Breast Cancer Diagnosis

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Background

Breast Cancer Diagnosis:

- In this project we study the breast cancer diagnosis problem
- The goal of the exercise is to build a predictive model based on logistic regression to facilitate cancer diagnosis
- We move towards the goal with the steps of task 1, 2, 3 and 4.

Background

Data Source:

- The data is the breast cancer medical data retrieved from "breast-cancer.csv", which has 569 rows and 32 columns
- The first column `ID` labels individual breast tissue images; The second column `Diagnosis` identifies if the image is coming from cancer tissue or benign cases (M = malignant, B = benign). There are 357 benign and 212 malignant cases
- The other 30 columns correspond to mean, standard deviation and the largest values (points on the tails) of the distributions of the following 10 features computed for the cellnuclei

Task Introduction

- Task 1: Build a logistic model to classify the images into malignant/benign, and write down your likelihood function, its gradient and Hessian matrix.
- Task 2: Develop a Newton-Raphson algorithm to estimate your model.
- Task 3: Build a logistic-LASSO model to select features, and implement a path-wise coordinate-wise optimization algorithm to obtain a path of solutions with a sequence of descending λ's.
- Task 4: Use 5-fold cross-validation to select the best λ. Compare the prediction performance between the 'optimal' model and 'full' model

Task 1 - Objective

Objective:

Build a logistic model to classify the images into malignant/benign, and write down your likelihood function, its gradient and Hessian matrix.

Task 1 - Build a logistic model

Define the "Diagnosis" variable will be coded as 1 for malignant cases and 0 for benign cases.

Given n i.i.d. observations with p predictors, we consider a logistic regression model

$$P(Y_i = 1 \mid \mathbf{x}_i) = \frac{e^{\mathbf{x}_i^\top \beta}}{1 + e^{\mathbf{x}_i^\top \beta}}, \ i = 1, \ldots, n \tag{1}$$

where $\beta = (\beta_0, \beta_1, \ldots, \beta_p)^\top \in \mathbb{R}^{p+1}$ is the parameter vector, $\mathbf{x}_i = (1, X_{i1}, \ldots, X_{ip})^\top$ is the vector of predictors in the i-th observation, and $Y_i \in \{0, 1\}$ is the binary response in the i-th observation.

Task 1 - Build a logistic model

Let $\mathbf{y} = (Y_1, Y_2, \ldots, Y_n)^\top$ denote the response vector, and $\mathbf{X} = (\mathbf{x}_1, \mathbf{x}_2, \ldots, \mathbf{x}_n)^\top \in \mathbb{R}^{n \times (p+1)}$ denote the design matrix.

The observed likelihood of $\{(Y_1, \mathbf{x}_1), (Y_2, \mathbf{x}_2) \ldots, (Y_n, \mathbf{x}_n)\}$ is

$$L(\beta; \mathbf{y}, \mathbf{X}) = \prod_{i=1}^{n} \left[\left(\frac{e^{\mathbf{x}_i^\top \beta}}{1 + e^{\mathbf{x}_i^\top \beta}} \right)^{Y_i} \left(\frac{1}{1 + e^{\mathbf{x}_i^\top \beta}} \right)^{1 - Y_i} \right]$$

Task 1 - Build a logistic model

Maximizing the likelihood is equivalent to maximizing the log-likelihood function:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[Y_i \mathbf{x}_i^\top \beta - \log \left(1 + e^{\mathbf{x}_i^\top \beta} \right) \right]. \tag{2}$$

The estimates of model parameters are

$$\widehat{\beta} = \arg\max_{\beta} \ f(\beta; \mathbf{y}, \mathbf{X}),$$

and the optimization problem is

$$\max_{\beta} \ f(\beta; \mathbf{y}, \mathbf{X}). \tag{3}$$

Task 1 - Build a logistic model

Denote $p_i = P(Y_i = 1 \mid \mathbf{x}_i)$ as given in (1) and $\mathbf{p} = (p_1, p_2, \ldots, p_n)^\top$. The gradient of f is:

$$\begin{aligned}
\nabla f(\beta; \mathbf{y}, \mathbf{X}) &= \mathbf{X}^\top (\mathbf{y} - \mathbf{p}) \\
&= \sum_{i=1}^{n} (Y_i - p_i)\mathbf{x}_i \\
&= \begin{pmatrix} \sum_{i=1}^{n} (Y_i - p_i) \\ \sum_{i=1}^{n} (Y_i - p_i) X_{i1} \\ \vdots \\ \sum_{i=1}^{n} (Y_i - p_i) X_{ip} \end{pmatrix}
\end{aligned}$$

Task 1 - Build a logistic model

Denote $w_i = p_i(1 - p_i) \in (0, 1)$ and $\mathbf{W} = \mathrm{diag}(w_1, \ldots, w_n)$. The Hessian matrix of f is given by

$$\begin{aligned}
\nabla^2 f(\beta; \mathbf{y}, \mathbf{X}) &= -\mathbf{X}^\top \mathbf{W} \mathbf{X} \\
&= -\sum_{i=1}^{n} w_i \mathbf{x}_i \mathbf{x}_i^\top \\
&= -\begin{pmatrix}
\sum_{i=1}^n w_i & \sum_{i=1}^n w_i X_{i1} & \cdots & \sum_{i=1}^n w_i X_{i1} \\
\sum_{i=1}^n w_i X_{i1} & \sum_{i=1}^n w_i X_{i1}^2 & \cdots & \sum_{i=1}^n w_i X_{i1} X_{ip} \\
\vdots & \vdots & \ddots & \vdots \\
\sum_{i=1}^n w_i X_{ip} & \sum_{i=1}^n w_i X_{in} X_{i1} & \cdots & \sum_{i=1}^n w_i X_{ip}^2
\end{pmatrix}
\end{aligned}$$

Task 1 - Build a logistic model

Next, we show that the Hessian matrix $\nabla^2 f(\beta; \mathbf{y}, \mathbf{X})$ is a negative-definite matrix if $\mathbf{X}$ has full rank.

Proof. For any $(p+1)$-dimensional nonzero vector $\boldsymbol{\alpha}$, given that $\mathbf{X}$ has full rank, $\mathbf{X}\boldsymbol{\alpha}$ is also a nonzero vector. Since $\mathbf{W}$ is positive-definite, we have

$$\begin{aligned}\boldsymbol{\alpha}^\top \nabla^2 f(\beta; \mathbf{y}, \mathbf{X})\boldsymbol{\alpha} &= \boldsymbol{\alpha}^\top(-\mathbf{X}^\top \mathbf{W}\mathbf{X})\boldsymbol{\alpha} \\ &= -(\mathbf{X}\boldsymbol{\alpha})^\top \mathbf{W}(\mathbf{X}\boldsymbol{\alpha}) \\ &< 0.\end{aligned}$$

Thus, $\nabla^2 f(\beta; \mathbf{y}, \mathbf{X})$ is negative-definite. □

Hence, the optimization problem (3) is a well-defined problem.

Task 1 - Logistic model R code

```
cancer <- read.csv("breast-cancer.csv") %>%
janitor::clean_names()%>%
select(-1,-33) %>%
mutate(diagnosis = recode(diagnosis, "M" = 1, "B" = 0))
cor()
ggcorrplot(corr, type = "upper", tl.cex = 8)

set.seed(1)
trainRows <- createDataPartition(y = cancer$diagnosis, p = 0.8, list = FALSE)
train <- cancer[trainRows, ]
test <-  cancer[-trainRows, ]
glm.fit <- glm(diagnosis ~ .,
               data = train,
               subset = trainRows,
               family = binomial(link = "logit"))
summary(glm.fit)
pred <- predict(glm.fit, newdata = test, type = "response")
y_test <- factor(test$diagnosis)
auc_full <- auc(y_test, pred)
auc_full
```

- Final AUC of full model reaches 0.9641.
- And if we remove correlated variables, the AUC is uplifted to 0.9962.

Task 2 - Objective

Objective:

Develop a Newton-Raphson algorithm to estimate your model

- Recall The target function f given in task 1 is:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[Y_i \mathbf{x}_i^\top \beta - \log\left(1 + e^{\mathbf{x}_i^\top \beta}\right) \right].$$

- We develop a modified Newton-Raphson algorithm including a step-halving step to maximize the target function.

Task 2 - Newton-Raphson algorithm design

Algorithm 1 Newton-Raphson algorithm

Require: $f(\beta)$ - target function as given in (4); β_0 - starting value
Ensure: $\widehat{\beta}$ such that $\widehat{\beta} \approx \arg\max_\beta \ f(\beta)$

$i \leftarrow 0$, where i is the current number of iterations
$f(\beta_{-1}) \leftarrow -\infty$
while convergence criterion is not met **do**
    $i \leftarrow i + 1$
    $\mathbf{d}_i \leftarrow -[\nabla^2 f(\beta_{i-1})]^{-1} \nabla f(\beta_{i-1})$, where $\mathbf{d}_i$ is the direction in the i-th iteration
    $\lambda_i \leftarrow 1$, where λ_i is the multiplier in the i-th iteration
    $\beta_i \leftarrow \beta_{i-1} + \lambda_i \mathbf{d}_i$
    while $f(\beta_i) \leq f(\beta_{i-1})$ **do**
        $\lambda_i \leftarrow \lambda_i / 2$
        $\beta_i \leftarrow \beta_{i-1} + \lambda_i \mathbf{d}_i$
    end while
end while
$\widehat{\beta} \leftarrow \beta_i$

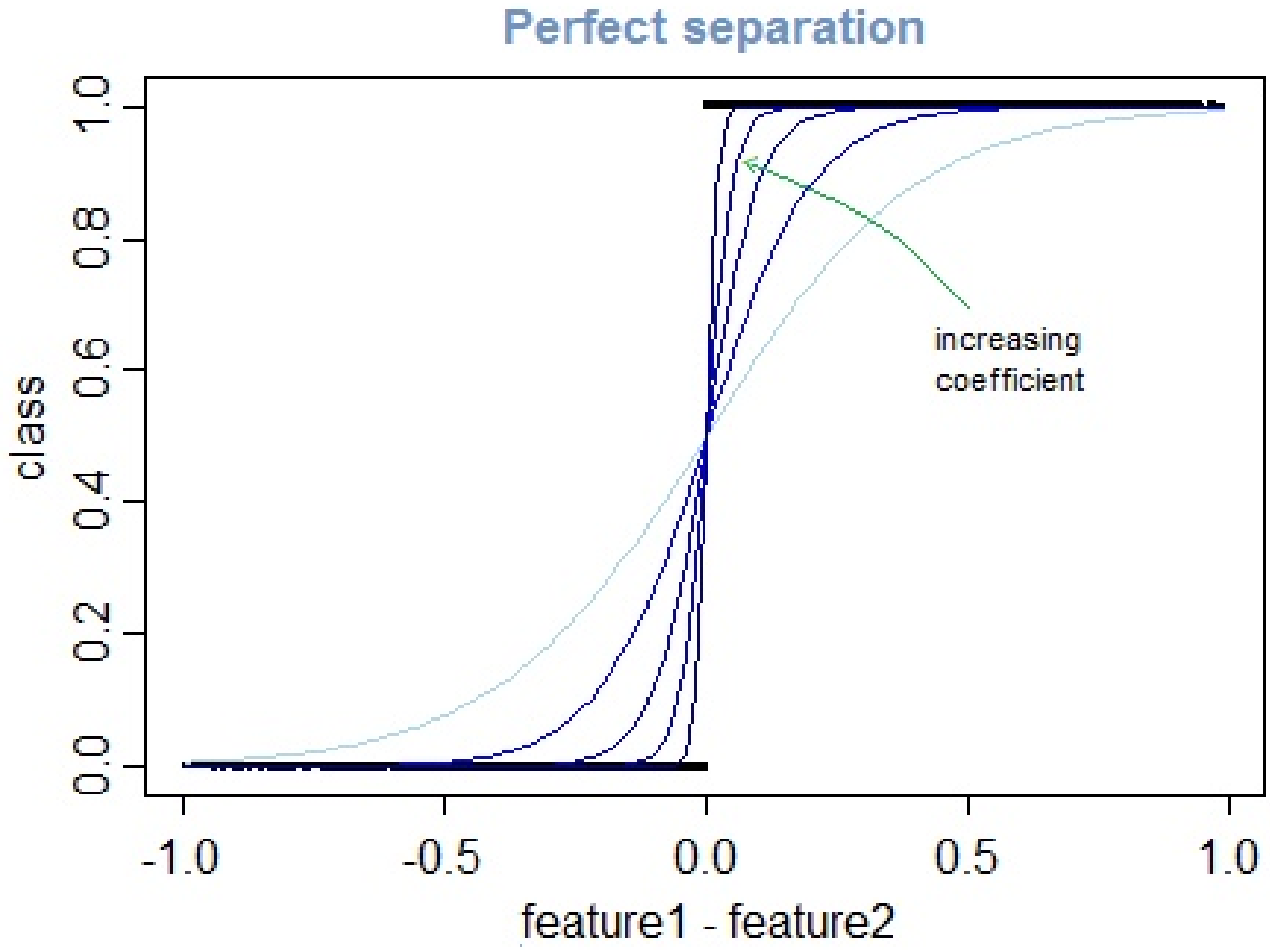
Task 2 - Newton-Raphson algorithm R code

```r
NewtonRaphson <- function(dat, func, start, tol = 1e-8, maxiter = 200) {
  i <- 0
  cur <- start
  stuff <- func(dat, cur)
  res <- c(0, stuff$f, cur)
  prevf <- -Inf
  X <- cbind(rep(1, nrow(dat)), as.matrix(dat[, -1]))
  y <- dat[, 1]
  warned <- 0
  while (abs(stuff$f - prevf) > tol && i < maxiter) {
    i <- i + 1
    prevf <- stuff$f
    prev <- cur
    d <- -solve(stuff$Hess) %*% stuff$grad
    cur <- prev + d
    lambda <- 1
    maxhalv <- 0
    while (func(dat, cur)$f < prevf && maxhalv < 50) {
      maxhalv <- maxhalv + 1
      lambda <- lambda / 2
      cur <- prev + lambda * d
    }
    stuff <- func(dat, cur)
    res <- rbind(res, c(i, stuff$f, cur))
    y_hat <- ifelse(X %*% cur > 0, 1, 0)
    if (warned == 0 && sum(y - y_hat) == 0) {
      warning("Complete separation occurs. Algorithm does not converge.")
      warned <- 1
    }
  }
  colnames(res) <- c("iter", "target_function", "(Intercept)", names(dat)[-1])
  return(res)
}
```

Task 2 - Complete separtion

- Sometimes our algorithm does not converge because of the complete separation.
- A complete separation in a logistic regression, sometimes also referred as perfect prediction, occurs whenever there exists some vector of coefficients β such that $Y_i = 1$ whenever $\mathbf{x}_i^\top \beta > 0$ and $Y_i = 0$ whenever $\mathbf{x}_i^\top \beta \leq 0$.
- Complete separation occur when a linear function of predictors can perfectly classify the response.

Task 2 - Complete separation



Task 2 - Complete separation

- We have proved that: when there exists a vector of coefficients $\hat{\beta}$ such that $Y_i = 1$ whenever $\mathbf{x}_i^\top \hat{\beta} > 0$ and $Y_i = 0$ whenever $\mathbf{x}_i^\top \hat{\beta} \leq 0$, there does not exist $\beta^* \in \mathbb{R}^{(p+1)}$ such that $\beta^* = \arg\max_\beta f(\beta)$, where f is given in (4). Thus our algorithm does not converge. (proof is attached in report appendix)
- If there is no complete separation, the parameters output by `glm` function and our algorithm are demonstrated to be the same
- In practice, if complete separation occurs, we randomly pick the parameters which satisfy complete separation as our full model

Task 2 - Comparison

Comparison of using `glm` function and our algorithm (part of)

predictor	ours	glm
(Intercept)	111.7230206	90.9690365
radius_mean	-3646.8235387	-2560.3938902
texture_mean	-2.8949199	0.8812037
perimeter_mean	1257.9481173	789.2724398
area_mean	2091.5001210	1539.8345650
smoothness_mean	180.4591375	128.8762247
compactness_mean	-471.3749334	-346.6691873
concavity_mean	-0.3063034	9.0810082
concave.points_mean	287.3555092	215.4808031
symmetry_mean	10.8459784	10.4772590
fractal_dimension_mean	-40.6413497	-28.6916549
radius_se	854.4290933	617.5687358
texture_se	-105.8920782	-79.9788638
perimeter_se	-1231.6514585	-917.3916527
area_se	789.0550146	628.6222313
smoothness_se	-83.3449730	-63.7660367
compactness_se	473.1970464	344.3180183
concavity_se	-440.6629325	-323.8533730
concave.points_se	503.3728692	374.7610875
symmetry_se	-144.4580558	-108.6211792

Task 3 - Objective

Objective:

Build a logistic-LASSO model to select features by implementing a path-wise coordinate-wise optimization algorithm

- Recall Log-likelihood f in task 1:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[Y_i \mathbf{x}_i^\top \beta - \log \left(1 + e^{\mathbf{x}_i^\top \beta} \right) \right]. \tag{4}$$

- LASSO estimates the logistic model parameters β by optimizing a penalized loss function:

$$\min_{\beta} \; -\frac{1}{n} f(\beta) + \lambda \sum_{k=1}^{p} |\beta_k|. \tag{5}$$

where $\lambda \geq 0$ is the tuning parameter. Note that the intercept is not penalized and all predictors are standardized.

Task 3 - Algorithm of Logistic-LASSO Model

Algorithm Structure:

- OUTER LOOP: Decrement λ.
- MIDDLE LOOP: Update $\tilde{w}_i$, $\tilde{p}_i$, and thus the quadratic approximation ℓ using the current parameters $\tilde{\beta}$.
- INNER LOOP: Run the coordinate descent algorithm on the penalized weighted-least-squares problem.

Task 3 - OUTER LOOP

OUTER LOOP:

Compute the solutions of the optimization problem (5) for a decreasing sequence of values for λ: $\{\lambda_1, \ldots, \lambda_m\}$, starting at the smallest value $\lambda_1 = \lambda_{max}$

$$\lambda_{max} = \frac{1}{n} \max_j |\langle \mathbf{x}_{\cdot j}, \mathbf{y} \rangle| , \tag{6}$$

where $\mathbf{x}_{\cdot j}$ is the j-th column of the design matrix $\mathbf{X}$, for $j = 1, \ldots, p$.

For tuning parameter value λ_{k+1}, we initialize coordinate descent algorithm at the computed solution for λ_k (warm start).

Task 3 - MIDDLE LOOP

MIDDLE LOOP:

- For a fixed λ, find the estimates of β by solving the optimization problem (5).

Based on current parameter estimates $\tilde{\beta}$, we we form a quadratic approximation to the log-likelihood f using a Taylor expansion:

$$f(\beta) \approx \ell(\beta) = f(\tilde{\beta}) + (\beta - \tilde{\beta})^\top \nabla f(\tilde{\beta}) + \frac{1}{2}(\beta - \tilde{\beta})^\top \nabla^2 f(\tilde{\beta})(\beta - \tilde{\beta})$$

$$= -\frac{1}{2}\sum_{i=1}^{n} \tilde{w}_i \left[\mathbf{x}_i^\top(\tilde{\beta} - \beta) + \frac{Y_i - \tilde{p}_i}{\tilde{w}_i}\right] + \frac{1}{2}\sum_{i=1}^{n} \tilde{w}_i \left(\frac{Y_i - \tilde{p}_i}{\tilde{w}_i}\right)^2 + f(\tilde{\beta}),$$

where $\tilde{\mathbf{p}} = (\tilde{p}_1, \ldots, \tilde{p}_n)^\top$ and $\tilde{\mathbf{W}} = \mathrm{diag}(\tilde{w}_1, \ldots, \tilde{w}_n)$ are the estimates of $\mathbf{p}$ and $\mathbf{W}$ based on $\tilde{\beta}$.

Task 3 - MIDDLE LOOP

- Let $\tilde{z}_i = \mathbf{x}_i^\top \tilde{\beta} + \frac{Y_i - \tilde{p}_i}{\tilde{w}_i}$, we have

$$\ell(\beta) = -\frac{1}{2} \sum_{i=1}^{n} \tilde{w}_i (\tilde{z}_i - \mathbf{x}_i^\top \beta)^2 + C(\tilde{\beta}), \tag{7}$$

where is $\tilde{z}_i$ the working response, $\tilde{w}_i$ is the working weight, and C is a function that does not depend on β.

Task 3 - INNER LOOP

INNER LOOP:

With fixed $\tilde{w}_i$'s, $\tilde{z}_i$'s, and a fixed form of ℓ based on the estimates of β in the previous iteration of the middle loop, we use coordinate descent to update β by solving

$$\min_{\beta} \; -\frac{1}{n}\ell(\beta) + \lambda \sum_{k=1}^{p} |\beta_k|, \tag{8}$$

Task 3 - INNER LOOP

- Coordinate-wise objective function Based the current estimates $\tilde{\beta}_k$ for $k \neq j$:

$$\min_{\beta_j} \frac{1}{2n} \sum_{i=1}^{n} \tilde{w}_i \left(\tilde{z}_i - x_{ij}\beta_j - \sum_{k \neq j} x_{ik}\tilde{\beta}_k \right)^2 + \lambda|\beta_j| + \lambda \sum_{k \neq j} |\beta_k|.$$

- Updates:

$$\tilde{\beta}_0 \leftarrow \frac{\sum_{i=1}^{n} \tilde{w}_i(\tilde{z}_i - \sum_{k=1}^{p} x_{ik}\tilde{\beta}_k)}{\sum_{i=1}^{n} \tilde{w}_i},$$

$$\tilde{\beta}_j \leftarrow \frac{S\left(\frac{1}{n}\sum_{i=1}^{n} \tilde{w}_i x_{ij}(\tilde{z}_i - \sum_{k \neq j} x_{ik}\tilde{\beta}_k), \lambda\right)}{\frac{1}{n}\sum_{i=1}^{n} \tilde{w}_i x_{ij}^2}, \; j = 1, \ldots, p$$

where $S(z, \gamma)$ is the soft-thresholding operator with value

$$S(z, \gamma) = \text{sign}(z)(|z| - \gamma)_+ = \begin{cases} z - \gamma, & \text{if } z > 0 \text{ and } \gamma < |z| \\ z + \gamma, & \text{if } z < 0 \text{ and } \gamma < |z| \\ 0, & \text{if } \gamma \geq |z| \end{cases}$$

Keep updating estimates of β_j's repeatedly for $j = 0, 1, 2, ..., p, 0, 1, 2, ...$ until convergence.

Task 4

Discussions

- There is much freedom when designing the simulations.
- In our algorithm, we have 5 parameters. n, p, ratio, c, corr
- More parameters can be adjusted.

Limitations and Future Work

- Limitation: We reproduced high-dimensional scenarios, but we still don't know the solution.
- Future Work: We may adjust other parameters to investigate further.

Reference

1. Li Y, Hong HG, Ahmed SE, Li Y. Weak signals in high-dimensional regression: Detection, estimation and prediction. Appl Stochastic Models Bus Ind. 2018;1–16. https://doi.org/10.1002/asmb.2340

Q&A

- Thanks for listening!
- Any questions?