Probability: Part III

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Bayes Theorem

Sometimes, we want to assess the probability that a particular event occurred on the basis of partial evidence. This can be achieved with Bayes Theorem.

Theorem 1

Suppose that a sample space S is the union of mutually disjoint events B_1 , B_2 , B_3 , ..., B_n , suppose A is an event in S, and suppose A and all the B_i have nonzero probabilities. If k is an integer with $1 \le k \le n$, then

$$p(B_k|A) = \frac{p(A|B_k)p(B_k)}{\sum_{i=1}^{n} p(A|B_i)p(B_i)}$$
(1)

Proof of Bayes' Theorem

By the definition of conditional probability, $p(A \cap B_k) = p(A)p(B_k|A)$. But

$$p(A \cap B_k) = p(B_k \cap A) = p(B_k)p(A|B_k).$$

Therefore,

$$p(B_k|A) = \frac{p(A|B_k)p(B_k)}{p(A)}.$$
 (2)

By assumption, $S = \bigcup_{i=1}^{n} B_i$, where $B_i \cap B_j = \emptyset$ for any (i,j) with $i \neq j$. Then

$$A = A \cap S = A \cap \left(\bigcup_{i=1}^{n} B_i \right) = \bigcup_{i=1}^{n} (A \cap B_i).$$

Note that $(A \cap B_i) \cap (A \cap B_j) = \emptyset$ for any (i,j) with $i \neq j$. We have

$$p(A) = p(\bigcup_{i=1}^{n} (A \cap B_i)) = \sum_{i=1}^{n} p(A \cap B_i) = \sum_{i=1}^{n} p(A|B_i)p(B_i).$$
 (3)

Combining (2) and (3) yields the desired equality.



Bayes Theorem

When k = 2, Bayes' theorem becomes the following.

Corollary 2

Let A and B be two events from a sample space S such that $p(A) \neq 0$ and $p(B) \neq 0$. Then

$$p(B|A) = \frac{p(A|B)p(B)}{p(A|B)p(B) + p(A|B^c)p(B^c)}.$$
 (4)

Bayes Theorem

Problem 3

We have two boxes. The first one contains 2 green balls and 7 red balls; the second one contains 4 green balls and 3 red balls. Bob selects a ball by first choosing one of the two boxes at random. He then selects one of the balls in this box at random. If Bob has selected a red ball, what is the probability that he selected a ball from the first box?

Solution

Let *B* be the event that Bob selected the ball from the first box, and *A* the event that he selected a red ball. Note that

$$p(B) = p(B^c) = \frac{1}{2}, \ p(A|B) = \frac{7}{9}, \ p(A|B^c) = \frac{3}{7}.$$

Then

$$p(B|A) = \frac{p(A|B)p(B)}{p(A|B)p(B) + p(A|B^c)p(B^c)} = \frac{49}{76}.$$

Random Variables

Definition 4

A <u>probability space</u> is a pair (S, p), where S is a sample space and p is probability function (or probability measure) on S.

Definition 5

A <u>random variable</u> is a **function** from the sample space S of a probability space (S, p) to the set \mathbb{R} of real numbers.

Remark

A random variable is a function, **not a variable.**

Random Variables

Example 6

Flip a fair coin three times. Let X(t) be the number of heads that occurs, where t is the outcome. Then

$$X(TTT) = 0,$$
 $X(TTH) = X(THT) = X(HTT) = 1,$
 $X(THH) = X(HHT) = X(HTH) = 2,$
 $X(HHH) = 3.$

Distribution of a Random Variable

Definition 7

The <u>distribution</u> of a random variable X on a probability space (S, p) is the set of pairs:

$$(r,p(X=r))$$
 for all $r \in X(S)$

where p(X = r) is the probability that X takes the value r.

A distribution is usually described by specifying p(X = r) for each $r \in X(S)$.

Example 8

For the random variable X in Example 6, the distribution is

$$p(X=0) = \frac{1}{8}, \ p(X=1) = p(X=2) = \frac{3}{8}, \ p(X=3) = \frac{1}{8}.$$

Expected Value of a Random Variable

Definition 9

The expected value (also called, expectation) of a random X(s) on a probability space (S, p) is defined by

$$E(X) = \sum_{s \in S} p(s)X(s) = \sum_{r} rp(X = r).$$

Remark

When there are infinitely many elements in S, the expectation is defined only when $\sum_{s \in S} p(s)X(s)$ is convergent.

Expected Value of a Random Variable

Example 10

Flip a fair coin three times. Let X(s) be the number of heads that occurs, where s is the outcome. For the random variable X, the distribution is

$$p(X=0)=\frac{1}{8}, \ p(X=1)=p(X=2)=\frac{3}{8}, \ p(X=3)=\frac{1}{8}.$$

Hence, the expected value is

$$E(X) = 0 \times \frac{1}{8} + 1 \times \frac{3}{8} + 2\frac{3}{8} + 3 \times \frac{1}{8} = \frac{3}{2}.$$

The Linearity of Expectation of a Random Variable

Theorem 11

Let X and X_i , i = 1, 2, ..., n, be random variables on a probability space (S, p), and let a and b be real numbers. Then

- ② E(aX + b) = aE(X) + b.

Proof.

By definition, we have

$$E\left(\sum_{i=1}^{n} X_{i}\right) = \sum_{s \in S} p(s) \sum_{i=1}^{n} X_{i}(s) = \sum_{i=1}^{n} \sum_{s \in S} p(s) X_{i}(s) = \sum_{i=1}^{n} E(X_{i}),$$

$$E(aX + b) = \sum_{s \in S} p(s) (aX + b)(s) = \sum_{s \in S} p(s) (aX)(s) + \sum_{s \in S} p(s)b.$$

Independent Random Variables

Definition 12

Two random variables X and Y on a probability space (S,p) are independent if

$$p(X(s) = r_1 \text{ and } Y(s) = r_2) = p(X(s) = r_1) \times p(Y(s) = r_2)$$

for all real numbers r_1 and r_2 .

Remark

The independence of two events in a sample space S is **usually** different from that of two random variables on S.

Independent Random Variables

Example 13

A pair of fair dice is rolled. Let X_1 and X_2 be the random variables denoting the numbers appearing on the first and second dice, respectively. Are X_1 and X_2 independent?

Solution

Since the pair of dice are fair and two rolling are independent of each other,

$$p(X_1 = r_1 \text{ and } X_2 = r_2) = \begin{cases} \frac{1}{36} & \text{if } r_1, r_2 \in \{1, 2, 3, 4, 5, 6\}, \\ 0 & \text{otherwise.} \end{cases}$$

On the other hand, $p(X_i = r) = 1/6$ if $r \in \{1,2,3,4,5,6\}$ and $p(X_i = r) = 0$ otherwise. Hence

$$p(X_1 = r_1 \text{ and } X_2 = r_2) = p(X_1 = r_1) \times p(X_2 = r_2)$$

for all real numbers r_1 and r_2 . Thus, X_1 and X_2 are independent.

Independent Random Variables

Proposition 14

Let X and Y be two independent random variables on a probability space (S,p). Then

$$E(XY) = E(X)E(Y).$$

Proof.

The proof is straightforward and left as an exercise.

Example 15

A pair of fair dice is rolled. Let X_1 and X_2 be the random variables denoting the numbers appearing on the first and second dice, respectively. Then $E(X_i) = \sum_{i=1}^6 \frac{1}{6}i = \frac{7}{2}$. Since X_1 and X_2 are independent,

$$E(X_1X_2) = E(X_1)E(X_2) = \frac{49}{4}.$$

Definition 16

Let X be a random variable on a probability space (S,p). The <u>variance</u> of X is

$$V(X) = \sum_{s \in S} [X(s) - E(X)]^2 p(s).$$

The standard deviation of X is defined to be $\sigma(X) = \sqrt{V(X)}$.

Remark

The expected value of a random variable informs us its average value, but does not give us information about how widely its values are distributed. The variance of a variable does this for us.

Proposition 17

If X is a random variable on a probability space (S,p), then $V(X) = E(X^2) - E(X)^2$.

Proof.

We have

$$V(X) = \sum_{s \in S} [X(s) - E(X)]^2 \rho(s)$$

$$= \sum_{s \in S} X(s)^2 \rho(s) - 2E(X) \sum_{s \in S} X(s) \rho(s) + E(X)^2 \sum_{s \in S} \rho(s)$$

$$= E(X^2) - 2E(X)^2 + E(X)^2$$

$$= E(X^2) - E(X)^2.$$

Example 18

Flip a fair coin three times. Let X(s) be the number of heads that occurs, where s is the outcome. For the random variable X, the distribution is

$$p(X=0) = \frac{1}{8}, \ p(X=1) = p(X=2) = \frac{3}{8}, \ p(X=3) = \frac{1}{8}.$$

Hence, the expected value is

$$E(X) = 0 \times \frac{1}{8} + 1 \times \frac{3}{8} + 2\frac{3}{8} + 3 \times \frac{1}{8} = \frac{3}{2}.$$

and

$$E(X^2) = 0 \times \frac{1}{8} + 1 \times \frac{3}{8} + 4\frac{3}{8} + 9 \times \frac{1}{8} = 3.$$

Thus, $V(X) = E(x^2) - E(X)^2 = \frac{3}{4}$ and $\sigma(X) = \frac{\sqrt{3}}{2}$.

Proposition 19

Let X and Y be two independent random variables on a probability space (S,p). Then

$$V(X+Y)=V(X)+V(Y).$$

Proof.

By Proposition 14, E(XY) = E(x)E(Y). It then follows from Proposition 17 and Theorem 11 that

$$V(X+Y) = E((X+Y)^2) - (E(X+Y))^2$$

$$= E(X^2 + Y^2 + 2XY) - (E(X) + E(Y))^2$$

$$= E(X^2) + E(Y^2) + 2E(XY) - E(X)^2 - E(Y)^2 - 2E(X)E(Y)$$

$$= V(X) + V(Y).$$

The Expected Value, Variance and Standard Deviation of Bernoulli Trials

Proposition 20

In n Bernoulli trials with probability p of success, the expected value is np, the variance is np(1-p) and the standard deviation is $\sqrt{np(1-p)}$.

Proof.

We consider the case n=1 and let X denote the number of successes in one Bernoulli trial. Then p(X=1)=p and p(X=0)=(1-p). Hence $E(X)=1\times p+0\times (1-p)=p$. Similarly, $E(X^2)=1^2\times p+0^2\times (1-p)=p$. It then follows that $V(X)=E(X^2)-E(X)^2=p-p^2=p(1-p)$. Let X_i denote the number of success in the ith Bernoulli trial. Then X_1, X_2, \cdots, X_n are independent random variables. It then follows from Proposition 19 that

$$E(X_1 + X_2 + \dots + X_n) = nE(X_1) = np,$$

 $V(X_1 + X_2 + \dots + X_n) = nV(X_1) = np(1-p).$

The Central Limit Theorem

Theorem 21

Let X_1, X_2, \dots, X_n be **independent and identically distributed** (in short, i.i.d.) random variables on a probability space (S,p) (i.e., they all have the same distribution and are mutually independent). Let μ and σ the expected value and the standard deviation of all X_i . Define

$$Z=\frac{X_1+X_2+\cdots+X_n-n\mu}{\sigma\sqrt{n}}.$$

Then

$$p(a \le Z \le b) = \frac{1}{\sqrt{2\pi}} \int_a^b e^{-x^2/2} dx.$$

Remark

A proof of this theorem can be found in a textbook on probability.



The Normal Distribution

Definition 22

A random variable X on a sample space S is said to have the **normal distribution** if its probability distribution is given by

$$p(X = x) = \phi_{(\mu,\sigma)}(x) := \frac{1}{\sigma\sqrt{2\pi}}e^{-(x-\mu)^2/2\sigma^2}, \ x \in \mathbb{R}.$$

Remarks

- ullet The normal distribution has expected value μ and standard deviation σ .
- It is called the **standard normal distribution** when $\mu = 0$ and $\sigma = 1$.
- $\bullet \int_{-\infty}^{\infty} \varphi_{(\mu,\sigma)}(x) dx = 1.$
- The Central Limit Theorem says that the probability distribution of

$$Z = \frac{X_1 + X_2 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

will tend to the standard normal distribution as *n* increases.

An Example of the Normal Distribution

Example 23

Consider n Bernoulli trials with probability p of success. Let X_i denote the number of success in the ith trial. Then X_1, X_2, \dots, X_n are i.i.d., and each has expected value p and standard deviation $\sqrt{p(1-p)}$.

One can verify that the probability distribution of the **standardized random variable**

$$Z = \frac{X_1 + X_2 + \dots + X_n - np}{\sqrt{np(1-p)}} = \frac{X_1 + X_2 + \dots + X_n - E(X_1 + X_2 + \dots + X_n)}{\sigma(X_1 + X_2 + \dots + X_n)}$$

converges to the standard normal distribution as n approaches to ∞ .