

Combinatorics 2018 Fall

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Recall. $cc(G) = \min \# \text{ cliques in a clique covering of } G$

Theorem. Let G be a graph on n vertices *s.t.* every vertex has at least one neighbor and at most d non-neighbors. Then

$$cc(G) \leq O(d^2 \ln n).$$

proof: Consider the following way of choosing cliques of $G = (V, E)$. First, pick $v \in V$ independently with probability $p = \frac{1}{1+d}$ to get a set $W \subset V$. Then remove from W all vertices having at least one non-neighbor in W . Then we get a clique of G . Apply this way independently t times to get t cliques H_1, \dots, H_t .

Let X be the number of edges not covered by any H_i . Let X_{uw} be the indicator random variable of the event that $u \sim w$ is not covered by any H_i . Then $X = \sum_{u \sim w} X_{uw}$.

Note that H_i covers $u \sim w$ if both u and w are chosen and none of their $\leq 2d$ non-neighbors are chosen. So

$$\begin{aligned} \Pr[u \sim w \text{ is covered by } H_i] &\geq p^2(1-p)^{2d} = \frac{1}{(1+d)^2} \left(\frac{d}{1+d}\right)^{2d} \\ &= \frac{1}{(1+d)^2} \frac{1}{(1+\frac{1}{d})^{2d}} \geq \frac{1}{(1+d)^2 e^2} \end{aligned}$$

So

$$\begin{aligned}
\Pr[u \sim w \text{ is not covered by any } H_i] &= \Pr\left[\bigcap_{i=1}^t \{u \sim w \text{ is not covered by } H_i\}\right] \\
&= \prod_{i=1}^t \Pr[u \sim w \text{ is not covered by } H_i] \\
&\leq \left(1 - \frac{1}{(1+d)^2 e^2}\right)^t \leq e^{-\frac{t}{(1+d)^2 e^2}}
\end{aligned}$$

Then, by linearity of expectation,

$$\begin{aligned}
\mathbb{E}[X] &= \sum_{u \sim w} \mathbb{E}[X_{uw}] = \sum_{u \sim w} \Pr[u \sim w \text{ is not covered by any } H_i] \\
&\leq \binom{n}{2} e^{-\frac{t}{(1+d)^2 e^2}} \leq \frac{e^{2 \ln n}}{2} e^{-\frac{t}{(1+d)^2 e^2}}
\end{aligned}$$

Choose $t \geq 2 \ln n (1+d)^2 e^2$, $t = O(d^2 \ln n)$, we have $\mathbb{E}[X] < 1$. Hence, there is at least one choice of $t = O(d^2 \ln n)$ cliques that form a clique covering of G . $\implies cc(G) \leq O(d^2 \ln n)$. \square

Markov's Inequality. Let $X \geq 0$ be a random variable, $a > 0$, then $\Pr[X \geq a] \leq \frac{\mathbb{E}[X]}{a}$.

proof: Let $I_{\{X \geq a\}}$ be the indicator random variable. Then $a I_{\{X \geq a\}} \leq X$. $\implies \mathbb{E}[a I_{\{X \geq a\}}] = a \Pr[X \geq a] \leq \mathbb{E}[X]$. \square

Corollary. Let $X_n \geq 0$ be integer-valued random variables in (Ω_n, \Pr_n) , $n \in \mathbb{Z}_{>0}$. If $\mathbb{E}[X_n] \rightarrow 0$ as $n \rightarrow \infty$, then $\Pr[X_n = 0] \rightarrow 1$ as $n \rightarrow \infty$.

proof: $\Pr[X_n \geq 1] \leq \mathbb{E}[X_n] \rightarrow 0$ as $n \rightarrow \infty$. \square

Definition. Let $X_n \geq 0$ be integer-valued random variables in (Ω_n, \Pr_n) , $n \in \mathbb{Z}_{>0}$. If $\Pr[X_n = 0] \rightarrow 1$ as $n \rightarrow \infty$, we say $X_n = 0$ almost surely occur.

Definition. The random graph $G(n, p)$ for $0 \leq p \leq 1$ is a graph with vertex set $[n]$, where each of the potential $\binom{n}{2}$ edges appears with probability p independently.

Theorem. $G(n, \frac{1}{2})$ almost surely is NOT bipartite.

proof: Let A_n be the event that $G(n, \frac{1}{2})$ is bipartite. Then it suffices to show $\Pr[A_n] \rightarrow 0$ as $n \rightarrow \infty$.

For $U \in 2^{[n]}$, let A_U be the event that all edges of G are between U and $[n] \setminus U$. Then $A_n = \cup_{U \in 2^{[n]}} A_U$.

$$\Pr[A_U] = \left(\frac{1}{2}\right)^{\binom{|U|}{2}} \left(\frac{1}{2}\right)^{\binom{n-|U|}{2}} = \frac{1}{2^{\binom{|U|}{2} + \binom{n-|U|}{2}}} \leq \frac{1}{2^{2\binom{n/2}{2}}} = 2^{-\frac{n^2}{4} + \frac{n}{2}}$$

So

$$\Pr[A_n] \leq 2^n \cdot 2^{-\frac{n^2}{4} + \frac{n}{2}} = 2^{-\frac{n^2}{4} + \frac{3n}{2}} \rightarrow 0 \text{ as } n \rightarrow \infty$$

□

Theorem. $G(n, p)$ for fixed $p \in (0, 1)$, then

$$\Pr[\alpha(G) \leq \left\lceil \frac{2 \ln n}{p} \right\rceil] \rightarrow 1 \text{ as } n \rightarrow \infty.$$

proof: Let $k = \left\lceil \frac{2 \ln n}{p} \right\rceil$. Let X_n be the number of independent sets

of size $k+1$. For $\forall S \in \binom{[n]}{k+1}$, let $I_{\{S \text{ is independent}\}}$ be the indicator

random variable, then $X_n = \sum_{S \in \binom{[n]}{k+1}} I_{\{S \text{ is independent}\}}$.

$$\mathbb{E}[I_{\{S \text{ is independent}\}}] = \Pr[S \text{ is independent}] = (1-p)^{\binom{k+1}{2}}$$

So,

$$\begin{aligned} \mathbb{E}[X_n] &= \sum_{S \in \binom{[n]}{k+1}} \mathbb{E}[I_{\{S \text{ is independent}\}}] = \binom{n}{k+1} (1-p)^{\binom{k+1}{2}} \\ &\leq \frac{n^{k+1}}{(k+1)!} e^{-p \binom{k+1}{2}} = \frac{(ne^{-p \frac{k}{2}})^{k+1}}{(k+1)!} < \frac{1}{(k+1)!} \rightarrow 0 \text{ as } n \rightarrow \infty \end{aligned}$$

By Corollary, $\Pr[X_n = 0] \rightarrow 1$ as $n \rightarrow \infty$.
 $\implies \Pr[\alpha(G) \leq k] \rightarrow 1$ as $n \rightarrow \infty$. \square

Definition. The girth of G , denoted by $g(G)$, is the length of the shortest cycle in G .

Theorem. For $\forall k \in \mathbb{N}^+$, there exists a graph G with $\chi(G) \geq k$ and $g(G) \geq k$.

proof: Consider $G = G(n, p)$ where p will be determined later.
Let X be the number of cycles of length less than k . Let X_i be the number of cycles of length i . Then $X = \sum_{i=3}^{k-1} X_i$.

$$\begin{aligned} \mathbb{E}[X_i] &= \sum_{(x_1 \dots x_i)} \mathbb{E}[I_{\{(x_1 \dots x_i) \text{ is a cycle}\}}] = \sum_{(x_1 \dots x_i)} \Pr[(x_1 \dots x_i) \text{ is a cycle}] \\ &= \frac{n(n-1) \dots (n-i+1)}{2i} p^i \end{aligned}$$

So

$$\mathbb{E}[X] = \sum_{i=3}^{k-1} \frac{n(n-1) \dots (n-i+1)}{2i} p^i \leq \sum_{i=0}^{k-1} (np)^i = \frac{(np)^k - 1}{np - 1}.$$

By Markov's Inequality,

$$\Pr[X \geq \frac{n}{2}] \leq \frac{\mathbb{E}[X]}{\frac{n}{2}} \leq \frac{2[(np)^k - 1]}{n(np - 1)}.$$

Let $p = n^{-\frac{k-1}{k}}$, then $\Pr[X \geq \frac{n}{2}] \leq \frac{2(n-1)}{n(n^{\frac{1}{k}} - 1)} \rightarrow 0$ as $n \rightarrow \infty$.

Delete a vertex from each cycle of length less than k , we have a graph $G' \subset G$ with no cycle of length less than k .

Note that $|G'| \geq n - \frac{n}{2} = \frac{n}{2}$, $g(G') \geq k$.

Recall that $\chi(G) \geq \frac{|G|}{\alpha(G)}$, and $\Pr[\alpha(G) \leq \left\lceil \frac{2 \ln n}{p} \right\rceil] \rightarrow 1$ as $n \rightarrow \infty$.

So $\alpha(G') \leq \alpha(G) \leq \left\lceil \frac{2 \ln n}{p} \right\rceil \leq 3 \ln n \cdot n^{\frac{k-1}{k}}$.

$\implies \chi(G') \geq \frac{|G'|}{\alpha(G')} \geq \frac{\frac{n}{2}}{3 \ln n \cdot n^{\frac{k-1}{k}}} = \frac{n^{\frac{1}{k}}}{6 \ln n} \gg k$ as $n \rightarrow \infty$. \square