### **Parallelising Glauber Dynamics**

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Joint Math Meetings 2024 http://www.arxiv.org/abs/2307.07131

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## **Glauber dynamics**

Problem: Approximately sample from

$$\mu(x) \propto e^{f(x)}$$
 on  $\prod_{i=1}^n \Omega_i$  (e.g.,  $\{\pm 1\}^n$ ).



Method (MCMC): Run Glauber dynamics. Given  $X_t$ ,

- Select coordinate:  $i \sim \text{Uniform}([n])$ .
- Resample coordinate:  $X_i | X_{-i}$ , i.e.,

$$X_{t+1,i} = z$$
 with probability  $\mu(X_i = z | X_{\sim i} = x_{\sim i}) = \frac{e^{f(x_{i \leftarrow z})}}{\sum_{z' \in \Omega_i} e^{f(x_{i \leftarrow z'})}}$ .

•  $\mu$  is **stationary distribution** (preserved under Markov chain).

For "nice"  $\mu$ , we have **rapid mixing**: letting  $\mu_T = \text{Law}(X_T)$ ,

$$t_{\mathsf{mix}}(\varepsilon) := \min \left\{ T : \mathrm{TV}(\mu_T, \mu) \leq \varepsilon \right\} = O\left(n \ln \left(\frac{n}{\varepsilon}\right)\right).$$

## Parallelising Glauber dynamics

**Question:** Glauber dynamics is sequential. Can we do better with parallel computation?

**Natural idea:** Resample k coordinates at a time.

1. k-Glauber dynamics mixes k times as fast.

**Problem:** How can we implement one step of *k*-Glauber?

2. Parallel algorithm for sampling from Ising (& p-spin) model.

### Related work

- 1. Various algorithms in  $\mathbb{R}^n$  for sampling from log-concave distributions using the gradient take o(n) steps; randomized midpoint method is fully parallelisable (RNC) [Shen and Lee, 2019]
- 2. Fast parallel algorithms under Dobrushin conditions [Feng et al., 2021, Liu and Yin, 2022]
  - Gives an alternate approach for the Ising model, but not the p-spin model
- 3. Fast parallel algorithms (RNC) when fast counting algorithms exist [Anari et al., 2023a, Anari et al., 2023b]

### Our focus: Get a fast parallel algorithm

- 1. in discrete setting
- 2. under general conditions for mixing
- 3. without fast parallel counting algorithms

### **Outline**

2 Parallel algorithm for Ising model

## k-Glauber dynamics

k-Glauber dynamics: Given  $x_t$ ,

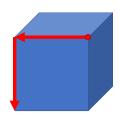
- Select random subset  $S \subset [n]$ , |S| = k.
- Resample subset  $x_S|x_{S^c}$ .  $X_{t+1,S}=z$  with probability

$$\mu(X_S = z | X_{S^c} = x_{S^c}) = \frac{e^{f(x_{S \leftarrow z})}}{\sum_{z' \in \{\pm 1\}^S} e^{f(x_{S \leftarrow z'})}}.$$

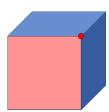


k-Glauber  $\neq k$  steps of Glauber...

2 steps of Glauber



2-Glauber



...but intuitively it should only be better!

### Down-up walks

A **Markov kernel**  $K: A \leadsto B$  is given by a transition matrix  $\mathbb{R}^{A \times B}$  with

 $K_{ab} = \text{probability of going from } a \text{ to } b.$ 

For a measure  $\mu$  on A,  $\mu K$  is the measure on B after applying K. Given a distribution  $\mu$  on  $\binom{[n]}{k}$ ,  $\ell < k$ , define...

Down operator	Up operator
$D_{k \to \ell} : \binom{[n]}{k} \leadsto \binom{[n]}{\ell}$	$U_{\ell  o k} : \binom{[n]}{\ell} \leadsto \binom{[n]}{k}$
Choose a uniform random subset	Choose a random superset
of $A$ of size $\ell$	of $B$ of size $k$ ,
	with probability $\mu(A A\supset B)$ .
$D_{k  o \ell}(A, B) = \mathbb{1}_{B \subset A} \frac{1}{{k \choose \ell}}$	$U_{\ell \to k}(B,A) = \mathbb{1}_{B \subset A} \frac{\mu(A)}{\sum_{A' \supset B} \mu(A')}$

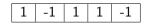
Given  $\mu = \mu_k$ , let  $\mu_\ell = \mu_k D_{k \to \ell}$ .

## Realizing Glauber dynamics as down-up walk

Let  $\mu$  be a measure on  $\{\pm 1\}^n$ .

The **homogenization** of  $\mu$  is the measure over  $\binom{[n] \times \{\pm 1\}}{n}$  where

 $(x_1,\ldots,x_n)$  is identified with  $\{(x_1,1)\ldots,(x_n,n)\}.$ 



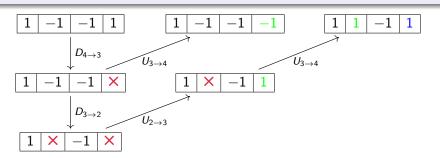
# Realizing Glauber dynamics as down-up walk

- $D_{k \to \ell}$ : Erase  $k \ell$  coordinates.
- $U_{\ell \to k}$ : Restore  $k \ell$  coordinates, using conditional distribution.

#### Lemma

$$P_{Glauber} = P_{n \leftrightarrow n-1}^{\nabla} = D_{n \rightarrow n-1} U_{n-1 \rightarrow n}$$

$$P_{k-Glauber} = P_{n \leftrightarrow n-k}^{\nabla} = D_{n \rightarrow n-1} \cdots D_{n-k+1 \rightarrow n-k} U_{n-k \rightarrow n-k+1} \cdots U_{n-1 \rightarrow n}.$$



### General framework: Diffusion models

A general way to form a Markov chain is by adding noise and denoising (using the posterior) [Montanari, 2023].

• Discrete setting  $\{\pm 1\}^n$ :

Noise: Erase coordinates.

1 step	k steps
Glauber	<i>k</i> -Glauber

• Continuous setting  $\mathbb{R}^n$ :

Noise: Add Gaussian N(0, t).

t  o 0	t > 0
Langevin dynamics	Proximal sampler [Lee et al., 2021]

Proximal sampler gives best known rates  $(\widetilde{O}(\sqrt{n}))$  for high-accuracy log-concave sampling [Fan et al., 2023].

Denoising process is also used in machine learning for generative modeling [Sohl-Dickstein et al., 2015, Song and Ermon, 2019, Song et al., 2020].

## Mixing for Markov chains

Often measure closeness of distributions in  $\chi^2$  or  $\mathcal{D}_{\mathsf{KL}}$ -divergence:

$$\chi^2(\nu\|\mu) = \int \left(\frac{d\nu}{d\mu} - 1\right)^2 \, d\mu \qquad \mathcal{D}_{\mathsf{KL}}(\nu\|\mu) = \int \frac{d\nu}{d\mu} \ln \frac{d\nu}{d\mu} \, d\mu.$$

These are examples of f-divergences  $D_f(\nu \| \mu) = \int f\left(\frac{d\nu}{d\mu}\right) d\mu$ .

#### Definition

 $P:A\leadsto B$  satisfies  $(1-\kappa)$ -contraction in f-divergence w.r.t.  $\mu$  if

$$D_f(\nu P \| \mu P) \leq (1 - \kappa) D_f(\nu \| \mu).$$

If  $P: A \leadsto A$  with stationary distribution  $\mu$ , we have exponential convergence of Markov chain:

$$D_f(\nu P^t || \mu) \leq (1 - \kappa)^t D_f(\nu || \mu).$$

# Mixing for Markov chains

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- For  $\chi^2$ , equivalent to spectral gap  $(\kappa)$  or Poincaré inequality  $(1/\kappa)$ .
- For  $\mathcal{D}_{\mathsf{KL}}$ , for  $D_{n \to n-1}$  called approximate tensorization of entropy, closely related to **log-Sobolev inequality**.

For Glauber dynamics, **rapid mixing** when  $\kappa = \Omega\left(\frac{1}{n}\right)$ : get constant contraction after O(n) steps.

### Main theorem

### Theorem (k-Glauber mixes k times faster)

Let  $\mu$  be a distribution on  $\Omega = \prod_{i=1}^n \Omega_i'$ ,  $1 \le k \le n$ , and  $C \ge 1$ . If  $D_{n \to n-1}$  satisfies  $(1 - \frac{1}{C_n})$ -contraction in  $\chi^2$  or  $\mathcal{D}_{\mathsf{KL}}$ -divergence, then  $P_{k\text{-}Glauber}$  satisfies  $(1 - \Omega\left(\frac{k}{C_n}\right))$ -contraction in  $\chi^2$  or  $\mathcal{D}_{\mathsf{KL}}$ -divergence.

#### Alternative phrasing:

- 1. If  $P_{\text{Glauber}}$  has Poincaré constant Cn then  $P_{k\text{-Glauber}}$  has Poincaré constant  $O\left(\frac{Cn}{k}\right)$ .
- 2. If  $\mu$  satisfies C-approximate tensorization of entropy, then  $\mu$  satisfies  $O\left(\frac{C}{k}\right)$ -approximate k-uniform block factorization of entropy.

### Main theorem

### Theorem (k-Glauber mixes k times faster)

Let  $\mu$  be a distribution on  $\Omega = \prod_{i=1}^n \Omega_i'$ ,  $1 \le k \le n$ , and  $C \ge 1$ . If  $D_{n \to n-1}$  satisfies  $(1 - \frac{1}{C_n})$ -contraction in  $\chi^2$  or  $\mathcal{D}_{\mathsf{KL}}$ -divergence, then  $P_{k\text{-}Glauber}$  satisfies  $(1 - \Omega\left(\frac{k}{C_n}\right))$ -contraction in  $\chi^2$  or  $\mathcal{D}_{\mathsf{KL}}$ -divergence.

Given contraction of  $D_{n \to n-1}$  in

$$P_{\mathsf{Glauber}} = P_{n \leftrightarrow n-1}^{\triangledown} = D_{n \rightarrow n-1} U_{n-1 \rightarrow n},$$

how to show (k times as much) contraction of

$$P_{k\text{-Glauber}} = P_{n \leftrightarrow n-k}^{\triangledown} = \underbrace{D_{n \rightarrow n-1} \cdots D_{n-k+1 \rightarrow n-k}}_{k} U_{n-k \rightarrow n-k+1} \cdots U_{n-1 \rightarrow n}?$$

**Sufficient to show:** for any m < n,  $D_{m \to m-1}$  is "at least as contractive" as  $D_{n \to n-1}$ .

### **Proof sketch**

**Need to show:** for any m < n,  $D_{m \to m-1}$  is "at least as contractive" as  $D_{n \to n-1}$ .

**Idea:** Write  $D_{m \to m-1}$  as projection of  $D_{n \to n-1}$  tensorized with noise!

- Tensorization with noise slightly degrades contraction.
- Projection only improves contraction.

**Bernoulli-Laplace model:** For  $\mu = \text{Uniform}\binom{[n]}{k}$ , suppose  $D_{k \to k-1}$  has contraction  $\kappa_{\text{BL},k}$  (known, [Salez, 2021]). Suffices to show:

#### Lemma

Let  $\kappa_k$  be the contraction of  $D_{k \to k-1}$  w.r.t.  $\mu_k$ . Then  $\kappa_k \geq \frac{n \kappa_n \kappa_{\text{BL},k}}{n \kappa_n + k \kappa_{\text{BL},k}}$ .

### **Tensorization**

### Proposition

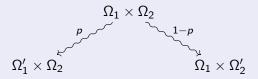
Let  $\kappa(\cdot)$  denote contraction in  $\chi^2$  or  $\mathcal{D}_{\mathsf{KL}}$ . Given kernels

$$P_1:\Omega_1\leadsto\Omega_1'$$

$$P_2:\Omega_2\leadsto\Omega_2',$$

define

$$P = p(P_1 \otimes I_2) + (1 - p)(I_1 \otimes P_2) : \Omega_1 \times \Omega_2 \leadsto \Omega_1' \times \Omega_2 \sqcup \Omega_1 \times \Omega_2'$$



Then

$$\kappa(P) \geq \min\{p\kappa_1, (1-p)\kappa_2\}.$$

When  $\Omega_i = \Omega'_i$  this is the product Markov chain.

## **Projection**

### Proposition

Given kernels making the following diagram commute:

$$(\Omega_1, \mu_1) \xrightarrow{P} (\Omega_2, \mu_2)$$

$$\downarrow^{\pi_1} \qquad \qquad \downarrow^{\pi_2}$$

$$(\Omega'_1, \mu'_1) \xrightarrow{P'} (\Omega'_2, \mu'_2)$$

Then

$$\kappa(P') \geq \kappa(P)$$
.

• Tensorization. Let  $\Omega = \{\pm 1\} \times [n]$ ,  $p = \frac{\kappa_{\text{BL},k}}{\kappa_n + \kappa_{\text{BL},k}}$ 

$$P = p(D_{n \to n-1} \otimes I_{\binom{[n]}{k}}) + (1-p)(I_{\binom{\Omega}{n}} \otimes D_{k \to k-1})$$
$$\binom{\Omega}{n} \times \binom{[n]}{k} \leadsto \binom{\Omega}{n-1} \times \binom{[n]}{k} \cup \binom{\Omega}{k} \times \binom{[n]}{k-1}$$
$$\kappa(P)^{-1} \leq \kappa_n^{-1} + \kappa_{\mathrm{BL},k}^{-1}.$$



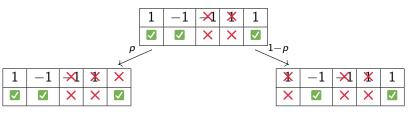
• Projection (only keep ✓ coordinates) is

$$P' = \left(1 - \frac{p(n-k)}{n}\right) D_{k \to k-1} + \frac{p(n-k)}{n} I$$
$$\binom{\Omega}{k} \leadsto \binom{\Omega}{k-1} \cup \binom{\Omega}{k}.$$

- ▶ Probability  $\frac{p(n-k)}{n}$  of trying to remove already-removed coordinate.
- Solve for  $\kappa_k$ .

• Tensorization. Let  $\Omega = \{\pm 1\} \times [n]$ ,  $p = \frac{\kappa_{\text{BL},k}}{\kappa_n + \kappa_{\text{BL},k}}$ 

$$P = p(D_{n \to n-1} \otimes I_{\binom{[n]}{k}}) + (1-p)(I_{\binom{\Omega}{n}} \otimes D_{k \to k-1})$$
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$$\kappa(P)^{-1} \leq \kappa_n^{-1} + \kappa_{\mathrm{BL},k}^{-1}.$$



• **Projection** (only keep **☑** coordinates) is

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- ▶ Probability  $\frac{p(n-k)}{n}$  of trying to remove already-removed coordinate.
- Solve for  $\kappa_k$ .

### **Outline**

1 k-Glauber mixes k times as fast

2 Parallel algorithm for Ising model

# Sampling from the Ising model

#### **Definition**

The **Ising model** with interaction matrix  $J \in \mathbb{R}^{n \times n}$  and external field  $h \in \mathbb{R}^n$  is the probability distribution on  $\{\pm 1\}^n$  given by

$$\mu_{J,h}(x) \propto \exp\left(\frac{1}{2}\langle x, Jx \rangle + \langle h, x \rangle\right), \quad x \in \{\pm 1\}^n.$$

- Model for interacting particles in statistical physics.
  - ▶  $J_{ij} > 0$  incentivizes  $x_i = x_j$ : **ferromagnetic** interaction.
  - ▶  $J_{ij}$  < 0 incentivizes  $x_i = -x_j$ : **antiferromagnetic** interaction.
- Connections to TCS, (Bayesian) statistics/machine learning,...
- When J is dense, cannot update different coordinates independently.

# Sampling from the Ising model

Glauber dynamics mixes rapidly under weak interactions.

## Theorem ([Anari et al., 2021])

Let  $J \in \mathbb{R}^{n \times n}$  be a symmetric matrix satisfying  $0 \leq J < I_n$ . Then for any  $\varepsilon > 0$ , Glauber dynamics mixes to  $\varepsilon$  in total variation distance in

$$O\left(\frac{n\log\left(\frac{n}{\varepsilon}\right)}{1-\|J\|_{op}}\right)$$
 steps.

- Their techniques can be used to show approximate tensorization of entropy.
- So if we could apply *k*-Glauber, then we get factor of *k* parallel speedup.
  - ▶ We will take  $k = \widetilde{\Theta}\left(\frac{n}{\|J\|_{F}}\right)$ .

# Parallel Sampling from the Ising model

#### Theorem

Fix c>0. With appropriate choice of constants depending only on c, if J is symmetric PSD with  $\|J\|\leq 1-c$ , then ParallellsingSampler outputs a sample  $\varepsilon$ -close in TV distance from  $\mu_{J,h}$  and, with probability  $\geq 1-\varepsilon$ , runs in time

$$O\left(\max\{\|J\|_{F},1\}\operatorname{poly}\log\left(\frac{n}{\varepsilon}\right)\right)$$

on a parallel machine with poly(n) processors.

Since 
$$||J||_F \leq \sqrt{n}$$
, this is a  $\widetilde{\Theta}\left(\frac{n}{||J||_F}\right) = \widetilde{\Omega}(\sqrt{n})$  speedup.

### **Algorithm 1** Parallel Ising Sampler (**ParallelIsingSampler**)

- 1: Input:  $J \in \mathbb{R}^{R \times R}$  ( $|R| \leq n$ ),  $h \in \mathbb{R}^R$ ,  $\varepsilon \in (0, \frac{1}{2})$ .
- 2: **if**  $\|J^{\aleph}\|_F \le c_3/\ln\left(\frac{n}{\varepsilon}\right)$  **then** ( $\aleph$  means zero out diagonal)
- 3: **Approximate rejection sample (ARS)** using  $\nu(x) \propto e^{\langle h+\widehat{h},x\rangle}$  where  $\widehat{h}$  solves  $\mathbb{E}_{\mu_{x+\widehat{x}}}J^{\otimes}x=\widehat{h}$ .
- 4: else
- 5: Initialize  $y \sim \nu_0(x) \propto e^{\langle h, x \rangle}$ .
- 6: **for** t from 1 to  $\Theta$  (poly  $\log \left(\frac{n}{\varepsilon}\right) ||J||_F$ ) **do**
- 7: Choose  $S \subseteq R$  a random subset of size  $\widetilde{\Theta}\left(\frac{|R|}{\|J\|_E}\right)$ .
- 8: Let  $y_S \leftarrow \mathbf{ParallellsingSampler}(J_S, J_{S \times R \setminus S} y_{R \setminus S} + h_S, \varepsilon)$ .
- 9: **end for**
- 10: **end if**

If ARS works perfectly, then achieves good error by **rapid mixing for Ising model** and **speedup of** *k***-Glauber**. Remains to show:

- 1. Approximate rejection sampling has small error when  $||J_{R\times R}||_F$  small.
- 2. Recursion is subcritical.

### **Algorithm 2** Approximate rejection sampler (**ApproxRejectionSampler**)

- 1: **Input:** Oracle sampler for Q, function g s.t.  $\frac{dP}{dQ} \propto e^g$ , parameter c.
- 2: repeat
- 3: Draw  $X, Z \sim Q$ .
- 4: Let  $R = \exp(g(X) g(Z))$ .
- 5: **until**  $U \leq \frac{1}{c}R$  where  $U \sim \text{Uniform}([0,1])$

# Theorem (Hanson-Wright Inequality)

For X with independent, mean-0, K-subgaussian coordinates,  $A \in \mathbb{R}^{n \times n}$ ,

$$\mathbb{P}\left(\left|\left\langle X,AX\right\rangle - \mathbb{E}\left\langle X,AX\right\rangle\right| \geq t\right) \leq 2\exp\left[-c\min\left\{\frac{t^2}{K^4\left\|A\right\|_F^2},\frac{t}{K^2\left\|A\right\|}\right\}\right].$$

When 
$$\left\|J^{\bigotimes}\right\|$$
 small, take  $Q(x) \propto e^{\langle h, x \rangle}$ ,  $g(x) = \frac{1}{2} \langle x, Jx \rangle$ ?

### Algorithm 2 Approximate rejection sampler (ApproxRejectionSampler)

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### Theorem ([Sambale and Sinulis, 2019])

For  $X \in \{\pm 1\}^n$  with independent coordinates and  $g: \{\pm 1\}^n \to \mathbb{R}$ ,

$$\mathbb{P}\left(\left|g - \mathbb{E}g\right| \geq t\right) \leq 2\exp\left[-c\min\left\{\frac{t^2}{\mathbb{E}[\left\|\nabla g\right\|^2]}, \frac{t}{\max_{\mathsf{x} \in \{\pm 1\}^n} \left\|\nabla^2 g\right\|_{\mathit{F}}}\right\}\right].$$

The discrete gradient is defined by  $(\nabla g)_i(x) = \frac{1}{2}[g(x_{i\leftarrow 1}) - g(x_{i\leftarrow -1})].$ 

$$Q = \mu_{h+\widehat{h}} = e^{\left\langle h+\widehat{h},x\right\rangle}$$

$$g = \ln \frac{dP}{dQ}(+\text{const.}) = \frac{1}{2} \left\langle x, Jx \right\rangle - \left\langle \widehat{h}, x \right\rangle$$

To make  $\mathbb{E}[\|\nabla g\|^2]$  small, need  $\nabla g$  to be centered.

$$\mathbb{E}_Q \nabla g = 0 \quad \iff \quad \mathbb{E}_{\mu_{h+\widehat{h}}} J^{\bigotimes} x = \widehat{h} \quad \iff \quad J^{\bigotimes} \tanh(h+\widehat{h}) = \widehat{h}.$$

Can solve approximately with fixed point iteration.

### Algorithm 2 Approximate rejection sampler (ApproxRejectionSampler)

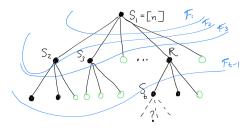
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- 2: repeat
- 3: Draw  $X, Z \sim Q$ .
- 4: Let  $R = \exp(g(X) g(Z))$ .
- 5: **until**  $U \leq \frac{1}{c}R$  where  $U \sim \text{Uniform}([0,1])$ 
  - For small enough  $||A|| \leq ||A||_F = O(1)$ ,
    - $g(X) = \langle X, AX \rangle \langle \widehat{h}, X \rangle$  has exponential tail (with large enough constant).
    - $ightharpoonup R = e^{g(X)-g(Z)}$  has power-law tail (large enough power).
  - If  $\widehat{P}$  is output of **ApproxRejectionSampler**,
    - ▶  $\mathrm{TV}(\widehat{P}, P) \leq \frac{\mathbb{E}[(R-c)\mathbb{1}_{R \geq c}]}{\mathbb{E}R}$  (tail of expectation)
    - acceptance probability is  $\geq \frac{1}{2c}$ .
  - Taking power to be  $\ln\left(\frac{n}{\varepsilon}\right)$ , we get  $\frac{\varepsilon}{n}$  error.

## 2. Analyzing the recursion

Need to bound total number of vertices in recursion tree.

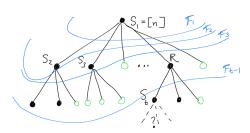
**Node:** Call of **ParallelIsingSampler**.

**Leaf:** Call to **ApproxRejectionSampler**.



- Tree depends on subsets chosen, not visited states  $y \in \{\pm 1\}^n$ .
- Expand one node at a time. Let  $\mathcal{F}_t =$  information revealed up to time t.
- Label nodes with subset  $S_t$ . Let  $D_t =$  number of children of  $S_t$ .
- Fix t. Let R be parent of  $S_t$ , |R| = m,  $|S_t| = s = \frac{c_1 m}{\ln(\frac{n}{\varepsilon}) \|J_{R \times R}\|_F}$ . If we recurse further, # children is  $T = O(\ln(\frac{n}{\varepsilon}) \frac{m}{s})$ .

# **Analyzing the recursion**



$$|R|=\textit{m,}\;|\textit{S}_t|=\textit{s}=\frac{c_1\textit{m}}{\ln\left(\frac{n}{\varepsilon}\right)||\textit{J}_{R\times R}||_F},\;\textit{T}=\textit{O}(\ln\left(\frac{n}{\varepsilon}\right)\frac{\textit{m}}{\textit{s}}),\;\textit{D}_t=\#\;\text{children of}\;\textit{S}_t.$$

$$\begin{split} \mathbb{E}[D_t|\mathcal{F}_{t-1}] \lesssim \mathbb{E}\left[\frac{\ln\left(\frac{n}{\varepsilon}\right)^2}{c_1} \left\|J_{S_t \times S_t}\right\|_F \mathbb{1}[\left\|J_{S_t \times S_t}\right\|_F > c|\mathcal{F}_{t-1}]\right] \\ &= O\left(\frac{1}{c_1}\ln\left(\frac{n}{\varepsilon}\right)^2 \mathbb{E}[\left\|J_{S_t \times S_t}\right\|_F^2|\mathcal{F}_{t-1}]\right) \quad \text{(Cauchy-Schwarz)} \end{split}$$

$$\mathbb{E}\left[\left\|J_{S_{t}\times S_{t}}\right\|_{F}^{2}\left|\mathcal{F}_{t-1}\right]=\mathbb{E}_{S\sim\mathsf{Uniform}\binom{R}{s}}\left[\left\|J_{S\times S}\right\|_{F}^{2}\right]=\left(\frac{s}{m}\right)^{2}\left\|J_{S\times S}\right\|_{F}^{2}\leq\frac{c_{1}^{2}}{\ln\left(\frac{n}{\varepsilon}\right)^{2}}$$

$$\mathbb{E}[D_{t}|\mathcal{F}_{t-1}]=O(c_{1}).$$

# **Analyzing the recursion**

$$|R| = m, |S_t| = s = \frac{c_1 m}{\ln(\frac{n}{\varepsilon}) \|J_{R \times R}\|_F}, T = O(\ln(\frac{n}{\varepsilon}) \frac{m}{s}), D_t = \# \text{ children of } S_t.$$

$$\mathbb{E}[D_t | \mathcal{F}_{t-s}] \leq \mathbb{E}\left[\frac{\ln(\frac{n}{\varepsilon})^2}{s} \|J_{S-s}\|_F + \frac{1}{s}\|J_{S-s}\|_F + \frac{1}{s}\|J_{S-s}\|$$

$$\begin{split} \mathbb{E}[D_t | \mathcal{F}_{t-1}] \lesssim \mathbb{E}\left[ \frac{\ln\left(\frac{n}{\varepsilon}\right)^2}{c_1} \left\| J_{S_t \times S_t} \right\|_F \mathbb{1}[\left\| J_{S_t \times S_t} \right\|_F > c | \mathcal{F}_{t-1}] \right] \\ &= O\left( \frac{1}{c_1} \ln\left(\frac{n}{\varepsilon}\right)^2 \mathbb{E}[\left\| J_{S_t \times S_t} \right\|_F^2 | \mathcal{F}_{t-1}] \right) \quad \text{(Cauchy-Schwarz)} \end{split}$$

$$\mathbb{E}\left[\left\|J_{S_{t}\times S_{t}}\right\|_{F}^{2}\left|\mathcal{F}_{t-1}\right] = \mathbb{E}_{S\sim\mathsf{Uniform}\binom{R}{s}}\left[\left\|J_{S\times S}\right\|_{F}^{2}\right] = \left(\frac{s}{m}\right)^{2}\left\|J_{S\times S}\right\|_{F}^{2} \leq \frac{c_{1}^{2}}{\ln\left(\frac{n}{\varepsilon}\right)^{2}}$$

$$\mathbb{E}[D_{t}|\mathcal{F}_{t-1}] = O(c_{1}).$$

- **Key:** When take subsets of R of size p|R|, expected Frobenius norm is  $O(p^2)$  but number of steps is only  $O\left(\frac{1}{n}\right)$ .
- For  $c_1 \ll 1$ ,  $\mathbb{E}[D_t | \mathcal{F}_{t-1}] < 1$ , this is a subcritical branching process.
- Martingale concentration: w.h.p. number of vertices (runtime) is bounded by poly  $\log \left(\frac{n}{n}\right) \|J\|_{F}$ .

# Extension: Mixed p-spin model

#### **Definition**

The **mixed** *p*-**spin model** with coefficients  $\beta_2, \beta_3, \ldots$  is the random measure

$$\mu(x) \propto \exp\left(\sum_{p=2}^{\infty} \frac{\beta_p}{n^{\frac{p-1}{2}}} \sum_{1 \leq i_1 \leq \cdots \leq i_p} g_{i_1, \cdots, i_p} x_{i_1} \cdots x_{i_p} + \sum_{i=1}^n h_i x_i\right), \quad x \in \{\pm 1\}^n$$

where  $g_{i_1,\cdots,i_p} \sim N(0,1)$ .

### Theorem ([Adhikari et al., 2022, Anari et al., 2023c])

There is an absolute constant A such that if  $\sum_{p=2}^{\infty} \sqrt{p^3 \ln p} \cdot \beta_p \leq A$  and  $\sum_{p=2}^{\infty} \sqrt{2^p p^3 \ln p} \cdot \beta_p = B < \infty$ , when with probability  $\geq 1 - \exp(-\Omega(n))$  over g,  $\mu$  satisfies approximate tensorization of entropy with constant  $O_B(1)$ .

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By using concentration of polynomials rather than quadratics, we get:

#### Theorem

In the above setting, w.h.p. there is an algorithm which outputs a sample  $\varepsilon$ -close in TV distance from  $\mu$  and, with probability  $\geq 1-\varepsilon$ , runs in time

$$O\left(\sqrt{n}\operatorname{poly}\log\left(\frac{n}{\varepsilon}\right)\right)$$

on a parallel machine with poly  $\left(\frac{n}{\varepsilon}\right)$  processors.

## **Conclusion and open questions**

- 1. *k*-Glauber dynamics gives a generic parallel speedup for discrete Markov chains, *if* we can implement each step.
- 2. Implemented for Ising (& p-spin) model in the regime of rapid mixing.

### Open questions

• Fast parallel sampling under generic smoothness assumptions?

# Theorem ([Anari et al., 2023c])

There is an absolute constant A>0 such that for  $\mu(x)\propto e^{H(x)}$ ,  $\beta:=\max_{x\in\{\pm 1\}^n}\left\|\nabla^2 H(x)\right\|_2\leq A$ , then  $\mu$  has spectral gap  $\geq\frac{1}{(1+O(\beta))n}$ .

- Analysis of "gradient-based" discrete sampling algorithms [Grathwohl et al., 2021, Zhang et al., 2022, Rhodes and Gutmann, 2022].
- Other settings where *k*-Glauber dynamics can be efficiently approximated?

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# Inspiration from the continuous analogue

#### Lemma

Suppose that  $X_1 \sim \mu_1, X_2 \sim \mu_2$  are distributions on  $\mathbb{R}^n$  with Poincaré constants  $C_1, C_2$ . Then  $X_1 + X_2 \sim \mu_1 * \mu_2$  has Poincaré constant bounded by  $C_1 + C_2$ .

**Example.**  $\mu_2 = N(0, C_2)$ .

Proof.

- 1. **Scaling.**  $m_i X_i$  has Poincaré constant  $m_i^2 C_i$ .
- 2. **Tensorization.**  $(m_1X_1, m_2X_2)$  has Poincaré constant  $\leq C = \max\{m_1^2C_1, m_2^2C_2\}.$
- 3. **Projection.**  $X + Y = (m_1 X_1, m_2 X_2) \cdot \left(\frac{1}{m_1}, \frac{1}{m_2}\right)$  has Poincaré constant  $\leq C$  when  $\left\|\left(\frac{1}{m_1}, \frac{1}{m_2}\right)\right\| = 1$ .

Choose 
$$\frac{1}{m_1^2} = \frac{C_1}{C_1 + C_2}$$
,  $\frac{1}{m_2^2} = \frac{C_2}{C_1 + C_2}$ .