Bayesian Learning (732A73) Lab 1

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Assisgnment 1

Daniel Bernoulli

Bernoulli ...again.

Let $y_1, \ldots, y_n|\theta \sim \mathrm{Bern}(\theta)$ and assume that you have obtained a sample with $s = 8$ successes in $n = 24$ trials. Assume a $\mathrm{Beta}(\alpha_0, \beta_0)$ prior for θ and let $\alpha_0 = \beta_0 = 3$

(a)Draw random numbers from the posterior $\theta|y \sim \mathrm{Beta}(\alpha_0 + s, \beta_0 + f)$, where $y = (y_1, \ldots, y_n)$,; and verify graphically that the posterior mean and standard deviation converges to the true values as the number of random draws grows large.

(b) Use simulation (nDraws = 10000) to compute the posterior probability $\Pr(\theta < 0.4|y)$ and compare with the exact value.

(c) Compute the posterior distribution of the log-odds $\phi = \log \frac{\theta}{1-\theta}$ (nDraws = 10000)

(a)

True mean for $Beta(\alpha, \beta)$ is:

$$E[\theta] = \frac{\alpha}{\alpha + \beta}$$

True σ^2 for $Beta(\alpha, \beta)$ is:

$$\sigma^2(\theta) = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

The prior function for this problem is:

$$p(\theta) \propto \theta^{\alpha_0 - 1}(1 - \theta^{\beta_0 - 1})$$

The likelihood function is :

$$p(y_1, \ldots, y_n|\theta) \propto \prod_{i=1}^{n} p(y_i|\theta) = \prod_{i=1}^{n} \theta^{y_i}(1-\theta)^{1-y_i} = \theta^s(1-\theta)^f$$

posterior is :

posterior

$$p(\theta|y) \propto Beta(\alpha_0 + s, \beta_0 + f) = Beta(11, 19)$$

In [1]:

```
s = 8
n = 24
f = 16
alpha_0 = beta_0 = 3
```

In [2]:

```
NDraws = 10000
res = matrix(,NDraws,3)
colnames(res) = c("iter","mean","sd")
alpha_n=alpha_0+s
beta_n = beta_0+f
TrueMean = (alpha_n)/(alpha_n+beta_n)
TrueVar = sqrt((alpha_n*beta_n)/((alpha_n+beta_n)^2*(alpha_n+beta_n+1)))
TrueMean
TrueVar
```

0.366666666666667

0.0865507910219387

In [3]:

```
PostB <- function(alpha= alpha_n,beta =beta_n ,n=NDraws){
    set.seed(43)

    for(i in 1:n){
        postB = rbeta(i,alpha,beta)
        res[i,] = c(i,mean(postB),sd(postB))

    }

    return(list(post=postB,mean_sd=res))
}
Posterior = PostB()
```

In [4]:

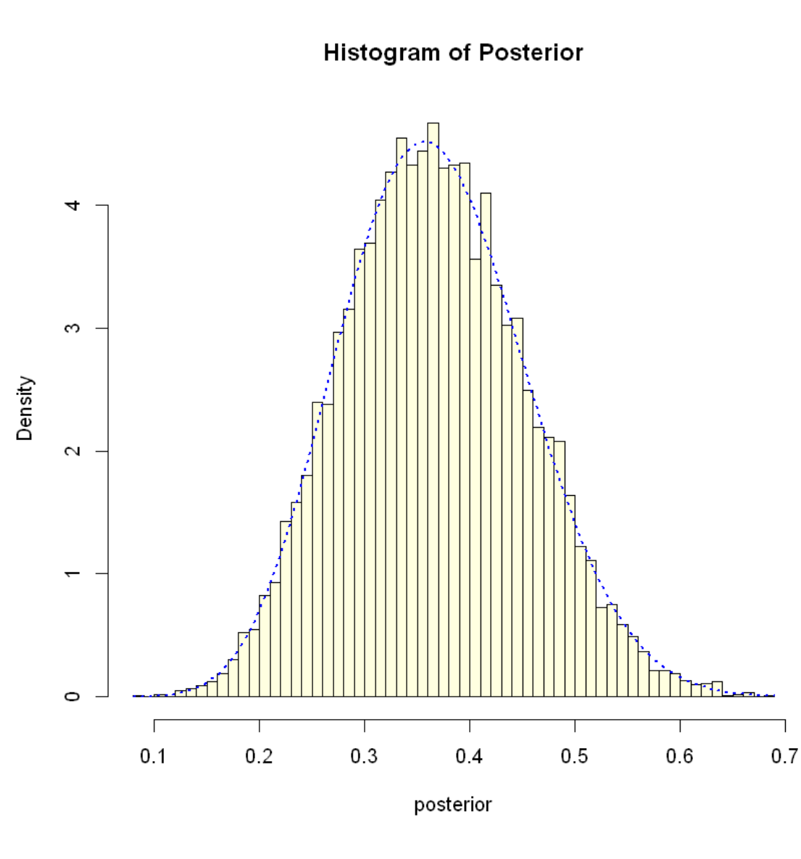
```
Posterior$mean_sd[1,3]=0
head(Posterior$mean_sd)
```

A matrix: 6 × 3 of type dbl

iter	mean	sd
1	0.3629400	0.00000000
2	0.2669733	0.07396499
3	0.3443858	0.06782267
4	0.3124737	0.09665964
5	0.4286403	0.13712276
6	0.3704016	0.11454616

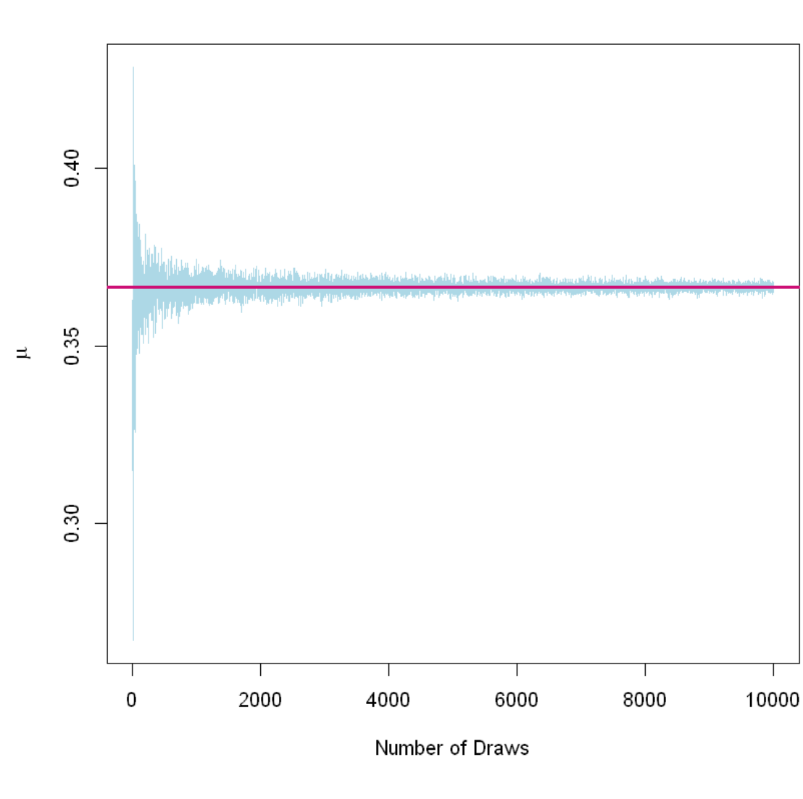
In [5]:

```
x <-seq( from =-5, to =10 , by =0.001)
hist(Posterior$post,col='lightyellow',freq = FALSE,breaks=50,xlab='posterior',main='H
curve(dbeta(x,alpha_n,beta_n),add=TRUE,col='blue', lwd =2,lty=3)
```



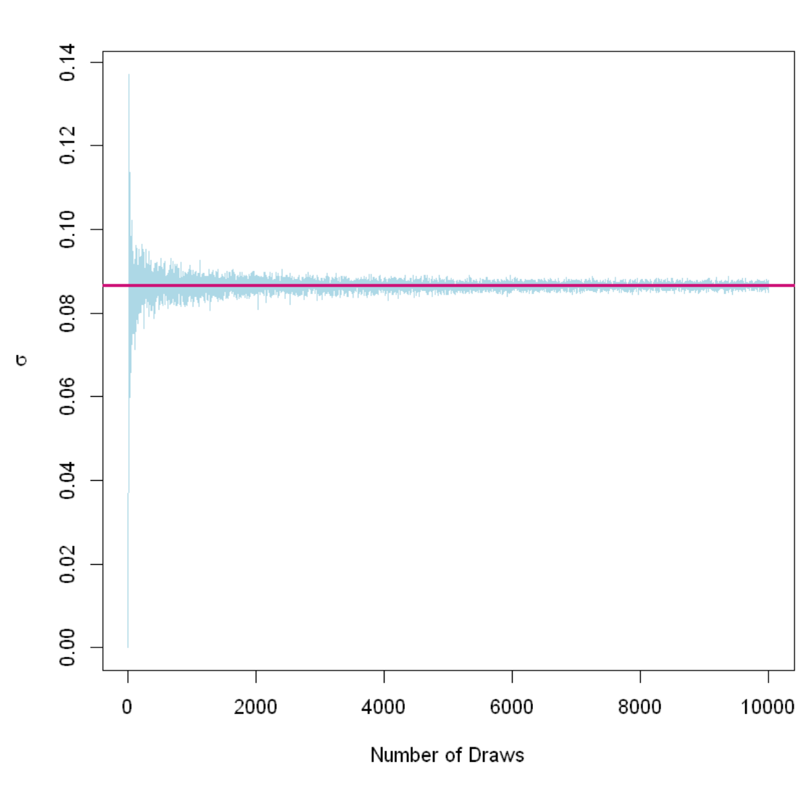
In [6]:

```
plot(x=1:NDraws,y=Posterior$mean_sd[,2],col='lightblue',lwd =1,type='l',xlab="Number c
abline(h=TrueMean,col='deeppink3',lwd=3)
```



In [7]:

```
plot(x=1:NDraws,y=Posterior$mean_sd[,3],col='lightblue',lwd =1,type='l',xlab="Number c
abline(h=TrueVar,col='deeppink3',lwd=3)
```



As we can see from the two plots above, as the number of draws increases, both μ and σ converge to the true values.

(b)

we can see that our estimation of $Pr(\theta > 0.4|y)$ is very close to true theoratical value.

In [8]:

```
exact_prb =1- pbeta(0.4,alpha_n,beta_n)
pos_prob = length(Posterior$post[Posterior$post > 0.4])/length(Posterior$post)
prb = data.frame(exact_prb,pos_prob)
colnames(prb) = c('excact_prob',"simulated_value")
prb
```

A data.frame: 1 × 2

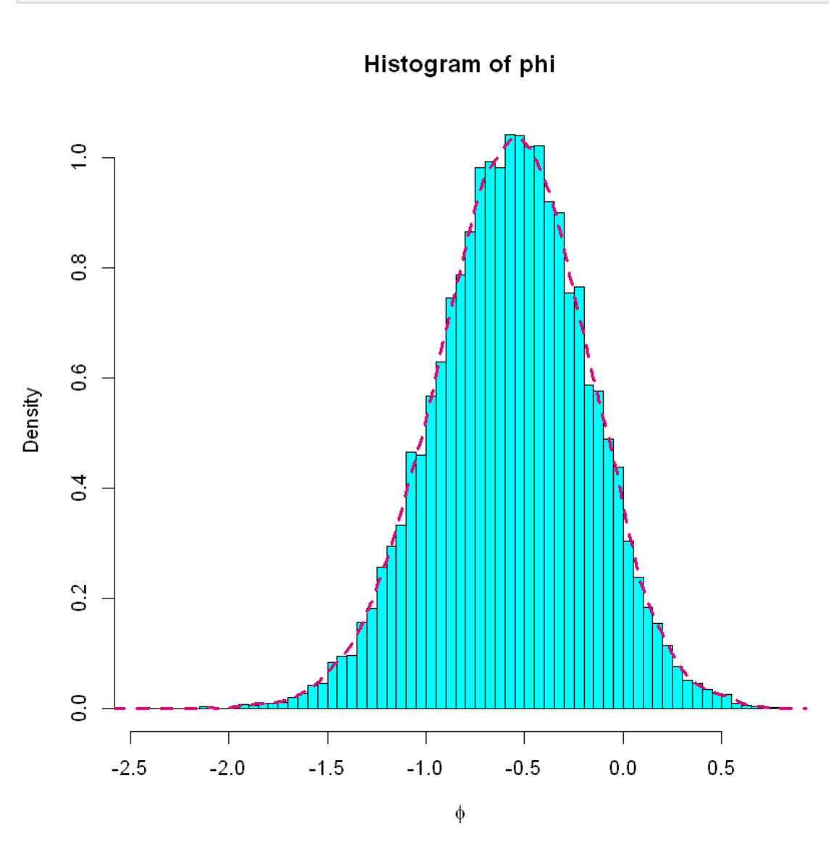
excact_prob	simulated_value
<dbl>	<dbl>
0.3426654	0.3402

(C)

Histogram and kernel density of of the data simulated from the posterior distribution of the $\phi = log(\theta\ (1-\theta))$ with 10000 draws.

In [9]:

```
phi = log(Posterior$post/(1- Posterior$post))
hist(phi,breaks=100,probability = TRUE,col='cyan',xlab=expression(phi))
lines(density(phi),col='deeppink3',lty=2,lwd=3)
```



Assignment 2

Log-normal distribution and the Gini coefficient.

Assume that you have asked 10 randomly selected persons about their monthly income (in thousands Swedish Krona) and obtained the following ten observations : 14, 25, 45, 25, 30, 33, 19, 50, 34 and 67. A common model for non-negative continuous variables is the log-normal distribution. The log-normal distribution $\log \mathcal{N}(\mu, \sigma^2)$

has density function

$$p(y|\mu, \sigma^2) = \frac{1}{y \cdot \sqrt{2\pi\sigma^2}} \exp(-\frac{1}{2\sigma^2}(\log y - \mu)^2)$$

where $y > 0$, $\mu > 0$ and $\sigma^2 > 0$

The log-normal distribution is related to the normal distribution as follows: $y \sim \log \mathcal{N}(\mu, \sigma^2)$ then

$\log y \sim \mathcal{N}(\mu, \sigma^2)$ Let $y_1, \ldots, y_n|\mu, \sigma^2 \overset{iid}{\sim} \log \mathcal{N}(\mu, \sigma^2)$, where $\mu = 3.8$

is assumed to be known but σ^2 is unknown with non-informative prior $p(\sigma^2) \propto 1/\sigma^2$ The posterior for σ^2 is the $Inv - \chi^2(n, \tau^2)$

distribution, where

$$\tau^2 = \frac{\sum_{i=1}^{n}(\log y_i - \mu)^2}{n}.$$

(a)

Simulate 10, 000 draws from the posterior of σ2 (assuming μ = 3.8) and com- pare it with the theoretical Inv − χ2(n, τ 2) posterior distribution.

In [10]:

```
num_draws = 10000 #sample size
mu = 3.8
observ = c(38,20, 49, 58, 31, 70, 18, 56, 25,78)
n_obser = length(observ)
```

In [11]:

```
tau2 = function(data,mu,n_obser){
    sum((log(data) -mu)^2 )/n_obser
}
```

In [12]:

```
tau2(observ,mu,n_obser )
```

0.261043665681099

In [13]:

```
rinvchisq <- function(num_draws, n_obser, tau_sq){
  set.seed(1234)
  x <- rchisq(num_draws,df = n_obser-1)
  x_inv <- ((n_obser-1)*tau_sq)/x
  return(x_inv)
}
```

In [14]:

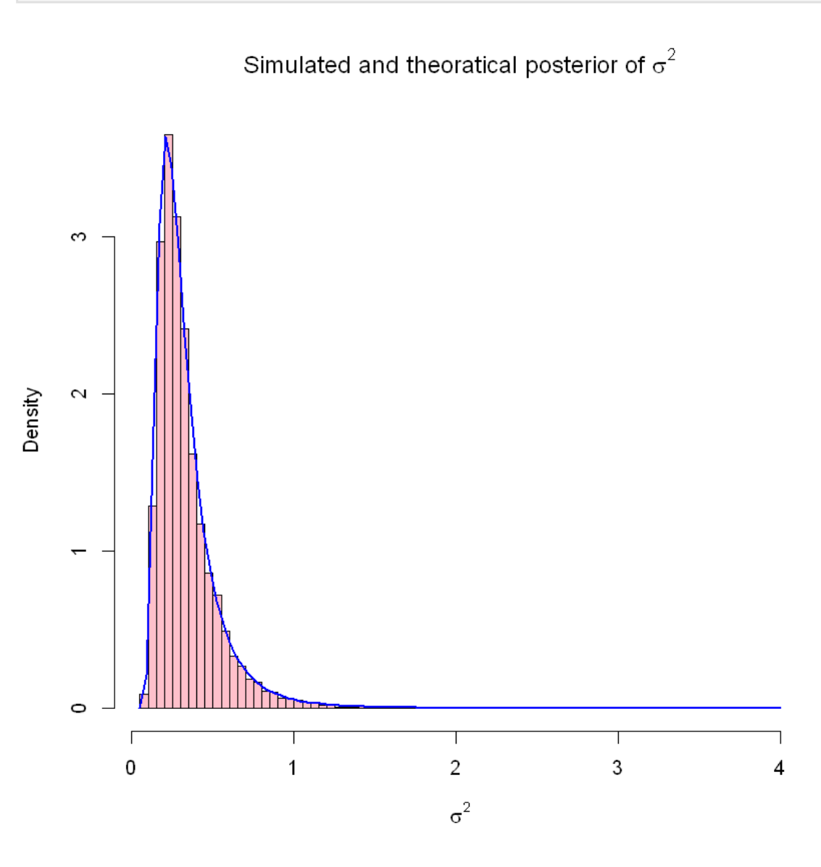
```
post_sigma2 <- function(m){
    set.seed(12345)
    rinvchisq(num_draws = num_draws,n_obser = n_obser,tau_sq = tau2(observ,mu,n_obser
}
```

In [15]:

```
dinvchisq <- function(x,n_obser,tau_sq){
  res <- (((tau_sq*(n_obser-1))/2)^(n_obser-1)/2 * exp(-(tau_sq*(n_obser-1))/(2*x)))/
  return(res)
}
```

In [16]:

```
x <-seq( from =0, to =10 , by =0.001)
hist(post_sigma2(m = num_draws ),probability = TRUE,col='pink',breaks = 100,
     main=expression(paste('Simulated and theoratical posterior of ', sigma^2) ),
    ,xlab = expression(paste(sigma ^2)))
    curve(dinvchisq(x ,n_obser= n_obser,tau_sq = tau2(observ,mu,n_obser )),add=TRUE,co
```

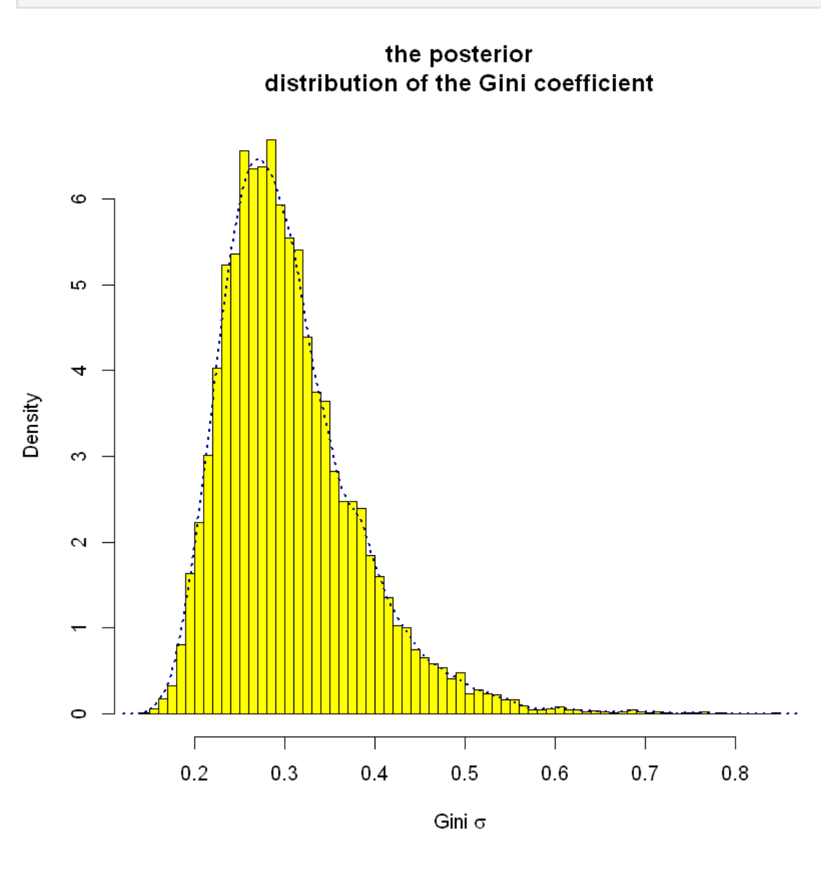


(b)

The most common measure of income inequality is the Gini coeffcient, G, where 0 < G < 1. $G = 0$ means a completely equal income distribution, whereas $G = 1$ means complete income inequality (see Wikipedia for more information about the Gini coeffcient). It can be shown that $G = 2\Phi\left(\sigma/\sqrt{2}\right) - 1$ when incomes follow a $\log \mathcal{N}(\mu, \sigma^2)$ distribution.$\Phi(z)$ is the cumulative distribution function (CDF) for the standard normal distribution with mean zero and unit variance. Use the posterior draws in a) to compute the posterior distribution of the Gini coeffcient G for the current data set

In [17]:

```
sigma2 = post_sigma2(m)
g_sigma = sqrt(sigma2)/sqrt(2)
gpdf =2* pnorm(q = g_sigma,mean = 0,sd = 1)-1
hist(gpdf,probability = TRUE,col='yellow',breaks=50,main='the posterior
distribution of the Gini coefficient',xlab=expression(paste('Gini ', sigma)))
lines(density(gpdf),col='darkblue',lty=3,lwd=2)
```



(c)

Use the posterior draws from b) to compute a 90% equal tail credible interval for G. A 90% equal tail interval (a; b) cuts off 5% percent of the posterior probability mass to the left of a, and 5% to the right of b. Also, do a kernel density estimate of the posterior of G using the density function in R with default settings, and use that kernel density estimate to compute a 90% Highest Posterior Density Interval (HPDI) for G. Compare the two intervals.

In [18]:

```
alpha_conf = 0.1
q_lower = quantile(gpdf,alpha_conf/2)
q_upper = quantile(gpdf,1-alpha_conf/2)
c(q_lower, q_upper)
true_mean = mean(gpdf)
```

5%: 0.208822072375841 **95%:** 0.445627779268848

In [19]:

```
true_mean
```

0.305486649437337

HPD: shortest possible interval that under the posterior has the siginificance probability (ie. 0.9)

In [20]:

```
KerDenG=density(gpdf)
xx = KerDenG$x
yy = sort(KerDenG$y,decreasing = TRUE,index.return=TRUE)
```

we will use the cumulative sum of kernel density values to find the shortest interval in data that contains 0.9 of the data. the reason we can use cumulative to find HPD interval is because our data is unimodal.

In [21]:

```
CumSum = cumsum(yy$x)
HPD_D = CumSum[length(CumSum)]*0.9
temp = which(CumSum<HPD_D)
HPD_interval = range(xx[which(CumSum<HPD_D)])
```

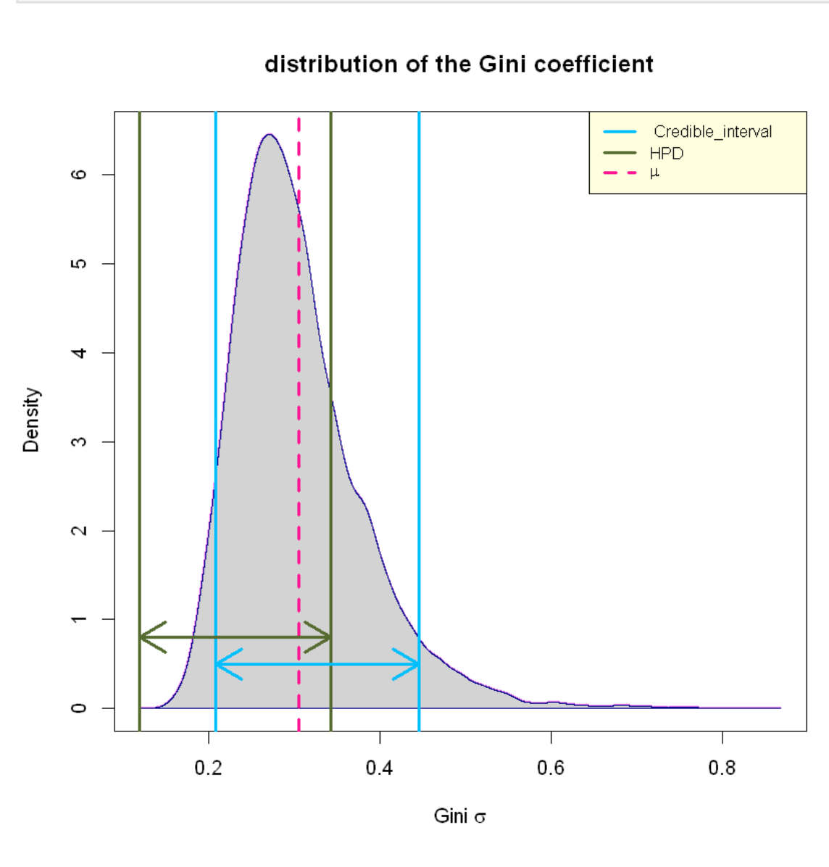
In [22]:

```
df = data.frame('HPD_interval'=HPD_interval,'credible_interval'=c(q_lower, q_upper))
```

In [24]:

```
y_temp=density(gpdf,n=10000)
plot(x=y_temp$x,y=y_temp$y,type = 'l', lwd = 2, col = 'violet',main= 'distribution of
     ,xlab=expression(paste('Gini ', sigma)),ylab='Density')
polygon(x=y_temp$x,y=y_temp$y, col = 'lightgrey',border = 'darkblue')
abline(v = true_mean, col="deeppink", lwd=3, lty=2)

segments(x0 =q_lower,y0 =0.5,x1 =q_upper,y1 =  0.5,col='deepskyblue1' ,lwd=3)
abline(v=q_lower,col='deepskyblue1' ,lwd=3)
abline(v=q_upper,col='deepskyblue1' ,lwd=3)
segments(x0=HPD_interval[1],y0=0.8,x1=HPD_interval[2],y=0.8,col='darkolivegreen',lwd=
abline(v=HPD_interval[1],col='darkolivegreen',lwd=3)
abline(v=HPD_interval[2],col='darkolivegreen',lwd=3)
arrows(x0 =q_lower,y0 =0.5,x1 =q_upper,y1 =  0.5,col='deepskyblue1' ,lwd=3,code = 3)
arrows(x0=HPD_interval[1],y0=0.8,x1=HPD_interval[2],y=0.8,col='darkolivegreen',lwd=3,
legend("topright",
  legend = c(" Credible_interval", "HPD",expression(mu)),
  col = c('deepskyblue1','darkolivegreen','deeppink'), lty=c(1,1,2), cex=0.8,lwd=3,bg=
```



Assignment 3

Bayesian inference for the concentration parameter in the von Mises distribution

distributions for circular data we convert the data into radians . The 10 observations in radians are

$$(-2.44, 2.14, 2.54, 1.83, 2.02, 2.33, -2.79, 2.23, 2.07, 2.02).$$

Assume that these data points are independent observations following the von Mises distribution

$$p(y|\mu, \kappa) = \frac{\exp[\kappa \cdot \cos(y - \mu)]}{2\pi I_0(\kappa)}, \quad -\pi \le y \le \pi,$$

where $I_0(\kappa)$ is the modified Bessel function of the first kind of order zero.The parameter $-\pi \le \mu \le \pi$ is the mean direction and κ > 0 is called the concentration parameter. Large κ gives a small variance around μ, and vice versa. Assume that μ is known to be 2.39. Let $\kappa \sim \mathrm{Exponential}(\lambda = 1)$ a priori, where λ is the rate parameter of the exponential distribution (so that the mean is $1/\lambda$).

(a) Plot the posterior distribution of κ for the wind direction data over a fine grid of κ values.

(b) Find the (approximate) posterior mode of κ from the information in a).

(a) Posterior :

$$p(\kappa \mid y_1, y_2, \ldots, y_n) \propto p(y_1, y_2, \ldots, y_n \mid \kappa) \cdot p(\kappa)$$

$$p(\kappa \mid y_1, y_2, \ldots, y_n) \propto \left[\frac{1}{I_0(\kappa)}\right]^n \cdot \exp\left[\sum_{i=1}^{n} \kappa \cdot cos(y_i - \mu) - \lambda\kappa\right]$$

$$p(\kappa \mid y_1, y_2, \ldots, y_n) \propto \left[\frac{1}{I_0(\kappa)}\right]^n \cdot \exp\left[\sum_{i=1}^{n} \kappa \cdot cos(y_i - 2.39) - \kappa\right]$$

In [25]:

```
y_data =c(-2.44, 2.14, 2.54, 1.83, 2.02, 2.33, -2.79, 2.23, 2.07, 2.02)
```

In [26]:

```
n =length(y_data)
mu = 2.39
n
```

10

In [27]:

```
# a function to compute prior for k (exponential with lambda = 1)
prior_k <- function(k){
    dexp(x = k,rate = 1)
}
```

In [28]:

```
likelikood_k <- function(y,mu ,n,k){

    exp(k*sum(cos(y-mu)))/((2*pi*besselI(k,nu = 0))^n)
}
```

In [29]:

```
k= seq(from = 0,to = 20,by = 0.01)
```

In [30]:

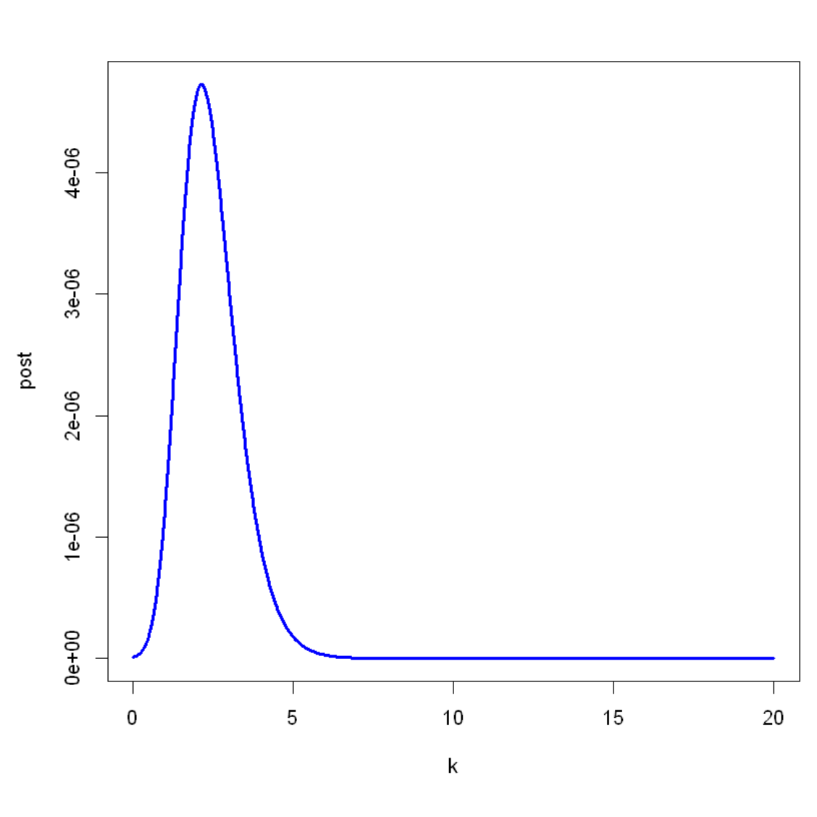
```
#
posterior_k <- function(y=y_data,mu=mu ,n=n,k=k){
    likelikood_k(y,mu ,n,k )*prior_k(k)
}
```

In [31]:

```
post=posterior_k(y=y_data,mu=mu ,n=n,k=k)
```

In [32]:

```
plot(y=post,x=k,col='blue',type='l',lwd=3)
```



(b) Mode

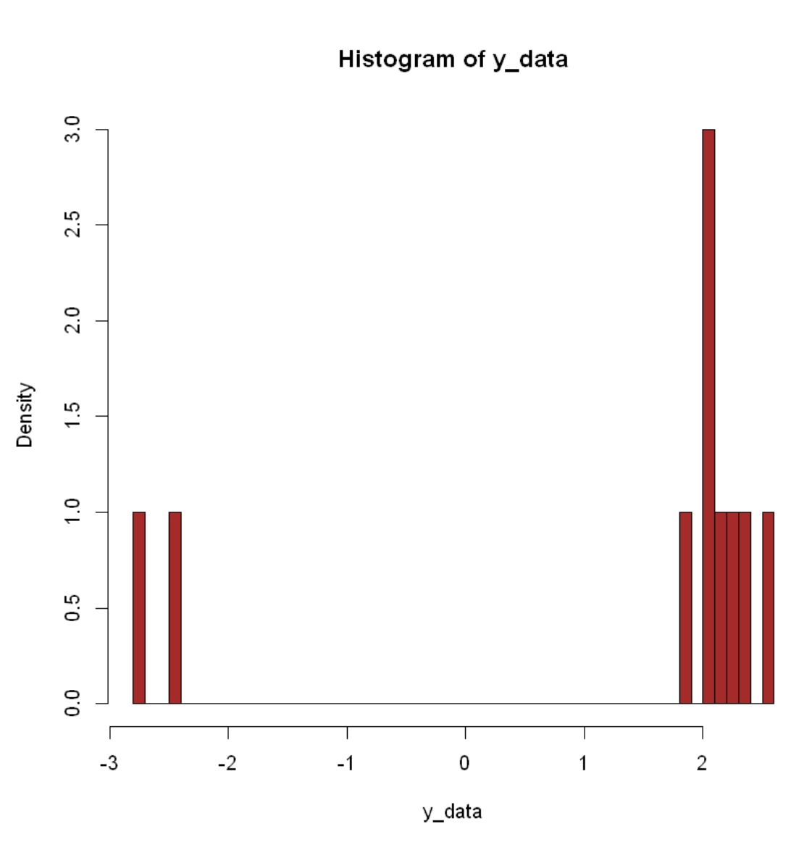
In [33]:

```
k[which.max(post)]
```

2.12

In [34]:

```
hist(y_data,probability = TRUE,breaks = 50,col='brown')
```



from the plot above we can see that $\kappa = 2.12$ is the actual mode of our data.

In [35]:

```
y_data[which(y_data<0)]
```

-2.44 · -2.79