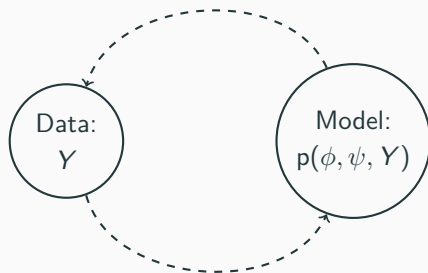


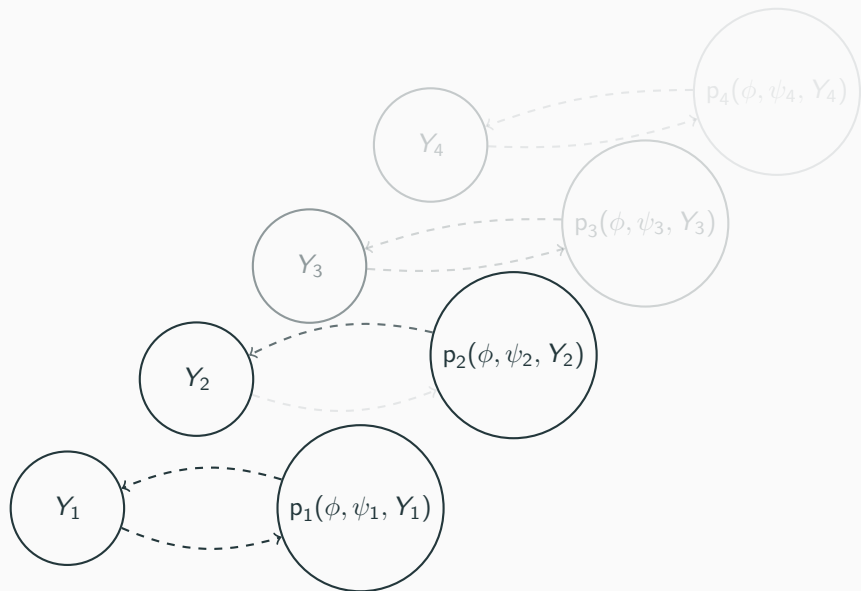
Markov melding, multi-stage sampling, and self-density ratios

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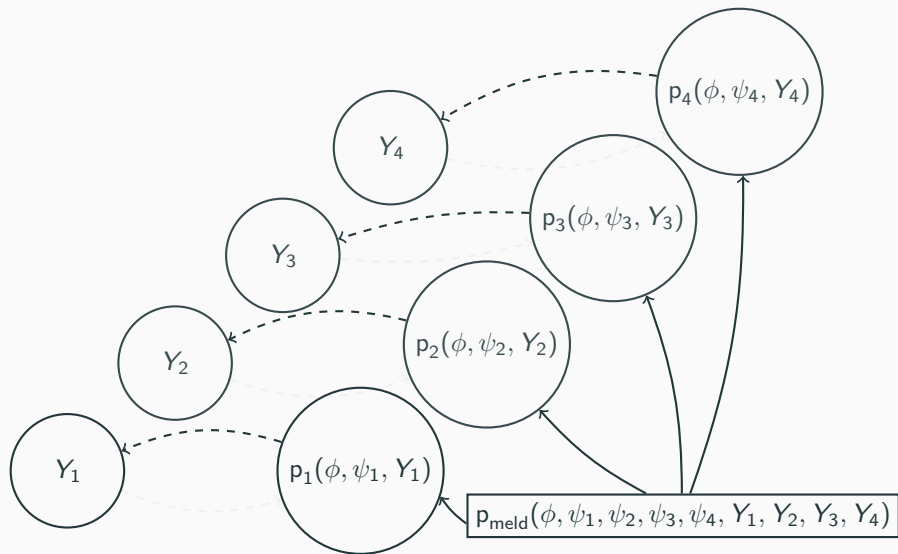
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Modern statistical modelling



Melding based statistical modelling?



Markov melding

- Markov melding (Goudie et al. 2019) is a method for joining models with common component: ϕ
- Ideally we would specify the generative model (For $M = 2$ submodels):

$$p_{\text{meld}}(\phi, \psi_1, \psi_2, Y_1, Y_2) = p_{\text{pool}}(\phi) p_1(\psi_1, Y_1 | \phi) p_2(\psi_2, Y_2 | \phi)$$

- Models developed in isolation, re-express as:

$$p_{\text{meld}}(\phi, \psi_1, \psi_2, Y_1, Y_2) = p_{\text{pool}}(\phi) \frac{p_1(\phi, \psi_1, Y_1)}{p_1(\phi)} \frac{p_2(\phi, \psi_2, Y_2)}{p_2(\phi)}$$

- Analytic form of $p_m(\phi)$ often unknown
 - Can cause numerical issues in the model fitting process
 - Requires accurate estimation in low probability regions
 - We will revisit

Multi-stage sampler

- Need to sample the melded posterior:

$$\begin{aligned} p_{\text{meld}}(\phi, \psi_1, \psi_2 \mid Y_1, Y_2) &\propto p_{\text{meld}}(\phi, \psi_1, \psi_2, Y_1, Y_2) \\ &= p_{\text{pool}}(\phi) \frac{p_1(\phi, \psi_1, Y_1)}{p_1(\phi)} \frac{p_2(\phi, \psi_2, Y_2)}{p_2(\phi)} \end{aligned}$$

- Submodels can be complex \rightarrow sample in stages
 - Multi-stage sampling & approximations are common in practice
 - Staged sampling targets components of $p_{\text{meld}}(\phi, \psi_1, \psi_2 \mid Y_1, Y_2)$ in a cumulative manner
 - Can reuse other implementations of $p_m(\phi, \psi_m, Y_m)$
 - Potentially more efficient

Stage one acceptance probability

- Say we choose to target, $p_{\text{meld},1}(\phi, \psi_1 \mid Y_1)$:

$$\begin{aligned} p_{\text{meld}}(\phi, \psi_1, \psi_2 \mid Y_1, Y_2) &\propto p_{\text{pool}}(\phi) \frac{p_1(\phi, \psi_1, Y_1)}{p_1(\phi)} \frac{p_2(\phi, \psi_2, Y_2)}{p_2(\phi)} \\ &\rightarrow p_{\text{meld},1}(\phi, \psi_1 \mid Y_1) \end{aligned}$$

- Given our stage one target, the stage one sampler has the following acceptance probability:

$$\alpha((\phi^*, \psi_1^*), (\phi, \psi_1)) = \frac{p_1(\phi^*, \psi_1^*, Y_1)p_1(\phi)Q(\phi, \psi_1 \mid \phi^*, \psi_1^*)}{p_1(\phi, \psi_1, Y_1)p_1(\phi^*)Q(\phi^*, \psi_1^* \mid \phi, \psi_1)}$$

- (asymptotically) Produces samples from the stage one target

Stage two acceptance probability

- Stage two target, the melded posterior:

$$p_{\text{meld}}(\phi, \psi_1, \psi_2 \mid Y_1, Y_2) \propto p_{\text{pool}}(\phi) \frac{p_1(\phi, \psi_1, Y_1)}{p_1(\phi)} \frac{p_2(\phi, \psi_2, Y_2)}{p_2(\phi)}$$

- Use stage one samples of ϕ as the proposal distribution in a Gibbs update for $\phi \mid \psi$:

$$\alpha(\phi^*, \phi) = \frac{p_{\text{pool}}(\phi^*)p_1(\phi^*, \psi_1, Y_1)p_2(\phi^*, \psi_2, Y_2)p_1(\phi)p_2(\phi)}{p_{\text{pool}}(\phi)p_1(\phi, \psi_1, Y_1)p_2(\phi, \psi_2, Y_2)p_1(\phi^*)p_2(\phi^*)} \frac{p_1(\phi, \psi_1, Y_1)p_1(\phi^*)}{p_1(\phi^*, \psi_1, Y_1)p_1(\phi)}$$

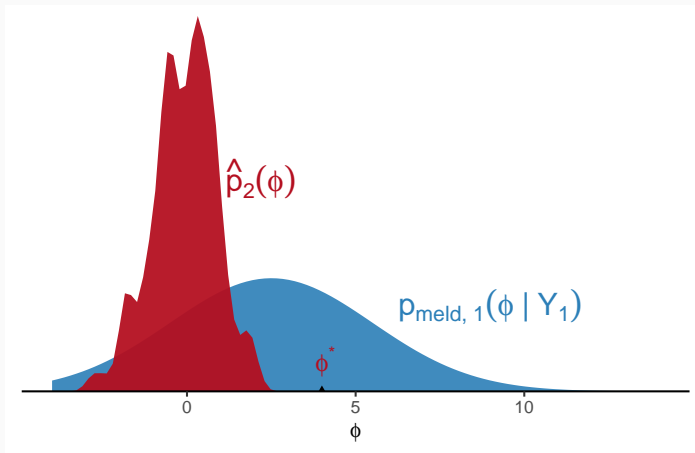
- Stage one terms cancel:

$$\alpha(\phi^*, \phi) = \frac{p_{\text{pool}}(\phi^*)p_2(\phi^*, \psi_2, Y_2)p_2(\phi)}{p_{\text{pool}}(\phi)p_2(\phi, \psi_2, Y_2)p_2(\phi^*)}$$

- $p_2(\phi)$ not known analytically \rightarrow substitute KDE: $\hat{p}_2(\phi)$

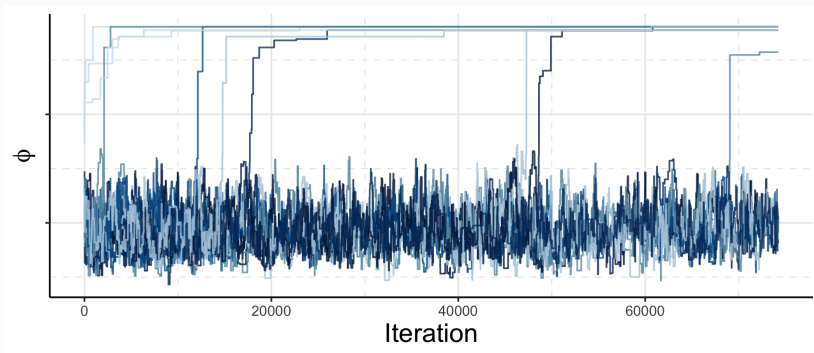
Conflict

$$\alpha(\phi^*, \phi) = \frac{p_{\text{pool}}(\phi^*)p_2(\phi^*, \psi_2, Y_2)}{p_{\text{pool}}(\phi)p_2(\phi, \psi_2, Y_2)} \cdot \frac{\hat{p}_2(\phi)}{\hat{p}_2(\phi^*)}$$



Conflict in action

$$\alpha(\phi^*, \phi) = \frac{p_{\text{pool}}(\phi^*)p_2(\phi^*, \psi_2, Y_2)}{p_{\text{pool}}(\phi)p_2(\phi, \psi_2, Y_2)} \cdot \frac{\hat{p}_2(\phi)}{\hat{p}_2(\phi^*)}$$



- H1N1 example from Goudie et al. (2019)

Self density ratios

We only interact with unknown marginal via self-density ratio (Hiraoka, Hamada, and Hori 2018):

$$r(\phi_{\text{nu}}, \phi_{\text{de}}) = \frac{p(\phi_{\text{nu}})}{p(\phi_{\text{de}})}$$

Weighted-sample self-density ratio estimation (WSRE) Intuition:

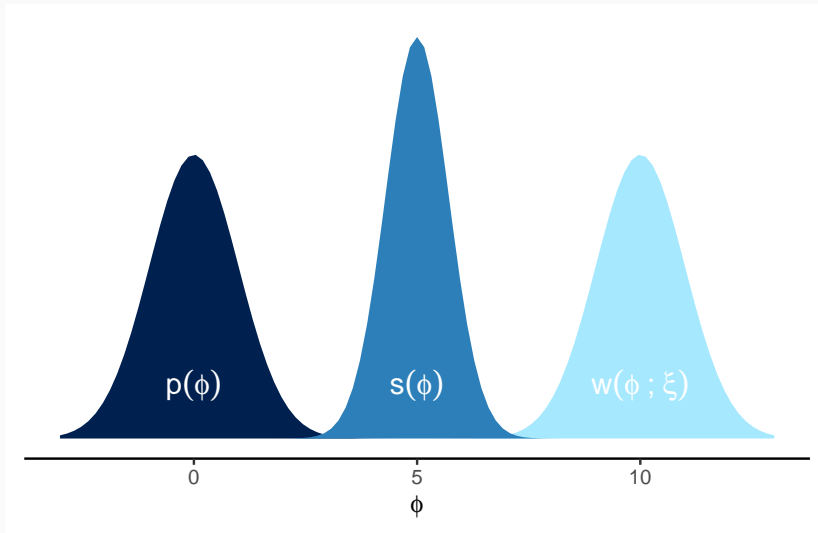
- Sampling $\phi \sim p(\phi, \psi, Y)$ admits $\phi \sim p(\phi)$
- Instead, sample:
 $\phi \sim p(\phi, \psi, Y)w(\phi; \xi) \rightarrow \phi \sim \frac{1}{Z}p(\phi)w(\phi; \xi) = s(\phi)$
- Use the weighted sample density estimator of Jones (1991):

$$\hat{p}(\phi) = \frac{1}{ZNh} \sum_{i=1}^N w(\phi; \xi)^{-1} K_h(\phi - \phi_i)$$

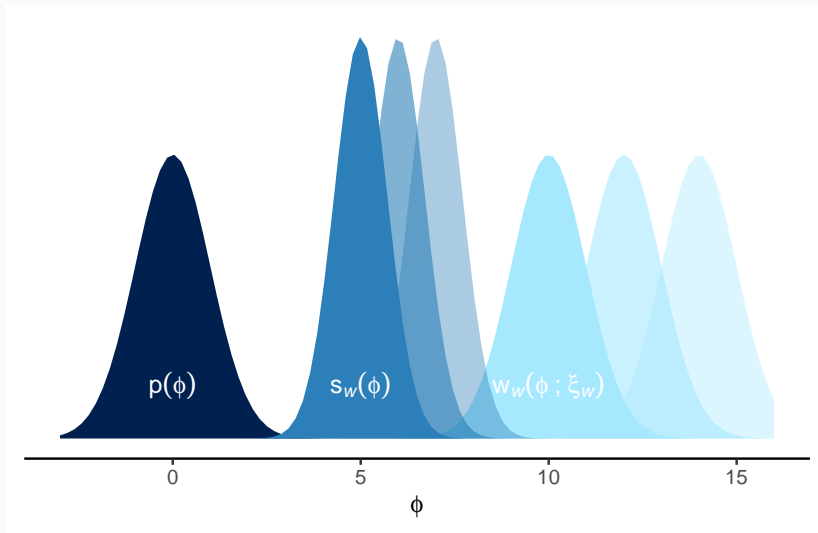
- Ratio estimator is then:

$$\hat{r}(\phi_{\text{nu}}, \phi_{\text{de}}) = \frac{\hat{p}(\phi_{\text{nu}})}{\hat{p}(\phi_{\text{de}})}$$

WSRE intuition - 1

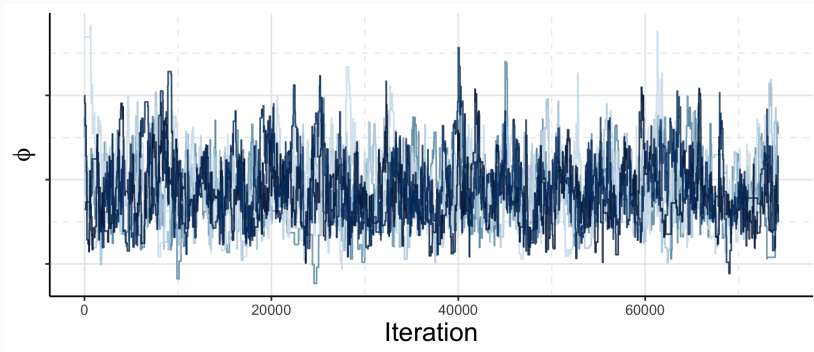


WSRE intuition - W



H1N1 example - 2

$$\alpha(\phi^*, \phi) = \frac{p_{\text{pool}}(\phi^*) p_2(\phi^*, \psi_2, Y_2)}{p_{\text{pool}}(\phi) p_2(\phi, \psi_2, Y_2)} \cdot \hat{r}(\phi, \phi^*)$$



Future plans

- Markov melding for ICU delirium, using electronic health record data
 - Observational data (Temperature, respiratory rate)
 - Blood test data
 - Daily ICU drug dose data
 - Each source is idiosyncratic \rightarrow requires substantial modelling effort (and domain expertise)
- Melding for multiple ϕ : $p_1(\phi_1, Y_1)$ $p_2(\phi_1, \phi_2, Y_2)$ $p_3(\phi_2, Y_3)$
- Addressing other kinds of internal difference in scale and location (conflict)

Should you also wish to estimate self-density ratios:

- <https://github.com/hhau/wsre>

The traceplots are from this example:

- <https://github.com/hhau/full-melding-example>

This talk:

- <https://github.com/hhau/bsu-together-20mins>

References

Goudie, Robert J. B., Anne M. Presanis, David Lunn, Daniela De Angelis, and Lorenz Wernisch. 2019. "Joining and Splitting Models with Markov Melding." *Bayesian Anal.* 14 (1). International Society for Bayesian Analysis: 81–109. doi:10.1214/18-BA1104.

Hiraoka, Kazuyuki, Toshihiko Hamada, and Gen Hori. 2018. "Necessary and Sufficient Conditions of Proper Estimators based on Self Density Ratio for Unnormalized Statistical Models." *Neural Networks* 98: 263–70. doi:10.1016/j.neunet.2017.11.018.

Jones, M. C. 1991. "Kernel Density Estimation for Length Biased Data." *Biometrika* 78 (3). Oxford University Press, Biometrika Trust: 511–19. <http://www.jstor.org/stable/2337020>.