### Breast Cancer Diagnosis

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### Background

### **Breast Cancer Diagnosis:**

- In this project we study the breast cancer diagnosis problem
- ► The goal of the exercise is to build a predictive model based on logistic regression to facilitate cancer diagnosis
- ▶ We move towards the goal with the steps of task 1, 2, 3 and 4.

### Background

#### **Data Source:**

- ► The data is the breast cancer medical data retrieved from "breast-cancer.csv", which has 569 rows and 32 columns
- ➤ The first column ID labels individual breast tissue images; The second column Diagnosis identifies if the image is coming from cancer tissue or benign cases (M = malignant, B = benign). There are 357 benign and 212 malignant cases
- ▶ The other 30 columns correspond to mean, standard deviation and the largest values (points on the tails) of the distributions of the following 10 features computed for the cellnuclei

#### Task Introduction

- ► Task 1: Build a logistic model to classify the images into malignant/benign, and write down your likelihood function, its gradient and Hessian matrix.
- Task 2: Develop a Newton-Raphson algorithm to estimate your model.
- ▶ Task 3: Build a logistic-LASSO model to select features, and implement a path-wise coordinate-wise optimization algorithm to obtain a path of solutions with a sequence of descending  $\lambda$ 's.
- Task 4: Use 5-fold cross-validation to select the best λ. Compare the prediction performance between the 'optimal' model and 'full' model

### Task 1 - Objective

#### **Objective:**

Build a logistic model to classify the images into malignant/benign, and write down your likelihood function, its gradient and Hessian matrix.

Define the "Diagnosis" variable will be coded as 1 for malignant cases and 0 for benign cases.

Given n i.i.d. observations with p predictors, we consider a logistic regression model

$$P(Y_i = 1 \mid \mathbf{x}_i) = \frac{e^{\mathbf{x}_i^\top \beta}}{1 + e^{\mathbf{x}_i^\top \beta}}, \ i = 1, \dots, n$$
 (1)

where  $\boldsymbol{\beta} = (\beta_0, \beta_1, \dots, \beta_p)^{\top} \in \mathbb{R}^{p+1}$  is the parameter vector,  $\mathbf{x}_i = (1, X_{i1}, \dots, X_{ip})^{\top}$  is the vector of predictors in the *i*-th observation, and  $Y_i \in \{0,1\}$  is the binary response in the *i*-th observation.

Let  $\mathbf{y} = (Y_1, Y_2, \dots, Y_n)^{\top}$  denote the response vector, and  $\mathbf{X} = (\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n)^{\top} \in \mathbb{R}^{n \times (p+1)}$  denote the design matrix.

The observed likelihood of  $\{(Y_1, \mathbf{x}_1), (Y_2, \mathbf{x}_2), \dots, (Y_n, \mathbf{x}_n)\}$  is

$$L(\beta; \mathbf{y}, \mathbf{X}) = \prod_{i=1}^{n} \left| \left( \frac{e^{\mathbf{x}_{i}^{\top} \beta}}{1 + e^{\mathbf{x}_{i}^{\top} \beta}} \right)^{Y_{i}} \left( \frac{1}{1 + e^{\mathbf{x}_{i}^{\top} \beta}} \right)^{1 - Y_{i}} \right|$$

Maximizing the likelihood is equivalent to maximizing the log-likelihood function:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[ Y_{i} \mathbf{x}_{i}^{\top} \beta - \log \left( 1 + e^{\mathbf{x}_{i}^{\top} \beta} \right) \right].$$
 (2)

The estimates of model parameters are

$$\widehat{m{eta}} = \arg\max_{m{eta}} \ f(m{eta}; \mathbf{y}, \mathbf{X}),$$

and the optimization problem is

$$\max_{\beta} f(\beta; \mathbf{y}, \mathbf{X}). \tag{3}$$

Denote  $p_i = P(Y_i = 1 \mid \mathbf{x}_i)$  as given in (1) and  $\mathbf{p} = (p_1, p_2, \dots, p_n)^{\top}$ . The gradient of f is:

$$\nabla f(\beta; \mathbf{y}, \mathbf{X}) = \mathbf{X}^{\top} (\mathbf{y} - \mathbf{p})$$

$$= \sum_{i=1}^{n} (Y_i - p_i) \mathbf{x}_i$$

$$= \begin{pmatrix} \sum_{i=1}^{n} (Y_i - p_i) \\ \sum_{i=1}^{n} (Y_i - p_i) X_{i1} \\ \vdots \\ \sum_{i=1}^{n} (Y_i - p_i) X_{ip} \end{pmatrix}$$

Denote  $w_i = p_i(1 - p_i) \in (0, 1)$  and  $\mathbf{W} = \operatorname{diag}(w_1, \dots, w_n)$ . The Hessian matrix of f is given by

$$\nabla^{2} f(\beta; \mathbf{y}, \mathbf{X}) = -\mathbf{X}^{\top} \mathbf{W} \mathbf{X} 
= -\sum_{i=1}^{n} w_{i} \mathbf{x}_{i} \mathbf{x}_{i}^{\top} 
= -\begin{pmatrix}
\sum_{i=1}^{n} w_{i} & \sum_{i=1}^{n} w_{i} X_{i1} & \cdots & \sum_{i=1}^{n} w_{i} X_{i1} \\
\sum_{i=1}^{n} w_{i} X_{i1} & \sum_{i=1}^{n} w_{i} X_{i1}^{2} & \cdots & \sum_{i=1}^{n} w_{i} X_{i1} X_{ip} \\
\vdots & \vdots & \ddots & \vdots \\
\sum_{i=1}^{n} w_{i} X_{ip} & \sum_{i=1}^{n} w_{i} X_{in} X_{i1} & \cdots & \sum_{i=1}^{n} w_{i} X_{ip}^{2}
\end{pmatrix}$$

Next, we show that the Hessian matrix  $\nabla^2 f(\beta; \mathbf{y}, \mathbf{X})$  is a negative-definite matrix if  $\mathbf{X}$  has full rank.

**Proof.** For any (p+1)-dimensional nonzero vector  $\alpha$ , given that **X** has full rank,  $\mathbf{X}\alpha$  is also a nonzero vector. Since **W** is positive-definite, we have

$$egin{aligned} oldsymbol{lpha}^{ op} 
abla^2 f(eta; \mathbf{y}, \mathbf{X}) &lpha = oldsymbol{lpha}^{ op} (-\mathbf{X}^{ op} \mathbf{W} \mathbf{X}) oldsymbol{lpha} \\ &= -(\mathbf{X} oldsymbol{lpha})^{ op} \mathbf{W} (\mathbf{X} oldsymbol{lpha}) \\ &< 0. \end{aligned}$$

Thus,  $\nabla^2 f(\beta; \mathbf{y}, \mathbf{X})$  is negative-definite.

Hence, the optimization problem (3) is a well-defined problem.

### Task 1 - Logistic model R code

```
cancer <- read.csv("breast-cancer.csv") %>%
ianitor::clean names()%>%
select(-1,-33) %>%
mutate(diagnosis = recode(diagnosis, "M" = 1, "B" = 0))
cor()
ggcorrplot(corr, type = "upper", tl.cex = 8)
set.seed(1)
trainRows <- createDataPartition(v = cancer$diagnosis, p = 0.8, list = FALSE)
train <- cancer[trainRows, ]</pre>
test <- cancer[-trainRows, ]
glm.fit <- glm(diagnosis ~ ..
               data = train,
               subset = trainRows,
               family = binomial(link = "logit"))
summary(glm.fit)
pred <- predict(glm.fit, newdata = test, type = "response")</pre>
v test <- factor(test$diagnosis)
auc full <- auc(v test, pred)
auc_full
```

- ► Final AUC of full model reaches 0.9641.
- ▶ And if we remove correlated variables, the AUC is uplifted to 0.9962.

### Task 2 - Objective

#### **Objective:**

Develop a Newton-Raphson algorithm to estimate your model

▶ Recall The target function *f* given in task 1 is:

$$f(eta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[ Y_i \mathbf{x}_i^{\top} eta - \log \left( 1 + e^{\mathbf{x}_i^{\top} eta} \right) \right].$$

We develop a modified Newton-Raphson algorithm including a step-halving step to maximize the target function.

# Task 2 - Newton-Raphson algorithm design

#### **Algorithm 1** Newton-Raphson algorithm

**Require:**  $f(\beta)$  - target function as given in (4);  $\beta_0$  - starting value

**Ensure:** 
$$\widehat{\boldsymbol{\beta}}$$
 such that  $\widehat{\boldsymbol{\beta}} \approx \arg \max_{\boldsymbol{\beta}} f(\boldsymbol{\beta})$ 

 $i \leftarrow 0$ , where i is the current number of iterations

$$f(\beta_{-1}) \leftarrow -\infty$$
  
**while** convergence criterion is not met **do**

$$i \leftarrow i + 1$$

 $\mathbf{d}_i \leftarrow -[\nabla^2 f(\beta_{i-1})]^{-1} \nabla f(\beta_{i-1})$ , where  $\mathbf{d}_i$  is the direction in

the *i*-th iteration 
$$\lambda_i \leftarrow 1$$
, where  $\lambda_i$  is the multiplier in the *i*-th iteration

$$oldsymbol{eta}_i \leftarrow oldsymbol{eta}_{i-1} + \lambda_i \mathbf{d}_i$$

while 
$$f(\beta_i) \le f(\beta_{i-1})$$
 do  $\lambda_i \leftarrow \lambda_i/2$ 

$$\lambda_i \leftarrow \lambda_i/2$$
 $\boldsymbol{\beta}_i \leftarrow \boldsymbol{\beta}_{i-1} + \lambda_i \mathbf{d}_i$ 

end while

#### end while

$$\widehat{\boldsymbol{\beta}} \leftarrow \boldsymbol{\beta}_i$$

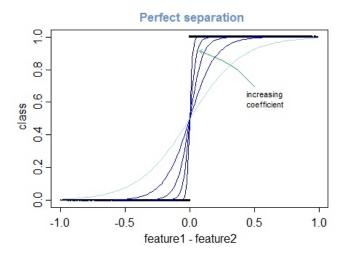
# Task 2 - Newton-Raphson algorithm R code

```
NewtonRaphson <- function(dat, func, start, tol = 1e-8, maxiter = 200) {
  i \leftarrow 0
 cur <- start
 stuff <- func(dat, cur)
 res <- c(0, stuff$f, cur)
 prevf <- -Inf
 X <- cbind(rep(1, nrow(dat)), as.matrix(dat[, -1]))</pre>
 v <- dat[, 1]</pre>
  warned <- 0
 while (abs(stuff$f - prevf) > tol && i < maxiter) {
    i < -i + 1
    prevf <- stuff$f
    prev <- cur
    d <- -solve(stuff$Hess) %*% stuff$grad
    cur <- prev + d
    lambda <- 1
   maxhalv <- 0
    while (func(dat, cur)$f < prevf && maxhalv < 50) {
      maxhalv <- maxhalv + 1
      lambda <- lambda / 2
      cur <- prev + lambda * d
    stuff <- func(dat, cur)
    res <- rbind(res, c(i, stuff$f, cur))
    v_hat \leftarrow ifelse(X %*% cur > 0, 1, 0)
    if (warned == 0 \&\& sum(v - v hat) == 0) {
      warning("Complete separation occurs. Algorithm does not converge.")
     warned <- 1
 colnames(res) <- c("iter", "target_function", "(Intercept)", names(dat)[-1])
 return(res)
```

### Task 2 - Complete separtion

- Sometimes our algorithm does not converge because of the complete separation.
- A complete separation in a logistic regression, sometimes also referred as perfect prediction, occurs whenever there exists some vector of coefficients  $\boldsymbol{\beta}$  such that  $Y_i = 1$  whenever  $\mathbf{x}_i^{\top} \boldsymbol{\beta} > 0$  and  $Y_i = 0$  whenever  $\mathbf{x}_i^{\top} \boldsymbol{\beta} \leq 0$ .
- Complete separation occur when a linear function of predictors can perfectly classify the response.

### Task 2 - Complete separation



### Task 2 - Complete separation

- We have proved that: when there exists a vector of coefficients  $\hat{\beta}$  such that  $Y_i=1$  whenever  $\mathbf{x}_i^{\top}\hat{\boldsymbol{\beta}}>0$  and  $Y_i=0$  whenever  $\mathbf{x}_i^{\top}\hat{\boldsymbol{\beta}}\leq 0$ , there does not exist  $\boldsymbol{\beta}^*\in\mathbb{R}^{(p+1)}$  such that  $\boldsymbol{\beta}^*=\arg\max_{\boldsymbol{\beta}}f(\boldsymbol{\beta})$ , where f is given in (4). Thus our algorithm does not converge. (proof is attached in report appendix)
- ▶ If there is no complete separation, the parameters output by glm function and our algorithm are demonstrated to be the same
- ▶ In practice, if complete separation occurs, we randomly pick the parameters which satisfy complete separation as our full model

# Task 2 - Comparison

### Comparison of using glm function and our algorithm (part of)

| predictor              | ours          | glm           |
|------------------------|---------------|---------------|
| (Intercept)            | 111.7230206   | 90.9690365    |
| radius_mean            | -3646.8235387 | -2560.3938902 |
| texture_mean           | -2.8949199    | 0.8812037     |
| perimeter_mean         | 1257.9481173  | 789.2724398   |
| area_mean              | 2091.5001210  | 1539.8345650  |
| smoothness_mean        | 180.4591375   | 128.8762247   |
| compactness_mean       | -471.3749334  | -346.6691873  |
| concavity_mean         | -0.3063034    | 9.0810082     |
| concave.points_mean    | 287.3555092   | 215.4808031   |
| symmetry_mean          | 10.8459784    | 10.4772590    |
| fractal_dimension_mean | -40.6413497   | -28.6916549   |
| radius_se              | 854.4290933   | 617.5687358   |
| texture_se             | -105.8920782  | -79.9788638   |
| perimeter_se           | -1231.6514585 | -917.3916527  |
| area_se                | 789.0550146   | 628.6222313   |
| smoothness_se          | -83.3449730   | -63.7660367   |
| compactness_se         | 473.1970464   | 344.3180183   |
| concavity_se           | -440.6629325  | -323.8533730  |
| concave.points_se      | 503.3728692   | 374.7610875   |
| symmetry_se            | -144.4580558  | -108.6211792  |

### Task 3 - Objective

#### **Objective:**

Build a logistic-LASSO model to select features by implementing a path-wise coordinate-wise optimization algorithm

▶ Recall Log-likelihood *f* in task 1:

$$f(\beta; \mathbf{y}, \mathbf{X}) = \sum_{i=1}^{n} \left[ Y_{i} \mathbf{x}_{i}^{\top} \beta - \log \left( 1 + e^{\mathbf{x}_{i}^{\top} \beta} \right) \right].$$
 (4)

ightharpoonup LASSO estimates the logistic model parameters eta by optimizing a penalized loss function:

$$\min_{\beta} \ -\frac{1}{n} f(\beta) + \lambda \sum_{k=1}^{p} |\beta_k|. \tag{5}$$

where  $\lambda \geq 0$  is the tuning parameter. Note that the intercept is not penalized and all predictors are standardized.

## Task 3 - Algorithm of Logistic-LASSO Model

#### **Algorithm Structure:**

- ▶ OUTER LOOP: Decrement  $\lambda$ .
- ▶ MIDDLE LOOP: Update  $\tilde{w}_i$ ,  $\tilde{p}_i$ , and thus the quadratic approximation  $\ell$  using the current parameters  $\tilde{\beta}$ .
- ► INNER LOOP: Run the coordinate descent algorithm on the penalized weighted-least-squares problem.

#### Task 3 - OUTER LOOP

#### OUTER LOOP:

Compute the solutions of the optimization problem (5) for a decreasing sequence of values for  $\lambda$ :  $\{\lambda_1,\ldots,\lambda_m\}$ , starting at the smallest value  $\lambda_1=\lambda_{max}$ 

$$\lambda_{max} = \frac{1}{n} \max_{j} \left| \langle \mathbf{x}_{.j}, \mathbf{y} \rangle \right|, \tag{6}$$

where  $\mathbf{x}_{\cdot j}$  is the j-th column of the design matrix  $\mathbf{X}$ , for  $j=1,\ldots,p$ .

For tuning parameter value  $\lambda_{k+1}$ , we initialize coordinate descent algorithm at the computed solution for  $\lambda_k$  (warm start).

Task 4

#### **Discussions**

- ▶ There is much freedom when designing the simulations.
- ▶ In our algorithm, we have 5 parameters. n, p, ratio, c, corr
- ▶ More parameters can be adjusted.

### Limitations and Future Work

- ► Limitation: We reproduced high-dimensional scenarios, but we still don't know the solution.
- ► Future Work: We may adjust other parameters to investigate further.

#### Reference

 Li Y, Hong HG, Ahmed SE, Li Y. Weak signals in high-dimensional regression: Detection, estimation and prediction. Appl Stochastic Models Bus Ind. 2018;1–16. https://doi.org/10.1002/asmb.2340 Q&A

- ► Thanks for listening!
- ► Any questions?