival, wealth, electrification and access to clean drinking water. The bands represent the 90% confidence interval and the standard errors are clustered on ethnic groups. All estimates include ethnic group, time and age fixed effects.

These results fit into the empirical results of [Kramon and Posner \(2013\)](#) and the theoretical work of [Francois et al. \(2015\)](#). In [Kramon and Posner \(2013\)](#), the authors found inconclusive evidence on the existence of ethnic favoritism. [Kramon and Posner \(2013\)](#) did not find any evidence of systematic favoritism in one area across the six countries they analyzed. Favoritism was country-dependent. For example, they found no evidence of ethnic favoritism on education in Benin and Zambia. They also found that there is ethnic favoritism in education in Kenya and Malawi and ethnic disadvantage in Senegal and Mali. While [Francois et al. \(2015\)](#) showed that Africa is more ethnically representative than Western democracies. Consequently, the empirical results of this paper fit with the theory. Leader aims to stay in power. They appoint cabinet members of many ethnic groups hoping that these cabinet members will share the wealth with their ethnic groups. Therefore, they can maximize their time in power by preventing coups and revolutions.

### ***V.B. Ethnic Favoritism and Democracy***

In this section, I will provide the results of the effect of ethnic favoritism on the outcomes of public goods provisions—education, infant health, wealth, access to electricity and clean drinking water—when a measure of democracy is introduced. The empirical specification is provided in equation 3. This specification allows for heterogeneity in treatment, in this case, exposure to democracy. Therefore, testing for differential co-ethnic effects between countries that are more or less democratic. The results are presented in table III and figure III. The coefficients provided are the heterogeneous effects of co-ethnicity in democracies/anocracies and autocracies.

In table III column 1, I present the estimates on primary school completion, column 2 on infant survival, column 3 on wealth quintile, column 4 on electrification and column 5 on access to clean drinking water. When a measure of democracy is introduced, the interaction between democratic bins and co-ethnicity are statistically significant when studying

**TABLE III**  
ETHNIC FAOVIRTISM AND DEMOCRACY RESULTS.

	(1) Schooling	(2) Infant Survival	(3) Wealth	(4) Electrification	(5) Water
<i>Democracy</i> $\times$ <i>Coethnic</i>	0.033 (0.051)	-0.004 (0.008)	0.079 (0.074)	0.154*** (0.043)	0.279* (0.142)
<i>Anocracy</i> $\times$ <i>Coethnic</i>	0.030** (0.014)	0.001 (0.005)	-0.078 (0.055)	0.067* (0.039)	0.297* (0.162)
Coethnic	-0.057** (0.028)	0.005 (0.007)	0.062 (0.074)	-0.077** (0.038)	-0.321** (0.150)
Observations	1,146,305	2,516,748	684,762	822,330	822,330

Standard errors in parentheses

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

<sup>1</sup> In this table, I am reporting the estimates of equation 3. I present the results of the interaction between the coethnic variable and *Polity IV* groups on primary school completions in column 1, infant survival in column 2, wealth quintile in column 3, electrification in column 4 and access to clean drinking water in column 4. Primary school completion is a dummy variable that is equal to one if a person completed primary school and zero otherwise. Infant survival is a dummy variable that is equal to one if an infant survived the first 12 months of life. Electrification is a dummy variable that is equal to one if a household has electricity. Finally, access to clean drinking water is an ordinal variable that has values from 1, worst water source, to 4.

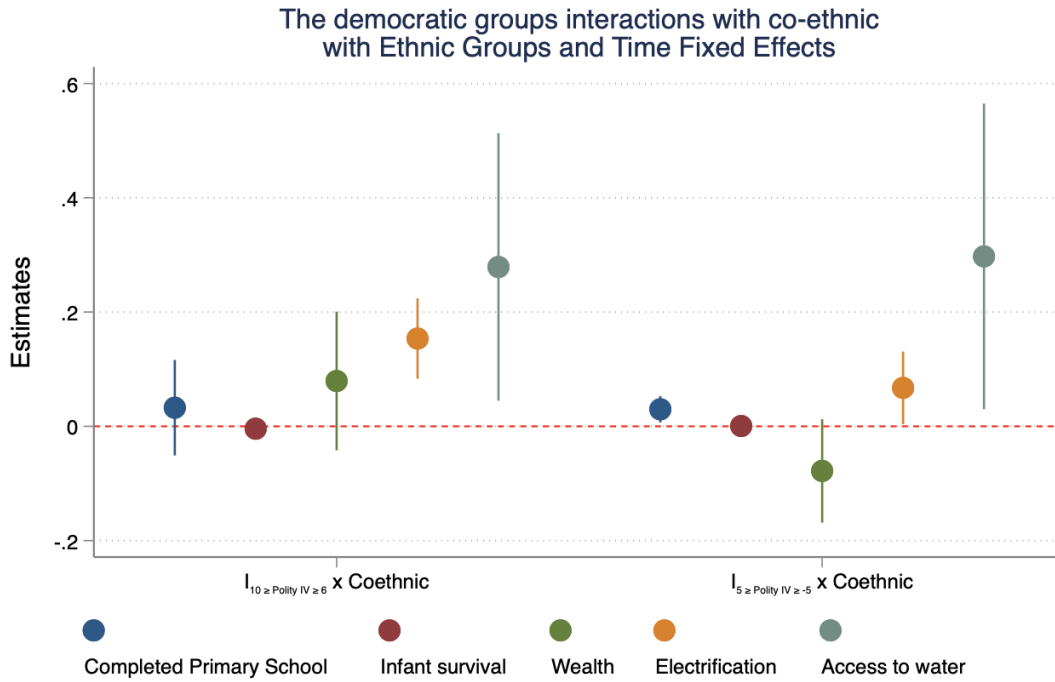
<sup>2</sup> Democracies have a *Polity IV*  $\in [10, 5]$ , anocracies have a *Polity IV*  $\in [4, -5]$  and autocracies, the omitted group, have a *Polity IV*  $< -5$ .

<sup>3</sup> Standard errors are clustered on ethnic groups. All results include ethnic group, time and age fixed effects.

their effect on schooling, electrification and access to clean drinking water.<sup>9</sup> The estimate of the interaction between co-ethnicity and the most democratic countries on primary school completion—table III column 1—ranges from a reduction of 6.7 to an increase of 13.2 percentage points. While the estimate of the interaction between co-ethnicity and anocracies has a point estimate of 3 percentage points. Indicating that a co-ethnic enjoys an increase of 3 percentage points in the likelihood of completing primary school. When studying the interaction between co-ethnicity on electrification (table III column 4), the estimate of the interaction between co-ethnicity and democracy and anocracy indicates an increase of 15.4 and 6.7 percentage points in the likelihood of having access to electricity. The point estimate of the effect of co-ethnicity on infant survival is small and close to zero (table III column 2). It ranges from a 2 percentage points reduction and 1.2 percentage points increase in democ-

9. Democratic bins include a bin for democracies with a polity score between 10 and 5, anocracies with a polity score between 5 and -5, and dictatorships with a polity score less than -5.

racies and a 1 percentage points reduction and 1.1 percentage points increase in anocracies. Ethnic kins in both democracies and anocracies are more likely to have better access to clean drinking water than a non-co-ethnic person (table III column 5. Finally, the wealth of ethnic kin in the most democratic countries is more likely to be in a higher wealth quintile, but the effect is insignificant. Ethnic kin in anocracies, however, is more likely to be in a lower wealth quintile, but the effect is insignificant (table III column 3).



**FIGURE III**  
WITH ETHNIC GROUPS AND TIME FIXED EFFECTS.

*Note.* In this figure, I present the effect of the interaction between co-ethnicity variable and Polity groups on primary school completion, infant survival, wealth, electrification and access to clean drinking water. Democracies have a *Polity IV*  $\in [10, 5]$ , anocracies have a *Polity IV*  $\in [4, -5]$  and autocracies, the omitted group, have a *Polity IV*  $< -5$ . The bands represent the 90% confidence interval and the standard errors are clustered on ethnic groups. All estimates include ethnic group, time and age fixed effects.

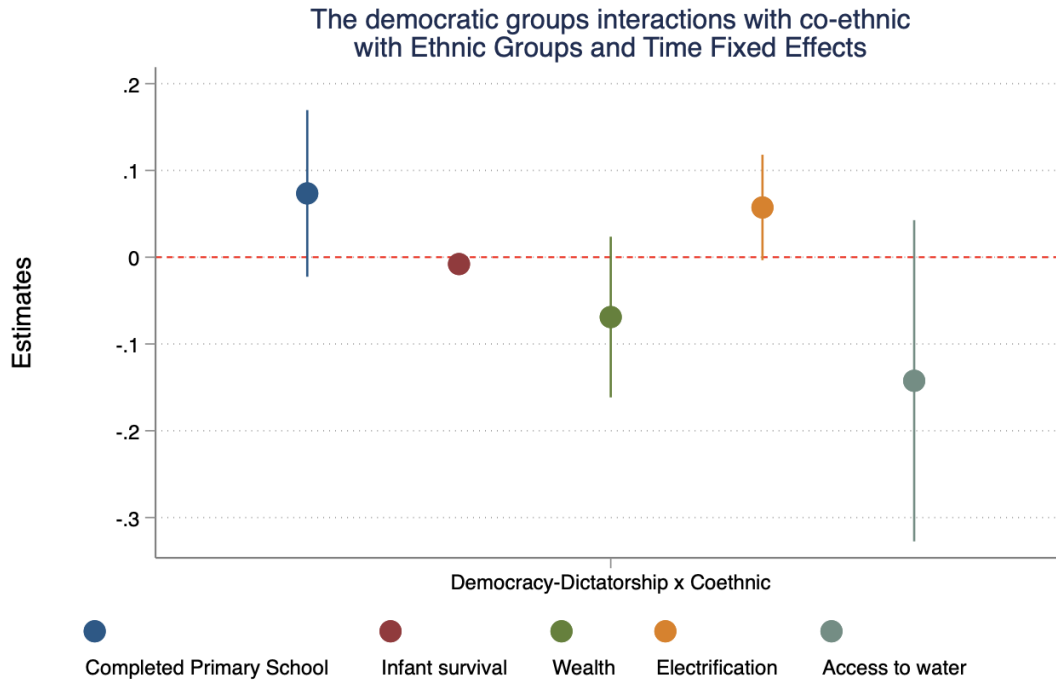
Putting the result in context— 36% of non-co-ethnic observations in the sample had electricity. An ethnic kin in the most democratic countries would have an electrification rate of 51.4%. Having more exposure to democracy increases ethnic favoritism in electrification. These results could be driven by the fact that leaders will now target funds to their ethnic

group to earn their support in elections. Especially since the support of ethnic kins could come at a lower bill. Moreover, 77.6% of non-co-ethnics finished primary school education while 80.6% of co-ethnics in anocracies finished primary school education.

### ***V.C. Discussion***

The results from the previous sections capture, *ceteris paribus*, the differences in outcomes between co-ethnic and non-co-ethnics. To interpret the results from the two previous sections as causal, I have to assume that the transitions between leaders of different ethnicities and in how democratic a country is are exogenous to the changes in an ethnic groups' education, health, electricity usage, wealth, and access to clean drinking water. A concern might arise from the fact that democracy is endogenous and is probably correlated with the outcomes I am studying. Even though the concern is correct, I am not interpreting the coefficient of democracy as a causal one. The coefficient of interest in this paper is that of the interaction between democracy and the co-ethnic variable. The coefficient of the interaction is less likely to suffer from an omitted variable bias. For the interaction between democracy and co-ethnicity to be endogenous, and thus have biased results, there must exist an omitted variable that is correlated with whether a country is a democracy, or not, and with how co-ethnics and non-co-ethnics are treated by their leaders. Such omitted variable is less likely to be found.

Another concern could arise from the measure of democracy, *Polity IV*, that I am using. Defining and quantifying the democratic institutions in a country is not straightforward and democracy differs from one country to another. My empirical strategy includes country and time-fixed effects. The inclusion of these fixed effects could alleviate the concern of the differences between democracy because it allows me to have a between and with-in country analysis. Moreover, the *Polity IV* measure is developed based on five categories that aim to disentangle the differences in how more, or less, democratic a country is. Furthermore, I ran equation 2 with a different measure of democracy. I used the *Democracy-Dictatorship*



**FIGURE IV**

EFFECT OF THE INTERACTION BETWEEN CO-ETHNICITY AND DEMOCRACY-DICTATORSHIP INDEX ON OUTCOMES OF INTEREST WITH ETHNIC GROUPS AND TIME FIXED EFFECTS.

*Note.* In this figure, I present the effect of the interaction between co-ethnicity variable and the Democracy-Dictatorship on primary school completion, infant survival, wealth, electrification and access to clean drinking water. In this specification, I used the Democracy-Dictatorship index as a measure of democracy. The Democracy-Dictatorship index is an indicator variable that is equal to 1 if a country is democratic and zero otherwise. The bands represent the 90% confidence interval and the standard errors are clustered on ethnic groups. All estimates include ethnic group, time and age fixed effects.

*Index* instead of *Polity IV*. The results to this regression are presented in table V and figure IV. Using the *Democracy-Dictatorship Index*, I retrieved point estimates that have the same sign as these from the *Polity IV* regression. The point estimates from the *Polity IV* regression were significant for schooling in anocracies and electrification and access to clean drinking water in democracies and access wealth in anocracies. None of the results using the *Democracy-Dictatorship Index* were significant. It could be the case that the coefficient of the effect of interaction between *Democracy-Dictatorship Index* and co-ethnicity became insignificant because the *Democracy-Dictatorship Index* lumps anocracies and democracies together, therefore the estimate will be averaging over both groups.

**TABLE IV**  
ETHNIC FAOVIRTISM AND CONTINUOUS DEMOCRACY MEASURE RESULTS.

	(1) Schooling	(2) Infant Survival	(3) Wealth	(4) Electrification	(5) Water
<i>PolityIV</i> × <i>Coethnic</i>	0.008* (0.004)	-0.000 (0.001)	0.002 (0.006)	0.012*** (0.005)	0.016* (0.009)
Coethnic	-0.019 (0.017)	0.005 (0.004)	0.066 (0.053)	0.000 (0.021)	-0.089* (0.052)
Observations	1,135,466	2,511,726	678,732	816,300	816,300

Standard errors in parentheses

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

<sup>1</sup> In this table, I am reporting the estimates of equation 2. I present the results of the interaction between the coethnic variable and *Polity IV* groups on primary school completions in column 1, infant survival in column 2, wealth quintile in column 3, electrification in column 4 and access to clean drinking water in column 4. Primary school completion is a dummy variable that is equal to one if a person completed primary school and zero otherwise. Infant survival is a dummy variable that is equal to one if an infant survived the first 12 months of life. Electrification is a dummy variable that is equal to one if a household has electricity. Finally, access to clean drinking water is an ordinal variable that has values from 1, worst water source, to 4.

<sup>2</sup> *Polity IV* score in this specification is continuous. It takes values that range from most autocratic -10 to most democratic 10.

<sup>3</sup> Standard errors are clustered on ethnic groups. All results include ethnic group, time and age fixed effects.

**TABLE V**  
ETHNIC FAOVIRTISM AND DEMOCRACY-DICTATORSHIP INDEX RESULTS.

	(1) Schooling	(2) Infant Survival	(3) Wealth	(4) Electrification	(5) Water
<i>D - DIndex</i> × <i>Coethnic</i>	0.073 (0.058)	-0.008 (0.007)	-0.069 (0.056)	0.057 (0.037)	-0.142 (0.112)
Coethnic	-0.066 (0.040)	0.008 (0.006)	0.135*** (0.042)	0.002 (0.015)	-0.006 (0.066)
Observations	988,238	2,070,049	380,253	453,655	453,655

Standard errors in parentheses

\*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$

<sup>1</sup> In this table, I am reporting the estimates of equation 1. I present the results of the interaction between the coethnic variable and *Polity IV* groups on primary school completions in column 1, infant survival in column 2, wealth quintile in column 3, electrification in column 4 and access to clean drinking water in column 4. Primary school completion is a dummy variable that is equal to one if a person completed primary school and zero otherwise. Infant survival is a dummy variable that is equal to one if an infant survived the first 12 months of life. Electrification is a dummy variable that is equal to one if a household has electricity. Finally, access to clean drinking water is an ordinal variable that has values from 1, worst water source, to 4.

<sup>2</sup> In this specification, I used the Democracy-Dictatorship index as a measure for checks and balances. The Democracy-Dictatorship index is an indicator variable that is equal to 1 if a country is democratic and zero otherwise.

<sup>3</sup> Standard errors are clustered on ethnic groups. All results include ethnic group, time and age fixed effects.



Another concern could be the “psychic effect” of having a co-ethnic person as the leader of a country. The “psychic effect” implies that a person will have better outcomes because ethnic kin became a leader without intervention by the leader herself. For example, [Marx et al. \(2009\)](#) find that test scores of African American children in the US improved following the election of Barack Obama. The authors argue that the increase in test scores does not reflect an allocation of resources that targeted African American pupils but the greater motivation of having a role model in an important position. The “psychic effect” is unlikely to cause a sudden increase in electrification and access to clean drinking water without the allocation of resources to benefit one ethnic group. Just because a leader is of a specific ethnic group will not magically create an infrastructure that delivers electricity and clean drinking water. The “psychic effect”, however, can influence education, health and wealth. To address this concern I use the nightlight used by [Hodler and Raschky \(2014b\)](#) to show the existence of regional favoritism. I restrict the dataset to the countries I use in my DHS sample. I run the following regression:

$$Light_{ict} = \alpha_{ic} + \lambda_{ct} + \beta_1 Leader_{ict} \times Polity_{ct} + \beta_2 Leader_{ict} + \beta_3 Polity_{ct} + \epsilon_{ict} \quad (4)$$

Where  $Light_{ict}$  is the log of the average nighttime light intensity in ethnographic region  $i$  in country  $c$  at time  $t$ <sup>10</sup>. The variable  $\alpha_{ic}$  is a region-fixed effect that controls for regional characteristics that vary across time, like historical and climatic factors.  $\lambda_{ct}$  is a vector of country-year dummy variables that would serve as a control for regional shocks and changes to satellites and sensory deterioration. [Henderson et al. \(2012\)](#) document a strong correlation between a country’s GDP and nighttime light intensity. Therefore,  $Light_{ict}$  could reflect the economic condition in a country, which is associated with education, health and wealth—the outcomes of interest I am studying.

10. following [Henderson et al. \(2012\)](#); [Michalopoulos and Papaioannou \(2014, 2013\)](#); [Hodler and Raschky \(2014b,a\)](#), the nighttime intensity variable is in logs after adding 0.01 because the distribution of light is right-skewed.

**TABLE VI**  
INSTITUTIONS' EFFECTS ON ETHNIC  
FAVORITISM: USING LIGHT DATA.

	Ethnologue <i>Light<sub>ict</sub></i>	GREG <i>Light<sub>ict</sub></i>
<b>Panel A: Ethnic Favoritism</b>		
<i>Coethnic<sub>ict</sub></i>	0.061 (0.051)	0.097** (0.036)
<b>Panel B: Ethnic favoritism and institutions</b>		
<i>Coethnic<sub>ict</sub> × Polity<sub>ct</sub></i>	-0.076 (0.204)	-0.041 (0.131)

Using the data from [Hodler and Raschky \(2014b\)](#) allows me to detect regional favoritism when a leader would allocate resources to regions where her ethnic group lives. The “psychic effect” or the “Obama effect” will not influence the results of regional favoritism because when a leader is targeting one region, they cannot exclude specific ethnic groups that live in said region. The results to equation 4 are presented in table VI. In panel A, I present the results of the existence of regional favoritism, while in panel B I present the results when I interact the leader’s ethnicity with *Polity IV*. The results are similar to what I found in my analysis of the interaction between *Polity IV* and co-ethnicity.

## VI. CONCLUSION

Ethnic favoritism and inequitable distribution of resources could be holding back the growth of some African countries ([Easterly and Levine, 1997](#); [Alesina and Ferrara, 2005](#)). Institutions could also play a part in explaining why African countries lagged behind other nations ([Acemoglu et al., 2001](#); [Alsan, 2015](#)). This paper’s results further show that ethnic diversity and corruption could be one of the causes of Africa’s slow growth.

When studying the effect of ethnic favoritism, I find that people that spent all of their primary school when a leader was co-ethnic were 3.8 percentage points less likely to finish primary school, though the result was statistically insignificant. A child whose mom shared

the same ethnicity as the leader of the country she lived in two years before birth is 1.3 percentage points more likely to survive the first 12 months of birth. Co-ethnics were as likely as their non-co-ethnic peers to be in the same wealth quintile and have electricity, though the results are statistically insignificant. Finally, co-ethnics have better access to clean drinking water than non-co-ethnics.

When I allow for the heterogeneous effect of democracy, the interaction between democratic bins and co-ethnicity are statistically significant when studying their effect on electrification and access to clean drinking water in democracies and anocracies and schooling among anocracies. Co-ethnicity and more democracy do not eliminate ethnic favoritism when it comes to primary infant survival and access to clean drinking water. Co-ethnics in democracies are as likely to finish primary schooling and survive the first year of birth as their non-co-ethnic peers. Co-ethnics in anocracies are 3 percentage points more likely to finish primary school. However, Co-ethnics in democracies are 15.4 percentage points more likely and anocracies are 6.7 percentage points more likely to have electricity. People living under the rule of co-ethnic kin in democracies and anocracies are more likely to have cleaner drinking waters. The results I find are consistent with the empirical work of [Kraimon and Posner \(2013\)](#) and the theoretical work of [Francois et al. \(2015\)](#). Where evidence of ethnic favoritism depends on the outcome one studies and that autocracies in Africa are more ethnically representative than western democracies. Leaders aim to stay in power, to do so they will have to share the benefits of being a leader.

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## APPENDIX A: DATA APPENDIX

TABLE VII

THE NUMBER OF CO-ETHNIC AND NON-COETHNIC OBSERVATIONS IN EACH POLITY SCORE.

Polity Score	Not a Co-ethnic Leader	Co-ethnic Leader
-9	130,381	26,648
-8	27,207	1,613
-7	241,663	37,388
-6	37,936	3,701
-5	17,965	5,928
-4	23,396	1,564
-3	12,863	3,147
-2	9,870	1,338
-1	44,950	19,604
0	16,269	1,336
1	6,624	340
2	6,345	200
3	7,604	273
4	31,825	1,689
5	4,550	539
6	13,091	3,329
7	18,290	4,054
8	5,890	1,674
9	37	1,762



**TABLE VIII**  
THE NUMBER OF CO-ETHNIC AND NON-COETHNIC OBSERVATIONS IN POLITY GROUPS.

Polity Score	Not a Co-ethnic Leader	Co-ethnic Leader
Democracy	37,308	10,819
Open Anocracy	56,948	3,041
Closed Anocracy	125,313	32,917
Autocracy	437,187	69,350
Failed or Occupied	1,229	92

***A.1. Countries included in the data-set***

I used data from twenty-one countries. These countries and the distribution of observations are presented in table [IX](#).

**TABLE IX**  
NUMBER OF OBSERVATIONS BY COUNTRY.

Country	Number of observations	Percent
Angola	12,145	1.57
Cameroon	31,777	4.10
Congo Democratic Republic	28,224	3.65
Benin	41,969	5.42
Ethiopia	29,437	3.80
Ghana	27,571	3.56
Guinea	22,107	2.86
Cote d'Ivoire	18,403	2.38
Kenya	58,278	7.53
Malawi	68,081	8.79
Mali	42,372	5.47
Mozambique	18,179	2.35
Namibia	2,872	0.37
Niger	29,963	3.87
Nigeria	90,445	11.68
Senegal	92,990	12.01
South Africa	20,249	2.62
Zimbabwe	16,987	2.19
Uganda	47,739	6.17
Burkina Faso	36,834	4.76
Zambia	37,582	4.85
Total	774,204	100