### Breast Cancer Diagnosis

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2023-02-27

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## Background

#### **Breast Cancer Diagnosis**

- ► The data is the breast cancer medical data retrieved from "breast-cancer.csv", which has 569 rows and 32 columns.
- ▶ The first column ID labels individual breast tissue images; The second column Diagnosis identifies if the image is coming from cancer tissue or benign cases (M = malignant, B = benign). There are 357 benign and 212 malignant cases.
- ▶ The other 30 columns correspond to mean, standard deviation and the largest values (points on the tails) of the distributions of the following 10 features computed for the cellnuclei;

## **Objectives**

- ► The goal of the exercise is to build a predictive model based on logistic regression to facilitate cancer diagnosis.
- ► We will move towards the goal with the steps of task 1, 2, 3 and 4.

### Task 1 - Objective

#### **Objective:**

Build a logistic model to classify the images into malignant/benign, and write down your likelihood function, its gradient and Hessian matrix.

Define the "Diagnosis" variable will be coded as 1 for malignant cases and 0 for benign cases.

Given n i.i.d. observations with p predictors, we consider a logistic regression model

$$P(Y_i = 1 \mid \mathbf{x}_i) = \frac{e^{\mathbf{x}_i^\top \beta}}{1 + e^{\mathbf{x}_i^\top \beta}}, \ i = 1, \dots, n$$
 (1)

where  $\boldsymbol{\beta} = (\beta_0, \beta_1, \dots, \beta_p)^{\top} \in \mathbb{R}^{p+1}$  is the parameter vector,  $\mathbf{x}_i = (1, X_{i1}, \dots, X_{ip})^{\top}$  is the vector of predictors in the *i*-th observation, and  $Y_i \in \{0,1\}$  is the binary response in the *i*-th observation.

Let  $\mathbf{y} = (Y_1, Y_2, \dots, Y_n)^{\top}$  denote the response vector, and  $\mathbf{X} = (\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n)^{\top} \in \mathbb{R}^{n \times (p+1)}$  denote the design matrix.

The observed likelihood of  $\{(Y_1, \mathbf{x}_1), (Y_2, \mathbf{x}_2), \dots, (Y_n, \mathbf{x}_n)\}$  is

$$L(\beta; \mathbf{y}, \mathbf{X}) = \prod_{i=1}^{n} \left| \left( \frac{e^{\mathbf{x}_{i}^{\top} \beta}}{1 + e^{\mathbf{x}_{i}^{\top} \beta}} \right)^{Y_{i}} \left( \frac{1}{1 + e^{\mathbf{x}_{i}^{\top} \beta}} \right)^{1 - Y_{i}} \right|$$

Maximizing the likelihood is equivalent to maximizing the log-likelihood function:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[ Y_{i} \mathbf{x}_{i}^{\top} \beta - \log \left( 1 + e^{\mathbf{x}_{i}^{\top} \beta} \right) \right].$$
 (2)

The estimates of model parameters are

$$\widehat{oldsymbol{eta}} = \arg\max_{oldsymbol{eta}} \ f(oldsymbol{eta}; \mathbf{y}, \mathbf{X}),$$

and the optimization problem is

$$\max_{\beta} f(\beta; \mathbf{y}, \mathbf{X}). \tag{3}$$

Denote  $p_i = P(Y_i = 1 \mid \mathbf{x}_i)$  as given in (1) and  $\mathbf{p} = (p_1, p_2, \dots, p_n)^{\top}$ . The gradient of f is:

$$\nabla f(\beta; \mathbf{y}, \mathbf{X}) = \mathbf{X}^{\top} (\mathbf{y} - \mathbf{p})$$

$$= \sum_{i=1}^{n} (Y_i - p_i) \mathbf{x}_i$$

$$= \begin{pmatrix} \sum_{i=1}^{n} (Y_i - p_i) \\ \sum_{i=1}^{n} (Y_i - p_i) X_{i1} \\ \vdots \\ \sum_{i=1}^{n} (Y_i - p_i) X_{ip} \end{pmatrix}$$

Denote  $w_i = p_i(1 - p_i) \in (0, 1)$  and  $\mathbf{W} = \operatorname{diag}(w_1, \dots, w_n)$ . The Hessian matrix of f is given by

$$\nabla^{2} f(\beta; \mathbf{y}, \mathbf{X}) = -\mathbf{X}^{\top} \mathbf{W} \mathbf{X} 
= -\sum_{i=1}^{n} w_{i} \mathbf{x}_{i} \mathbf{x}_{i}^{\top} 
= -\begin{pmatrix}
\sum_{i=1}^{n} w_{i} & \sum_{i=1}^{n} w_{i} X_{i1} & \cdots & \sum_{i=1}^{n} w_{i} X_{i1} \\
\sum_{i=1}^{n} w_{i} X_{i1} & \sum_{i=1}^{n} w_{i} X_{i1}^{2} & \cdots & \sum_{i=1}^{n} w_{i} X_{i1} X_{ip} \\
\vdots & \vdots & \ddots & \vdots \\
\sum_{i=1}^{n} w_{i} X_{ip} & \sum_{i=1}^{n} w_{i} X_{in} X_{i1} & \cdots & \sum_{i=1}^{n} w_{i} X_{ip}^{2}
\end{pmatrix}$$

Next, we show that the Hessian matrix  $\nabla^2 f(\beta; \mathbf{y}, \mathbf{X})$  is a negative-definite matrix if  $\mathbf{X}$  has full rank.

**Proof.** For any (p+1)-dimensional nonzero vector  $\alpha$ , given that **X** has full rank,  $\mathbf{X}\alpha$  is also a nonzero vector. Since **W** is positive-definite, we have

$$egin{aligned} oldsymbol{lpha}^{ op} 
abla^2 f(eta; \mathbf{y}, \mathbf{X}) &lpha = oldsymbol{lpha}^{ op} (-\mathbf{X}^{ op} \mathbf{W} \mathbf{X}) oldsymbol{lpha} \\ &= -(\mathbf{X} oldsymbol{lpha})^{ op} \mathbf{W} (\mathbf{X} oldsymbol{lpha}) \\ &< 0. \end{aligned}$$

Thus,  $\nabla^2 f(\beta; \mathbf{y}, \mathbf{X})$  is negative-definite.

Hence, the optimization problem (3) is a well-defined problem.

### Task 1 - Logistic model R code

Task 1 - Logistic model R code

## Task 2 - Object

#### Objective:

Develop a Newton-Raphson algorithm to estimate your model

Recall The target function f given in task 1 is:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[ Y_{i} \mathbf{x}_{i}^{\top} \beta - \log \left( 1 + e^{\mathbf{x}_{i}^{\top} \beta} \right) \right]. \tag{4}$$

We develop a modified Newton-Raphson algorithm including a step-halving step to maximize the target function.

# Task 2 - Newton-Raphson algorithm design

#### **Algorithm 1** Newton-Raphson algorithm

**Require:**  $f(\beta)$  - target function as given in (4);  $\beta_0$  - starting value

**Ensure:** 
$$\widehat{\boldsymbol{\beta}}$$
 such that  $\widehat{\boldsymbol{\beta}} \approx \arg \max_{\boldsymbol{\beta}} f(\boldsymbol{\beta})$ 

 $i \leftarrow 0$ , where i is the current number of iterations

$$f(\beta_{-1}) \leftarrow -\infty$$
  
**while** convergence criterion is not met **do**

while convergence criterion is not met do  $i \leftarrow i + 1$ 

$$\mathbf{d}_i \leftarrow -[\nabla^2 f(\beta_{i-1})]^{-1} \nabla f(\beta_{i-1})$$
, where  $\mathbf{d}_i$  is the direction in

the i-th iteration

$$\lambda_i \leftarrow 1$$
, where  $\lambda_i$  is the multiplier in the *i*-th iteration  $eta_i \leftarrow eta_{i-1} + \lambda_i \mathbf{d}_i$  while  $f(eta_i) \leq f(eta_{i-1})$  do  $\lambda_i \leftarrow \lambda_i/2$ 

$$eta_i \leftarrow eta_{i-1} + \lambda_i \mathbf{d}_i$$
 end while

end while

$$\hat{\boldsymbol{\beta}} \leftarrow \boldsymbol{\beta}_i$$

### Task 2 - Newton-Raphson algorithm R code

### The corresponding R-function for the algorithm

```
NewtonRaphson <- function(dat, func, start, tol = 1e-8, maxiter = 200) {
 i <- 0
 cur <- start
 stuff <- func(dat, cur)
 res <- c(0, stuff$f, cur)
 prevf <- -Inf
 X <- cbind(rep(1, nrow(dat)), as.matrix(dat[, -1]))</pre>
 v <- dat[, 1]
 warned <- 0
 while (abs(stuff$f - prevf) > tol && i < maxiter) {
    i <- i + 1
    prevf <- stuff$f
    prev <- cur
    d <- -solve(stuff$Hess) %*% stuff$grad
    cur <- prev + d
   lambda <- 1
    maxhalv <- 0
    while (func(dat, cur)$f < prevf && maxhalv < 50) {
      maxhalv <- maxhalv + 1
     lambda <- lambda / 2
     cur <- prev + lambda * d
    stuff <- func(dat, cur)
    res <- rbind(res, c(i, stuff$f, cur))
    v_hat <- ifelse(X %*% cur > 0, 1, 0)
    if (warned == 0 && sum(y - y_hat) == 0) {
      warning("Complete separation occurs, Algorithm does not converge,")
      warned <- 1
    }
  colnames(res) <- c("iter", "target function", "(Intercept)", names(dat)[-1])
 return(res)
```

### Task 2 - Complete separtion

Sometimes our algorithm does not converge because of the complete separation.

A complete separation in a logistic regression, sometimes also referred as perfect prediction, occurs whenever there exists some vector of coefficients  $\boldsymbol{\beta}$  such that  $Y_i=1$  whenever  $\mathbf{x}_i^{\top}\boldsymbol{\beta}>0$  and  $Y_i=0$  whenever  $\mathbf{x}_i^{\top}\boldsymbol{\beta}\leq 0$ .

Complete separation occur when a linear function of predictors can perfectly classify the response.

## Task 2 - Comparison

When there is no complete separation, the parameters output by glm function and our algorithm are the same.

## Task 2 - Comparison

### Comparison of using glm function and our algorithm (part1)

predictor	ours	glm
(Intercept)	111.7230206	90.9690365
radius_mean	-3646.8235387	-2560.3938902
texture_mean	-2.8949199	0.8812037
perimeter_mean	1257.9481173	789.2724398
area_mean	2091.5001210	1539.8345650
smoothness_mean	180.4591375	128.8762247
compactness_mean	-471.3749334	-346.6691873
concavity_mean	-0.3063034	9.0810082
concave.points_mean	287.3555092	215.4808031
symmetry_mean	10.8459784	10.4772590
fractal_dimension_mean	-40.6413497	-28.6916549
radius_se	854.4290933	617.5687358
texture_se	-105.8920782	-79.9788638
perimeter_se	-1231.6514585	-917.3916527
area_se	789.0550146	628.6222313

## Task 2 - Comparison

### Comparison of using glm function and our algorithm (part2)

predictor	ours	glm
smoothness_se	-83.34497	-63.76604
compactness_se	473.19705	344.31802
concavity_se	-440.66293	-323.85337
concave.points_se	503.37287	374.76109
symmetry_se	-144.45806	-108.62118
fractal_dimension_se	-462.42425	-339.16464
radius_worst	338.04088	197.93147
texture_worst	211.28178	155.07873
perimeter_worst	1406.29842	1068.10688
area_worst	-851.74931	-606.16579
smoothness_worst	-42.11738	-26.47595
compactness_worst	-480.79091	-346.40518
concavity_worst	518.45952	373.20892
concave.points worst	-216.94150	-161.61771
symmetry_worst	98.05014	71.94885
fractal_dimension_worst	384.83305	282.82535

Task 3

Task 4

#### **Discussions**

- ▶ There is much freedom when designing the simulations.
- ▶ In our algorithm, we have 5 parameters. n, p, ratio, c, corr
- ► More parameters can be adjusted.

#### Limitations and Future Work

- ► Limitation: We reproduced high-dimensional scenarios, but we still don't know the solution.
- ► Future Work: We may adjust other parameters to investigate further.

#### Reference

 Li Y, Hong HG, Ahmed SE, Li Y. Weak signals in high-dimensional regression: Detection, estimation and prediction. Appl Stochastic Models Bus Ind. 2018;1–16. https://doi.org/10.1002/asmb.2340 Q&A

- ► Thanks for listening!
- ► Any questions?