

# Notes on PI model estimation

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## 0th and 1st order expansion

We assume that the equilibrium model solutions form a class of  $n$  dimensional stochastic processes  $\{X_t(\mathbf{q})\}$  indexed by the perturbation parameter  $\mathbf{q}$  satisfying the recursion (law of motion):

$$X_{t+1}(\mathbf{q}) = \psi[X_t(\mathbf{q}), \mathbf{q}W_{t+1}, \mathbf{q}] \quad (11)$$

Although we typically don't have quasi-analytical solution of (11), we can approximate it with first- and second-order small noise expansions.

A first-order expansion of  $X_t(\mathbf{q})$  around  $\mathbf{q} = 0$  takes the form

$$X_t \approx X_t^0 + \mathbf{q}X_t^1$$

The 0th order component (steady state),  $X_t^0$ , is just a constant.

For example, in the permanent income model<sup>1</sup> that is of interest to us:

- The 0th order component of  $\frac{K}{Y}$  is a free initial condition that we can freely<sup>2</sup> impose, and we would like to choose the  $\frac{K^0}{Y^0}$  that can match the investment/consumption data<sup>3</sup>.
- The 0th order component of  $\log\left(\frac{Y_{t+1}}{Y_t}\right)$  is  $\mathbf{g}$ , which is a number that we can compute by matching it with log consumption growth data<sup>4</sup>.
- The 0th order component of  $\frac{C}{Y}$  is:

$$\log\left(\frac{C^0}{Y^0}\right) = \log\left\{[1 + \mathbf{a} - \exp(\mathbf{g})]\left(\frac{K^0}{Y^0}\right) + 1\right\}$$

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<sup>1</sup>For a full description of the model, see another file PI model notes.

<sup>2</sup>This means that each different choice of  $\frac{K^0}{Y^0}$  can result in unique model dynamics

<sup>3</sup>How? See the explanation below on  $\frac{I_t^0}{C_t^0}$

<sup>4</sup>Why log consumption growth? See the explanation below on  $\log\left(\frac{C_{t+1}^0}{C_t^0}\right)$ .

- The 0th order component of  $\log\left(\frac{C_{t+1}}{C_t}\right)$  is (recall balanced growth of  $\log\left(\frac{C}{Y}\right)$ ):

$$\log\left(\frac{C_{t+1}^0}{C_t^0}\right) = \log\left(\frac{C_{t+1}^0}{Y_{t+1}^0}\right) + \log\left(\frac{Y_{t+1}^0}{Y_t^0}\right) - \log\left(\frac{C_t^0}{Y_t^0}\right) = \log\left(\frac{Y_{t+1}^0}{Y_t^0}\right) = \mathbf{g}$$

- The 0th order component of  $\frac{I_t}{C_t} \doteq \frac{K_{t+1}-K_t}{C_t}$  is (see PI model notes):

$$\begin{aligned} \frac{I_t^0}{C_t^0} &= \frac{K_{t+1}^0}{Y_{t+1}^0} \exp \log\left(\frac{Y_{t+1}^0}{Y_t^0}\right) \exp \log\left(\frac{Y_t^0}{C_t^0}\right) - \frac{K_t^0}{Y_t^0} \exp \log\left(\frac{Y_t^0}{C_t^0}\right) \\ &= \frac{K^0}{Y^0} [\exp(\mathbf{g}) - 1] \exp \log\left(\frac{Y^0}{C^0}\right) \\ &= \frac{\frac{K^0}{Y^0} [\exp(\mathbf{g}) - 1]}{[1 + \mathbf{a} - \exp(\mathbf{g})] \left(\frac{K^0}{Y^0}\right) + 1} \end{aligned}$$

The 1st order component,  $X_t^1$ , has a state space representation:

$$\begin{aligned} \log \frac{C_{t+1}^1}{C_t^1} &= D_1 \left[ \log\left(\frac{H_t^1}{Y_t^1}\right) \quad \frac{K_t^1}{Y_t^1} \quad Z_{1,t}^1 \quad Z_{2,t}^1 \quad Z_{2,t-1}^1 \right]' + F_1 W_{t+1} + H_1 \\ \frac{I_{t+1}^1}{C_{t+1}^1} \doteq \frac{K_{t+2}^1 - K_{t+1}^1}{C_{t+1}^1} &= D_2 \left[ \log\left(\frac{H_t^1}{Y_t^1}\right) \quad \frac{K_t^1}{Y_t^1} \quad Z_{1,t}^1 \quad Z_{2,t}^1 \quad Z_{2,t-1}^1 \right]' + F_2 W_{t+1} + H_2 \\ \left[ \log\left(\frac{H_{t+1}^1}{Y_{t+1}^1}\right) \quad \frac{K_{t+1}^1}{Y_{t+1}^1} \quad Z_{1,t+1}^1 \quad Z_{2,t+1}^1 \quad Z_{2,t}^1 \right]' &= A \left[ \log\left(\frac{H_t^1}{Y_t^1}\right) \quad \frac{K_t^1}{Y_t^1} \quad Z_{1,t}^1 \quad Z_{2,t}^1 \quad Z_{2,t-1}^1 \right]' + B W_{t+1} \end{aligned}$$

Combined together the 0th + 1st order state space representation makes the **Kalman Filter** applicable for estimation, i.e. we are actually estimating the 0th + 1st order linear state space approximation of the original (non-linear) permanent income model.

## Estimation problem

We want to estimate 5 parameters that govern the exogenous hidden states  $Z_1$  and  $Z_2$ :

$$\begin{bmatrix} Z_{1,t+1} \\ Z_{2,t+1} \\ Z_{2,t} \end{bmatrix} = \begin{bmatrix} \phi_1 & 0 & 0 \\ 0 & \phi_{21} & -\phi_{22} \\ 0 & 1 & 0 \end{bmatrix} \begin{bmatrix} Z_{1,t} \\ Z_{2,t} \\ Z_{2,t-1} \end{bmatrix} + \begin{bmatrix} \sigma_1 & 0 \\ 0 & \sigma_2 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} W_{1,t+1} \\ W_{2,t+1} \end{bmatrix}$$

We also have parameters  $\rho, \chi, \alpha, \epsilon$  that govern preference, and we would potentially like to include them into estimation:

$$V_t = \left[ (1 - \beta) U_t^{1-\rho} + \beta R_t^{1-\rho} \right]^{\frac{1}{1-\rho}}$$

$$R_t = \mathbb{E} \left[ V_{t+1}^{1-\gamma} \mid \mathfrak{F}_t \right]^{\frac{1}{1-\gamma}}$$

$$U_t = \left[ (1 - \alpha) C_t^{1-\epsilon} + \alpha H_t^{1-\epsilon} \right]^{\frac{1}{1-\epsilon}}$$

$$H_{t+1} = \chi H_t + (1 - \chi) C_t$$

The log income growth process would be

$$\log\left(\frac{Y_{t+1}}{Y_t}\right) = .01(Z_{1,t+1} + Z_{2,t+1} - Z_{2,t}) + \mathbf{g} = .01 \left( \begin{bmatrix} \phi_1 \\ \phi_{21} - 1 \\ -\phi_{22} \end{bmatrix} \cdot \begin{bmatrix} Z_{1,t} \\ Z_{2,t} \\ Z_{2,t-1} \end{bmatrix} + \begin{bmatrix} \sigma_1 & \sigma_2 \end{bmatrix} \begin{bmatrix} W_{1,t+1} \\ W_{2,t+1} \end{bmatrix} \right) + \mathbf{g}$$

Here are some “old numbers” of the exogenous processes - these numbers are good for illustration purposes, but can potentially be far away from what is implied by the data, and that’s the reason why we’re doing the estimation.

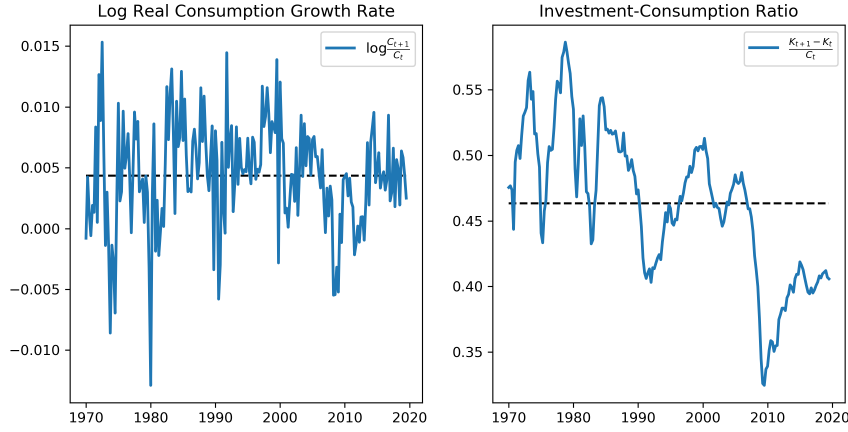
$$\begin{bmatrix} Z_{1,t+1} \\ Z_{2,t+1} \\ Z_{2,t} \end{bmatrix} = \begin{bmatrix} .704 & 0 & 0 \\ 0 & 1 & -.154 \\ 0 & 1 & 0 \end{bmatrix} \begin{bmatrix} Z_{1,t} \\ Z_{2,t} \\ Z_{2,t-1} \end{bmatrix} + \begin{bmatrix} .144 & 0 \\ 0 & .206 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} W_{1,t+1} \\ W_{2,t+1} \end{bmatrix}$$

## Some observations in numerical experiments

- If we allow `scipy.optimize.minimize` to freely search for the preference parameters, the outcome will usually be ugly, but things are usually not bad when we only search for the parameters in the exogenous state equations. This is potentially because the preference parameters governs the dynamics in a very non-linear manner. Therefore, we decide to only let `scipy.optimize.minimize` search over the exogenous parameters’ space, and for the preference parameters, we do a grid search instead.
- We notice that  $\chi$  is not moving around very much, usually very close to 0.6. Small  $\alpha$  and small (lower than 1)  $\rho$  tend to have higher likelihood.  $\epsilon$  has a big magnitude (a few hundreds).

## Data, algorithm, and code

As stated previously, we use  $\log\left(\frac{C_{t+1}}{C_t}\right)$  and  $\frac{I_{t+1}}{C_{t+1}}$  as observables in writing the state space. The data (obtained from FRED) will look like:



The algorithm works like this:

1. Back out the relevant  $\mathbf{g}$  and  $\frac{K^0}{Y^0}$  from data<sup>5</sup>;
2. Divide the preference parameter space using grids. On each grid point (i.e. for each combination of fixed preference parameters), let `scipy.optimize.minimize` search for the combination of exogenous parameters that can give the maximized log likelihood, where log likelihood is obtained using Kalman Filter.
3. Find the point on the preference parameter space that gives the maximized log likelihood, and also plot the region on the preference parameter space that gives log likelihood close<sup>6</sup> to the maximized value.

The relevant code and results are contained in `estimation_preference_2d.ipynb`, as well as other relevant `.py` scripts.

Very roughly speaking, we have the highest log likelihood when the preference parameters are given by  $\rho = .67$ ,  $\alpha = .1$ ,  $\epsilon = 218$ ,  $\chi = .61$ , and the exogenous parameters that maximize the log likelihood in this case are:

$$\begin{bmatrix} Z_{1,t+1} \\ Z_{2,t+1} \\ Z_{2,t} \end{bmatrix} = \begin{bmatrix} .796 & 0 & 0 \\ 0 & 1.029 & -.074 \\ 0 & 1 & 0 \end{bmatrix} \begin{bmatrix} Z_{1,t} \\ Z_{2,t} \\ Z_{2,t-1} \end{bmatrix} + \begin{bmatrix} .770 & 0 \\ 0 & 1.298 \\ 0 & 0 \end{bmatrix} \begin{bmatrix} W_{1,t+1} \\ W_{2,t+1} \end{bmatrix}$$

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<sup>5</sup>For some reason that I don't remember exactly, we have fixed  $\mathbf{a}$  to be 0.663% rather than estimating it.

<sup>6</sup>Usually, we look at the region with log likelihood decrease of 3.5 or less.

