

Adaptive Prediction for Functional Time Series

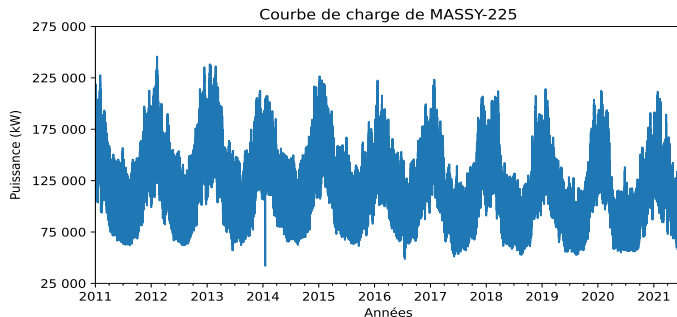
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Introduction (1/3)

Example of a connection point for the extraction and injection of electricity

- ▶ A set of N time-dependent curves, $X_n : [0, 1] \rightarrow \mathbb{R}$, $n = 1 \dots N$.



- ▶ The trajectories are **irregular**.
- ▶ We observe each curve with **measurement errors**.
- ▶ **Regularity** and **final goal** should be considered in reconstruction.

Introduction (2/3)

Observation scheme

For $n = 1, \dots, N$, X_n is measured with error at discrete, randomly sampled points :

$$Y_{n,k} = X_n(T_{n,k}) + \sigma(T_{n,k})\varepsilon_{n,k}, \quad 1 \leq k \leq M_n,$$

- ▶ $\{X_n\}$ is a stationary process of $\mathcal{H} = \mathbb{L}^2[0, 1]$,
- ▶ $M_1, \dots, M_N \stackrel{i.i.d.}{\sim} M$ with expectation λ ,
- ▶ the observation times $T_{n,k} \sim T$ are i.i.d.,
- ▶ $\varepsilon_{n,k} \sim \epsilon$ are independent centred errors with unit variance,
- ▶ $\{X_n\}$, $\{M_n\}$, $\{\varepsilon_{n,k}\}$, and $\{T_{n,k}\}$ are mutually independent.

Introduction (3/3)

Motivation

We aim to build a procedure for **curve prediction** that adapts to the **local regularity** of the trajectories for **FTS** in the context of **weak dependence**.

Using dependent curves measured with noise at random discrete points, our goal is to perform **adaptive estimation** of :

- ▶ the Best Linear Unbiased Predictor (BLUP) that is a combination of
 - ▶ mean, covariance and autocovariance functions.
-
- ▶ For FTS, a functional data recovery has already been considered by RUBÌN AND PANARETOS (2020) under the hypothesis that these functions admit at least one derivative.
 - ▶ For irregular curves, MAISSORO ET AL. (2024) proposed new estimators of the mean and autocovariance functions.

Outline

1 Introduction

2 Adaptive linear predictor

- Definition of the BLUP
- Estimation of the BLUP
- Application

3 Take home message

Adaptive linear predictor (1/6)

Let $\mu(t) = \mathbb{E}(X_n(t))$ and $\Gamma_\ell(s, t) = \mathbb{E} \{ [X_0(s) - \mu(s)][X_\ell(t) - \mu(t)] \}$, for all $s, t \in I$ and $\ell \geq 0$.
Moreover,

$$\mathbb{Y}_n = (Y_{n,1}, \dots, Y_{n,M_n})^\top, \quad \mathcal{Y}_{n_0,1} = (\mathbb{Y}_{n_0-1}^\top, \mathbb{Y}_{n_0}^\top)^\top, \quad \Sigma_n = \text{diag}(\sigma^2(T_{n,1}), \dots, \sigma^2(T_{n,M_n})),$$

$$\mathcal{M}_{n_0,1} = (\mu(T_{n_0-1,1}), \dots, \mu(T_{n_0-1,M_{n_0-1}}), \mu(T_{n_0,1}), \dots, \mu(T_{n_0,M_{n_0}}))^\top.$$

Mettre des pauses Definition. Let $t_0 \in I$ and $n_0 \in \{1, \dots, N\}$ be fixed. Following ROBINSON (1991), the BLUP of $X_{n_0}(t_0)$ given $\mathcal{Y}_{n_0,1}$ is :

$$\hat{X}_{n_0}(t_0) = \hat{\mu}(t_0) + \hat{B}_{n_0,1}^\top (\mathcal{Y}_{n_0,1} - \hat{\mathcal{M}}_{n_0,1}),$$

$$\text{where } B_{n_0,1} = \begin{pmatrix} G_0^{(n_0-1, n_0-1)} + \Sigma_{n_0-1} & G_1^{(n_0-1, n_0)} \\ G_1^{(n_0, n_0-1)} & G_0^{(n_0, n_0)} + \Sigma_{n_0} \end{pmatrix}^{-1} \begin{pmatrix} \Gamma_1(T_{n_0-1,1}, t_0) \\ \vdots \\ \Gamma_1(T_{n_0-1, M_{n_0-1}}, t_0) \\ \Gamma_0(T_{n_0,1}, t_0) \\ \vdots \\ \Gamma_0(T_{n_0, M_{n_0}}, t_0) \end{pmatrix},$$

and $G_\ell^{(n, n')} = (\Gamma_\ell(T_{n,i}, T_{n',j}))_{1 \leq i \leq M_n, 1 \leq j \leq M_{n'}}$.

Estimation. Put a hat on to get an estimate...

Adaptive linear predictor (2/6)

Local Regularity Parameters

Definition. The process X admits a *local regularity* at $t \in I$, with *local exponent* $H_t \in (0, 1)$ and *Hölder constant* $L_t > 0$, if

$$\mathbb{E} [(X(u) - X(v))^2] \approx L_t^2 |u - v|^{2H_t},$$

for all u, v satisfying $t - \Delta/2 \leq u \leq t \leq v \leq t + \Delta/2$ for some $\Delta > 0$. Donner la philo

Adaptive linear predictor (2/6)

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Estimation. We use some nonparametric estimates \tilde{X}_n to recover the X_n 's. For any u, v close to t , let

$$\hat{\theta}(u, v) = \frac{1}{N} \sum_{n=1}^N \left\{ \tilde{X}_n(v) - \tilde{X}_n(u) \right\}^2.$$

Our estimators of H_t and L_t^2 are defined as empirical counterparts of their respective definition. Let $t_1 = t - \Delta/2$, $t_3 = t + \Delta/2$. The estimators of H_t and L_t^2 are

$$\hat{H}_t = \frac{\log(\hat{\theta}(t_1, t_3)) - \log(\hat{\theta}(t_1, t))}{2 \log(2)} \quad \text{and} \quad \hat{L}_t^2 = \frac{\hat{\theta}(t_1, t_3)}{\Delta^{2\hat{H}_t}}.$$

Concentration bounds. Under \mathbb{L}_C^p – m-approximability by MAISSORO ET AL. (2024).

Adaptive linear predictor (3/6)

Weak dependency assumption

Let $\{X_n\}_{n \in \mathbb{Z}}$ be a stationary FTS, with **continuous paths**, on $I = [0, 1]$:

- ▶ $(\mathcal{H}, \langle \cdot, \cdot \rangle_{\mathcal{H}})$: space of square integrable functions ;
- ▶ $(\mathcal{C}, \|\cdot\|_{\infty})$: space of continuous functions on I .

The space $\mathbb{L}_{\mathcal{C}}^p$ is the space of \mathcal{C} -valued random element X such that

$$\nu_p(\|X\|_{\infty}^p) = (\mathbb{E}[\|X\|_{\infty}^p])^{1/p} < \infty.$$

Weak dependency assumption : $\{X_n\}_n$ is $\mathbb{L}_{\mathcal{C}}^p$ – **m-approximable**.

\mathbb{L}^p – **m-approximation** for \mathcal{H} -valued functional data was introduced by HÖRMANN and KOKOSZKA (2010).

Example. $FAR(1)$ is $\mathbb{L}_{\mathcal{C}}^p$ – *m-approximable*.

Adaptive linear predictor (4/6)

Adaptive mean estimation. Let $\mu(t) = \mathbb{E}(X_n(t))$.

- ▶ A naive estimator of $\mu(t)$ is an average of nonparametric estimator of the curves.
- ▶ **The objective** : estimation of $\mu(t)$ by adaptive smoothing of X_n .
- ▶ The proposed estimator is $\hat{\mu}_N(t; h_\mu^*)$, with

$$\hat{\mu}_N(t; h) = \sum_{n=1}^N \frac{\pi_n(t; h)}{P_N(t; h)} \hat{X}_n(t; h), \quad \text{where}$$

- $\pi_n(t; h) = 1$ if there is at least one $T_{n,i} \in [t - h, t + h]$ and 0 otherwise,
 - $P_N(t; h) = \sum_{n=1}^N \pi_n(t; h)$,
 - and $\hat{X}_n(t; h)$ is Nadaraya-Watson estimator with bandwidth h .
- ▶ h_μ^* minimises a sharp bound of the quadratic risk of $\hat{\mu}_N(t; h)$.

Adaptive linear predictor (5/6)

Adaptive mean estimation. More precisely, we consider

$$\mathbb{E}_{M,T} \left[(\hat{\mu}_N(t; h) - \mu(t))^2 \right] \leq 2R_\mu(t; h), \quad \text{where}$$

$$R_\mu(t; h) = L_t^2 h^{2H_t} \mathbb{B}(t; h, 2H_t) + \sigma^2(t) \mathbb{V}_\mu(t; h) + \mathbb{D}_\mu(t; h) / P_N(t; h),$$

and define $h_\mu^* \in \arg \min_{h \in \mathcal{H}_N} \hat{R}_\mu(t; h)$ with $\hat{R}_\mu(t; h) = R_\mu(t; h, \hat{H}_t, \hat{L}_t^2, \hat{\sigma}^2(t))$.

Let $t \in I$. Under some assumptions we have

$$\begin{aligned} \hat{R}_\mu(t; h) &= \mathcal{O}_{\mathbb{P}} \left\{ h^{2H_t} + (N\lambda h)^{-1} + N^{-1} \right\}, \\ h_\mu^* &= \mathcal{O}_{\mathbb{P}} \left\{ (N\lambda)^{-\frac{1}{1+2H_t}} \right\}, \end{aligned}$$

and the estimator $\hat{\mu}_N(t; h_\mu^*)$ satisfies

$$\hat{\mu}_N^*(t) - \mu(t) = \mathcal{O}_{\mathbb{P}} \left\{ (N\lambda)^{-\frac{H_t}{1+2H_t}} + N^{-1/2} \right\}.$$

Adaptive linear predictor (6/6)

Adaptive lag— $\ell(\ell > 0)$ **autocovariance estimation**. As for the mean,

- ▶ the 'first smooth, then estimate' estimator is $\widehat{\Gamma}_{N,\ell}(s, t; h_s^*, h_t^*)$, where
- ▶ (h_s^*, h_t^*) minimises a sharp bound of the quadratic risk of $\widehat{\Gamma}_{N,\ell}(s, t; h_s, h_t)$.

Let $s, t \in I$. Under some assumptions we have

$$h_s^* = \mathcal{O}_{\mathbb{P}} \left(\max \left\{ (N\lambda^2)^{-\frac{H_t}{H_s\{2H_t+1\}+H_t}}, (N\lambda)^{-\frac{1}{2H_s+1}} \right\} \right),$$

$$h_t^* = \mathcal{O}_{\mathbb{P}} \left(\max \left\{ (N\lambda^2)^{-\frac{H_s}{H_s\{2H_t+1\}+H_t}}, (N\lambda)^{-\frac{1}{2H_t+1}} \right\} \right),$$

and the estimator $\Gamma_{N,\ell}^*(s, t) = \widehat{\Gamma}_{N,\ell}(s, t; h_s^*, h_t^*)$ satisfies

$$\widehat{\Gamma}_{N,\ell}^*(s, t) - \Gamma_{\ell}(s, t) = \mathcal{O}_{\mathbb{P}} \left((N\lambda^2)^{-\frac{H_s H_t}{H_s\{2H_t+1\}+H_t}} + (N\lambda)^{-\frac{H_s}{2H_s+1}} + (N\lambda)^{-\frac{H_t}{2H_t+1}} + N^{-1/2} \right).$$

Better rates than those using a single bandwidth, as in GOLOVKINE ET AL. (2021).

Application (1/5)

We simulate a FAR(1) where the WN are i.i.d. *multifractional Brownian motion* (see STOEV and TAQQU (2006)) paths with :

- ▶ a logistic H_t function and $L_t^2 = 1$, **Descriptif à la de "Figure"**
- ▶ a kernel $\Psi(s, t)$ estimated from voltage curves (see HEBRAIL and BERARD (2012)).

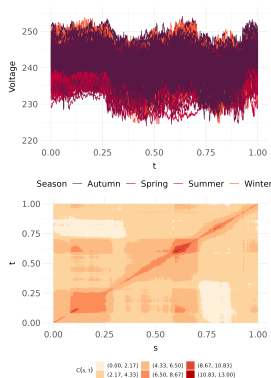


Figure – Cov

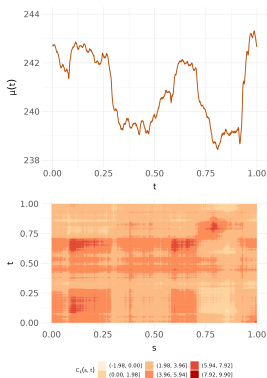


Figure – lag-1 Autocov

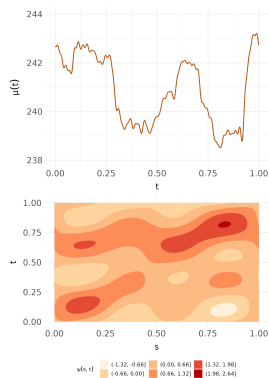
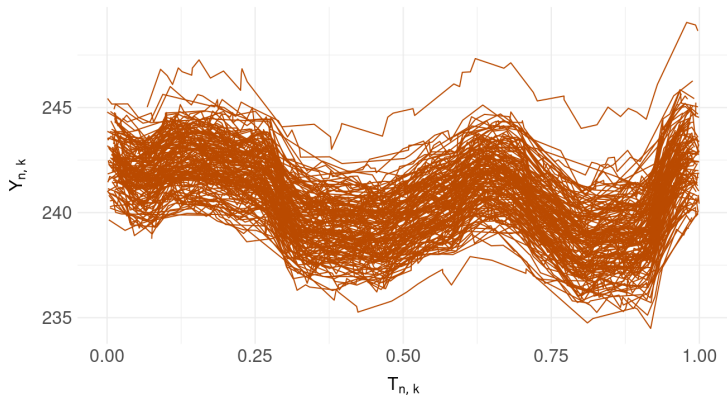


Figure – Kernel $\Psi(s, t)$

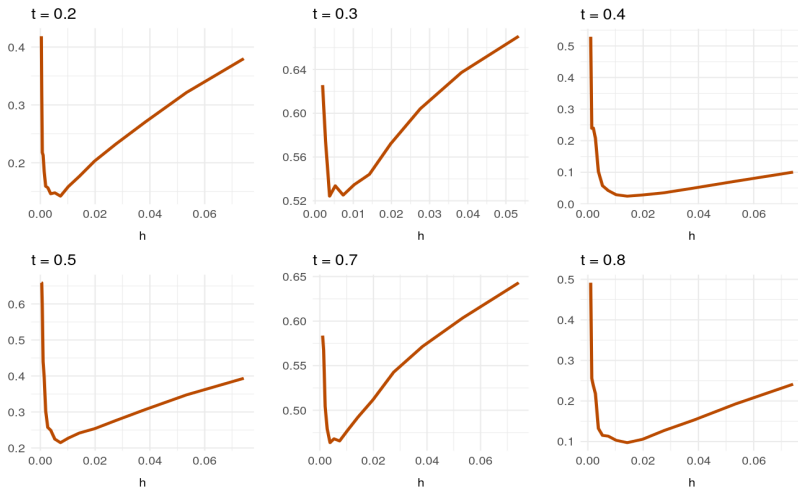
Application (2/5)

Generate curves for $N = 150$ and $\lambda = 40$ with additional normal noise of standard deviation 0.25.



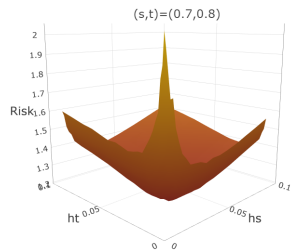
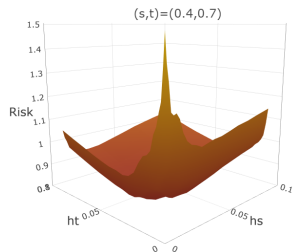
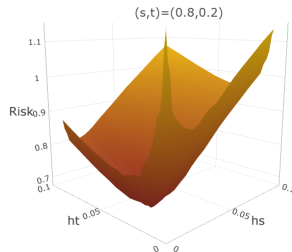
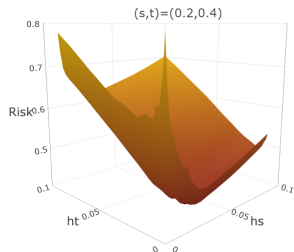
Applications (3/5)

Adaptive mean function estimation. $\hat{R}_\mu(t; h)$ at some locations :



Application (4/5)

Adaptive mean function estimation. $\hat{R}_r(s, t; h_s, h_t)$ at some locations :

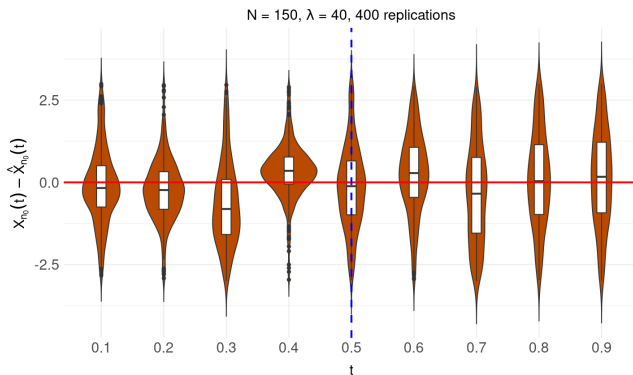


Applications (5/5)

Application : Adaptive BLUP estimation.

Generate a FTS with $N = 150$ and $\lambda = 40$ and regenerate the 150-th curve 400 times

- ▶ Estimate the BLUP using only points observed before 0.5.
- ▶ Obtaining satisfactory results.



Take home message

Adaptive predictor which combines

- 1 The best Linear Unbiased Predictor (BLUP) estimator.
- 2 The estimation of local regularity parameters for FTS.
- 3 The adaptive optimal estimates of mean, covariance and autocovariance.

Work in progress...

- ▶ Advanced empirical study on BLUP,
- ▶ Establishing confidence bands for the estimates, etc.

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Thanks for your attention !