# 机器学习统计学习方法

主讲: 蔡 波

武汉大学网络安全学院

## 第六章 补充

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### 【超平面的性质】

$$w^T(x_1-x_2)=0$$

$$x = x_{\rho} + r \frac{w}{\|w\|}$$

#### 式中:

xp--x在H上的投影

r--x到H的垂直距离

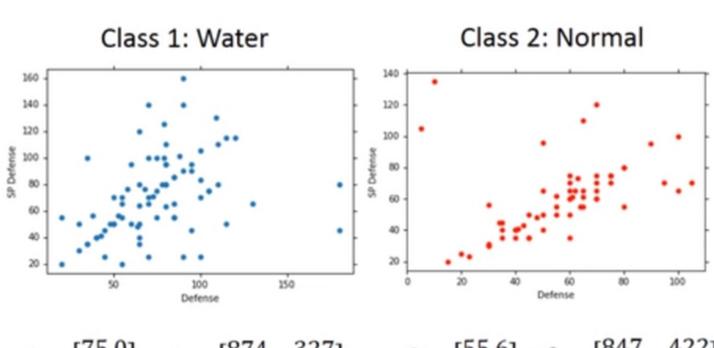
若x为原点,则: g (x) = w 0为原点到超平面的距离。

若 $w_0>0$ ,则原点在H的正侧;

若 $w_0$ <0,则原点在H的负侧;

若 $w_0$ =0,则H通过原点,具有齐次形式,超平面过原点。

#### **Maximum Likelihood**



$$\mu^1 = \begin{bmatrix} 75.0 \\ 71.3 \end{bmatrix} \quad \Sigma^1 = \begin{bmatrix} 874 & 327 \\ 327 & 929 \end{bmatrix} \qquad \mu^2 = \begin{bmatrix} 55.6 \\ 59.8 \end{bmatrix} \quad \Sigma^2 = \begin{bmatrix} 847 & 422 \\ 422 & 685 \end{bmatrix}$$

#### Now we can do classification ©

$$f_{\mu^{1},\Sigma^{1}}(x) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^{1}|^{1/2}} exp\left\{-\frac{1}{2}(x-\mu^{1})^{T}(\Sigma^{1})^{-1}(x-\mu^{1})\right\} = 79/(79+61) = 0.56$$

$$\mu^{1} = \begin{bmatrix} 75.0 \\ 71.3 \end{bmatrix} \quad \Sigma^{1} = \begin{bmatrix} 874 & 327 \\ 327 & 929 \end{bmatrix}$$

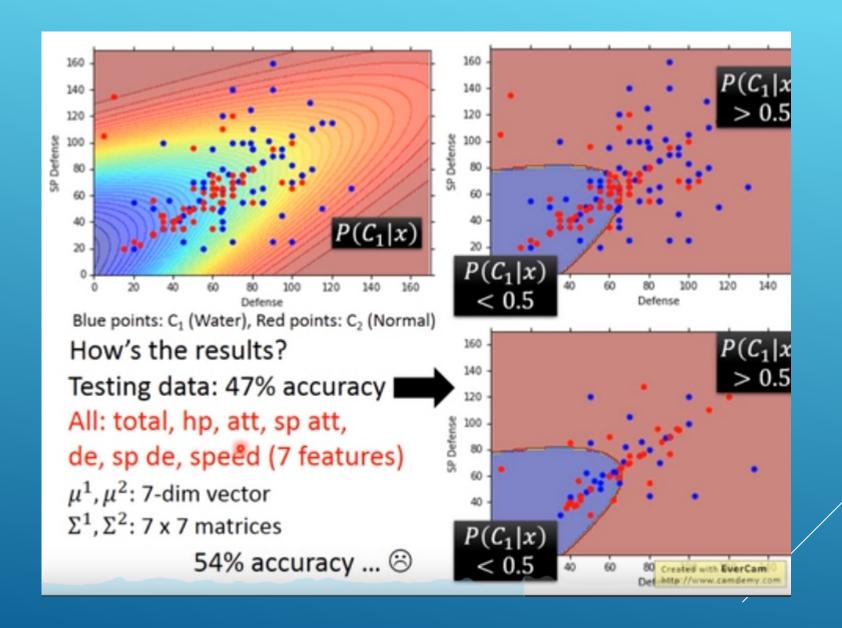
$$P(C_{1}|x) = \frac{P(x|C_{1})P(C_{1})}{P(x|C_{1})P(C_{1}) + P(x|C_{2})P(C_{2})}$$

$$f_{\mu^{2},\Sigma^{2}}(x) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^{2}|^{1/2}} exp\left\{-\frac{1}{2}(x-\mu^{2})^{T}(\Sigma^{2})^{-1}(x-\mu^{2})\right\} = \frac{9(C2)}{61/(79+61)}$$

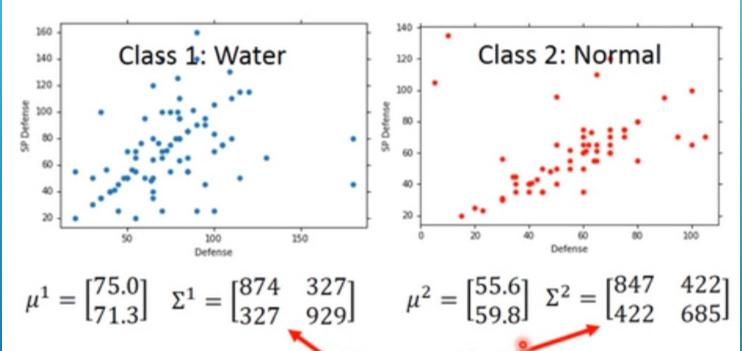
$$\mu^{2} = \begin{bmatrix} 55.6 \\ 59.8 \end{bmatrix} \quad \Sigma^{2} = \begin{bmatrix} 847 & 422 \\ 422 & 685 \end{bmatrix}$$



If  $P(C_1|x) > 0.5$   $\blacksquare$  x belongs to class 1 (Water)



#### Modifying Model



• The same  $\Sigma$ 

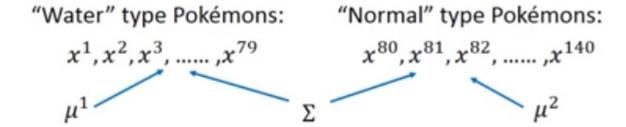
Less parameters

Created with EverCam.

#### Modifying Model

Ref: Bishop, chapter 4.2.2

#### Maximum likelihood



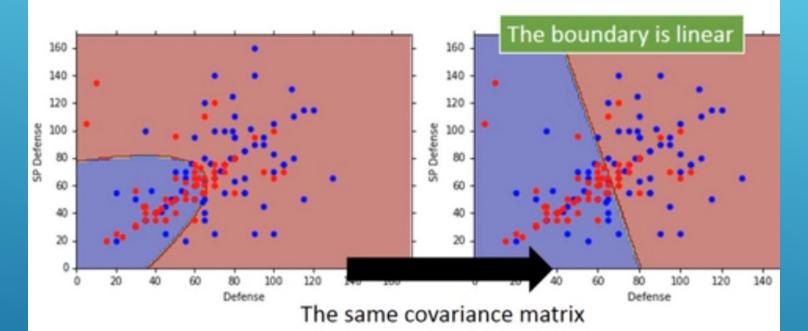
Find  $\mu^1$ ,  $\mu^2$ ,  $\Sigma$  maximizing the likelihood  $L(\mu^1, \mu^2, \Sigma)$ 

$$\begin{split} L(\mu^1, & \mu^2, \Sigma) = f_{\mu^1, \Sigma}(x^1) f_{\mu^1, \Sigma}(x^2) \cdots f_{\mu^1, \Sigma}(x^{79}) \\ & \times f_{\mu^2, \Sigma}(x^{80}) f_{\mu^2, \Sigma}(x^{81}) \cdots f_{\mu^2, \Sigma}(x^{140}) \end{split}$$

$$\mu^1$$
 and  $\mu^2$  is the same  $\Sigma = \frac{79}{140} \Sigma^1 + \frac{61}{140} \Sigma^2$ 



#### Modifying Model



All: total, hp, att, sp att, de, sp de, speed



#### Three Steps

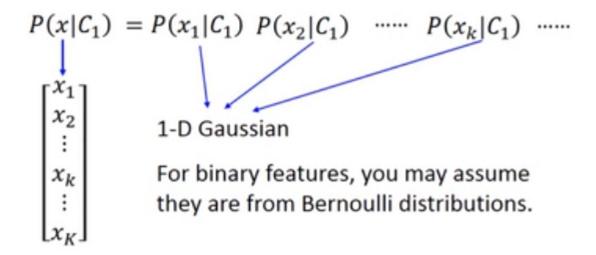
Function Set (Model):

$$P(C_1|x) = \frac{P(x|C_1)P(C_1)}{P(x|C_1)P(C_1) + P(x|C_2)P(C_2)}$$
If  $P(C_1|x) > 0.5$ , output: class 1
Otherwise, output: class 2

- · Goodness of a function:
  - The mean  $\mu$  and covariance  $\Sigma$  that maximizing the likelihood (the probability of generating data)
- Find the best function: easy

#### **Probability Distribution**

• You can always use the distribution you like ©



If you assume all the dimensions are independent, then you are using Naive Bayes Classifier.

#### Posterior Probability

$$P(C_{1}|x) = \frac{P(x|C_{1})P(C_{1})}{P(x|C_{1})P(C_{1}) + P(x|C_{2})P(C_{2})}$$

$$= \frac{1}{1 + \frac{P(x|C_{2})P(C_{2})}{P(x|C_{1})P(C_{1})}} = \frac{1}{1 + exp(-z)} = \sigma(z)$$
Sigmoid function
$$z = ln \frac{P(x|C_{1})P(C_{1})}{P(x|C_{2})P(C_{2})}$$

#### Posterior Probability

$$P(C_1|x) = \sigma(z)$$
 sigmoid  $z = ln \frac{P(x|C_1)P(C_1)}{P(x|C_2)P(C_2)}$ 

$$z = ln \frac{P(x|C_1)}{P(x|C_2)} + ln \frac{P(C_1)}{P(C_2)} \longrightarrow \frac{\frac{N_1}{N_1 + N_2}}{\frac{N_2}{N_1 + N_2}} = \frac{N_1}{N_2}$$

$$P(x|C_1) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^1|^{1/2}} exp\left\{-\frac{1}{2} (x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1)\right\}$$

$$P(x|C_2) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^2|^{1/2}} exp \left\{ -\frac{1}{2} (x - \mu^2)^T (\Sigma^2)^{-1} (x - \mu^2) \right\}$$

$$z = \ln \frac{P(x|C_1)}{P(x|C_2)} + \ln \frac{P(C_1)}{P(C_2)} = \frac{N_1}{N_2}$$

$$P(x|C_1) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^1|^{1/2}} exp \left\{ -\frac{1}{2} (x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1) \right\}$$

$$P(x|C_2) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^2|^{1/2}} exp \left\{ -\frac{1}{2} (x - \mu^2)^T (\Sigma^2)^{-1} (x - \mu^2) \right\}$$

$$\ln \frac{1}{(2\pi)^{D/2}} \frac{1}{|\Sigma^1|^{1/2}} exp \left\{ -\frac{1}{2} (x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1) \right\}$$

$$= \ln \frac{|\Sigma^2|^{1/2}}{|\Sigma^1|^{1/2}} exp \left\{ -\frac{1}{2} (x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^2) \right\}$$

$$= \ln \frac{|\Sigma^2|^{1/2}}{|\Sigma^1|^{1/2}} exp \left\{ -\frac{1}{2} [(x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1) - (x - \mu^2)^T (\Sigma^2)^{-1} (x - \mu^2)] \right\}$$

$$= \ln \frac{|\Sigma^2|^{1/2}}{|\Sigma^1|^{1/2}} - \frac{1}{2} [(x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1) - (x - \mu^2)^T (\Sigma^2)^{-1} (x - \mu^2)]$$
Constituting the form constant co

$$z = \ln \frac{P(x|C_1)}{P(x|C_2)} + \ln \frac{P(C_1)}{P(C_2)} = \frac{N_1}{N_2}$$

$$= \ln \frac{|\Sigma^2|^{1/2}}{|\Sigma^1|^{1/2}} - \frac{1}{2} [(x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1) - (x - \mu^2)^T (\Sigma^2)^{-1} (x - \mu^2)]$$

$$(x - \mu^1)^T (\Sigma^1)^{-1} (x - \mu^1)$$

$$= x^T (\Sigma^1)^{-1} x - x^T (\Sigma^1)^{-1} \mu^1 - (\mu^1)^T (\Sigma^1)^{-1} x + (\mu^1)^T (\Sigma^1)^{-1} \mu^1$$

$$= x^T (\Sigma^1)^{-1} x - 2(\mu^1)^T (\Sigma^1)^{-1} x + (\mu^1)^T (\Sigma^1)^{-1} \mu^1$$

$$(x - \mu^2)^T (\Sigma^2)^{-1} (x - \mu^2)$$

$$= x^T (\Sigma^2)^{-1} x - 2(\mu^2)^T (\Sigma^2)^{-1} x + (\mu^2)^T (\Sigma^2)^{-1} \mu^2$$

$$z = \ln \frac{|\Sigma^2|^{1/2}}{|\Sigma^1|^{1/2}} - \frac{1}{2} x^T (\Sigma^1)^{-1} x + (\mu^1)^T (\Sigma^1)^{-1} x - \frac{1}{2} (\mu^1)^T (\Sigma^1)^{-1} \mu^1$$

$$+ \frac{1}{2} x^T (\Sigma^2)^{-1} x - (\mu^2)^T (\Sigma^2)^{-1} x + \frac{1}{2} (\mu^2)^T (\Sigma^2)^{-1} \mu^2 + \ln \frac{N_1}{N_2}$$

$$P(C_1|x) = \sigma(z)$$

$$\begin{split} z &= ln \frac{|\Sigma^2|^{1/2}}{|\Sigma^1|^{1/2}} - \frac{1}{2} x^T (\Sigma^1)^{-1} x + (\mu^1)^T (\Sigma^1)^{-1} x - \frac{1}{2} (\mu^1)^T (\Sigma^1)^{-1} \mu^1 \\ &+ \frac{1}{2} x^T (\Sigma^2)^{-1} x - (\mu^2)^T (\Sigma^2)^{-1} x + \frac{1}{2} (\mu^2)^T (\Sigma^2)^{-1} \mu^2 + ln \frac{N_1}{N_2} \end{split}$$

$$\Sigma_1 = \Sigma_2 = \Sigma$$

$$z = \underbrace{(\mu^1 - \mu^2)^T \Sigma^{-1} x}_{\bullet} - \underbrace{\frac{1}{2} (\mu^1)^T (\Sigma^1)^{-1} \mu^1 + \frac{1}{2} (\mu^2)^T (\Sigma^2)^{-1} \mu^2 + ln \frac{N_1}{N_2}}_{b}$$

$$P(C_1|x) = \sigma(w \cdot x + b)$$
 How about directly find **w** and b?

In generative model, we estimate  $N_1$ ,  $N_2$ ,  $\mu^1$ ,  $\mu^2$ ,  $\Sigma$ 

Then we have w and b

softmax函数,也称指数归一化函数,它是一种logistic 函数的归一化,可以将[公式]维实数向量压缩成范围(0~1)的[公式]维实数向量函数形式为

$$\sigma(z)_j = rac{e^{z_j}}{\Sigma_{k=1}^K e^{z_k}}$$

其中分母指归一化的作用,取指数的原因,第一是模拟 max的行为,即使得大的数值更大,第二是,方便运算求 导

https://www.codercto.com/a/47514.html

#### Softmax回归

- = softmax回归是logistic回归的一般化,适用于 $\underline{K}$ 分类的问题,第k类的参数为向量 $\theta_k$ ,组成的二维矩阵为 $\theta_{k*n}$ ;
- ■softmax函数的本质就是将一个K维的任意实数向量压缩(映射)成另一个K维的实数向量,其中向量中的每个元素取值都介于(0,1)之间。
- ■softmax回归概率函数为:

$$p(y = k \mid x; \theta) = \frac{e^{\theta_k^T x}}{\sum_{l=1}^K e^{\theta_l^T x}}, k = 1, 2 \dots, K$$

#### Softmax算法原理

$$p(y = k \mid x; \theta) = \frac{e^{\theta_k^T x}}{\sum_{l=1}^K e^{\theta_l^T x}}, k = 1, 2 \dots, K$$

$$h_{\theta}(x) = \begin{bmatrix} p(y^{(i)} = 1 \mid x^{(i)}; \theta) \\ p(y^{(i)} = 2 \mid x^{(i)}; \theta) \\ \dots \\ p(y^{(i)} = k \mid x^{(i)}; \theta) \end{bmatrix} = \frac{1}{\sum_{j=1}^{k} e^{\theta_{j}^{T} x^{(i)}}} \begin{bmatrix} e^{\theta_{1}^{T} x} \\ e^{\theta_{2}^{T} x} \\ \dots \\ e^{\theta_{k}^{T} x} \end{bmatrix} \Longrightarrow \theta = \begin{bmatrix} \theta_{11} & \theta_{12} & \dots & \theta_{1n} \\ \theta_{21} & \theta_{22} & \dots & \theta_{2n} \\ \dots & \dots & \dots & \dots \\ \theta_{k1} & \theta_{k2} & \dots & \theta_{kn} \end{bmatrix}$$

#### Softmax算法损失函数

$$J(\theta) = -\frac{1}{m} \sum_{i=1}^{m} \sum_{j=1}^{k} I(y^{(i)} = j) \ln \left( \frac{e^{\theta_{j}^{T} x^{(i)}}}{\sum_{l=1}^{k} e^{\theta_{l}^{T} x^{(i)}}} \right) \qquad I(y^{(i)} = j) = \begin{cases} 1, & y^{(i)} = j \\ 0, & y^{(i)} \neq j \end{cases}$$

#### Softmax算法梯度下降法求解

$$J(\theta) = -\frac{1}{m} \sum_{i=1}^{m} \sum_{j=1}^{k} I(y^{(i)} = j) \ln \left(\frac{e^{\theta_{j}^{T} x^{(i)}}}{\sum_{l=1}^{k} e^{\theta_{l}^{T} x^{(i)}}}\right)$$

$$\frac{\partial}{\partial \theta_{i}} J(\theta) = \frac{\partial}{\partial \theta_{j}} - I(y^{(i)} = j) \ln \left(\frac{e^{\theta_{j}^{T} x^{(i)}}}{\sum_{l=1}^{k} e^{\theta_{l}^{T} x^{(i)}}}\right)$$

$$I(y^{(i)} = j) = \begin{cases} 1, & y^{(i)} = j \\ 0, & y^{(i)} \neq j \end{cases}$$

$$\frac{\partial}{\partial \theta_{i}} J(y^{(i)} = j) = \begin{cases} 1, & y^{(i)} = j \\ 0, & y^{(i)} \neq j \end{cases}$$

$$= \frac{\partial}{\partial \theta_{j}} - I(y^{(i)} = j) \left( \theta_{j}^{T} x^{(i)} - \ln \left( \sum_{l=1}^{k} e^{\theta_{l}^{T} x^{(i)}} \right) \right)$$

$$= -I(y^{(i)} = j) \left(1 - \frac{e^{\theta_j^T x^{(i)}}}{\sum_{l=1}^k e^{\theta_l^T x^{(i)}}}\right) x^{(i)}$$

#### Softmax算法梯度下降法求解

$$\frac{\partial}{\partial \theta_{j}} J(\theta) = -I(y^{(i)} = j) \left( 1 - \frac{e^{\theta_{j}^{T} x^{(i)}}}{\sum_{l=1}^{k} e^{\theta_{l}^{T} x^{(i)}}} \right) x^{(i)}$$

$$\theta_j = \theta_j + \alpha \sum_{i=1}^m I(y^{(i)} = j) (1 - p(y^{(i)} = j | x^{(i)}; \theta)) x^{(i)}$$

$$\theta_j = \theta_j + \alpha I(y^{(i)} = j)(1 - p(y^{(i)} = j|x^{(i)}; \theta))x^{(i)}$$

