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## Abstract

*Keywords:*

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## 1. Introduction

The source localization technique has been studied for several decades and applied in many practical scenarios to provide the location information, for instance, the Location-Based Services (LBS) in communication systems, objects tracking in Internet of Things (IoT) and targets detection in radar system. The classical localization methods are two-step processing. Firstly, intermediate parameters that rely on the location of the sources, like the angle of arrival (AOA), time difference of arrival (TDOA), Doppler frequency shift (DFS) or received signal strength (RSS), are estimated from the measurements of different receivers. The locations of sources are, then, estimated based on these intermediate parameters. The two-step methods are sub-optimal since they measure the intermediate parameters at each receiver separately and independently, which ignore the constraint that the measurements from different receivers have to correspond to the same source position. The Direct Position Determination (DPD) approaches have recently been proposed as a single-step localization

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technique where the source locations are estimated by minimizing a single cost function, into which all received data enter jointly. It has been verified that the DPD methods outperform the two-step methods especially at low Signal to Noise Ratio (SNR) and the two methods' performance converge together as  
20 SNR increases. Furthermore, an additional data association step (to partition the intermediate parameters corresponding to the same source into one set) is required in multiple sources scenario, which is avoided inherently in DPD methods.

In multiple sources scenario, the exact Maximum Likelihood (ML) estimator  
25 can be derived but it requires a multi-dimensional search which is usually impractical. To sidestep the multi-dimensional search in multiple sources scenario, several existing DPD methods abandon the exact ML estimator and resort to the subspace-based technique or the filtering technique such as the well-known Multiple Signal Classification (MUSIC) [1], the Match Filter (M-  
30 F) [2] and the Minimum Variance Distortionless Response (MVDR) filter [3]. These techniques transform the multiple sources localization problem into single source localization problem which requires only two- or three-dimensional space search. These methods, however, require enough observation samples to achieve the asymptotically optimal performance at high SNR. To be specific,  
35 the MUSIC-DPD [1] and MVDR-DPD [3] both need a lots of observation samples to approximate the covariance matrix by the sample covariance matrix and the MF-DPD [2] also need enough observation samples to decouple the problem based on the independence among transmitted signals. While the number of observation samples is often limited for some practical reasons such as the  
40 request for real-time processing or the assumption that sources are stationary during the observation. The performance of these estimators will deteriorate when there are only a few observation samples available. Besides, the iterative ML methods [4] is proposed to achieve the exact ML solution efficiently. As in many iterative algorithms, the convergence to the global minimum of the cost  
45 function, however, depends on the initial estimate.

In this paper, we resort to the optimization framework [5], consisting of the

Pincus' theorem [6] and the Importance Sampling (IS) technique, to efficiently approximate the global solution of the multi-dimensional ML problem, using a few observation samples. The major difficulty to use IS technique consists in  
50 generating the alternative parameter realizations (i.e., sampling) according to a given (multi-dimensional) pdf. Much like all the IS-based works mentioned in section 4, we design a factorable pdf on all involved parameters, which allows a very easy generation of the required parameter realizations. Numerical simulations show that the proposed IS-based multidimensional ML estimator  
55 outperforms the subspace-based DPD techniques and the filtering-based DPD techniques in multiple sources scenario.

## 2. Problem Formulation

Given  $Q$  adjacent narrow-band emitters at  $\mathbf{p}_q \in \mathbb{R}^{D \times 1}, q = 1, \dots, Q$  (In general,  $D = 2$  for plane geometry or  $D = 3$  for solid geometry) and  $N_r$  space separated receivers which intercept the transmitted signals. Every receiver is equipped with an antenna array consisting of  $M$  elements. To avoid the angle ambiguity, the carrier frequency is small than the inverse of the propagation time over the array aperture. The complex envelopes of the signals observed by the  $j$ th receiver are given by

$$\mathbf{y}_j(t) = \sum_{q=1}^Q \alpha_{q,j} \mathbf{a}_j(\mathbf{p}_q) s_q(t - \tau_j(\mathbf{p}_q) - \hat{t}_q) + \mathbf{w}_j(t), \quad 0 \leq t \leq T \quad (1)$$

where  $\mathbf{y}_j(t)$  is the  $j$ th received signal,  $\alpha_{q,j}$  is an unknown complex scalar representing the channel attenuation between the  $q$ th emitter and the  $j$ th receiver,  $\mathbf{a}_j(\mathbf{p}_q)$  is the  $j$ th array response to the signal transmitted from position  $\mathbf{p}_q$  and  $s_q(t - \tau_j(\mathbf{p}_q) - \hat{t}_q)$  is the copy of the  $q$ th transmitted signal waveform, which is transmitted at time  $\hat{t}_q$  and is received by the  $j$ th receiver after being delayed by  $\tau_j(\mathbf{p}_q)$ . The vector  $\mathbf{w}_j(t)$  represents noise and interference observed by the  $j$ th array. The received signal can be partitioned into  $L$  sections and the length of each section satisfies  $T/L \gg \max_{j,q} \tau_j(\mathbf{p}_q)$  to avoid the delay ambiguity. The emitters are assumed to be stationary during the total observation time  $T$ . Each

section can be Discrete Fourier Transformed (DFT) into

$$\mathbf{y}_j(n, l) = \sum_{q=1}^Q \alpha_{q,j} \mathbf{a}_j(\mathbf{p}_q) e^{-j2\pi f_n(\tau_j(\mathbf{p}_q) + \hat{t}_q)} s_q(n, l) + \mathbf{w}_j(n, l) \quad (2)$$

$$j = 1, \dots, N_r; n = 1, \dots, N; l = 1, \dots, L$$

where  $\mathbf{y}_j(n, l)$  is the  $n$ th Fourier coefficient of the  $l$ th section of the  $j$ th received signal,  $s_q(n, l)$  is the  $n$ th Fourier coefficient of the  $l$ th section of the  $q$ th transmitted signal waveform, and  $\mathbf{w}_j(n, l)$  represents the  $n$ th Fourier coefficient of the  $l$ th section of the noise waveform.

For compact representation, we define the following vectors:

$$\boldsymbol{\gamma}_{j,q,n,l}(\mathbf{p}_q, \hat{t}_q | s_q(n, l)) \triangleq \mathbf{a}_j(\mathbf{p}_q) e^{-j2\pi f_n(\tau_j(\mathbf{p}_q) + \hat{t}_q)} s_q(n, l) \quad (3)$$

Here we assume that the spectrum of transmitted signals, like the synchronization and training sequences in cellular systems, are known to location system but the transmitted time  $\hat{t}_q$ , which is assumed known in some DPD works [1, 4], is assumed unknown here (for the precise clock synchronization between emitter and receiver is a challenging task). We observe that all information about the emitter's position is embedded in the manifold vector  $\boldsymbol{\gamma}_{j,q,n,l}$ . The compact harmonic-superposition form of (2) is given by:

$$\mathbf{y}_j(n, l) = \sum_{q=1}^Q \alpha_{q,j} \boldsymbol{\gamma}_{j,q,n,l} + \mathbf{w}_j(n, l) \quad (4)$$

which can be rewritten as the matrix form of the *Nonlinear regression* model described in [5],

$$\mathbf{y}_j(n, l) = \boldsymbol{\Gamma}_{j,n,l}(\mathbf{P}, \hat{\mathbf{t}}) \boldsymbol{\alpha}_j + \mathbf{w}_j(n, l) \quad (5)$$

where

$$\boldsymbol{\Gamma}_{j,n,l}(\mathbf{P}, \hat{\mathbf{t}}) \triangleq [\boldsymbol{\gamma}_{j,q,n,l}, \dots, \boldsymbol{\gamma}_{j,q,n,l}] \in \mathbb{C}^{M \times Q} \quad (6)$$

$$\boldsymbol{\alpha}_j \triangleq [\alpha_{1,j}, \dots, \alpha_{Q,j}]^T \in \mathbb{C}^{Q \times 1}$$

$\mathbf{P}, \hat{\mathbf{t}}$  are the sets of emitters' positions and transmitted times respectively. Collecting and stacking all the  $N$  coefficients of the  $l$ th section of the  $j$ th received

signal,

$$\mathbf{y}_j(l) = \mathbf{\Gamma}_{j,l}(\mathbf{P}, \mathring{\mathbf{t}}) \boldsymbol{\alpha}_j + \mathbf{w}_j(l) \quad (7)$$

where

$$\begin{aligned} \mathbf{y}_j(l) &\triangleq [\mathbf{y}_j^T(1, l), \dots, \mathbf{y}_j^T(N, l)]^T \in \mathbb{C}^{NM \times 1} \\ \mathbf{\Gamma}_{j,l}(\mathbf{P}, \mathring{\mathbf{t}}) &\triangleq [\mathbf{\Gamma}_{j,1,l}^T(\mathbf{P}, \mathring{\mathbf{t}}), \dots, \mathbf{\Gamma}_{j,N,l}^T(\mathbf{P}, \mathring{\mathbf{t}})]^T \in \mathbb{C}^{NM \times Q} \\ \mathbf{w}_j(l) &\triangleq [\mathbf{w}_j^T(1, l), \dots, \mathbf{w}_j^T(N, l)]^T \in \mathbb{C}^{NM \times 1} \end{aligned} \quad (8)$$

The problem that we address now is stated as follow: Given the measurements  $\{\mathbf{y}_j(l)\}_{j,l=1}^{N_r, L}$ , how to efficiently and accurately estimate the locations of the emitters  $\mathbf{P}$ .

### 65 3. the Concentrated Likelihood Function

In this subsection, we will derive the *Concentrated* Likelihood Function (CLF) that depends on the parameters of interest  $\mathbf{P}$  and the auxiliary parameters  $\mathring{\mathbf{t}}$ . The transmitted times  $\mathring{\mathbf{t}}$  are conducive to acquiring the location information from the propagation delay  $\tau_j(\mathbf{p})$  since the receivers are assumed to be synchronous. In fact, since  $\mathbf{w}_j(l)$  in (7) are independent among different sections of different received signals and have the identity distribution  $\mathbf{w}_j(l) \sim N(\mathbf{0}, \sigma^2 \mathbf{I}_{NM})$ , it can be shown that the Log Likelihood Function (LLF) of which the constant terms are dropped is given by,

$$\mathcal{L}(\mathbf{P}, \mathring{\mathbf{t}}, \boldsymbol{\alpha}) \triangleq - \sum_{j=1}^{N_r} \sum_{l=1}^L \|\mathbf{y}_j(l) - \mathbf{\Gamma}_{j,l}(\mathbf{P}, \mathring{\mathbf{t}}) \boldsymbol{\alpha}_j\|^2 \quad (9)$$

To get a compact CLF, we define the manifold matrix for the  $j$ th received signal by collecting and stacking all the  $L$  signal sections,

$$\mathbf{\Gamma}_j(\mathbf{P}, \mathring{\mathbf{t}}) \triangleq [\mathbf{\Gamma}_{j,1}^T, \dots, \mathbf{\Gamma}_{j,L}^T]^T \in \mathbb{C}^{LNM \times Q} \quad (10)$$

The CLF (11) can be rewritten as the compact one as follows,

$$\mathcal{L}(\mathbf{P}, \mathring{\mathbf{t}}, \boldsymbol{\alpha}) = - \sum_{j=1}^{N_r} \|\mathbf{y}_j - \mathbf{\Gamma}_j(\mathbf{P}, \mathring{\mathbf{t}}) \boldsymbol{\alpha}_j\|^2 \quad (11)$$

where  $\mathbf{y}_j \triangleq [\mathbf{y}_j^T(1), \dots, \mathbf{y}_j^T(L)]^T$ . Maximizing (11) jointly with respect to all parameters is extremely challenging. Yet, significant computational savings follow from the observation that for any given  $\mathbf{P}, \mathring{\mathbf{t}}$ , the problem of finding the optimal  $\{\boldsymbol{\alpha}_j\}_{j=1}^{N_r}$  becomes a linear least squares (LS) problem [7] whose solution is given by

$$\hat{\boldsymbol{\alpha}}_j = \mathbf{\Gamma}_j^\dagger(\mathbf{P}, \mathring{\mathbf{t}})\mathbf{y}_j \quad (12)$$

where  $\mathbf{\Gamma}_j^\dagger$  is the Moore-Penrose pseudo-inverse of  $\mathbf{\Gamma}_j$  given by  $(\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j)^{-1} \mathbf{\Gamma}_j^H$ . Note that  $(\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j)^{-1}$  always exists since  $\mathbf{\Gamma}_j$  has full column rank. By substituting the estimations  $\{\hat{\boldsymbol{\alpha}}_j\}_{j=1}^{N_r}$  back in (11) and resorting to some straightforward algebraic operations, we obtain the so-called *concentrated* likelihood function,

$$\mathcal{L}_c(\mathbf{P}, \mathring{\mathbf{t}}) = \sum_{j=1}^{N_r} \mathbf{y}_j^H \mathbf{\Gamma}_j (\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j)^{-1} \mathbf{\Gamma}_j^H \mathbf{y}_j \quad (13)$$

The joint ML estimates of  $\mathbf{P}$  and  $\mathring{\mathbf{t}}$  are then obtained as,

$$[\hat{\mathbf{P}}, \hat{\mathring{\mathbf{t}}}] = \arg \max_{\mathbf{P}, \mathring{\mathbf{t}}} \mathcal{L}_c(\mathbf{P}, \mathring{\mathbf{t}}) \quad (14)$$

The reduced-dimension optimization problem (14) still has  $Q(D+1)$  unknown parameters (including  $Q$   $D$ -dimensional real-valued location vectors  $\{\mathbf{p}_q\}_{q=1}^Q$  and  $Q$  transmitted times scalar  $\{\mathring{t}_q\}_{q=1}^Q$ ) remain to estimate. The ML estimation which takes exponential complexity will, therefore, requires a complex  
70 multidimensional search over the parameter space, which is impractical.

To solve the multidimensional optimization problem (14) efficiently, the approach proposed in [1] is to decouple the problem into several  $D$ -dimensional optimization problems, under the assumption that the number of sections  $L \rightarrow \infty$  or  $L$  is sufficiently large in practice. Considering the length of section has low  
75 limit  $T/L \gg \max \tau_j(\mathbf{p}_q)$ , the number of sections  $L$  is proportional to the total observation time  $T$ . The observation time  $T$  is, however, limited in some practical reasons such as the request for real-time processing or the assumption that emitters are stationary during the observation. The decoupling approach proposed in [1] may, therefore, be failure for the assumption is difficult to be

80 satisfied in some practical cases, especially when the emitters are adjacent to one another.

#### 4. Global Maximization of The CLF

Like the previous works within other estimation problems ([8, 5, 9, 10, 11]), we resort to the optimization framework which consisting of the Pincus' theorem  
85 and the powerful Importance Sampling (IS) method to efficiently and accurately estimate the adjacent emitters' positions using finite observation resources such as the observation time, the signal bandwidth and the array aperture. First of all, we introduce the theorem of Pincus [6] as follows.

**Theorem 1.** Let  $F(x_1, \dots, x_n) = F(\mathbf{x})$  be a continuous function on the closure of a bounded domain  $S$  in  $n$ -dimensional Euclidean space of  $\mathbb{R}^n$ . Assume that  $F$  attains a global maximum at exactly on point  $\hat{\mathbf{x}} = [\hat{x}_1, \dots, \hat{x}_n]^T$  of  $S$ . Then for  $i = 1, \dots, n$ ,

$$\hat{x}_i = \lim_{\rho \rightarrow \infty} \frac{\int \dots \int x_i e^{\rho F(\mathbf{x})} dx_1 \dots dx_n}{\int \dots \int e^{\rho F(\mathbf{x})} dx_1 \dots dx_n} \quad (15)$$

Applying this general result to our multidimensional optimization problem (14) with  $\mathbf{x} \triangleq [\mathbf{P}, \mathbf{t}]$  and  $F(\mathbf{x}) \triangleq \mathcal{L}_c(\mathbf{P}, \mathbf{t})$  leads to the following required ML estimation for the emitter locations and the transmitted times respectively,

$$\hat{\mathbf{p}}_q(d) = \int \dots \int \mathbf{p}_q(d) \mathcal{F}(\mathbf{P}, \mathbf{t}) d\mathbf{p}_1(1) \dots dt_Q \quad (16)$$

$$\hat{t}_q^{(0)} = \int \dots \int t_q \mathcal{F}(\mathbf{P}, \mathbf{t}) d\mathbf{p}_1(1) \dots dt_Q \quad (17)$$

$$q = 1, \dots, Q; d = 1, \dots, D$$

where

$$\mathcal{F}(\mathbf{P}, \mathbf{t}) \triangleq \lim_{\rho \rightarrow \infty} \frac{e^{\rho \mathcal{L}_c(\mathbf{P}, \mathbf{t})}}{\int \dots \int e^{\rho \mathcal{L}_c(\mathbf{P}, \mathbf{t})} d\mathbf{p}_1(1) \dots dt_Q} \quad (18)$$

The Pincus' theorem can be interpreted intuitively as follows: as  $\rho$  tends to  
90 infinity,  $\mathcal{F}(\mathbf{P}, \mathbf{t})$  becomes a Dirac-delta function centered at the maximum of

$\mathcal{L}_c(\mathbf{P}, \mathring{\mathbf{t}})$ , which eliminates the integrals in (16) and (17) and chooses its center (i.e., the position of the maximum of  $\mathcal{L}_c(\mathbf{P}, \mathring{\mathbf{t}})$ ) as the ML solution directly.

Nevertheless,  $\rho$  taking a sufficiently large value in practice, the multidimensional grid search we attempt to avoid is transformed into several multidimensional integrations in (16) and (17) bearing the same computational load. By closely analyzing the properties of function (18), it, fortunately, turns out that  $\mathcal{F}(\mathbf{P}, \mathring{\mathbf{t}})$  can be seen as a Probability Density Function (pdf) since it is nonnegative and integrates to one. As a pdf,  $\mathcal{F}(\mathbf{P}, \mathring{\mathbf{t}})$  indicates the probability of the candidate parameter values being true. Furthermore, the estimators in (16) and (17) can be regarded as statistical expectations, i.e., we have,

$$\hat{\mathbf{p}}_q(d) \triangleq \mathbb{E}_{\mathbf{P}, \mathring{\mathbf{t}}} \{\mathbf{p}_q(d)\} \quad \hat{t}_q \triangleq \mathbb{E}_{\mathbf{P}, \mathring{\mathbf{t}}} \{\mathring{t}_q\} \quad (19)$$

Consequently, the multidimensional integrations are transformed into the statistical expectations, which can be approximated by the sample mean. We will then explain the method to generate  $R$  realizations,  $\{\mathbf{P}^{(r)}\}_{r=1}^R$  and  $\{\mathring{\mathbf{t}}^{(r)}\}_{r=1}^R$ , using the joint pdf  $\mathcal{F}(\mathbf{P}, \mathring{\mathbf{t}})$  and approximate the expectations in (19) using the following sample means,

$$\hat{\mathbf{p}}_q(d) = \frac{1}{R} \sum_{r=1}^R \mathbf{p}_q^{(r)}(d) \quad \hat{t}_q = \frac{1}{R} \sum_{r=1}^R \mathring{t}_q^{(r)} \quad (20)$$

Clearly, as the number of realizations  $R$  increases, the two sample mean in (20) approach the statistical expectations in (19). Unfortunately, the extremely non-linear property of the pdf  $\mathcal{F}(\mathbf{P}, \mathring{\mathbf{t}})$  makes it impractical to generate  $\{\mathbf{P}^{(r)}\}_{r=1}^R$  and  $\{\mathring{\mathbf{t}}^{(r)}\}_{r=1}^R$  using itself. To sidestep this problem, we can resort to the Importance Sampling (IS) method [8, 5] and rewrite (16) and (17) in the following equivalent forms,

$$\hat{\mathbf{p}}_q(d) = \int \cdots \int \mathbf{p}_q(d) \frac{\mathcal{F}(\mathbf{P}, \mathring{\mathbf{t}})}{\mathcal{G}(\mathbf{P}, \mathring{\mathbf{t}})} \mathcal{G}(\mathbf{P}, \mathring{\mathbf{t}}) d\mathbf{p}_1(1) \cdots d\mathring{t}_Q \quad (21)$$

$$\hat{t}_q = \int \cdots \int \mathring{t}_q \frac{\mathcal{F}(\mathbf{P}, \mathring{\mathbf{t}})}{\mathcal{G}(\mathbf{P}, \mathring{\mathbf{t}})} \mathcal{G}(\mathbf{P}, \mathring{\mathbf{t}}) d\mathbf{p}_1(1) \cdots d\mathring{t}_Q \quad (22)$$



where  $\mathcal{G}(\mathbf{P}, \mathbf{t})$  is another pdf called *importance function* (IF). By introducing the IF, the multidimensional integrations in (21) and (22) are interpreted as statistical expectations of two transformed random variates (RVs), i.e.,

$$\hat{\mathbf{p}}_q(d) \triangleq \mathbb{E}_{\mathbf{P}, \mathbf{t}}\{\eta(\mathbf{P}, \mathbf{t}) \mathbf{p}_q(d)\} \quad \hat{t}_q \triangleq \mathbb{E}_{\mathbf{P}, \mathbf{t}}\{\eta(\mathbf{P}, \mathbf{t}) t_q\} \quad (23)$$

where  $\eta(\mathbf{P}, \mathbf{t})$  is called *importance weight* (IW) corresponding to the realizations of  $\mathbf{P}, \mathbf{t}$  and defined as,

$$\eta(\mathbf{P}, \mathbf{t}) \triangleq \frac{\mathcal{F}(\mathbf{P}, \mathbf{t})}{\mathcal{G}(\mathbf{P}, \mathbf{t})} \quad (24)$$

If  $\mathcal{G}(\mathbf{P}, \mathbf{t})$  is carefully designed, the expectations in (23) can be computed at any desired degree of accuracy (by increasing  $R$ ) using the corresponding sample mean estimates. In fact, the choice of  $\mathcal{G}(\mathbf{P}, \mathbf{t})$  is a tradeoff between the ease and efficiency of sampling (i.e., the generation of the required realizations). The appropriate choice of the IF will be discussed in the following section.

## 5. Appropriate Choice of Importance Function

To generate the realizations efficiently, the appropriate  $\mathcal{G}(\mathbf{P}, \mathbf{t})$  must be designed as close as possible to  $\mathcal{F}(\mathbf{P}, \mathbf{t})$  while it should also take the ease of generating the required realizations  $\{\mathbf{P}^{(r)}\}_{r=1}^R$  and  $\{\mathbf{t}^{(r)}\}_{r=1}^R$  into account. In order to find the appropriate  $\mathcal{G}(\mathbf{P}, \mathbf{t})$ , we start with constraining IF's form separable in terms of the  $Q$  position-transmitted time pairs  $\{(\mathbf{p}_q, t_q)\}_{q=1}^Q$  for it is relative easy to generate realizations from pdf with the factorized form as follows,

$$\mathcal{G}(\mathbf{P}, \mathbf{t}) \triangleq \prod_{q=1}^Q g_q(\mathbf{p}_q, t_q) \quad (25)$$

This allow us to interpret the original pdf of  $Q(D+1)$ -dimensional random vector as the joint pdf of  $Q$  independent  $(D+1)$ -dimensional random vectors. Hence, we can independently generate  $Q$  groups of realizations for  $(D+1)$ -dimensional random vectors  $\{(\mathbf{p}_q^{(r)}, t_q^{(r)})\}_{r,q=1}^{R,Q}$ , instead of generating realizations for  $Q(D+1)$ -dimensional random vector  $(\mathbf{P}, \mathbf{t})$  directly using  $\mathcal{G}(\mathbf{P}, \mathbf{t})$ .

According to the definition of  $\mathcal{F}(\mathbf{P}, \hat{\mathbf{t}})$  (18), the separable form (25) requests that the CLF in (18) has summation form with respect to  $q$ . Note that the product  $\mathbf{y}_j^H \mathbf{\Gamma}_j$  in CLF is a  $Q$ -dimensional vector. Diagonalizing the inverse matrix  $(\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j)^{-1}$  is, therefore, the key step to the separable form (25). According to the definition of  $\mathbf{\Gamma}_j$  in (10),  $\mathbf{\Gamma}_{j,l}$  in (8),  $\mathbf{\Gamma}_{j,n,l}$  in (6) and  $\gamma_{j,q,n,l}$  in (3), the  $(u, v)$ th element of the normalized matrix product  $\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j$  is given by,

$$[\frac{1}{MNL} \mathbf{\Gamma}_j^H \mathbf{\Gamma}_j]_{u,v} \quad (26)$$

$$= \frac{1}{MNL} \sum_{l=1}^L \sum_{n=1}^N \gamma_{j,q,n,l}^H \gamma_{j,q,n,l} \quad (27)$$

$$= \underbrace{\frac{1}{M} \mathbf{a}_j^H(\mathbf{p}_u) \mathbf{a}_j(\mathbf{p}_u)}_{r_{doa,j}^{(u,v)}} \underbrace{\frac{1}{N} \sum_{n=1}^N e^{j2\pi f_n(\tau_j(\mathbf{p}_u) + \hat{t}_u - \tau_j(\mathbf{p}_v) - \hat{t}_v)}}_{r_{toa,j}^{(u,v)}} \underbrace{\frac{1}{L} \sum_{l=1}^L s_u^*(n, l) s_v(n, l)}_{r_s^{(u,v)}} \quad (28)$$

where  $r_{doa,j}^{(u,v)}, r_{toa,j}^{(u,v)}, r_s^{(u,v)}$  denote the  $(u, v)$ th normalized correlation coefficients in aspects of Direction Of Arrival (DOA), Time Of Arrival (TOA) and the transmitted signal, respectively. The matrix  $\frac{1}{MNL} \mathbf{\Gamma}_j^H \mathbf{\Gamma}_j$  is, therefore, a normalized correlation matrix among all emitters. To avoid additional computational load, we directly assume, instead of strictly diagonalizing,  $\frac{1}{MNL} \mathbf{\Gamma}_j^H \mathbf{\Gamma}_j$  as identity matrix, at the cost of losing the correlation information. As a consequence, the diagonal matrix, for  $j = 1, \dots, N_r$ , is given by,

$$\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j = MNL \cdot \mathbf{I} \quad (29)$$

where  $\mathbf{I}$  is the identity matrix, the coefficient  $MNL$  is the diagonal element of  $\mathbf{\Gamma}_j^H \mathbf{\Gamma}_j$ .

substituting (29) into (13) and ignoring the constant coefficient, we get its separated form as follows,

$$\tilde{\mathcal{L}}_c(\mathbf{P}, \hat{\mathbf{t}}) = \sum_{j=1}^{N_r} \|\mathbf{\Gamma}_j^H \mathbf{y}_j\|^2 \quad (30)$$

$$= \sum_{q=1}^Q \tilde{\ell}_c(\mathbf{p}_q, \hat{t}_q) \quad (31)$$

where

$$\tilde{\ell}_c(\mathbf{p}_q, \mathring{t}_q) \triangleq \sum_{j,n,l=1}^{N_r, N, L} |\gamma_{j,q,n,l}^H \mathbf{y}_j(n, l)|^2 \quad (32)$$

Replacing  $\mathcal{L}_c(\mathbf{P}, \mathring{\mathbf{t}})$  and  $\rho$  in (18) with  $\tilde{\mathcal{L}}_c(\mathbf{P}, \mathring{\mathbf{t}})$  and an appropriate  $\rho_2$  respectively, we get the separated form of  $\mathcal{G}(\mathbf{P}, \mathring{\mathbf{t}})$  of which the  $q$ th component is given by,

$$g_q(\mathbf{p}_q, \mathring{t}_q) \triangleq \frac{e^{\rho_2 \tilde{\ell}_c(\mathbf{p}_q, \mathring{t}_q)}}{\int \int e^{\rho_2 \tilde{\ell}_c(\mathbf{p}_q, \mathring{t}_q)} d\mathbf{p}_q d\mathring{t}_q} \quad (33)$$

The generation of  $q$ th group of realizations  $\{(\mathbf{p}_q^{(r)}, \mathring{t}_q^{(r)})\}_{r=1}^R$  follows, furthermore, the well-known general result from probability theory, which is that the joint pdf,  $g_q(\mathbf{p}_q, \mathring{t}_q)$ , can be factorized as the product of marginal and conditional pdfs,

$$g_q(\mathbf{p}_q, \mathring{t}_q) = g_q^{\text{mar}}(\mathbf{p}_q) g_q^{\text{con}}(\mathring{t}_q | \mathbf{p}_q) \quad (34)$$

The usual method to obtain the marginal pdf  $g_q^{\text{mar}}(\mathbf{p}_q)$  is integrating the joint pdf  $g_q(\mathbf{p}_q, \mathring{t}_q)$  with respect to the transmitted time  $\mathring{t}_q$ . The analytical integration is, however, difficult since  $g_q(\mathbf{p}_q, \mathring{t}_q)$  is extremely non-linear about  $\mathring{t}_q$ . The discrete approximation also involves  $D+1$ -dimensional grid search. We have another alternative method to construct  $g_q^{\text{mar}}(\mathbf{p}_q)$  directly and efficiently,

$$g_q^{\text{mar}}(\mathbf{p}_q) \triangleq \frac{e^{\rho_1 \ell(\mathbf{p}_q)}}{\int e^{\rho_1 \ell(\mathbf{p}_q)} d\mathbf{p}_q} \quad (35)$$

where  $\ell(\mathbf{p}_q)$  is the *single* ( $Q=1$ ) Likelihood Function (SLF) of the position  $\mathbf{p}_q$ , which is derived from (32) by eliminating the transmitted time  $\mathring{t}_q$ . We will, then, describe the derivation in details.

Our derivation starts with a vector definition as follows,

$$\gamma_{q,j,l} \triangleq \Phi_{q,j,l}(\mathbf{p}_q) \mathbf{b}(\mathring{t}_q) \quad (36)$$

where

$$\mathbf{\Phi}_{q,j,l}(\mathbf{p}_q) \triangleq \mathbf{\Lambda}_j(\mathbf{p}_q, s_q(l)) \otimes \mathbf{a}_j(\mathbf{p}_q) \quad (37)$$

$$\mathbf{b}(\mathring{t}_q) \triangleq [e^{-j2\pi f_1 \mathring{t}_q}, \dots, e^{-j2\pi f_N \mathring{t}_q}]^T \quad (38)$$

$$\mathbf{\Lambda}_j(\mathbf{p}_q, s_q(l)) \triangleq \begin{bmatrix} s_q(1, l) e^{-j2\pi f_1 \tau_j(\mathbf{p}_q)} & \dots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \dots & s_q(N, l) e^{-j2\pi f_N \tau_j(\mathbf{p}_q)} \end{bmatrix} \quad (39)$$

substituting (36) and (37) into (32), we get

$$\tilde{\ell}_c(\mathbf{p}_q, \mathring{t}_q) = \sum_{j=1}^{N_r} \sum_{l=1}^L \gamma_{q,j,l}^H \mathbf{y}_j(l) \mathbf{y}_j^H(l) \gamma_{q,j,l} \quad (40)$$

$$= \mathbf{b}^H(\mathring{t}_q) \underbrace{\left( \sum_{j=1}^{N_r} \sum_{l=1}^L \mathbf{\Phi}_{q,j,l}^H(\mathbf{p}_q) \mathbf{y}_j(l) \mathbf{y}_j^H(l) \mathbf{\Phi}_{q,j,l}(\mathbf{p}_q) \right)}_{\triangleq \mathbf{\Pi}_q(\mathbf{p}_q)} \mathbf{b}(\mathring{t}_q) \quad (41)$$

The right hand of the second equation is a non-normalized Rayleigh Quotient with respect to  $\mathring{t}_q$ . Consequently, maximizing  $\tilde{\ell}_c(\mathbf{p}_q, \mathring{t}_q)$  with respect to  $\mathring{t}_q$ , we get,

$$\ell_q(\mathbf{p}_q) \triangleq \lambda_{\max}(\mathbf{\Pi}_q(\mathbf{p}_q)) \quad (42)$$

where the notation  $\lambda_{\max}(\mathbf{X})$  denotes the maximum eigenvalue of the matrix  $\mathbf{X}$ . Substituting (42) into (35), we get the marginal pdf of the  $q$ th emitter's position,

$$g_q^{\text{mar}}(\mathbf{p}_q) = \frac{e^{\rho_1 \lambda_{\max}(\mathbf{\Pi}_q(\mathbf{p}_q))}}{\int e^{\rho_1 \lambda_{\max}(\mathbf{\Pi}_q(\mathbf{p}_q))} d\mathbf{p}_q} \quad (43)$$

Given position realization  $\mathbf{p}_q^{(r)}$  generated by  $g_q^{\text{mar}}(\mathbf{p}_q)$ , the conditional transmitted time pdf is defined as,

$$g_q^{\text{con}}(\mathring{t}_q | \mathbf{p}_q^{(r)}) = \frac{g_q(\mathbf{p}_q^{(r)}, \mathring{t}_q)}{\int_{\mathring{t}} g_q(\mathbf{p}_q^{(r)}, \mathring{t}_q) d\mathring{t}} \quad (44)$$

Substituting (33) into (44), we have,

$$g_q^{\text{con}}(\mathring{t}_q | \mathbf{p}_q^{(r)}) = \frac{e^{\rho_2 \tilde{\ell}_c(\mathbf{p}_q^{(r)}, \mathring{t}_q)}}{\int_{\mathring{t}} e^{\rho_2 \tilde{\ell}_c(\mathbf{p}_q^{(r)}, \mathring{t}_q)} d\mathring{t}} \quad (45)$$

Note that we use  $\int_{\tilde{t}} e^{\rho_2 \tilde{\ell}_c(\mathbf{p}_q^{(r)}, \tilde{t}_q)} d\tilde{t}$ , instead of  $g_q^{\text{mar}}(\mathbf{p}_q^{(r)})$ , to normalize the conditional pdf (44) correctly.

The process of generating the required realizations  $\{(\mathbf{p}_q^{(r)}, \tilde{t}_q^{(r)})\}_{r,q=1}^{R,Q}$  are as follows: for  $q = 1, \dots, Q$ th group of realizations, generate  $\{\mathbf{p}_q^{(r)}\}_{r=1}^R$  using the marginal pdf  $g_q^{\text{mar}}(\mathbf{p}_q)$  (43) and then use the conditional pdf  $g_q^{\text{con}}(\tilde{t}_q | \mathbf{p}_q^{(r)})$  (45) to generate  $\{\tilde{t}_q^{(r)}\}_{r=1}^R$ .

## 6. Generations Of The Required Realizations

We recall the following lemma [12] that will be used to generate the required realizations:

**Lemma 1.** Let  $X \in \mathcal{X}$  be any RV with pdf  $f_X(x)$  and the Cumulative Distribution Function (CDF)  $F_X(x)$  and denote the inverse CDF as  $F_X^{-1}(\cdot) : [0, 1] \rightarrow \mathcal{X}, u \rightarrow x$  s.t.  $F_X(x) = u$ . Then, for any uniform RV,  $U \in [0, 1]$ , the RV  $\tilde{X} = F_X^{-1}(U)$  is distributed according to  $f_X(\cdot)$ .

According to the generating process described before, firstly, we should use the marginal pdf  $g_q^{\text{mar}}(\mathbf{p}_q)$  (43) to generate  $\{\mathbf{p}_q^{(r)}\}_{r=1}^R$ .  $g_q^{\text{mar}}(\mathbf{p}_q)$  is, however, unsuitable to be used along with the result of Lemma 1 since  $\mathbf{p}_q$  is a  $D$ -dimensional vector (defining  $\mathbf{p}_q = (x_q, y_q)$  for  $D = 2$ ) instead of a scalar. To acquire the  $D = 2$  scalar pdfs, consequently,  $g_q^{\text{mar}}(\mathbf{p}_q)$  should be factorized as follows,

$$g_q^{\text{mar}}(x_q, y_q) \triangleq g_x^{\text{mar}}(x_q) g_{y|x}^{\text{con}}(y_q | x_q) \quad (46)$$

where

$$g_x^{\text{mar}}(x_q) \triangleq \int g_q^{\text{mar}}(x_q, y_q) dy_q \quad (47)$$

$$g_{y|x}^{\text{con}}(y_q | x_q) \triangleq \frac{g_q^{\text{mar}}(x_q, y_q)}{g_x^{\text{mar}}(x_q)} \quad (48)$$

In conclusion, the approach of generating the  $q$ th group of realizations  $\{(\mathbf{p}_q^{(r)}, \tilde{t}_q^{(r)})\}_{r=1}^R$  can be stated as follows: generate  $R$  realizations  $\{u_x^{(r)}\}_{r=1}^R \sim U[0, 1]$  and obtain  $x_q^{(r)} = G_{x_q}^{-1}(u_x^{(r)})$  where  $G_{x_q}(\cdot)$  is the CDF associated to  $g_x^{\text{mar}}(x_q)$  in (47); For  $r = 1, \dots, R$ , then, generate two realizations  $u_y^{(r)}, u_{\tilde{t}}^{(r)} \sim U[0, 1]$  and obtain

$y_q^{(r)} = G_{y_q|x_q}^{-1}(u_y^{(r)})$  and  $\hat{t}_q^{(r)} = G_{\hat{t}_q|\mathbf{p}_q}^{(-1)}(u_{\hat{t}}^{(r)})$  where  $G_{y_q|x_q}(\cdot)$  and  $G_{\hat{t}_q|\mathbf{p}_q}(\cdot)$  are the CDF associated to  $g_{y|x}^{\text{con}}(y_q|x_q)$  in (47) and  $g_q^{\text{con}}(\hat{t}_q|\mathbf{p}_q^{(r)})$  in (44) respectively;

## 7. Implementation Details

In this section, we give all the necessary details for an efficient implemen-  
 130 tation of our proposed IS-based ML DPD algorithm. To narrow the range of  
 sampling, we would like to generate the realizations around the true paramet-  
 ers. The area of interest, where emitters may be put in, is confined within  
 $x \in [x_{\min}, x_{\max}]$  and  $y \in [y_{\min}, y_{\max}]$  and the range of transmitted time is  
 $\hat{t} \in [0, \hat{t}_{\max}]$ , where  $\hat{t}_{\max}$  can be freely chosen as high as desired. The process  
 135 of generating the required  $Q$  groups of realizations amounts to performing the  
 following steps for every  $q = 1, \dots, Q$ :

STEP 1: Evaluate the  $q$ th SLF (42) at multiple two-dimensional grid points  
 $\{(x_q^{(i)}, y_q^{(j)})\}_{i,j=1}^{I,J}$  with relatively large step  $\Delta x$  and  $\Delta y$ . Then, by approximating  
 integrals with discrete sums, we evaluate the  $q$ th marginal position pdf with  
 respect to  $\mathbf{p}_q = (x_q, y_q)$  in (35),  $g_q^{\text{mar}}(x_q, y_q)$ , at every grid point in the whole  
 area of interest  $[x_{\min}, x_{\max}] \times [y_{\min}, y_{\max}]$ , as follows,

$$g_q^{\text{mar}}(x_q^{(i)}, y_q^{(j)}) \simeq \frac{e^{\rho_1 \ell_q(x_q^{(i)}, y_q^{(j)})}}{\sum_{i=1}^I \sum_{j=1}^J e^{\rho_1 \ell_q(x_q^{(i)}, y_q^{(j)})} \Delta x \Delta y} \quad (49)$$

STEP 2: Initialize the estimation of the  $q$ th emitter position as the grid  
 point corresponding to the maximum peak of  $g_q^{\text{mar}}(x_q^{(i)}, y_q^{(j)})$ ,

$$[\hat{x}_q, \hat{y}_q] = \arg \max_{x^{(i)}, y^{(j)}} g_q^{\text{mar}}(x^{(i)}, y^{(j)}) \quad (50)$$

STEP 3: Substitute the initial position estimation  $\hat{\mathbf{p}}_q = (\hat{x}_q, \hat{y}_q)$  into the  
 conditional transmitted time pdf (44),  $g_q^{\text{con}}(\hat{t}_q|\hat{\mathbf{p}}_q)$ , and then evaluate it at mul-  
 tiple grid points  $\{\hat{t}_q^{(s)}\}_{s=1}^S$  with relatively large step  $\Delta \hat{t}$  in the interval  $[0, \hat{t}_{\max}]$ ,  
 as follows,

$$g_q^{\text{con}}(\hat{t}_q^{(s)}|\hat{\mathbf{p}}_q) = \frac{g_q(\hat{\mathbf{p}}_q, \hat{t}_q^{(s)})}{\sum_{s=1}^S g_q(\hat{\mathbf{p}}_q, \hat{t}_q^{(s)}) \Delta \hat{t}} \quad (51)$$

STEP 4: Initialize the estimation of the  $q$ th transmitted time  $\hat{t}_q$  as the grid point corresponding to the maximum peak of the conditional transmitted time pdf  $g_q^{\text{con}}(\hat{t}_q^{(s)} | \hat{\mathbf{p}}_q)$ .

The distance between the initial estimation and the ML solution in (14), in the multidimensional parameter space, however, is non-negligible since the decouple processing in (30) lose the correlation information among emitters. To force  $\{(\mathbf{p}_q^{(r)}, \hat{t}_q^{(r)})\}_{r=1}^R$  to be generated in the vicinity of the ML solutions  $(\hat{\mathbf{p}}_q, \hat{t}_q)$ , we fix the local parameter space to include the ML solutions as follows,

$$D_{\hat{x}_q} \triangleq [\hat{x}_q - \Delta_x, \hat{x}_q + \Delta_x] \quad (52)$$

$$D_{\hat{y}_q} \triangleq [\hat{y}_q - \Delta_y, \hat{y}_q + \Delta_y] \quad (53)$$

$$D_{\hat{t}_q} \triangleq [\hat{t}_q - \Delta_t, \hat{t}_q + \Delta_t] \quad (54)$$

STEP 5: Evaluate the  $q$ th marginal position pdf  $g_q^{\text{mar}}(x_q, y_q)$  at new grid points with relatively small step  $\delta x < \Delta x$  and  $\delta y < \Delta y$ , in the local area  $D_{\hat{x}_q} \times D_{\hat{y}_q}$ , as follows,

$$g_q^{\text{mar}}(x_q^{(i')}, y_q^{(j')}) \simeq \frac{e^{\rho_1 \ell_q(x_q^{(i')}, y_q^{(j')})}}{\sum_{i'=1}^{I'} \sum_{j'=1}^{J'} e^{\rho_1 \ell_q(x_q^{(i')}, y_q^{(j')})} \delta x \delta y} \quad (55)$$

STEP 6: Compute the marginal coordinate  $x_q$  pdf by approximating integrals with discrete sums in (47),

$$g_x^{\text{mar}}(x_q^{(i')}) \simeq \sum_{j'=1}^{J'} g_q^{\text{mar}}(x_q^{(i')}, y_q^{(j')}) \delta y, \quad \forall x_q^{(i')} \in D_{\hat{x}_q} \quad (56)$$

STEP 7: Compute the coordinate  $x_q$ 's CDF as follows:

$$G_{x_q}(x_q^{(r)}) = \sum_{i' \leq r} g_x^{\text{mar}}(x_q^{(i')}) \delta x, \quad \forall x_q^{(r)} \in D_{\hat{x}_q} \quad (57)$$

140 The following steps are designed to generate the  $r$ th realization  $(x_q^{(r)}, y_q^{(r)}, \hat{t}_q^{(r)})$  corresponding to the  $q$ th emitter, i.e., for  $r = 1, \dots, R$ :

STEP 8: Generate the  $r$ th RV  $u_{x_q}^{(r)} \sim U[0, 1]$  and invert  $G_{x_q}(\cdot)$  via linear interpolation in order to obtain the  $r$ th coordinate  $x_q$  realization  $x_q^{(r)} = G_{x_q}^{-1}(u_{x_q}^{(r)})$ .

STEP 9: Given the  $r$ th realization  $x_q^{(r)}$ , obtain the conditional coordinate  $y_q$  pdf from the marginal position pdf (already evaluated in "STEP 5") as follows:

$$g_{y|x}^{\text{con}}(y_q^{(j')}|x_q^{(r)}) \simeq \frac{g_q^{\text{mar}}(x_q^{(r)}, y_q^{(j')})}{g_x^{\text{mar}}(x_q^{(r)})} \quad (58)$$

STEP 10: Evaluate the coordinate  $y_q$ 's CDF,  $G_{y_q}(y_q^{(r)})$  similarly to  $G_{x_q}(x_q^{(r)})$  in (57) and generate the  $r$ th coordinate  $y_q$  realization,  $y_q^{(r)} = G_{y_q}^{-1}(u_{y_q}^{(r)})$ , using linear interpolation as well.

STEP 11: Given the  $r$ th position realization  $\mathbf{p}_q^{(r)} = (x_q^{(r)}, y_q^{(r)})$ , evaluate the  $q$ th conditional transmitted time pdf (44) at new grid points with relatively small step  $\delta \hat{t} < \Delta \hat{t}$ , in the local interval  $D_{\hat{t}_q}$ , as follows,

$$g_q^{\text{con}}(\hat{t}_q^{(s')}|\mathbf{p}_q^{(r)}) = \frac{g_q(\mathbf{p}_q^{(r)}, \hat{t}_q^{(s')})}{\sum_{s'=1}^{S'} g_q(\mathbf{p}_q^{(r)}, \hat{t}_q^{(s')}) \delta \hat{t}} \quad (59)$$

STEP 12: Evaluate the transmitted time  $\hat{t}_q$ 's CDF,  $G_{\hat{t}_q|\mathbf{p}_q}(\cdot)$  similarly to  $G_{x_q}(x_q^{(r)})$  in (57) and generate the  $r$ th transmitted time realization,  $\hat{t}_q^{(r)} = G_{\hat{t}_q|\mathbf{p}_q}^{-1}(u_{\hat{t}_q}^{(r)})$ , using linear interpolation as well.

All the required realizations  $\{(\mathbf{p}_q^{(r)}, \hat{t}_q^{(r)})\}_{r,q=1}^{R,Q}$  can be generated according to the steps above. The  $q$ th expectations in (23), which is also the ML solutions, can be approximated based on these realizations as follows,

$$\hat{\mathbf{p}}_q = \frac{1}{R} \sum_{r=1}^R \eta(\mathbf{P}^{(r)}, \hat{\mathbf{t}}^{(r)}) \mathbf{p}_q^{(r)}, \quad \hat{t}_q = \frac{1}{R} \sum_{r=1}^R \eta(\mathbf{P}^{(r)}, \hat{\mathbf{t}}^{(r)}) \hat{t}_q^{(r)} \quad (60)$$

According to equations (24),(18),(25) and (33), the IW can be written as,

$$\eta(\mathbf{P}^{(r)}, \hat{\mathbf{t}}^{(r)}) \triangleq A e^{\phi(\mathbf{P}^{(r)}, \hat{\mathbf{t}}^{(r)})} \quad (61)$$

$$\phi(\mathbf{P}^{(r)}, \hat{\mathbf{t}}^{(r)}) \triangleq \rho_0 \mathcal{L}_c(\mathbf{P}^{(r)}, \hat{\mathbf{t}}^{(r)}) - \rho_2 \sum_{q=1}^Q \tilde{\ell}_c(\mathbf{p}_q^{(r)}, \hat{t}_q^{(r)}) \quad (62)$$

where  $\rho_0$ , instructed by the Pincus' theorem (15), is selected as a sufficiently high value in practice. The scale factor  $A$ , which has absorbed the coefficient  $R^{-1}$  in (60), involves the multidimensional integration. To greatly reduce the computational load and the memory with no changes in the final results, we can



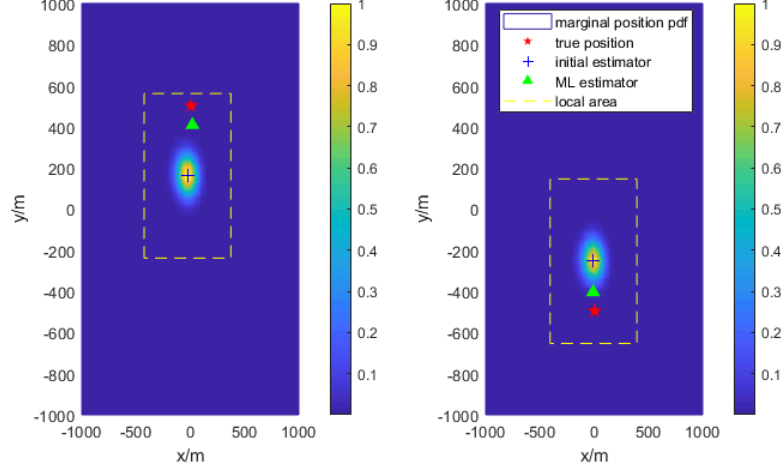


Figure 1: the marginal position pdf, with too large  $\rho_1 = 500$ , of two adjacent emitters at  $\mathbf{p}_1 = (0, 500)\text{m}$  and  $\mathbf{p}_2 = (0, -500)\text{m}$ , the main simulation conditions are SNR=15dB, the number of sections  $L = 10$ , the number of elements  $M = 3$ , the bandwidth  $B_w = 20\text{kHz}$  and the range  $\Delta_x = \Delta_y = 400\text{m}$ .

use the following *normalized* IW,

$$\bar{\eta}(\mathbf{P}^{(r)}, \mathbf{t}^{(r)}) = \frac{e^{\phi(\mathbf{P}^{(r)}, \mathbf{t}^{(r)}) - \max_r \phi(\mathbf{P}^{(r)}, \mathbf{t}^{(r)})}}{\sum_{r=1}^R e^{\phi(\mathbf{P}^{(r)}, \mathbf{t}^{(r)}) - \max_r \phi(\mathbf{P}^{(r)}, \mathbf{t}^{(r)})}} \quad (63)$$

The final IS based ML estimator is, therefore, given by

$$\hat{\mathbf{p}}_q = \sum_{r=1}^R \bar{\eta}(\mathbf{P}^{(r)}, \mathbf{t}^{(r)}) \mathbf{p}_q^{(r)}, \quad \hat{t}_q = \sum_{r=1}^R \bar{\eta}(\mathbf{P}^{(r)}, \mathbf{t}^{(r)}) t_q^{(r)} \quad (64)$$

150 Another detail about the super-parameter  $\rho_1$  and  $\rho_2$  remained to be stated.

The distribution of probability mass in the local area  $D_{\hat{x}_q} \times D_{\hat{y}_q}$  (or  $D_{\hat{t}_q}$ ) is governed by the super-parameter  $\rho_1$  in the marginal position pdf  $g_q^{\text{mar}}(x_q, y_q)$  (or  $\rho_2$  in the conditional transmitted time pdf  $g_q^{\text{con}}(t_q | \mathbf{p}_q^{(r)})$ ). Similarly to (18), the probability mass of  $g_q^{\text{mar}}(x_q, y_q)$  (or  $g_q^{\text{con}}(t_q | \mathbf{p}_q^{(r)})$ ) would concentrate at the  
155 location of the maximum peak,  $(\hat{x}_q, \hat{y}_q)$  (or  $\hat{t}_q$ ), as  $\rho_1$  (or  $\rho_2$ ) increases. The too large  $\rho_1$  (or  $\rho_2$ ) would result in that the points, included in local area but far away from the maximum peak point, are hardly to be sampled (as seen in Fig.1). The range parameter  $\Delta_x, \Delta_y$  (or  $\Delta_{\hat{t}}$ ) should, therefore, match with  $\rho_1$  (or  $\rho_2$ ),

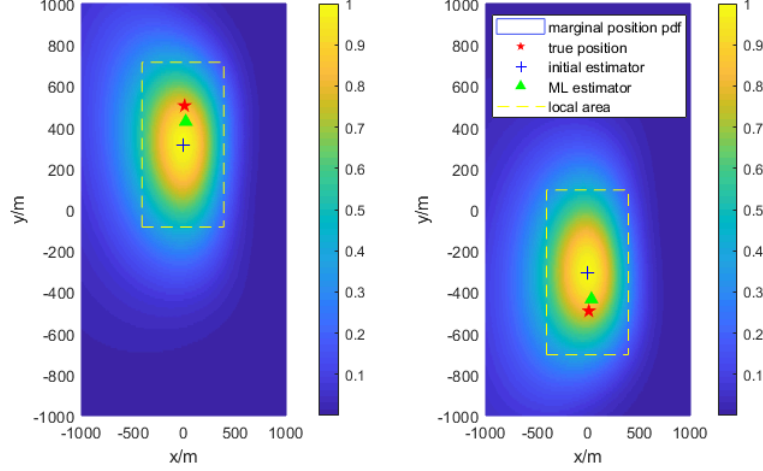


Figure 2: the marginal position pdf with appropriate  $\rho_1 = 20$ . Other conditions are the same as the above.

in order to generate the required realizations efficiently (as seen in Fig.2).

## 160 8. Complexity analysis

The main computation of IS-based ML DPD algorithm concentrates in the generation of the required realizations and the computation of the IWs. The generation process involves evaluating the marginal position pdf  $g_q^{\text{mar}}(x_q^{(i)}, y_q^{(j)})$  and the conditional transmitted time pdf  $g_q^{\text{con}}(\hat{t}_q^{(s)} | \hat{\mathbf{p}}_q)$  at grids with different  
165 ranges and steps. The computational complexity, therefore, relies on the complexity of pdf and the number of grid points. According to the definition of  $g_q^{\text{mar}}(x_q^{(i)}, y_q^{(j)})$  (43), its complexity is  $\mathcal{O}(N^3)$  for the operation  $\lambda_{\text{max}}$  involves the eigen-decomposition. The marginal position pdf  $g_q^{\text{mar}}(x_q^{(i)}, y_q^{(j)})$  is evaluated in (49), (55) and (58), at  $IJ$ ,  $I'J'$  and  $RJ'$  grid points respectively. The  
170 complexity of the generation of  $Q$  groups of  $\{(\mathbf{p}_q^{(r)})\}_{r=1}^R$  is, therefore, about  $\mathcal{O}(Q(IJ + I'J' + RJ')N^3)$ . Analyzing the complexity of generation of  $Q$  groups of  $\{(\hat{t}_q^{(r)})\}_{r=1}^R$  similarly, the complexity is about  $\mathcal{O}(Q(S + RS')LNM)$ . According to the definition of IWs in (63) and (61), the computation of IWs involves evaluating the CLF (13) and (30) at the  $Q$  group of realizations  $\{(\mathbf{p}_q^{(r)}, \hat{t}_q^{(r)})\}_{q,r=1}^{Q,R}$ ,

175 of which the complexity is about  $\mathcal{O}(RQ(LNMN_r + Q^2))$ . The total complexity of IS-based ML DPD is about  $\mathcal{O}(Q(G_p N^3 + (G_t + RN_r)LNM + RQ^2))$  where  $G_p = IJ + I'J' + RJ'$  and  $G_t = S + RS'$ .

## 9. Numerical Simulation and Discussion

In this section, we design several numerical simulations to verify and evaluate  
 180 the performance of the proposed IS based ML estimator, denoted by IS-ML-DPD, and compare it with the decoupled DPD estimator proposed in [1] and the MVDR based DPD estimator proposed in [3], denoted by single-ML-DPD and MVDR-DPD respectively. To give an explicit comparison with the ML estimator (14), we also give its performance curve that are obtained by an  
 185 iterative local optimization method using the true emitters' positions as the initial point, denoted by optimal-ML-DPD.

The decoupled DPD estimator in [1], single-ML-DPD, decouples the  $Q(D + 1)$ -dimensional optimization problem into  $Q$  independent  $D + 1$ -dimensional optimization problems, through approximating the emitters' correlation matrix  
 190 using a diagonal matrix (similar to the equation (29)). The off-diagonal elements of the normalized correlation matrix, i.e., the product of correlation coefficients in aspects of DOA, TOA and transmitted signal (shown in (26)), tend to zero as the number of sections  $L \rightarrow \infty$  since the transmitted signals among emitters are uncorrelated. The approximation in [1], consequently, may become inaccurate  
 195 when the number of sections is limited in practice, which badly worsen the performance of single-ML-DPD estimator.

According to the multiplier effect, The off-diagonal elements (26) can also tend to zero as the DOA or TOA correlation coefficient  $r_{doa,j}^{(u,v)}, r_{toa,j}^{(u,v)} \rightarrow 0$ , even though the number of sections is limited in practice. It means that some estimators with high resolution, like the MVDR filter, may acquire better performance.  
 200 The MVDR based estimator in [3], MVDR-DPD, constructs the MVDR filter on the location domain directly. It is well known that the MVDR principle requires the Distortionless Response vector (3) keeps the same among data samples (sig-

nal sections), which means that the prior waveform information  $s_q(n, l)$  has to  
 205 be eliminated from the characteristic vector (3). The MVDR-DPD estimator  
 can, therefore, only use the DOA and TOA information to distinguish emitters.

We use the Root Mean Square Error (RMSE) of the first emitter's location  
 estimation to measure the performance of each algorithm,

$$\text{RMSE} = \sqrt{\frac{\sum_{s=1}^S \|\mathbf{p}_1^{(s)} - \mathbf{p}_1\|^2}{S}} \quad (65)$$

where the  $S$  is the number of Monte Carlo simulations and we use  $S = 300$  to  
 obtain the statistic results. As illustrated in Fig.3, we consider the scenario in  
 which There are three Base Stations (BS) at  $(2, 2)\text{km}$ ,  $(2, 0)\text{km}$  and  $(2, -2)\text{km}$   
 210 and  $Q = 2$  adjacent (relative to system's resolution) emitters at  $\mathbf{p}_1 = (0, 500)\text{m}$   
 and  $\mathbf{p}_2 = (0, -500)\text{m}$ . Each BS is equipped with an Uniform Linear Array (U-  
 LA) of  $M = 3$  elements. The signals transmitted by the two emitters are two  
 independent Gaussian random process. The Fourier coefficients of the transmit-  
 ted signals are, consequently, complex Gaussian RVs,  $s_q(n, l) \sim \mathcal{N}(0, \frac{10^{(SNR/10)}}{\sigma^2})$ .  
 215 The signal bandwidth is 20kHz. The number of sections  $L = 10$  and the number  
 of samples in each section is  $N = 5$ . The transmitted time is selected from a  
 Uniform Distribution  $U[0, 100\mu\text{s}]$ .

The super-parameter  $\rho_0 = 666$  in IW' expressions (61) is set as large as  
 possible while  $\rho_1 = 20$  and  $\rho_2 = 8$  are designed appropriately to make the ML  
 220 solutions  $\hat{\mathbf{p}}_q$  and  $\hat{t}_q$  included in the main lobe of  $g_q^{\text{mar}}(x_q, y_q)$  and  $g_q^{\text{con}}(\hat{t}_q | \mathbf{p}_q^{(r)})$   
 respectively. To match the size of the main lobe, the range parameters are  
 set as  $\Delta_x = \Delta_y = 400\text{m}$  and  $\Delta_{\hat{t}} = 10\mu\text{s}$ . These super-parameter values can  
 be, statistically, acquired by plenty of simulations.

Note that the transmitted signal correlation coefficient  $r_s$  in (26) would de-  
 225 crease until zero (the correlation information can be ignored) as the number  
 of sections  $L$  increases sufficiently, for the independence among  $Q$  transmitted  
 signals. Under the condition that  $L$  is sufficiently large, the single-ML-DPD  
 estimator would, therefore, achieve the performance of the optimal-ML-DPD  
 estimator approximately. To indicate the IS-ML-DPD estimator's advantage  
 230 in the case of few signal sections, we start with varying the number of sec-

tions  $L$  between 10 and 100 in the first simulation. Fig.4 shows that all three IS-ML-DPD performance curves (corresponding to the number of realizations  $R = 100, 500, 2000$  respectively) locate in the gap between the performance curves of single-ML-DPD and optimal-ML-DPD. The localization accuracy of IS-ML-DPD improves as the number of realizations  $R$  increase and achieves the accuracy of optimal-ML-DPD approximately at  $R = 2000$ . The reason is that generating sufficient realizations can guarantee that the realizations in vicinity of the ML solutions can be sampled. It is worthy to note that  $R = 2000$  realizations is just a few in the  $Q(D+1) = 6$ -dimensional parameter space. In another words, the complexity of brute searching in  $Q(D+1) = 6$ -dimensional parameter space is about  $\mathcal{O}(N_g^{Q(D+1)})$  ( $N_g$  is the number of grid points for each dimension), which is significantly larger than the complexity of IS-ML-DPD. The MVDR-DPD estimator's localization RMSE is around half of the distance between the two emitters at all  $L$  candidate values, which indicates that it fails to distinguish the two adjacent emitters at  $SNR = 5dB, Bw = 20kHz$  since the beforehand waveform information is not used sufficiently.

Fig.5 shows that the RMSE of each algorithm varies with SNR in the second simulation. The SNR varies between  $-5dB$  and  $15dB$ . It can be seen that the IS-ML-DPD (The number of realizations  $R = 2000$ ) also achieves, approximately, the performance of the optimal-ML-DPD at all SNR. The reason why there is a performance gap between single-ML-DPD and the IS-ML-DPD is that there exist correlation among emitters for the limited observation resources ( $B_w = 20kHz, L = 10$ ), which results in the interferences with one another. In such case, assuming the independence among emitters to decouple the multidimensional optimization problem is sub-optimal. On the condition of limited observation resources, the MVDR-DPD estimator is also unable to distinguish the two adjacent emitters until  $SNR = 15dB$ . The MVDR-DPD shows the higher resolution than single-ML-DPD and approximates the IS-ML-DPD at relatively high SNR.

The first two simulations indicates that the single-ML-DPD and MVDR-DPD are sub-optimal in the case of limited observation resources while the

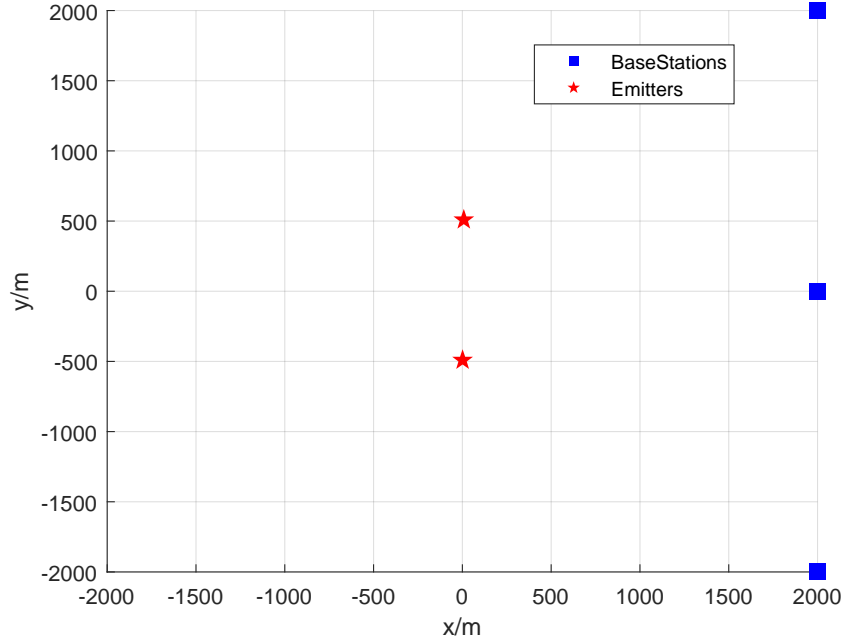


Figure 3: the simulation scenario

IS-ML-DPD is able to achieve the performance of the optimal-ML-DPD approximately in the same case. To further validate this result, we gradually increase the signal bandwidth  $B_w$  from 20kHz to 800kHz in the third simulation. As  
265 expected, the IS-ML-DPD outperforms the single-ML-DPD and MVDR-DPD, achieving the performance of the optimal-ML-DPD, at  $B_w \leq 400\text{kHz}$  and all estimators' performance converge as the  $B_w$  increases.

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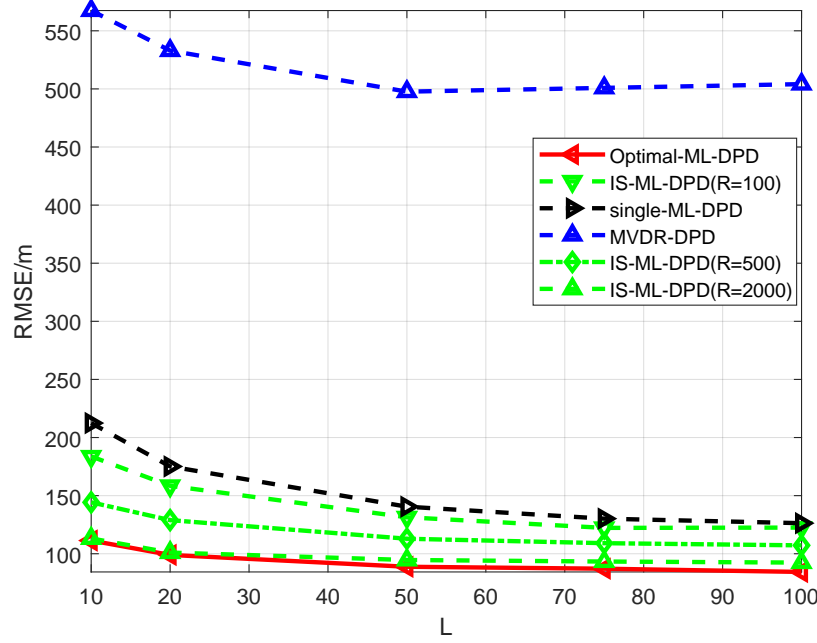


Figure 4: the RMSE of each estimator varies with the number of sections  $L$ . The signal bandwidth is 20kHz and the  $SNR = 5\text{dB}$ .

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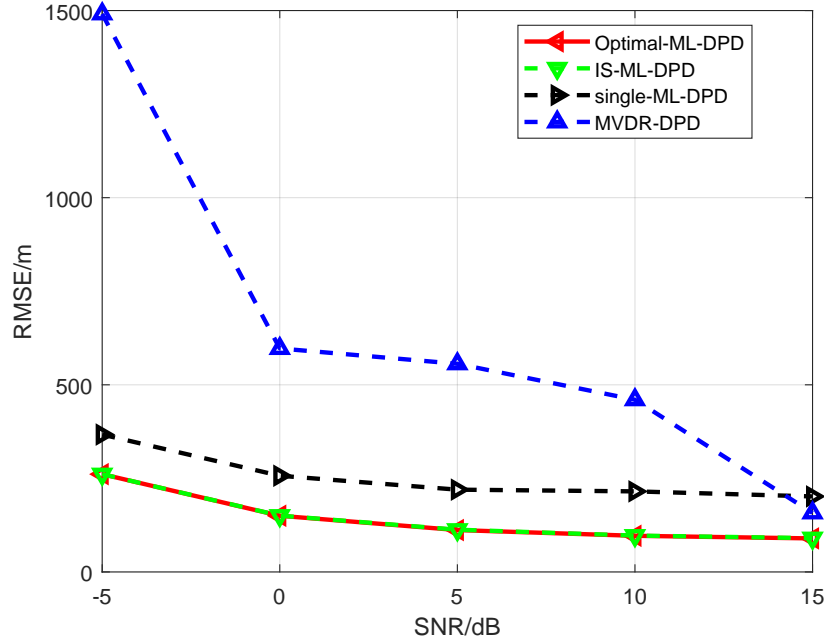


Figure 5: the RMSE of each estimator varies with the SNR. The signal bandwidth is 20kHz and the number of signal sections is  $L = 10$

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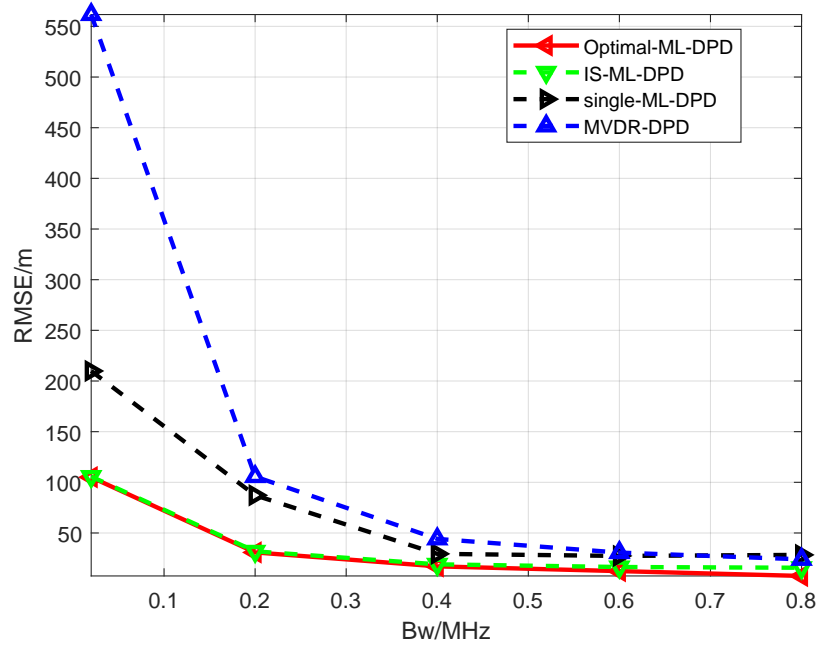


Figure 6: the RMSE of each estimator varies with the bandwidth.

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