ST502 Project

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Part I

Framingham Heart Study

The Framingham data set contains the diastolic blood pressure of 300 smokers and nonsmokers. For this study we will assume the following:

H₀: μ_1 - μ_2 = 0

 $H_A: \mu_1 - \mu_2 \neq 0$

The null hypothesis being there is no difference in the blood pressure between smokers and non smokers. We believe that there is a difference and this paper will prove if we have enough evidence to reject the null hypothesis.

Normality of Data

Plotting the density of total data (see Appendix I: Figure 1A), we can that it closely follows a normal distribution, Density plots of the split data exhibit similar forms (see Appendix I: Figure 1B). We can also calculate the kurtosis and the skew of the data, which are 3.812, 0.88 for kurtosis and skew respectively, and see that they are quite close to values typically found in normally distributed data. Additionally, the calculated values of sample standard deviation, $\frac{IQR}{1.35}$, and $\frac{MAD}{0.675}$ are 22.89, 22.22, and 21.48. The similarities between the three values indicate that the data is not influenced by outliers. Therefore, we will assume the data and the subsequent split data to be normal.

Statistical Analysis

In the first analysis we will be assuming equal variances, thereby allowing us to use pooled sample variance (see Appendix I: Equation 1).

The pooled sample variance of the data is 510 using a degree of freedom value 298. In this case, we reject the null hypothesis because the calculated p value, 0.0041, is smaller than the chosen α value of 0.05. In terms of t values, our observed t value of -3.04 is smaller than the t value, -1.97 for a two sided α of 0.05.

When computing the observed t value using the assumption that the population variances are not equivalent (variance smoker is 352.2 and variance nonsmoker is 562.1), a value of -2.9 is obtained. Comparing that the two sided α of -1.98 with 158 degrees of freedom (as computed using the Satterthwaite Approximation see Appendix I: Equation 4). The observed t value is less than the chosen α value of 0.05. In case, there is sufficient evidence to reject the null hypothesis in favor of the alternative.

The observed 95% confidence intervals are -15.08 to -3.23 for pooled sample variance and -15.4 to -2.91 for non pooled sample variance. It can be seen that 0 does not fall into the 95% confidence interval in either case. Therefore the null hypothesis can be rejected.

In all three case, there is sufficient evidence to reject the null hypothesis and to support the alternative at an α level of 0.05.

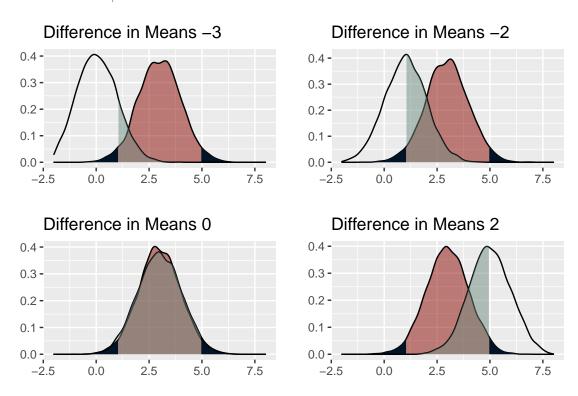
Part II

Introduction

In Part I we were able to see how hypothesis testing allowed to find enough evidence to challenge the "status quo" argument that smoking has no affect on the blood pressure of a person. What we found, through a number of robust statistical methods, that the difference in means, calculated to be 9.1577778, was significant to not have arisen by chance. That within an α of 5%, we are confident that the two groups do exhibit measurable differences. In the following sections, we will explore, through simulated data, the concept of randomness in data and what steps we can take to mitigate that randomness. All simulated date will be generated using the built in rnorm() function in R.

Power of Hypothesis Testing

It is important to briefly discuss the concept of power in hypothesis testing. Power describes the probability of not committing a Type II error, which is to not reject the null hypothesis when there was in actuality enough evidence to do so. The probability of a Type II error is represented by β . Power is the complement to that probability $(1 - \beta)$. The value of β is the portion of the alternate distribution that is within the null hypothesis non rejection region limits. Below are a number of plots to help depict this concept. In each of the plots the shaded red is the null distribution and its location stays constant (mean = 3). The dark tails represent the rejection region of the null distribution. The outline distribution represents the "true" alternative distribution. Within that distribution is the shaded light blue region which is equal to β . It can be seen that as the difference between the two distributions shrinks the value of β increases and the power shrinks. Power reaches minimum and β reaches maximum when the two distributions are the same.



Simulation Study

Below, various scenarios were simulated by creating 2 normally distributed data sets of various mean, variances, and sample sizes. Each scenario was repeated a thousand times. For each of those repeats a hypothesis test was conducted and the number of times the null hypothesis was rejected was recorded. See Appendix III for results. All hypothesis testing was conducted assuming different values of variance, regardless of the actual variance values.

The following table includes the values of each parameter used in the simulation:

Parameter	Possible Values
$\overline{\mu_0}$	0, 4, 5, 6, 10
σ_0^2	1, 4, 9
n_0	10, 30, 70, 100
μ_A	5
σ_A^2	1
n_A	10, 30, 70, 100

It was noted that the smaller difference between the alternative and null mean values were, the less likely the null hypothesis was rejected. Larger values of variance also reduced the number of rejected null hypothesis. In both cases, the more overlap between the two distributions, the less powerful the test. One way that can increase the power of the t-test is to increase the sample size. For a specific example, we can look at test cases #12 and #27. In both cases the distribution means were 6 and 5 for the null and alternative hypotheses respectively, along with variances of 1 for both distributions. The only difference between the two cases was the sample size ($n_{12} = 10$ and $n_{27} = 100$). This increase in sample size resulted in almost a 2 fold increase in power (570 vs 1000 rejected nulls). It stands to reason that large variances and small deltas in mean can be mitigated by increasing the sample size accordingly.

Appendix I: Equations and Figures

Figures

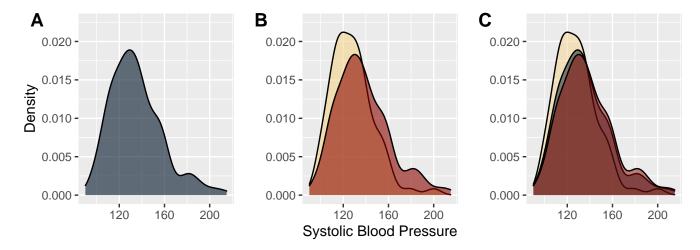


Figure 1: Data plots

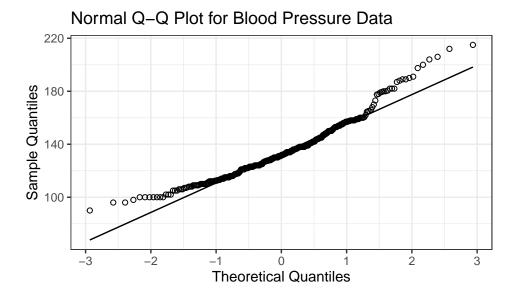


Figure 2: Q-Q Plot

Equations:

Equation 1: Pooled Sample Variance:

$$S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$$

Equation 2: Observed T Statistic (pooled variance):

$$t_{obs} = \frac{\mu_1 - \mu_2}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

Equation 3: Observed T Statistic (distinct variance):

$$t_{obs} = \frac{\mu_1 - \mu_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

Equation 4: Satterthwaite Approximation:

$$\nu = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)^2}{\frac{\left(\frac{S_1^2}{n_1}\right)^2}{n_1 - 1} + \frac{\left(\frac{S_2^2}{n_2}\right)^2}{n_2 - 1}}$$

Appendix II: Results

Data Mean: 134.935

Data Sample Variance: 524.0852258

Smoker Mean: 128.0666667 Smoker Variance: 352.2117117 Nonsmoker Mean: 137.2244444 Nonsmoker Variance: 562.1447123

Difference in mean: -9.1577778

Pooled Sample Variance: 510.0136987 Nonpooled Sample Variance: 9.9936838

CI Pooled: -15.08, -3.23 CI Nonpooled: -15.4, -2.91

Appendix III: Part II Results

Test#	μ_0	σ_0^2	n_0	μ_a	σ_a^2	n_a	Test Results
1	0	1	10	5	1	10	Test Results 1: 1000
2	4	1	10	5	1	10	Test Results 2: 574
3	5	1	10	5	1	10	Test Results 3: 50
4	6	1	10	5	1	10	Test Results 4: 565
5	10	1	10	5	1	10	Test Results 5: 1000
6	5	1	10	5	1	10	Test Results 6: 39
7	5	4	10	5	1	10	Test Results 7: 55
8	5	9	10	5	1	10	Test Results 8: 61
9	6	1	10	5	1	10	Test Results 9: 570
10	6	4	10	5	1	10	Test Results 10: 262
11	6	9	10	5	1	10	Test Results 11: 186
12	6	1	10	5	1	10	Test Results 12: 570
13	6	1	30	5	1	10	Test Results 13: 764
14	6	1	70	5	1	10	Test Results 14: 821
15	6	1	100	5	1	10	Test Results 15: 861
16	6	1	10	5	1	30	Test Results 16: 743
17	6	1	30	5	1	30	Test Results 17: 971
18	6	1	70	5	1	30	Test Results 18: 993
19	6	1	100	5	1	30	Test Results 19: 997
20	6	1	10	5	1	70	Test Results 20: 838
21	6	1	30	5	1	70	Test Results 21: 997
22	6	1	70	5	1	70	Test Results 22: 1000
23	6	1	100	5	1	70	Test Results 23: 1000
24	6	1	10	5	1	100	Test Results 24: 830
25	6	1	30	5	1	100	Test Results 25: 996
26	6	1	70	5	1	100	Test Results 26: 1000
27	6	1	100	5	1	100	Test Results 27: 1000
28	6	1	30	5	1	100	Test Results 28: 999
29	6	4	30	5	1	100	Test Results 29: 899
30	6	9	30	5	1	100	Test Results 30: 721
31	6	4	70	5	1	100	Test Results 31: 980
32	6	4	100	5	1	100	Test Results 32: 995
33	6	9	70	5	1	100	Test Results 33: 847
34	6	9	100	5	1	100	Test Results 34: 885

Appendix IV: Code

```
knitr::opts chunk$set(echo = TRUE)
#Use required packages
library(tidyverse) #for plots and data manipulation
library(cowplot) #aligning plots
library(gridExtra)
library(scales)
df_data <- read_csv("framingham_data.csv") # Read in data</pre>
df_data$index <- seq(nrow(df_data)) # Add an index column</pre>
#df_data %>% summary # Summarize Data
# Split data into smoker and nonsmoker
df_smoker <- df_data %>% filter(currentSmoker == 1)
df_nonsmoker <- df_data %>% filter(currentSmoker == 0)
#Create a sample variance function to ensure proper calculation
sample_variance <- function(x, sampling = TRUE){</pre>
    if (sampling == TRUE){
        sum((x - mean(x))^2) / (length(x) - 1)
    } else if(sampling == FALSE) {
        sum((x - mean(x))^2) / (length(x))
}
#Create pooled sample variance function
f_pooled_variance <- function(x, y){</pre>
    ((length(x) - 1) * sample_variance(x) +
     (length(y) - 1) * sample_variance(y)) /
    (length(x) + length(y) - 2)
}
# Skewness function
skew_function <- function(x){</pre>
    mean((x - mean(x))^3) / sqrt(sample_variance(x))^3
}
# kurtosis function
kurt_function <- function(x){</pre>
    mean((x - mean(x))^4) / sqrt(sample_variance(x))^4
}
# Create a Satterthawaite Approximation Function
satterth <- function(s1, s2, n1, n2){</pre>
    term1 <- s1/n1
    term2 \leftarrow s2/n2
    nu \leftarrow (term1 + term2)^2 / ((term1^2/(n1 - 1)) + (term2^2/(n2 - 1)))
    return(floor(nu))
}
#Plot and compare split data
#options(repr.plot.width = 6, repr.plot.height = 4, repr.plot.res = 150)
```

```
plot_colors <- c("#001427","#708d81","#f4d58d","#bf0603","#8d0801")
y_{limits} \leftarrow c(0, 0.0225)
total_data <- ggplot(df_data) + geom_density(aes(sysBP),</pre>
                                                fill = plot_colors[1],
                                                alpha = 0.6) +
                        ylim(y_limits) + ylab("Density") + xlab("")
sep_data <- ggplot() + geom_density(data = df_smoker, aes(sysBP),</pre>
                                      fill = plot_colors[3], alpha = 0.6) +
                        geom_density(data = df_nonsmoker, aes(sysBP),
                                      fill = plot_colors[5], alpha = 0.6) +
                        ylim(y_limits) + ylab("") + xlab("Systolic Blood Pressure")
plot_3 <- ggplot() + geom_density(data = df_smoker, aes(sysBP),</pre>
                                      fill = plot_colors[3], alpha = 0.5) +
                        geom_density(data = df_data, aes(sysBP),
                                      fill = plot_colors[1], alpha = 0.5) +
                        geom_density(data = df_nonsmoker, aes(sysBP),
                                      fill = plot_colors[5], alpha = 0.5) +
                        ylim(y_limits) + ylab("") + xlab("")
#plot_grid(total_data, sep_data, plot_3, align = 'vh',
           hjust = -1, nrow = 1, ncol = 3
data_kurtosis <- kurt_function(df_data$sysBP)</pre>
data_skew <- skew_function(df_data$sysBP)</pre>
data_IQR <- as.numeric(quantile(df_data$sysBP, probs = 0.75)) -</pre>
       as.numeric(quantile(df_data$sysBP, probs = 0.25))
data_MAD <- median(abs(df_data$sysBP - median(df_data$sysBP)))</pre>
data_samVar <- sample_variance(df_data$sysBP)</pre>
eIQR <- data_IQR / 1.35
eMAD <- data_MAD / 0.675
# Q-Q Plot
data_qqplot <-
ggplot(df_data, aes(sample = sysBP)) +
stat_qq(shape = 1) + stat_qq_line() +
ggtitle("Normal Q-Q Plot for Blood Pressure Data") +
xlab("Theoretical Quantiles") +
ylab("Sample Quantiles")
# Common values for analysis
alpha <- 0.05
mu_smoker <- mean(df_smoker$sysBP)</pre>
var_smoker <- sample_variance(df_smoker$sysBP)</pre>
n_smoker <- length(df_smoker$sysBP)</pre>
mu_nonsmoker <- mean(df_nonsmoker$sysBP)</pre>
var_nonsmoker <- sample_variance(df_nonsmoker$sysBP)</pre>
n_nonsmoker <- length(df_nonsmoker$sysBP)</pre>
# Two Sample T-test - Pooled Sample Variance - P-value
```

```
dof_1 <- (n_smoker + n_nonsmoker - 2)</pre>
p_sample_var_1 <- f_pooled_variance(df_smoker$sysBP,</pre>
                                      df_nonsmoker$sysBP)
t_obs_1 <- (mu_smoker - mu_nonsmoker) / (sqrt(p_sample_var_1) * sqrt(1/n_smoker + 1/n_nonsmoker))
t_stat_1 <- qt(alpha / 2, dof_1)
p_value_obs_1 <- dt(t_obs_1, dof_1)</pre>
#Two Sample T-test - Difference Variance Sample Variance - P-value
dof_2 <- satterth(var_smoker, var_nonsmoker, n_smoker, n_nonsmoker)</pre>
np_sample_var_2 <- (var_nonsmoker/n_smoker + var_nonsmoker/n_nonsmoker)</pre>
t_obs_2 <- (mu_smoker - mu_nonsmoker) / (sqrt(var_nonsmoker/n_smoker + var_nonsmoker/n_nonsmoker))
t_stat_2 <- qt(alpha / 2, dof_2)</pre>
p_value_obs_2 <- dt(t_obs_2, dof_2)</pre>
# Confidence Limits
diff_mu <- mu_smoker - mu_nonsmoker
#Pooled Sample variance
CL_pooled <- t_stat_1 * (sqrt(p_sample_var_1/n_smoker + p_sample_var_1/n_nonsmoker))
#Non pooled Sample variance
CL_nonpooled <- t_stat_2 * (sqrt(var_nonsmoker/n_smoker + var_nonsmoker/n_nonsmoker))
CI_pooled <- round(c(diff_mu + CL_pooled, diff_mu - CL_pooled), 2)</pre>
CI_nonpooled <-round(c(diff_mu + CL_nonpooled, diff_mu - CL_nonpooled), 2)
#Part II
#Introduction
options(repr.plot.width = 12, repr.plot.height = 5, repr.plot.res = 150)
set.seed(100)
null_mean <- 3</pre>
alt_means <- c(0, 1, 3, 5)
plot_list <- list()</pre>
#plot_colors <- c("#072ac8", "#1e96fc", "#a2d6f9", "#fcf300", "#ffc600")</pre>
for(i in 1:length(alt_means)){
    sim1 <- rnorm(5000, null_mean, sqrt(1))</pre>
    sim2 <- rnorm(5000, alt_means[i], sqrt(1))</pre>
```

```
alpha1 <- qnorm(0.025, null_mean, sqrt(1))
    alpha2 <- qnorm(0.975, null_mean, sqrt(1))</pre>
    df_set <- tibble("HO" = sim1, "HA" = sim2)</pre>
    title_string <- sprintf("Difference in Means %i", (alt_means[i] - null_mean))
    plot_list[[i]] <-</pre>
    ggplot(data = df_set) + geom_density(aes(H0), alpha = 0.5, fill = plot_colors[5]) +
                             geom_area(
                                  aes(x = stage(H0, after_scale = oob_censor(x, c(-Inf, alpha1))
                                                )
                                     ),
                                  stat = "density", fill = plot_colors[1]
                                   ) +
                             geom_area(
                                  aes(x = stage(H0, after_scale = oob_censor(x, c(alpha2, Inf)
                                                )
                                     ),
                                  stat = "density", fill = plot_colors[1]
                              geom_density(aes(HA), alpha = 0.5) +
                              geom_area(
                                  aes(x = stage(HA, after_scale = oob_censor(x, c(alpha1, alpha2))
                                     ),
                                  stat = "density", fill = plot_colors[2], alpha = 0.5
                             xlim(-2, 8) + xlab("") + ylab("") + ggtitle(title_string)
                                  }
do.call(grid.arrange, plot_list)
#Part II
set.seed(1)
alpha <- 0.05
test_function <- function (x, y, pooled = FALSE){</pre>
    mu_1 \leftarrow mean(x)
    var_1 <- sample_variance(x, sampling = TRUE)</pre>
    mu_2 \leftarrow mean(y)
    var_2 <- sample_variance(y, sampling = TRUE)</pre>
    #Calculate the pooled sample variance
    pooled\_sample \leftarrow ((length(x) - 1) * var_1 + (length(y) - 1) * var_2) / (length(x) + length(y) - 2)
    #calculate the observed t statistic
    if (pooled == TRUE){
        cal_sigma <- (sqrt(pooled_sample/length(x) + pooled_sample/length(y)))</pre>
```

```
ttest <- (mu_1 - mu_2) / cal_sigma
        dof <- length(x) + length(y) - 2 #Determine degrees of freedom</pre>
        } else {
        cal_sigma <- (sqrt(var_1/length(x) + var_2/length(y)))</pre>
        ttest <- (mu_1 - mu_2) / cal_sigma
        dof <- satterth(var_1, var_2, length(x), length(y))</pre>
    }
    #Determine whether or not the null hypothesis can be rejected (1 = rejected, 0 = not rejected)
    verdict <- !between(ttest, qt(alpha / 2, dof), qt(1 - alpha / 2, dof))</pre>
    #Return calculated values
    return(c(mu_1, var_1, mu_2, var_2, ttest, cal_sigma, dof,verdict))
}
mu1 \leftarrow c(0, 4, 5, 6, 10)
var1 \leftarrow c(1, 4, 9)
n1 \leftarrow c(10, 30, 70, 100)
mu2 <- 5
var2 <- 1
n2 < c(10, 30, 70, 100)
sim_test <- function(x_mu, x_var, x_n, y_mu, y_var, y_n, pooled = TRUE){</pre>
    sim_data_results <- matrix(rep(0, 8), ncol = 8)</pre>
    for (i in 1:1000){
        sim_set1 <- rnorm(x_n, x_mu, sqrt(x_var))</pre>
        sim_set2 <- rnorm(y_n, y_mu, sqrt(y_var))</pre>
        sim_data_results <- rbind(sim_data_results, test_function(sim_set1, sim_set2, pooled))</pre>
        #print(sim_data_results)
    }
    df_sim_data <- data.frame(sim_data_results[2 : nrow(sim_data_results),])</pre>
    colnames(df_sim_data) = c("Null Mean", "Null Variance", "Alternate Mean", "Alternate Variance",
                                    "T statistic", "Calculated Variance", "DoF", "Null Reject")
    return(df_sim_data)
}
# HA: mean = 5, var = 1
test_results_1 <- sim_test(mu1[1], var1[1], n1[1], mu2, var2, n2[1])
test_results_2 <- sim_test(mu1[2], var1[1], n1[1], mu2, var2, n2[1])
test_results_3 <- sim_test(mu1[3], var1[1], n1[1], mu2, var2, n2[1])
test_results_4 <- sim_test(mu1[4], var1[1], n1[1], mu2, var2, n2[1])
test_results_5 <- sim_test(mu1[5], var1[1], n1[1], mu2, var2, n2[1])
test_results_6 <- sim_test(mu1[3], var1[1], n1[1], mu2, var2, n2[1])
test_results_7 <- sim_test(mu1[3], var1[2], n1[1], mu2, var2, n2[1])
test_results_8 <- sim_test(mu1[3], var1[3], n1[1], mu2, var2, n2[1])
```

```
test_results_9 <- sim_test(mu1[4], var1[1], n1[1], mu2, var2, n2[1])
test_results_10 <- sim_test(mu1[4], var1[2], n1[1], mu2, var2, n2[1])
test_results_11 <- sim_test(mu1[4], var1[3], n1[1], mu2, var2, n2[1])
test_results_12 <- sim_test(mu1[4], var1[1], n1[1], mu2, var2, n2[1])
test_results_13 <- sim_test(mu1[4], var1[1], n1[2], mu2, var2, n2[1])
test_results_14 <- sim_test(mu1[4], var1[1], n1[3], mu2, var2, n2[1])
test_results_15 <- sim_test(mu1[4], var1[1], n1[4], mu2, var2, n2[1])
test_results_16 <- sim_test(mu1[4], var1[1], n1[1], mu2, var2, n2[2])
test_results_17 <- sim_test(mu1[4], var1[1], n1[2], mu2, var2, n2[2])
test_results_18 <- sim_test(mu1[4], var1[1], n1[3], mu2, var2, n2[2])
test_results_19 <- sim_test(mu1[4], var1[1], n1[4], mu2, var2, n2[2])
test_results_20 <- sim_test(mu1[4], var1[1], n1[1], mu2, var2, n2[3])
test_results_21 <- sim_test(mu1[4], var1[1], n1[2], mu2, var2, n2[3])
test_results_22 <- sim_test(mu1[4], var1[1], n1[3], mu2, var2, n2[3])
test_results_23 <- sim_test(mu1[4], var1[1], n1[4], mu2, var2, n2[3])
test_results_24 <- sim_test(mu1[4], var1[1], n1[1], mu2, var2, n2[4])
test_results_25 <- sim_test(mu1[4], var1[1], n1[2], mu2, var2, n2[4])
test_results_26 <- sim_test(mu1[4], var1[1], n1[3], mu2, var2, n2[4])
test_results_27 <- sim_test(mu1[4], var1[1], n1[4], mu2, var2, n2[4])
test_results_28 <- sim_test(mu1[4], var1[1], n1[2], mu2, var2, n2[4])
test_results_29 <- sim_test(mu1[4], var1[2], n1[2], mu2, var2, n2[4])
test_results_30 <- sim_test(mu1[4], var1[3], n1[2], mu2, var2, n2[4])
test_results_31 <- sim_test(mu1[4], var1[2], n1[3], mu2, var2, n2[4])
test_results_32 <- sim_test(mu1[4], var1[2], n1[4], mu2, var2, n2[4])
test_results_33 <- sim_test(mu1[4], var1[3], n1[3], mu2, var2, n2[4])
test_results_34 <- sim_test(mu1[4], var1[3], n1[4], mu2, var2, n2[4])
# Plotting
options(repr.plot.width = 6, repr.plot.height = 4, repr.plot.res = 150)
plot_grid(total_data, sep_data, plot_3, align = 'vh',
          hjust = -1, nrow = 1, ncol = 3, labels = c("A", "B", "C"))
data_qqplot + theme_bw()
```