Math 475: Final Exam

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Problem 1: The Peppered Moths

The goal of this problem is to find and create and EM Algorithm for looking at Peppered moths species in an area/from a sample population. ### Step 1 (E): Find the Q function

$$Q(\theta|\theta^t) = E[logL(\theta|Y)|x, \theta^t]$$

Following the notation, we first have our observed data x. $x = (n_c, n_I, n_T)$ where n is the number of moths based on one of three phenotype (C, I, T). The complete data $y = (n_{CC}, n_{CI}, n_{CT}, n_{II}, n_{IT}, n_{TT})$ where y represents the number of moths based on allele frequency (genotype). For this problem, we wish to estimate the allele probabilities: $p = (p_C, p_I, p_T)$. Because the allele T is recessive to I and I is recessive to C, p only needs to depend on allele counts for C and I. This means $p_T = 1 - p_C - p_I$ and $p = (p_C, p_I, 1)$. From there we can compute the complete log likelihood:

 $log f_Y(y|p) = n_{CC} log(p_C^2) + n_{CI} log(2p_C p_I) + n_{CT} log(2p_C p_T) + n_{II} log(p_I^2) + n_{IT} log(2p_I p_T) + n_{TT} log(p_T^2) +$

$$+log {n \atop n_{CC}} n_{CI} \quad n_{CT} \quad n_{II} \quad n_{IT} \quad n_{TT}$$

The only complete data variable entirely observed is n_{TT} because it is a recessive allele. Therefore the complete data becomes $Y = (N_{CC}, N_{CI}, N_{CT}, N_{II}, N_{IT}, n_{TT})$. Using this, we can calculate the expectation for each variable in the complete data set Y:

$$\begin{split} E[N_{CC}|n_C,n_I,n_T,p^{(t)}] &= n_{CC}^{(t)} = \frac{n_C(p_C^{(t)})^2}{(p_C^{(t)})^2 + 2p_C^{(t)}p_I^{(t)} + 2p_C^{(t)}p_T^{(t)}} \\ E[N_{CC}|n_C,n_I,n_T,p^{(t)}] &= n_{CI}^{(t)} = \frac{2n_C p_C^{(t)}p_I^{(t)}}{(p_C^{(t)})^2 + 2p_C^{(t)}p_I^{(t)} + 2p_C^{(t)}p_T^{(t)}} \\ E[N_{CC}|n_C,n_I,n_T,p^{(t)}] &= n_{CT}^{(t)} = \frac{2n_C p_C^{(t)}p_I^{(t)} + 2p_C^{(t)}p_T^{(t)}}{(p_C^{(t)})^2 + 2p_C^{(t)}p_I^{(t)} + 2p_C^{(t)}p_T^{(t)}} \\ E[N_{CC}|n_C,n_I,n_T,p^{(t)}] &= n_{II}^{(t)} = \frac{n_I(p_I^{(t)})^2}{(p_I^{(t)})^2 + 2p_I^{(t)}p_T^{(t)}} \\ E[N_{CC}|n_C,n_I,n_T,p^{(t)}] &= n_{IT}^{(t)} = \frac{2n_I p_I^{(t)}p_T^{(t)}}{(p_I^{(t)})^2 + 2p_I^{(t)}p_T^{(t)}} \end{split}$$

We then get a Q function like the following:

 $Q(p|p^{(t)}) = n_{CC}log(p_C^2) + n_{CI}log(2p_Cp_I) + n_{CT}log(2p_Cp_T) + n_{II}log(p_I^2) + n_{IT}log(2p_Ip_T) + n_{TT}log(p_T^2) + k(n_C, n_I, n_T, p^{(t)})$ where k is a conditional expectation. This will become unimportant in the M step because it does not depend on p.

Step 2 (M): Maximize the Q function and set next variable to equal maximizer.

To Maximize the Q function, we will take the derivative of the Q function. Because the Q function is multinomial, we will need to take two separate derivatives in regard to p_C and p_I .

$$\begin{split} \frac{dQ(p|p^{(t)})}{dp_C} &= \frac{2n_{CC}^{(t)} + n_{CI}^{(t)} + n_{CT}^{(t)}}{p_C} - \frac{2n_{TT}^{(t)} + n_{CT}^{(t)} + n_{IT}^{(t)}}{1 - p_C - p_I} \\ & p_C^{(t)} = \frac{2n_{CC}^{(t)} + n_{CI}^{(t)} + n_{CT}^{(t)}}{2(n_C + n_I + n_T)} \\ \frac{dQ(p|p^{(t)})}{dp_I} &= \frac{2n_{II}^{(t)} + n_{IT}^{(t)} + n_{CI}^{(t)}}{p_I} - \frac{2n_{TT}^{(t)} + n_{CT}^{(t)} + n_{IT}^{(t)}}{1 - p_C - p_I} \\ & p_I^{(t)} &= \frac{2n_{II}^{(t)} + n_{IT}^{(t)} + n_{CI}^{(t)}}{2(n_C + n_I + n_T)} \\ & p_T^{(t)} &= \frac{2n_{TT}^{(t)} + n_{CT}^{(t)} + n_{IT}^{(t)}}{2(n_C + n_I + n_T)} \end{split}$$

Step 3: Return to step 1 (E step) unless stopping criterion has been

Implementation in R

```
set.seed(475)
#initialize variables and initial values
x \leftarrow c(85, 196, 341) #observed data x = (nC, nI, nT)
n \leftarrow rep(0,6) #complete data (y)
#1:nCC,2:nCI,3:nCT,4:nII,5:nIT,6:nTT
n.old < rep(0,6,3)
p <- rep(1/3,3) #starting probability values for each allele frequency 1:pC, 2:pI, 3:pT
p.old <- p
I <- 8 #Number of EM algorithm iterations
#E step: Updated n(t) value
moth.e <- function(x,p){
  n.cc \leftarrow (x[1]*p[1]*p[1])/(p[1]*p[1] + 2*p[1]*p[2] + 2*p[1]*p[3])#1
  n.ci \leftarrow (2*x[1]*p[1]*p[2])/(p[1]*p[1] + 2*p[1]*p[2] + 2*p[1]*p[3])#2
  n.ct <- (2*x[1]*p[1]*p[3])/(p[1]*p[1] + 2*p[1]*p[2] + 2*p[1]*p[3])#3
  n.ii \leftarrow (x[2]*p[2]*p[2])/(p[2]*p[2] + 2*p[2]*p[3])#4
  n.it \leftarrow (2*x[2]*p[2]*p[3])/(p[2]*p[2] + 2*p[2]*p[3])#5
  n.tt <- x[3] #6
  return(c(n.cc,n.ci,n.ct,n.ii,n.it,n.tt))
}
\#M step: Updated p(t+1) value
moth.m <- function(x,n){</pre>
  p.c \leftarrow (2*n[1] + n[2] + n[3])/(2*sum(x))
  p.i \leftarrow (2*n[4] + n[2] + n[5])/(2*sum(x))
  p.t <- (2*n[6] + n[3] + n[5])/(2*sum(x))
  return(c(p.c,p.i,p.t))
#Main Loop
```

```
for(i in 1:I)
{
    n.old <- n
    p.old <- p
    n <- moth.e(x,p)
    p <- moth.m(x,n)
}
p</pre>
```

[1] 0.07083691 0.18873679 0.74042630

```
set.seed(475)
#initialize values
x <- c(85, 196, 341) #observed data x = (nC, nI, nT)
n <- rep(0,6) #complete data (y)
#1:nCC,2:nCI,3:nCT,4:nII,5:nIT,6:nTT
p <- rep(1/3,3) #starting probability values for each allele frequency 1:pC, 2:pI, 3:pT
it <- 8 #Number of

#part a: MH Algorithm to est alpha, beta, mu, lambda (sigma)
m.h <- function(x,p,n){</pre>
```