# High Dimensional Probability Notes

### Nong Minh $\mathrm{Hieu^1}$

 $^{1}$  School of Physical and Mathematical Sciences, Nanyang Technological University (NTU - Singapore)

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#### 1 Random variables

#### 1.1 Basic inequalities

First, we revisit the definition of a random variable as well as some basic inequalities that we learned in introductory statistics.

#### Definition 1.1.

Random variable Let  $(\Omega, \Sigma.\mathbb{P})$  be a probability space. A random variable X is defined as a mapping from the sample space  $\Omega$  to  $\mathbb{R}$ :

$$X: \Omega \to \mathbb{R} \tag{1}$$

 $\Sigma$  is the  $\sigma$ -algebra containing the possible events (collection of subsets of  $\Omega$ ) and  $\mathbb{P}$  is a probability measure that assigns events with probabilities:

$$\mathbb{P}: \Sigma \to [0, 1] \tag{2}$$

For a given probability space  $(\Omega, \Sigma, \mathbb{P})$  and a random variable  $X : \Omega \to \mathbb{R}$ , we will use the following basic notations throughout this note:

•  $||X||_{L^p}$  - The  $p^{th}$  root of the  $p^{th}$  moment of the random variable X.

$$||X||_{L^p} = (\mathbb{E}|X|^p)^{1/p}, \ p \in (0, \infty)$$
 (3)

$$||X||_{L^{\infty}} = \operatorname{ess\,sup}|X| \tag{4}$$

•  $L^p(\Omega, \Sigma, \mathbb{P})$  - The space of random variables X satisfying:

$$L^{p}(\Omega, \Sigma, \mathbb{P}) = \left\{ X : \Omega \to \mathbb{R} \middle| \|X\|_{L^{p}} < \infty \right\}$$
 (5)

Some basic inequalities and identities:

• 1. Jensen's Inequality - For a random variable X and a convex function  $\varphi : \mathbb{R} \to \mathbb{R}$ , we have:

$$\varphi(\mathbb{E}X) \leqslant \mathbb{E}\varphi(X) \tag{6}$$

• 2. Monotonicity of  $L^p$  norm - For a random variable X:

$$||X||_{L^p} \leqslant ||X||_{L^q}, \ 0 \leqslant p \leqslant q \leqslant \infty \tag{7}$$

• 3. Minkowski's Inequality - For  $1 \le p \le \infty$  and two random variables X, Y in  $L^p(\Omega, \Sigma, \mathbb{P})$  space:

$$||X + Y||_{L^p} \le ||X||_{L^p} + ||Y||_{L^p} \tag{8}$$

• 4. Holder's Inequality - For  $p, q \in [1, \infty]$  such that 1/p + 1/q = 1. Then, for random variables  $X \in L^p(\Omega, \Sigma, \mathbb{P})$  and  $Y \in L^q(\Omega, \Sigma, \mathbb{P})$ , we have:

$$|\mathbb{E}XY| \leqslant ||X||_{L^p} \cdot ||Y||_{L^q} \tag{9}$$

• 5. Markov's Inequality - For a non-negative random variable X and t > 0, we have:

$$\mathbb{P}(X \geqslant t) \leqslant \frac{\mathbb{E}X}{t} \tag{10}$$

• 6. Chebyshev's Inequality - For a random variable X with mean  $\mu$  and variance  $\sigma^2$ . Then, for any t > 0, we have:

$$\mathbb{P}(|X - \mu| \ge t) \le \frac{\sigma^2}{t^2} \tag{11}$$

• 7. Integral Identity - Let X be a non-negative random variable, we have:

$$\mathbb{E}X = \int_0^\infty \mathbb{P}(X > t)dt \tag{12}$$

#### **Exercises**

#### Exercise 1.1.1: Generalized Integral Identity

Let X be a random variable (not necessarily non-negative). Prove the following identity:

$$\mathbb{E}X = \int_0^\infty \mathbb{P}(X > t)dt - \int_{-\infty}^0 \mathbb{P}(X < t)dt \tag{13}$$

Solution (Exercise 1.1.1).

For  $x \in \mathbb{R}$ , using the basic integral indentity, we have:

$$|x| = \int_0^\infty \mathbf{1}\{t < |x|\}dt$$

We consider the following cases:

• When  $x < 0 \implies x = -|x|$ :

$$x = -\int_0^\infty \mathbf{1}\{t < |x|\}dt = -\int_0^\infty \mathbf{1}\{t < -x\}dt = -\int_0^\infty \mathbf{1}\{-t > x\}dt = -\int_{-\infty}^0 \mathbf{1}\{t > x\}dt$$

• When  $x \ge 0 \implies x = |x|$ :

$$x = \int_0^\infty \mathbf{1}\{t < |x|\}dt = \int_0^\infty \mathbf{1}\{t < x\}dt$$

Therefore, for  $x \in \mathbb{R}$ , we can write:

$$x = \int_0^\infty \mathbf{1}\{t < x\}dt - \int_{-\infty}^0 \mathbf{1}\{t > x\}dt$$

Therefore, for a random variable X not necessarily non-negative, we have:

$$\begin{split} \mathbb{E}X &= \mathbb{E}\Bigg[\int_0^\infty \mathbf{1}\{t < X\}dt - \int_{-\infty}^0 \mathbf{1}\{t > X\}dt\Bigg] \\ &= \mathbb{E}\int_0^\infty \mathbf{1}\{t < X\}dt - \mathbb{E}\int_{-\infty}^0 \mathbf{1}\{t > X\}dt \\ &= \int_0^\infty \mathbb{E}\mathbf{1}\{t < X\}dt - \int_{-\infty}^0 \mathbb{E}\mathbf{1}\{t > X\}dt \\ &= \int_0^\infty \mathbb{P}(t < X)dt - \int_{-\infty}^0 \mathbb{P}(t > X)dt \end{split}$$

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#### Exercise 1.1.2: $p^{th}$ -moments via tails

Let X be a random variable and  $p \in (0, \infty)$ . Show that:

$$\mathbb{E}|X|^p = \int_0^\infty pt^{p-1}\mathbb{P}(|X| > t)dt \tag{14}$$

Solution (Exercise 1.1.2).

Let X be a random variable that is not necessarily non-negative. Using the integral identity, we have:

$$\mathbb{E}|X|^p = \int_0^\infty \mathbb{P}(u < |X|^p) du$$

Let  $t^p = u \implies pt^{p-1}dt = du$ . Since we integrate u from  $0 \to \infty$ , we also integrate t from  $0 \to \infty$  when changing the variables. Hence, we have:

$$\mathbb{E}|X|^p = \int_0^\infty \mathbb{P}(t^p < |X|^p) pt^{p-1} dt = \int_0^\infty \mathbb{P}(t < |X|) pt^{p-1} dt$$

Hence, we obtained the desired identity.

#### 1.2 Limit Theorems

#### Theorem 1.1: Strong Law of Large Numbers

Let  $X_1, \ldots, X_N$  be *i.i.d* random variables with mean  $\mu$ . Consider the sum:

$$S_N = X_1 + \dots + X_N \tag{15}$$

□.

Then, the sample mean converges to  $\mu$  almost surely  $(S_N/N \xrightarrow{a.s} \mu)$ :

$$\mathbb{P}\left(\limsup_{N\to\infty}|S_N/N-\mu|>\epsilon\right)=0, \ \forall \epsilon>0$$
 (16)

**Proof** (Strong Law of Large Numbers).

We revisit the **Borel-Cantelli** lemma: Given a sequence of events  $\{E_n\}_{n=1}^{\infty}$  such that its sum of probabilities  $\sum_{n=1}^{\infty} \mathbb{P}(E_n) < \infty$ . Then, we have:

$$P\Big(\limsup_{n\to\infty} E_n\Big) = 0 \tag{17}$$

Let  $\epsilon > 0$  be given. Denote the sequence of events  $\{E_n\}_{n=1}^{\infty}$  as follows:

$$E_n = \left\{ |S_n/n - \mu| > \epsilon \right\}$$

One of our goal in this proof is to prove that  $\sum_{n=1}^{\infty} \mathbb{P}(E_n) < \infty$  to apply the Borel-Cantelli lemma.

**1.** Attempt 1: Suppose that  $X_i$  have finite variance. Denote  $\bar{S}_n = S_n/n$ . Then,  $\bar{S}_n$  is a random variable with  $\mathbb{E}\bar{S}_n = \mu$  and  $Var(\bar{S}_n) = \sigma^2/n$ . For  $n \ge 1$ , we have:

$$\mathbb{P}(E_n) = \mathbb{P}\left(|\bar{S}_n - \mu| > \epsilon\right) \leqslant \frac{\sigma^2}{n\epsilon^2} \quad (Chebyshev's Inequality)$$

However, we cannot use the above inequality because we can only conclude:

$$\sum_{n=1}^{\infty} \mathbb{P}(E_n) \leqslant \frac{\sigma^2}{\epsilon^2} \sum_{n=1}^{\infty} n^{-1} \quad (Divergent \ sum)$$

Hence, we move on to our next attemp.

**2.** Attempt 2: Now, we change the strategy to bound  $\mathbb{P}(E_n)$ . Specifically, for any s > 0, using the Chernoff bound, we have:

$$\begin{split} P\Big(|\bar{S}_n - \mu| > \epsilon\Big) &\leqslant M_{|\bar{S}_n - \mu|}(s)e^{-s\epsilon} \\ &= \mathbb{E}\Big[e^{s|\bar{S}_n - \mu|}\Big]e^{-s\epsilon} \\ &\leqslant e^{s\mathbb{E}|\bar{S}_n - \mu|}e^{-s\epsilon} \end{split}$$

Now, we have to bound  $\mathbb{E}|\bar{S}_n - \mu|$ . Using the integral identity, we have:

$$\begin{split} \mathbb{E}|\bar{S}_n - \mu| &= \int_0^\infty \mathbb{P}(|\bar{S}_n - \mu| > t) dt \\ &\leq \frac{\sigma^2}{n} \int_0^\infty t^{-2} dt \quad (Chebyshev's Inequality) \end{split}$$

This does not work either because the above integral is divergent.

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