




# Poisson

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CS109, Stanford University



A satellite image of a tropical storm or hurricane over the Gulf of Mexico. The storm features a prominent eye and a dense, swirling cloud structure. The surrounding ocean is visible in shades of blue, and the landmasses of North and Central America are partially seen at the top and bottom of the frame.

**Probability for Extreme Weather?**

Review

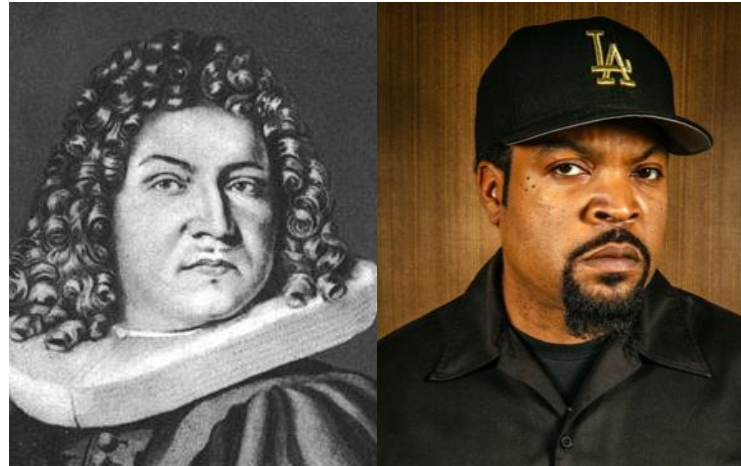
# Natural Exponent Definition

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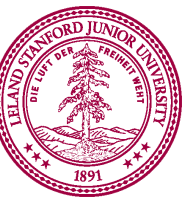
Natural Exponent def:

$$\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n = e^{-\lambda}$$

Jacob  
Bernoulli



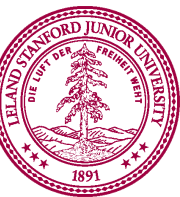
[https://en.wikipedia.org/wiki/E\\_\(mathematical\\_constant\)](https://en.wikipedia.org/wiki/E_(mathematical_constant))



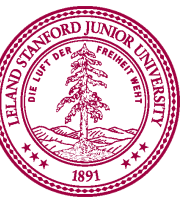
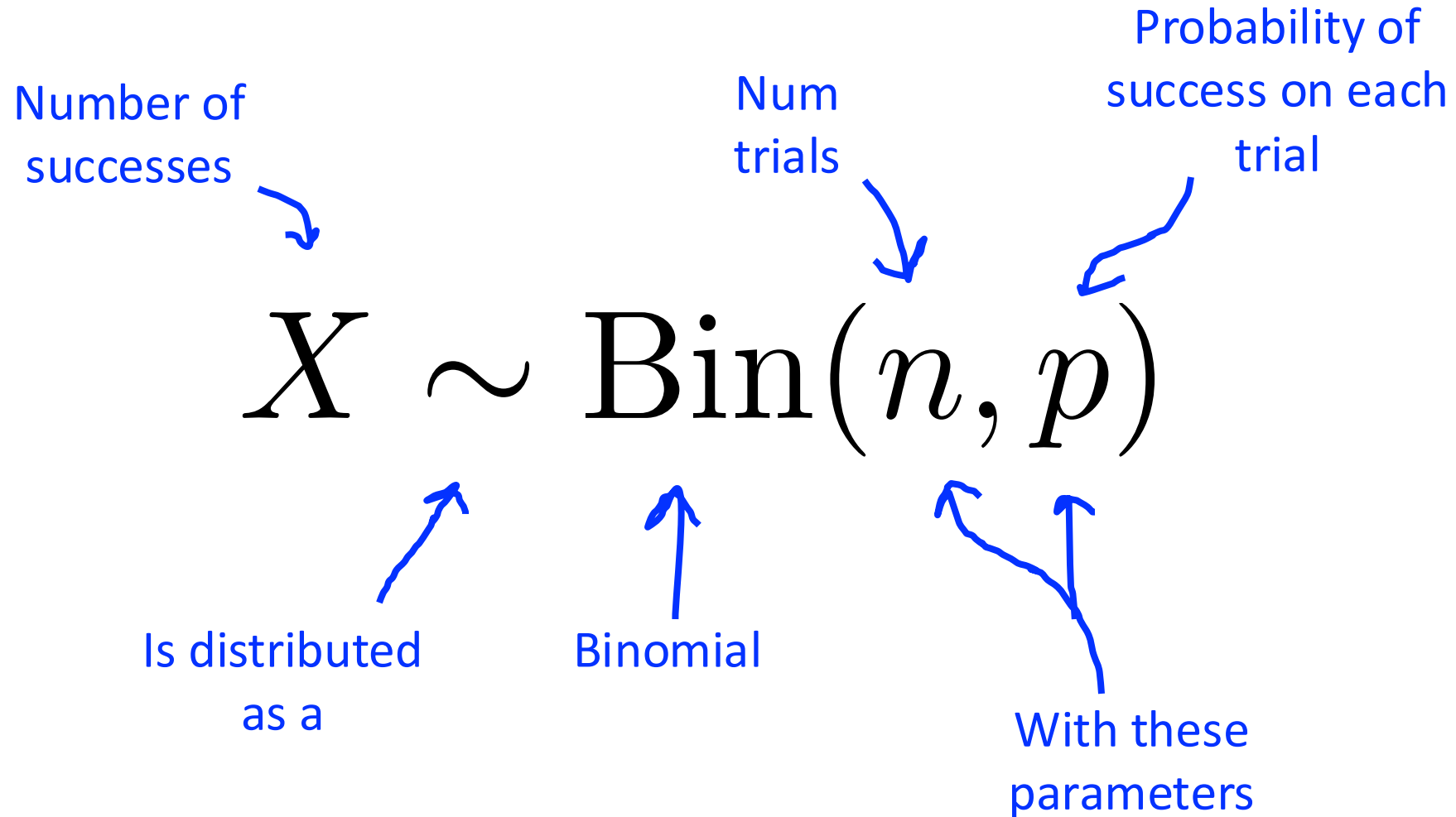
# Binomial Random Variable

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The number of **successes**, in  $n$  independent **trials**, where each **trial** is a **success** with probability  $p$ :



# Declare a Random Variable to be Binomial



# Automatically Know the PMF

Probability Mass Function for a  
Binomial

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

↑  
Probability that there are  
k successes

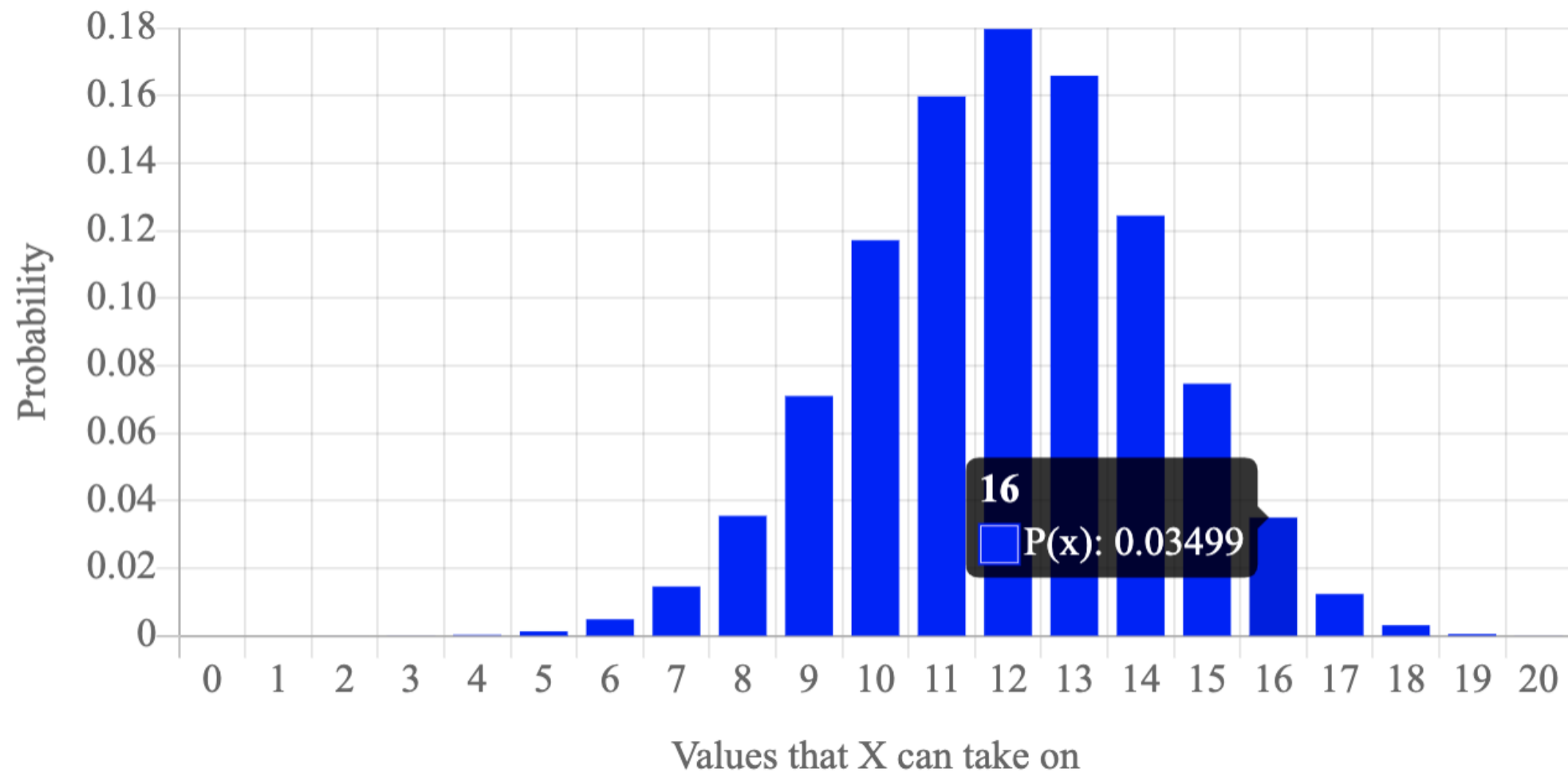
↑  
\* This is also called the  
binomial term





# The PMF as a Graph: $X \sim \text{Bin}(n = 20, p = 0.6)$

Parameter  $n$ :  Parameter  $p$ :





# You Get So Much For Free!

## Binomial Random Variable

**Notation:**  $X \sim \text{Bin}(n, p)$

**Description:** Number of "successes" in  $n$  identical, independent experiments each with probability of success  $p$ .

**Parameters:**  $n \in \{0, 1, \dots\}$ , the number of experiments.  
 $p \in [0, 1]$ , the probability that a single experiment gives a "success".

**Support:**  $x \in \{0, 1, \dots, n\}$

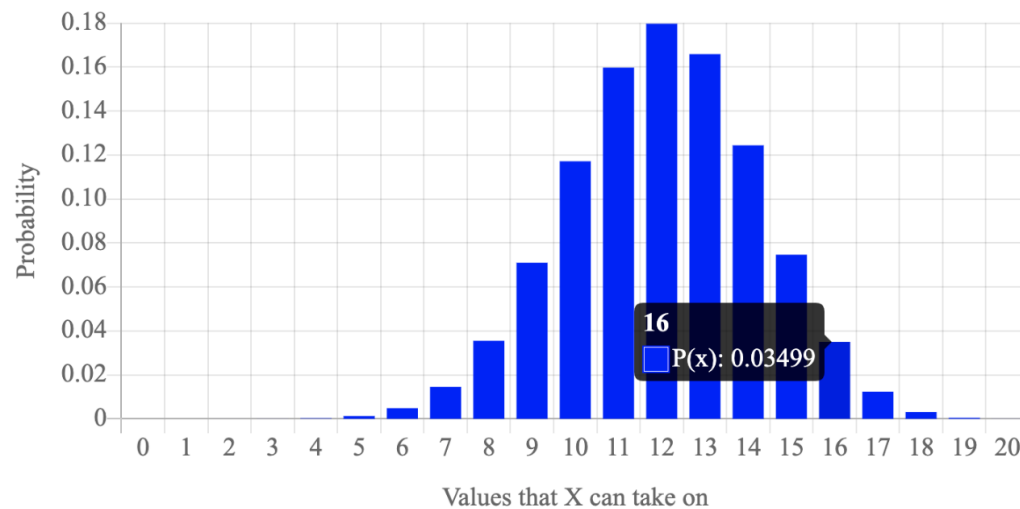
**PMF equation:**  $\Pr(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$

**Expectation:**  $E[X] = n \cdot p$

**Variance:**  $\text{Var}(X) = n \cdot p \cdot (1 - p)$

**PMF graph:**

Parameter  $n$ :  Parameter  $p$ :



## Bernoulli Random Variable

**Notation:**  $X \sim \text{Bern}(p)$

**Description:** A boolean variable that is 1 with probability  $p$

**Parameters:**  $p$ , the probability that  $X = 1$ .

**Support:**  $x$  is either 0 or 1

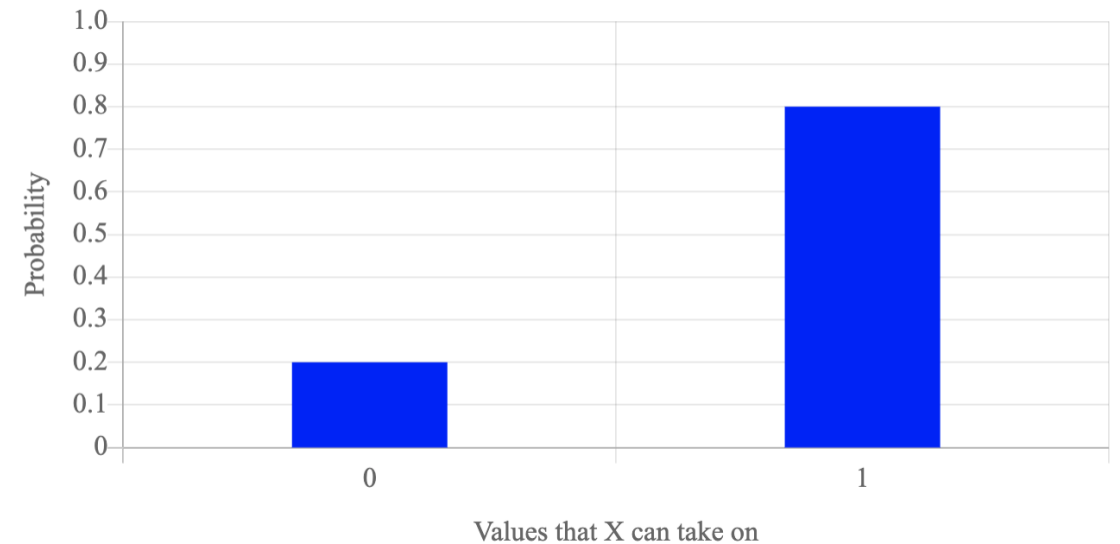
**PMF equation:**  $\Pr(X = x) = \begin{cases} p & \text{if } x = 1 \\ 1 - p & \text{if } x = 0 \end{cases}$

**Expectation:**  $E[X] = p$

**Variance:**  $\text{Var}(X) = p(1 - p)$

**PMF graph:**

Parameter  $p$ :



What if we could summarize the whole beautiful PMF into a single number?

# Expected Value

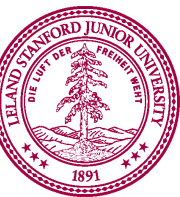
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The value

The probability of that value

$$E[X] = \sum_x x \cdot P(X = x)$$

Loop over all values  $x$  that  $X$  can take on



# St. Petersburg Paradox

The Game:

- We have a fair coin (lands on heads with  $p = 0.5$ )
- Let  $n$  = number of coin flips to get the first heads
- You will win:  $\$2^n$

How much would you pay to play?

Let  $X$  be your winnings.

$$E[X] = 2^1 \left(\frac{1}{2}\right)^1 + 2^2 \left(\frac{1}{2}\right)^2 + 2^3 \left(\frac{1}{2}\right)^3 + \dots = \sum_{i=1}^{\infty} 1 = \infty$$

What if you could play this game for only \$1000...but just once?



# St. Petersburg Paradox

The Game:

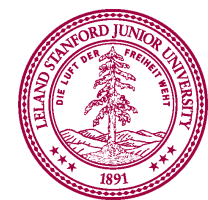
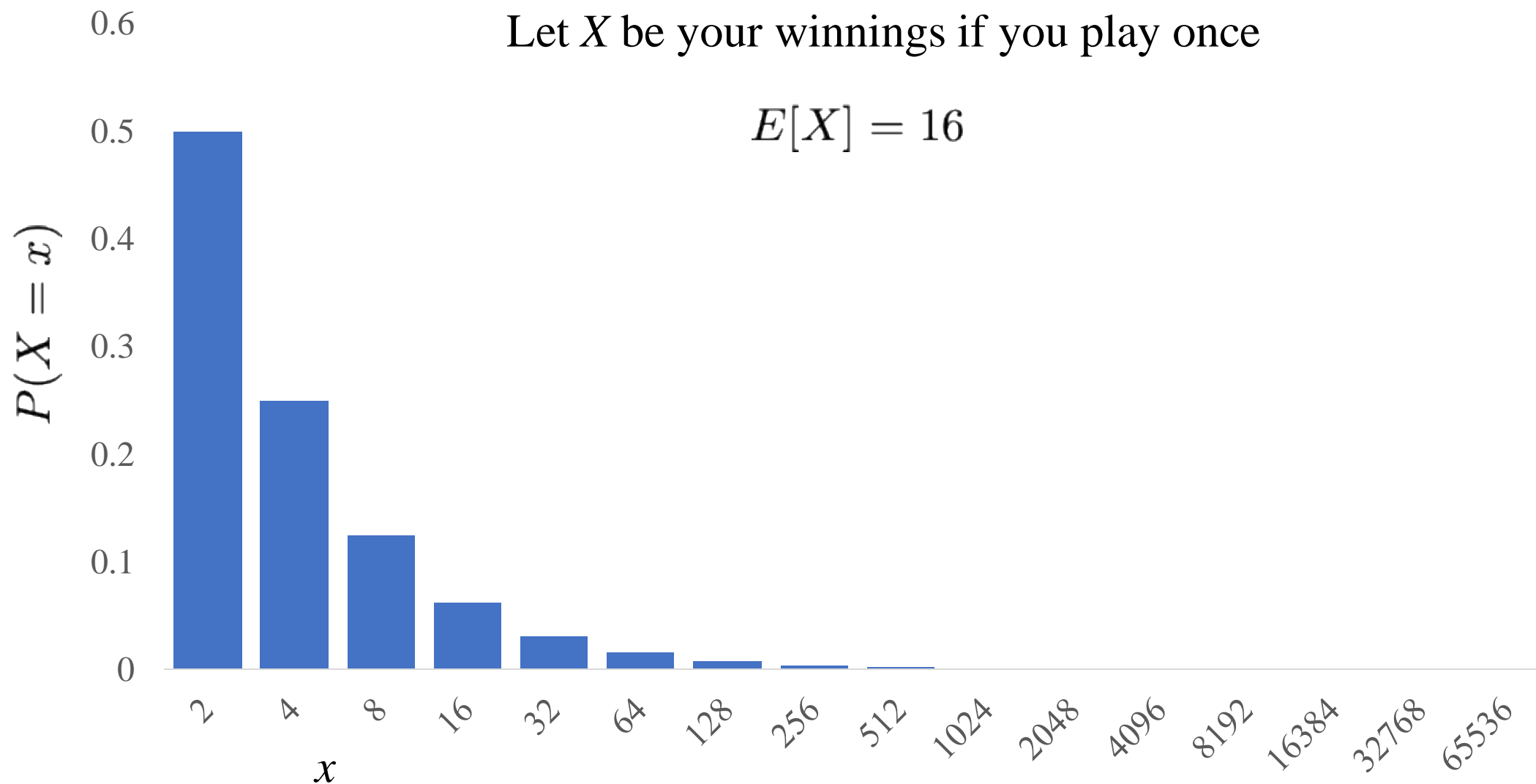
- We have a fair coin (lands on heads with  $p = 0.5$ )
- Let  $n$  = number of coin flips to get the first heads
- You will win:  $\$2^n$
- If you win over \$65,536 **I leave the country.**

How much would you pay to play?

Let  $X$  be your winnings.

$$E[X] = 2^1 \left(\frac{1}{2}\right)^1 + 2^2 \left(\frac{1}{2}\right)^2 + \cdots + 2^{16} \left(\frac{1}{2}\right)^{16} = \sum_{i=1}^{16} 1 = 16$$

# St Petersburg Probability Mass Function



# Properties of Expectation (more on this later)

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## Linearity:

$$E[aX + b] = aE[X] + b$$

**Expectation of a sum** is the sum of expectations

$$E[X + Y] = E[X] + E[Y]$$

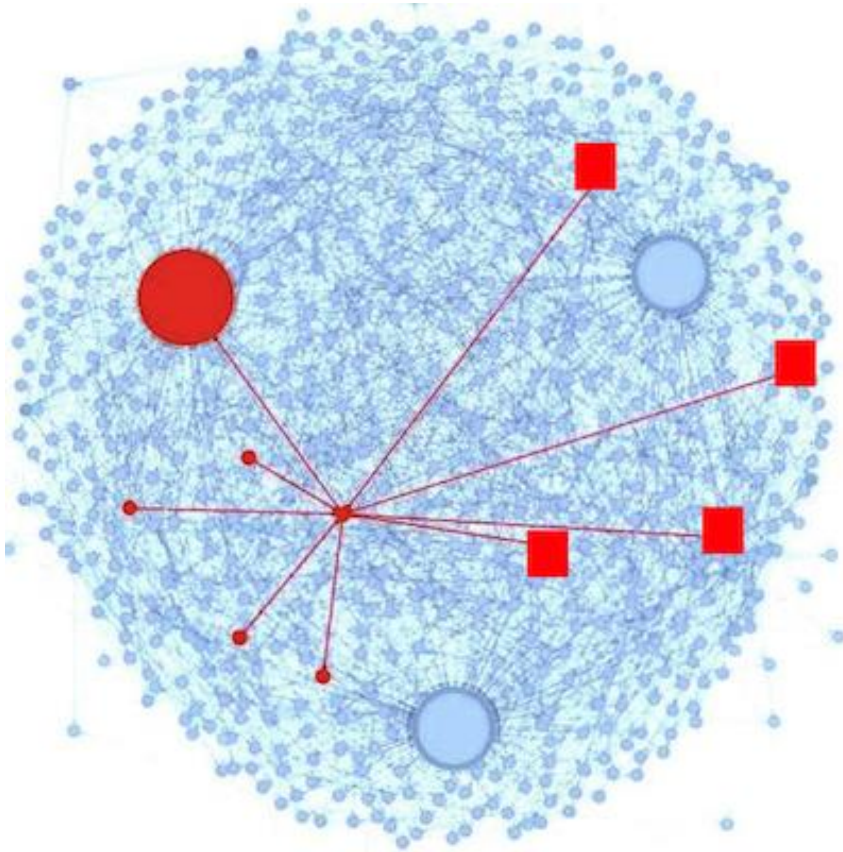
## Unconscious statistician:

$$E[g(x)] = \sum_{x \in X} g(x)P(X = x)$$



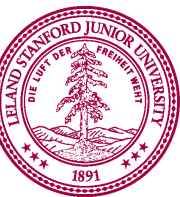
# Intuition: Peer Grading

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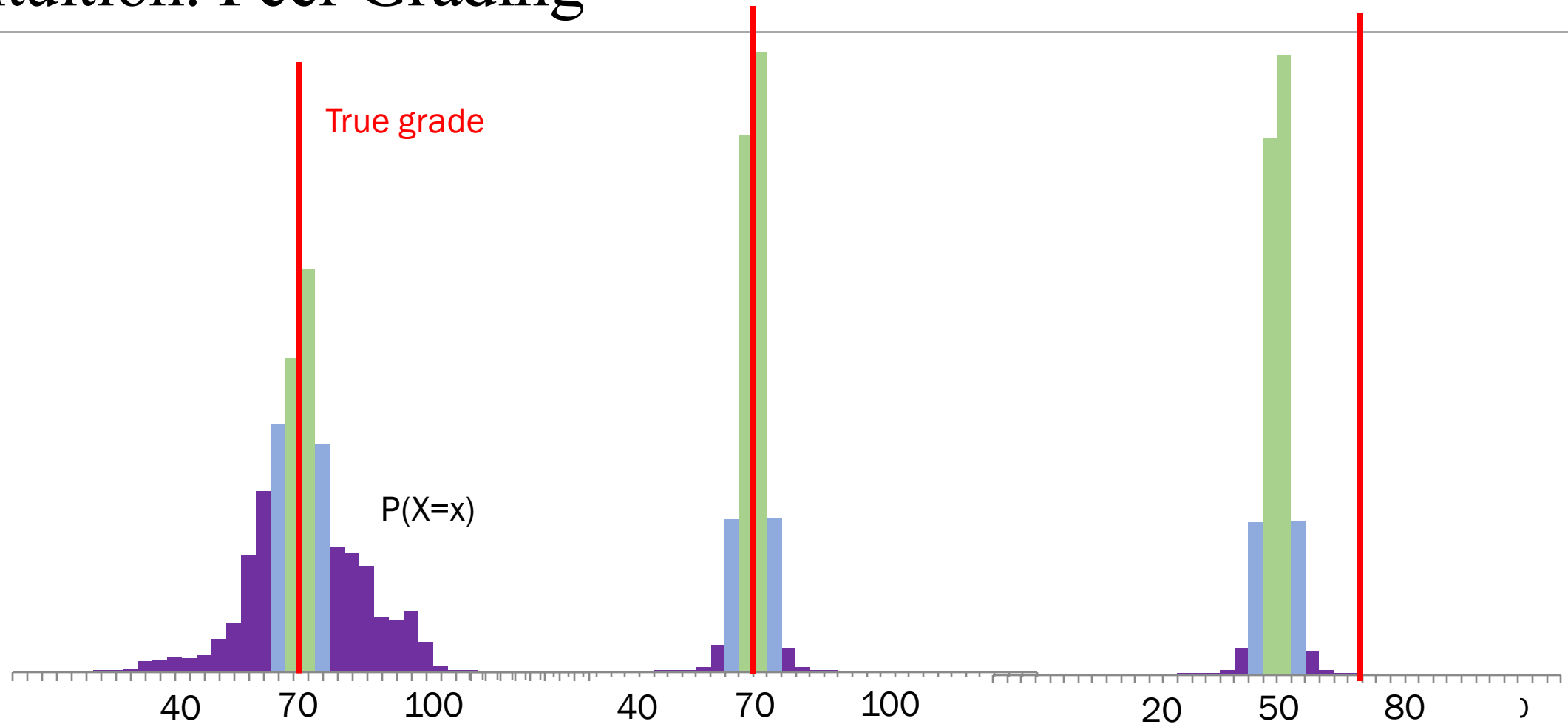
Peer Grading on Coursera HCI.

31,067 peer grades for 3,607 students.





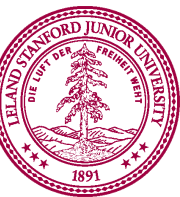
# Intuition: Peer Grading



A

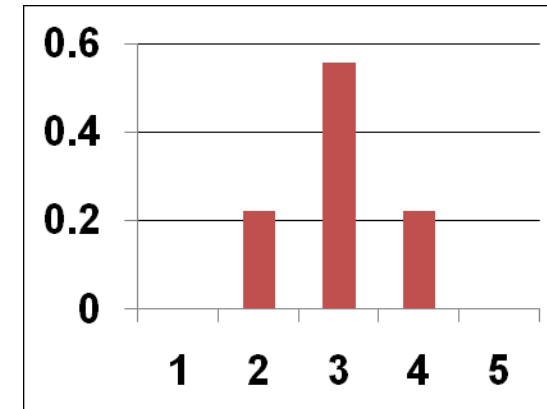
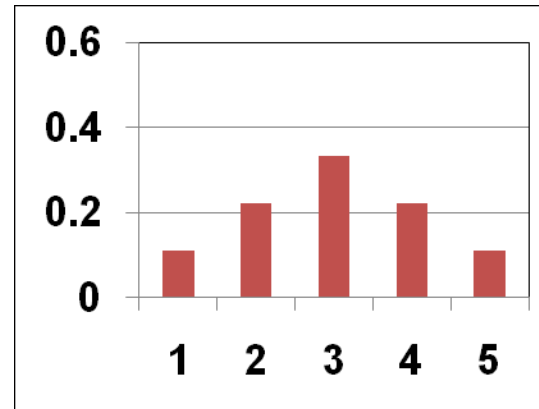
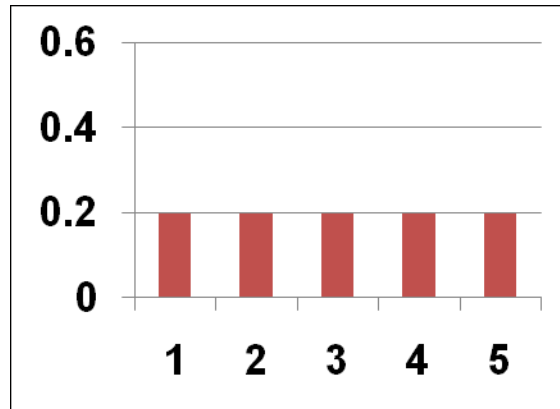
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# Intuition: Measure of Spread

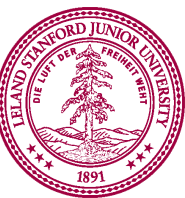
Consider the following 3 distributions (PMFs)



All have the same expected value,  $E[X] = 3$

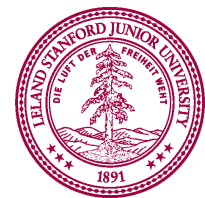
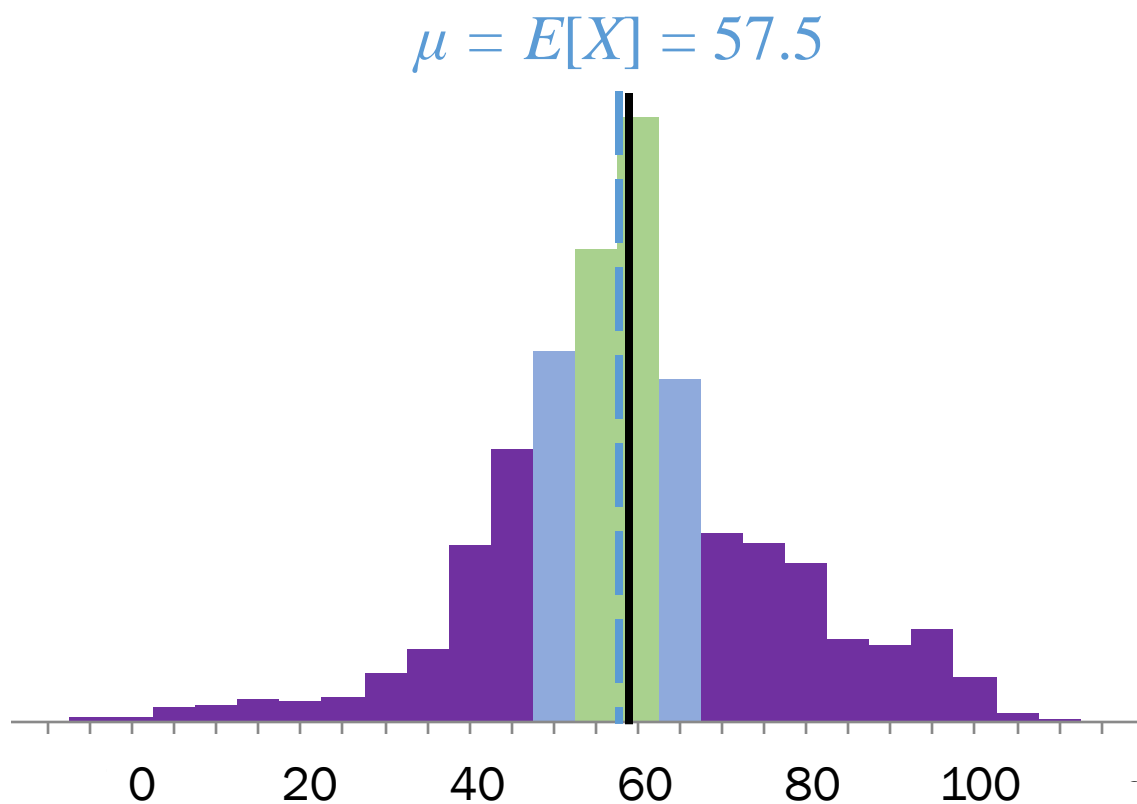
But “spread” in distributions is different

Invent a formal quantification of “spread”?



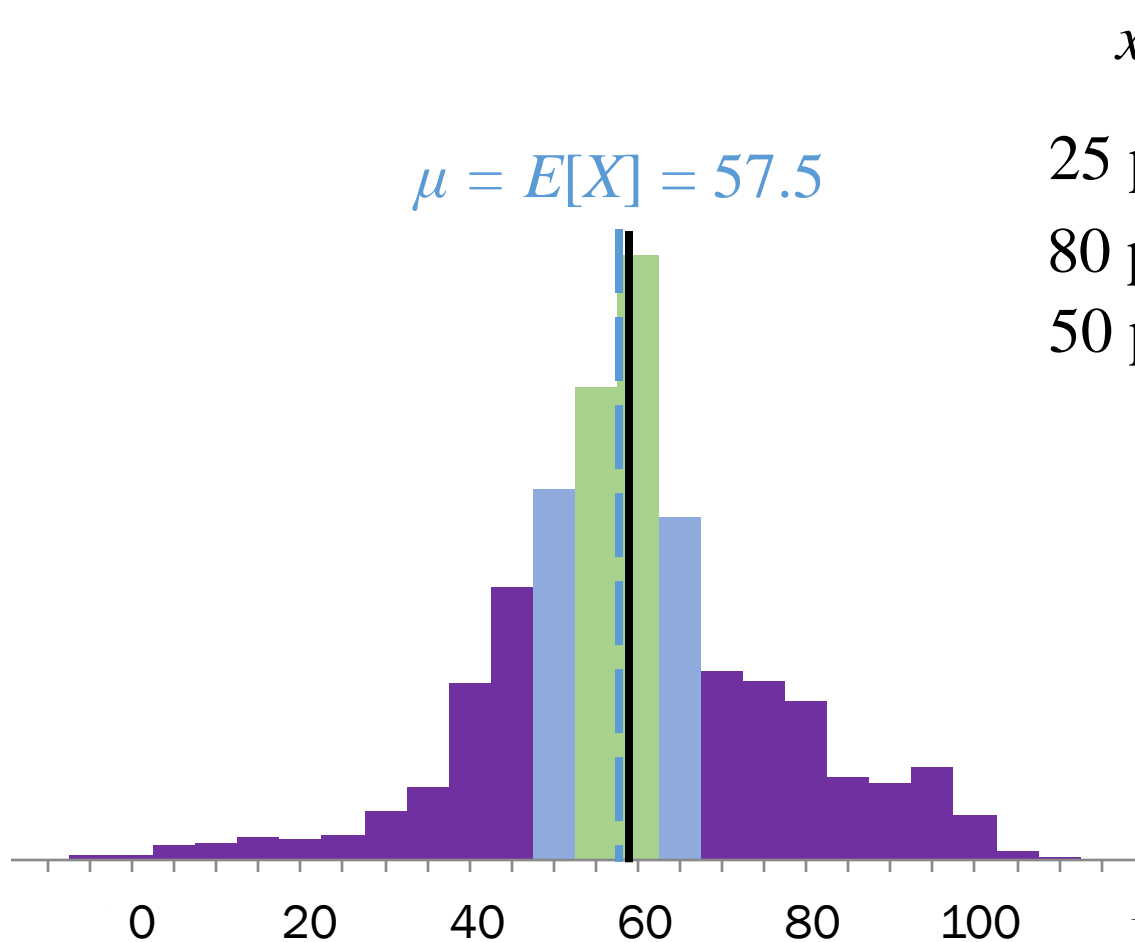
# Peer grading in Coursera HCI

Let  $X$  be a random variable that represents a peer grade



# Peer grading in Coursera HCI

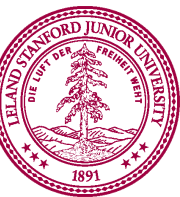
Let  $X$  be a random variable that represents a peer grade



$x$	$(x - \mu)^2$	$P(X = x)$
25 points	1056 points <sup>2</sup>	0.02
80 points	506 points <sup>2</sup>	0.09
50 points	56 points <sup>2</sup>	0.12
...		

$$E[(X - \mu)^2] = 52 \text{ points}^2$$

$$\text{Std}(X) = 7.2 \text{ points}$$





# How Should We Measure Spread?

Let  $X$  be a random variable

$$\mu = E[X] = 57.5$$

Spread stat.

On average..

distance

$$\text{Var}(X) = E[(X - E[X])^2]$$

The random  
variable  $X$

The mean  
of  $X$

Different Possibility:  $E[|X - E[X]|]$ ?



End Review

Finishing up on Variance

# Variance

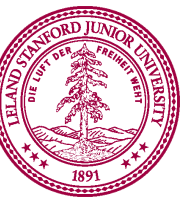
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If  $X$  is a random variable with mean  $\mu$  then the **variance** of  $X$ , denoted  $\text{Var}(X)$ , is:

$$\text{Var}(X) = E[(X - \mu)^2]$$

Variance is a formal definition of the **spread** of a random variable.

Also known as the 2nd **Central** Moment, or square of the Standard Deviation





# Computing Variance

$$\begin{aligned}\text{Var}(X) &= E[(X - \mu)^2] \\ &= \sum_x (x - \mu)^2 P(X = x)\end{aligned}$$

Law of unconscious statistician

$$\begin{aligned}&= \sum_x (x^2 - 2\mu x + \mu^2) P(X = x) \\ &= \sum_x x^2 P(X = x) - 2\mu \sum_x x P(X = x) + \mu^2 \sum_x P(X = x) \\ &= E[X^2] - 2\mu E[X] + \mu^2 \\ &= E[X^2] - 2\mu^2 + \mu^2 \\ &= E[X^2] - \mu^2 \\ &= E[X^2] - (E[X])^2\end{aligned}$$

Notation  
 $\mu = E[X]$



How do you get  $E[X^2]$ ?

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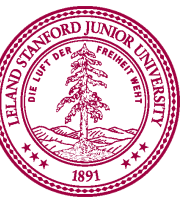
$$\text{Var}(X) = E[X^2] - E[X]^2$$

**Unconscious statistician:**

$$E[g(X)] = \sum_x g(x) P(X = x)$$

**$E[X^2]$ :**

$$E[X^2] = \sum_x x^2 \cdot P(X = x)$$



# Standard Deviation?

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$$\text{Std}(X) = \sqrt{\text{Var}(X)}$$

Units are in points



Units are in points squared

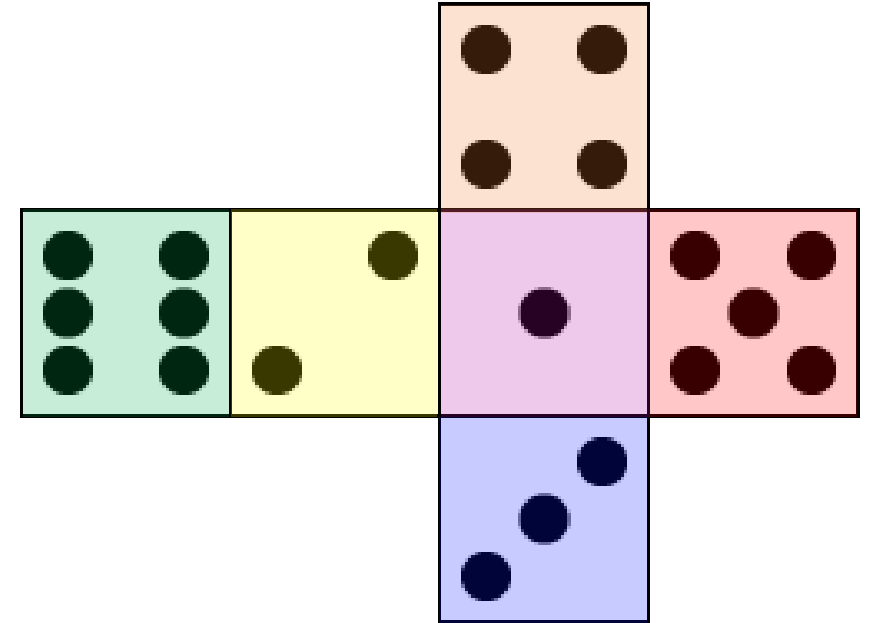


# Example: Variance of a Dice Roll

Let  $X$  be the result of rolling a 6 sided dice.

What is  $\text{Var}(X)$ ?

$$\text{Var}(X) = E[X^2] - E[X]^2$$



# Example: Variance of a Dice Roll

$$\text{Var}(X) = E[X^2] - E[X]^2$$

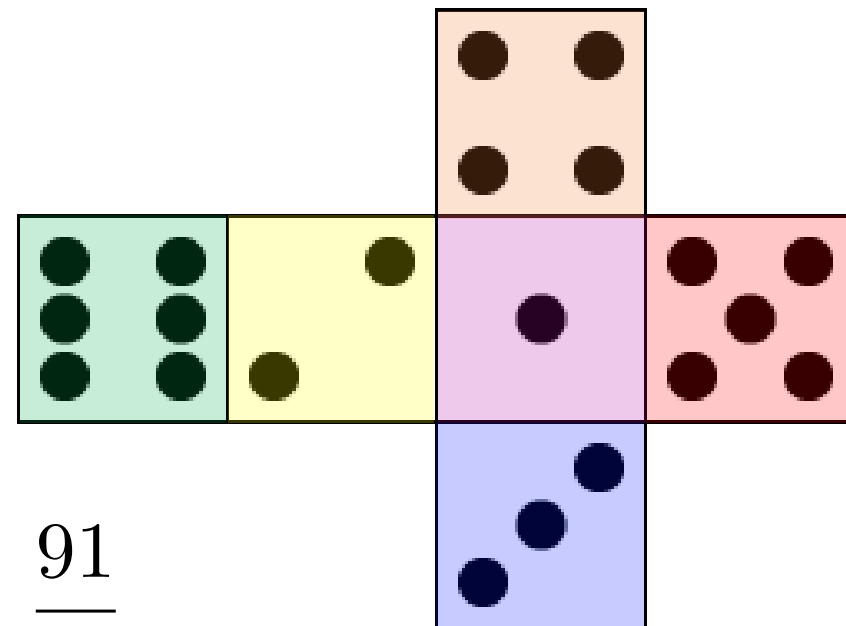
Let  $X$  be the result of rolling a 6 sided dice.

What is  $\text{Var}(X)$ ?

$$E[X] = 3.5$$

$$E[X^2] = 1^2 \frac{1}{6} + 2^2 \frac{1}{6} + 3^2 \frac{1}{6} + 4^2 \frac{1}{6} + 5^2 \frac{1}{6} = \frac{91}{6}$$

$$\begin{aligned}\text{Var}(X) &= E[X^2] - E[X]^2 \\ &= \frac{91}{6} - (3.5)^2 = 2.91\end{aligned}$$

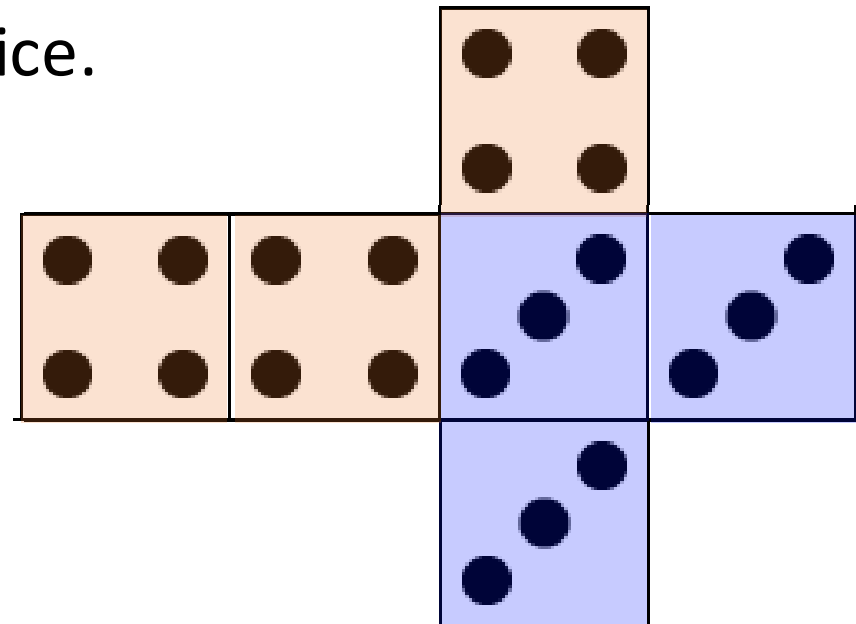


# Example: Variance of a Dice Roll

$$\text{Var}(X) = E[X^2] - E[X]^2$$

Let  $X$  be the result of rolling **this weird** 6 sided dice.

What is  $\text{Var}(X)$ ?





# Example: Variance of a Dice Roll

$$\text{Var}(X) = E[X^2] - E[X]^2$$

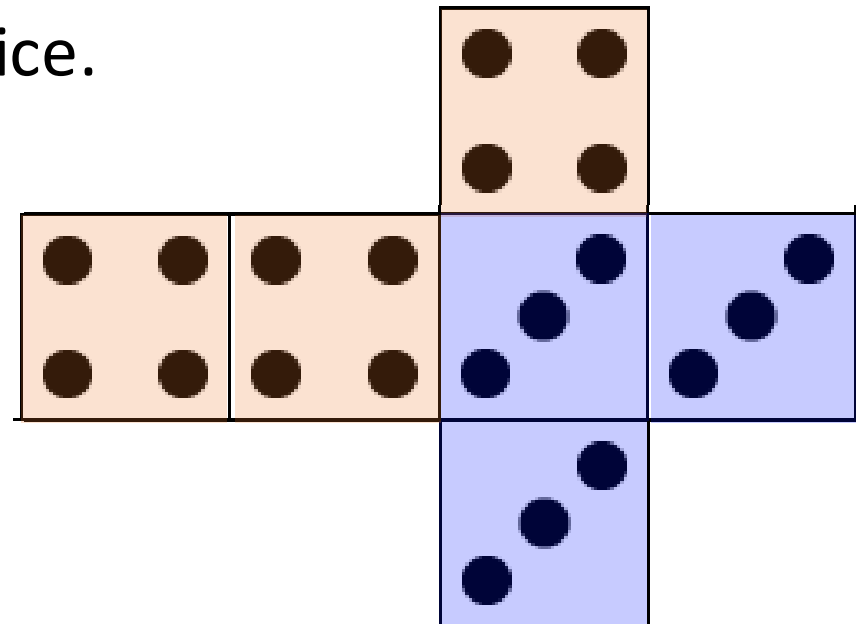
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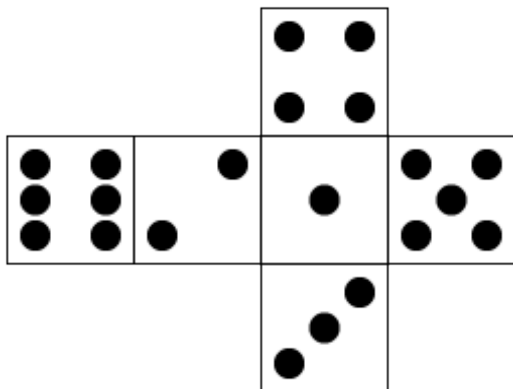
$$E[X] = 3.5$$

$$E[X^2] = 3^2 \cdot \frac{3}{6} + 4^2 \cdot \frac{3}{6} = 12.5$$

$$\begin{aligned}\text{Var}(X) &= E[X^2] - E[X]^2 \\ &= 12.5 - (3.5)^2 = 0.25\end{aligned}$$

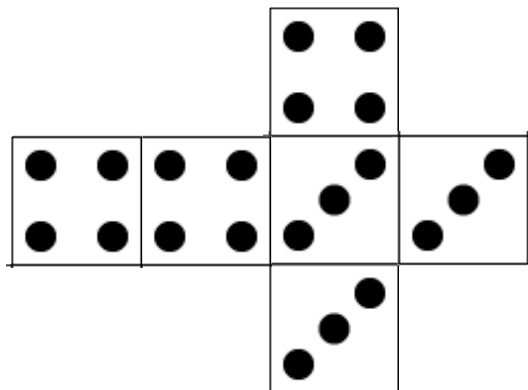


# Variance of a 6 Sided Dice



$$\text{Var}(X) = 2.91$$

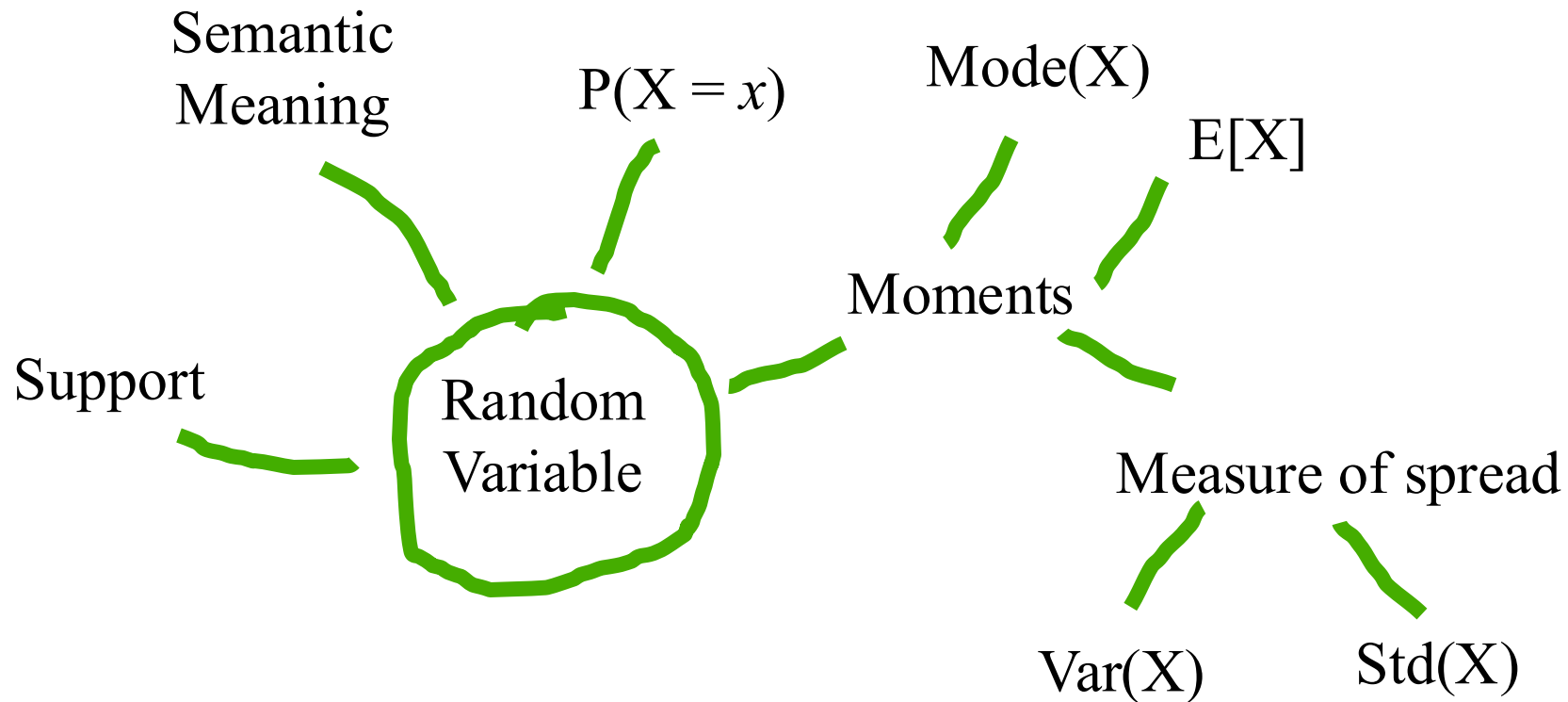
$$\text{Std}(X) = 1.7$$



$$\text{Var}(X) = 0.25$$

$$\text{Std}(X) = 0.5$$

# Fundamental Properties of Random Variables



# You Get So Much For Free!

## Binomial Random Variable

**Notation:**  $X \sim \text{Bin}(n, p)$

**Description:** Number of "successes" in  $n$  identical, independent experiments each with probability of success  $p$ .

**Parameters:**  $n \in \{0, 1, \dots\}$ , the number of experiments.  
 $p \in [0, 1]$ , the probability that a single experiment gives a "success".

**Support:**  $x \in \{0, 1, \dots, n\}$

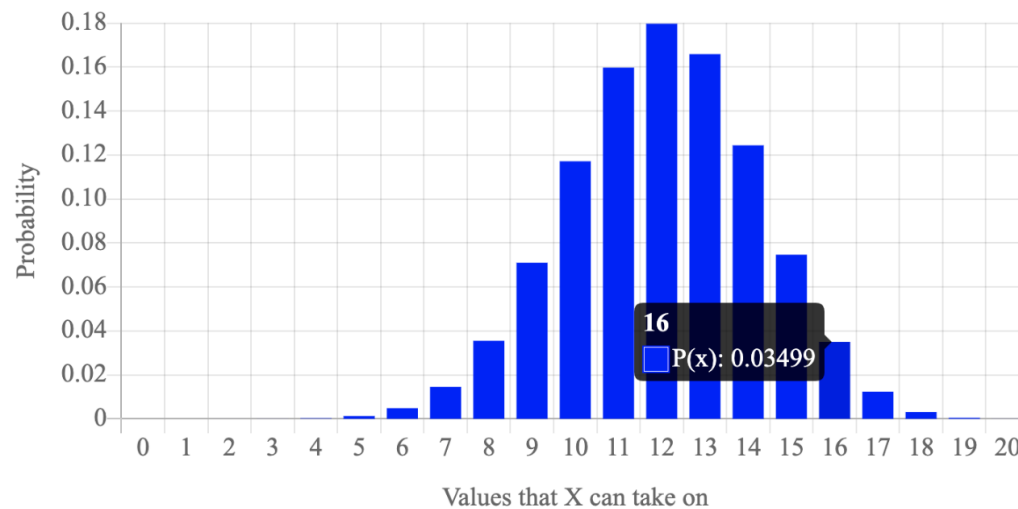
**PMF equation:**  $\Pr(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$

**Expectation:**  $E[X] = n \cdot p$

**Variance:**  $\text{Var}(X) = n \cdot p \cdot (1 - p)$

**PMF graph:**

Parameter  $n$ : 20      Parameter  $p$ : 0.60



## Bernoulli Random Variable

**Notation:**  $X \sim \text{Bern}(p)$

**Description:** A boolean variable that is 1 with probability  $p$

**Parameters:**  $p$ , the probability that  $X = 1$ .

**Support:**  $x$  is either 0 or 1

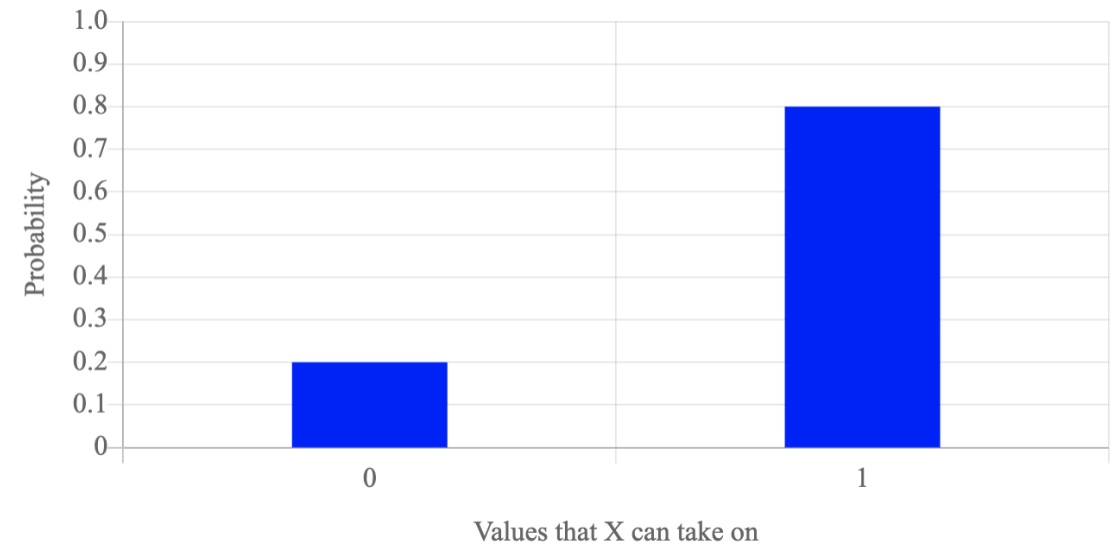
**PMF equation:**  $\Pr(X = x) = \begin{cases} p & \text{if } x = 1 \\ 1 - p & \text{if } x = 0 \end{cases}$

**Expectation:**  $E[X] = p$

**Variance:**  $\text{Var}(X) = p(1 - p)$

**PMF graph:**

Parameter  $p$ : 0.80



# Curious? Proof of Variance for a Binomial

$$\begin{aligned}
 E(X^2) &= \sum_{k=0}^n k^2 \binom{n}{k} p^k q^{n-k} \\
 &= \sum_{k=0}^n kn \binom{n-1}{k-1} p^k q^{n-k} \\
 &= np \sum_{k=1}^n k \binom{n-1}{k-1} p^{k-1} q^{(n-1)-(k-1)} \\
 &= np \sum_{j=0}^m (j+1) \binom{m}{j} p^j q^{m-j} \\
 &= np \left( \sum_{j=0}^m j \binom{m}{j} p^j q^{m-j} + \sum_{j=0}^m \binom{m}{j} p^j q^{m-j} \right) \\
 &= np \left( \sum_{j=0}^m m \binom{m-1}{j-1} p^j q^{m-j} + \sum_{j=0}^m \binom{m}{j} p^j q^{m-j} \right) \\
 &= np \left( (n-1)p \sum_{j=1}^m \binom{m-1}{j-1} p^{j-1} q^{(m-1)-(j-1)} + \sum_{j=0}^m \binom{m}{j} p^j q^{m-j} \right) \\
 &= np \left( (n-1)p(p+q)^{m-1} + (p+q)^m \right) \\
 &= np((n-1)p + 1) \\
 &= n^2 p^2 + np(1-p)
 \end{aligned}$$

Definition of **Binomial Distribution**:  $p + q = 1$

**Factors of Binomial Coefficient**:  $k \binom{n}{k} = n \binom{n-1}{k-1}$

Change of limit: term is zero when  $k - 1 = 0$

putting  $j = k - 1, m = n - 1$

splitting sum up into two

**Factors of Binomial Coefficient**:  $j \binom{m}{j} = m \binom{m-1}{j-1}$

Change of limit: term is zero when  $j - 1 = 0$

**Binomial Theorem**

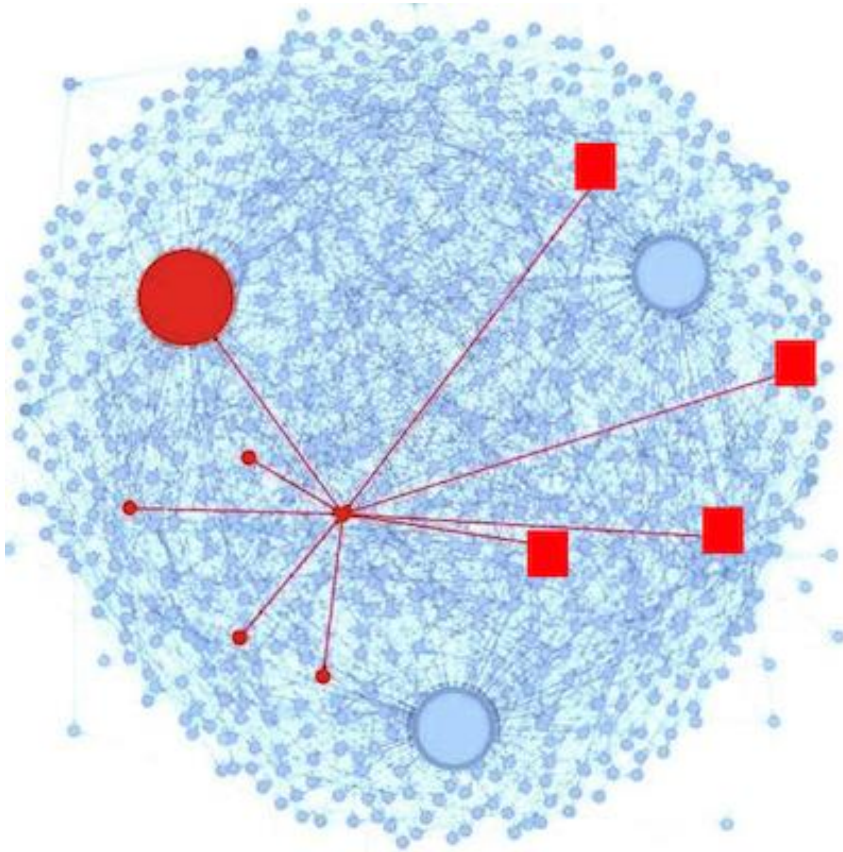
as  $p + q = 1$

by algebra



# Is Peer Grading Accurate Enough?

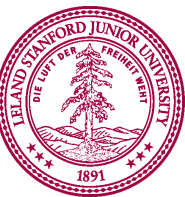
*Looking ahead*



Peer Grading on Coursera HCI.

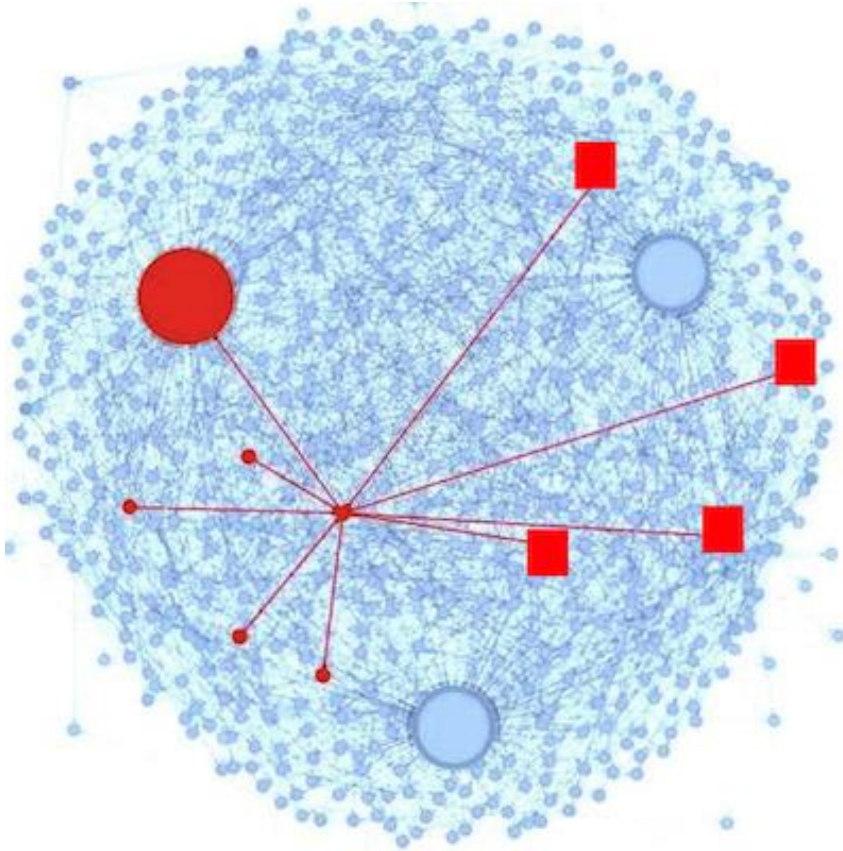
31,067 peer grades for 3,607 students.

Tuned Models of Peer Assessment. C Piech, J Huang, A Ng, D Koller



# Is Peer Grading Accurate Enough?

*Looking ahead*



1. Defined random variables for:
  - True grade ( $s_i$ ) for assignment  $i$
  - Observed ( $z_i^j$ ) score for assign  $i$
  - Bias ( $b_j$ ) for each grader  $j$
  - Variance ( $r_j$ ) for each grader  $j$
2. Designed a probabilistic model that defined the distributions for all random variables

$$s_i \sim \text{Bin}(\text{points}, \theta)$$

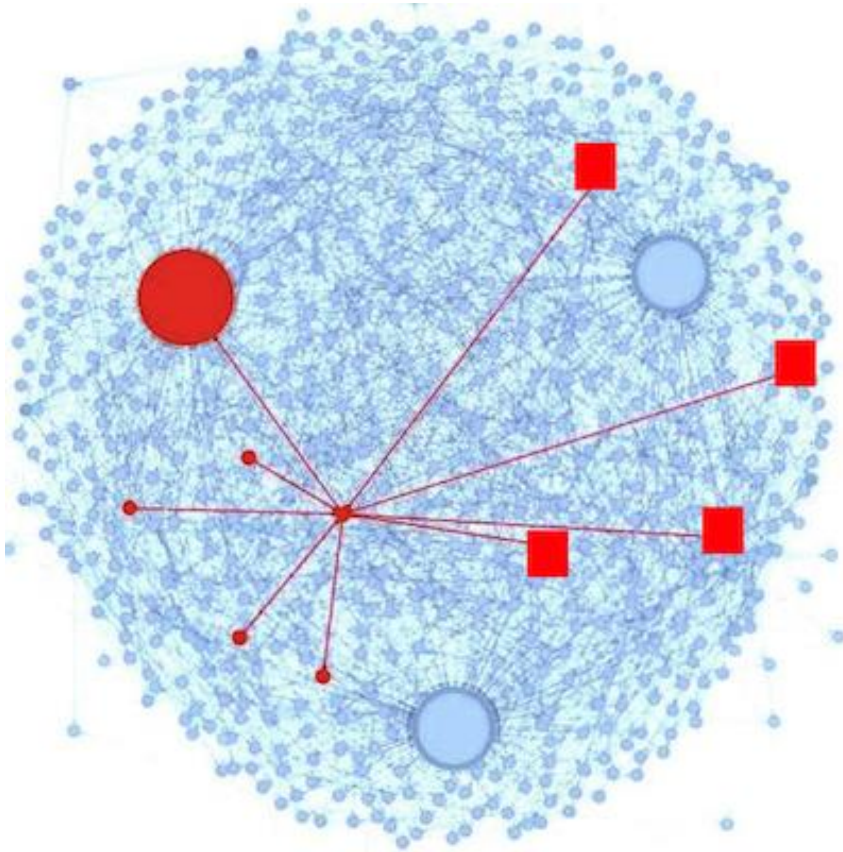
$$z_i^j \sim \mathcal{N}(\mu = s_i + b_j, \sigma = \sqrt{r_j})$$

Problem param  
↙



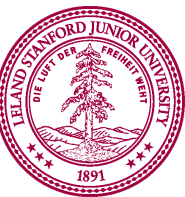
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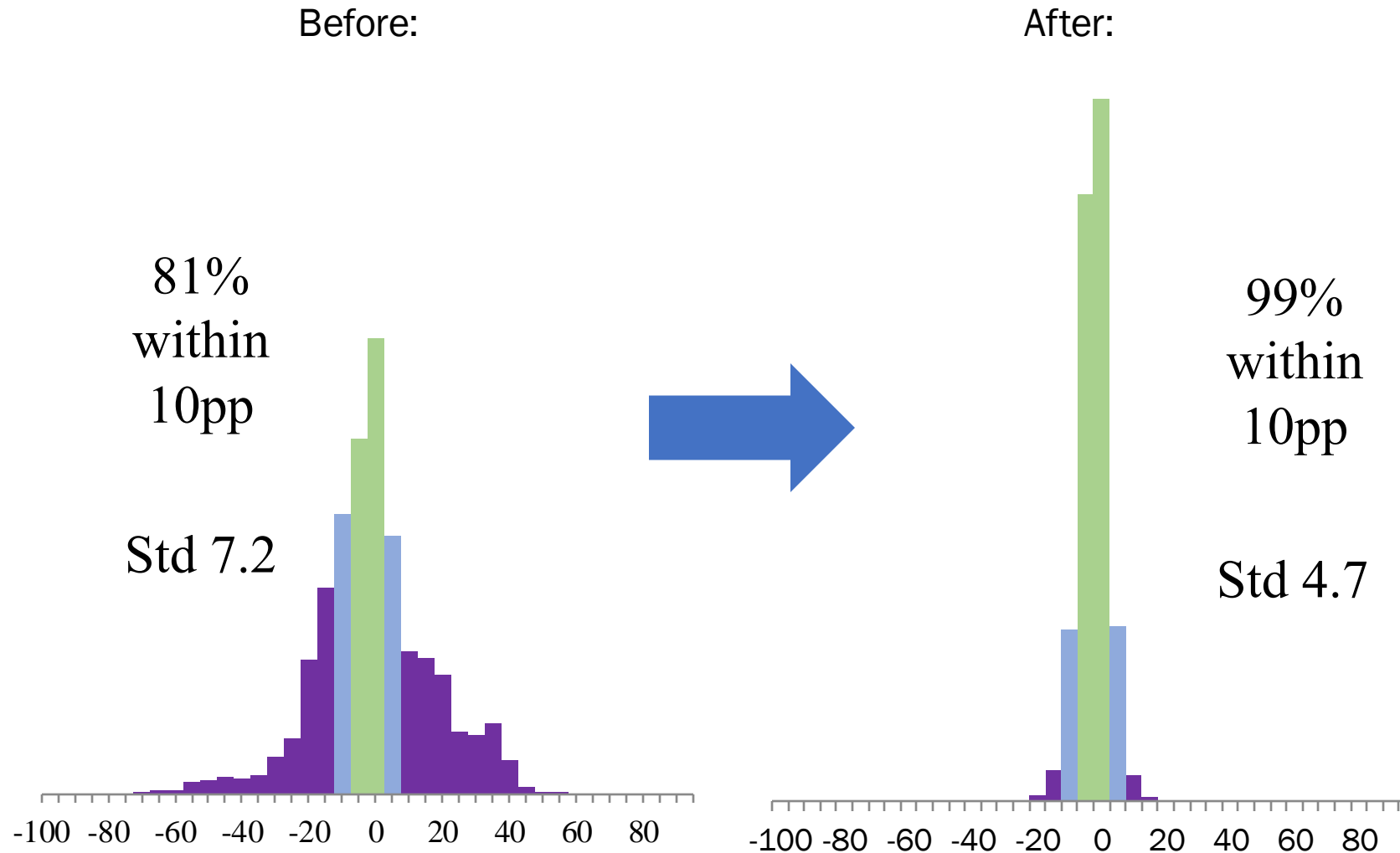


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2. Designed a probabilistic model that defined the relationship between all the random variables
3. Found the variable assignments that maximized the probability of our observed data

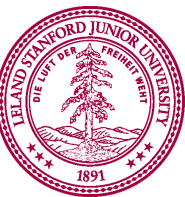
↑  
Inference or Machine Learning



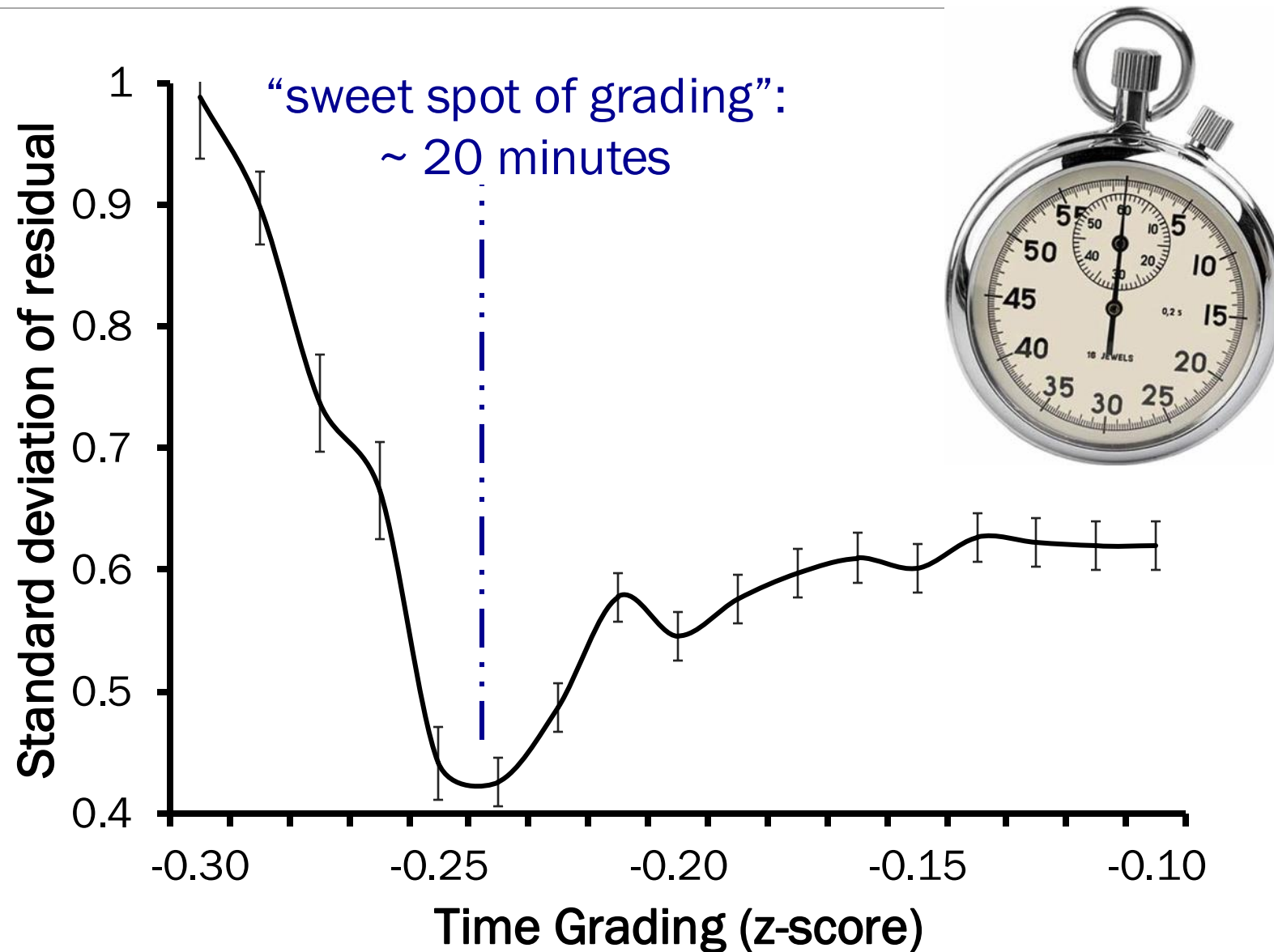
# Yes, With Probabilistic Modelling



Tuned Models of Peer Assessment. C Piech, J Huang, A Ng, D Koller



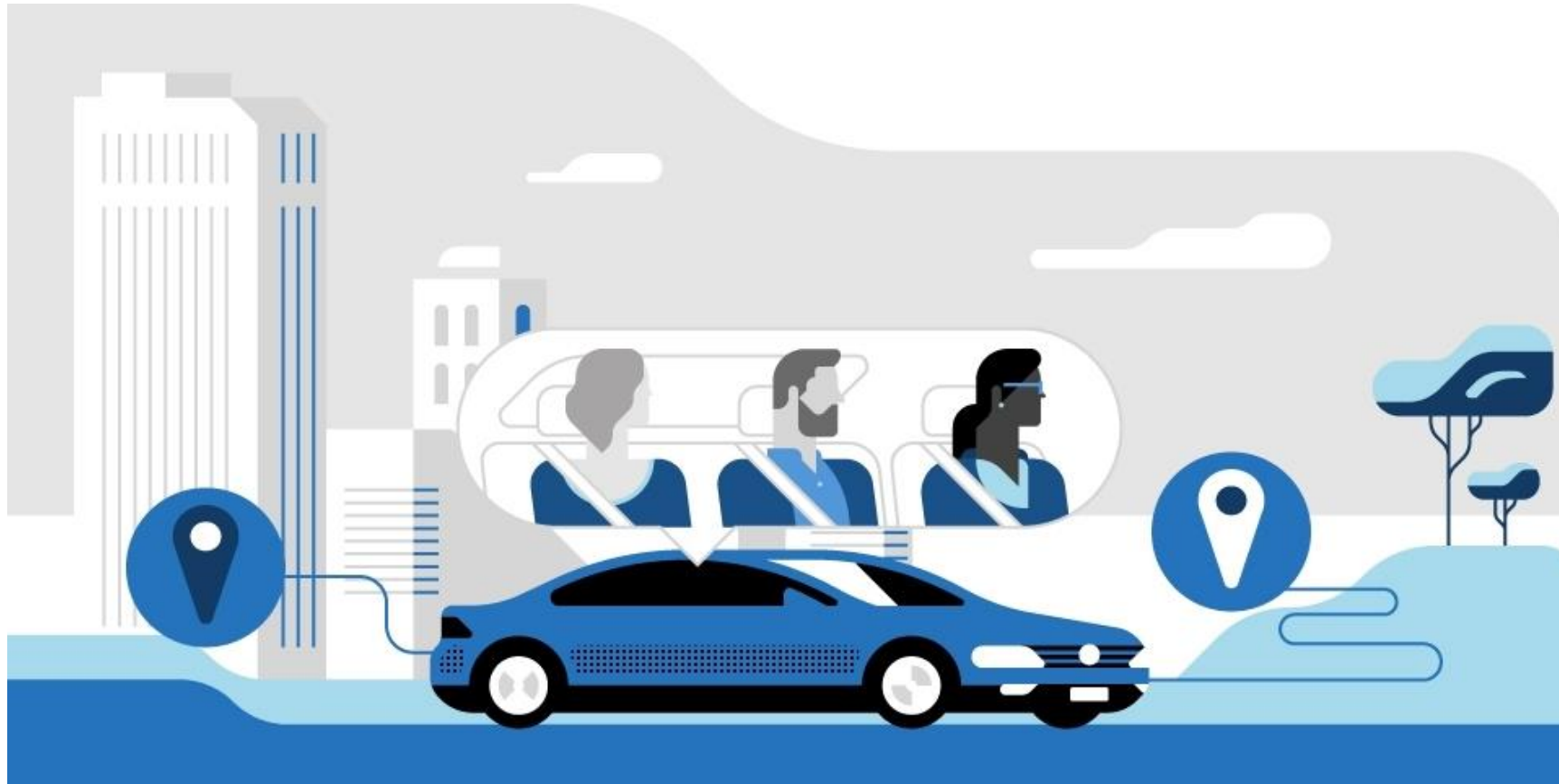
# Grading Sweet Spot



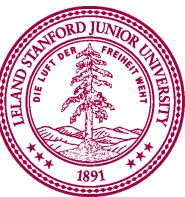
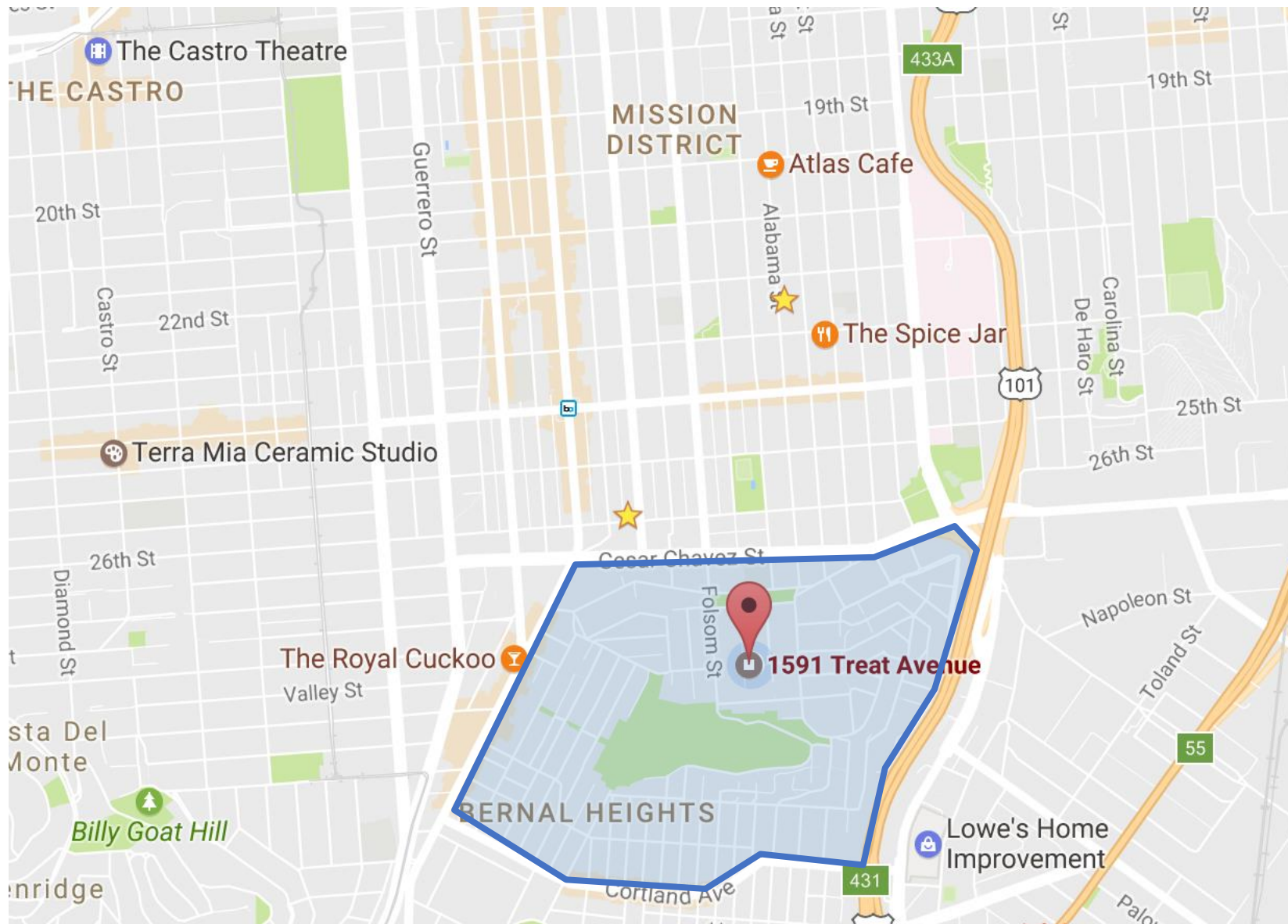
Ready..

# Algorithmic Ride Sharing

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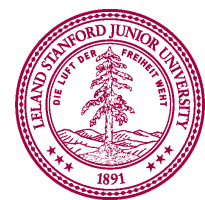
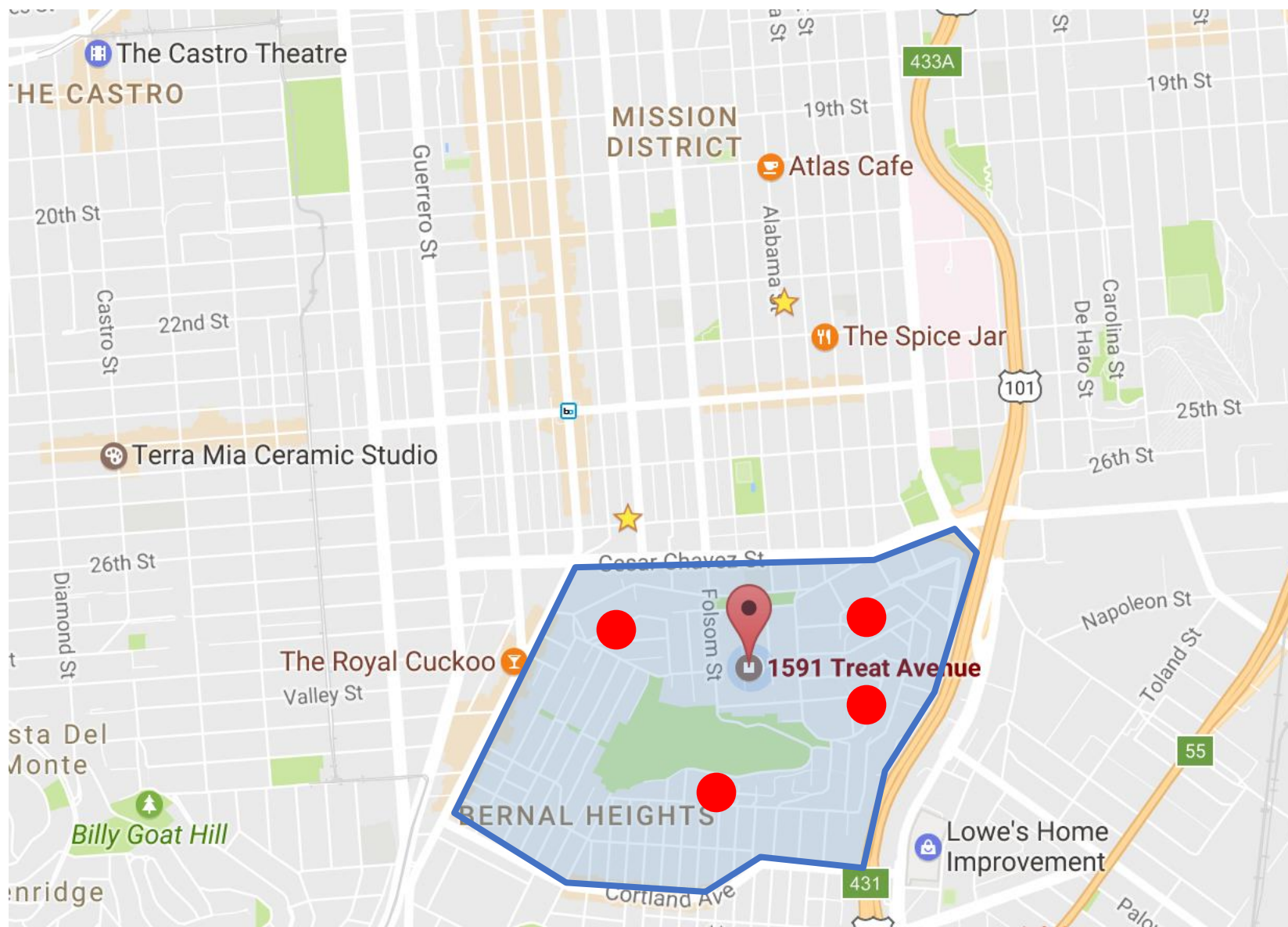


# Probability of $k$ requests from this area in the next 1 min





# Probability of $k$ requests from this area in the next 1 min





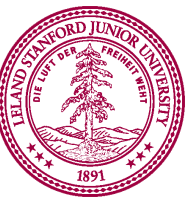
# Probability of $k$ requests from this area in the next 1 min



# Probability of **$k$ requests** from this area in the next 1 min

On average  $\lambda = 5$  requests per minute

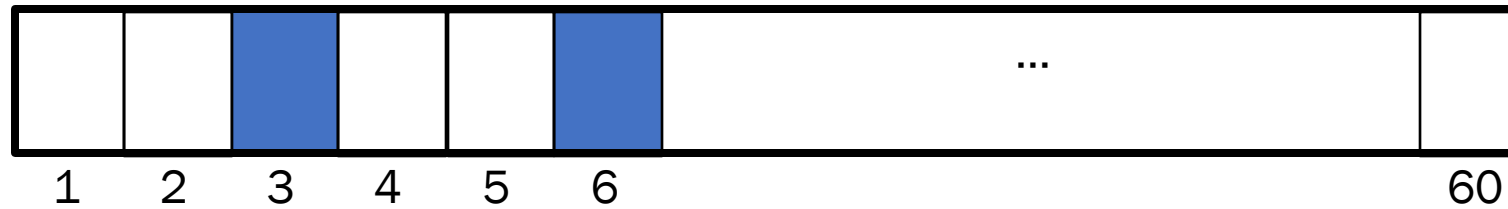
We can break the next minute down into seconds



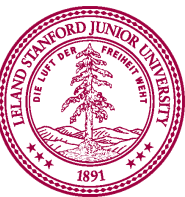
# Probability of **$k$ requests** from this area in the next 1 min

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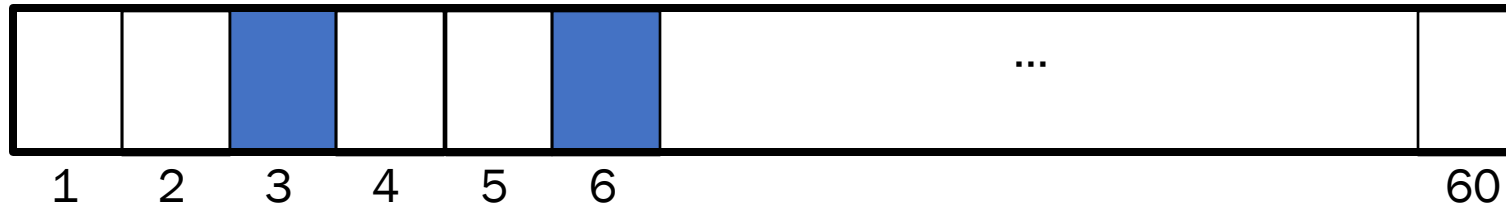
At each second either get a request or you don't.



# Probability of **$k$ requests** from this area in the next 1 min

On average  $\lambda = 5$  requests per minute

We can break the next minute down into seconds



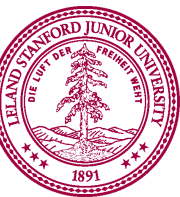
At each second either get a request or you don't.

Let  $X$  = Number of requests in the minute

$$X \sim \text{Bin}(n = 60, p = 5/60)$$

$$P(X = k) = \binom{n}{k} (p)^k (1 - p)^{n-k}$$

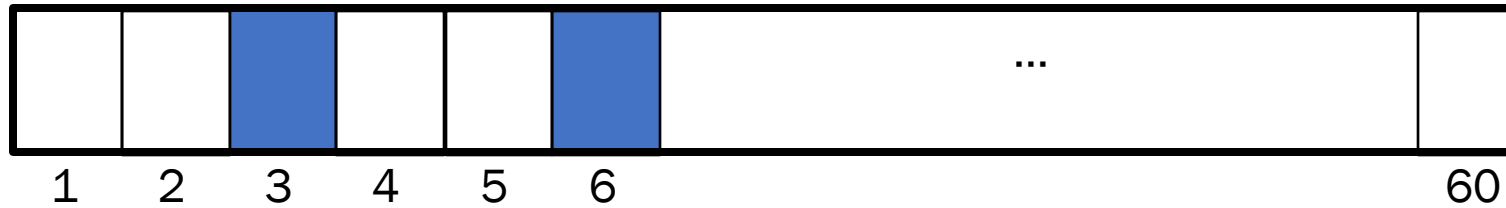
$$P(X = 3) = \binom{60}{3} (5/60)^3 (1 - 5/60)^{57}$$



# Probability of $k$ requests from this area in the next 1 min

On average  $\lambda = 5$  requests per minute

We can break the next minute down into seconds



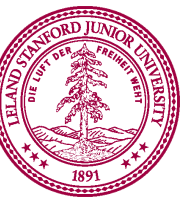
At each second either get a request or you don't.

Let  $X$  = Number of requests in the minute

$$X \sim \text{Bin}(n = 60, p = 5/60)$$

$$P(X = k) = \binom{n}{k} (p)^k (1 - p)^{n-k}$$

But what if there are two requests in the same second?



# Probability of $k$ requests from this area in the next 1 min

On average  $\lambda = 5$  requests per minute

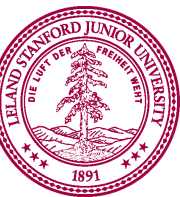
We can break that next minute down into *milli*-seconds



At each *milli*-second either get a request or you don't.

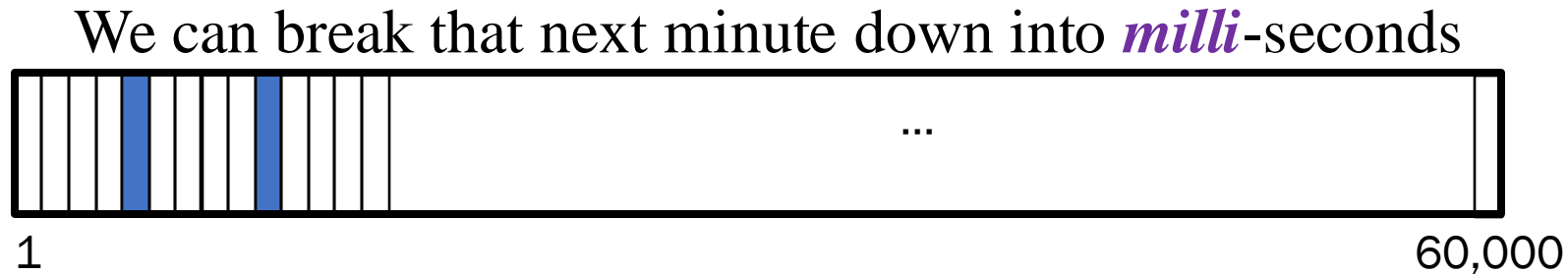
Let  $X$  = Number of requests in the minute

But what if there are two requests in the same second?



# Probability of $k$ requests from this area in the next 1 min

On average  $\lambda = 5$  requests per minute



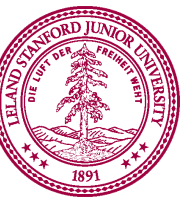
At each *milli*-second either get a request or you don't.

Let  $X$  = Number of requests in the minute

$$X \sim \text{Bin}(n = 60000, p = \lambda/n)$$

$$P(X = k) = \binom{n}{k} (\lambda/n)^k (1 - \lambda/n)^{n-k}$$

Can we do any better than milli-seconds?





# Probability of ***k* requests** from this area in the next 1 min

On average  $\lambda = 5$  requests per minute

We can break that minute down into *infinitely small* buckets

OMG so small

1

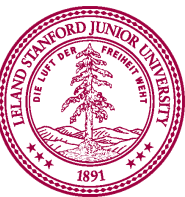
$\infty$

Let  $X$  = Number of requests in the minute

$$X \sim \text{Bin}(n, p = \lambda/n)$$

$$P(X = k) = \lim_{n \rightarrow \infty} \binom{n}{k} (\lambda/n)^k (1 - \lambda/n)^{n-k}$$

Who wants to see some cool math?



# Probability of **$k$ requests** from this area in the next 1 min

---

$$P(X = k) = \lim_{n \rightarrow \infty} \binom{n}{k} (\lambda/n)^k (1 - \lambda/n)^{n-k}$$

$$= \lim_{n \rightarrow \infty} \frac{n!}{(n-k)!k!} \cdot \frac{\lambda^k}{n^k} \cdot \frac{(1 - \lambda/n)^n}{(1 - \lambda/n)^k}$$

By expanding each term

$$= \lim_{n \rightarrow \infty} \frac{n!}{(n-k)!k!} \cdot \frac{\lambda^k}{n^k} \cdot \frac{e^{-\lambda}}{1}$$

By definition of natural exp

$$= \lim_{n \rightarrow \infty} \frac{n!}{(n-k)!n^k} \cdot \frac{\lambda^k}{k!} \cdot \frac{e^{-\lambda}}{1}$$

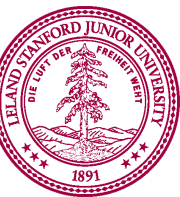
Rearranging terms

$$= \lim_{n \rightarrow \infty} \frac{n^k}{n^k} \cdot \frac{\lambda^k}{k!} \cdot \frac{e^{-\lambda}}{1}$$

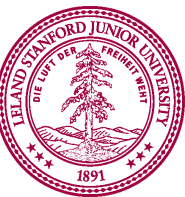
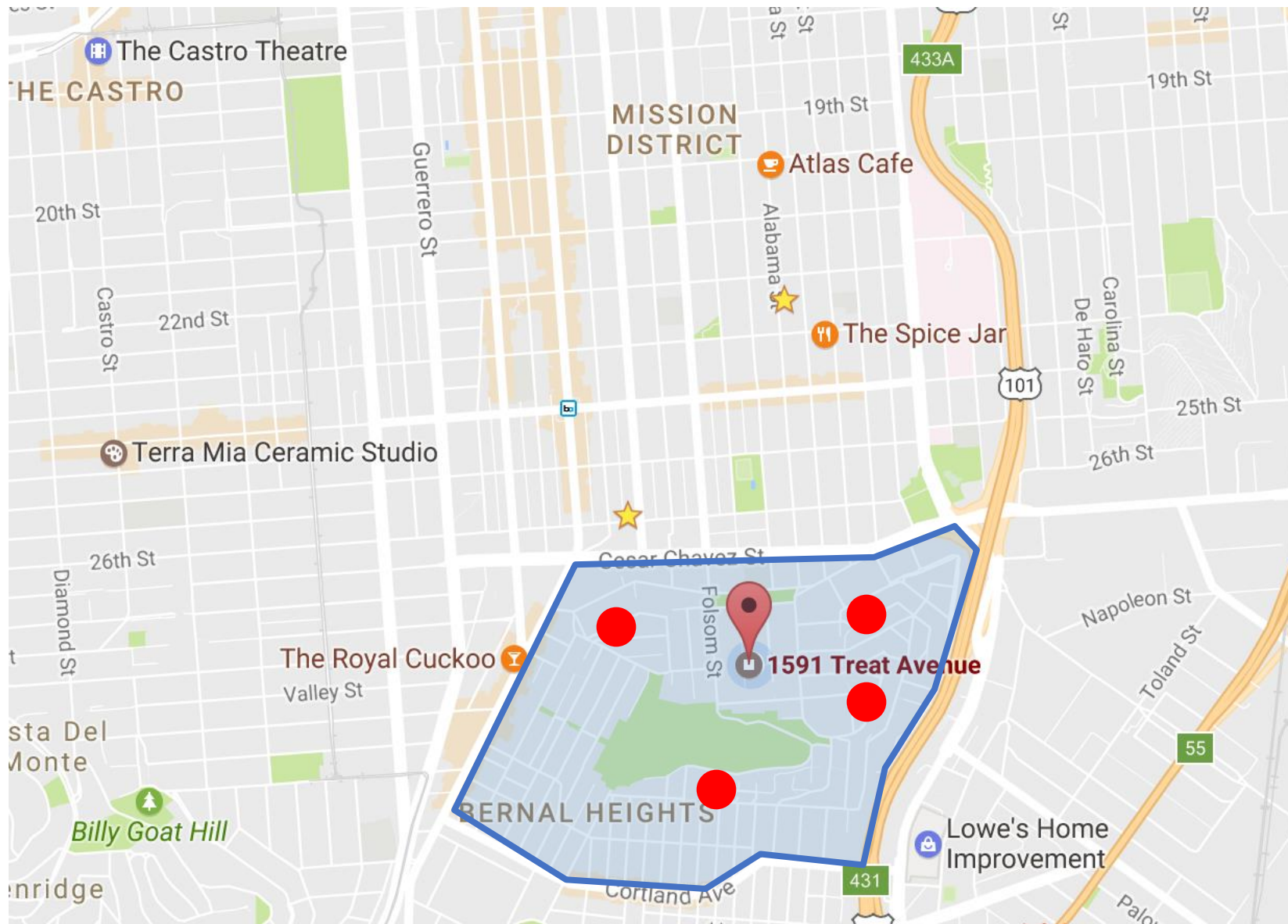
Limit analysis

$$= \frac{\lambda^k e^{-\lambda}}{k!}$$

Simplifying



# Probability of $k$ requests from this area in the next 1 min



# Simeon-Denis Poisson

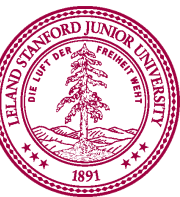
Simeon-Denis Poisson (1781-1840) was a prolific French mathematician



Published his first paper at 18, became professor at 21, and published over 300 papers in his life

- He reportedly said “*Life is good for only two things, discovering mathematics and teaching mathematics.*”

I’m going with French Martin Freeman



# Poisson Random Variable

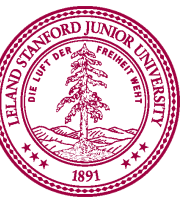
---

$X$  is a **Poisson** Random Variable: the number of occurrences in a fixed interval of time.

$$X \sim \text{Poi}(\lambda)$$

- $\lambda$  is the “rate”
- $X$  takes on values 0, 1, 2...
- has distribution (PMF):

$$P(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}$$



# Poisson Process

---

1

Consider events that occur over time

- Earthquakes, radioactive decay, hits to web server, etc.
- Have time interval for events (1 year, 1 sec, whatever...)
- Events arrive at rate:  $\lambda$  events per interval of time

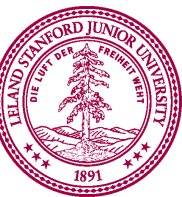
2

Split time interval into  $n \rightarrow \infty$  sub-intervals

- Assume at most one event per sub-interval
- Event occurrences in sub-intervals are independent
- With many sub-intervals, probability of event occurring in any given sub-interval is small

3

# events in original time interval  $\sim \text{Poi}(\lambda)$



# To the reader!

## Poisson Random Variable

**Notation:**  $X \sim \text{Poi}(\lambda)$

**Description:** Number of events in a fixed time frame if (a) the events occur with a constant mean rate and (b) they occur independently of time since last event.

**Parameters:**  $\lambda \in \{0, 1, \dots\}$ , the constant average rate.

**Support:**  $x \in \{0, 1, \dots\}$

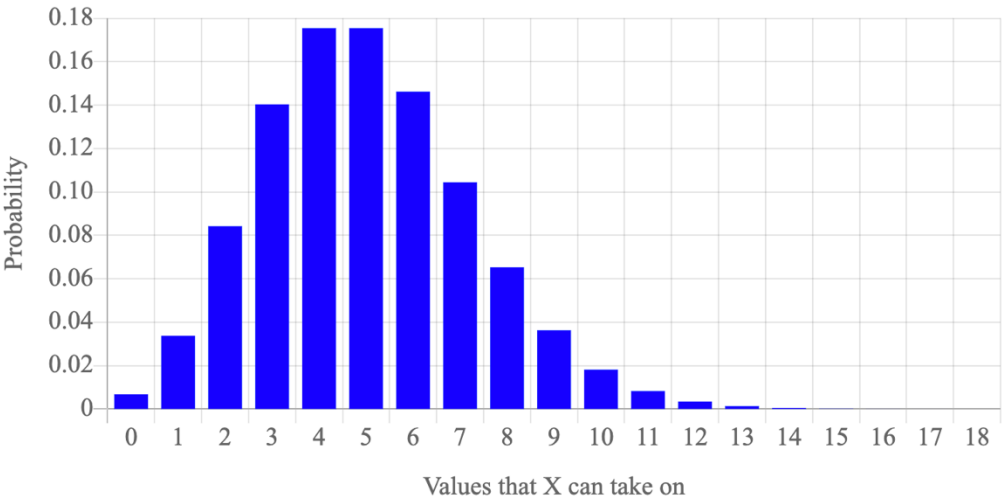
**PMF equation:**  $\Pr(X = x) = \frac{\lambda^x e^{-\lambda}}{x!}$

**Expectation:**  $E[X] = \lambda$

**Variance:**  $\text{Var}(X) = \lambda$

**PMF graph:**

Parameter  $\lambda$ :





Poisson is great when you  
have a rate!

---





Poisson is great when you  
have a rate and you care  
about # of occurrences!

---



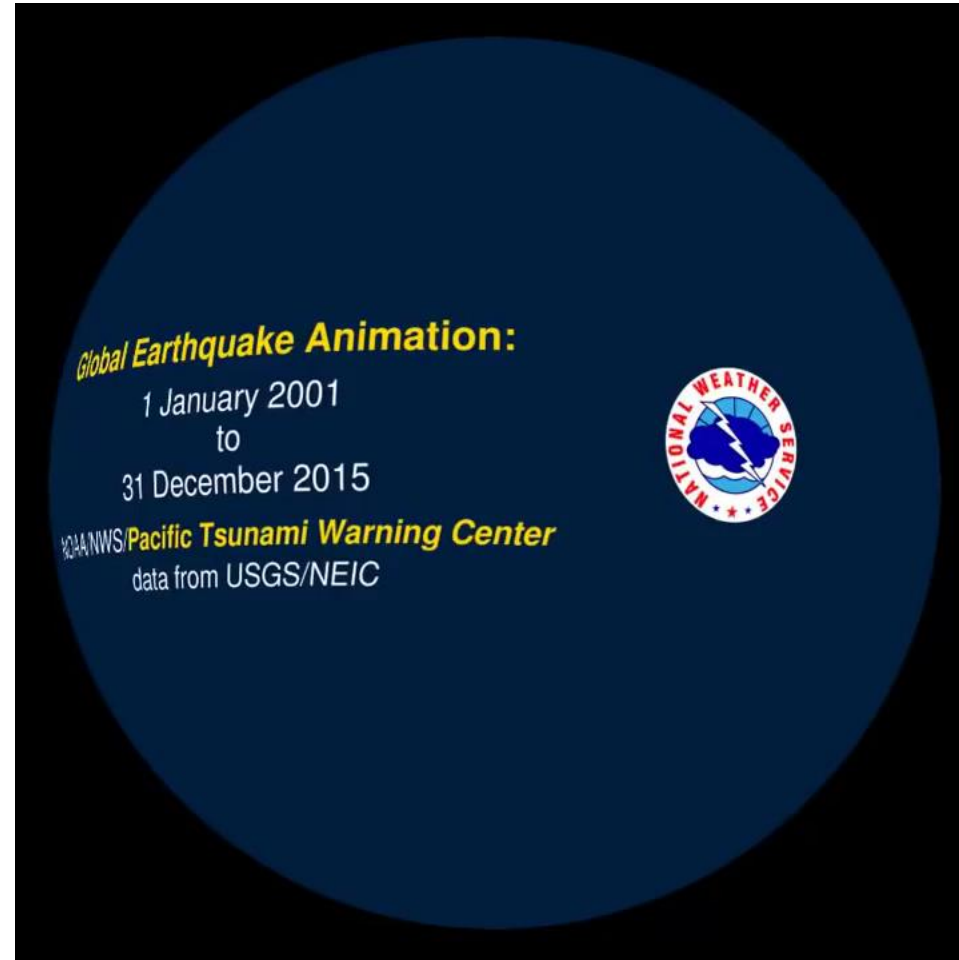
Make sure that the time unit for “rate” and match the probability question

---

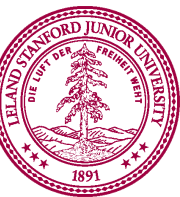
Two quick examples!

# Earthquakes

---



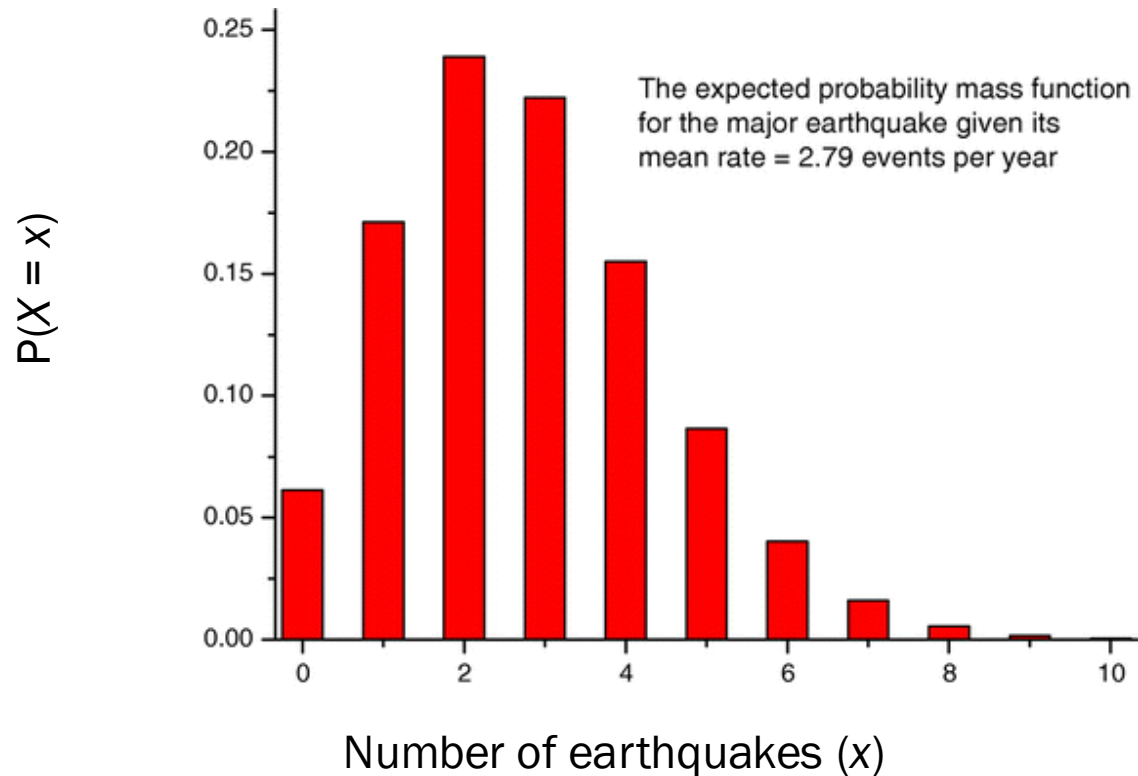
Average of 2.79 major earthquakes per year.  
What is the probability of 3 major earthquakes next year?



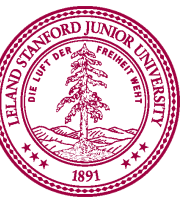
# Earthquake Probability Mass Function

Let  $X$  = number of earthquakes next year

$$X \sim \text{Poi}(2.79)$$



$$P(X = 3) = \frac{\lambda^k e^{-\lambda}}{k!} = \frac{2.79^3 e^{-2.79}}{3!} \approx 0.23$$



## Bulletin of the Seismological Society of America

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Vol. 64

October 1974

No. 5

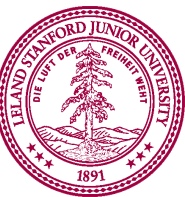
---

IS THE SEQUENCE OF EARTHQUAKES IN SOUTHERN CALIFORNIA,  
WITH AFTERSHOCKS REMOVED, POISSONIAN?

BY J. K. GARDNER and L. KNOPOFF

ABSTRACT

**Yes.**



# Fruit in Class!

Students ask on average **15 questions per class**.

Chris only brought **10 mandarins**!

What is the probability he runs out of fruit?

\* Assume: (a) question rate is constant (b) questions don't impact one another.



Let  $X$  be the number of questions asked in class.  $X \sim \text{Poi}(\lambda = 15)$

$$\begin{aligned} P(X \leq 10) &= \sum_{i=0}^{10} P(X = i) \\ &= \sum_{i=0}^{10} \frac{\lambda^i e^{-\lambda}}{i!} \quad \text{PMF of Poisson} \\ &= \sum_{i=0}^{10} \frac{15^i e^{-15}}{i!} \quad \lambda = 15 \end{aligned}$$

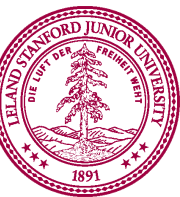
```
from scipy import stats
def main():
    lam = int(input("Questions per class: "))
    num_fruit = int(input("Number of fruits: "))
    X = stats.poisson(lam)
    prob_enough_fruit = 0
    for i in range(0, num_fruit+1):
        pr_i_questions = X.pmf(i)
        prob_enough_fruit += pr_i_questions
    print(prob_enough_fruit)
```

# Poisson in Python

---

```
from scipy import stats # great package
X = stats.poisson(2.5) #  $X \sim \text{Poi}(\lambda = 2.5)$ 
print(X.pmf(2))      #  $P(X = 2)$ 
```

Function	Description
<code>X.pmf(k)</code>	$P(X = k)$
<code>X.cdf(k)</code>	$P(X \leq k)$
<code>X.mean()</code>	$E[X]$
<code>X.var()</code>	$\text{Var}(X)$
<code>X.std()</code>	$\text{Std}(X)$





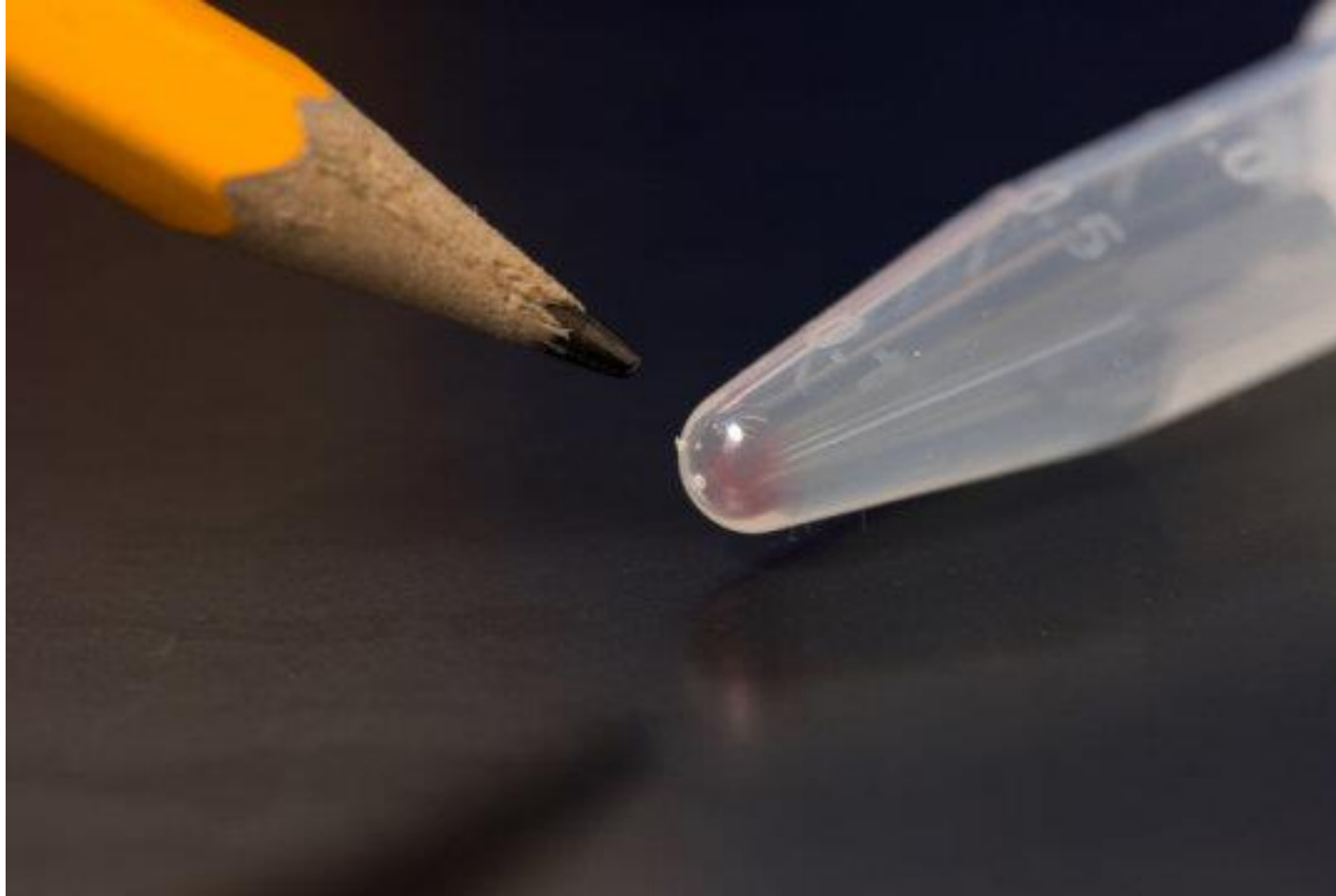
# Poisson can approximate a Binomial!

Wait why would you want to do that?

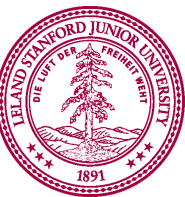
- 1) Binomial can be expensive to compute.
- 2) Connections help build math intuition.

# Storing Data in DNA

---



All the movies, images, emails and other digital data from more than 600 smartphones (10,000 gigabytes) can be stored in the faint pink smear of DNA at the end of this test tube.



# Storing Data in DNA

---

Will more than 1% of DNA storage become corrupt?

- In DNA (and real networks) store large strings
- Length  $n \approx 10^4$
- Probability of corruption of each base pair is very small  $p \approx 10^{-6}$
- $X \sim \text{Bin}(10^4, 10^{-6})$  is unwieldy to compute

Extreme  $n$  and  $p$  values arise in many cases

- # bit errors in stream sent over a network
- # of servers crashes in a day in giant data center



# Storing Data in DNA

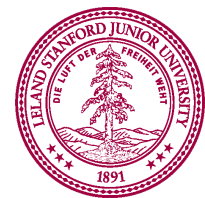
---

Will the DNA storage become corrupt?

- In DNA (and real networks) store large strings
- Length  $n \approx 10^4$
- Probability of corruption of each base pair is very small  $p \approx 10^{-6}$
- $X \sim \text{Poi}(\lambda = 10^4 * 10^{-6} = 0.01)$

$$P(X = k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

$$\begin{aligned} P(X = 0) &= e^{-\lambda} \frac{1}{0!} \\ &= e^{-0.01} \approx 0.99 \end{aligned}$$



# Poisson is a Binomial in the Limit

---

Poisson approximates Binomial where  $n$  is large,  $p$  is small, and  $\lambda = np$  is “moderate”

Different interpretations of "moderate"

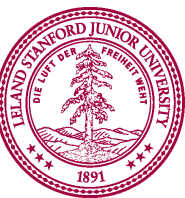
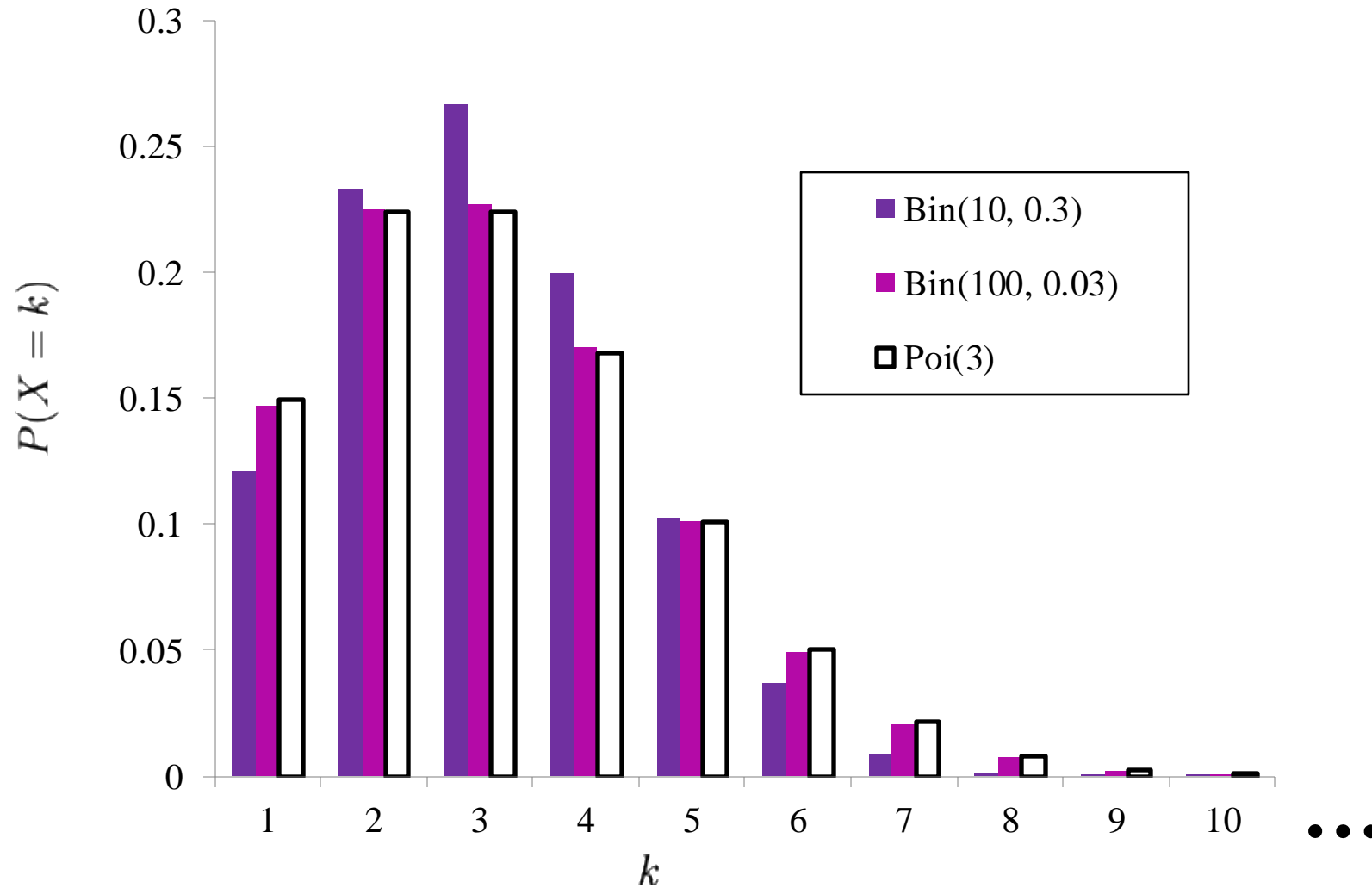
- $n > 20$  and  $p < 0.05$
- $n > 100$  and  $p < 0.1$

Really, Poisson is Binomial as

$$n \rightarrow \infty \text{ and } p \rightarrow 0, \text{ where } np = \lambda$$



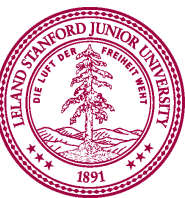
# Bin(10,0.3) vs Bin(100,0.03) vs Poi(3)



# A Real License Plate Seen at Stanford



No, it's not mine...  
but I kind of wish it was.







Poisson can be used  
to approximate a  
Binomial where  $n$  is  
large and  $p$  is small.

---

# Tender (Central) Moments with Poisson

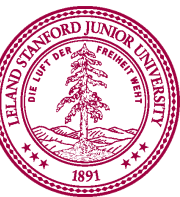
---

Recall:  $Y \sim \text{Bin}(n, p)$

- $E[Y] = np$
- $\text{Var}(Y) = np(1 - p)$

$X \sim \text{Poi}(\lambda)$  where  $\lambda = np$  ( $n \rightarrow \infty$  and  $p \rightarrow 0$ )

- $E[X] = np = \lambda$
- $\text{Var}(X) = np(1 - p) = \lambda(1 - 0) = \lambda$
- Yes, expectation and variance of Poisson are same
- It brings a tear to my eye...



# Poisson Paradigm

---

Poisson can still provide a good way to model an event, even when assumptions are “mildly” violated.  
Can apply Poisson approximation when...

1

**“Successes” in trials are not entirely independent.**

- Example: # entries in each bucket in large hash table

2

**Probability of “Success”  $p$  in each trial varies slightly.**

- Example: average # requests to web server/sec. may fluctuate slightly due to load on network



# Web Server Load

Consider requests to a web server in 1 second

- In past, server load averages 2 hits/second
- $X = \#$  hits server receives in a second
- What is  $P(X < 5)$ ?

Solution

$$X \sim \text{Poi}(\lambda = 2)$$

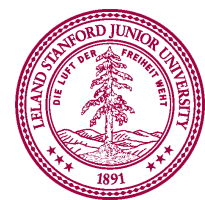
$$P(X < 5) = \sum_{i=0}^4 P(X = i)$$

$$= \sum_{i=0}^4 e^{-\lambda} \frac{\lambda^i}{i!}$$

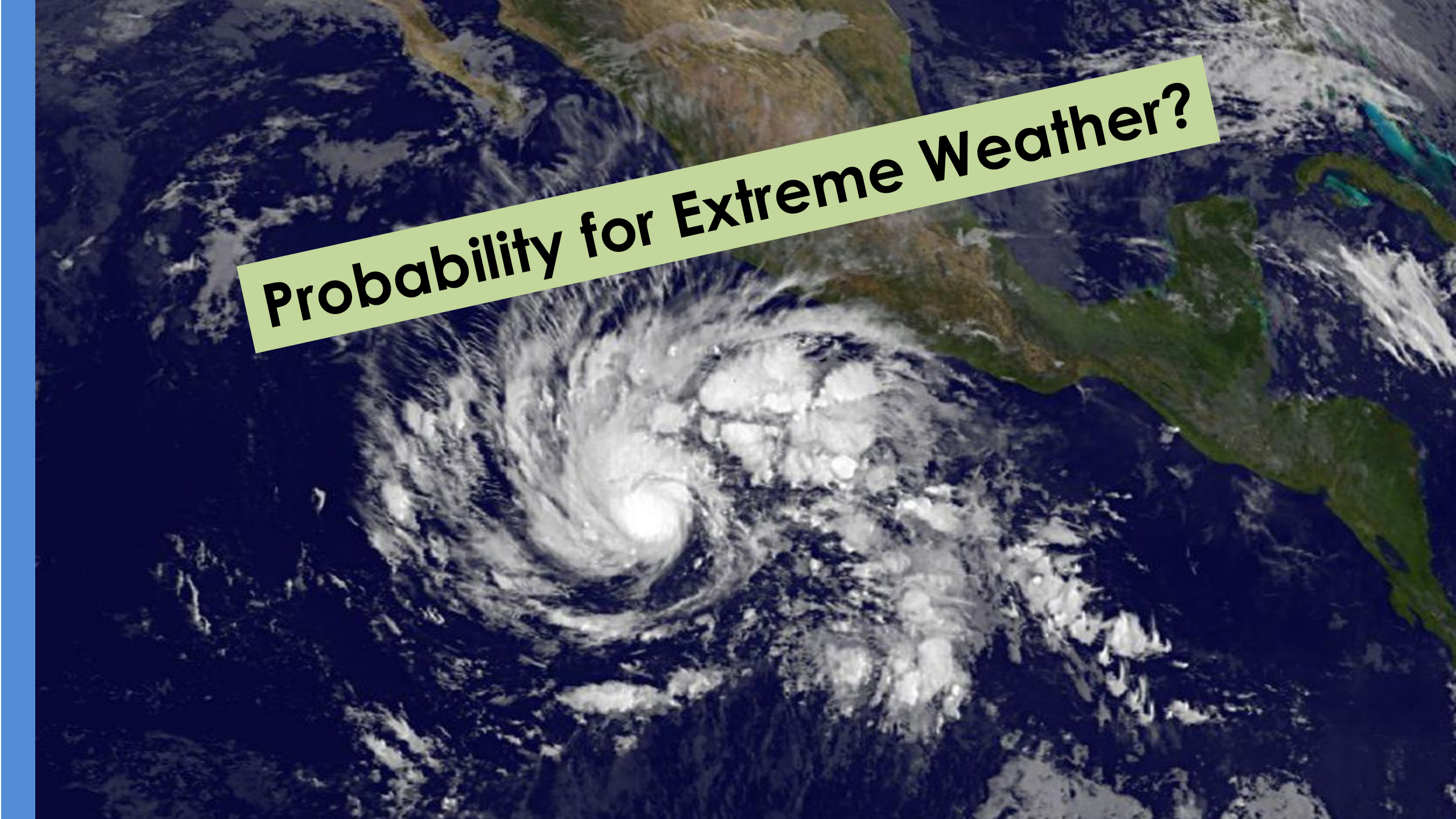
Since  $X$  is Poisson

$$= \sum_{i=0}^4 e^{-2} \frac{2^i}{i!} \approx 0.95$$

Since  $\lambda = 2$

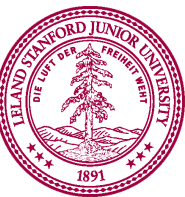
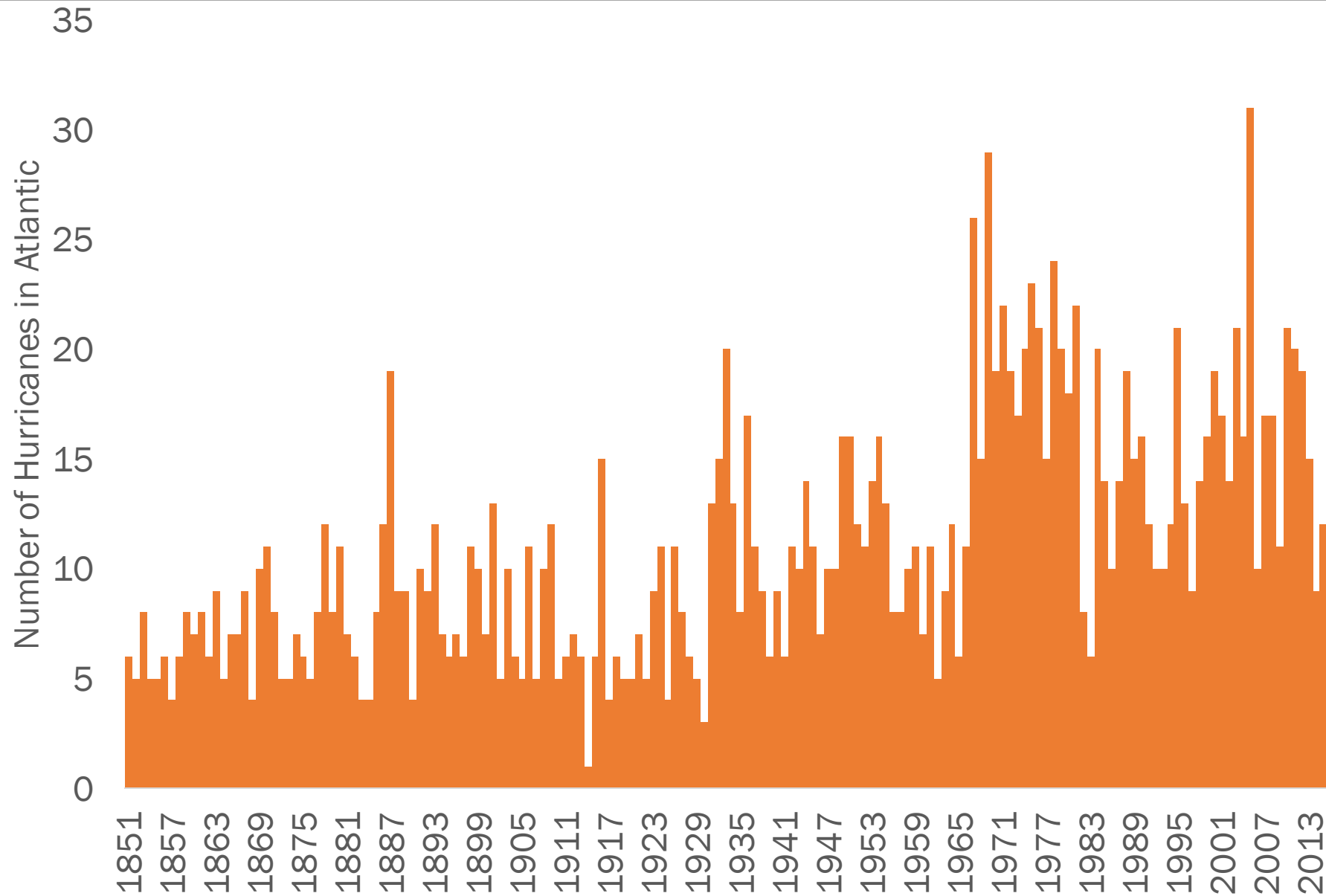




A satellite image of a hurricane over the Gulf of Mexico, with the text "Probability for Extreme Weather?" overlaid in a yellow banner. The hurricane shows a clear eye and dense cloud structure. The surrounding ocean is dark blue, and the landmasses of North and Central America are visible in shades of green and brown.

**Probability for Extreme Weather?**

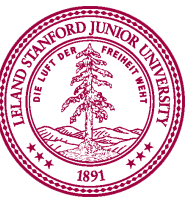
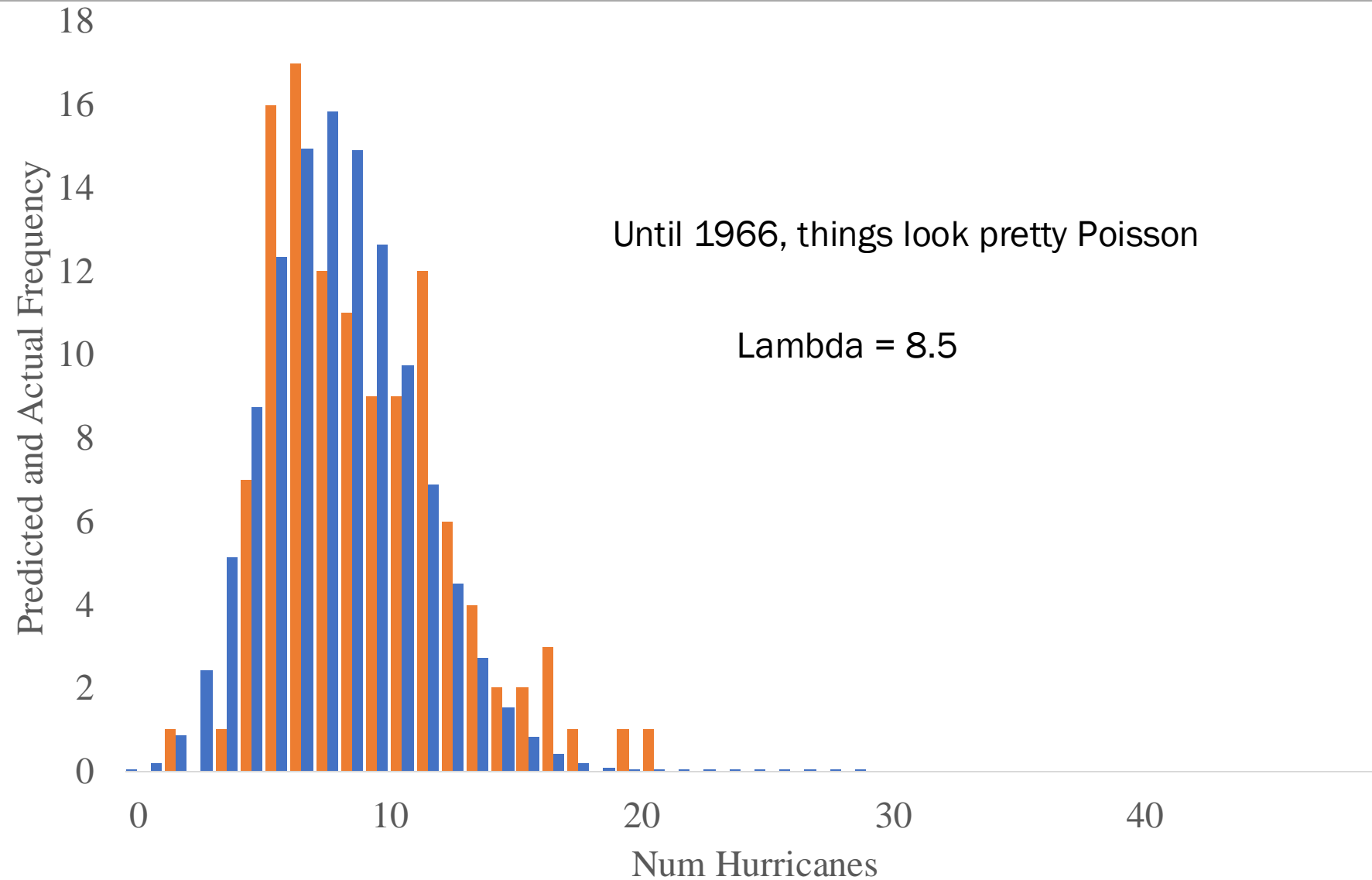
# Hurricanes per Year since 1851



To the code!



# Historically ~ Poisson(8.5)



# Improbability Drive

What is the probability of over 15 hurricanes in a season given that the distribution doesn't change?

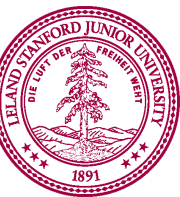
- Let  $X$  = # hurricanes in a year.  $X \sim \text{Poi}(8.5)$

Solution:

$$\begin{aligned} P(X > 15) &= 1 - P(X \leq 15) \\ &= 1 - \sum_{i=0}^{15} P(X = i) \end{aligned}$$

This is the pmf of a Poisson. Your favorite programming language has a function for it

$$= 0.0135$$



Twice since 1966 there have been two  
years with over 30 hurricanes

# Improbability Drive

What is the probability of over 30 hurricanes in a season given that the distribution doesn't change?

- Let  $X$  = # hurricanes in a year.  $X \sim \text{Poi}(8.5)$

Solution:

$$P(X > 30) = 1 - P(X \leq 30)$$

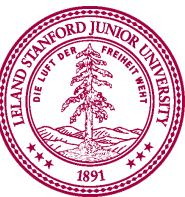
$$= 1 - \sum_{i=0}^{30} P(X = i)$$

$$= 1 - 0.9999999997823$$

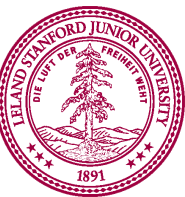
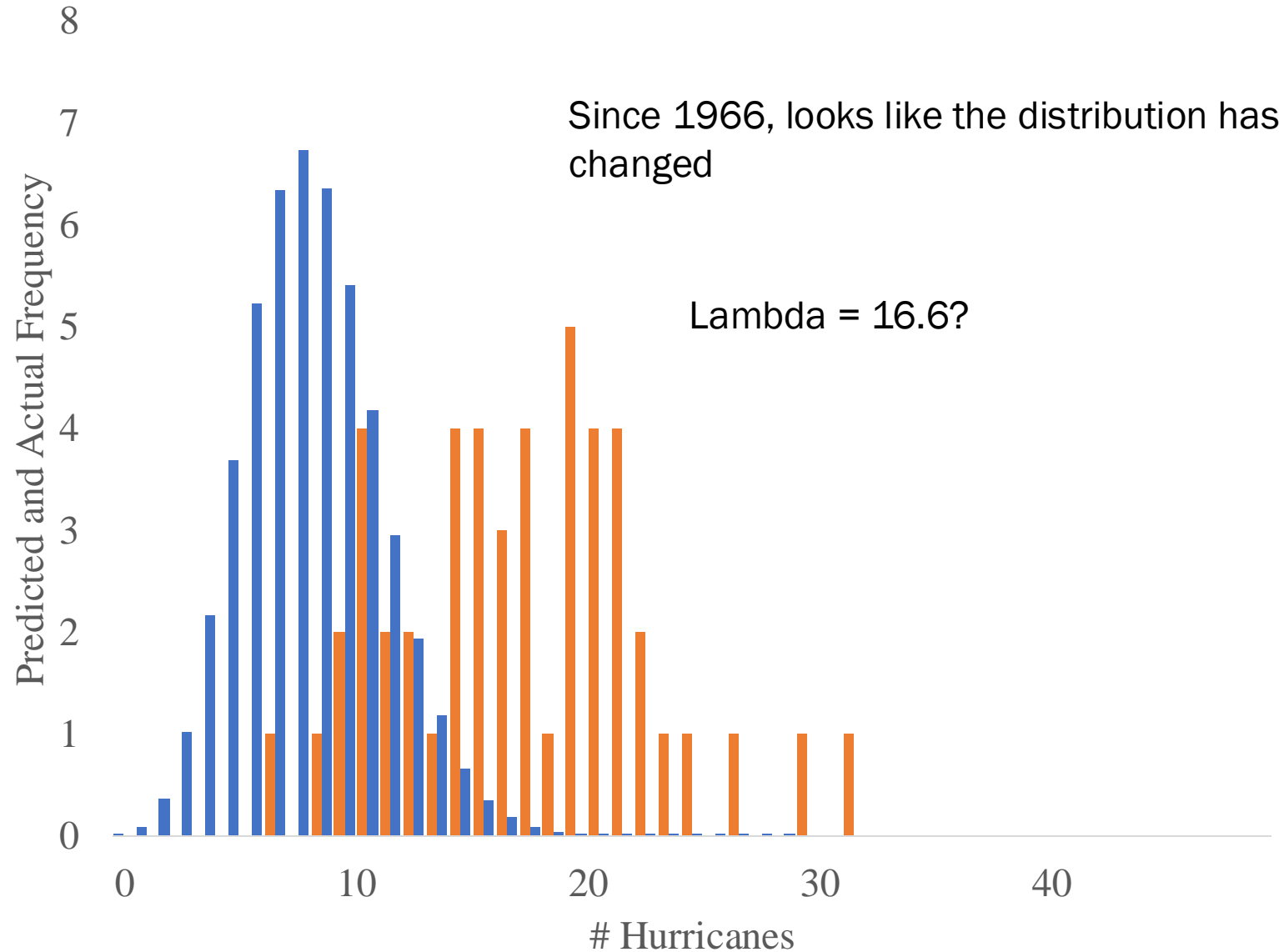
$$= 2.2e - 09$$

This is the pdf of a Poisson. Your favorite programming language has a function for it

\* Challenge: Calculate the probability of two years with over 30 hurricanes

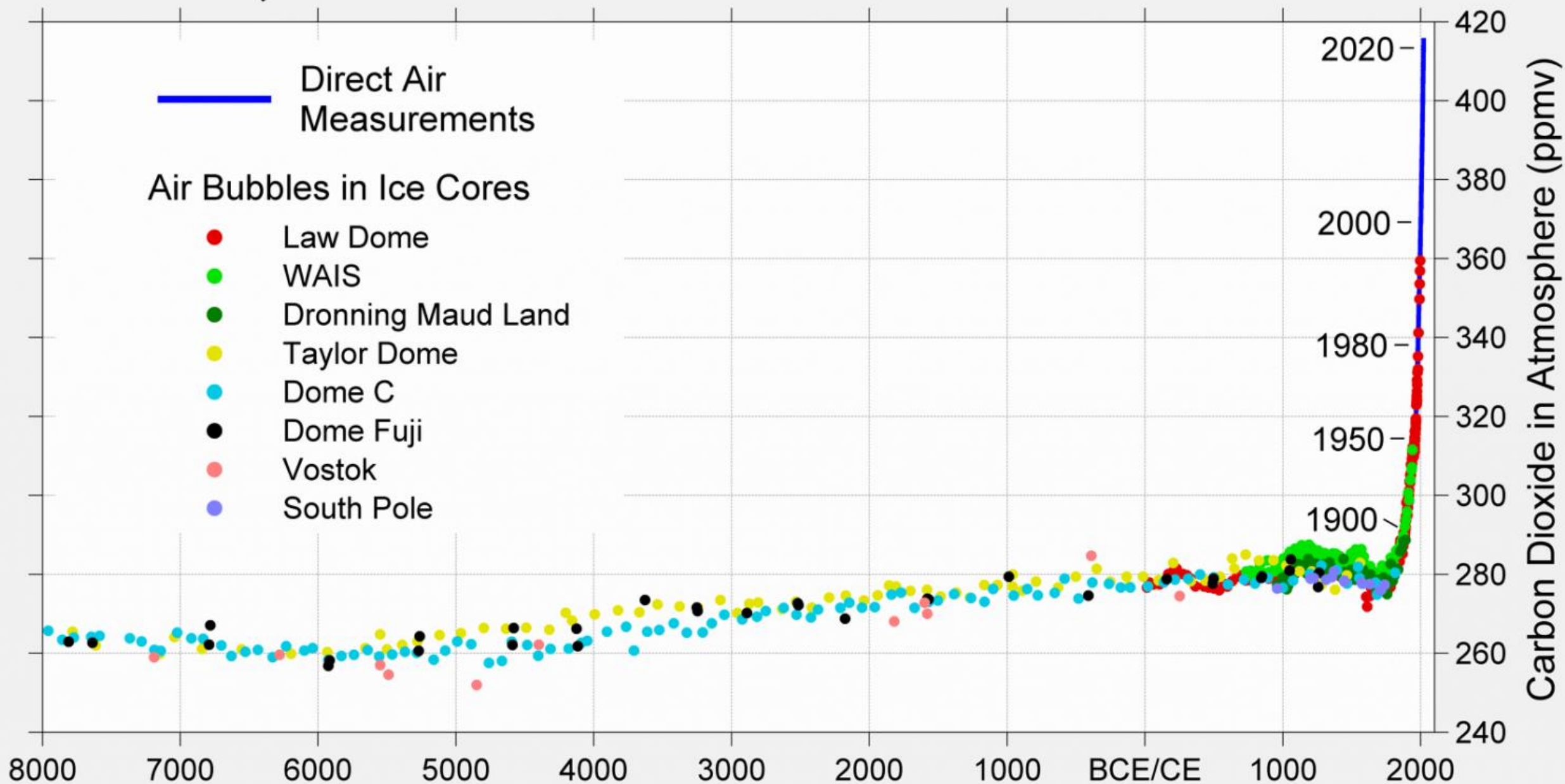


# The Distribution has Changed



What's up?

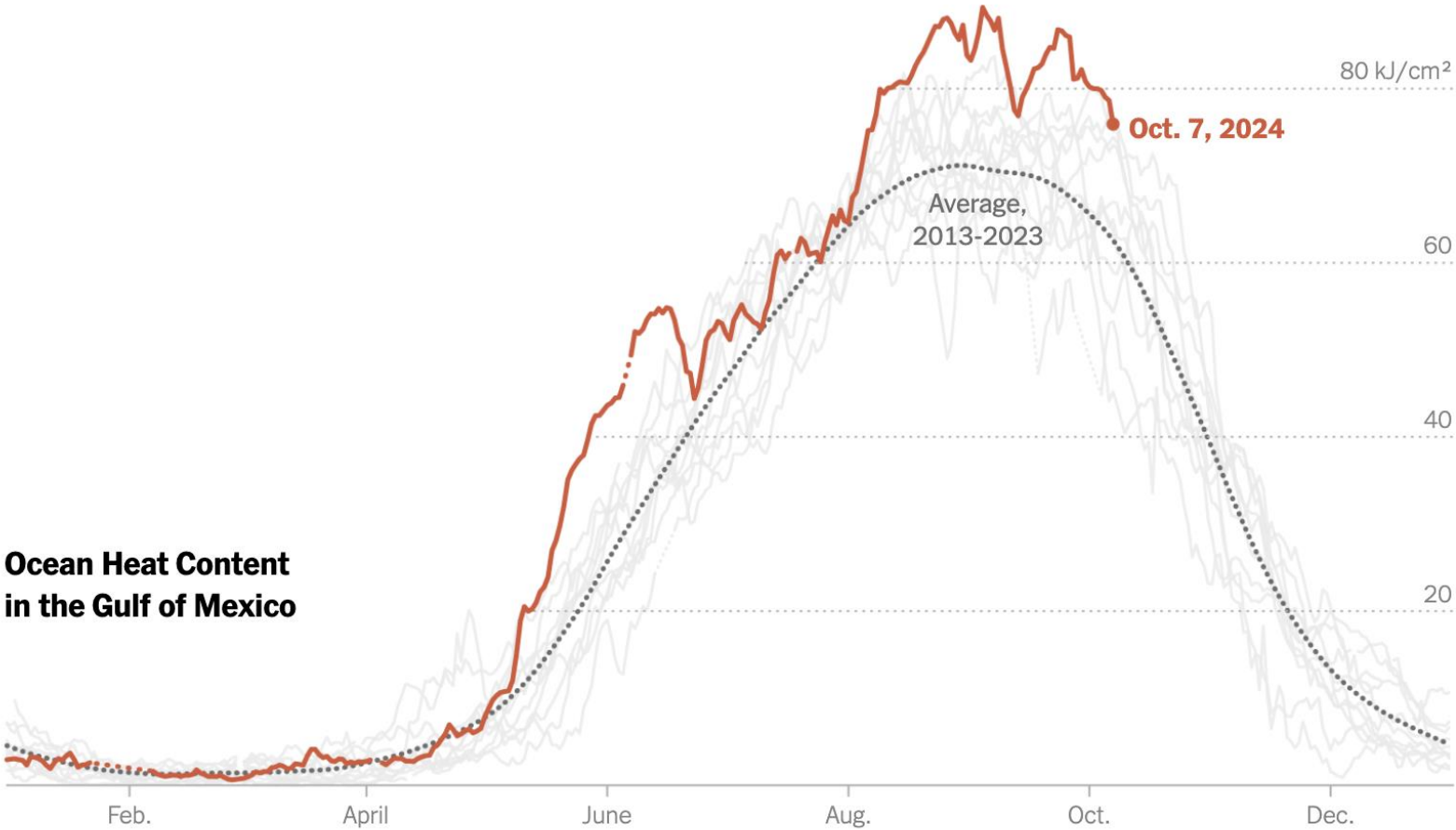
# 10,000 Years of Carbon Dioxide





# CO2 leads to Hotter Oceans

Ocean Heat Content  
in the Gulf of Mexico



Source: Brian McNoldy; University of Miami Upper Ocean Dynamics Lab



# What's Up?



Next Time



# Discrete Distributions

## **Bernoulli:**

- indicator of coin flip  $X \sim \text{Ber}(p)$

## **Binomial:**

- # successes in  $n$  coin flips  $X \sim \text{Bin}(n, p)$

## **Poisson:**

- # successes in  $n$  coin flips  $X \sim \text{Poi}(\lambda)$

## **Geometric:**

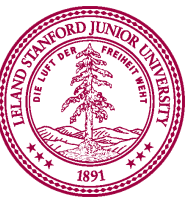
- # coin flips until success  $X \sim \text{Geo}(p)$

## **Negative Binomial:**

- # trials until  $r$  successes  $X \sim \text{NegBin}(r, p)$

## **Zipf:**

- The popularity rank of a random word, from a natural language
- $X \sim \text{Zipf}(s)$



# The Poisson Common Path

