# Computer Lab 1

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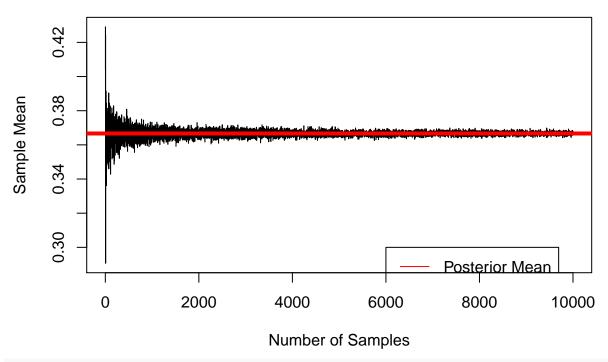
```
# Data distribution parameters
n <- 24
s <- 8
f <- n-s
# Prior distribution parameters
alpha_0 <- 3
beta_0 <- 3</pre>
```

a. Generating random numbers from posterior distribution and verifying graphically

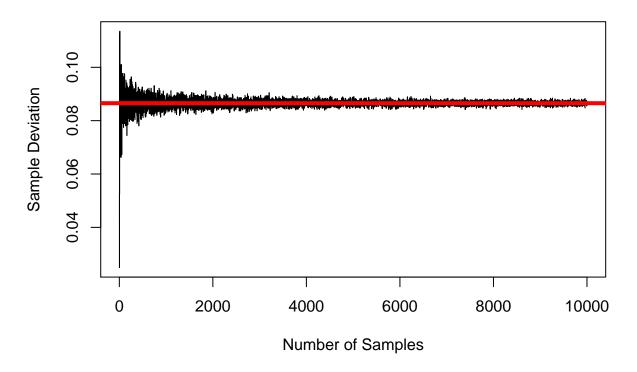
```
alpha_n <- alpha_0 + s
beta_n <- beta_0 + f
posterior_mean <- alpha_n/(alpha_n + beta_n)</pre>
posterior_variance <- (alpha_n * beta_n)/((alpha_n + beta_n)^2 *
                                                 (alpha_n + beta_n + 1))
posterior_sd <- sqrt(posterior_variance)</pre>
posteriorDrawParams <- function(num_draws,alpha_n,beta_n){</pre>
  #set.seed(1234)
  theta n <- rbeta(num draws, shape1 = alpha n, shape2 = beta n)
  sample_mean <- mean(theta_n)</pre>
  sample_sd <- sd(theta_n)</pre>
  return(c(sample_mean,sample_sd))
}
num_samples <- 1</pre>
mu_vector <- c()</pre>
sample_mu <- 0.3</pre>
sample_dev <- 1</pre>
#plot(num_samples,sample_mu)
\#while(abs(posterior\_mean - sample\_mu) > 1e-7){}
while(num_samples < 10000){</pre>
  num_samples <- num_samples + 1</pre>
  sample_mu <- posteriorDrawParams(num_samples,alpha_n,beta_n)[1]</pre>
  mu_vector <- c(mu_vector,sample_mu)</pre>
plot(c(2:num_samples),mu_vector,'l',
     xlab = 'Number of Samples',ylab = 'Sample Mean',
```

```
main = 'Sample Mean vs Number of Samples')
legend(6000,0.30,legend = c('Posterior Mean'),col = c('red'),lty=1)
abline(h = posterior_mean,lwd=4,col = 'red')
```

# **Sample Mean vs Number of Samples**



### Sample SD vs Number of Samples

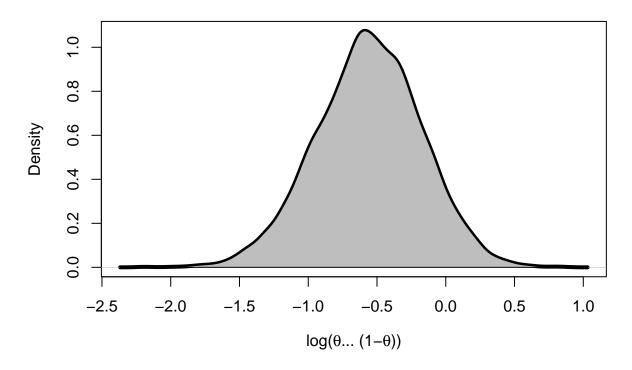


b.

```
posteriorDraw <- function(num_draws,alpha_n,beta_n){</pre>
  set.seed(1234)
  theta_n <- rbeta(num_draws, shape1 = alpha_n, shape2 = beta_n)</pre>
  sample_mean <- mean(theta_n)</pre>
  sample_sd <- sd(theta_n)</pre>
  return(theta_n)
}
\# P(theta > 0.4/y) from sampled data
simulated_data <- posteriorDraw(num_draws = 10000,alpha_n,beta_n)</pre>
posterior_prob <- length(simulated_data[simulated_data > 0.4])/length(simulated_data)
print(posterior_prob)
## [1] 0.3473
# Actual value of P(theta > 0.4/y)
x < -0.4
actual_prob <- pbeta(x,alpha_n,beta_n,lower.tail = FALSE)</pre>
print(actual_prob)
## [1] 0.3426654
c.
theta_post <- posteriorDraw(num_draws = 10000,alpha_n,beta_n)</pre>
phi <- log(theta_post) - log(1 - theta_post)</pre>
plot(density(phi), lwd=4,
     main = 'Posterior distribution of log odds of theta',
```

```
xlab = expression(paste('log(',theta, '/ (1-',theta,'))')))
polygon(density(phi),col = 'grey')
```

## Posterior distribution of log odds of theta

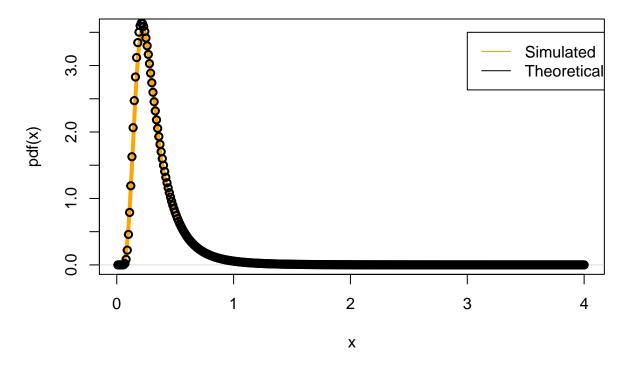


#### 2. Log-normal distribution and Gini coefficient

```
a.
```

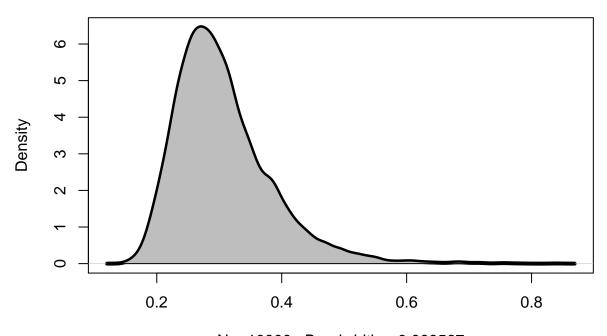
```
y \leftarrow c(38, 20, 49, 58, 31, 70, 18, 56, 25, 78)
mu <- 3.8
n <- 10
tau_sq \leftarrow sum((log(y)-mu)^2)/n
rinvchisquare <- function(num_draws, n, tau_sq){</pre>
  set.seed(1234)
  x <- rchisq(num_draws,df = n-1)</pre>
  x_{inv} \leftarrow ((n-1)*tau_{sq})/x
  return(x_inv)
# Scaled Inverse chi square distribution
pdfinvchisquare <- function(x,n,tau_sq){</pre>
  res \leftarrow (((tau_sq*(n-1))/2)^{(n-1)/2} * exp(-(tau_sq*(n-1))/(2*x)))/((x^{(1+(n-1)/2)}) * gamma((n-1)/2))
  return(res)
post_sample <- rinvchisquare(10000, n, tau_sq)</pre>
plot(density(post_sample),
     col = 'orange', lwd=4,
     main = 'Posterior Variance: Simulated vs Theoretical',
```

#### **Posterior Variance: Simulated vs Theoretical**



```
b.
```

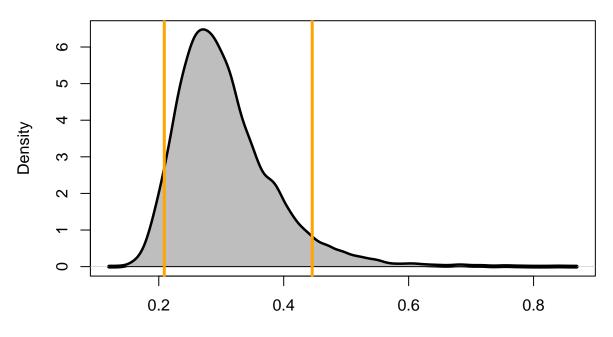
### **Posterior Distribution of Gini Coefficient**



N = 10000 Bandwidth = 0.009567

 $\mathbf{c}.$ 

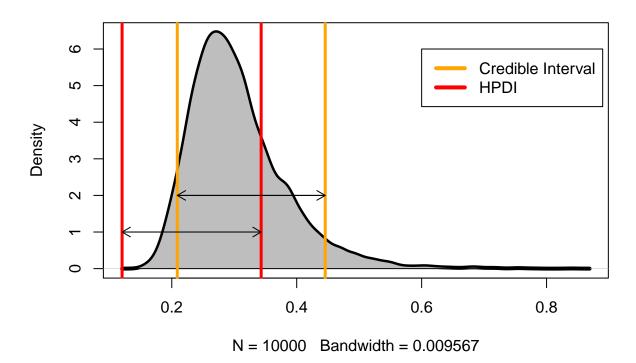
#### Posterior Distribution of Gini Coefficient with Credible Intervals



N = 10000 Bandwidth = 0.009567

```
kernel_gini_coeff <- density(gini_coeff)</pre>
# Highest Posterior Density Interval (HDPI)
sorted_kgc <- sort(kernel_gini_coeff$y, decreasing = TRUE)</pre>
cumm_kgc <- cumsum(sorted_kgc)</pre>
hpd <- cumm_kgc[length(cumm_kgc)]*0.9
hpdi <- range(kernel gini coeff$x[which(cumm kgc<hpd)])</pre>
plot(density(gini_coeff), lwd = 4,
     main = '90% Credible Interval and HDPI of Gini Coefficient')
polygon(density(gini_coeff),col = 'grey')
abline(v = lower_interval,col = 'orange', lwd = 3)
abline(v = upper_interval,col = 'orange', lwd = 3)
abline(v = hpdi[1],col = 'red',lwd = 3)
abline(v = hpdi[2],col = 'red',lwd = 3)
arrows(lower_interval,2,upper_interval,2,length = 0.1,col = 'black')
arrows(upper_interval,2,lower_interval,2,length = 0.1,col = 'black')
arrows(hpdi[1],1,hpdi[2],1,length = 0.1,col = 'black')
arrows(hpdi[2],1,hpdi[1],1,length = 0.1,col = 'black')
legend(0.6, 6,
       legend = c('Credible Interval', 'HPDI'),
       col = c('Orange','Red'),
       lty=c(1,1), lwd = 4)
```

#### 90% Credible Interval and HDPI of Gini Coefficient



3. Bayesian inference for the concentration parameter in the von Mises distribution

```
likelihood <- function(y, mu, kappa){</pre>
  return_val <- exp(kappa * sum(cos(y - mu))) / ((2*pi*besselI(kappa,nu=0))^10)
  return(return_val)
}
prior_kap <- function(kapa,lambda = 1){</pre>
  return_val <- lambda*exp(-1*lambda*kapa)
  return(return_val)
}
posterior_kap <- function(y, mu, kappa){</pre>
  return_val <- likelihood(y, mu, kappa) * prior_kap(kappa)</pre>
  return(return_val)
y \leftarrow c(-2.44, 2.14, 2.54, 1.83, 2.02, 2.33, -2.79, 2.23, 2.07, 2.02)
kap \leftarrow seq(from = 0, to = 10, by = 0.01)
#posterior_kappa <- rep(0,length(kap))</pre>
#for(k in 1:length(kap)){
# posterior_kappa[k] <- (posterior_kap(y,mu,kap[k]))</pre>
posterior_kappa <- (posterior_kap(y, mu, kappa = kap))</pre>
plot(kap,prior_kap(kap),'1',
     xlab = expression(paste(kappa)),
     ylab = expression(paste('p(',kappa,'|y)')),col = 'red',lwd = 2)
```

