

Moments

4760

$$m_n = E[X^n] = \int_{-\infty}^{\infty} x^n f_X(x) dx$$

n 'th moment of r.v. X .

$$\{m_0, m_1, \dots\} = \begin{matrix} f_X(x) \\ \text{or} \\ F_X(x) \end{matrix}$$

$$N_n = E[\overbrace{(X - \bar{X})^n}^{\text{zero mean}}] \longrightarrow E[\overbrace{X - \bar{X}}^{\text{zero mean}}] = E[X] - E[\bar{X}] =$$

$$\bar{X} = E[X]$$



n 'th centralized moment.

$$E[c] = \int_{-\infty}^{\infty} c \cdot f_X(x) dx = c \cdot \int_{-\infty}^{\infty} f_X(x) dx = c$$

$$\bar{X} - \bar{X} = 0 \Rightarrow E[Y] = 0$$

$$N_n = \int_{-\infty}^{\infty} (x - \bar{X})^n f_X(x) dx$$

we take interest in $n=1, 2$ for m_n (m_1, m_2)

and $n=2$ for N_n (N_2)

$m_0 = E[\underbrace{X^0}_1] = 1$ does not give any information about r.v. X .

$m_1 = E[X^1] = E[X]$ is simply the mean value that we already know.

$m_2 = E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx$ referred to as second moment.

$$\left. \begin{aligned} \mu_0 &= E[(X - \bar{X})^0] = 1 \\ \mu_1 &= E[(X - \bar{X})^1] = 0 \end{aligned} \right\} \text{ does not give any info about } X.$$

$$\mu_2 = E[(X - \bar{X})^2] = E(X^2 - 2X\bar{X} + \bar{X}^2) = E[X^2] - 2\bar{X} \underbrace{E[X]}_{\bar{X}} + \underbrace{E[\bar{X}^2]}_{\bar{X}^2}$$

$$= E[X^2] - \bar{X}^2 = E[X^2] - E^2[X] \quad \text{is called} \quad \text{"the variance of the r.v. } X" = \sigma_X^2$$

$$\sigma_X^2 = \mu_2 = \int_{-\infty}^{\infty} \underbrace{(X - \bar{X})^2}_{\geq 0} \underbrace{f_X(X)}_{\geq 0} dX \geq 0$$

σ_X : standard deviation of r.v. X .
(has the same units as r.v. X)

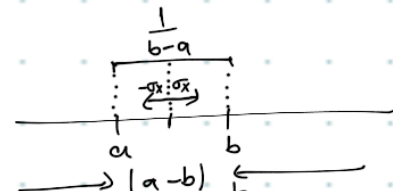
Ex Consider uniform r.v. uniformly distributed between a and b .

$$f_X(x) = \begin{cases} \frac{1}{b-a}, & a \leq x \leq b \\ 0, & \text{otherwise} \end{cases}$$

$$\sigma_X^2 = \underbrace{E[X^2]}_{\mu_2} - \underbrace{E^2[X]}_{\mu_1^2} \Rightarrow \sigma_X^2 = \mu_2 - \mu_1^2$$

$$\mu_1 = \frac{a+b}{2}$$

$$\Rightarrow \sigma_X^2 = \frac{a^2 + ab + b^2}{3} - \left(\frac{a+b}{2}\right)^2 = \frac{a^2 - 2ab + b^2}{12} = \frac{(a-b)^2}{12} \Rightarrow$$

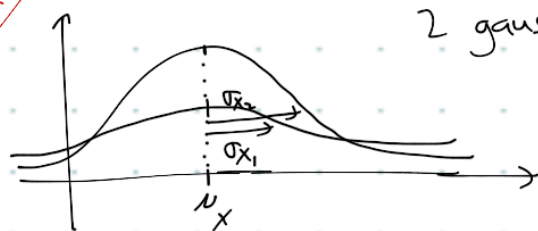


$$\mu_2 = E[X^2] = \int_a^b x^2 \frac{1}{b-a} dx = \frac{b^3 - a^3}{(b-a) \cdot 3}$$

$$= \frac{a^2 + ab + b^2}{3}$$

$$\sigma_X = \frac{|a-b|}{\sqrt{12}}$$

Ex/



2 gaussian r.v.'s

Systematic Determination of Higher Order Moments

- a) Moment generating function (like the Laplace transform of $f_X(x)$)
b) Characteristic function (like Fourier transform of $f_X(x)$)

$$F(s) = \mathcal{L}\{f(t)\} = \int_{-\infty}^{\infty} e^{-st} f(t) dt$$

$s = j\omega$

$$F(j\omega) = \int_{-\infty}^{\infty} e^{-j\omega t} f(t) dt$$

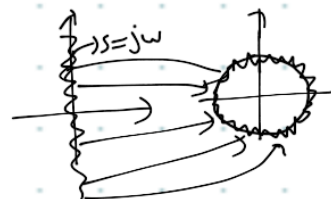
a) $M_X(s) \triangleq \mathbb{E}[e^{sX}] = \int_{-\infty}^{\infty} e^{sx} f_X(x) dx$
 $g(x) = e^{sx}$

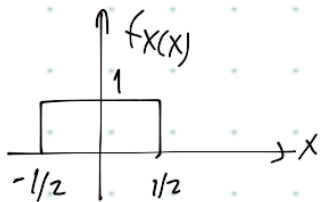
$$\Rightarrow \underline{M_X(-s)} = \underline{\mathcal{L}\{f_X(x)\}}$$

b) $\phi_X(\omega) = M_X(s) \Big|_{s=j\omega} = \mathbb{E}[e^{j\omega X}]$

$$= \int_{-\infty}^{\infty} e^{j\omega x} f_X(x) dx$$

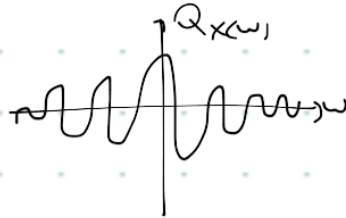
$$\Rightarrow \underline{\phi_X(-\omega)} = \underline{\text{F.T. of } \{f_X(x)\}}$$





$$\Phi_X(\omega) = \int_{-\infty}^{\infty} e^{j\omega x} f_X(x) dx = \int_{-1/2}^{1/2} e^{j\omega x} dx = \frac{e^{j\omega/2} - e^{-j\omega/2}}{j\omega}$$

$$= \frac{(e^{j\omega/2} - e^{-j\omega/2}) \cdot 2j}{2j \cdot \omega} = \frac{\sin \omega/2}{\omega/2}$$



$$\Phi_X(\omega) = \int_{-\infty}^{\infty} e^{j\omega x} f_X(x) dx$$

$$\left. \frac{d\Phi_X(\omega)}{j\omega} \right|_{\omega=0} = \int_{-\infty}^{\infty} x \cdot e^{j\omega x} \cdot f_X(x) dx \Big|_{\omega=0} = \int_{-\infty}^{\infty} x f_X(x) dx = m_1$$

$$\left. \frac{d^2 \Phi_X(\omega)}{j^2 \omega^2} \right|_{\omega=0} = \int_{-\infty}^{\infty} x^2 \cdot e^{j\omega x} \cdot f_X(x) dx \Big|_{\omega=0} = \int_{-\infty}^{\infty} x^2 f_X(x) dx = m_2$$

$$\left. \frac{d^n \Phi_X(\omega)}{j^n \omega^n} \right|_{\omega=0} = \int_{-\infty}^{\infty} x^n \cdot e^{j\omega x} \cdot f_X(x) dx \Big|_{\omega=0} = \int_{-\infty}^{\infty} x^n f_X(x) dx = \underline{\underline{m_n}}$$

Ex/

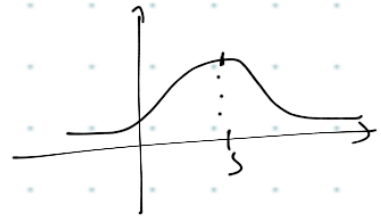
$$X \sim N(0,1) \quad \mu_X = E[X] = \bar{x}$$

$\downarrow \quad \downarrow$
 $\mu_X \quad \sigma_X$

$$M_X(s) = \int_{-\infty}^{\infty} e^{sx} f_X(x) dx = \int_{-\infty}^{\infty} e^{sx} \frac{1}{\sqrt{2\pi}\sigma_X} \exp\left(-\frac{1}{2} \cdot \left(\frac{x-\mu_X}{\sigma_X}\right)^2\right) dx$$

$$= \int_{-\infty}^{\infty} e^{sx} \cdot \frac{1}{\sqrt{2\pi}} \cdot \exp\left(-\frac{1}{2}x^2\right) dx = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \cdot \exp\left(-\frac{1}{2}x^2 + sx\right) dx$$

$$= e^{s^2/2} \underbrace{\int_{-\infty}^{\infty} \frac{e^{-1/2(x-s)^2}}{\sqrt{2\pi}} dx}_{\text{like a r.v. } \sim N(s,1)}$$



$$= e^{s^2/2} \cdot 1$$

$$M_X(s) = \int_{-\infty}^{\infty} e^{sx} f_X(x) dx = \left. \frac{d}{ds} M_X(s) \right|_{s=0} = \int_{-\infty}^{\infty} \underbrace{x \cdot e^{sx}}_{s=0} \bigg|_{s=0} f_X(x) dx$$

$$m_1 = \left. \frac{d}{ds} e^{s^2/2} \right|_{s=0} = e^{s^2/2} \cdot s \bigg|_{s=0} = 0$$

$$m_2 = \left. \frac{d^2 M_X(s)}{ds^2} \right|_{s=0} \quad m_2 = \left[\frac{d}{ds} (e^{s^2/2} \cdot s) \right] \bigg|_{s=0} = (e^{s^2/2} \cdot s^2 + e^{s^2/2} \cdot 1) \bigg|_{s=0}$$

$$= 1 \Rightarrow \sigma_X^2 = \mu_2 = m_2 - m_1^2 = E[X^2] - E^2[X] = \sigma_X^2 = 1 - 0^2 = \underline{1}$$

Mean Value of Poisson r.v.

Ex: $f_X(x) = \sum_{n=0}^{\infty} \underbrace{\frac{e^{-a} \cdot a^n}{n!}}_{\Pr\{X=n\}} \delta(x-n)$

$$p_n = \Pr\{X=n\} = \frac{e^{-a} \cdot a^n}{n!}$$

$$E[X] = \sum_{n=0}^{\infty} \lambda_n p_n = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} \cdot n$$

$$M_X(s) = \int_{-\infty}^{\infty} e^{sx} \underbrace{f_X(x)}_{\sum_{n=0}^{\infty} \frac{e^{-a} a^n}{n!} \delta(x-n)} dx = \int_{-\infty}^{\infty} e^{sx} \sum_{n=0}^{\infty} \frac{e^{-a} a^n}{n!} \cdot \delta(x-n) dx$$

$$= \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} \cdot \int_{-\infty}^{\infty} \underbrace{e^{sx} \delta(x-n)}_{e^{sn} \delta(x-n)} dx = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} e^{sn} \cdot \underbrace{\int_{-\infty}^{\infty} \delta(x-n) dx}_1$$

$$= \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} \cdot e^{sn}$$

$$m_1 = \left. \frac{d}{ds} M_X(s) \right|_{s=0} = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} \underbrace{n e^{sn}}_{s=0} = a$$

$$\Rightarrow M_X(s) = e^{-a} \sum_{n=0}^{\infty} \frac{a^n a^{\log_a e^{sn}}}{n!} \dots \dots \dots$$

$$m_1 = E[X] = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} \cdot n = e^{-a} \cdot \sum_{n=0}^{\infty} \frac{a^n \cdot n}{n!}$$

$$\Rightarrow 1 = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n}{n!} \Rightarrow e^a = \sum_{n=0}^{\infty} \frac{a^n}{n!} \Rightarrow e^a = \sum_{n=0}^{\infty} \frac{n \cdot a^{n-1}}{n!} \stackrel{(a \cdot e^{-a})}{\Rightarrow} a = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot n}{n!}$$

$$m_2 = E[X^2] = \sum_{n=0}^{\infty} \frac{e^{-a} \cdot a^n \cdot n^2}{n!}$$

$$e^a = \sum_{n=0}^{\infty} \frac{n \cdot (n-1) \cdot a^{n-2}}{n!} \xrightarrow{(a^2 \cdot e^{-a})} a^2 = \sum_{n=0}^{\infty} \frac{n \cdot (n-1) \cdot a^n \cdot e^{-a}}{n!}$$

$$= \underbrace{\sum_{n=0}^{\infty} \frac{n^2 a^n e^{-a}}{n!}}_{m_2} - \underbrace{\sum_{n=0}^{\infty} \frac{n \cdot a^n \cdot e^{-a}}{n!}}_{-m_1 = -a} \Rightarrow m_2 - m_1 = a^2$$

$$\boxed{m_2 = a + a^2}$$

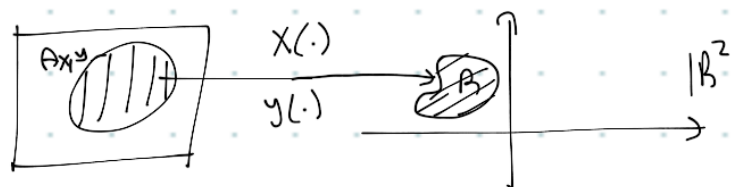
$$\sigma_X^2 = m_2 = E[X^2] - E^2[X] = m_2 - m_1^2 = a^2 + a - (a)^2 = \underline{a}$$

$$\boxed{m_1 = m_2 = a} \text{ for poisson.}$$

$$a = n \cdot p \quad \left(\begin{matrix} n \rightarrow \infty \\ p \rightarrow 0 \end{matrix} \right)$$

Multiple Random Variables

Joint event $A_{X,Y} = \{ \omega : (X(\omega), Y(\omega)) \in B \}$ ↑ outcomes of ω



Focus on special event:

$$\text{! } A_X = \{ \omega : X(\omega) \leq x \}$$

$$A_{X,Y} = \{ \omega : X(\omega) \leq x, Y(\omega) \leq y \}$$

$$Pr\{ \omega : X(\omega) \leq x, Y(\omega) \leq y \} = F_{X,Y}(x,y) \rightarrow \text{joint distribution function.}$$

$$= Pr\{ X \leq x, Y \leq y \} \quad (\text{a region})$$

Properties

1.) $F_{X,Y}(-\infty, y) = F_{X,Y}(x, -\infty) = \emptyset$

$$F_{X,Y}(-\infty, y) = \Pr\{X \leq -\infty, Y \leq y\}$$

$$\Pr\{\omega: X(\omega) \leq -\infty, Y(\omega) \leq y\}$$

$$\subseteq \{\omega: X(\omega) \leq -\infty\} \rightarrow \Pr\{X \leq -\infty\} = 0$$

$$\Pr\{X \leq -\infty\} \geq \Pr\{X \leq -\infty, Y \leq y\}$$

$0 \rightarrow$ has to be 0.

2) More like a definition.

$$F_{X,Y}(-\infty, y) = \Pr\{\omega: X(\omega) \leq -\infty, Y(\omega) \leq y\} = \Pr\{Y \leq y\} = F_Y(y)$$

$$F_{X,Y}(x, \infty) = F_X(x)$$

\hookrightarrow marginal dist. func. of X .

\hookrightarrow marginal dist. function of Y .

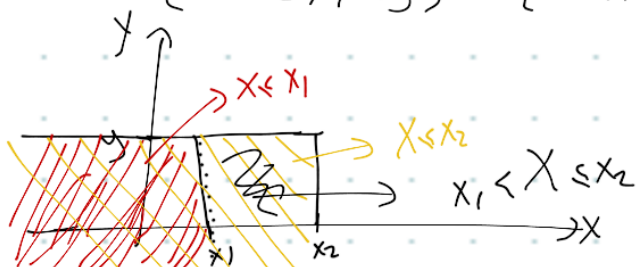
3) $F_{X,Y}(x, y)$ is non-decreasing function in either x or y when the other variable is held fixed.

$$x_1 < x_2$$

$$F_{X,Y}(x_1, y) \leq F_{X,Y}(x_2, y)$$

$$\{\omega: X(\omega) \leq x_2, Y(\omega) \leq y\} = \{\omega: X(\omega) \leq x_1, Y(\omega) \leq y\} \cup \{\omega: x_1 < X(\omega) \leq x_2, Y(\omega) \leq y\}$$

$$\{X \leq x_2, Y \leq y\} = \{X \leq x_1, Y \leq y\} \cup \{x_1 < X \leq x_2, Y \leq y\}$$

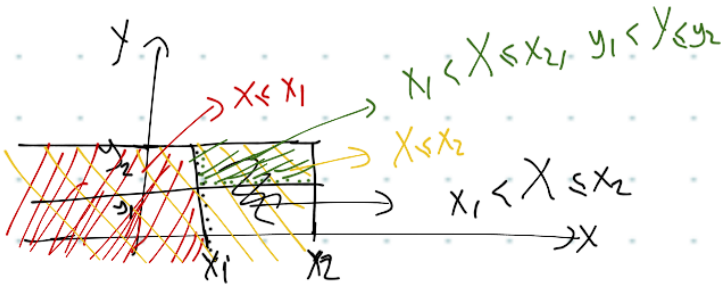


$$\Pr\{X \leq x_2, Y \leq y\} = \Pr\{X \leq x_1, Y \leq y\} + \Pr\{x_1 < X \leq x_2, Y \leq y\}$$

$$F_{X,Y}(x_2, y) - F_{X,Y}(x_1, y) = \Pr\{x_1 < X \leq x_2, Y \leq y\} \rightarrow \geq 0$$

$$Pr\{x_1 < X \leq x_2, Y \leq y_1\} = F_{X,Y}(x_2, y_1) - F_{X,Y}(x_1, y_1)$$

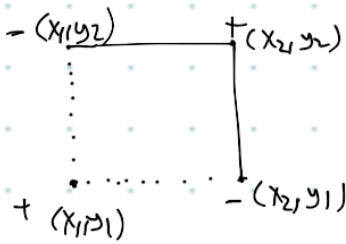
$$Pr\{x_1 < X \leq x_2, Y \leq y_2\} = F_{X,Y}(x_2, y_2) - F_{X,Y}(x_1, y_2)$$



$$\{x_1 < X \leq x_2, Y \leq y_2\} = \{x_1 < X \leq x_2, Y \leq y_1\} \cup \{x_1 < X \leq x_2, y_1 < Y \leq y_2\}$$

$$Pr\{x_1 < X \leq x_2, Y \leq y_2\} = Pr\{x_1 < X \leq x_2, Y \leq y_1\} + Pr\{x_1 < X \leq x_2, y_1 < Y \leq y_2\}$$

$$Pr\{x_1 < X \leq x_2, y_1 < Y \leq y_2\} = F_{X,Y}(x_2, y_2) - F_{X,Y}(x_1, y_2) - F_{X,Y}(x_2, y_1) + F_{X,Y}(x_1, y_1)$$



$$Pr\{(X(t), Y(t)) \in A\}$$

$$f_X(x) = \frac{d}{dx} F_X(x)$$

$$\longrightarrow f_{X,Y}(x,y) = \frac{d^2}{dx dy} F_{X,Y}(x,y)$$

$$F_{X,Y}(x,y) = \int_{-\infty}^y \int_{-\infty}^x f_{X,Y}(a,b) da db$$

$g(B)$