

Lecture 11

Principal Component Analysis

As the no. of features or dimensions in a dataset ↑ so; the amount of data required to obtain a statistically significant result ↑ so; as a result, this, is increased

a statistically significant reason. This results in issues like overfitting, increased computation time, and reduced accuracy.

As the no. of dimensions ↑_{se}, the no. of possible combinations of features ↑_{se} exponentially. Making it computationally difficult to obtain a representative sample of the data and it becomes expensive to perform tasks such as clustering or classification.

In some cases, one need more data to achieve the same level of accuracy as lower-dimensional data. This is known as "Curse of dimensionality".

This is known as Principal Component Analysis (PCA). PCA is a statistical procedure that uses orthogonal transformation to convert a set of correlated variables to a set of uncorrelated variables. Widely used for dimensionality reduction, lossy data compression, feature extraction, and data visualization. It is also known as Karhunen-Loeve transformation.

PCA is a unsupervised learning algorithm that is used to examine interrelations among a set of variables.

It is also known as a general factor analysis where regression determines a line of best fit

Let say a student want admission to a grad school. Look at survey

	Academician	Transport	Safety	Placement
G1	16	12	18	14
G2	12	14	17	20
G3	10	14	11	10
G4	12	14	16	14
G5	15	14	17	19
G6	16	13	17	15
G7	10	14	16	18

more weightage
difficult or soft

$\downarrow V_1$

$\downarrow V_2$

$$G_1 = \{16, 12, 18, 14\} \in \mathbb{R}^4 \rightarrow$$

Wants to reduce to 3 feature y_1, y_2, y_3

$$\mathbb{R}^4 \rightarrow \mathbb{R}^3$$

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \end{pmatrix} = \begin{pmatrix} a_{11} & a_{12} & a_{13} & a_{14} \\ a_{21} & a_{22} & a_{23} & a_{24} \\ a_{31} & a_{32} & a_{33} & a_{34} \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \end{pmatrix}$$

$$y_1 = a_{11}x_1 + a_{12}x_2 + a_{13}x_3 + a_{14}x_4$$

and so on

" linear combinations

We just need to find which convert ones

4D into 3D. in such a way that it
preserves the max. information

PCA does this kind of work.

Easy way is that column with less variation is
not that much useful for our decision

$$\text{Mean: } \mu = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\text{Standard deviation: } \sigma = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (x_i - \mu)^2}$$

Covariance

It is a measure of how two variables change
together

$$\sum_{xy} = \frac{1}{n-1} \sum_{i=1}^n (x_i - \mu_x)(y_i - \mu_y)^T$$

Tells about the relationship b/w multiple variables

in a data set.

ence Σ

Tell us about which direction it is maximum.

f Diagonal element \rightarrow represents variance of each variable.

Variance tell us how spread out the values of a single variable are from the mean.

off diagonal \rightarrow covariance b/w pairs of variables.

for a k dimensional dataset $\{X_1, X_2, \dots, X_k\}$ the covariance Σ is defined as

$$\Sigma = \begin{bmatrix} \Sigma_{x_1^2} & \Sigma_{x_1 x_2} & \Sigma_{x_1 x_3} & \dots & \Sigma_{x_1 x_k} \\ \Sigma_{x_2 x_1} & \Sigma_{x_2^2} & \Sigma_{x_2 x_3} & \dots & \Sigma_{x_2 x_k} \\ \vdots & & & & \vdots \\ \Sigma_{x_k x_1} & \Sigma_{x_k x_2} & \dots & \dots & \Sigma_{x_k^2} \end{bmatrix}$$

It is a $k \times k$ Σ symmetric Σ dataset

$$\Sigma = \frac{1}{n-1} \sum (x_i - \mu_x)(y_i - \mu_y)^T$$

	X_1	X_2	X_3	\dots	X_k
1	-	-	-	-	-
2	-	-	-	-	-
3	-	-	-	-	-
n	-	-	-	-	-

The principal components are the eigenvectors of the covariance Σ of the data

if dataset of 5 variables

$$\Sigma_{5 \times 5} \text{ matrix}$$

The first principal component is the eigenvector corresponding to the largest eigenvalue of the covariance matrix.

Covariance matrix.

- If one wants to transform a m -dimensional dataset to a d -dimensional dataset, then one will simply select the first d principal components.

This somewhat reminds us of the SVD.

- The objective of PCA is to perform dimensionality reduction while keeping as much of the information as possible in the high dimensional space.

If we are projecting from M to D dimensions, PCA will define D vectors, Φ_D , each of which is N -dimensional.

The d^{th} element of projection x_{nd} (where $x_n = [x_{n1}, \dots, x_{nD}]^T$) is computed as

$$x_{nd} = \Phi_d^T y_n$$

learning task is therefore to choose how many dimensions one wants to project into D & then pick a projection vector ϕ_D , for each.

PCA uses variance in projected space as the criteria to choose ϕ_D .

$\phi_1 \rightarrow$ be projection that makes variance in x_{n1} as high as possible

Second projected dimension is also chosen to max. Variance but ϕ_2 must be orthogonal to ϕ_1 .

$$\phi_1^T \phi_2 = 0$$

Third component, ϕ_3 again max. variance & orthogonal

$$\therefore \phi_i^T \phi_j = 0 \quad \forall i \neq j$$

Further PCA $\rightarrow \phi_i$ must have a length of 1

$$\phi_i^T \phi_i = 1$$

A PCA solver is order to find projection ϕ_1, \dots, ϕ_d and it is derived in no. of ways

$$\bar{y} = \frac{1}{N} \sum_{n=1}^N y_n = 0$$

→ force mean to zero.

This is forced by subtracting mean \bar{y} from each y

Finding projection into $D=1$ dim.

We are only interested in finding one ϕ vector

$$x_n = \phi^T y_n$$

Variance $\sigma_x^2 = \frac{1}{N} \sum_{n=1}^N (x_n - \bar{x})^2$ - ①

$$\begin{aligned}\bar{x} &= \frac{1}{N} \sum_{n=1}^N \phi^T y_n \\ &= \phi^T \left(\frac{1}{N} \sum_{n=1}^N y_n \right) = \phi^T \bar{y} = 0\end{aligned}$$

= n ① becomes

$$\sigma_x^2 = \frac{1}{N} \sum_{n=1}^N x_n^2$$

Now substituting $x_n = \phi^T y_n$ gives

$$\begin{aligned}\sigma_x^2 &= \frac{1}{N} \sum_{n=1}^N (\phi^T y_n)^2 \\ &= \frac{1}{N} \sum_{n=1}^N (\phi^T y_n)(y_n^T \phi)\end{aligned}$$

$$\sigma_x^2 = \mathbb{E}[\phi^T \left(\frac{1}{N} \sum_{n=1}^N \phi y_n y_n^T \right) \phi]$$

$$= \phi^T C \phi$$

C is sample covariance matrix.

$$C = \frac{1}{N} \sum_{n=1}^N (y_n - \bar{y})(y_n - \bar{y})^T$$

but here $\bar{y} = 0$ so.

$$C = \frac{1}{N} \sum_{n=1}^N y_n y_n^T$$

As our aim is to find the value ϕ that maximizes

σ^2 and σ^2 also maximizes $\phi^T C \phi$.

One can say that let keep ϕ and max. $\phi^T C \phi$.
But remember we have constrained ϕ to have length of 1. $\phi^T \phi = 1$

One can optimize it through the use of Lagrange

multiplier for maximizing function

$$L = \phi^T C \phi - \lambda (\phi^T \phi - 1)$$

condition / constraint
of $\phi^T \phi = 1$
 $\phi^T \phi^{-1} = 0$

Taking partial derivatives

$$\frac{\partial L}{\partial \phi} = 2 C \phi - 2 \lambda' \phi = 0$$

$$C\phi = \lambda \phi$$

$$\boxed{\lambda = \lambda'}$$

$$C\phi = \lambda \phi$$

This $= n$ is that of eigenvector

Eigenvalue $= n$,

ϕ is eigenvector of the covariance n^T
& λ is eigenvalue.

$$\lambda v = A v$$

Eigenvalue $= n$

This suggest us that projection ϕ that maximizes Variance is one of the eigenvectors of the Covariance $n^T C$.

However there will be M of them, how do we know which one corresponds to the highest variance?

$$\sigma^2 = \phi^T C \phi$$

$$\phi^T \phi = 1 \text{ so.}$$

$$\sigma^2 \phi^T \phi = \phi^T C \phi$$

Removing ϕ^T from both sides.

~~$$\sigma^2 \phi = C \phi$$~~

$$\sigma^2 \phi = C \phi$$

tell us that given an eigenvalue / eigenvector pair (λ, ϕ) , λ corresponds to variance of data in the projected space defined by ϕ .

If we find M eigenvector / eigenvalue pairs of the covariance matrix C , the pair with the highest eigenvalue corresponds to the projection with maximal variance ϕ_1 . The second highest eigenvalue corresponds to ϕ_2 + so on.

Eigenvectors of the covariance matrix represent the direction (principal component) that capture the most variance in data.

Corresponding eigenvalue \rightarrow amount of variance captured by each principal component.

Once we have eigenvectors (P.C.); one can transform original data X by projecting it onto new basis formed by eigenvectors

$$X_{\text{new}} = X V, \quad V \text{ is mt of eigenvectors}$$

Another way X be m -dim vector such that

$$X = \sum_{i=1}^m y_i \phi_i$$

where $(\phi_1, \phi_2, \dots, \phi_m)$ form an orthonormal basis of m -dimensional space & the coordinates ~~of~~ y_i are given as

$$y_i = \langle X, \phi_i \rangle = X^T \phi_i \quad \forall i = 1, 2, \dots, m$$

Want to represent X with few basis vectors (let say d where $d < m$)

We can attempt to do so by replacing y_{d+1}, \dots, y_m with some constant basis

$$\hat{X}(d) = \sum_{i=1}^d y_i \phi_i + \sum_{i=d+1}^m b_i \phi_i$$

Representation error is

$$\Delta X(d) = X - \hat{X}(d) = \sum_{i=d+1}^m (y_i - b_i) \phi_i$$

$$\mathbb{E}(|\Delta X|^2) = \mathbb{E}\left[\left(\sum_{i=d+1}^m (y_i - b_i) \phi_i\right)^2\right]$$

$$= \mathbb{E}\left[\sum_{i=d+1}^m (y_i - b_i) \phi_i \cdot \sum_{j=d+1}^m (y_j - b_j) \phi_j\right]$$

$$= \mathbb{E}\left[\sum_{i=d+1}^m \sum_{j=d+1}^m (y_i - b_i)(y_j - b_j) \phi_i^T \phi_j\right]$$

$$\text{as } \phi_i^T \phi_j = 0 \text{ for } i \neq j$$

$$= 1 \text{ for } i=j$$

$$\mathbb{E}[(\Delta x)^2] = \sum_{i=d+1}^m \mathbb{E}[y_i - b_i]^2$$

on add $\frac{\partial \mathbb{E}[(\Delta x)^2]}{\partial b_i} = 0$
and get $b_i = \mathbb{E}[y_i]$

$$\mathbb{E}[(\Delta x)^2] = \sum_{i=d+1}^m \mathbb{E}[y_i - \mathbb{E}[y_i]]^2$$

$$\Rightarrow y_i = x^\top \phi_i$$

$$= \sum_{i=d+1}^m \mathbb{E}[(x^\top \phi_i - \mathbb{E}[x^\top \phi_i])^2]$$

$$= \sum_{i=d+1}^m \phi_i^\top \Sigma_x \phi_i$$

argm subject to max. $\frac{\partial \mathbb{E}[(\Delta x)^2]}{\partial \phi_i} = 0$ s.t.
 $\phi_i^\top \phi_i = 1$

one get

$$\mathbb{E}[(\Delta x)^2] = \sum_{i=d+1}^m \lambda_i$$

In order to minimize the representation error
 λ_1 is need to be smallest eigenvalue.

In PCA, we choose ~~the~~ eigenvectors corresponding
to d largest eigenvalues λ_i of the covariance
matrix Σ_x as the principal direction.

$$\frac{|\lambda_1| + |\lambda_2| + \dots + |\lambda_d|}{|\lambda_1| + |\lambda_2| + \dots + |\lambda_m|} \text{ closer to } 1 \text{ gives less errors.}$$

PCA + SVD

data matrix C of size $n \times p$

$$\boxed{\begin{matrix} X_1 & X_2 & \dots & X_p \\ \vdots & \vdots & & \vdots \\ X_n & & & \end{matrix}}$$

Then principal components are coming from eigenvectors of

$$\sum_x = \frac{1}{n-1} C^T C \quad p \times p \text{ matrix}$$

$$\text{SVD of } C \text{ is } C = U S V^T$$

$$\begin{aligned} \sum_x &= \frac{1}{(n-1)} C^T C = \frac{1}{(n-1)} (U S V^T)^T U S V^T \\ &= \frac{1}{(n-1)} V S^T U S V^T = \frac{1}{(n-1)} V S^2 V^T. \end{aligned}$$

The columns of V are the eigenvectors of \sum_x

\therefore If SVD of data matrix is $C = U S V^T$.

Column of V gives principal component/direction

PCA

or $\Sigma + \lambda I$

- ① Calculate covariance matrix of data points
- ② Calculate eigenvectors & corresponding eigenvalues.
- ③ Sort eigenvectors in decreasing value.
- ④ Choose first k eigenvectors & that will be new k dimensions.
- ⑤ Transform original n-dim to k-dimensions.

$$\Sigma = V \Lambda V^T$$

$$\Sigma = V \Lambda V^T \quad \text{or} \quad \Sigma = V \Lambda V^T = V \Lambda^{\frac{1}{2}} V^T$$

$$\Sigma = V \Lambda^{\frac{1}{2}} V^T$$

for information will be contained

$V \Lambda V^T$ - 3 dimensions from 2D

information lost by discarding 3rd dimension