

December 16, 2022

Problem 1. [8] Let $X \sim U(0, 2)$ and $Y \sim U(0, 2)$ be two INDEPENDENT uniform random variables and let $Z = \max\{X, Y\}$.

(i) Compute $P[Z \leq z]$ for $z \in \mathbb{R}$;

(ii) Compute $E[Z]$;

(iii) Compute $E[1/Z]$.

$$(i) \quad P[Z \leq z] = P[\max\{X, Y\} \leq z] = P[X \leq z, Y \leq z]$$

$$\stackrel{(IND.)}{=} P[X \leq z] \cdot P[Y \leq z] = \begin{cases} 0 & z < 0 \\ z^2/4 & 0 \leq z < 2 \\ 1 & z \geq 2 \end{cases}$$

$$(ii) \quad \text{Since } z \geq 0, \quad E[Z] = \int_0^{+\infty} P[Z > z] dz = \int_0^2 (1 - z^2/4) dz =$$

$$= \left[z - \frac{z^3}{12} \right]_0^2 = 2 - \frac{8}{12} = 2 - \frac{2}{3} = \frac{6-2}{3} = \frac{4}{3}$$

$$(iii) \quad f_z(z) = z/2 \cdot \frac{1}{(0,2)}(z)$$

$$E[1/Z] = \int_0^2 \frac{1}{z} \cdot \frac{z}{2} dz = \int_0^2 \frac{1}{2} dz = \left[\frac{z}{2} \right]_0^2 = 1$$

Problem 2. [8] Let $X \sim \text{Bin}(1, p)$ and Y be a Poisson random variable with (random) parameter $X + 1$, i.e. $P[Y = k | X = x] = e^{-(x+1)} \frac{(x+1)^k}{k!}$ for $x = 0, 1$ and $k \in \mathbb{N}$.

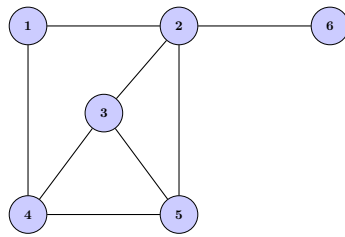
(i) Compute $P[Y = k]$ for any $k \in \mathbb{N}$;

(ii) Compute $E[Y]$.

$$\begin{aligned} \text{(i)} \quad P[Y = k] &= P[Y = k | X = 0] \cdot P[X = 0] + \\ &\quad + P[Y = k | X = 1] \cdot P[X = 1] \\ &= e^{-1} \frac{1}{k!} (1-p) + e^{-2} \frac{2^k}{k!} \cdot p \end{aligned}$$

$$\begin{aligned} \text{(ii)} \quad E[Y] &= \sum_{k=0}^{+\infty} k \left(e^{-1} \frac{1}{k!} (1-p) + e^{-2} \frac{2^k}{k!} p \right) \\ &= (1-p) \sum_{k=0}^{+\infty} k e^{-1} \frac{1}{k!} + p \sum_{k=0}^{+\infty} k \cdot e^{-2} \frac{2^k}{k!} \\ &= (1-p) 1 + p \cdot 2 = 1 + p \end{aligned}$$

Problem 3. [8] Define a simple Random Walk $\{X_n, n \geq 0\}$ on the undirected graph:



- (i) Compute the probability to go from 6 to 3 in three steps.
- (ii) Is the chain irreducible? Is the chain aperiodic?
- (iii) Find the invariant distribution.
- (iv) Starting from state 6, how many steps are needed on average to go back to 6?

$$(i) \quad \mathbb{P}[X_3 = 3 \mid X_0 = 6] = P_{62} \cdot P_{25} \cdot P_{53} = 1 \cdot \frac{1}{3} \cdot \frac{1}{4} = \frac{1}{12}$$

(ii) irreducible (the graph is connected), aperiodic
 $(P_{21}^2 > 0 \text{ and } P_{22}^3 > 0)$.

$$(iii) \quad P_1 = 2, P_2 = 4, P_3 = 3, P_4 = 3, P_5 = 3, P_6 = 1, \quad \sum_i P_i = 16$$

$$(\pi_1, \dots, \pi_6) = \left(\frac{2}{16}, \frac{4}{16}, \frac{3}{16}, \frac{3}{16}, \frac{3}{16}, \frac{1}{16} \right)$$

$$(iv) \quad m_6 = \mathbb{E}[\tau_6 \mid X_0 = 6] = \frac{1}{\pi_6} = 16$$

Problem 4. [10] Let $(X_i)_{1 \leq i \leq n}$ be a family of i.i.d. $\text{Poisson}(\lambda)$ random variables.

- (i) Compute the moment generating function of X_1 ;
- (ii) Prove that if $\lambda = 1$, then $X_1 + \dots + X_n$ is a $\text{Poisson}(n)$ random variable;
- (iii) Defined $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ and taking again $\lambda = 1$, determine an exponential decay for the "lower tail" of $\bar{X}_n - 1$.

$$\begin{aligned} \text{(i)} \quad m_{X_1}(t) &= \mathbb{E}[e^{tX_1}] = \sum_{k=0}^{\infty} e^{tk} e^{-\lambda} \frac{\lambda^k}{k!} = \\ &= e^{-\lambda} \cdot \sum_{k=0}^{\infty} \frac{(\lambda e^t)^k}{k!} = e^{-\lambda} \cdot e^{\lambda e^t} = e^{\lambda(e^t - 1)} \end{aligned}$$

$$\begin{aligned} \text{(ii)} \quad \text{if } \lambda = 1, \quad m_{X_1}(t) = \dots = m_{X_n}(t) &= e^{(e^t - 1)} \\ m_{X_1 + \dots + X_n}(t) &= \prod_{i=1}^n m_{X_i}(t) = e^{n(e^t - 1)} \sim \text{mgf-Poisson}(n) \end{aligned}$$

$\text{(iii)} \quad \text{let } \bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i, \quad \mathbb{E}[\bar{X}_n] = 1.$ The lower tail estimate for $\bar{X}_n - 1 = \bar{X}_n - \mathbb{E}[\bar{X}_n]$ is

$$\mathbb{P}[\bar{X}_n - 1 \leq -\varepsilon] = \mathbb{P}[\bar{X}_n \leq 1 - \varepsilon] \leq e^{-n \cdot \varepsilon^2 / 2} \quad \forall 0 < \varepsilon < 1$$

Indeed,

$$\mathbb{P}[\bar{X}_n \leq 1 - \varepsilon] \leq e^{-n h(t)}$$

$$\text{where } h(t) = -t(1 - \varepsilon) - \log m_{X_1}(-t) = -t(1 - \varepsilon) - (e^{-t} - 1)$$

$$h'(t) = 0 \Leftrightarrow t = t^* = -\log(1 - \varepsilon). \quad \text{Since } \varepsilon + (1 - \varepsilon)\log(1 - \varepsilon) \geq \varepsilon^2 / 2$$

for $0 < \varepsilon < 1$, we get $h(t^*) \geq \varepsilon^2 / 2$ and

$$\mathbb{P}[\bar{X}_n \leq 1 - \varepsilon] \leq e^{-n \varepsilon^2 / 2} \quad \forall 0 < \varepsilon < 1$$