

# Homework 1 solution

**Due: 2/4/2019 before 11pm. Submit in Canvas (file upload). Rmd file and the html output file (submit both files) are strongly recommended, but not required.**

## 1. (10 pts)

Let  $X = (X_1, \dots, X_n)'$  be the  $n \times p$  data matrix, where  $X_i = (X_{i1}, \dots, X_{ip})'$  is the  $i$ th onservation. Let  $\bar{X} = n^{-1} \sum_{i=1}^n X_i$  be the sample mean. Let  $s_{j_1 j_2} = n^{-1} \sum_{i=1}^n (X_{ij_1} - \bar{X}_{j_1})(X_{ij_2} - \bar{X}_{j_2})$  be the sample covariance between the  $j_1$ th and  $j_2$ th variables. Let  $S = (s_{j_1 j_2})$  be the sample covariance matrix. Show that

$$S = \frac{1}{n} X'X - \bar{X}'\bar{X}.$$

**Proof:** Let  $A = \frac{1}{n} X'X - \bar{X}'\bar{X}$ . The  $(j_1, j_2)$ th element of  $A$  is

$$a_{j_1, j_2} = \frac{1}{n} \sum_{i=1}^n X'_{j_1 i} X_{ij_2} - \bar{X}_{j_1} \bar{X}_{j_2} = \frac{1}{n} \sum_{i=1}^n X_{ij_1} X_{ij_2} - \bar{X}_{j_1} \bar{X}_{j_2}.$$

On the other hand,

$$s_{j_1 j_2} = \frac{1}{n} \sum_{i=1}^n (X_{ij_1} - \bar{X}_{j_1})(X_{ij_2} - \bar{X}_{j_2}) = \frac{1}{n} \sum_{i=1}^n X_{ij_1} X_{ij_2} - \frac{1}{n} \sum_{i=1}^n \bar{X}_{j_1} \bar{X}_{j_2} = \frac{1}{n} \sum_{i=1}^n X_{ij_1} X_{ij_2} - \bar{X}_{j_1} \bar{X}_{j_2}.$$

Therefore,  $a_{j_1, j_2} = s_{j_1 j_2}$ , and  $S = \frac{1}{n} X'X - \bar{X}'\bar{X}$ .

## 2. (10 pts)

Find ALL the eigenvalues and their eigenvectors for the following matrices

- $\Sigma = \sigma \mathbb{1} \mathbb{1}'$  where  $\mathbb{1} = (1, 1, \dots, 1)'$  is the  $p$  dimensional vector of 1.

**Solution:** Since  $\text{rank}(\Sigma) = 1$ , we have  $\ker(\Sigma) = p - 1$ . Thus,  $\lambda = 0$  is a root of the characteristic polynomial with multiplicity at least  $p - 1$ . On the other hand,  $\text{tr}(\Sigma) = p\sigma \neq 0$ . Thus,  $\Sigma$  has an eigenvalue  $p\sigma$ . So  $\Sigma$  has eigenvalue 0 with multiplicity  $p - 1$  and eigenvalue  $p\sigma$  with multiplicity 1. The eigenvectors corresponding to 0 solves the equation  $\{\mathbf{x} = (x_1, x_2, \dots, x_p)' | x_1 + x_2 + \dots + x_p = 0\}$ . One possible solution system would be

$$\left\{ \begin{bmatrix} 1 \\ -1 \\ 0 \\ \vdots \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 \\ 0 \\ -1 \\ \vdots \\ 0 \\ 0 \end{bmatrix}, \dots, \begin{bmatrix} 1 \\ 0 \\ 0 \\ \vdots \\ 0 \\ -1 \end{bmatrix} \right\}.$$

The eigenvector corresponding to  $p\sigma$  can be chosen as  $\mathbb{1}$ .

- $\Sigma = \text{diag}(\sigma_1, \dots, \sigma_p)$  is a diagonal matrix.

**Solution:** We have  $\Sigma = \text{diag}\{\sigma_1, \dots, \sigma_p\} = \sigma_1 \mathbf{e}_1 \mathbf{e}_1' + \sigma_2 \mathbf{e}_2 \mathbf{e}_2' + \dots + \sigma_p \mathbf{e}_p \mathbf{e}_p'$  where  $\mathbf{e}_i$  is the  $p$  dimensional vector with the  $i$ th coordinate being 1 and others being 0 for  $i = 1, 2, \dots, p$ . Since  $\{\mathbf{e}_i\}$  are orthogonal, the eigenvalues of  $\Sigma$  are  $\{\sigma_1, \sigma_2, \dots, \sigma_p\}$  and the corresponding eigenvectors are  $\{\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_p\}$ .

### 3. (10 pts)

Show that  $\text{tr}(AB) = \text{tr}(BA)$ .

**Proof:** Let  $A = (a_{ij})_{p \times q}$  and  $B = (b_{ij})_{q \times p}$ . (Note that  $A$  and  $B$  don't require to be square matrices.) Without loss generality, assume  $p \leq q$ .

$$\text{tr}(AB) = \sum_{i=1}^p \sum_{j=1}^q a_{ij} b_{ji} = \sum_{j=1}^q \sum_{i=1}^p b_{ji} a_{ij} = \text{tr}(BA).$$

### 4. (10 pts)

Given two variables in the data matrix  $X$  (say the  $j_1$ th and  $j_2$ th variables). Show that their sample correlation will not change by standardization.

**Proof:** Let  $x_{(j_1)}, x_{(j_2)}$  be the  $j_1$ th and  $j_2$ th variables and  $\tilde{x}_{(j_1)}, \tilde{x}_{(j_2)}$  be the standardization of  $x_{(j_1)}$  and  $x_{(j_2)}$ . Then,  $\tilde{x}_{(j_1)} = \frac{1}{\sqrt{s_{j_1 j_1}}}(x_{(j_1)} - \bar{x}_{j_1} \mathbf{1})$  and  $\tilde{x}_{(j_2)} = \frac{1}{\sqrt{s_{j_2 j_2}}}(x_{(j_2)} - \bar{x}_{j_2} \mathbf{1})$ . The sample correlation between  $\tilde{x}_{(j_1)}$  and  $\tilde{x}_{(j_2)}$  is

$$\begin{aligned} \tilde{r}_{j_1 j_2} &= \frac{1}{n} \sum_{i=1}^n \tilde{x}_{ij_1} \tilde{x}_{ij_2} = \frac{1}{n} \sum_{i=1}^n \frac{x_{ij_1} - \bar{x}_{j_1}}{\sqrt{s_{j_1 j_1}}} \frac{x_{ij_2} - \bar{x}_{j_2}}{\sqrt{s_{j_2 j_2}}} \\ &= \frac{1}{\sqrt{s_{j_1 j_1}} \sqrt{s_{j_2 j_2}}} \frac{1}{n} \sum_{i=1}^n (x_{ij_1} - \bar{x}_{j_1})(x_{ij_2} - \bar{x}_{j_2}) \\ &= \frac{s_{j_1 j_2}}{\sqrt{s_{j_1 j_1}} \sqrt{s_{j_2 j_2}}} = r_{j_1 j_2}. \end{aligned}$$

### 5. (10 pts)

Given a data matrix  $X$  as in Question 1. Assume the means of the  $p$  variables are zero. Let  $S = \frac{1}{n} X'X$  be the sample covariance matrix. Let  $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_p$  be the ordered eigenvalues of  $S$ . Let  $e_1, \dots, e_p$  be their corresponding orthogonal eigenvectors with unit length. In multivariate analysis, we usually want to use the first few eigenvalues and eigenvectors to represent the original data, as a tool of dimension reduction.

- On one aspect, let

$$S_m = \lambda_1 e_1 e_1' + \dots + \lambda_m e_m e_m'$$

be an approximate of  $S$  for  $m < p$ . Calculate  $\text{tr}\{(S - S_m)^2\}$  and  $\text{tr}\{(S - S_m)^2\}/\text{tr}(S^2)$ , where  $\text{tr}(S^2)$  can be regarded as the total variation of the data.

**Solution:** Note that  $\text{tr}\{(S - S_m)^2\} = \text{tr}(S^2) + \text{tr}(S_m^2) - 2\text{tr}(SS_m)$ , and  $\text{tr}(S^2) = \sum_{i=1}^p \lambda_i^2$ ,  $\text{tr}(S^2) = \sum_{i=1}^m \lambda_i^2$ . We also have  $\text{tr}(SS_m) = \text{tr}(S_m^2)$ . Therefore,

$$\frac{\text{tr}\{(S - S_m)^2\}}{\text{tr}(S^2)} = \frac{\sum_{i=m+1}^p \lambda_i^2}{\sum_{i=1}^p \lambda_i^2}.$$

- On another aspect,  $\{Xe_1, \dots, Xe_m\}$  are the transformed data by the eigenvectors. Calculate the sample covariance  $S_t$  of  $\{Xe_1, \dots, Xe_m\}$  (regard the sample mean as 0). What is  $\text{tr}(S_t^2)$  comparing to  $\text{tr}(S^2)$ ?

**Solution:** Note that  $\{Xe_1, \dots, Xe_m\} = X(e_1, \dots, e_m)$ . This leads to

$$\begin{aligned}
 S_t &= \frac{1}{n} (X(e_1, \dots, e_m))' X(e_1, \dots, e_m) \\
 &= \frac{1}{n} \begin{pmatrix} e'_1 \\ e'_2 \\ \vdots \\ e'_m \end{pmatrix} X' X(e_1, e_2, \dots, e_m) = \begin{pmatrix} e'_1 \\ e'_2 \\ \vdots \\ e'_m \end{pmatrix} S(e_1, e_2, \dots, e_m) \\
 &= \begin{pmatrix} e'_1 \\ e'_2 \\ \vdots \\ e'_m \end{pmatrix} (\lambda_1 e_1 e'_1 + \dots + \lambda_p e_p e'_p)(e_1, \dots, e_m) = \begin{pmatrix} \lambda_1 e'_1 \\ \lambda_2 e'_2 \\ \vdots \\ \lambda_m e'_m \end{pmatrix} (e_1, e_2, \dots, e_m) \\
 &= \text{diag}(\lambda_1, \dots, \lambda_m).
 \end{aligned}$$

Therefore,  $\text{tr}(S_t^2) = \sum_{i=1}^m \lambda_i^2 = \text{tr}(S_m^2)$ .

- What can you conclude on the dimension reduction by eigenvectors from the above two points?

**Solution:** If there is a few leading eigenvalues dominating the summation  $\sum_{i=1}^p \lambda_i^2$ , dimension reduction by eigenvectors can preserve most of the total variation of the original data while decrease the dimensions of the variables.

## 6. (10 pts)

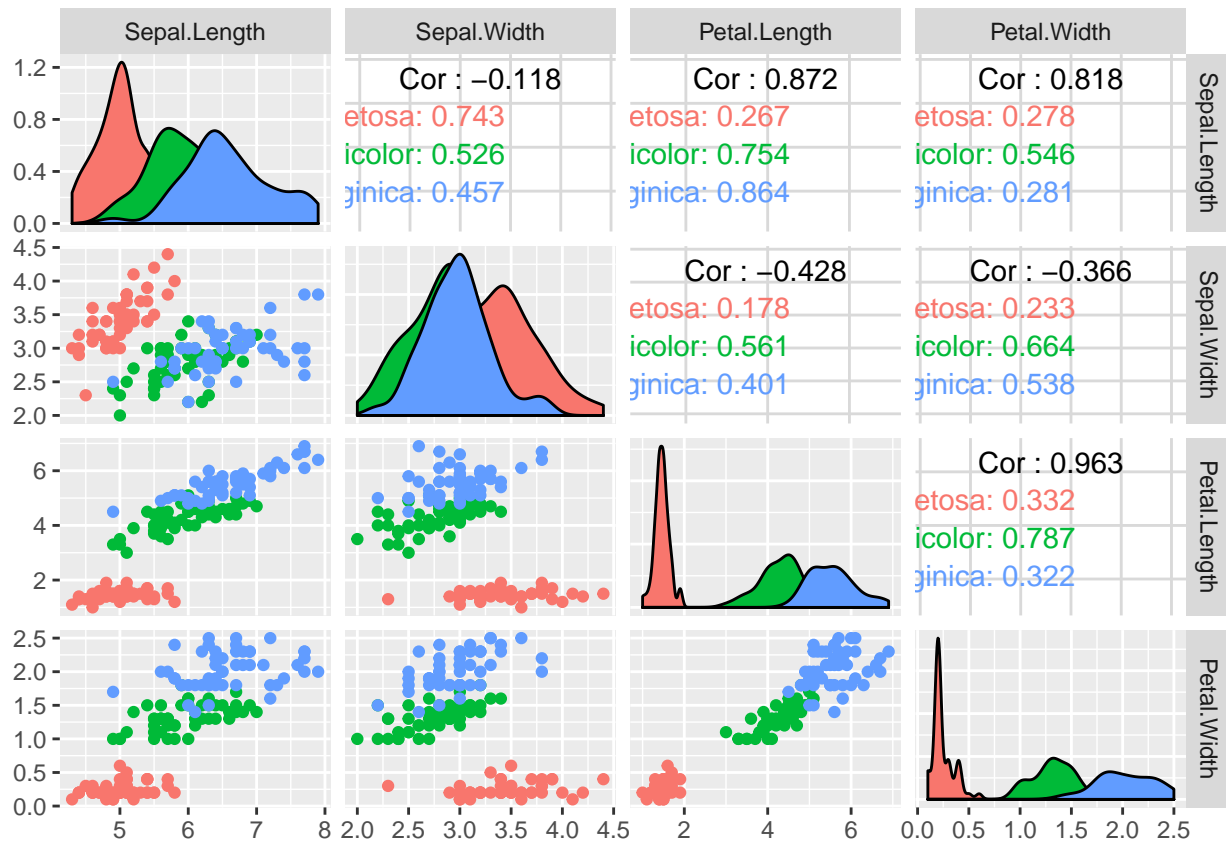
Use the **ggplot2** package to visualize the data `iris` in R.

- Make a scatter plot for the variables `Sepal.Length` and `Sepal.Width` colored by `Species`. What can you see?
- Make a scatter matrix for every pairs of the variables, colored by `Species`. What can you see?
- Calculate the sample mean for each species.
- Make a star plot for the sample means of each species to illustrate their potential differences. Comment.

```
library(tidyverse)
ggplot(data = iris, aes(x = Sepal.Length, y = Sepal.Width, colour = Species)) +
  geom_point()
```



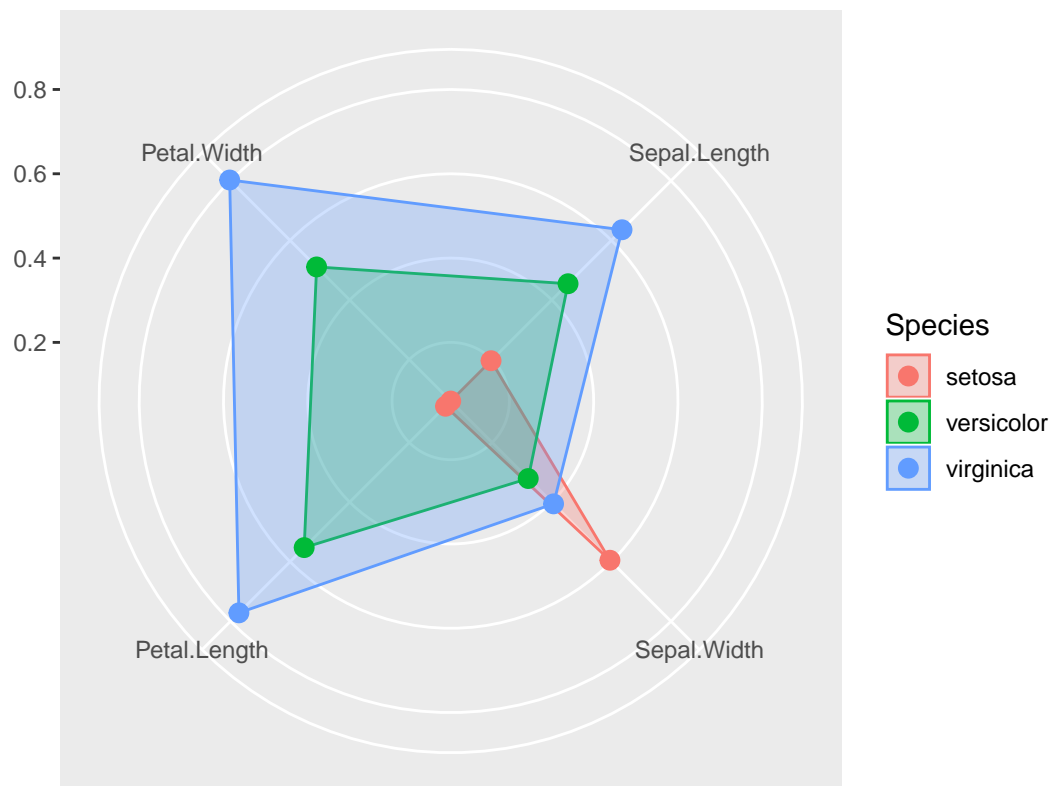
```
GGally::ggpairs(iris, columns = 1:4, aes(colour = Species))
```



```
iris %>% group_by(Species) %>% summarise_all(mean)
```

```
## # A tibble: 3 x 5
##   Species   Sepal.Length Sepal.Width Petal.Length Petal.Width
##   <fct>         <dbl>         <dbl>         <dbl>         <dbl>
## 1 setosa         5.01           3.43           1.46           0.246
## 2 versicolor    5.94           2.77           4.26           1.33
## 3 virginica     6.59           2.97           5.55           2.03
```

```
ggiraphExtra::ggRadar(iris, aes(colour = Species), legend.position = "right")
```



```
# Another star plot function.
# devtools::install_github("ricardo-bion/ggradar", dependencies = TRUE)
# iris_mean %>%
#   mutate_at(vars(-Species), scales::rescale) %>%
#   ggradar::ggradar(axis.label.size = 3, group.line.width = .5,
#                     group.point.size = 2, legend.text.size = 12)
```