

#### 1 Uniform Random Numbers

Let U be a uniform random variable between 0 and 1.

1.1 Generate  $10^6$  samples of U using a C program and save into a file called uni.dat.

**Solution:** Download the following files and execute the C program.

wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/ques
\_\_1/1.1.c
wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/coeffs.h

1.2 Load the uni.dat file into python and plot the empirical CDF of *U* using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr\left(U \le x\right) \tag{1.1}$$

**Solution:** The CDF of U is plotted in Fig. 0 using the below python code

wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/ques
\_1/1.2.py

1.3 Find a theoretical expression for  $F_U(x)$ .

**Solution:** The CDF of U is given by

$$F_U(x) = \Pr\left(U \le x\right) \tag{1.2}$$

$$= \int_{-\infty}^{x} f_U(x) dx \qquad (1.3)$$

(1.4)

1

Now, for x < 0

$$f_U(x) = 0 \tag{1.5}$$

$$\therefore F_U(x) = 0 \tag{1.6}$$

,for  $0 \le x \le 1$ 

$$f_U(x) = 1 \tag{1.7}$$

$$\therefore F_U(x) = \int_{-\infty}^x dx \tag{1.8}$$

$$=0+\int_{0}^{x}dx$$
 (1.9)

$$= x \tag{1.10}$$

and for x > 1

$$f_U(x) = 0 \tag{1.11}$$

$$\therefore F_U(x) = \int_{-\infty}^x dx \tag{1.12}$$

$$= 0 + \int_0^1 dx + 0 \quad (1.13)$$

$$= 1 \tag{1.14}$$

$$So, F_U(x) = \begin{cases} 0 & x < 0 \\ x & 0 \le x \le 1 \\ 1 & x > 1 \end{cases}$$

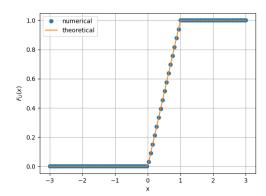


Fig. 0: The CDF of U

1.4 The mean of U is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^{N} U_i$$
 (1.15)

and its variance as

$$var[U] = E[U - E[U]]^2$$
 (1.16)

Write a C program to find the mean and variance of U.

**Solution:** Download and run the following C program.

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/ques \_1/1.4.c

1.5 Verify your result theoretically given that

$$E\left[U^{k}\right] = \int_{-\infty}^{\infty} x^{k} dF_{U}(x) \qquad (1.17)$$

**Solution:** The mean of U is given by

$$E[U] = \int_{-\infty}^{\infty} x dF_U(x)$$
 (1.18)

$$= 0 + \int_0^1 x dx + 0 \qquad (1.19)$$

$$= \left[ \frac{x^2}{2} \right]_0^1 \tag{1.20}$$

$$=1/2\tag{1.21}$$

and variance is given by,

$$E[U^{2}] - E[U]^{2} = \int_{-\infty}^{\infty} x^{2} dF_{U}(x) - \left(\frac{1}{2}\right)^{2}$$

$$= 0 + \int_{0}^{1} x^{2} dx + 0 - \left(\frac{1}{2}\right)^{2}$$

$$= \left[\frac{x^{2}}{3}\right]_{0}^{1} - \left(\frac{1}{2}\right)^{2}$$

$$= 1/12 \qquad (1.25)$$

## 2 Central Limit Theorem

2.1 Generate 10<sup>6</sup> samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \tag{2.1}$$

using a C program, where  $U_i$ , i = 1, 2, ..., 12 are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat **Solution:** Download the following files and run the C program.

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/

2.2 Load gau.dat in python and plot the empirical CDF of *X* using the samples in gau.dat. What properties does a CDF have?

**Solution:** The CDF of *X* is plotted in Fig. 0 using the below python code

# Properties of CDF:

- a) It is always continuous.
- b) It is monotonic increasing.
- c) It is bounded between 0 to 1.
- d) At infinity, it tends to 1.
- e) At negative infinity, it tends to 0.
- f) Since it is bounded and have limit so it is convergent also.
- g) Q function is defined as:

$$Q_X(x) = \Pr(x < X) \tag{2.2}$$

$$F_X(x) = 1 - Q_X(x)$$
 (2.3)

2.3 Load gau.dat in python and plot the empirical PDF of *X* using the samples in gau.dat. The PDF of *X* is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{2.4}$$

What properties does the PDF have? **Solution:** The PDF of *X* is plotted in Fig. 0 using the below python code

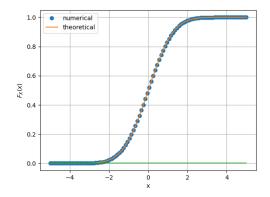


Fig. 0: The CDF of X

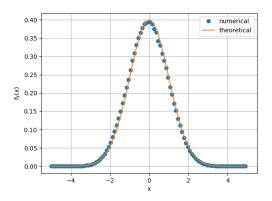


Fig. 0: The PDF of X

blob/master/ques 2/2.2.py

## Properties of PDF:

- a) It is always non-negative
- b) It is always bounded between 0 to 1
- c) Always symmetric about the mean.
- d) Maxima is observed at mean
- e) Area under the curve of PDF is always equal to 1

2.4 Find the mean and variance of *X* by writing a C program.

**Solution:** download the following files and run the C program.

wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/ques
\_2/2.4.c
wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/coeffs.h

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty,$$
(2.5)

repeat the above exercise theoretically. **Solution:** The mean of X is given by

$$E[X] = \int_{-\infty}^{\infty} x p_X(x) dx \qquad (2.6)$$

$$= \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \qquad (2.7)$$

$$= \left[-\frac{1}{\sqrt{2\pi}} exp\left(-\frac{x^2}{2}\right)\right]_{-\infty}^{+\infty} = 0 \qquad (2.8)$$

and variance is given by

$$E[X^{2}] - E[X]^{2} = \int_{-\infty}^{\infty} x^{2} f_{X}(x) dx$$

$$= \int_{-\infty}^{\infty} x^{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx \quad (2.10)$$

$$= 2 \int_{0}^{\infty} x^{2} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^{2}}{2}\right) dx \quad (2.11)$$

$$= \frac{2}{\sqrt{\pi}} \int_{0}^{\infty} t^{-1/2} \exp(-t) dt \quad (2.12)$$

$$= \frac{2}{\sqrt{\pi}} \Gamma(1 + 1/2) \quad (2.13)$$

$$= 1 \quad (2.14)$$

- 3 From Uniform to Other
- 3.1 Generate samples of

$$V = -2\ln(1 - U) \tag{3.1}$$

and plot its CDF.

**Solution:** Download and run the python and C file where you get samples of V and fig. ??

wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/ques
\_3/3.1.py
wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/ques
\_3/3.1.c

3.2 Find a theoretical expression for  $F_V(x)$ .

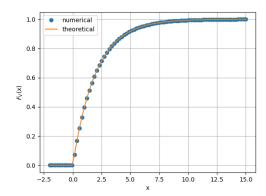


Fig. 0: The CDF of V

**Solution:** The CDF of V is given by

$$F_{V}(x) = \Pr(V \le x)$$

$$= \Pr(-2\ln(1 - U) \le x)$$

$$= \Pr\left(U \le 1 - \exp\left(-\frac{x}{2}\right)\right)$$
(3.4)

$$=F_U\left(1-\exp\left(-\frac{x}{2}\right)\right) \qquad (3.5)$$

let  $a = \exp(-x/2)$  .So,

$$F_V(x) = \begin{cases} 0 & 1 - a < 0 \\ 1 - a & 0 \le 1 - a < 1 \\ 1 & 1 - a \ge 1 \end{cases}$$

on solving, we get

$$F_V(x) = \begin{cases} 0 & x < 0\\ 1 - \exp\left(-\frac{x}{2}\right) & x \ge 0 \end{cases}$$

#### 4 Triangular Distribution

#### 4.1 Generate

$$T = U_1 + U_2 \tag{4.1}$$

**Solution:** Download the following code and run the C file

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/ques \_4/4.1.c

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/coeffs.h

### 4.2 Find the CDF of T.

**Solution:** The CDF of T is given by

$$F_T(t) = \Pr(T \le t) \tag{4.2}$$

Since T is sum of 2 uniform random variable between 0 and 1 So, any T will be always less than 2 and greater than 0 So,

$$F_T(t) = 0$$
 for  $X < 0$  (4.3)

$$F_T(t) = 1$$
 for  $X \ge 0$  (4.4)

Now, for  $0 \le t < 2$ ,

let  $U_1$  as a constant x and later we integrate it for all possible values of x So,

$$F_T(t) = \Pr(T \le t) \tag{4.5}$$

$$= \Pr(U_1 + U_2 \le t) \tag{4.6}$$

$$= \Pr\left(u_2 \le t - x\right) \tag{4.7}$$

$$= \int_0^t F_{U_2}(t-x) p_{u_1}(x) dx$$
(4.8)

Now, we make 2 cases

**case-1:**  $0 \le t < 1$ 

$$F_{T}(t) = \int_{0}^{t} F_{U_{2}}(t - x) dx \qquad (4.9)$$

$$= \int_{0}^{t} F_{U_{2}}(m) dm \qquad (4.10)$$

$$= \int_{0}^{t} m dm \qquad (4.11)$$

$$= \left[\frac{m^2}{2}\right]_0^t \tag{4.12}$$

$$=\frac{t^2}{2}$$
 (4.13)

**case-2:** 1 < t < 2

$$F_{T}(t) = \int_{0}^{1} F_{U_{2}}(t - x) dx \qquad (4.14)$$

$$= \int_{0}^{1} F_{U_{2}}(t - x) dx \qquad (4.15)$$

$$= \int_{t-1}^{t} F_{U_{2}}(m) dm \qquad (4.16)$$

$$= \int_{t-1}^{1} m dm + \int_{1}^{t} dm \qquad (4.17)$$

$$= \left[\frac{m^{2}}{2}\right]_{t-1}^{1} + [1]_{1}^{t} \qquad (4.18)$$

$$= \frac{-t^{2}}{2} + 2t - 1 \qquad (4.19)$$

4.3 Find the PDF of T.

**Solution:** The PDF of T is given by

$$p_T(t) = \frac{d}{dt} F_T(t) \tag{4.20}$$

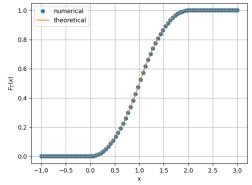


Fig. 0: The CDF of T

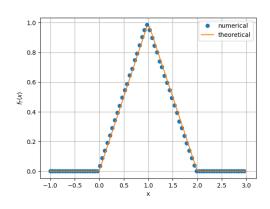


Fig. 0: The PDF of T

$$\therefore F_T(t) = \begin{cases} 0 & t < 0 \\ t & 0 \le t < 1 \\ 2 - t & 1 \le t < 2 \\ 0 & 2 \le t \end{cases}$$

4.4 Find the theoretical expressions for the PDF and CDF of T.

**Solution:** Already done in 4.2 and 4.3.

4.5 Verify your results through a plot. **Solution:** Plot of CDF and PDF is plotted in fig.

(0) and (0) respectively using following python code.

## 5 Maximul Likelihood

5.1 Generate equiprobable  $X \in \{1, -1\}$ . Solution: Download the following files and execute the C file

wget https://github.com/himanshukumargupta
11012/Random-Numbers/blob/master/ques
\_5/5.1.c
wget https://github.com/himanshukumargupta
11012/Random-Numbers/blob/master/coeffs.h

## 5.2 Generate

$$Y = AX + N, (5.1)$$

where A = 5 dB, and  $N \sim \mathcal{N}(0, 1)$ . **Solution:** Download the following files and run the C code

wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/ques
\_5/5.2.c
wget https://github.com/
himanshukumargupta
11012/Random-Numbers/
blob/master/coeffs.h

5.3 Plot *Y* using a scatter plot. **Solution:** The scatter plot of *Y* is plotted in fig. 0 using below python code.

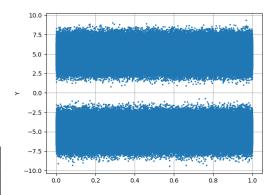


Fig. 0: The scatter plot of Y

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/ques \_5/5.3.py

5.4 Guess how to estimate *X* from *Y*.

Solution: After seeing the scatter plot we can conclude that if Y is positive then X will be 1 and if Y is negative then X will be -1. This conclusion also depends on value of A because if we decrease A the two parts of graph intermix and then we can't say anything. If A is approximately greater than 3 then only we can conclude this.

5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.2)

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.3)

**Solution:**  $\hat{X}$  is defined as

$$\hat{X} = \begin{cases} 1 & Y \ge 0 \\ -1 & Y < 0 \end{cases} \tag{5.4}$$

So,

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.5)

$$= \Pr(Y < 0 | X = 1) \tag{5.6}$$

$$= \Pr(A \times 1 + N < 0)$$
 (5.7)

$$= \Pr\left(N < -A\right) \tag{5.8}$$

$$=F_N\left(-A\right) \tag{5.9}$$

$$=0 (5.10)$$

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.11)

$$= \Pr(Y \ge 0 | X = -1) \tag{5.12}$$

$$= \Pr(A \times -1 + N \ge 0) \quad (5.13)$$

$$= \Pr\left(N \ge A\right) \tag{5.14}$$

$$= 1 - \Pr(N < A) \tag{5.15}$$

$$= 1 - F_N(A) (5.16)$$

$$=0 (5.17)$$

Download and run the following file to get coding proof

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/ques 5/5.5.py

5.6 Find  $P_e$  assuming that X has equiprobable symbols.

**Solution:**  $P_e$  can be defined as

$$P_e = \Pr(X = -1) P_{e|1} + \Pr(X = 1) P_{e|0}$$
(5.18)

$$P_{e|1} = 0, P_{e|0} = 0 (5.19)$$

$$\therefore P_e = 0 \tag{5.20}$$

5.7 Verify by plotting the theoretical  $P_e$  with respect to A from 0 to 10 dB.

**Solution:** X is equiprobable.So

$$Pr(X = 1) = Pr(X = -1), (5.21)$$

$$Pr(X = 1) + Pr(X = -1) = 1$$
(5.22)

$$\therefore \Pr(X = 1) = \Pr(X = -1) = 1/2$$
(5.23)

Now, $P_e$  is given by

$$P_e = \Pr\left(X = -1\right) P_{e|1} + \Pr\left(X = 1\right) P_{e|0}$$

(5.24)

$$P_e = \frac{1}{2} (1 - F_N(A)) + \frac{1}{2} (F_N(-A))$$
(5.25)

Now, the plot of  $P_e$  w.r.t. A is in fig. 0 by plotting the above function using following code

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/ques \_5/5.7.py

5.8 Now, consider a threshold  $\delta$  while estimating X from Y. Find the value of  $\delta$  that maximizes the theoretical  $P_e$ . **Solution:** Taking threshold as  $\delta$  we

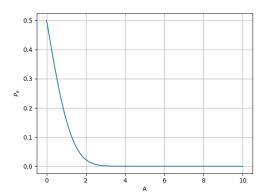


Fig. 0:  $P_e$  graph wrt A

get

$$P_{e|0} = \Pr(Y < \delta | X = 1)$$
 (5.26)  
=  $\Pr(A \times 1 + N < \delta)$  (5.27)

$$= \Pr\left(N < -A + \delta\right) \tag{5.28}$$

$$=F_N\left(-A+\delta\right) \tag{5.29}$$

and

$$P_{e|1} = \Pr(Y \ge \delta | X = -1)$$
 (5.30)

$$= \Pr(A \times -1 + N \ge \delta) \quad (5.31)$$

$$= \Pr\left(N \ge A + \delta\right) \tag{5.32}$$

$$= 1 - \Pr(N < A + \delta)$$
 (5.33)

$$=1-F_N(A+\delta) \tag{5.34}$$

and  $P_e$  is defined as

$$P_e = \Pr(X = -1) P_{e|1} + \Pr(X = 1) P_{e|0}$$
(5.35)

$$= \frac{1}{2} (1 - F_N (A + \delta) + F_N (-A + \delta))$$
(5.36)

Now we differentiate  $P_e$  and equate it

to 0 to get  $\delta$  that minimizes the  $P_e$ . So,

$$\frac{d}{d\delta}P_{e} = \frac{1}{2} \left( \frac{d}{d\delta}F_{N} \left( -A + \delta \right) - \frac{d}{d\delta}F_{N} \left( A + \delta \right) \right) = 0$$

$$(5.37)$$

$$\implies p_{N} \left( -A + \delta \right) - p_{N} \left( A + \delta \right) = 0$$

$$(5.38)$$

$$\implies \exp\left( -\frac{\left( -A + \delta \right)^{2}}{2} \right) - \exp\left( -\frac{\left( A + \delta \right)^{2}}{2} \right) = 0$$

$$(5.39)$$

$$\implies \delta = 0$$

$$(5.40)$$

Now we have to check that  $\delta$ =0 is for maxima or minima by double differentiating it. So,

$$\frac{d^2}{d\delta^2}|_{\delta=0} = 2A \exp\left(\frac{-A^2}{2}\right)$$
 (5.41)  
=  $10 \exp\left(\frac{-A^2}{2}\right) > 0$  (5.42)

Double derivative is positive. So, $\delta$ =0 will give minima of  $P_e$  Plot of  $P_e$  w.r.t.  $\delta$  is plotted in fig. 0 using following code

5.9 Repeat the above exercise when

$$p_X(0) = p$$
 (5.43)

**Solution:** 

$$Pr(X = 1) = p_X(0) = p \quad (5.44)$$

$$Pr(X = 1) = 1 \quad Pr(X = 1)$$

$$\therefore \Pr(X = -1) = 1 - \Pr(X = 1)$$
(5.45)

$$= 1 - p$$
 (5.46)

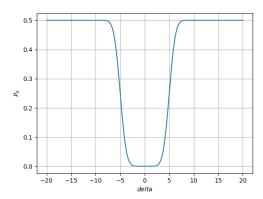


Fig. 0:  $P_e$  graph wrt  $\delta$ 

and  $P_e$  is given by

$$P_{e} = \Pr(X = -1) P_{e|1} + \Pr(X = 1) P_{e|0}$$

$$(5.47)$$

$$= (1 - p) (1 - F_{N}(A + \delta)) \quad (5.48)$$

$$+ pF_{N}(-A + \delta) \quad (5.49)$$

Then.

$$\frac{d}{d\delta}P_e = p\frac{d}{d\delta}F_N(-A+\delta) \qquad (5.50)$$

$$-(1-p)\frac{d}{d\delta}F_N(A+\delta) = 0 \qquad (5.51)$$

$$\implies p \times p_N(-A+\delta) \qquad (5.52)$$

$$-(1-p)p_N(A+\delta) = 0 \qquad (5.53)$$

$$\implies p \times \exp\left(-\frac{(-A+\delta)^2}{2}\right) \qquad (5.54)$$

$$-(1-p)\exp\left(-\frac{(A+\delta)^2}{2}\right) = 0 \qquad (5.55)$$

$$\implies \delta = \frac{\log\left(\frac{1}{p}-1\right)}{24} \qquad (5.56)$$

(5.56)

5.10 Repeat the above exercise using the MAP criterion.

**Solution:** From bayes theorem we can write,

$$\Pr(X = 1|Y = y) \qquad (5.57)$$

$$= \frac{\Pr(Y = y|X = 1)\Pr(X = 1)}{\Pr(Y = y)} \qquad (5.58)$$

$$= \frac{\Pr(N = y - A)p}{\Pr(N = y - A)p + \Pr(N = y + A)(1 - p)} \qquad (5.59)$$

$$= \frac{p_N(y - A)p}{p_N(y - A)p + p_N(y + A)(1 - p)} \qquad (5.60)$$

$$= \frac{p}{p + (1 - p)\exp(-2Ay)} \qquad (5.61)$$

$$Pr(X = -1|Y = 0) (5.62)$$

$$= 1 - Pr(X = 1|Y = 0) (5.63)$$

$$= 1 - \frac{p}{p + (1 - p)\exp(-2Ay)} (5.64)$$

$$= \frac{(1 - p)\exp(-2Ay)}{p + (1 - p)\exp(-2Ay)} (5.65)$$

$$= \frac{1 - p}{p + (1 - p)\exp(-2Ay)} (5.65)$$

$$= \frac{1 - p}{p \exp(2Ay) + 1 - p}$$
 (5.66)

6 Gaussian to Other

6.1 Let  $X_1 \sim \mathcal{N}(0,1)$  and  $X_2 \sim \mathcal{N}(0,1)$ . Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \tag{6.1}$$

**Solution:** Download the following files and run C and python program. CDF and PDF of V is plotted in fig. 0 and 0 respectively using below python code.

wget https://github.com/ himanshukumargupta

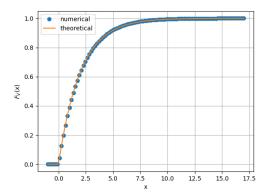


Fig. 0: The CDF of V

11012/Random-Numbers/blob/master/coeffs.h
wget https://github.com/himanshukumargupta
11012/Random-Numbers/blob/master/ques
\_6/6.1.c
wget https://github.com/himanshukumargupta
11012/Random-Numbers/blob/master/ques
\_6/6.1.py

6.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \ge 0 \\ 0 & x < 0, \end{cases}$$
 (6.2)

find  $\alpha$ .

**Solution:** Let

 $X_1 = R \cos \theta$ ,  $X_2 = R \sin \theta$ So, PDF of R and  $\theta$  is given by

$$p_{R,\theta}(r,\phi) = p_{X_1,X_2}(x_1,x_2)|J|$$
 (6.3)

where J is jacobian matrix transforming  $R,\theta$  to  $X_1,X_2$ 

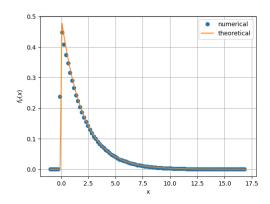


Fig. 0: The PDF of V

On soving we get jacobian,J=R. So,

$$p_{R,\theta}(r,\phi) = p_{X_1,X_2}(x_1,x_2)r$$
 (6.4)

since  $X_1$  and  $X_2$  are i.i.d. and gaussian random variable,  $\mathcal{N}(0, 1)$ . So,

$$p_{R,\theta}(r,\phi) = p_{X_1}(x_1) p_{X_2}(x_2) r \qquad (6.5)$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x_1^2}{2}\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x_2^2}{2}\right) r \qquad (6.6)$$

$$= \frac{1}{2\pi} \exp\left(-\frac{r^2}{2}\right) r \qquad (6.7)$$

Now PDF of R,

$$p_{R}(r) = \int_{0}^{2\pi} p_{R,\theta}(r,\phi) d\phi \qquad (6.8)$$

$$= \int_{0}^{2\pi} \frac{1}{2\pi} \exp\left(-\frac{r^{2}}{2}\right) r d\phi \quad (6.9)$$

$$= \exp\left(-\frac{r^{2}}{2}\right) r \qquad (6.10)$$

and

$$V = X_1^2 + X_2^2 = R^2 (6.11)$$

Since  $V = R^2 > 0$ 

So,CDF and PDF would be zero for v<0

Then CDF of V for  $v \ge 0$ ,

$$F_V(v) = \Pr(V \le v) \tag{6.12}$$

$$= \Pr\left(R^2 \le v\right) \tag{6.13}$$

$$= \Pr\left(R \le \sqrt{v}\right) \tag{6.14}$$

$$=F_R\left(\sqrt{v}\right) \tag{6.15}$$

$$= \int_{-\infty}^{\sqrt{\nu}} p_R(r) dr \qquad (6.16)$$

$$= 0 + \int_0^{\sqrt{\nu}} \exp\left(-\frac{r^2}{2}\right) r dr$$
(6.17)

$$= 1 - \exp\left(-\frac{v}{2}\right) \tag{6.18}$$

Comparing the 2 equations we get

$$\alpha = \frac{1}{2} \tag{6.19}$$

and PDF is given by

$$p_V(v) = \frac{d}{dx} F_V(v) \tag{6.20}$$

$$\therefore p_V(v) = \begin{cases} \frac{\exp(-\frac{v}{2})}{2} & v \ge 0\\ 0 & v < 0 \end{cases}$$
 (6.21)

6.3 Plot the CDF and PDf of

$$A = \sqrt{V} \tag{6.22}$$

**Solution:** CDF and PDF of A is plotted in fig. ?? and ?? respectively using below python code.

wget https://github.com/ himanshukumargupta 11012/Random-Numbers/ blob/master/ques

$$6/6.3.$$
 py

We know that

$$V = R^2 \tag{6.23}$$

$$\therefore A = \sqrt{V} = R \tag{6.24}$$

7 CONDITIONAL PROBABILITY

7.1 Plot

$$P_e = \Pr(\hat{X} = -1|X = 1)$$
 (7.1)

for

$$Y = AX + N. (7.2)$$

where A is Raleigh with  $E[A^2] = \gamma, N \sim \mathcal{N}(0, 1), X \in (-1, 1)$  for  $0 \le \gamma \le 10$  dB.

- 7.2 Assuming that *N* is a constant, find an expression for  $P_e$ . Call this  $P_e(N)$
- 7.3 For a function g,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)p_X(x) dx \quad (7.3)$$

Find  $P_e = E[P_e(N)]$ .

7.4 Plot  $P_e$  in problems 7.1 and 7.3 on the same graph w.r.t  $\gamma$ . Comment.

8 Two Dimensions

Let

$$\mathbf{y} = A\mathbf{x} + \mathbf{n},\tag{8.1}$$

where

$$x \in (\mathbf{s}_0, \mathbf{s}_1), \mathbf{s}_0 = \begin{pmatrix} 1 \\ 0 \end{pmatrix}, \mathbf{s}_1 = \begin{pmatrix} 0 \\ 1 \end{pmatrix}$$
 (8.2)

$$\mathbf{n} = \begin{pmatrix} n_1 \\ n_2 \end{pmatrix}, n_1, n_2 \sim \mathcal{N}(0, 1).$$
 (8.3)

8.1 Plot

$$\mathbf{y}|\mathbf{s}_0$$
 and  $\mathbf{y}|\mathbf{s}_1$  (8.4)

on the same graph using a scatter plot.

- 8.2 For the above problem, find a decision rule for detecting the symbols  $\mathbf{s}_0$  and  $\mathbf{s}_1$ .
- 8.3 Plot

$$P_e = \Pr(\hat{\mathbf{x}} = \mathbf{s}_1 | \mathbf{x} = \mathbf{s}_0) \tag{8.5}$$

with respect to the SNR from 0 to 10 dB.

8.4 Obtain an expression for  $P_e$ . Verify this by comparing the theory and simulation plots on the same graph.