CSE 291-D: Homework 2

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Problem 1

(a)

According to Metropolis-Hastings update rule, a proposal will be accepted only if

$$\alpha = \frac{P(\mathbf{X}') \cdot Q(\mathbf{X}^{(t)}|\mathbf{X}')}{P(\mathbf{X}^{(t)}) \cdot Q(\mathbf{X}'|\mathbf{X}^{(t)})} >= 1.$$

If it is applied to Gibbs sampling, I have

$$\begin{split} \frac{P(\boldsymbol{X}') \cdot Q(\boldsymbol{X}^{(t)}|\boldsymbol{X}')}{P(\boldsymbol{X}^{(t)}) \cdot Q(\boldsymbol{X}'|\boldsymbol{X}^{(t)})} &= \frac{P(X_i'|X_{-i}')P(X_{-i}') \cdot Q(\boldsymbol{X}^{(t)}|\boldsymbol{X}')}{P(X_i^{(t)}|X_{-i}')P(X_{-i}^{(t)}) \cdot Q(\boldsymbol{X}'|\boldsymbol{X}^{(t)})} \\ &= \frac{P(X_i'|X_{-i}')P(X_{-i}') \cdot P(X_i^{(t)}|X_{-i}')}{P(X_i^{(t)}|X_{-i}')P(X_{-i}') \cdot P(X_i'|X_{-i}')} \\ &= \frac{P(X_i'|X_{-i}')P(X_{-i}') \cdot P(X_i'|X_{-i}')}{P(X_i^{(t)}|X_{-i}')P(X_{-i}') \cdot P(X_i'|X_{-i}')} \\ &= 1. \end{split}$$

and thus, I prove proposal by Gibbs sampling will always be accepted.

(b)

If I integral X_a at both sides of detailed balance formula, I can derive

$$\int_{X_a} T(X_a|X_b)P(X_b) \ d(X_a) = \int_{X_a} T(X_b|X_a)P(X_a) \ d(X_a)$$

$$\Rightarrow \int_{X_a} T(X_a, X_b) \ d(X_a) = \int_{X_a} T(X_b|X_a)P(X_a) \ d(X_a)$$

$$\Rightarrow X_b = \int_{X_a} T(X_b|X_a)P(X_a) \ d(X_a).$$

Thus, I prove the stationary distribution requirement.

Problem 2

(a): update of μ

Assume $\mu \sim \text{Gaussian}(\mu_0, \tau_0)$, the update of μ can be derived as follows.

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$$\begin{split} P(\mu|\boldsymbol{x},\tau) &= \frac{P(\mu,\boldsymbol{x},\tau)}{\sum_{\mu}P(\mu,\boldsymbol{x},\tau)} \\ &= \frac{P(\boldsymbol{x}|\mu,\tau) \cdot P(\mu|\tau) \cdot P(\tau)}{\sum_{\mu}P(\boldsymbol{x}|\mu,\tau) \cdot P(\mu|\tau) \cdot P(\tau)} \\ &= \frac{P(\boldsymbol{x}|\mu,\tau) \cdot P(\mu)}{\sum_{\mu}P(\boldsymbol{x}|\mu,\tau) \cdot P(\mu)} \\ &\propto P(\boldsymbol{x}|\mu,\tau) \cdot P(\mu) \\ &= P(\mu) \cdot \prod_{i}P(x_{i}|\mu,\tau) \\ &= \sqrt{\frac{\tau_{0}}{2\pi}}e^{-\frac{1}{2}\tau_{0}(\mu-\mu_{0})^{2}} \cdot \prod_{i}\left(\sqrt{\frac{\tau}{2\pi}}e^{-\frac{1}{2}\tau(x_{i}-\mu)^{2}}\right) \\ &= \sqrt{\frac{\tau_{0}\tau^{n}}{2\pi}}e^{-\frac{1}{2}\left[\tau_{0}(\mu-\mu_{0})^{2}+\tau \cdot \left(\sum_{i}(x_{i}-\bar{x})^{2}+n(\bar{x}-\mu)^{2}\right)\right]} \\ &= \sqrt{\frac{\tau_{0}\tau^{n}}{2\pi}}e^{-\frac{1}{2}\left[\tau \cdot \sum_{i}(x_{i}-\bar{x})^{2}+(\tau_{0}+\tau n)\left(\mu-\frac{\tau_{0}\mu_{0}+\tau n\bar{x}}{\tau_{0}+\tau n}\right)^{2}+\frac{\tau_{0}\tau n}{\tau_{0}+\tau n}(\mu_{0}-\bar{x})^{2}\right]} \\ &\propto \sqrt{\frac{\tau_{0}+\tau n}{2\pi}}e^{-\frac{1}{2}(\tau_{0}+\tau n)\left(\mu-\frac{\tau_{0}\mu_{0}+\tau n\bar{x}}{\tau_{0}+\tau n}\right)^{2}} \\ &= \text{Gaussian}\left(\frac{\tau_{0}\mu_{0}+\tau n\bar{x}}{\tau_{0}+\tau n}, \frac{1}{\tau_{0}+\tau n}\right) \end{split}$$

(b): update of τ

Assume $\tau \sim \text{Gamma}(\alpha, \beta)$, the update of τ can be derived as follows.

$$\begin{split} P(\tau|\boldsymbol{x},\boldsymbol{\mu}) &= \frac{P(\boldsymbol{\mu},\boldsymbol{x},\tau)}{\sum_{\tau}P(\boldsymbol{\mu},\boldsymbol{x},\tau)} \\ &= \frac{P(\boldsymbol{x}|\boldsymbol{\mu},\tau) \cdot P(\tau|\boldsymbol{\mu}) \cdot P(\boldsymbol{\mu})}{\sum_{\tau}P(\boldsymbol{x}|\boldsymbol{\mu},\tau) \cdot P(\tau|\boldsymbol{\mu}) \cdot P(\boldsymbol{\mu})} \\ &= \frac{P(\boldsymbol{x}|\boldsymbol{\mu},\tau) \cdot P(\tau)}{\sum_{\tau}P(\boldsymbol{x}|\boldsymbol{\mu},\tau) \cdot P(\tau)} \\ &\propto P(\boldsymbol{x}|\boldsymbol{\mu},\tau) \cdot P(\tau) \\ &= P(\tau) \cdot \prod_{i}P(x_{i}|\boldsymbol{\mu},\tau) \\ &= \frac{\beta^{\alpha}}{\Gamma(\alpha)}\tau^{\alpha-1}e^{-\beta\tau} \cdot \prod_{i}\left(\sqrt{\frac{\tau}{2\pi}}e^{-\frac{1}{2}\tau(x_{i}-\boldsymbol{\mu})^{2}}\right) \\ &= \frac{1}{\sqrt{2\pi}}\tau^{\alpha+\frac{n}{2}-1}e^{-\tau\cdot\left(\beta+\frac{1}{2}\sum_{i}(x_{i}-\boldsymbol{\mu})^{2}\right)} \\ &\approx \operatorname{Gamma}\left(\alpha+\frac{n}{2},\beta+\frac{1}{2}\sum_{i}(x_{i}-\boldsymbol{\mu})^{2}\right) \end{split}$$

Problem 3

(a)

The figure is shown as follows.

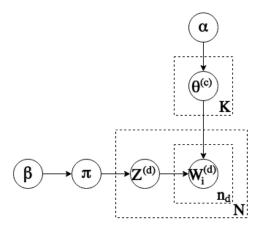


Figure 1: Directed Graphical Model Diagram

(b)

The joint probability of model can be written as follows.

$$P(\boldsymbol{\pi},\boldsymbol{Z},\boldsymbol{\theta},\boldsymbol{w}|\alpha=0.1,\beta=1) = P(\boldsymbol{\pi}|\beta=1) \cdot \prod_{d} P(Z^{(d)}|\boldsymbol{\pi}) \cdot \prod_{c} P(\theta^{(c)}|\alpha=0.1) \cdot \prod_{d} \prod_{i} P(w_{i}^{(d)}|Z^{(d)},\boldsymbol{\theta})$$

(c)

From Fig. 1, one can easy tell the follows.

- Markov blanket for π includes
 - β , - $Z^{(d)}$, $\forall d$.
- Markov blanket for $\theta^{(c)}$ includes
 - $\begin{array}{l} \ \alpha, \\ \ w_i^{(d)}, \ \forall i, \ \forall Z^{(d)} = c, \\ \ Z^{(d)}, \ \forall Z^{(d)} = c. \end{array}$
- ullet Markov blanket for $Z^{(d)}$ includes
 - $$\begin{split} &-\pi,\\ &-w_i^{(d)}, \forall i,\\ &-\theta^{Z^{(d)}}. \end{split}$$

(**d**)

The update for π is derived as follows. It it noticeable that $Z_i^{(d)}$ is a boolean notation to indicate the latent class for document d, i.e. $\sum_i Z_i^{(d)} = 1$.

$$\begin{split} P(\pmb{\pi}|\beta = 1, \pmb{Z}) &= \frac{P(\pmb{\pi}, \beta = 1, \pmb{Z})}{\sum_{\pmb{\pi}} P(\pmb{\pi}, \beta = 1, \pmb{Z})} \\ &\approx P(\pmb{\pi}|\beta = 1) \cdot \prod_{d} P(Z^{(d)}|\pmb{\pi}) \\ &= \frac{\Gamma(K)}{\Gamma(1)^K} \cdot \prod_{i=1}^K \pi_i^0 \cdot \prod_{d} \left(\frac{\Gamma(2)}{\prod_{i=1}^K \Gamma(Z_i^{(d)} + 1)} \prod_{i=1}^K \pi_i^{Z_i^{(d)}} \right) \\ &\approx \prod_{d} \left(\prod_{i=1}^K \pi_i^{Z_i^{(d)}} \right) \\ &= \prod_{i=1}^K \prod_{d} \pi_i^{Z_i^{(d)}} \\ &= \prod_{i=1}^K \pi_i^{\sum_{d} Z_i^{(d)}} \\ &\approx \text{Dirichlet} \left(\pi_1, \cdots, \pi_K \, \middle| \, \sum_{d} Z_1^{(d)} + 1, \cdots, \sum_{d} Z_K^{(d)} + 1 \right) \end{split}$$

The update for $\theta^{(c)}$ is derived as follows. Notice that I use ' to indicate every variable with document class c for simplicity, for example: Z', w', and d'. On top of that, $Z^{(d)}$ is again boolean notation here; while $w_i^{(d)}$ indicates the count of words i in document d.

$$\begin{split} P(\boldsymbol{\theta}^{(c)}|\boldsymbol{\alpha} &= 0.1, \boldsymbol{w}, \boldsymbol{Z}) = P(\boldsymbol{\theta}^{(c)}|\boldsymbol{\alpha} = 0.1, \boldsymbol{w}', \boldsymbol{Z}') \\ &= \frac{P(\boldsymbol{\theta}^{(c)}, \boldsymbol{\alpha} = 0.1, \boldsymbol{w}', \boldsymbol{Z}')}{\sum_{\boldsymbol{\theta}^c} P(\boldsymbol{\theta}^{(c)}, \boldsymbol{\alpha} = 0.1, \boldsymbol{w}', \boldsymbol{Z}')} \\ &\approx P(\boldsymbol{\theta}^{(c)}|\boldsymbol{\alpha} = 0.1) \cdot \prod_{\boldsymbol{d}'} P(\boldsymbol{w}^{(d')}|\boldsymbol{\theta}^{(c)}) \\ &= \frac{\Gamma(0.1V)}{\Gamma(0.1)^V} \cdot \prod_{i=1}^V \left(\boldsymbol{\theta}_i^{(c)}\right)^{-0.9} \cdot \prod_{\boldsymbol{d}'} \frac{\Gamma(\boldsymbol{n}_d' + 1)}{\prod_{i=1}^V \Gamma\left(\boldsymbol{w}_i^{(d')} + 1\right)} \prod_{i=1}^V \left(\boldsymbol{\theta}_i^{(c)}\right)^{\boldsymbol{w}_i^{(d')}} \\ &\approx \prod_{i=1}^V \left(\boldsymbol{\theta}_i^{(c)}\right)^{-0.9} \cdot \prod_{\boldsymbol{d}'} \prod_{i=1}^V \left(\boldsymbol{\theta}_i^{(c)}\right)^{\boldsymbol{w}_i^{(d')}} \\ &= \prod_{i=1}^V \left(\boldsymbol{\theta}_i^{(c)}\right)^{-0.9 + \sum_{\boldsymbol{d}'} \boldsymbol{w}_i^{(d')}} \\ &\approx \operatorname{Dirichlet} \left(\boldsymbol{\theta}_1^{(c)}, \cdots, \boldsymbol{\theta}_V^{(c)} \,\middle|\, \sum_{\boldsymbol{d}'} \boldsymbol{w}_1^{(d')} + 0.1, \cdots, \sum_{\boldsymbol{d}'} \boldsymbol{w}_V^{(d')} + 0.1 \right) \end{split}$$

The update for \mathbf{Z}^d is derived as follows. Again, $Z^{(d)}$ and $w_i^{(d)}$ follow the definition in previous update formulas.

$$\begin{split} P(\mathbf{Z}^{(d)}|\boldsymbol{\pi}, w^{(d)}, \boldsymbol{\theta}) &= \frac{P(\mathbf{Z}^{(d)}, \boldsymbol{\pi}, w^{(d)}, \boldsymbol{\theta})}{\sum_{\mathbf{Z}^{(d)}} P(\mathbf{Z}^{(d)}, \boldsymbol{\pi}, w^{(d)}, \boldsymbol{\theta})} \\ &= \frac{\Gamma(2)}{\prod_{i=1}^{K} \Gamma\left(Z_{i}^{(d)} + 1\right)} \prod_{i=1}^{K} \pi_{i}^{Z_{i}^{(d)}} \cdot \frac{\Gamma(n_{d} + 1)}{\prod_{i=1}^{V} \Gamma\left(w_{i}^{(d)} + 1\right)} \prod_{i=1}^{V} \left(\boldsymbol{\theta}_{i}^{Z^{(d)}}\right)^{w_{i}^{(d)}} \\ &\approx \frac{\prod_{i=1}^{K} \pi_{i}^{Z_{i}^{(d)}} \cdot \prod_{i=1}^{V} \left(\boldsymbol{\theta}_{i}^{Z^{(d)}}\right)^{w_{i}^{(d)}}}{\prod_{i=1}^{K} \Gamma\left(Z_{i}^{(d)} + 1\right)} \\ &= \pi_{Z_{i}^{(d)}} \cdot \prod_{i=1}^{V} \left(\boldsymbol{\theta}_{i}^{Z^{(d)}}\right)^{w_{i}^{(d)}} \end{split}$$

(e)

According to my implementation, shown as Code 1, the top ten words for each latent class can be listed as follows. It is noticeable that I filter out most common English stop words through one of the *CountVectorizer* parameters.

- $P(\pi = 0) = 0.35005771$: ['edu', u'lines', u'subject', u'organization', u'com', u'space', u'writes', u'article', u'posting', u'just']
- $P(\pi=1)=0.25799032$: [u'edu', u'image', u'jpeg', u'com', u'subject', u'organization', u'lines', u'space', u'graphics', u'software']
- $P(\pi=2)=0.39195198$: [u'edu', u'lines', u'subject', u'organization', u'com', u'space', u'writes', u'article', u'like', u'posting']

As one can tell, there are many stop words among them, since these words account for a great portion in each document.

Code Listing 1: Code for Homework 2

```
# coding: utf-8
# In[1]:
import numpy as np
from collections import defaultdict
import itertools
from sklearn.datasets import fetch_20newsgroups
from sklearn.feature_extraction.text import CountVectorizer
from sklearn.metrics import accuracy_score
seed = 514
np.random.seed(seed)
# In[2]:
# select targets
categories = [ 'talk.religion.misc', 'comp.graphics', 'sci.space']
K = len(categories)
# In[3]:
# load data
```

```
newsgroups = fetch_20newsgroups(subset='all', categories=categories,
                                     shuffle=True, random_state=seed)
N = len(newsgroups.data)
# In[4]:
# create vectorizer
vectorizer = CountVectorizer(stop_words='english')
tdata = vectorizer.fit_transform(newsgroups.data, )
V = len(vectorizer.vocabulary_)
# In[5]:
# prepare words
wordLst = defaultdict(list)
idxLst = zip(*tdata.nonzero())
for idx in idxLst:
    wordLst[idx[0]].append((idx[1], tdata[idx]))
# In[6]:
# given a, b
a = 0.1
b = 1.0
\# initialize pZ
pZ = np.zeros((N, K))
for d in range(N):
    for k in range(K):
        pZ[d,k] = 1.0 / K
# initialize Z
np.random.seed(514)
Z = np.zeros((N,), dtype=np.int)
for d in range(N):
    c = np.random.choice(range(K), 1, p=pZ[d])[0]
    Z[d] = c
# In[7]:
# main process
for i in xrange(200):
    if i % 10 == 0:
        print 'Now is working on Iteration #%d...' % i
    np.random.seed(514)
    # update pP
    sD = np.zeros((K,))
    for d in range(N):
        sD[Z[d]] += 1
    pP = np.random.dirichlet(sD + b, 1)[0]
    # update pT
    sW = np.zeros((K, V))
    for d, vcLst in wordLst.iteritems():
        for v, c in vcLst:
            sW[Z[d], v] += c
```

```
pT = np.zeros((K, V))
    for k in range(K):
        pT[k] = np.random.dirichlet(sW[k] + a, 1)[0]
    # update pZ
    for d in range(N):
        for k in range(K):
            pZ[d,k] = np.log(pP[k])
            for v, c in wordLst[d]:
                pZ[d,k] += c * np.log(pT[k,v])
        # trick
        maxv = max(pZ[d])
        pZ[d] = np.exp(pZ[d] - maxv)
        pZ[d] /= sum(pZ[d])
    # update Z
    for d in range(N):
        c = np.random.choice(range(K), 1, p=pZ[d])[0]
        Z[d] = c
# In[8]:
# maximize prediction
print pP
argm = np.argmax(pZ, axis=1)
def calACC(perm):
    pred = []
    for v in argm:
        pred.append(perm[v])
    return accuracy_score(newsgroups.target, pred)
maxc = 0
for perm in itertools.permutations(range(K), K):
    pred = []
    for v in argm:
        pred.append(perm[v])
    acc = accuracy_score(newsgroups.target, pred)
    if acc > maxc:
       maxc = acc
       maxp = pred
# In[9]:
# ten most frequent words
words = vectorizer.get_feature_names()
for k in range(K):
    sortLst = sorted(list(enumerate(pT[k])), key=lambda x: x[1],
                                        reverse=True)
    print [words[p[0]] for p in sortLst[:10]]
```