1          Unravelling the time-course of listener adaptation to an unfamiliar talker

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10    Abstract

11    YOUR ABSTRACT GOES HERE. All data and code for this study are shared via OSF,

12    including the R markdown document that this article is generated from, and an R library that

13    implements the models we present.

14        *Keywords:* speech perception; adaptation; incremental changes; distributional learning

15        Word count: X

16  # Unravelling the time-course of listener adaptation to an unfamiliar talker

17  TO-DO

18  ## 0.1  Highest priority

19  - MARYANN

20  ### 0.1.1  Lower Priority

21  - Decide on PSE vs. category boundary
22  - standardize BE vs. AE spelling (categoriz/sation, label(l)ed, synthesiz/sed etc.)

23  ## 0.2  To do later

24  - Everyone: Eat ice-cream and perhaps have a beer.

## 1  Introduction

One of the hallmarks of human speech perception is its adaptivity. Listeners' interpretation of acoustic input can change within minutes of exposure to an unfamiliar talker, supporting robust speech recognition across talkers (Bradlow & Bent, 2008; Clarke & Garrett, 2004; Xie, Liu, & Jaeger, 2021; Xie et al., 2018). Recent reviews have identified distributional learning of marginal cue statistics ('normalization,' Apfelbaum & McMurray, 2015; McMurray & Jongman, 2011) or the statistics of cue-to-category mappings as an important mechanisms affording this adaptivity ('representational learning,' Clayards, Tanenhaus, Aslin, & Jacobs, 2008; D. F. Kleinschmidt & Jaeger, 2015; **idemaru-hold2011?**; **davis-sohoglu2020?**; for review, **schertz-clare2019?**; **xie2023?**). This hypothesis has gained considerable influence over the past decade, with findings that changes in listener perception are qualitatively predicted by the statistics of exposure stimuli (Bejjanki, Clayards, Knill, & Aslin, 2011; Clayards et al., 2008; Nixon, Rij, Mok, Baayen, & Chen, 2016; Tan, Xie, & Jaeger, 2021; **idemaru2021?**; **kleinschmidt2012?**; **munson2011-thesis?**; **theodore2019distributional?**; for important caveats, see **harmon2018?**).

We investigate an important constraints on this type of adaptivity that is suggested by recent findings. D. F. Kleinschmidt and Jaeger (2016) exposed L1-US English listeners to recordings of /b/-/p/ minimal pair words like *beach* and *peach* that were acoustically manipulated. Separate groups of listeners were exposed to distributions of voice onset times (VOTs)—the primary cue distinguishing between *beach* and *peach*—that were shifted by XXX to XXX msecs, respectively, relative to what one might expect from a 'typical' talker (Figure **??**A). In line with the distributional learning hypothesis, listeners' category boundary or point of subjective equality (PSE)—i.e., the VOT for which listeners are equally likely to respond "d" or "t"—shifted in the same direction as the exposure distribution (Figure **??**B). Also in line with the distributional learning hypothesis, these shifts were larger the further the exposure distributions were shifted. However, Kleinschmidt and Jaeger also observed a previously undocumented property of these adaptive changes: shifts in the exposure distribution had less than proportional (sublinear) effect on shifts in PSE (Figure **??**C). While this finding—recently replicated in one more experiment (Dave Kleinschmidt, 2020, Experiment 4)—is broadly compatible with the hypothesis of distributional learning, it points to important not well-understood constraints on

54  adaptive speech perception.

55      For example, influential *models* of adaptive speech perception predict proportional, rather

56  than sublinear, shifts (for proof, see SI **??**). This is the case both for incremental Bayesian

57  belief-updating model (D. F. Kleinschmidt & Jaeger, 2011) and general purpose normalization

58  accounts (McMurray & Jongman, 2011)—models that have been found to explain listeners'

59  behavior well in experiments with less substantial changes in exposure. There are, however,

60  proposals that can accommodate this finding. Some proposals distinguish between two types of

61  mechanisms that might underlie representational changes, *model learning* and *model selection*

62  (**xie2018?**). The former refers to the learning of a new category representations—for example,

63  learning a new generative model for the talker (D. F. Kleinschmidt & Jaeger, 2015, pt. II) or

64  storage of new talker-specific exemplars (Sumner, 2011; **johnson1997?**). Xie and colleagues

65  hypothesized that this process might be much slower than is often assumed in the literature,

66  potentially requiring multiple days of exposure and memory consolidation during sleep (see also

67  **fenn2013?**; **tamminen2012?**; **xie2018sleep?**). Rapid adaptation that occurs within minutes of

68  exposure might instead be achieved by selecting between *existing* talker-specific representations

69  that were learned from previous speech input—e.g., previously learned talker-specific generative

70  models (see mixture model in D. F. Kleinschmidt & Jaeger, 2015, pp. 180–181) or previously

71  stored exemplars from other talkers (**johnson1997?**). Model learning and model selection both

72  offer explanations for the sublinear effects observed in D. F. Kleinschmidt and Jaeger (2016). But

73  they suggest different predictions for the evolution of this effect over the course of exposure.

74      Under the hypothesis of model learning, sublinear shifts in PSEs can be explained by

75  assuming a hierarchical prior over talker-specific generative models ($p(\Theta)$ in D. F. Kleinschmidt &

76  Jaeger, 2015, p. 180). This prior would 'shrink' adaptation towards listeners' priors—similar to

77  the effect of random by-subject or by-item effects in generalized linear mixed-effect models, which

78  shrink group-level effect estimates towards the population mean of the data (**bates?**). Critically,

79  as long as these priors attribute non-zero probability to even extreme shifts (e.g., the type of

80  Gaussian prior used in mixed-effects models), this predicts listeners' PSEs will continue to change

81  with increasing exposure until they have converged against the PSE that is ideal for the exposure

82  statistics. In contrast, the hypothesis of model selection predicts that rapid adaptation is more

83 strongly constrained by previous experience: listeners can only adapt their categorisation

84 functions up to a point that corresponds to (a mixture of) previously experienced talker-specific

85 generative models. Figure 1 visualizes the contrasting predictions of model learning and selection

86 for incremental adaptation.
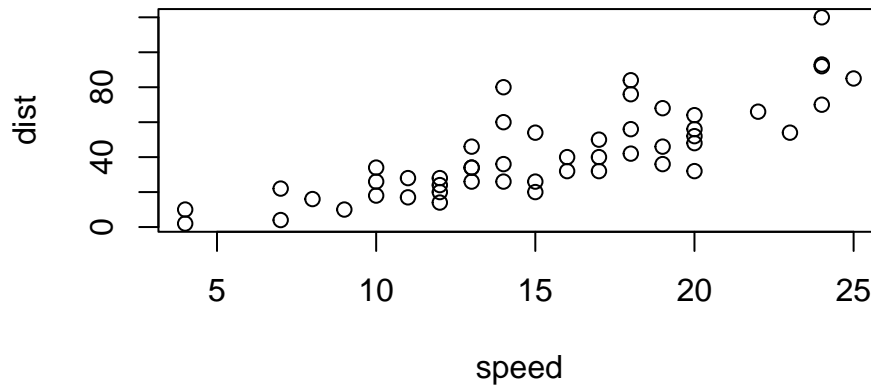


*Figure 1.* Contrasting predictions of model learning and model selection hypotheses about the incremental effects of exposure on listeners' categorisation function. Both hypothesis predict incremental adaptation towards the statistics of the input, as well as constraints on this adaptation. The two hypotheses differ, however, in that model selection predicts a hard limit on how far listeners' can adapt during initial encounters with an unfamiliar talker.

87 To test these predictions, we revise the standard paradigm used to investigate distributional

88 learning in speech perception. Previous work has employed 'batch testing' designs, in which

89 changes in categorisation responses are assessed only after extended exposure to hundreds of trials

90 or by averaging over extended exposure (e.g., Clayards et al., 2008; Idemaru & Holt, 2011; D. F.

91 Kleinschmidt & Jaeger, 2016; Munson, 2011; Nixon et al., 2016; Theodore & Monto, 2019;

92 **harmon2018?**; **idemaru2021?**). These designs are well-suited to investigate cumulative effects

93 of exposure but are less so to identify constraints on rapidly unfolding incremental adaptation. To

94 be able to detect both incremental and cumulative effects of exposure, within and across exposure

95 conditions, we employed the repeated exposure-test design shown in Figure 2.

96 A secondary aim of the present study was to ameliorate possible concerns about the

97 ecological validity of research on distributional learning. The pioneering works that inspired the
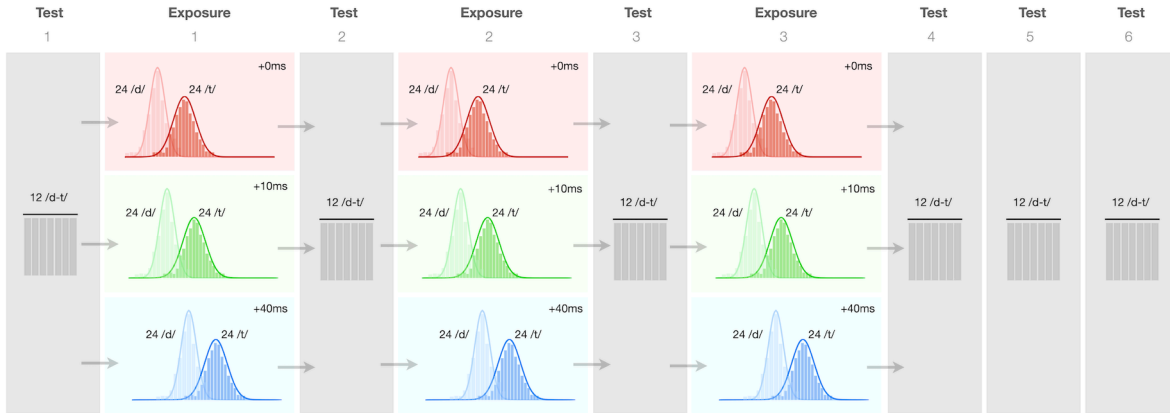
*Figure 2.* Exposure-test design of the experiment. Test blocks presented identical stimuli within and across conditions

present study employed highly unnatural sounding stimuli that were clearly identifiable as robotic speech (Clayards et al., 2008; D. F. Kleinschmidt & Jaeger, 2016). These studies also followed the majority of research on distributional learning in language (e.g., **maye2003?**; **pajak2012?**) and *designed* rather than *sampled* the exposure distributions. As a consequence, exposure distributions in these experiments tend to be symmetrically balanced around the category means—unlike in everyday speech input. Indeed, all of the works we follow here further used categories with *identical* variances (e.g., identical variance along VOT for /b/ and /p/, Clayards et al., 2008; D. F. Kleinschmidt & Jaeger, 2016; or /g/ and /k/, Theodore & Monto, 2019). This, too, is highly atypical for everyday speech input (**chodroff2017structure?**; **lisker-abrahamson1964?**). The present study takes several modest steps to address these issues, with the goal to improve the ecological validity of our stimuli and exposure distributions.

All data and code for this article can be downloaded from XXX. The article is written in R markdown, allowing readers to replicate our analyses with the press of a button using freely available software (R, R Core Team, 2021; RStudio Team, 2020), while changing any of the parameters of our models (see SI, **??**).

## 2   Experiment

The use of test blocks that repeat the same stimuli across blocks and exposure conditions deviates from previous work (Clayards et al., 2008; D. F. Kleinschmidt & Jaeger, 2016; Theodore &

Monto, 2019). This design feature allowed us to assess how increasing exposure affects listeners'
perception without making strong assumptions about the nature of these changes (e.g., linear
changes across trials). We kept test blocks short for two reasons. First, previous work has found
that repeated testing over uniform test continua can reduce or undo the effects of informative
exposure (Liu & Jaeger, 2018, 2019; **cummings202X?**). Second, since we held test stimuli
constant across exposure conditions, the distribution—and thus the relative unexpectedness—of
test stimuli differed to different degrees from the three exposure distributions. By keeping tests
short relative exposure (12 vs. 48 trials), we aimed to minimize the influence of test trials on
adaptation. The final three test blocks were intended to ameliorate the potential risks of this novel
design: in case adaptation remains stable despite repeated testing, those additional test blocks
were meant to provide additional statistical power to detect the effects of cumulative exposure.

## 2.1   Methods

### 2.1.1   Participants

We recruited 126 participants from the Prolific crowdsourcing platform. We used Prolific's
pre-screening to limit the experiment to participants (1) of US nationality, (2) who reported to be
English speaking monolinguals, and (3) had not previously participated in any experiment from
our lab on Prolific. Prior to the start of the experiment, participants had to confirm that they (4)
had spent the first 10 years of their life in the US, (5) were in a quiet place and free from
distractions, and (6) wore in-ear or over-the-ears headphones that cost at least $15. An additional
115 participants loaded the experiment but did not start or complete it.[1]

Participants took an average of 31.6 minutes to complete the experiment (SD = 20
minutes) and were remunerated $8.00/hour. An optional post-experiment survey recorded
participant demographics using NIH prescribed categories, including participant sex (59 = female,
60 = male, 3 = NA), age (mean = NA years; 95% quantiles = 20-62.1 years), race (6 = Black, 31
= White, 85 = NA), and ethnicity (6 = Hispanic, 113 = Non-Hispanic, 3 = NA).

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[1] Unlike in lab-based experiments, for which participants' right to stop the experiment at any point is costly (both
in terms of physical effort and perceived social cost), exercising this right in web-based experiments is essentially
cost free—in particular, if exercised early in the experiment.

141 Participants' responses were collected via Javascript developed by the Human Language

142 Processing Lab at the University of Rochester (**JSEXP?**) and stored via Proliferate developed at,

143 and hosted by, the ALPs lab at Stanford University (**schuster?**).

### 2.1.2 Materials

145 We recorded 8 tokens each of four minimal word pairs (*dill/till, dim/tim, din/tin*, and *dip/tip*)

146 from a 23-year-old, female L1-US English talker from New Hampshire, judged to have a "general

147 American" accent. In addition to these critical minimal pairs we also recorded three words that

148 did not did not contain any stop consonant sounds ("flare", "share", and "rare"). These word

149 recordings were used for catch trials. Stimulus intensity was normalized to 70 dB sound pressure

150 level for all recordings.

151 The critical minimal pair recordings were used to create four VOT continua using a script

152 (Winn, 2020) in Praat (**praat?**). This approach resulted in continuum steps that sound natural

153 (unlike the highly robotic-sounding stimuli employed in Clayards et al., 2008; D. F. Kleinschmidt

154 & Jaeger, 2016). A post-experiment survey asked participants: "*Did you notice anything in*

155 *particular about how the speaker pronounced the different words (e.g. till, dill, etc.)?*" No

156 participant reported that the stimuli sounded unnatural. The procedure also maintained the

157 natural correlations between the most important cues to word-initial stop-voicing in L1-US

158 English (VOT, F0, and vowel duration). Specifically, the F0 at vowel onset of each stimulus was

159 set to respect the linear relation with VOT observed in the original recordings of the talker. The

160 duration of the vowel was set to follow the natural trade-off relation with VOT (Allen & Miller,

161 1999). Further details on the recording and resynthesis procedure are provided in the

162 supplementary information (SI, **??**).

163 The VOTs generated for each continuum ranged from -100 to +130 msec in 5 msec steps.[2]

164 A norming experiment (N = 24 participants) reported in the SI (**??**) was used to select the three

165 minimal pair continua that elicited the most similar categorization responses (*dill-till, din-tin*, and

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[2] We follow previous work (Dave Kleinschmidt, 2020; **OTHERS?**) and refer to prevoicing as negative VOTs though we note that prevoicing is perhaps better conceived of as a separate phonetic feature (for discussion, see **REF?**). In L1-US English, prevoicing occurs on about 20%-48% of word-initial voiced stops and ~0% of voiceless stops (**lisker-abramson1967?**; **smith1978?**).

166  *dip-tip*). These three continua were used to create the exposure conditions shown in Figure 2.

167  **2.1.3  Procedure**

168  At the start of the experiment, participants acknowledged that they met all requirements and

169  provided consent, as per the Research Subjects Review Board of the University of Rochester.

170  Participants also had to pass a headphone test (**REF?**), and were instructed to not change the

171  volume throughout the experiment. Following instructions, participants completed 234

172  two-alternative forced-choice categorisation trials (Figure 3). Participants were instructed that

173  they would hear a female talker say a single word on each trial, and were asked to select which

174  word they heard. Participants were asked to listen carefully and answer as quickly and as

175  accurately as possible. They were also alerted to the fact that the recordings were subtly different

176  and therefore may sound repetitive.
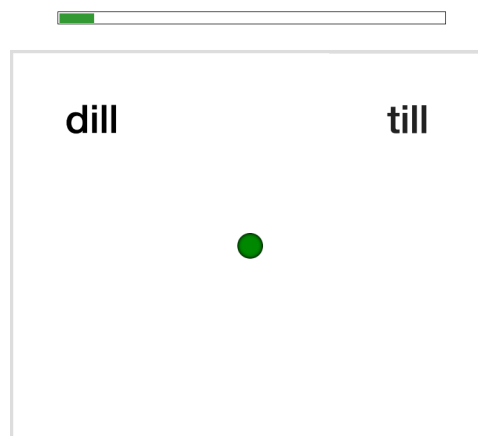


*Figure 3.* Example trial display. When the green button turned bright green, participants had to
click on it to play the recording.


177      Unbeknownst to participants, the 234 trials were split into exposure (54 trials each) and

178  test blocks (12 trials each). Participants were given the opportunity to take breaks after every 60

179  trials, which was always during an exposure block. Finally, participants completed an exit survey

180  and an optional demographics survey.

181      *Test blocks.*   The experiment started with a test block. Test blocks were identical within

182  and across conditions, always including 12 minimal pair trials assessing participants'

categorization at 12 different VOTs (-5, 5, 15, 25, 30, 35, 40, 45, 50, 55, 65, 70 msec). A uniform distribution over VOTs was chosen to maximize the statistical power to determine participants' categorisation function. The assignment of VOTs to minimal pair continua was randomized for each participant, while counter-balancing it within and across test blocks. Each minimal pair appear equally often within each test block (four times), and each minimal pair appear with each VOT equally often (twice) across all six test blocks (and no more than once per test block).

Each trial started with a dark-shaded green fixation dot being displayed. At 500ms from trial onset, two minimal pair words appeared on the screen, as shown in Figure **??**. At 1000ms from trial onset, the fixation dot would turn bright green and participants had to click on the dot to play the recording. This was meant to reduce trial-to-trial correlations by resetting the mouse pointer to the center of the screen at the start of each trial. Participants responded by clicking on the word they heard and the next trial would begin.

*Exposure blocks.* Each exposure block consisted of 24 /d/ and 24 /t/ trials, as well as 6 catch trials that served as a check on participant attention throughout the experiment (2 instances for each of three combinations of the three catch recordings). With a total of 144 trials, exposure was substantially shorter than in similar previous experiments (cf. 228 trials in Clayards et al., 2008; 222 trials in Dave Kleinschmidt, 2020; 2 x 236 trials, Theodore & Monto, 2019; 456 trials, Nixon et al., 2016).

The distribution of VOTs across the 48 /d/-/t/ trials depended on the exposure condition. Specifically, we first created a *baseline* condition. Although not critical to the purpose of the experiment, we aimed for the VOT distribution in this condition to closely resemble participants' prior expectations for a 'typical' female talker of L1-US English (for details, see SI, **??**). The mean and standard deviations for /d/ along VOT were set 5 msecs and 50 msecs, respectively. The mean and standard deviations for /t/ were set 80 msecs and 270 msecs, respectively. To create more realistic VOT distributions, we *sampled* from the intended VOT distribution (top row of Figure 4). This creates distributions that more closely resemble the type of distributional input listeners experience in everyday speech perception, deviating from previous work, which exposed listeners to highly unnatural fully symmetric samples (Clayards et al., 2008; Dave Kleinschmidt, 2020; D. F. Kleinschmidt & Jaeger, 2016).

Half of the /d/ and half of the /t/ trials were labeled, the other half was unlabeled (paralleling one of the conditions in D. Kleinschmidt, Raizada, & Jaeger, 2015). Unlabeled trials were identical to test trials except that the distribution of VOTs across those trials was bimodal (rather than uniform), and determined by the exposure condition.[3] Labeled trials instead presented two response options with identical stop onsets (e.g., *din* and *dill*). This effectively labeled the input as belonging to the intended category (e.g., /d/).

Next, we created the two additional exposure conditions by shifting these VOT distributions by +10 or +40 msecs (see Figure 4). This approach exposes participants to heterogenous approximations of normally distributed VOTs for /d/ and /t/ that varied across blocks, while holding all aspects of the input constant across conditions except for the shift in VOT. The order of trials was randomized within each block and participant, with the constraint that no more than two catch trials would occur in a row. Participants were randomly assigned to one of 3 (exposure condition) x 3 (block order) x 2 (placement of response options) lists.

### 2.1.4 Exclusions

Due to data transfer errors 4 participants' data were not stored and therefore excluded from analysis. We further excluded from analysis participants who committed more than 3 errors out of the 18 catch trials (<83% accuracy, N = 1), participants who committed more than 4 errors out of the 72 labelled trials (<94% accuracy, N = 0), participants with an average reaction time more than three standard deviations from the mean of the by-participant means (N = ), participants who had atypical categorisation functions at the start of the experiment (N = 2, see SI, **??** for details), and participants who reported not to have used headphones (N = 0). This left for analysis 17,136 exposure and 8,568 test observations from 119 participants (94% of total), evenly split across the three exposure conditions.

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[3] Previous studies have estimated changes in participants' categorisation responses by analyzing responses on unlabeled exposure trials (e.g., Clayards et al., 2008; D. F. Kleinschmidt & Jaeger, 2016; Theodore & Monto, 2019). This approach compares responses across different values of acoustic-phonetic cues (since the exposure inputs differed by exposure condition), so that assumptions baked into the analysis approach (e.g., linearity along the acoustic-phonetic continuum) can potentially bias the results. Here we avoid this issue by holding test stimuli constant (see also Dave Kleinschmidt, 2020, Experiment 4).
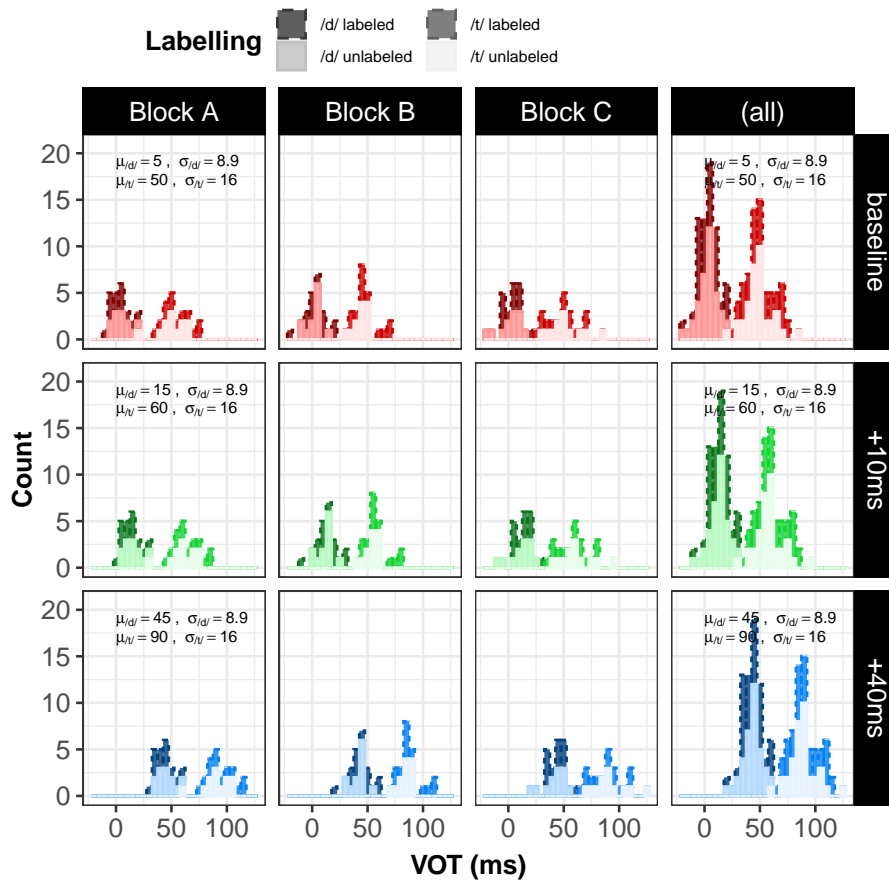
*Figure 4.* Histogram of VOTs across the 48 trials of all three exposure blocks by exposure condition. The order of blocks was counter-balanced across participants.

## 2.2  Results

Figure 5 summarizes the results. Panels A and B show participants' categorisation responses during exposure and test blocks, along with the categorisation function estimated from those responses via a mixed-effects psychometric model that we describe next. These panels facilitate comparison between exposure conditions within each block. Panels C and D show the slope and point of subject equality (PSE)—i.e., the point at which participants are equally likely to respond "d" and "t"—of the categorisation function across blocks and conditions. These panels facilitate comparison across blocks within each exposure condition.

### 2.2.1 Analysis approach

We analyzed participants' categorisation responses during exposure and test blocks in two separate Bayesian mixed-effects psychometric models, fit using brms (Bürkner, 2017) in R (R Core Team, 2021; RStudio Team, 2020, for details, see SI, **??**). These models account for attentional lapses while estimating participants' categorisation functions. Failing to account for attentional lapses—while commonplace in research on speech perception (but see Clayards et al., 2008; D. F. Kleinschmidt & Jaeger, 2016)—can lead to biased estimates of categorization boundaries (Prins, 2012; Wichmann & Hill, 2001). For the present experiment, however, lapse rates were negligible (0.9%, 95%-CI: 0.4 to 1.5%), and all results replicate in simple mixed-effects logistic regressions (Jaeger, 2008).

### 2.2.2 Does exposure affect participants' categorisations (averaging across all blocks)?

Here we focus on the test blocks, which were identical within and across exposure conditions. Analyses of the exposure blocks are reported in the SI (**??**), and replicate all effects found in the test blocks. We first assessed whether the exposure conditions had the expected effects across all test blocks *relative to each other*. Unsurprisingly, participants were more likely to respond "t" the larger the VOT ($\hat{\beta} = 15.68$, $90\%-\text{CI} = [13.149, 18.4]$, $BF = 7999$, $p_{posterior} = 1$). Critically, exposure affects participants' categorisation responses in the expected direction. Marginalizing across all blocks, participants in the +40 condition were less likely to respond "t" than participants in the +10 condition ($\hat{\beta} = -2.43$, $90\%-\text{CI} = [-3.541, -1.363]$, $BF = 443.4$, $p_{posterior} = 0.998$) or the baseline condition ($\hat{\beta} = -3.39$, $90\%-\text{CI} = [-4.969, -1.93]$, $BF = 332.3$, $p_{posterior} = 0.997$). There was also evidence—albeit less decisive—that participants in the +10 condition were less likely to respond "t" than participants in the baseline condition ($\hat{\beta} = -0.97$, $90\%-\text{CI} = [-2.241, 0.298]$, $BF = 9.2$, $p_{posterior} = 0.902$). That is, the +10 and +40 conditions resulted in categorisation functions that were shifted rightwards compared to the baseline condition, as also visible in Figures 5.

270      This replicates previous findings that exposure to changed VOT distributions changes

271  listeners' categorization responses (for /b/-/p/: Clayards et al., 2008; Dave Kleinschmidt, 2020;

272  D. F. Kleinschmidt & Jaeger, 2016; for /g/-/k/, Theodore & Monto, 2019). Having established

273  that exposure affected categorization, we turn to the questions of primary interest.

274  1. How quickly does exposure begin to affect participants' responses? And do they dissipate

275      with repeated testing?

276  2. Are the shifts in categorisation behaviour proportional to the differences between the

277      exposure conditions?

278  3. [model learning vs. selection]

279      [MORE HERE]
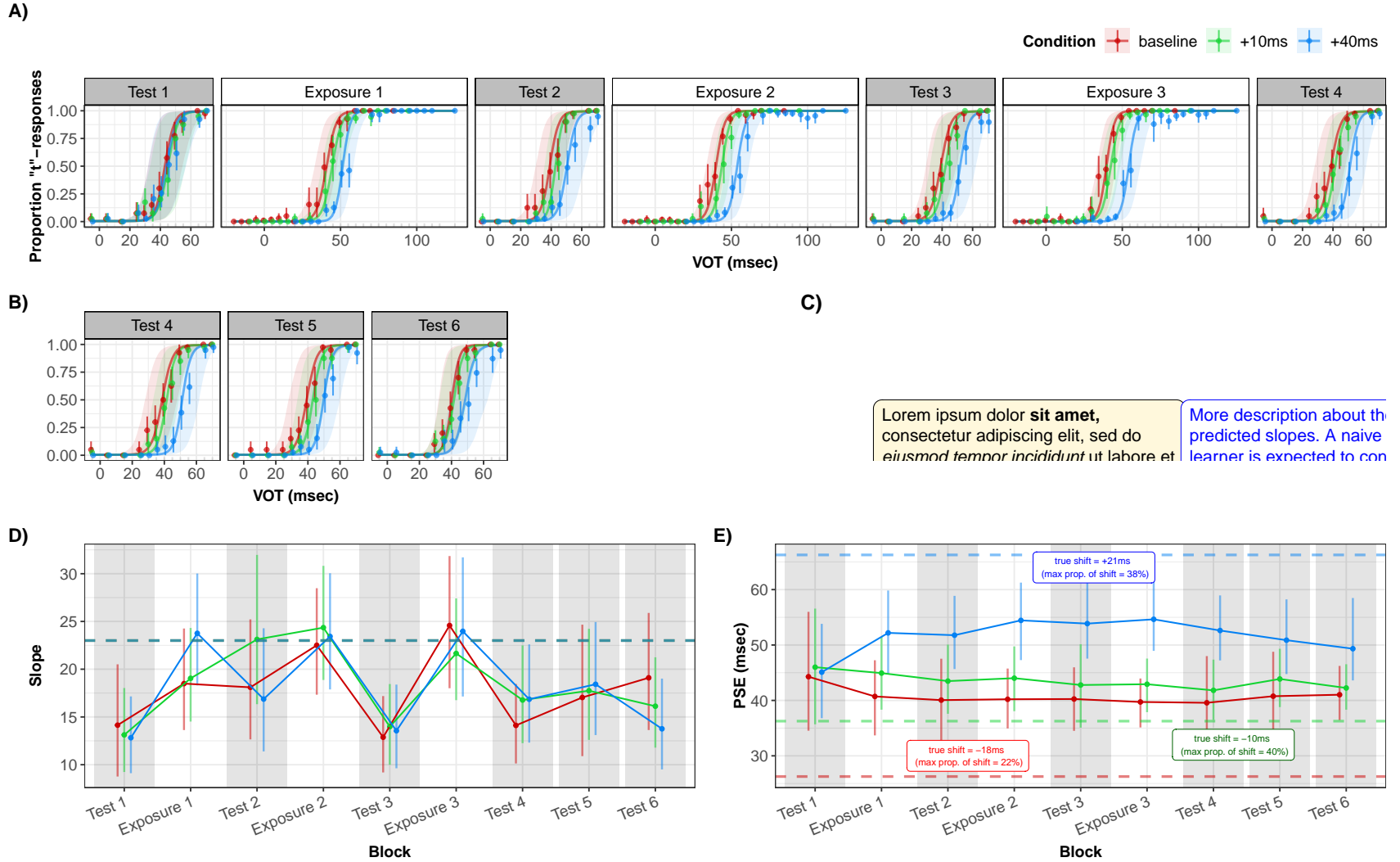
*Figure 5*. Summary of results. **Panel A:** Changes in listeners psychometric categorisation functions as a function of exposure, from Test 1 to Test 4 with all intervening exposure blocks (only unlabelled trials were included in the analysis of exposure blocks since labelled trials provide no information about listeners' categorization function). Point ranges indicate the mean proportion of "t"-responses and their 95% bootstrapped CI. Lines and shaded intervals show the MAP predictions and 95% posterior CIs of a Bayesian mixed-effects psychometric model fit to participants' responses. **Panel B:** Same as Panel A but for the final three test blocks without intervening exposure. Test 4 is shown as part of both Panels A and B. **Panels C & D:** Changes across blocks in the slope and boundary (point-of-subjective-equality, PSE) of the categorisation functions shown in Panels A-B. Point ranges represent the posterior means and their 95% CI. Dashed reference lines show the intercepts and PSEs that naive (non-rational) learner would be expected to converge against after sufficient exposure (an ideal observer model that knows the exposure distributions).

280  Next, we compared the effects of exposure conditions for all test blocks with intervening

281  exposure. These effects can be compared from three complementing perspectives. First, we can

282  compare how exposure affects listeners' categorisation responses relative to other exposure

283  conditions. Second, we can compare how exposure affects listeners' categorisation responses

284  within each condition relative to listeners' responses prior to any exposure. Third, we can

285  compare changes in listeners' responses to those expected from an ideal observer that has fully

286  learned the exposure distributions.

### 2.2.3  Comparing across exposure conditions: How quickly does exposure begin to affect participants' responses?

289  Figure 5A suggests that differences between exposure conditions emerged early in the experiment:

290  already in Test 2, listener's categorisation functions seem to be shifted rightwards (larger PSEs)

291  in the +40 condition compared to the +10 condition, and in the +10 condition compared to the

292  baseline condition. This is confirmed by the Bayesian hypothesis tests summarized in Table 1.

293  Prior to any exposure, during Test 1, participants' responses did not differ across exposure

294  condition. After exposure to only 24 /d/ and 24 /t/ stimuli, during Test 2, participants'

295  responses differed between exposure conditions. The difference between the +40 condition and

296  the +10 or baseline condition kept increasing with exposure up to Test 4. Additional hypothesis

297  tests in Table 2 show that the change from Test 1 to 2 was largest (BF = 27.8), followed by the

298  change from Test 2 to 3 (BF = 19.2), with only minimal changes from Test 3 to 4 (BF = 1.7).

299  Qualitatively paralleling the changes across blocks for the +40 condition, the change in the

300  difference between the +10 and baseline conditions was largest from Test 1 to 2 (BF = 13.5), and

301  then somewhat decreased from Test 2 to Test 4 (BFs < 4).

302  Tables 1 and 2 also summarize the consequences of repeated testing. The difference in the

303  PSE decreased from Test 4 to 6, both for the +40 compared to the +10 condition

304  ($\hat{\beta} = 1.98, \ 90\%-\text{CI} = [-0.418, 4.338], \ BF = 12.2, \ p_{posterior} = 0.924$) and the +10 compared to

305  the baseline condition ($\hat{\beta} = 0.93, \ 90\%-\text{CI} = [-0.921, 2.908], \ BF = 4.3, \ p_{posterior} = 0.811$, see

306  also Figure 5B & D). On the final test block, the +10 condition did not differ any longer from the

307  baseline condition. Only the differences between the +40 condition relative to the +10 and

308 baseline conditions persisted, albeit substantially reduced compared to Test 4.

309      This replicates previous findings that repeated testing over uniform test continua can undo

310 the effects of exposure (Cummings & Theodore, 2023; Liu & Jaeger, 2018, 2019), and extends

311 them from perceptual recalibration paradigms to distributional learning paradigms (see also Dave

312 Kleinschmidt, 2020). One important methodological consequence of these findings is that longer

313 test phases do not necessarily increase the statistical power to detect effects of adaptation (unless

314 analyses take the effects of repeated testing into account, as in the approach developed in Liu &

315 Jaeger, 2018). Analyses that average across all test tokens—as remains the norm—are bound to

316 systematically underestimate the adaptivity of human speech perception.

### 2.2.4  Comparing within exposure conditions: How quickly does exposure begin to

317

318       affect participants' responses?

319 Second, we can compare how exposure affects listeners' categorisation reponses within each

320 condition relative to listeners' responses prior to any exposure. These changes are summarised for

321 the slope and PSE in Figure 5C & D, respectively. This visualization makes apparent an aspect of

322 participants' behavior that were not readily apparent in the statistical comparisons we have

323 summarized so far: while the PSEs for the +40 and +10 conditions were shifted rightwards

324 compared to the baseline condition, both the +10 and the baseline condition actually shift

325 leftwards relative to their pre-exposure starting point in Test 1. To understand this pattern, it is

326 helpful to relate our exposure conditions to the distribution of VOT in listeners' prior experience.

327 Figure 6 shows the mean and covariance of our exposure conditions relative to the distribution of

328 VOT by talkers of L1-US English (based on Chodroff & Wilson, 2017). This comparison offers an

329 explanation as to why the baseline condition (and to some extent the +10 condition) shift

330 leftwards with increasing exposure, whereas the +40 condition shifts rightwards: relative to

331 listeners' prior experience our baseline condition actually presented lower-than-expected category

332 means; of our three exposure conditions, only the +40 condition presented larger-than-expected

333 category means. That is, once we take into account how our exposure conditions relate to

334 listeners' prior experience, both the direction of changes from Test 1 to 4 *within* each exposure

335 condition, and the direction of differences *between* exposure conditions receive an explanation.

### 2.2.5  Constraints on cumulative changes

Most relevant to the purpose of the present study, Figure 5D also suggests that there are strong constraints on listeners' ability to adapt. To some extent, this already follows from the analysis of how the differences between exposure conditions change with exposure.

XXX

Second, the reason for the slight decrease in the difference between the +10 and baseline conditions observed in Tables **??** and **??** (visible in Figure 5D as the decreasing difference between the green and red line) is *not* due to a reversal of the effects in the +10 condition. Rather, both conditions are changing in the same direction but the baseline condition stops changing after Test 2, which brings the +10 condition increasingly closer to the baseline condition.
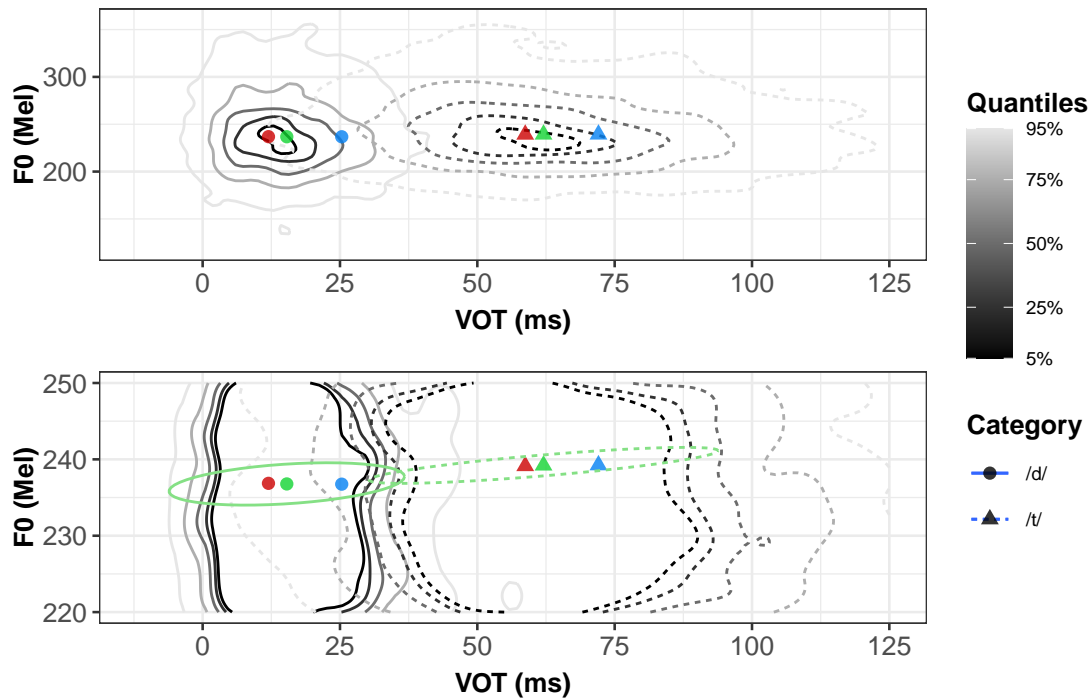


*Figure 6*. Placement of exposure stimuli relative to an estimate of typical phonetic distributions for XXX word-initial /d/ and /t/ productions in L1-US English (based on 92 talkers in Chodroff & Wilson, 2017). The outermost contour of each category shows the 95% density quantile. Points show the category means of the exposure condition, the green ellipsis shows the covariance of the +10 exposure condition (covariance was identical across conditions).

Table 1

*When did exposure begin to affect participants' categorization responses? When, if ever, were these changes undone with repeated testing? This table summarizes the simple effects of the exposure conditions for each test block.*

| Hypothesis | Estimate | SE | 90%-CI | BF | $p_{posterior}$ |
|---|---|---|---|---|---|
| **Test block 1 (pre-exposure)** | | | | | |
| +10 vs. baseline | -0.39 | 0.94 | [-2.096, 1.403] | 1.99 | 0.66 |
| +40 vs. +10 | 0.20 | 0.86 | [-1.359, 1.849] | 0.68 | 0.40 |
| +40 vs. baseline | -0.19 | 1.11 | [-2.377, 2.041] | 1.32 | 0.57 |
| **Test block 2** | | | | | |
| +10 vs. baseline | -2.12 | 1.12 | [-4.334, -0.109] | 22.12 | 0.96 |
| +40 vs. +10 | -2.10 | 1.21 | [-4.333, 0.071] | 17.35 | 0.95 |
| +40 vs. baseline | -4.22 | 1.47 | [-7.048, -1.624] | 80.63 | 0.99 |
| **Test block 3** | | | | | |
| +10 vs. baseline | -0.88 | 0.69 | [-2.244, 0.417] | 7.98 | 0.89 |
| +40 vs. +10 | -3.26 | 0.96 | [-5.164, -1.624] | 169.21 | 0.99 |
| +40 vs. baseline | -4.15 | 1.11 | [-6.371, -2.226] | 162.26 | 0.99 |
| **Test block 4** | | | | | |
| +10 vs. baseline | -1.08 | 0.99 | [-3.017, 0.947] | 5.46 | 0.84 |
| +40 vs. +10 | -4.02 | 1.09 | [-6.043, -2.284] | 420.05 | 1.00 |
| +40 vs. baseline | -5.10 | 1.43 | [-7.839, -2.542] | 132.33 | 0.99 |
| **Test block 5 (no additional exposure)** | | | | | |
| +10 vs. baseline | -1.50 | 0.86 | [-3.08, 0.086] | 16.24 | 0.94 |
| +40 vs. +10 | -2.98 | 1.08 | [-5.01, -1.205] | 130.15 | 0.99 |
| +40 vs. baseline | -4.10 | 1.52 | [-6.811, -1.436] | 73.77 | 0.99 |
| **Test block 6 (no additional exposure)** | | | | | |
| +10 vs. baseline | -0.14 | 0.88 | [-1.829, 1.456] | 1.28 | 0.56 |
| +40 vs. +10 | -2.03 | 0.91 | [-3.852, -0.396] | 34.71 | 0.97 |
| +40 vs. baseline | -3.15 | 1.39 | [-5.754, -0.515] | 31.39 | 0.97 |

## 3  General discussion

- discuss consequences of findings for other accounts (decision-making; normalization)

- discuss fact that test stimuli deviate from exposure stimuli to different extent. on the one hand, it's just 1/4 of all trials. on the other hand, we do see relatively systematic changes in slopes each time we test. so there is evidence that even these 12 trials can affect categorisation slopes (though it is worth keeping in mind that this is a comparison across different sets of stimuli). could this explain shrinkage? unlikely since it wasn't the case in kleinschmidt and jaeger. could it explain the constraint on adaptation? that's less clear. we

Table 2

*Was there incremental change from test block 1 to 4? Did these changes dissipate with repeated testing from block 4 to 6? This table summarizes the interactions between exposure condition and block, whether the differences between exposure conditions changed from test block to test block.*

| Hypothesis | Estimate | SE | 90%-CI | BF | $p_{posterior}$ |
|---|---|---|---|---|---|
| **Difference in +10 vs. baseline** | | | | | |
| Block 1 to 2: increased $\Delta_{PSE}$ | -1.40 | 0.92 | [-3.065, 0.199] | 13.52 | 0.93 |
| Block 2 to 3: increased $\Delta_{PSE}$ | 0.85 | 0.98 | [-1.113, 2.775] | 0.25 | 0.20 |
| Block 3 to 4: increased $\Delta_{PSE}$ | -0.01 | 0.92 | [-1.838, 1.885] | 1.02 | 0.50 |
| *Block 1 to 4: increased $\Delta_{PSE}$* | -0.58 | 1.54 | [-3.652, 2.483] | 1.82 | 0.64 |
| Block 4 to 5: decreased $\Delta_{PSE}$ | -0.38 | 0.71 | [-1.734, 1.091] | 0.42 | 0.30 |
| Block 5 to 6: decreased $\Delta_{PSE}$ | 1.25 | 0.77 | [-0.143, 2.723] | 13.95 | 0.93 |
| *Block 4 to 6: decreased $\Delta_{PSE}$* | 0.86 | 0.97 | [-0.921, 2.908] | 4.30 | 0.81 |
| **Difference in +40 vs. +10** | | | | | |
| Block 1 to 2: increased $\Delta_{PSE}$ | -2.05 | 1.03 | [-3.89, -0.231] | 27.78 | 0.96 |
| Block 2 to 3: increased $\Delta_{PSE}$ | -1.79 | 1.06 | [-3.688, -0.001] | 19.15 | 0.95 |
| Block 3 to 4: increased $\Delta_{PSE}$ | -0.39 | 1.18 | [-2.629, 1.624] | 1.70 | 0.63 |
| *Block 1 to 4: increased $\Delta_{PSE}$* | -4.28 | 1.53 | [-7.158, -1.722] | 101.56 | 0.99 |
| Block 4 to 5: decreased $\Delta_{PSE}$ | 1.41 | 1.07 | [-0.541, 3.319] | 8.66 | 0.90 |
| Block 5 to 6: decreased $\Delta_{PSE}$ | 0.60 | 0.94 | [-1.271, 2.311] | 2.79 | 0.74 |
| *Block 4 to 6: decreased $\Delta_{PSE}$* | 1.99 | 1.25 | [-0.418, 4.338] | 12.24 | 0.92 |

354     can, however, compare the relative mean of exposure and test.

355  • could some form of moving window with historical decay explain the findings? On the one

356     hand if the moving window is very small, that would not explain why we do see some

357     *cumulative* changes across blocks (window must be at least $48 + 12 = 60$ trials). on the

358     other hand, the qualitative changes in the PSEs and slopes suggest that 12 trials can be

359     enough to change some aspects of the categorisation function. it's thus *possible* that

360     something that ways recent input much more strongly but also considers less recent input

361     beyond 48 trials might explain the overall pattern.

362  • discuss potential that observed adaptation maximizes accuracy under the choice rule. use

363     psychometric function fit during unlabeled exposure trials to calculate *accuracy* (not

364     likelihood) on labeled trials under criterion and under proportional matching decision rules.

365     compare against accuracy if ideal observers categorization functions are used instead.

Table 3

*Effects of repeated testing (test blocks 4 to 6)*

| Hypothesis | Estimate | SE | 90%-CI | BF | $p_{posterior}$ |
|---|---|---|---|---|---|
| **Difference in +10 vs. baseline** | | | | | |
| Predicted change from test block 4 to 5 | -0.38 | 0.71 | [-1.734, 1.091] | 0.42 | 0.30 |
| Predicted change from test block 5 to 6 | 1.25 | 0.77 | [-0.143, 2.723] | 13.95 | 0.93 |
| Predicted total change from test block 4 to 6 | 0.86 | 0.97 | [-0.921, 2.908] | 4.30 | 0.81 |
| **Difference in +40 vs. +10** | | | | | |
| Predicted change from test block 4 to 5 | 1.41 | 1.07 | [-0.541, 3.319] | 8.66 | 0.90 |
| Predicted change from test block 5 to 6 | 0.60 | 0.94 | [-1.271, 2.311] | 2.79 | 0.74 |
| Predicted total change from test block 4 to 6 | 1.99 | 1.25 | [-0.418, 4.338] | 12.24 | 0.92 |

## 3.1 Methodological advances that can move the field forward

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Table 4

*Is the shift in +40 from baseline proportional to the magnitude of shift in the exposure distribution?*

| Hypothesis | Estimate | SE | 90%-CI | BF | $p_{posterior}$ |
|---|---|---|---|---|---|
| **Test block 2** | | | | | |
| +40 vs. baseline < 4x +10 vs. baseline | 4.24 | 4.0 | [-2.563, 11.755] | 6.00 | 0.86 |
| **Test block 3** | | | | | |
| +40 vs. baseline < 4x +10 vs. baseline | -0.66 | 2.6 | [-5.316, 4.068] | 0.66 | 0.40 |
| **Test block 4** | | | | | |
| +40 vs. baseline < 4x +10 vs. baseline | -0.84 | 3.4 | [-7.448, 5.528] | 0.67 | 0.40 |

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