

# Data association techniques in multiple object tracking

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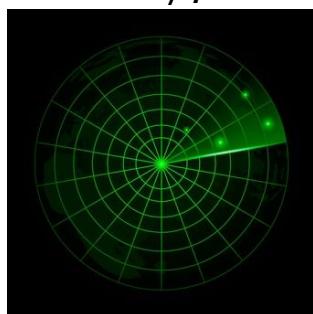
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# Introduction

- **What is data association?**
  - Process of associating uncertain measurements to known tracks.
  - **Tracking = Data association + Filtering**

- **When is it used?**

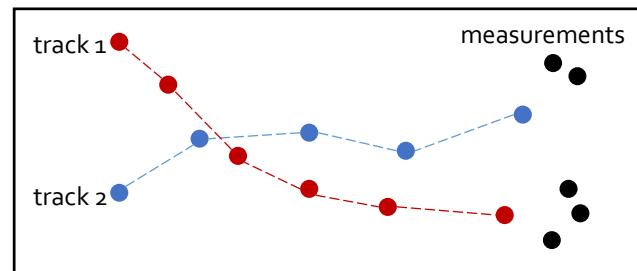
- Initially, radars → now, all kinds of tracking problems.



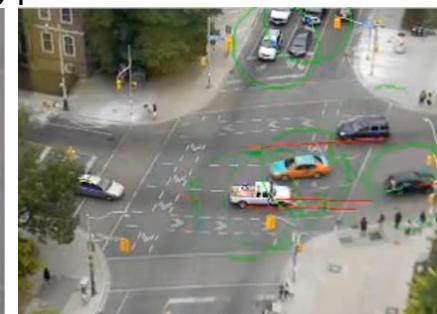
radar



laboratory



track 1  
measurements  
track 2



surveillance



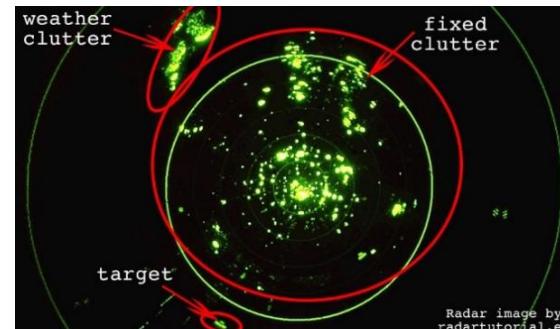
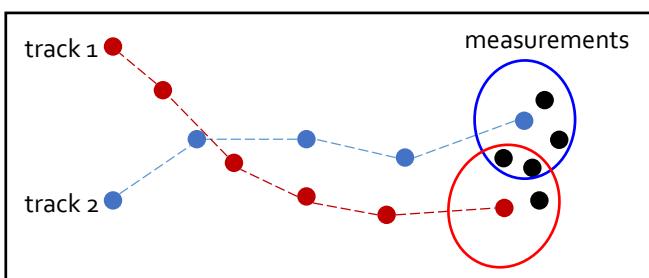
motion capture



fleet management

## • Problem?

- multiple tracks → association ambiguity
- false alarm/clutter → shouldn't detect, but detected (false positive)
- detection uncertainty → should detect, but not detected (false negative)



# Introduction

- System is assumed to behave as:

- state vector:  $x(k) = F(k-1)x(k-1) + v(k-1) \in \mathbb{R}^{n_1}$
- measurement vector:  $y(k) = H(k)x(k) + w(k) \in \mathbb{R}^{n_2}$
- $v(k-1) \sim \mathcal{N}[0, Q(k-1)]$ ,  $w(k) \sim \mathcal{N}[0, R(k)]$ : Gaussian noise

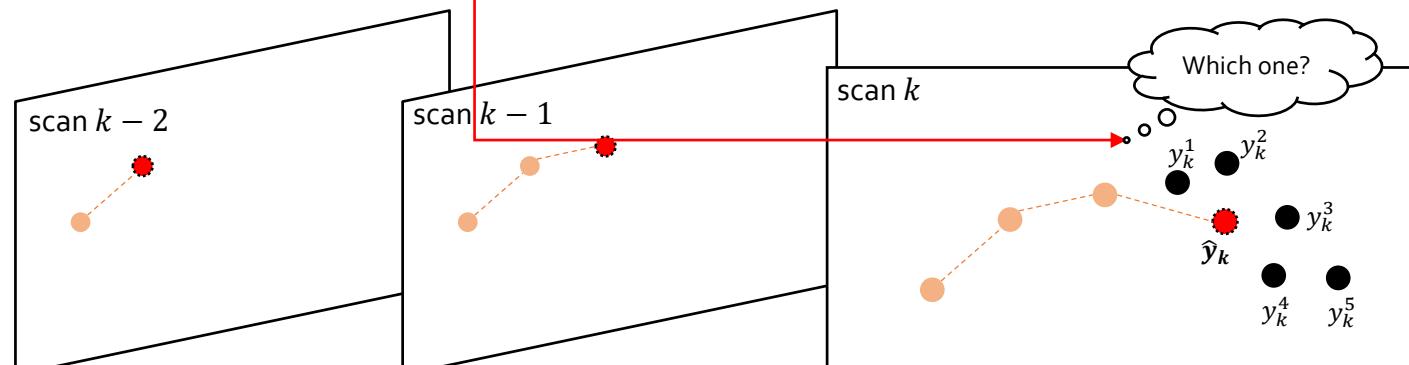
1. Predict state, measurement, and state covariance

- a *prior* state estimate:  $\hat{x}_{k|k-1} = F_{k-1}\hat{x}_{k-1|k-1}$
- predicted measurement:  $\hat{y}_k = H_k\hat{x}_{k|k-1}$
- a *priori* state covariance:  $P_{k|k-1} = F_{k-1}P_{k-1|k-1}F_{k-1}^T + Q_{k-1}$

## Validation gating and data association process

2. Update state using measurement

- innovation vector:  $z_k = y_k - \hat{y}_k$
- a *posterior* state estimate:  $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k z_k$



# Introduction

- **Validation Gating**

→ Rule out measurements that are geometrically unlikely from the start.

- Assuming Gaussian probability density for the residual  $z_k$  follow:

$$p(z_k) = p(y_k - \hat{y}_k) = \mathcal{N}(z_k; 0, \hat{S}_k) = \frac{1}{2\pi\sqrt{|\hat{S}_k|}} e^{-\frac{1}{2}z_k^T \hat{S}_k^{-1} z_k}$$

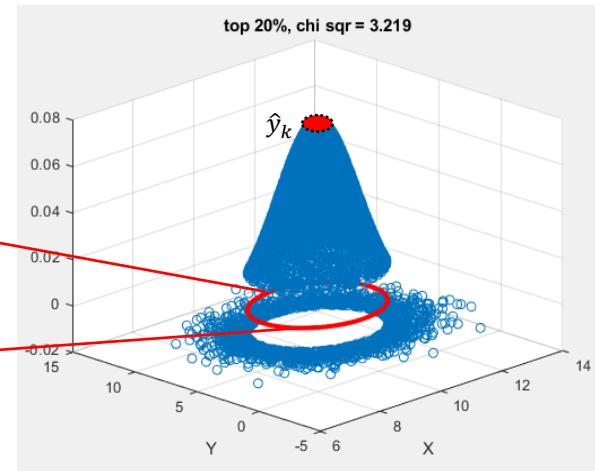
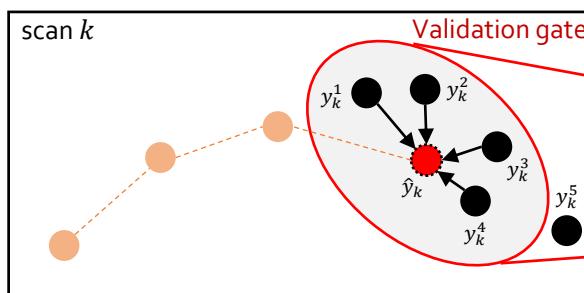
- Validation gate using **Mahalanobis distance**:

$$(y_k^i - \hat{y}_k)^T \hat{S}_k^{-1} (y_k^i - \hat{y}_k) < \chi_k^2$$

where,

- $\hat{S}_k = H_k P_{k|k-1} H_k^T$  = residual covariance (uncertainty in  $\hat{y}_k$ )

- $\chi_k^2$ : chi-square value at 2dof (look-up value)



- In implementation perspective, use **Euclidean distance** to avoid computing  $\hat{S}_k^{-1}$ .
- Gate calculation costs a great part of the total MHT calculation time. (discussed later)

# Data association techniques

## I. Non-Bayesian

1. Nearest Neighbor Standard Filter (NNSF)
2. Global Nearest Neighbor (GNN)

## II. Bayesian

1. Probabilistic Data Association Filter (PDAF)
2. Joint Probabilistic Data Association Filter (JPDAF)
3. Multiple Hypothesis Tracking (MHT)

# Data association techniques

## I. Non-Bayesian

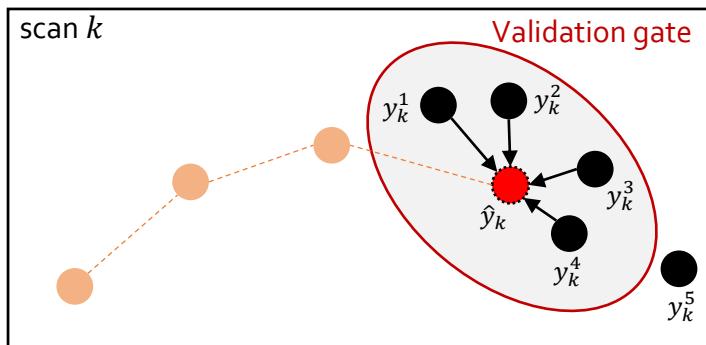
1. **Nearest Neighbor Standard Filter (NNSF)**
2. Global Nearest Neighbor (GNN)

## II. Bayesian

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# Nearest Neighbor Standard Filter (NNSF)

- Key idea
  - choose one **closest** measurement.



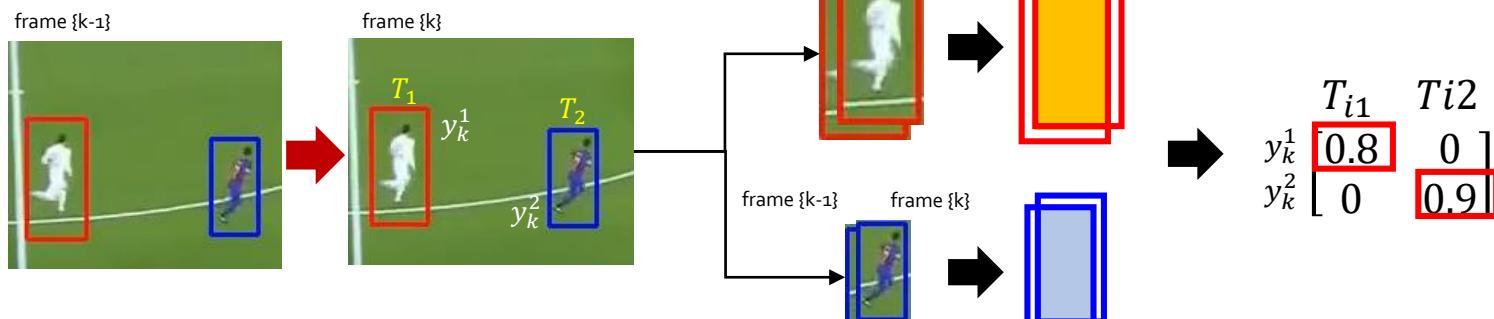
Assignment matrix

$$\begin{matrix} T_{i1} \\ y_k^1 \\ y_k^2 \\ y_k^3 \\ y_k^4 \\ y_k^5 \end{matrix}$$

$$\leftarrow (y_k^i - \hat{y}_k)^T \hat{S}_k^{-1} (y_k^i - \hat{y}_k)$$

- update state estimate using  $y_k^4$   
 $\rightarrow$  a posterior state estimate:  $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k (y_k^4 - \hat{y}_k)$

- Can use other scoring criteria other than distance.
  - Affinity scores such as size, shape, appearance, etc.
  - e.g. location, size and shape similarity can be scored based on bounding box **overlap** (if targets move slowly)



# Nearest Neighbor Standard Filter (NNSF)

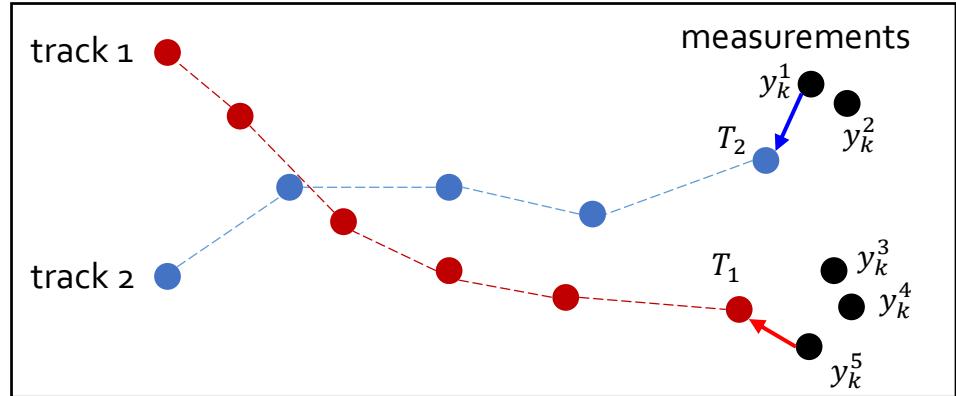
- Implementation:

- Build assignment matrix, each element:

$$d_{ij}^2 = (y_k^i - \hat{y}_k)^T \hat{S}_k^{-1} (y_k^i - \hat{y}_k)$$

Initialization

$$A = \begin{bmatrix} d_{11}^2 & d_{12}^2 & d_{13}^2 & d_{14}^2 & d_{15}^2 \\ d_{21}^2 & d_{22}^2 & d_{23}^2 & d_{24}^2 & d_{25}^2 \end{bmatrix}_{\text{tracks}} = \begin{bmatrix} 200 & 170 & 36.25 & 52 & 28.5 \\ 25 & 25 & 56.25 & 89 & 127.25 \end{bmatrix}$$



Loop

- Find the closest pairing in  $A$ .

$$= \begin{bmatrix} 200 & 170 & 36.25 & 52 & 28.5 \\ 25 & 26 & 56.25 & 89 & 127.25 \end{bmatrix}$$

- Remove row & column of that pairing.

$$= \begin{bmatrix} 200 & 170 & 36.25 & 52 & 28.5 \\ 25 & 26 & 56.25 & 89 & 127.25 \end{bmatrix}$$

- Find next closest pairing in  $A$ .

$$= \begin{bmatrix} 200 & 170 & 36.25 & 52 & 28.5 \\ 25 & 26 & 56.25 & 89 & 127.25 \end{bmatrix}$$

- Remove row & column of that pairing.

$$= \begin{bmatrix} 200 & 170 & 36.25 & 52 & 28.5 \\ 25 & 26 & 56.25 & 89 & 127.25 \end{bmatrix} \begin{matrix} T_1 \\ T_2 \end{matrix}$$

Update

- No elements left  $\rightarrow$  done.
- Update all tracks with associated measurements.

- Generally does **not** find the global minimum.
  - Cannot handle association ambiguities.
- $\rightarrow$  Instead of NNSF, for multiple tracks use **Global Nearest Neighbor(GNN)**

# Data association techniques

## I. Non-Bayesian

1. Nearest Neighbor Standard Filter (NNSF)
2. **Global Nearest Neighbor (GNN)**

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# Global Nearest Neighbor (GNN)

- Problem: both tracks want  $y_k^3$
- Linear Assignment Problem  
→ joint data association, finds global minimum.
- “Best assignment of  $n$  different workers to  $n$  jobs”

*Given a cost matrix  $C = \{c_{ij}\}$ ,*

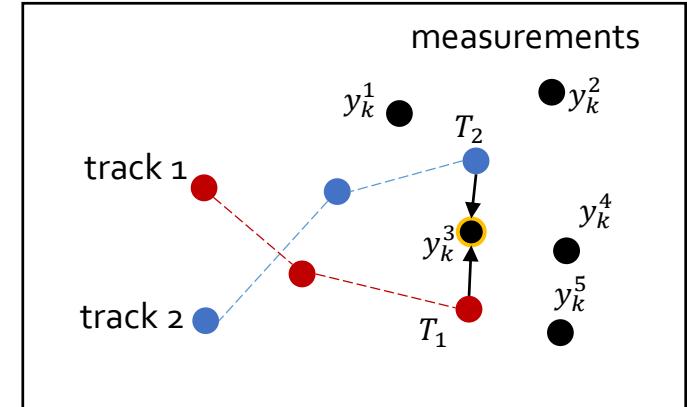
$$\min \sum c_{ij}x_{ij}, \quad x_{ij} \in \{1, 0\}$$

$$s. t. \boxed{\sum_i x_{ij} = 1}, \boxed{\sum_j x_{ij} = 1}$$

sum of each row = 1

sum of each column = 1

e.g.  $\begin{bmatrix} T_1 & T_2 & T_3 \\ 108 & 125 & 150 \\ 150 & 135 & 175 \\ 122 & 148 & 250 \end{bmatrix} \begin{bmatrix} y_k^1 \\ y_k^2 \\ y_k^3 \\ y_k^4 \\ y_k^5 \end{bmatrix} \rightarrow \begin{bmatrix} 0 & 0 & 1 \\ 0 & 1 & 0 \\ 1 & 0 & 0 \end{bmatrix}$ , with minimum cost:  $\sum c_{ij}x_{ij} = 122 + 135 + 150 = 407$



- **Hungarian Algorithm** to solve.  
For non-square cost matrix, augment it with zeros to make it square.
- Developed by American mathematician Harold Kuhn in 1955.
- Time complexity  $O(n^4)$

## Global Nearest Neighbor (GNN)

- Implementation:

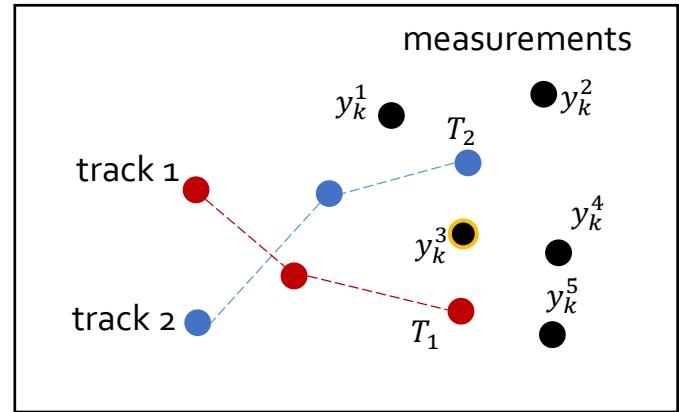
1. Build cost(assignment) matrix, each element:

$$c_{ij} = \left( y_k^i - \hat{y}_k \right)^T \hat{S}_k^{-1} \left( y_k^j - \hat{y}_k \right)$$

- $$2. \quad C = \begin{bmatrix} 100 & 105 & 20 & 32 & 28 \\ 30 & 39 & 20 & 47 & 93 \end{bmatrix} \text{ tracks}$$

measurements

war



3. Subtract row minima  $\rightarrow$  reduce the rows

$$\left[ \begin{array}{cccccc} 100 & 105 & 20 & 32 & 28 \\ 30 & 39 & 20 & 47 & 93 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{array} \right] \xrightarrow{\quad} \left[ \begin{array}{ccccc} 100 - 20 & 105 - 20 & 20 - 20 & 32 - 20 & 28 - 20 \\ 30 - 20 & 39 - 20 & 20 - 20 & 47 - 20 & 93 - 20 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{array} \right] \xrightarrow{\quad} \left[ \begin{array}{cccccc} 80 & 85 & 0 & 12 & 8 \\ 10 & 19 & 0 & 27 & 73 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{array} \right]$$

4. Subtract column minima → reduce the columns

$$\begin{bmatrix} 80 & 85 & 0 & 12 & 8 \\ 10 & 19 & 0 & 27 & 73 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 80 & 85 & 0 & 12 & 8 \\ 10 & 19 & 0 & 27 & 73 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix}$$

# Global Nearest Neighbor (GNN)

5. Cover zeros using minimum  $N$  number of lines.

$$\begin{bmatrix} 80 & 85 & 0 & 12 & 8 \\ 10 & 19 & 0 & 27 & 73 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 80 & 85 & 0 & 12 & 8 \\ 10 & 19 & 0 & 27 & 73 \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \end{bmatrix} \rightarrow N = 4 \text{ (i.e. 4 lines used)}$$

6.  $N = 4 == 5$ ? No  $\rightarrow$  Create more zeros by:

- 1) Subtracting minima of uncovered elements from uncovered elements.

$$\begin{bmatrix} 80 & 85 & 0 & 12 & 8 \\ 10 & 19 & 0 & 27 & 73 \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \end{bmatrix} \rightarrow \begin{bmatrix} 80 - 8 & 85 - 8 & 0 & 12 - 8 & 8 - 8 \\ 10 - 8 & 19 - 8 & 0 & 27 - 8 & 73 - 8 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix}$$

- 2) Leave once-covered elements as they are.

$$\begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} & \cancel{0} \end{bmatrix} \rightarrow \begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ \color{red}{0} & \color{red}{0} & \color{blue}{0} & \color{red}{0} & \color{red}{0} \\ \color{red}{0} & \color{red}{0} & \color{blue}{0} & \color{red}{0} & \color{red}{0} \\ \color{red}{0} & \color{red}{0} & \color{blue}{0} & \color{red}{0} & \color{red}{0} \end{bmatrix}$$

- 3) Add minima to all twice covered elements.

$$\begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ \color{red}{0} & \color{red}{0} & \color{blue}{0} & \color{red}{0} & \color{red}{0} \\ \color{red}{0} & \color{red}{0} & \color{blue}{0} & \color{red}{0} & \color{red}{0} \\ \color{red}{0} & \color{red}{0} & \color{blue}{0} & \color{red}{0} & \color{red}{0} \end{bmatrix} \rightarrow \begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ 0 & 0 & \color{blue}{8} & 0 & 0 \\ 0 & 0 & \color{blue}{8} & 0 & 0 \\ 0 & 0 & \color{blue}{8} & 0 & 0 \end{bmatrix}$$

# Global Nearest Neighbor (GNN)

7. Cover zeros using minimum  $N$  number of lines

$$\begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ \cancel{0} & \cancel{0} & \cancel{8} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{8} & \cancel{0} & \cancel{0} \\ \cancel{0} & \cancel{0} & \cancel{8} & \cancel{0} & \cancel{0} \end{bmatrix} \rightarrow N = 5 \text{ (i.e. 5 lines used)}$$

8.  $N = 5 == 5$ ? Yes  $\rightarrow$  Do association.

9. Assign task  $k \rightarrow$  Start with lines with only one zero.

$$\begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 72 & 77 & 0 & 4 & 0 \\ 2 & 11 & 0 & 19 & 65 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \\ 0 & 0 & 8 & 0 & 0 \end{bmatrix}$$

10. Calculate the corresponding cost

$$C = \begin{bmatrix} 100 & 105 & 20 & 32 & 28 \\ 30 & 39 & 20 & 47 & 93 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} 100 & 105 & 20 & 32 & 28 \\ 30 & 39 & 20 & 47 & 93 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \rightarrow 20 + 28 + 0 + 0 + 0 = 48 \text{ (minimum)}$$

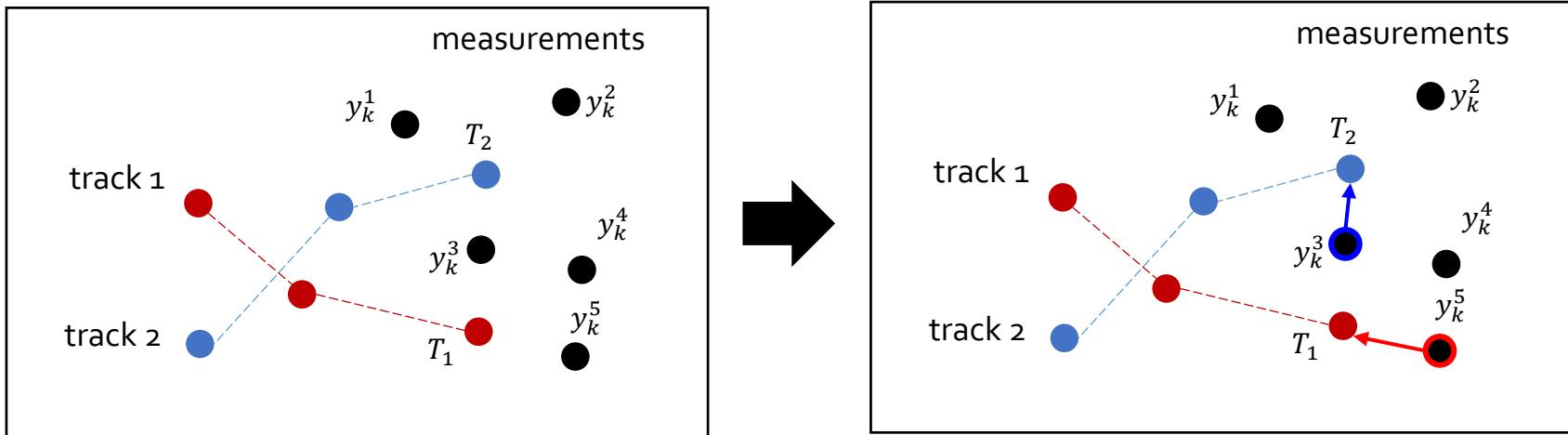
# Global Nearest Neighbor (GNN)

$$C = \begin{bmatrix} 100 & 105 & 20 & 32 & 28 \\ 30 & 39 & 20 & 47 & 93 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \rightarrow \begin{bmatrix} y_k^1 & y_k^2 & y_k^3 & y_k^4 & y_k^5 \\ 100 & 105 & 20 & 32 & 28 \\ 30 & 39 & 20 & 47 & 93 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{bmatrix} \xrightarrow{\text{highlighted}} \begin{bmatrix} T_1 & T_2 \\ \text{---} & \text{---} \end{bmatrix} \rightarrow \text{output: } \begin{bmatrix} 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 1 & 0 & 0 \\ 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \end{bmatrix} = \text{permutation matrix}$$

$\rightarrow \text{cost: } 20 + 28 + 0 + 0 + 0 = 48 \text{ (minimum)}$

Given a cost matrix  $C = \{c_{ij}\}$ ,  
 $\min \sum c_{ij} x_{ij}, \quad x_{ij} \in \{0, 1\}$   
s.t.  $\sum_i x_{ij} = 1, \sum_j x_{ij} = 1$

Result



Solution may not be unique  $\rightarrow$  same cost, but different matching pairs.

- **State update**

- update **track 1** state estimate using  $y_k^5$

$\rightarrow$  a posterior state estimate:  $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k(y_k^5 - \hat{y}_k)$

- update **track 2** state estimate using  $y_k^3$

$\rightarrow$  a posterior state estimate:  $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k(y_k^3 - \hat{y}_k)$

# Data association techniques

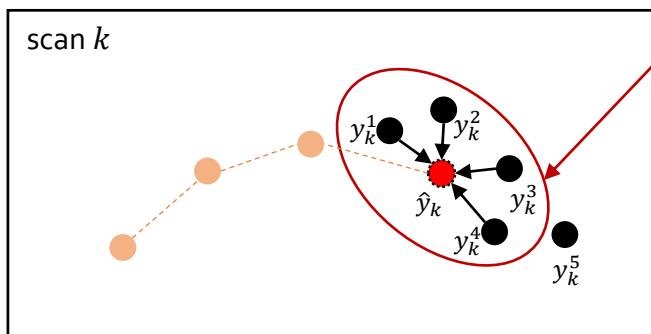
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# Probabilistic Data Association Filter (PDAF)

- Key idea

- Update state using ***all*** validated measurements weighted by their likelihoods.
- ①                           ②
- $\hat{x}_{k|k} = \int x_k p(x_k | Y^k) dx_k$ , where  $Y_k = \{y_k^i\}_{i=1}^{m_k}$ : a set of measurements at  $k^{th}$  scan:  
The best estimate, in MMSE sense, considers all the measurements and their probability that they are originated from the same track.
  - $\hat{x}_{k|k} = \sum_{i=0}^{m_k} E\{x_k | \chi_{k,i}, Y^k\} P(\theta_{k,i} | Y^k)$

$P(\theta_{k,i} | Y^k) \triangleq \beta_k^i$   
= Association probability  
= Probability that  $y_k^i$  is associated with track.



① **Validated measurements:**

→ Validation gate using Mahalanobis distance:

$$(y_k^i - \hat{y}_k)^T \hat{S}_k^{-1} (y_k^i - \hat{y}_k) < \chi_k^2$$

where,

$$\hat{S}_k = H_k P_{k|k-1} H_k^T$$

= Residual variance (uncertainty in  $\hat{y}_k$ )

② **weighted by their likelihoods:**

→ condition the update on the association events.

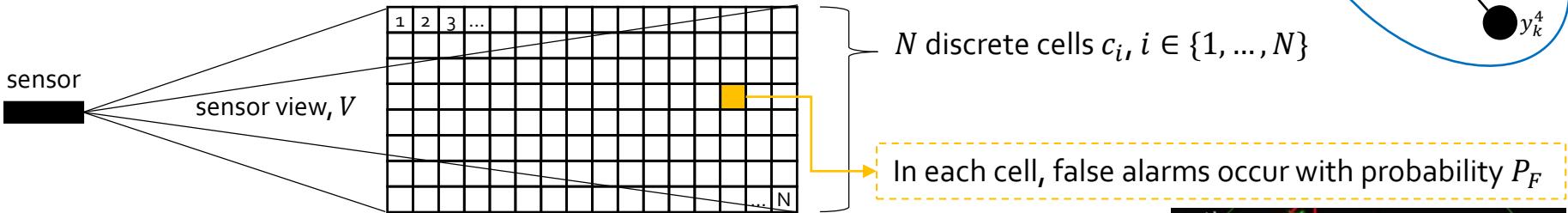
- 1) False alarm model
- 2) Association events

# Probabilistic Data Association Filter (PDAF)

1) **False alarm (false positive)** due to sensor imperfections, backgrounds, etc.

- Assume we know spatial density  $\lambda$  of false alarm.

$$\text{Uniformly distributed across } V: \text{spatial density } \lambda = \frac{N \cdot P_F}{V} = \frac{N_p}{V}$$



- Follows **Bernoulli process** (appear / not appear) with probability:  $P_F$ .

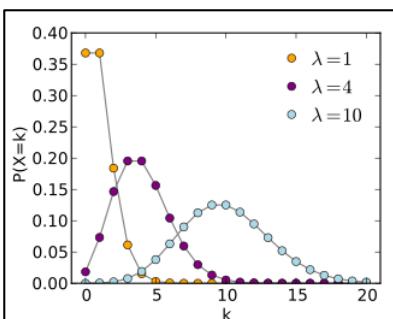
$$\rightarrow \text{Binomial distribution: } P(m_F) = \binom{N}{m_F} P_F^{m_F} (1 - P_F)^{N-m_F}$$

- Average number of false alarms in the sensor view  $V$  each scan:  $N_p$ .

- Let  $N \rightarrow \infty$ , then **Binomial distribution** with  $P_F \ll 1$  approximates to:

**Poisson distribution** (mean = variance =  $\lambda V_k$ )  $\rightarrow$  easier to compute for large  $N$ .

$$\lim_{N \rightarrow \infty} \binom{N}{m_F} P_F^{m_F} (1 - P_F)^{N-m_F} = e^{-\lambda V_k} \frac{(\lambda V_k)^{m_F}}{m_F!} =: \mu_F(m_F)$$

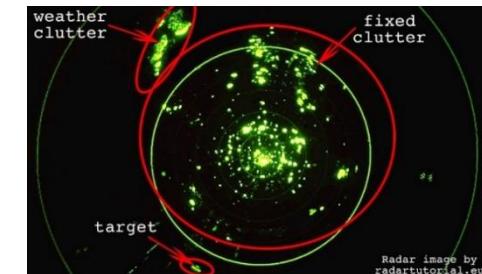
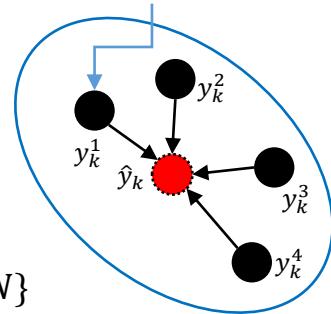


**Poisson distribution** is a model for  $k$  number of events:

$$P(k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

$\rightarrow$  mean = variance =  $\lambda$ .

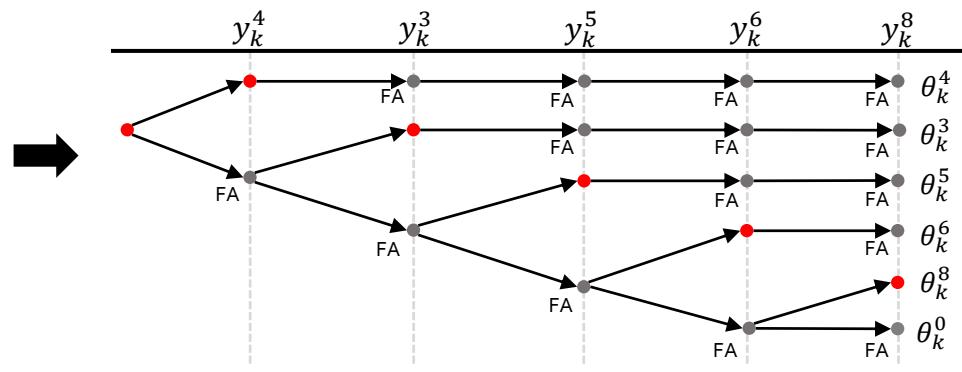
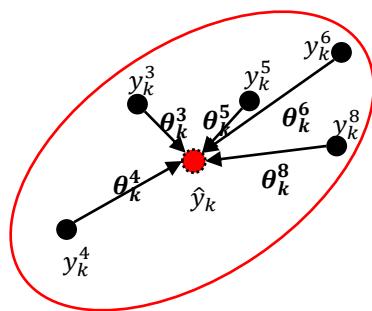
actual measurement  
or false alarm?



# Probabilistic Data Association Filter (PDAF)

## 2) Association event $\theta_k^i$

- $\theta_k^i = \begin{cases} \text{measurement } y_k^i \text{ belongs to the track.} & i = 1, \dots, m_k \\ \text{none of the measurement belongs to the track.} & i = 0 \end{cases}$
- $\beta_k^i$  is the **association probability**, i.e. a posteriori probability  $P(\theta_k^i | Y^k, m_k) = \beta_k^i$ , such that  $\sum_{i=0}^{m_k} \beta_k^i = 1$ .



- $P_D$  = detection probability = the target from the track is detected.
  - e.g. "the target we are tracking has emitted a signal, and we detected it and is among the measurements inside the gate."
  - assumed to be known.
    - dependent on sensors and scenarios
    - can be tested via offline experiments.
- $P_G$  = validation probability = probability corresponding to  $\chi_k^2$ .
- Assumptions:**
  - at most one of the validated observations is target-originated.
  - remaining measurements are false alarms.

# Probabilistic Data Association Filter (PDAF)

$\beta_k^i$  = Association probability given measurements up to k,  $Y^k$ :

$$\beta_k^i = P(\theta_k^i | Y^k) = P(\theta_k^i | Y_k, m_k, Y^{k-1}) = \frac{1}{c} f(Y_k | \theta_k^i, m_k, Y^{k-1}) P(\theta_k^i | m_k, Y^{k-1})$$

where,  $i = 0, \dots, m(k)$

Baye's formula:

$$P(\theta | Y) = \frac{f(Y|\theta)P(\theta)}{\sum f(Y|\theta_i)P(\theta_i)}$$

1.  $f(Y_k | \theta_k^i, m_k, Y^{k-1})$  = pdf of the measurements.

$$\begin{cases} \frac{1}{V_k} m_{k-1} \frac{\mathcal{N}(z_k^i; 0, \hat{S}_k)}{\sum_{j=1}^{m_k} \mathcal{N}(z_k^j; 0, \hat{S}_k)}, & i = 1, \dots, m_k \\ \frac{1}{V_k} m_k, & i = 0 \end{cases}$$

← pdf of incorrect measurements assumed *uniform* in the validation region  $V_k$ . i.e.  $\frac{1}{V_k}$

2.  $P(\theta_k^i | m_k, Y^{k-1})$  = prior probability of the association events, conditioned on the number of measurements.

$$= \begin{cases} P(\theta_k^i | \text{when 1 measurement is to be associated, } m_k - 1 \text{ are FA}) \times P(m_k - 1 \text{ are FA}), & i = 1, \dots, m_k \\ P(\theta_k^i | \text{when 0 measurement is to be associated, } m_k \text{ are FA}) \times P(m_k \text{ are FA}), & i = 0 \end{cases}$$

$$= \begin{cases} P(\theta_k^i | FA = m_k - 1) P(FA = m_k - 1 | Y_k), & i = 1, \dots, m_k \\ P(\theta_k^i | FA = m_k) P(FA = m_k | Y_k), & i = 0 \end{cases}$$

where  $\begin{cases} P(FA = m_k - 1 | Y_k) = \frac{P(Y_k | FA = m_k - 1) P(FA = m_k - 1)}{P(Y_k)} = \frac{P_D P_G \times \mu_F(m_k - 1)}{P(Y_k)}, & i = 1, \dots, m_k \\ P(FA = m_k | Y_k) = \frac{P(Y_k | FA = m_k) P(FA = m_k)}{P(Y_k)} = \frac{(1 - P_D P_G) \times \mu_F(m_k)}{P(Y_k)}, & i = 0 \end{cases}$

where,  $P(Y_k) = P_D P_G \times \mu_F(m_k - 1) + (1 - P_D P_G) \times \mu_F(m_k)$

$$= \begin{cases} \frac{1}{m_k} P_D P_G \times \mu_F(m_k - 1) P(Y_k)^{-1}, & i = 1, \dots, m_k \\ (1 - P_D P_G) \times \mu_F(m_k) P(Y_k)^{-1}, & i = 0 \end{cases}$$

# Probabilistic Data Association Filter (PDAF)

$$\beta_k^i = \begin{cases} \frac{1}{c_k} \left[ \frac{1}{V_k} m_{k-1} \frac{\mathcal{N}(z_k^i; 0, \hat{S}_k)}{\sum_{j=1}^{m_k} \mathcal{N}(z_k^j; 0, \hat{S}_k)} \frac{P_D P_G}{m_k} \mu_F(m_k - 1) \right], & i = 1, \dots, m_k \\ \frac{1}{c_k} \left[ \frac{1}{V_k} m_k (1 - P_D P_G) \mu_F(m_k) \right], & i = 0 \end{cases}$$

where,  $c_k = \frac{1}{V_k} m_k (1 - P_D P_G) \mu_F(m_k) + \sum_{i=1}^{m_k} \frac{1}{V_k} m_{k-1} \frac{\mathcal{N}(z_k^i; 0, \hat{S}_k)}{\sum_{j=1}^{m_k} \mathcal{N}(z_k^j; 0, \hat{S}_k)} \frac{P_D P_G}{m_k} \mu_F(m_k - 1)$

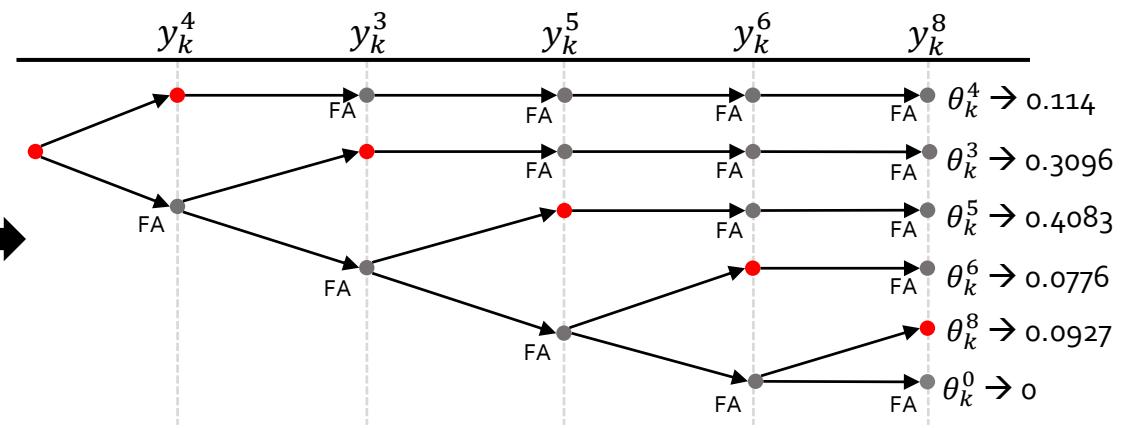
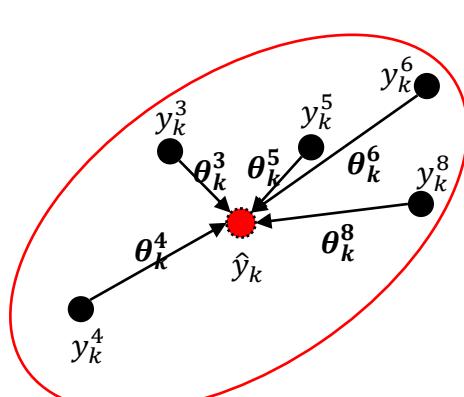
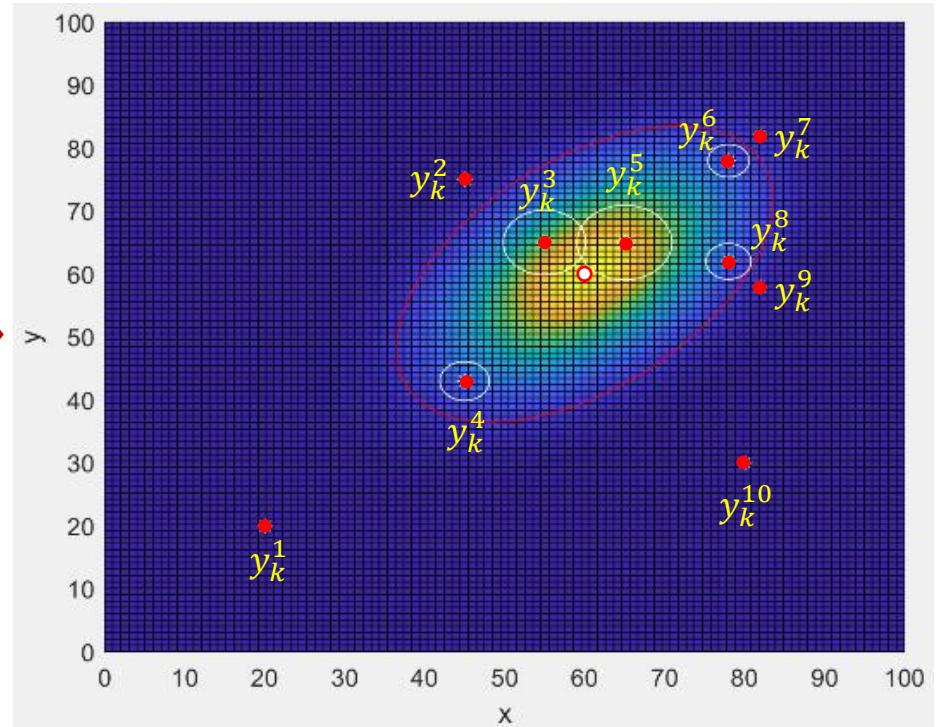
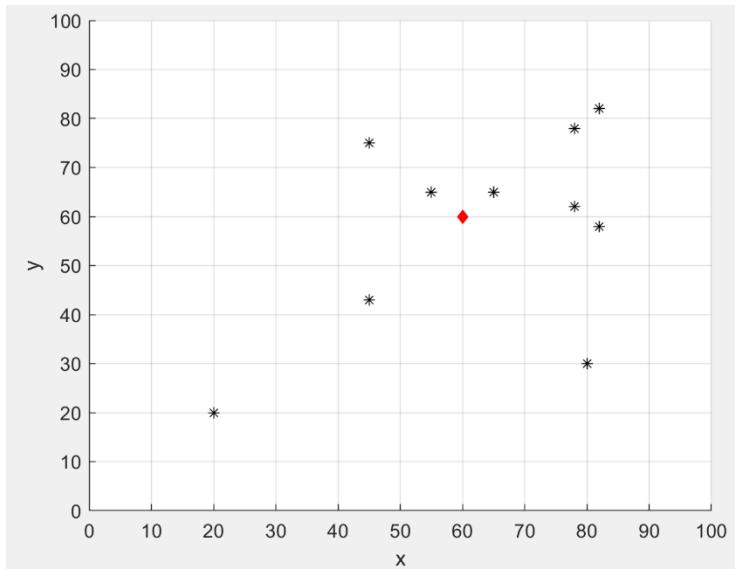
$z_k^i$  is from the track:

- pdf of 1 measurement and  $m_k - 1$  false alarms.
- Probability of 1 measurement being associated to the track x probability of  $m_k - 1$  false alarms.

None of the measurements are from the track:

- pdf of  $m_k$  false alarms.
- Probability that no measurements being associated to the track x probability of  $m_k$  false alarms.

	$y_k^0$	$y_k^1$	$y_k^2$	$y_k^3$	$y_k^4$	$y_k^5$	$y_k^6$	$y_k^7$	$y_k^8$	$y_k^9$	$y_k^{10}$
$\beta_k^i$	0	0	0	0.3096	0.1114	0.4083	0.0776	0	0.0927	0	0



# Probabilistic Data Association Filter (PDAF)

- Key idea
  - Update state using all validated measurements weighted by their likelihoods.

(1)

(2)

- Use weighted combination of the measurements:

- posterior state estimate:

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k \sum_{i=1}^{m_k} \beta_k^i (y_k - \hat{y}_k)$$

- posterior state uncertainty covariance:

$$P_{k|k} = \beta_k^0 P_{k|k-1} + (1 - \beta_k^0) [(I - K_k H_k) P_{k|k-1} (I - K_k H_k)^T + K_k \hat{S}_k K_k^T] + K_k \left[ \sum_{i=1}^{m_k} (\beta_k^i z_k^i z_k^{i^T}) - Z_k Z_k^T \right] K_k^T$$

- $K_k$  = Kalman gain
- $H_k$  = state-to-measurement mapping matrix
- $\hat{S}_k = H_k P_{k|k-1} H_k^T$  = Residual variance
- $Z_k = \sum_{i=1}^{m_k} \beta_k^i z_k = \sum_{i=1}^{m_k} \beta_k^i (y_k - \hat{y}_k)$  = innovation mtx

if  $\beta_k^0 = 1$ , i.e. 100% sure that **none** of the measurements belong to the track  
 → covariance of state does not get updated.

=  $P_{k|k}$ : cov. of the update if only 1 measurement is used.

Effects of incorrect measurement assignments  
 → increases covariance of the update. ( $\geq 0$ , PSD)

c.f.) When using a single best measurement (e.g. GNN):

- $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k (y_k - \hat{y}_k)$

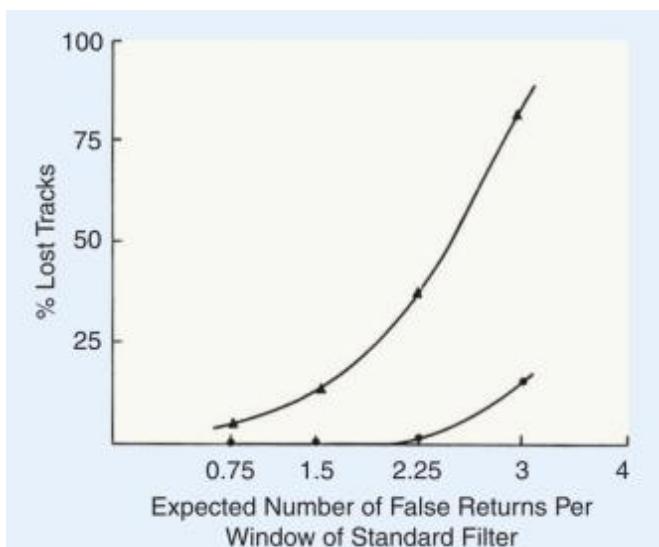
- $P_{k|k} = (I - K_k H_k) P_{k|k-1} (I - K_k H_k)^T + K_k \hat{S}_k K_k^T$  ← compare  
 or

$$P_{k|k} = (I - K_k H_k) P_{k|k-1} \text{ with optimal Kalman gain for MMSE estimates, } K_k = P_{k|k-1} H_k^T S_k^{-1}$$

- Although state estimation is linear, the filter is nonlinear because in state estimation update the association probability,  $\beta_k^i$ , is nonlinearly defined.

# Probabilistic Data Association Filter (PDAF)

	Nearest Neighbor Standard Filter (NNSF)	Probabilistic Data Association Filter (PDAF)
state prediction $\hat{x}_{k k-1}$		$\hat{x}_{k k-1} = F_{k-1}\hat{x}_{k-1 k-1}$
state cov. prediction $P_{k k-1}$		$P_{k k-1} = F_{k-1}P_{k-1 k-1}F_{k-1}^T + Q_{k-1}$
innovation covariance $\hat{S}_k$		$\hat{S}_k = H_k P_{k k-1} H_k^T$
Kalman gain $K_k$ (optimal)		$K_k = P_{k k-1} H_k^T \hat{S}_k^{-1}$
innovation $z_k$	$z_k = y_k - H_k \hat{x}_{k k-1}$	$z_k = \sum_{i=1}^{m_k} \beta_k^i (y_k^i - H_k \hat{x}_{k k-1})$
state cov. update $P_{k k}$	$P_{k k} = (I - K_k H_k) P_{k k-1}$	$P_{k k} = \beta_k^0 P_{k k-1} + (1 - \beta_k^0) P_{k k} + K_k \left[ \sum_{i=1}^{m_k} (\beta_k^i z_k^i z_k^{i^T}) - Z_k Z_k^T \right] K_k^T$



- **Computational requirement:**  
PDAF  $\approx 2 \times$  Standard KF
- **Robustness against false alarm in validation gate:**  
PDAF  $\gg$  NNSF
- **Suboptimal strategy**  
never totally wrong, but never totally right.
- **Multiple tracks?** Joint PDAF (JPDAF)

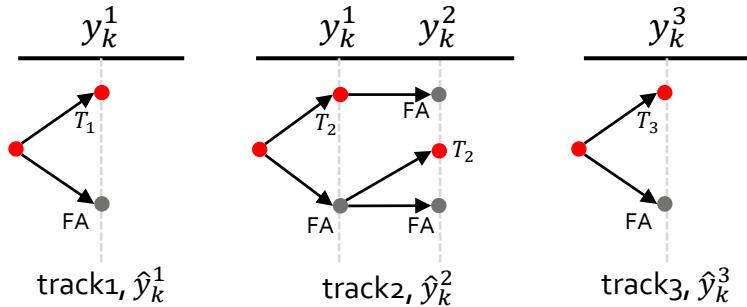
▲ NNSF (Nearest-Neighbor Standard Filter)  
● PDAF (Probabilistic Data-Association Filter)

# Data association techniques

- I. Non-Bayesian
  - 1. Nearest Neighbor Standard Filter (NNSF)
  - 2. Global Nearest Neighbor (GNN)
- II. Bayesian
  - 1. Probabilistic Data Association Filter (PDAF)
  - 2. **Joint Probabilistic Data Association Filter (JPDAF)**
  - 3. Multiple Hypothesis Tracking (MHT)

# Joint Probabilistic Data Association Filter (JPDAF)

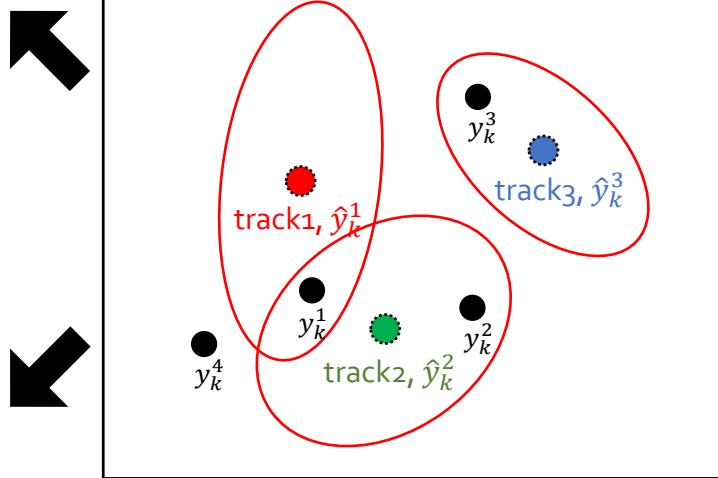
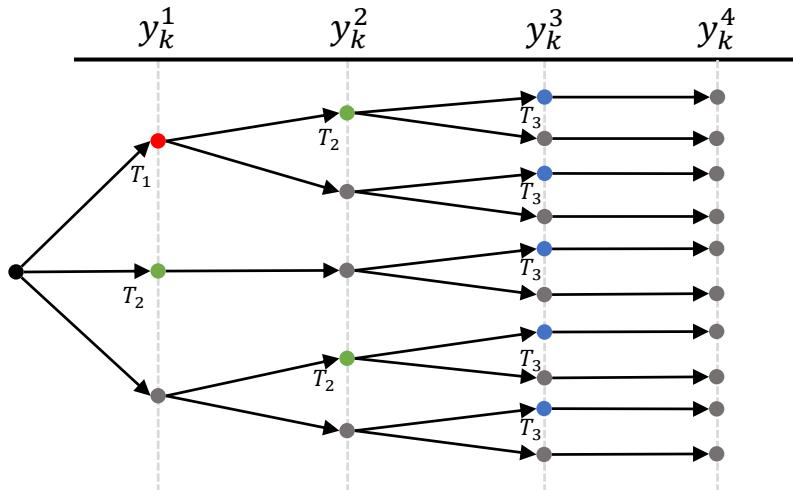
- Differs in the definition of association events and association probability.
- PDAF: disjoint trees for each track



Association events  $\theta_k^i$  decoupled.

$$\theta_k^i = \begin{cases} y_k^i \text{ belongs to the track.} & i = \{1, \dots, m_k\} \\ \text{none belongs to the track.} & i = 0 \end{cases}$$

- JPDAF: single tree for all measurements



Association events  $\theta_k^{T_j^i}$  coupled.

$$\theta_k = \bigcap_{i=1}^{m_k} \theta_k^{T_j^i}, j = 1, \dots, N_T$$

# Joint Probabilistic Data Association Filter (JPDAF)

- Probability of joint association event  $\theta_k$

$$P(\theta_k|Y^k) = P(\theta_k|Y_k, m_k, Y^{k-1}) = \frac{1}{c} \underbrace{f(Y_k|\theta_k, m_k, Y^{k-1})}_{1.} \underbrace{P(\theta_k|m_k)}_{2.}$$

Baye's formula:

$$P(\theta|Y) = \frac{f(Y|\theta)P(\theta)}{\sum f(Y|\theta_i)P(\theta_i)}$$

- $f(Y_k|\theta_k, m_k, Y^{k-1})$  = conditional pdf of measurements

- $V^{-m_k^F} \prod_{T \in \mathcal{T}_D} \mathcal{N}(z_k^{j_T}; 0, \hat{S}_k)$ ,  
where  $m_k^F$  = number of false alarms,  $V$  = scan area

- $P(\theta_k|m_k)$  = prior probability of a joint association event  $\theta_k$

$$\frac{\prod_{T \in \mathcal{T}_D} P_D^T \prod_{T \in \mathcal{T}_{ND}} (1 - P_D^T) \mu_F(m_k^F)}{m_k! \mathbf{P}_{m_k - m_k^F}} = \frac{m_k^F!}{m_k!} \prod_{T \in \mathcal{T}_D} P_D^T \prod_{T \in \mathcal{T}_{ND}} (1 - P_D^T) \mu_F(m_k^F)$$

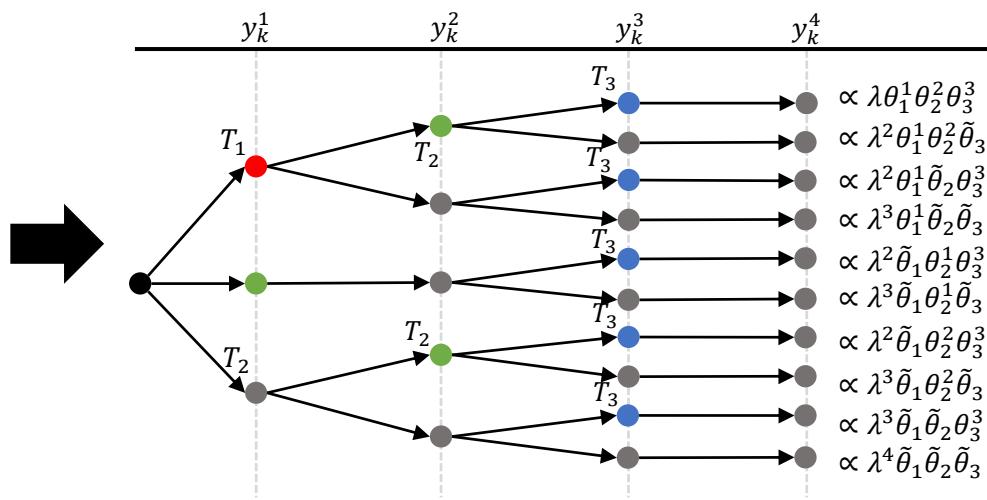
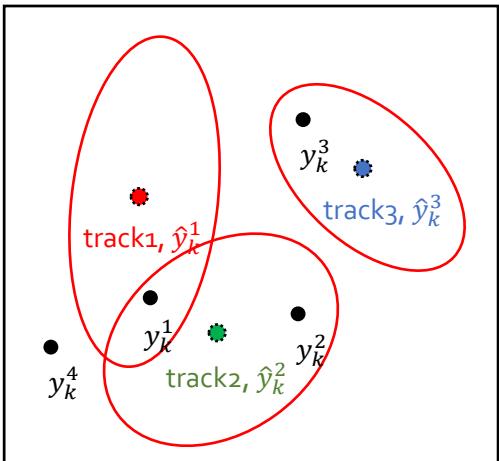
Number of combinations where non-false-alarm ( $m_k - m_k^F$ ) measurements out of total  $m_k$  measurements assigned to existing tracks, order matters.

$$\rightarrow m_k! \mathbf{P}_{m_k - m_k^F} = \frac{m_k!}{(m_k - (m_k - m_k^F))!} = \frac{m_k!}{m_k^F!}$$

- Result:

$$P(\theta_k|Y^k) = \frac{1}{c} \frac{m_k^F!}{m_k!} \mu_F(m_k^F) V^{-m_k^F} \prod_{T \in \mathcal{T}_D} P_D^T \mathcal{N}(z_k^{j_T}; 0, \hat{S}_k) \prod_{T \in \mathcal{T}_{ND}} (1 - P_D^T)$$

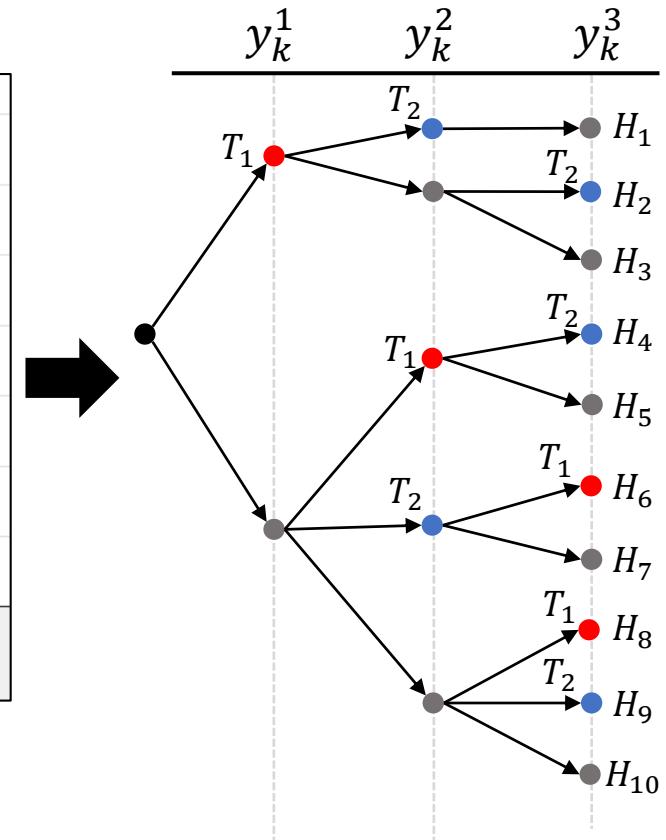
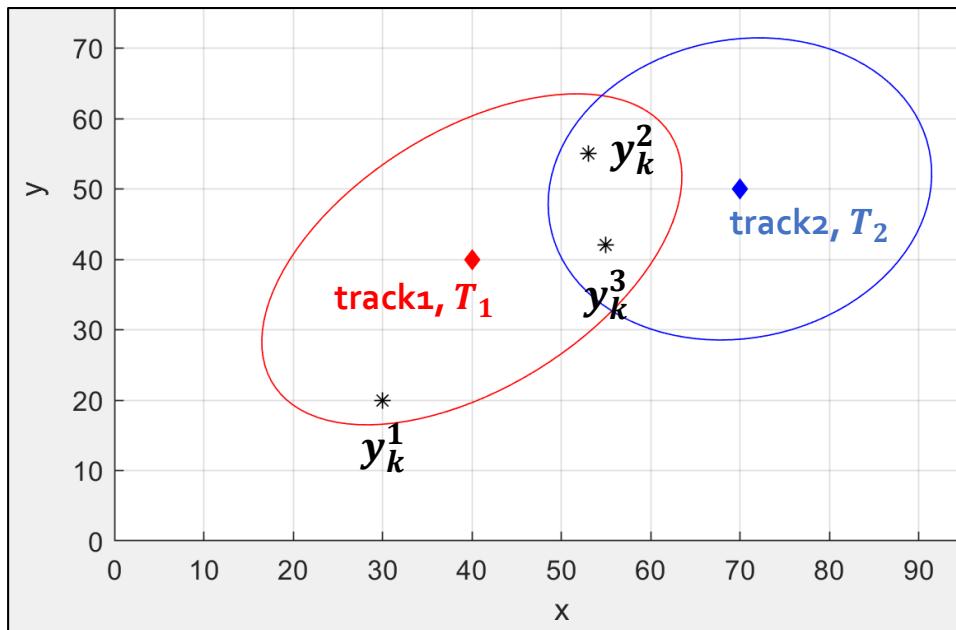
- $c$  = normalizing factor
- $P_D^T$  = track-specific detection probability of track  $T$
- $P_G^T$  = known gate probability of track  $T$
- $\mathcal{N}(z_k^{j_T}; 0, \hat{S}_k)$  =  $j_T^{th}$  measurement likelihood given track  $T$



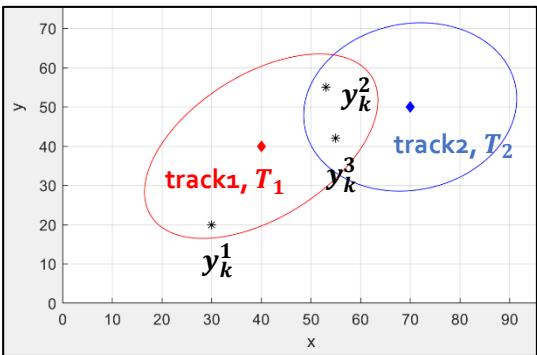
# Joint Probabilistic Data Association Filter (JPDAF)

$$P(\theta_k | Y^k) = \frac{1}{c} \frac{m_k^F!}{m_k!} \mu_F(m_k^F) V^{-m_k^F} \prod_{T \in \mathcal{T}_D} P_D^T \mathcal{N}(z_k^{j_T}; 0, \hat{S}_k) \prod_{T \in \mathcal{T}_{ND}} (1 - P_D^T)$$

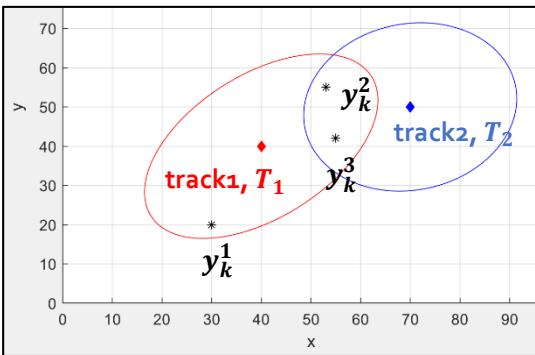
- Example



$y_k^1$	$y_k^2$	$y_k^3$		$H_i$	$T_1$	$T_2$	$p\{H_i\}$	$p\{H_i\}$ normalized
				$H_1$	1	2	$\lambda\theta_1^1\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_1^1\mathcal{N}_2^2P_{D,1}P_{D,2}$
				$H_2$	1	3	$\lambda\theta_1^1\theta_2^3$	$\mu_F(1)V^{-1}\mathcal{N}_1^1\mathcal{N}_2^3P_{D,1}P_{D,2}$
				$H_3$	1	-	$\lambda^2\theta_1^1\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^1P_{D,1}(1-P_{D,2})$
				$H_4$	2	3	$\lambda\theta_1^2\theta_2^3$	$\mu_F(1)V^{-1}\mathcal{N}_1^2\mathcal{N}_2^3P_{D,1}P_{D,2}$
				$H_5$	2	-	$\lambda^2\theta_1^2\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^2P_{D,1}(1-P_{D,2})$
				$H_6$	3	2	$\lambda\theta_1^3\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_1^3\mathcal{N}_2^2P_{D,1}P_{D,2}$
				$H_7$	-	2	$\lambda\tilde{\theta}_1\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_2^2(1-P_{D,1})P_{D,2}$
				$H_8$	3	-	$\lambda^2\theta_1^3\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^3P_{D,1}(1-P_{D,2})$
				$H_9$	-	3	$\lambda^2\tilde{\theta}_1\theta_2^3$	$\mu_F(2)V^{-2}\mathcal{N}_2^3(1-P_{D,1})P_{D,2}$
				$H_{10}$	-	-	$\lambda^3\tilde{\theta}_1\tilde{\theta}_2$	$\mu_F(3)V^{-3}(1-P_{D,1})(1-P_{D,2})$



$y_k^1$	$y_k^2$	$y_k^3$		$H_i$	$T_1$	$T_2$	$p\{H_i\}$	$p\{H_i\}$ normalized
				$H_1$	1	2	$\lambda\theta_1^1\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_1^1\mathcal{N}_2^2P_{D,1}P_{D,2}$ <b>0.333889</b>
				$H_2$	1	3	$\lambda\theta_1^1\theta_2^3$	$\mu_F(1)V^{-1}\mathcal{N}_1^1\mathcal{N}_2^3P_{D,1}P_{D,2}$ <b>0.267313</b>
				$H_3$	1	-	$\lambda^2\theta_1^1\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^1P_{D,1}(1-P_{D,2})$ <b>0.0</b>
				$H_4$	2	3	$\lambda\theta_1^2\theta_2^3$	$\mu_F(1)V^{-1}\mathcal{N}_1^2\mathcal{N}_2^3P_{D,1}P_{D,2}$ <b>0.177318</b>
				$H_5$	2	-	$\lambda^2\theta_1^2\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^2P_{D,1}(1-P_{D,2})$ <b>0.0</b>
				$H_6$	3	2	$\lambda\theta_1^3\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_1^3\mathcal{N}_2^2P_{D,1}P_{D,2}$ <b>0.22148</b>
				$H_7$	-	2	$\lambda\tilde{\theta}_1\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_2^2(1-P_{D,1})P_{D,2}$ <b>0.0</b>
				$H_8$	3	-	$\lambda^2\theta_1^3\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^3P_{D,1}(1-P_{D,2})$ <b>0.0</b>
				$H_9$	-	3	$\lambda^2\tilde{\theta}_1\theta_2^3$	$\mu_F(2)V^{-2}\mathcal{N}_2^3(1-P_{D,1})P_{D,2}$ <b>0.0</b>
				$H_{10}$	-	-	$\lambda^3\tilde{\theta}_1\tilde{\theta}_2$	$\mu_F(3)V^{-3}(1-P_{D,1})(1-P_{D,2})$ <b>0.0</b>



### Track 1

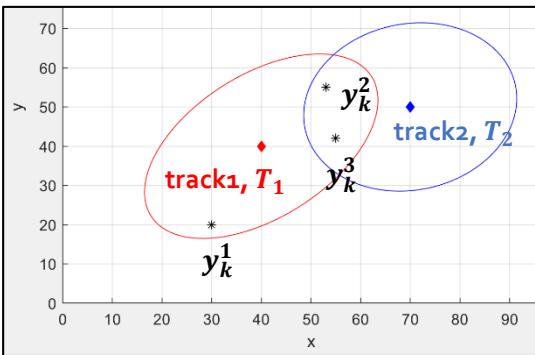
$$P_1^0 = \text{track 1 not assigned} = 0 + 0 + 0 = \mathbf{0}$$

$$P_1^1 = \text{track 1 to } y_k^1 = 0.333889 + 0.267313 + 0 = \mathbf{0.6012}$$

$$P_1^2 = \text{track 1 to } y_k^2 = 0.177318 + 0 = \mathbf{0.177318}$$

$$P_1^3 = \text{track 1 to } y_k^3 = 0.22148 + 0 = \mathbf{0.22148}$$

$y_k^1$	$y_k^2$	$y_k^3$		$H_i$	$T_1$	$T_2$	$p\{H_i\}$	$p\{H_i\}$ normalized
				$H_1$	1	2	$\lambda\theta_1^1\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_1^1\mathcal{N}_2^2P_{D,1}P_{D,2}$ <b>0.333889</b>
				$H_2$	1	3	$\lambda\theta_1^1\theta_2^3$	$\mu_F(1)V^{-1}\mathcal{N}_1^1\mathcal{N}_2^3P_{D,1}P_{D,2}$ <b>0.267313</b>
				$H_3$	1	-	$\lambda^2\theta_1^1\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^1P_{D,1}(1-P_{D,2})$ <b>0.0</b>
				$H_4$	2	3	$\lambda\theta_1^2\theta_2^3$	$\mu_F(1)V^{-1}\mathcal{N}_1^2\mathcal{N}_2^3P_{D,1}P_{D,2}$ <b>0.177318</b>
				$H_5$	2	-	$\lambda^2\theta_1^2\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^2P_{D,1}(1-P_{D,2})$ <b>0.0</b>
				$H_6$	3	2	$\lambda\theta_1^3\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_1^3\mathcal{N}_2^2P_{D,1}P_{D,2}$ <b>0.22148</b>
				$H_7$	-	2	$\lambda\tilde{\theta}_1\theta_2^2$	$\mu_F(1)V^{-1}\mathcal{N}_2^2(1-P_{D,1})P_{D,2}$ <b>0.0</b>
				$H_8$	3	-	$\lambda^2\theta_1^3\tilde{\theta}_2$	$\mu_F(2)V^{-2}\mathcal{N}_1^3P_{D,1}(1-P_{D,2})$ <b>0.0</b>
				$H_9$	-	3	$\lambda^2\tilde{\theta}_1\theta_2^3$	$\mu_F(2)V^{-2}\mathcal{N}_2^3(1-P_{D,1})P_{D,2}$ <b>0.0</b>
				$H_{10}$	-	-	$\lambda^3\tilde{\theta}_1\tilde{\theta}_2$	$\mu_F(3)V^{-3}(1-P_{D,1})(1-P_{D,2})$ <b>0.0</b>



## Track 2

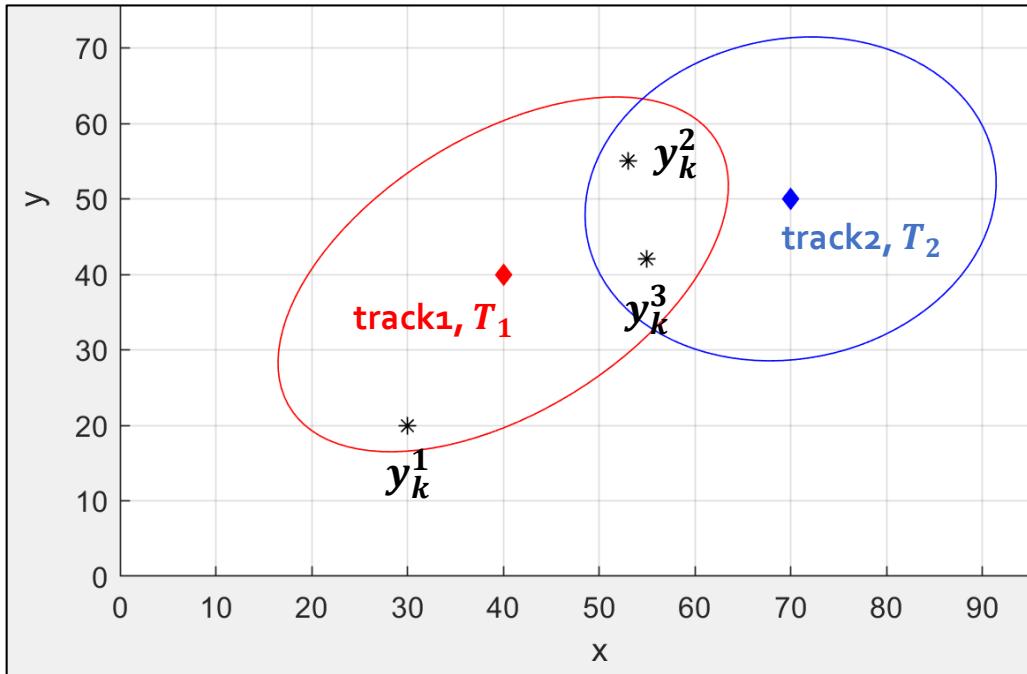
$$P_2^0 = \text{track 2 not assigned} = 0 + 0 + 0 + 0 = \mathbf{0}$$

$$P_2^1 = \text{track 2 to } y_k^1 = \mathbf{0} \leftarrow \text{outside validation region}$$

$$P_2^2 = \text{track 2 to } y_k^2 = 0.333889 + 0.22148 + 0 = \mathbf{0.5554}$$

$$P_2^3 = \text{track 2 to } y_k^3 = 0.22148 + 0 = \mathbf{0.4446}$$

# Joint Probabilistic Data Association Filter (JPDAF)



$$\begin{aligned} P_2^0 &= 0 \\ P_2^1 &= 0.6012 \\ P_2^2 &= 0.177318 \\ P_2^3 &= 0.22148 \end{aligned}$$

PDAF for track 1

$$\begin{aligned} P_2^0 &= 0 \\ P_2^1 &= 0 \\ P_2^2 &= 0.5554 \\ P_2^3 &= 0.4446 \end{aligned}$$

PDAF for track 2

PDAF for each track.

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k \sum_{i=1}^{m_k} \beta_k^i (y_k - \hat{y}_k)$$

$$P_{k|k} = \beta_k^0 P_{k|k-1} + (1 - \beta_k^0) [(I - K_k H_k) P_{k|k-1} (I - K_k H_k)^T + K_k \hat{S}_k {K_k}^T] + K_k \left[ \sum_{i=1}^{m_k} (\beta_k^i z_k^i {z_k^i}^T) - \mathbf{Z}_k \mathbf{Z}_k^T \right] {K_k}^T$$

## • Implementation perspective

1. Build Validation Gate matrix  $\bar{\Omega}$ .

$$(y_k^i - \hat{y}_k)^T \hat{S}_k^{-1} (y_k^i - \hat{y}_k) < \chi^2$$

$$\rightarrow \bar{\Omega} = \begin{bmatrix} FA & T_1 & T_2 \\ 1 & 1 & 0 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} y_k^1 \\ y_k^2 \\ y_k^3 \end{bmatrix}$$

2. Expand to a set of event matrices  $\Omega_X$ .

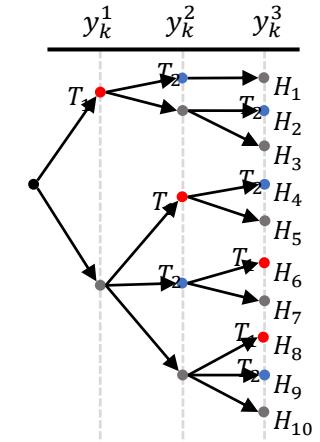
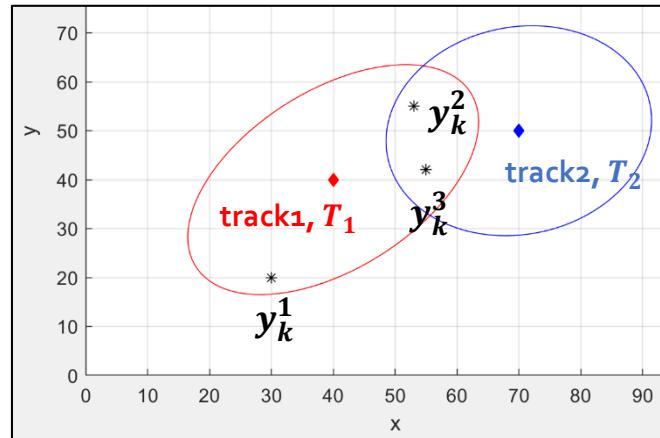
$$\Omega_1 = \begin{bmatrix} FA & T_1 & T_2 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix} y_k^1 \quad \Omega_6 = \begin{bmatrix} FA & T_1 & T_2 \\ 1 & 0 & 0 \\ 0 & 0 & 1 \\ 0 & 1 & 0 \end{bmatrix} y_k^1$$

$$\Omega_2 = \begin{bmatrix} 0 & 1 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 1 \end{bmatrix} y_k^1 \quad \Omega_7 = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix} y_k^1$$

$$\Omega_3 = \begin{bmatrix} 0 & 1 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \end{bmatrix} y_k^1 \quad \Omega_8 = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix} y_k^1$$

$$\Omega_4 = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} y_k^1 \quad \Omega_9 = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 1 \end{bmatrix} y_k^1$$

$$\Omega_5 = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 1 & 0 & 0 \end{bmatrix} y_k^1 \quad \Omega_{10} = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \end{bmatrix} y_k^1$$



3. For each event matrix, compute joint probability using

$$P(\theta_k | Y^k) = \frac{1}{c} \frac{m_k^F}{m_k!} \mu_F(m_k^F) V^{-m_k^F} \prod_{T \in \mathcal{T}_D} P_D^T \mathcal{N}(z_k^{j_T}; 0, \hat{S}_k) \prod_{T \in \mathcal{T}_{ND}} (1 - P_D^T)$$

→ Results:  $P(\Omega_X) =$

$$[0.333889, 0.267313, 0, 0.177318, 0, 0.22148, 0, 0, 0, 0]$$

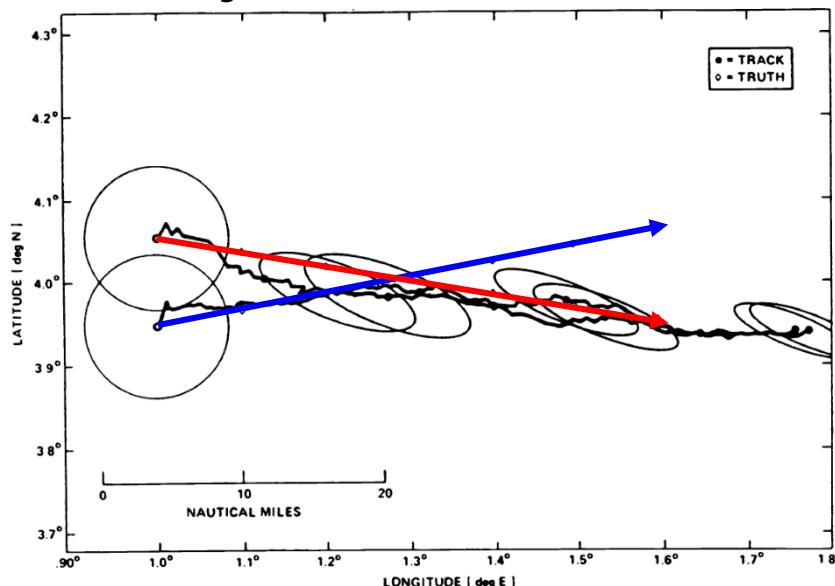
4. Update state estimation for each track as in PDAF.

$$\beta_{jT}(k) = P\{\theta_{jT}(k) | Y^k\} = \sum_{\theta: \theta_{jT} \in \theta} P\{\theta(k) | Y^k\}$$

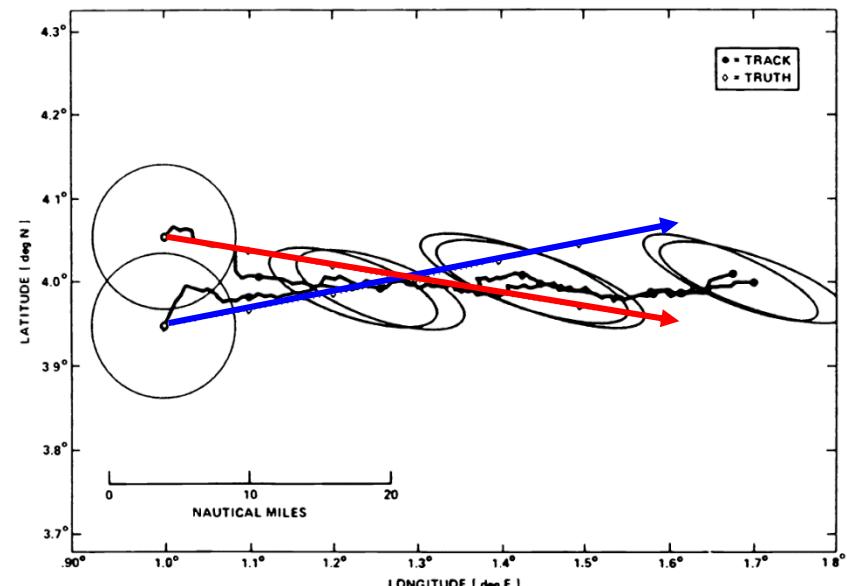
- $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k \sum_{i=1}^{m_k} \beta_k^i (y_k - \hat{y}_k)$
- $P_{k|k} = \beta_k^0 P_{k|k-1} + (1 - \beta_k^0) [(I - K_k H_k) P_{k|k-1} (I - K_k H_k)^T + K_k \hat{S}_k {K_k}^T] + K_k \left[ \sum_{i=1}^{m_k} (\beta_k^i z_k^i {z_k^i}^T) - \mathbf{Z}_k \mathbf{Z}_k^T \right] {K_k}^T$

- Comparisons

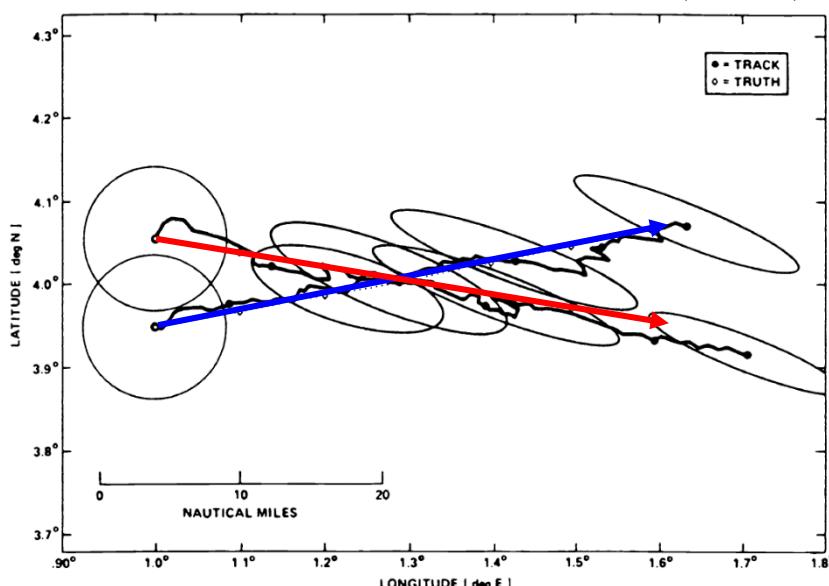
Nearest Neighbor Standard Filter (NNSF)



Probabilistic Data Association Filter (PDAF)



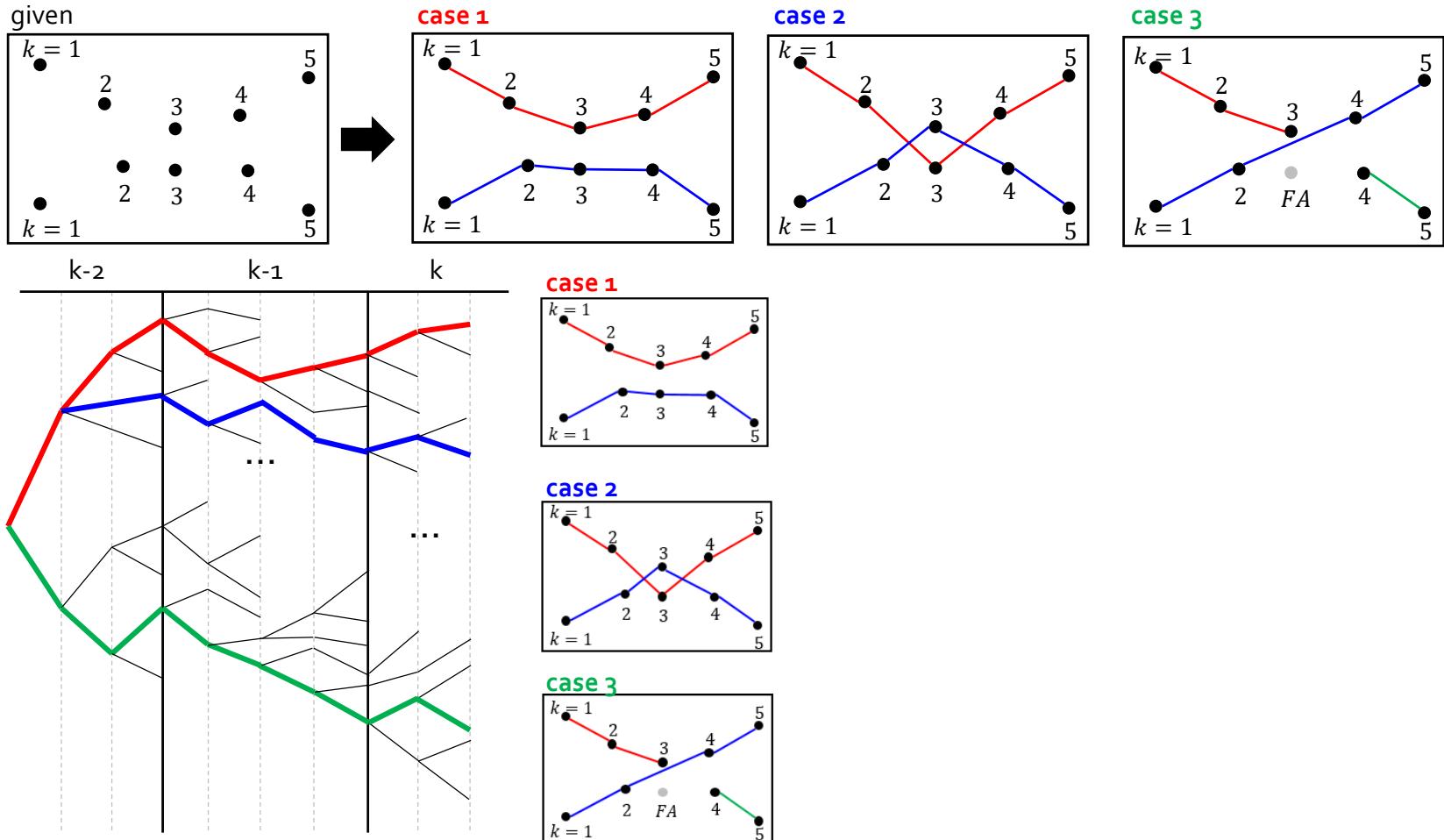
Joint Probabilistic Data Association Filter (JPDAF)



- JPDAF more robust for multiple object tracking.

# Multiple Hypothesis Tracking (MHT)

- Resolves assignment ambiguities by delaying measurement-to-track decisions.



- Two main methods:
  - Hypotheses Oriented MHT (HOMHT)
  - Target Oriented MHT (TOMHT) → multi-dimensional assignment problem.

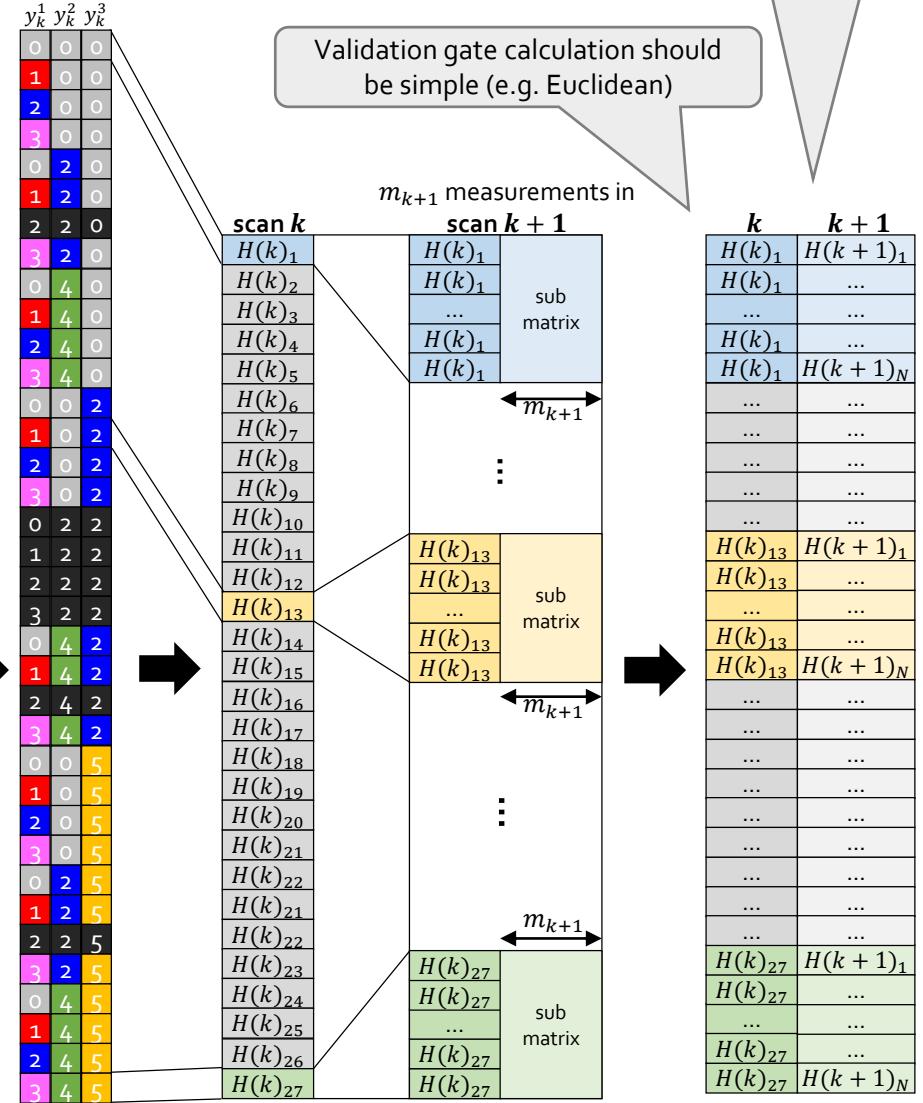
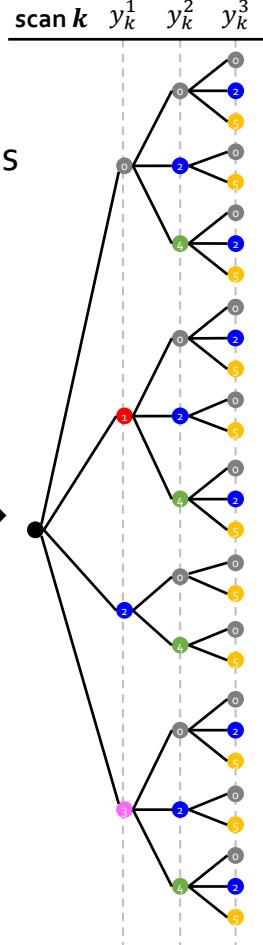
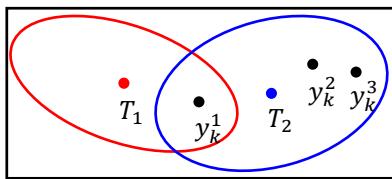
# Multiple Hypothesis Tracking (MHT)

## 1. Hypotheses Oriented MHT (HOMHT)

- Expand existing hypotheses.
- Form different hypotheses considering each measurement as:

1. new track
2. false alarm
3. existing track

- Each measurement makes up at least 2 hypotheses:
  1. new track
  2. false alarm



# Multiple Hypothesis Tracking (MHT)

- For each hypothesis,

- State and state covariance update is performed in parallel.

- $\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k(y_k - \hat{y}_k)$
- $P_{k|k} = (I - K_k H_k) P_{k|k-1} (I - K_k H_k)^T + K_k \hat{S}_k K_k^T$

- Hypothesis probability is calculated as well.

$$P(\Theta^{k,H} | Y^k)$$

$$= \frac{1}{c} \frac{m_k^F! m_k^V!}{m_k!} \mu_F(m_k^F) \mu_V(m_k^V) V^{-m_k^F - m_k^V} \prod_{i=1}^{m_k} \mathcal{N}(z_k^{i_T}; 0, \hat{S}_k) \left\{ \text{assigned} \right\} \prod_{T \in \mathcal{T}_D} P_D^T \left\{ \text{detected} \right\} (1 - P_D^T) \left\{ \text{~detected} \right\} P(\Theta^{k-1,H} | Y^{k-1})$$

Resulting from

$$m_k P_{(m_k - m_k^F - m_k^V)} \times (m_k^F + m_k^V) C_{m_k^V}$$

pdf of new track appearing (assumption)

Hypothesis probability from the last (parent) scan

- $m_k^V$  = number of new target (based on hypothesis)
- $m_k^V$  = number of false alarms

- Make decision using past  $N$  hypotheses ( $N$  = window size, e.g.  $N = 3$ )

- Reduce hypotheses  $\rightarrow$  if corresponding hypothesis probability is low, discard it.
- For each hypotheses from each past scan, loop each column and make decision on track assignment
  - case i)
  - case ii)

$y_k^1$	$y_k^2$	$y_k^3$	$H_k^1$
1	2	0	$H_k^1$
1	4	0	$H_k^2$
1	4	2	$H_k^3$
1	0	5	$H_k^4$
1	2	5	$H_k^5$
1	4	5	$H_k^6$

$\rightarrow$  can confirm  $y_k^1$  belongs to track 1

$y_k^1$	$y_k^2$	$y_k^3$	$H_k^1$
1	4	2	$H_k^1$
1	2	5	$H_k^2$
1	0	5	$H_k^3$
1	0	2	$H_k^4$
2	0	2	$H_k^5$
2	0	5	$H_k^6$

$P_1 = \sum_{i=1}^4 P(H_k^i)$   
 $P_2 = \sum_{i=5}^6 P(H_k^i)$   
 $\rightarrow$  If  $P_1 > P_2$  and  $P_1 > \text{thres}$   
 $\rightarrow$  can confirm  $y_k^1$  belongs to track 1

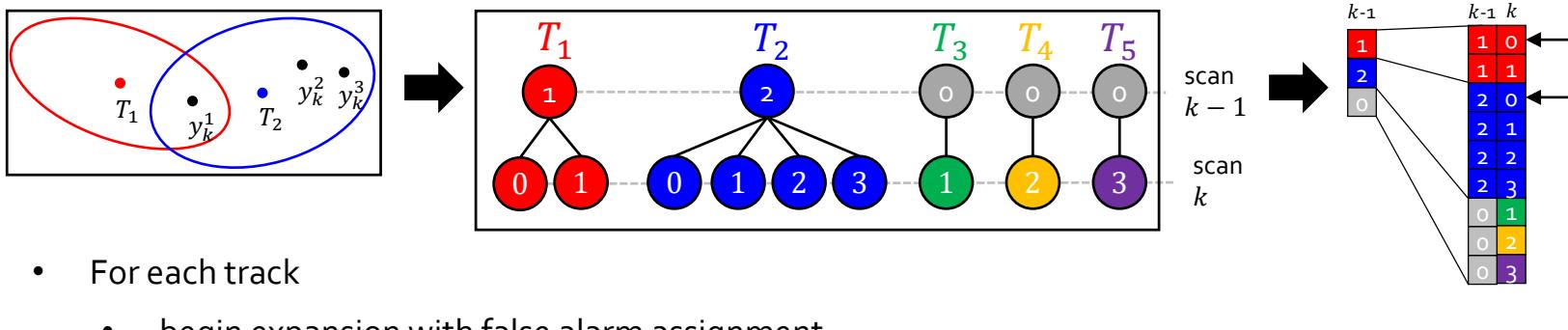
- Update track with:

- most probable hypothesis  $\rightarrow$  GNN approach
- use weighted sum using track probabilities, in order to combine the estimates from all the hypotheses that contain the track  $\rightarrow$  JPDAF approach

# Multiple Hypothesis Tracking (MHT)

## 2. Target Oriented MHT (TOMHT) → door to advanced multiple object tracking techniques

- Instead of expanding on all hypotheses, expand on all tracks.



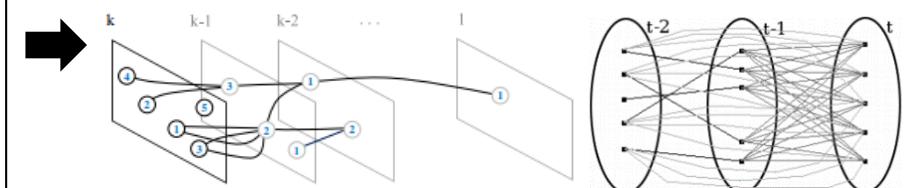
- For each track
  - begin expansion with false alarm assignment
  - Calculate track probability in parallel  
→ similar to HOMHT
- Find the best measurement-to-track match across a window size:  
→ Since 1 measurement cannot be assigned to 2 different tracks at each scan,  
**multi-dimensional assignment problem** needs to be solved:

For example,

$k-4$	$k-3$	$k-2$	$k-1$	$k$	
1	1	1	0	0	$P(T_1^1)$
1	1	1	0	2	$P(T_1^2)$
1	1	1	1	1	$P(T_1^3)$
1	1	2	1	3	$P(T_1^4)$
1	2	3	3	0	$P(T_1^5)$
1	2	3	3	4	$P(T_1^6)$
2	2	2	4	0	$P(T_2^1)$
2	2	2	4	5	$P(T_2^2)$
2	3	2	4	0	$P(T_2^3)$
2	3	2	4	5	$P(T_2^4)$

Assume, the top two highest track probabilities are:  
 $\mathbf{P}(T_1^4)$  and  $\mathbf{P}(T_2^2)$   
→ conflict occurs in measurement-to-track matching at scan  $k-2$ .  
→ Highest track probability does not mean the best measurement-to-track assignment.

- "The Lagrangian relaxation method for solving integer programming problems.", Fisher, 1981
- "The maximum weight independent set problem for data association in multiple hypothesis tracking.", Papageorgiou, 2009



# Further Researches

- View Multiple Objet Tracking as:
  - **Linear Programming:**
    1. "Application of 0-1 integer programming to multitarget tracking problems.", Morefield, 1977
    2. "An efficient implementation of Reid's multiple hypothesis tracking algorithm and its evaluation for the purpose of visual tracking.", Cox, 1996
    3. "An LP-based algorithm for the data association problem in multitarget tracking.", Storms, , 2003
    4. "A linear programming approach for multiple object tracking.", Jiang, 2007.
  - **Multi-dimensional assignment problem**
    1. "An applications oriented guide to Lagrangian relaxation.", Fisher, 1985
    2. "The maximum weight independent set problem for data association in multiple hypothesis tracking", Dimitri, 2009
  - **Dynamic Programming**
    1. "Finding the best set of K paths through a trellis with application to multitarget tracking.", Wolf, 1989
    2. "Robust people tracking with global trajectory optimization.", Berclaz, 2006.
    3. "Globally-optimal greedy algorithms for tracking a variable number of objects.", Pirsavash, 2011
  - **Algorithmic**
    1. "Markov chain Monte Carlo data association for multi-target tracking.", Oh, 2009

# MHT Application

- **"Spatially grounded multi-hypothesis tracking of people."** [ICRA2009]

Luber, Matthias, Gian Diego Tipaldi, and Kai O. Arras. "Spatially grounded multi-hypothesis tracking of people." *Proceedings of the ICRA Workshop on People Detection and Tracking*. Vol. 104. 2009.

- Key idea:

- MHT assumes that new track and false alarm events are uniformly distributed in the sensor field of view  $V$  with fixed Poisson rates,  $\lambda$ .  
→ only justifiable for settings which MHT has been originally developed (e.g., radar or underwater sonar). However, in the context of people tracking with vision or laser these models are overly simplified.

$$P(\boldsymbol{\Theta}^{k,H} | Y^k) = \frac{1}{c} \frac{m_k^F! m_k^V!}{m_k!} u_F(m_k^F) \mu_V(m_k^V) V^{-m_k^F - m_k^V} \prod_{i=1}^{m_k} \mathcal{N}(z_k^{i_T}; 0, \hat{S}_k)^{\{\text{assigned}\}} \prod_{T \in \mathcal{T}_D} P_D^T \{\text{detected}\} (1 - P_D^T)^{\{\sim \text{detected}\}} P(\boldsymbol{\Theta}^{k-1,H} | Y^{k-1})$$

- Incorporate learned distributions to MHT
- How?
  1. Re-design Poisson distribution from  $\lambda$  to  $\lambda(\vec{x}, t)$ , where  $\vec{x} \in \mathbb{R}^2$ .  
→  $\lambda(\vec{x}, t)$  = spatio-temporal distribution of events
  2. Learn spatio-temporal distribution,  $\lambda(\vec{x}, t)$ , of events as tracking continues.

# MHT Application

1. Re-design Poisson distribution from  $\lambda$  to  $\lambda(\vec{x}, t)$ , where  $\vec{x} \in X$ .

- Let  $N(t)$  be the number of events occurring up to time  $t$  with rate  $\lambda$ , then  $N(t)$  follows Poisson distribution:

$$\mu(N(t) = k) = e^{-\lambda t} \frac{(\lambda t)^k}{k!}$$

- Here, let  $\lambda$  be the generalized *rate function*  $\lambda(t)$

- Then, expected number of events between time  $a$  and  $b$  is:

$$\lambda_{a,b} = \int_a^b \lambda(t) dt$$

- Then, introduce a spatial dependency on the *rate function*  $\lambda(\vec{x}, t)$

- Then, expected number of events at time  $t$  in subspace  $S \in X$  is:

$$\lambda_S(t) = \int_S \lambda(\vec{x}, t) d\vec{x}$$

- In which case,  $\vec{x}$  and  $t$  are decoupled in the generalized rate function  $\rightarrow \lambda(\vec{x}, t) = f(\vec{x})\lambda(t)$

- $\lambda(t)$  = occurrence rate of events

- $f(\vec{x})$  = probability distribution on where the event occurs in space.

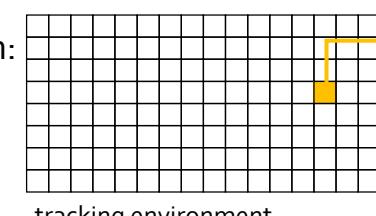
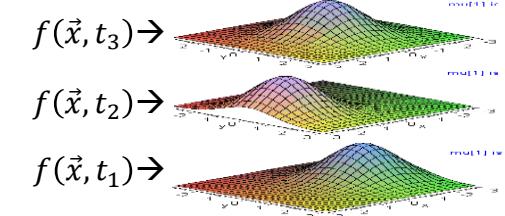
- Also, give constraint on  $f(\vec{x}) \rightarrow \int_X f(\vec{x})d\vec{x} = 1$

- Subdivide tracking environment into discrete cells

- Each cell represents a local homogeneous Poisson process with a fixed-rate over time,  $P_{ij}(k) = e^{-\lambda_{ij}} \frac{(\lambda_{ij})^k}{k!}$  where  $\lambda_{ij}$  is constant over time.

- Finally, the generalized rate function  $\lambda(\vec{x}, t)$  using grid approximation:

$$\lambda(\vec{x}, t) \cong \sum_{(i,j) \in X} \lambda_{ij} \mathbf{1}_{ij}(\vec{x}), \text{ where } \mathbf{1}_{ij}(\vec{x}) = \begin{cases} 1, & x \in \text{cell}_{ij} \\ 0, & x \notin \text{cell}_{ij} \end{cases}$$



# MHT Application

2. Learn Poisson distribution parameters as tracking goes on  
= learn rate function  $\lambda(\vec{x}, t)$

- Model  $\lambda$  using *Gamma* distribution, as it is a *conjugate prior* of Poisson distribution.  
i.e.  $\lambda$  is distributed according to *Gamma* density:  $\text{Gamma}(\lambda; \alpha, \beta) = \frac{\beta^\alpha}{(\alpha-1)!} \lambda^{\alpha-1} e^{-\beta\lambda}$
- Learning  $\lambda$  = estimating parameters of Gamma distribution,  $\alpha$  and  $\beta$ .
- Bayes' rule  $\rightarrow P(\lambda_i | K_{1..i}) \propto \underset{\text{posterior}}{P(k_i | \lambda_{i-1})} \underset{\text{likelihood}}{P(\lambda_{i-1})} \underset{\text{prior}}{P(\lambda_i)}, i = \text{time index}$ 
  - $P(\lambda_{i-1}) = \text{Gamma}(\alpha_{i-1}, \beta_{i-1}), \text{ prior}$
  - $P(k_i | \lambda_{i-1}) = P(k_i) = e^{-\lambda_i} \frac{(\lambda_i)^{k_i}}{k_i!}, \text{ likelihood}$

*conjugate prior*: Poisson distribution and Gamma distribution is in the same family of distributions  
e.g. if  $\alpha = k + 1, \beta = 1 \rightarrow$  Poisson distribution:  $e^{-\lambda} \frac{(\lambda)^k}{k!}$

- By substitution, turns out that
  - $\alpha_i = \alpha_{i-1} + k_i$
  - $\beta_i = \beta_{i-1} + 1$
  - start out with  $\alpha_0 = \beta_0 = 1$
- Expected value (mean) of *Gamma* distribution,  $\text{Gamma}(\alpha, \beta)$ , is:  $\frac{\alpha}{\beta}$   
 $\rightarrow$  i.e.  $E(\lambda) = \text{event occurring ratio} = \frac{\text{number of events occurring}}{\text{number of measurements}} = \frac{\alpha}{\beta}$

# Thank you

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