

Code ▾

Statistical methods for de novo mutation analysis

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This note describes some approaches for gene-level de novo mutation (DNM) analysis. We will use the dataset from [1]. To demonstrate these methods, only loss-of-function DNMs (dn.LoF) are used.

This is a draft, there would be errors/typos inside the note. We will update this note, and will also add current methods.

We will summarize two tests: the Poisson test, and a Bayesian test (a Mixture-model based approach).

Dataset

We will use the dataset from De Rubeis et al., (2015) [1]

[1]. <https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4402723/> (<https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4402723/>)

- Read the data frame into R.

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```
x <- read.table("test_Data_from_ASD.txt", header = TRUE, as.is = TRUE)
##https://www.nature.com/articles/nature13772#Sec9
Ntrio = 3871
```

- Print some lines of this dataset

Hide

```
head(x)
```

	Gene <chr>	mut.rate <dbl>	dn.LoF <int>
1	SCN2A	0.00007400	4
2	SYNGAP1	0.00006600	5
3	CHD8	0.00009200	3
4	ARID1B	0.00009000	4
5	ANK2	0.00014923	3
6	SUV420H1	0.00003500	3
6 rows			

Statistical analysis

Poisson test

Let q, μ be the gene-level DNM count and the mutation rate of the tested gene, and let N_{trio} be the number of trios/families. Let X be a random variable denoting DNM counts. As described in [1], X follows a Poisson distribution with mean = $2 * N_{trio} * \mu$, and the p-value is calculated as $P(X \geq q)$.

Usually, there are only some DNMs for a gene, we can write a simple function to calculate the p-value by using $P(X \geq q) = 1 - P(X < q)$. However, the function `ppois` in R can be used to calculate the p-values.

[1] Ware JS, Samocha KE, Homsy J, Daly MJ. Interpreting de novo Variation in Human Disease Using denovolyzeR. Curr Protoc Hum Genet. 2015 Oct 6;87:7.25.1-7.25.15. doi: 10.1002/0471142905.hg0725s87. PMID: 26439716; PMCID: PMC4606471. <https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4606471/> (<https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4606471/>).

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```
x.poisson <- ppois(q = x$dn.LoF - 1, ##observed number - 1 (X - 1)
                  lambda = 2*Ntrio*x$mut.rate, ##expected
                  lower.tail = FALSE)
```

Take a look at the result.

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```
x1 <- data.frame(x, x.poisson = x.poisson)
head(x1[x.poisson < 0.05, ])
```

	Gene <chr>	mut.rate <dbl>	dn.LoF <int>	x.poisson <dbl>
1	SCN2A	7.4e-05	4	0.0028513034
2	SYNGAP1	6.6e-05	5	0.0001901230
3	CHD8	9.2e-05	3	0.0356506317
4	ARID1B	9.0e-05	4	0.0056625306
6	SUV420H1	3.5e-05	3	0.0027099340
7	DYRK1A	3.1e-05	4	0.0001141829
6 rows				

For this test, we can remove genes with DNM counts = 0 (p-value = 1).

Mixture model

Another approach is to use a mixture model. Based on how many hypotheses (n) you plan to test, you can build a mixture model of n distributions. There are some published methods for this approach, we will review a simple model and go through some of these methods.

Simple mixture model.

We can compare two hypotheses: the gene is a risk gene (H_1) and the gene is not a risk gene (H_0)[2]. Let M_1 and M_0 be the two models for the two hypotheses. We assume that X follows two Poisson distributions for the two models (If we can find better distributions, we can use those distributions.).

[2] Statistical approaches only help to prioritize genes. To better understand the prioritized genes, other approaches (e.g., wet-lab/dry-lab approaches + results from independent datasets) should be used in downstream analyses.

For example, we can model X as follows:

Hypothesis	Model	Distribution
H_1	M_1	$X \sim \text{Poisson}(2 * N_{\text{trio}} * \mu * \gamma)$
H_0	M_0	$X \sim \text{Poisson}(2 * N_{\text{trio}} * \mu)$
—	—	—

The parameter for M_0 is the same as the Poisson test above ($2 * N_{\text{trio}} * \mu$). We assume another parameter for M_1 : $2 * N_{\text{trio}} * \mu * \gamma$. **The question here is how to find γ to compare two hypotheses.**

If you have any reliable information for γ , you can use that information. However, a simple way is to find γ from the tested dataset by using a likelihood function.

We will describe more for this part. However, you can write your own approach by reading mixture-model lectures.

TADA uses a similar approach, but add a prior for γ .

TADA method

We will calculate posterior probabilities (PPs) for genes from this method, and then compare with the Poisson test above.

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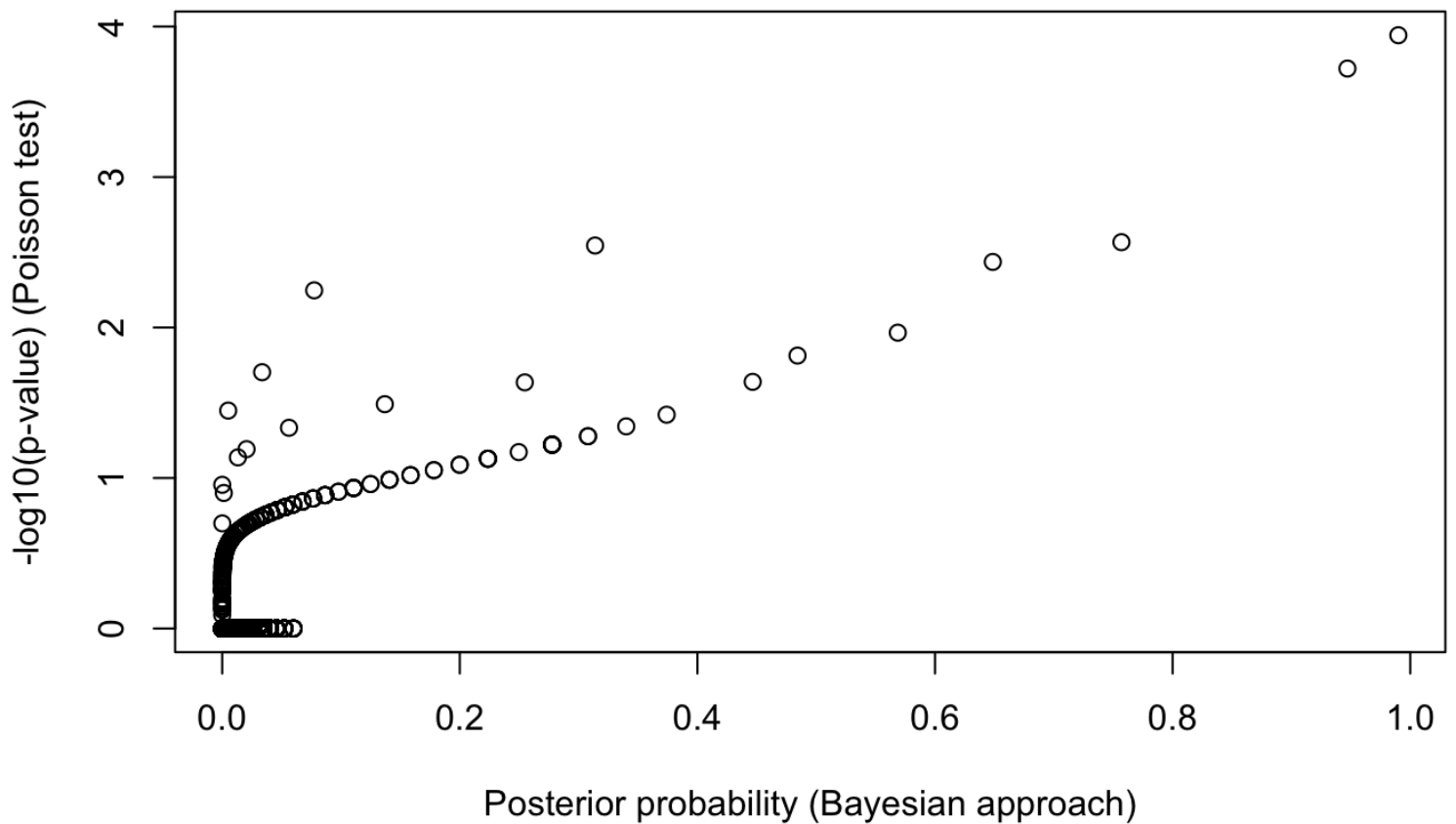
```
source("TADA.R")
gamma0 = 20
beta0 = 1
pi0 = 0.06
x.bf = bayes.factor.denovo(x$dn.LoF, N = Ntrio,
                           mu = x$mut.rate,
                           gamma.mean = gamma0, beta = beta0)
##Convert Bayes Factors to PPs

x.pp <- pi0*x.bf/(1 - pi0 + pi0*x.bf)
```

Compare results from a Poisson test and a Bayesian approach

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```
plot(x.pp, -log(x.poisson, base = 10), xlab = 'Posterior probability (Bayesian approach)',
     ylab = '-log10(p-value) (Poisson test)')
```



Questions and answers

Do the Poisson test and a Bayesian approach generate different results?

A Bayesian method's results rely on its priors. Frequently, results are not much different if you use prior information from the tested data.

We would suggest that you take a look at a Poisson test described in [1] to better understand this test.

[1] Ware JS, Samocha KE, Homsy J, Daly MJ. Interpreting de novo Variation in Human Disease Using denovolyzeR. *Curr Protoc Hum Genet*. 2015 Oct 6;87:7.25.1-7.25.15. doi: 10.1002/0471142905.hg0725s87. PMID: 26439716; PMCID: PMC4606471 <https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4606471/> (<https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4606471/>).

For the TADA approach, can we use a different distribution?

Yes, of course. The Gamma distribution is a simple prior which can help us analytically calculate the marginal probability of the data. For example, in TADA, an analytic approach is used. However, you can use numerical integration to calculate Bayes Factors.

Example

In TADA, the marginal distribution is a negative binomial distribution. **If you are not familiar with *Calculus 1***, you can use **numerical integration** in R or other programming languages.

Note: Numerical integration is an approximation, and might not be the same as analytical integration.

In the TADA model, $X \sim \text{Poisson}(2 * N_{\text{trio}} * \mu * \gamma)$ and $\gamma \sim \text{Gamma}(\bar{\gamma} * \beta, \beta)$, we can create a *Gamma* function in R.

Below is an approach using numerical integration. We will also compare this approach with the negative binomial distribution based approach of TADA.

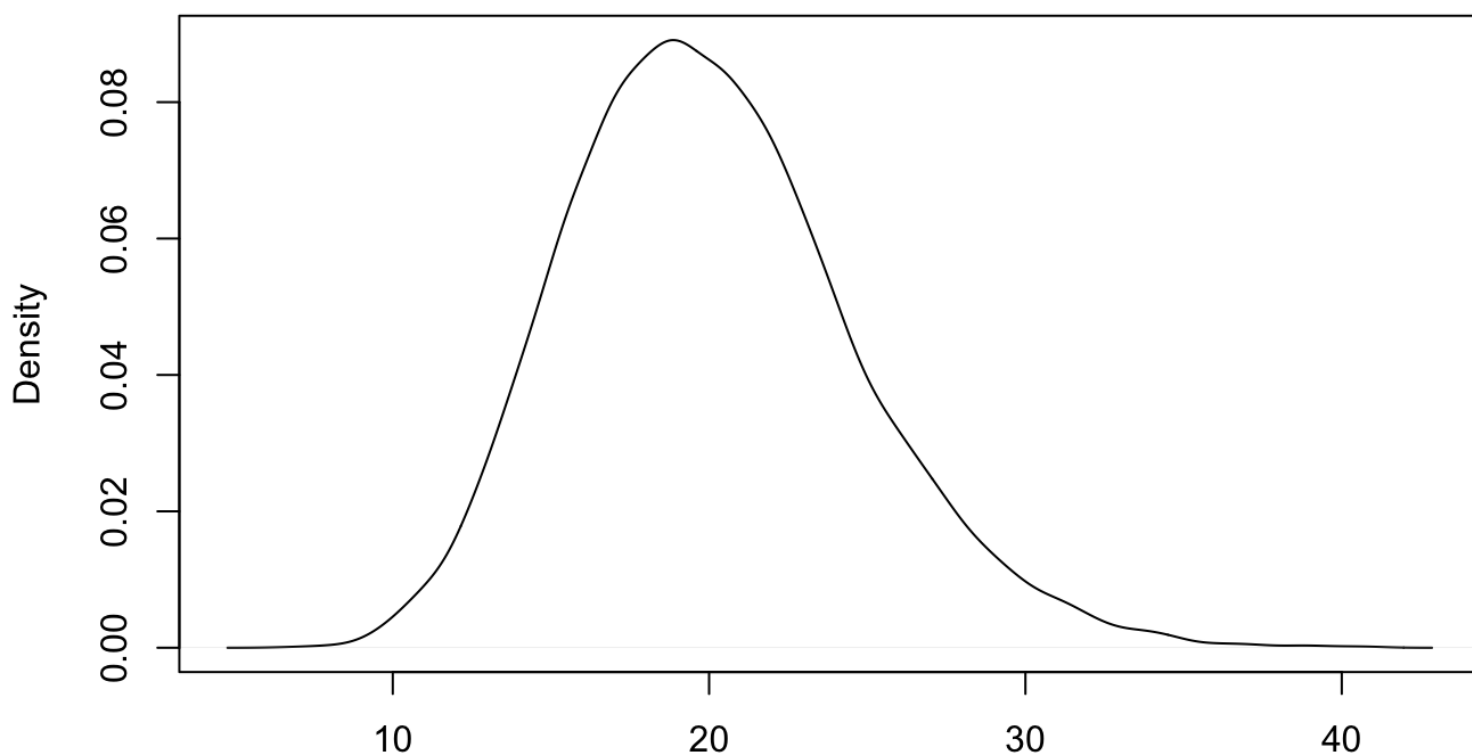
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```
gamma0 = 20
beta0 = 1
f1 <- function(x.gamma){dgamma(x.gamma, shape = gamma0*beta0, rate = beta0)}
```

and have a look at its density distribution with parameters from the paper.

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```
xGamma <- rgamma(10000, shape = gamma0*beta0, rate = beta0)
plot(density(xGamma), main = '', xlab = '')
```



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```
gammaMin = min(xGamma)
gammaMax = max(xGamma)
```

Now, we can calculate Bayes Factors $\frac{P(X|H_1)}{P(X|H_0)}$ in which $P(X|H_1 = \int Poisson(X|2 * Ntrio * \mu * \gamma)Gamma(\gamma|\bar{\gamma} * \beta, \beta)d\gamma$ and $P(X|H_0) = Poisson(X|2 * Ntrio * \mu)$.

- We test for $X = 0$.

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```
xdn = 0
(dpois(xdn, lambda = 2*Ntrio*10^-6*gamma0)*integrate(f1, gammaMin, gammaMax, subdivisions = 200L)
$value)/dpois(xdn, lambda = 2*Ntrio*10^-6)
```

```
[1] 0.8630973
```

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```
x.bf.test = bayes.factor.denovo(xdn, N = Ntrio,
                                mu = 10^-6,
                                gamma.mean = gamma0, beta = beta0)

x.bf.test
```

```
[1] 0.8637243
```

- We test for $X = 2$.

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```
xdn = 2
(dpois(xdn, lambda = 2*Ntrio*10^-6*gamma0)*integrate(f1, gammaMin, gammaMax, subdivisions = 1000L)
$value)/dpois(xdn, lambda = 2*Ntrio*10^-6)
```

```
[1] 345.2389
```

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```
x.bf.test = bayes.factor.denovo(xdn, N = Ntrio,
                                mu = 10^-6,
                                gamma.mean = gamma0, beta = beta0)

x.bf.test
```

```
[1] 357.2117
```

- We test for all genes from the dataset in [1].

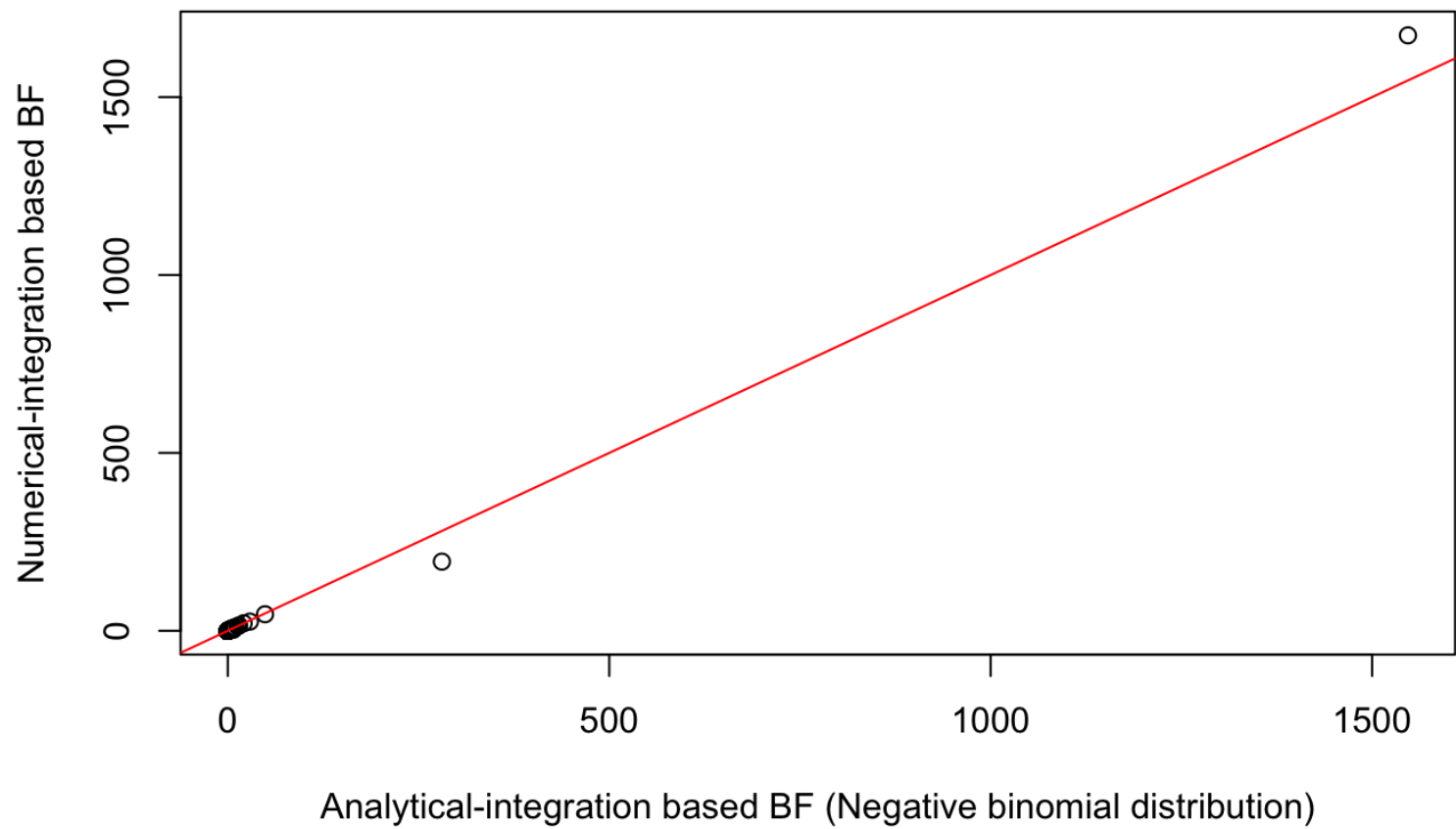
Hide

```
x.bf.numericalIn <- apply(data.frame(x$mut.rate, x$dn.LoF), 1, function(y){
  (dpois(y[2], lambda = 2*Ntrio*y[1]*gamma0)*integrate(f1, gammaMin, gammaMax, subdivisions = 200L)
  )$value)/dpois(y[2], lambda = 2*Ntrio*y[1])
})
```

- Compare between the two methods.

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```
plot(x.bf, x.bf.numericalIn, xlab = 'Analytical-integration based BF (Negative binomial distributi
on)', ylab = 'Numerical-integration based BF')
abline(a = 0, b = 1, col = 'red')
```



Hide

```
x[x.bf > 250, ]
```

	Gene<chr>	mut.rate<dbl>	dn.LoF<int>
2	SYNGAP1	6.6e-05	5
7	DYRK1A	3.1e-05	4
2 rows			

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NA

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```
pT = 0.001
x[x.poisson < pT, ]
```

	Gene<chr>	mut.rate<dbl>	dn.LoF<int>
2	SYNGAP1	6.6e-05	5
7	DYRK1A	3.1e-05	4
2 rows			

```
x.poisson[x.poisson < pT]
```

```
[1] 0.0001901230 0.0001141829
```