# Web-based Supplementary Materials for "Testing Calibration of Survival

## Models at Extremes of Event Risk"

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## CONTENTS

- 1 Web Appendix A: R code
- 2 Web Figures
- 3 Web Tables

#### 1. Web Appendix A: R code

Although not directly available as output from the coxph() function in the "survival" R software package [R Core Team (2016). R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. URL https://www.R-project.org/],  $\widehat{\Sigma}$  can be obtained as follows: (1) Fit a Cox model corresponding to model (Eq. 1) using coxph() to obtain estimates of the coefficients,  $\widehat{\beta}$ . (2) Substitute these fixed estimates for  $\beta$  in a Cox model corresponding to model (Eq. 2) while also specifying  $\gamma = 0$ ; coefficients can be fixed in a coxph() fit by specifying "iter.max=0". (3) Use the vcov() function to return the inverse of the observed information,  $\widetilde{I}$ , and obtain  $\widehat{\Sigma}$  by taking the inverse of the submatrix with rows and columns corresponding to  $\gamma$ . The following is example R code for this procedure.

#### 

```
### Example R code for obtaining \widehat{\Sigma} in the implementation of the ER and GB tests. ### See Section 2.3 \cos 0 < -\cosh(Surv(time, event) \sim z) \cos 1 < -\cosh(Surv(time, event) \sim z + K, iter.max=0,init=c(cox0$coef,rep(0,D-1))) I.inv <-vcov(cox1) Sigma.hat <-solve(I.inv[row.gamma,col.gamma])
```

### 2. Web Figures

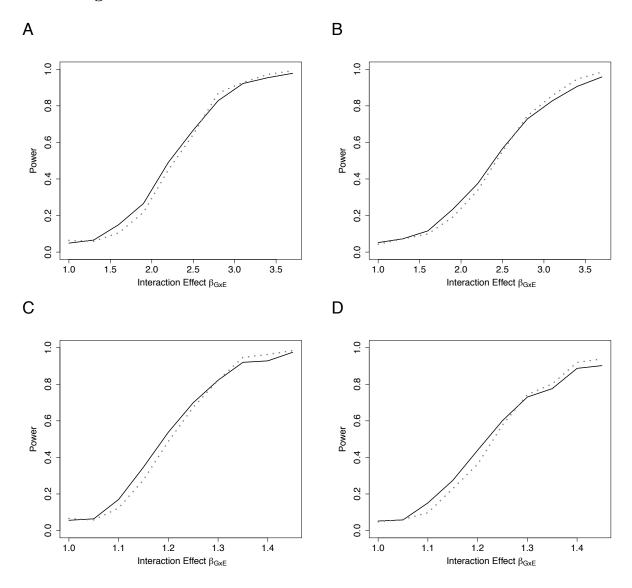


Figure 1. Power of tests under missing interaction models  $g_{02}$ ; n = 1500. Power of the ER (solid line) and GB (dotted line) tests when GxE interaction terms are missing from the fitted Cox model (Eq. 3), but exposure E is included, using  $g_{02}$ . n = 1500 for the 10-SNP model with 1 interaction (**A** and **B**) or 10 interactions (**C** and **D**) in the underlying true model  $g_{A1}$ , and either 0% (**A** and **C**) or 50% (**B** and **D**) lost to follow-up censoring.

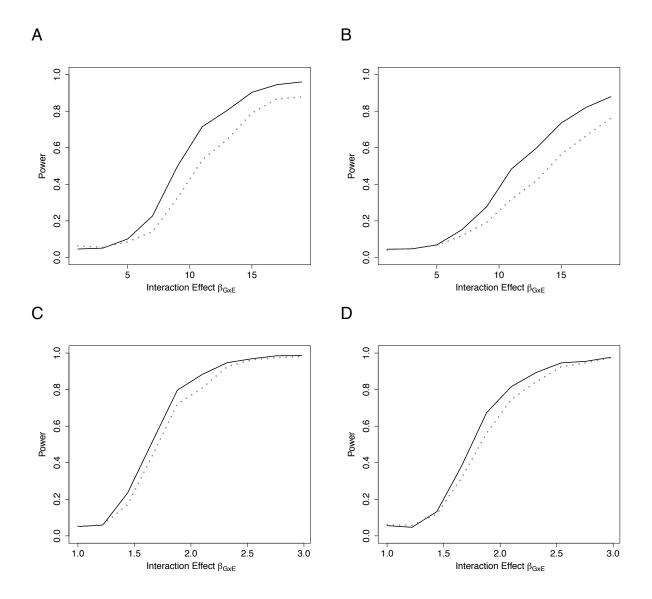


Figure 2. Power of tests under missing interaction models  $g_{01}$ ; n = 1500. Power of the ER (solid line) and GB (dotted line) tests when main effect of E and GxE interaction terms are missing from the fitted Cox model (Eq. 3), using  $g_{01}$ . n = 1500 for the 10-SNP model with 1 interaction (**A** and **B**) or 10 interactions (**C** and **D**) in the underlying true model  $g_{A1}$ , and either 0% (**A** and **C**) or 50% (**B** and **D**) lost to follow-up censoring.

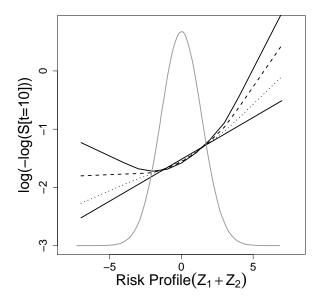


Figure 3. log-log survival probability [t = 10 years] versus risk profile under missing interaction models with Gaussian covariates. Time to event data were simulated under model (Eq. 3) with 2 standard Gaussian covariates  $(Z_1, Z_2)$  and an interaction  $(g_{A3}(\boldsymbol{\beta}, \mathbf{Z}_i) = 1.15Z_1 + 1.15Z_2 + \beta_{Z_1}Z_2Z_1Z_2$  for  $\beta_{Z_1}Z_2 = 1.05$  (dotted line), 1.1 (dashed line) and 1.15 (curved solid line)). Cox models were fit using the null model  $g_{02}(\boldsymbol{\beta}, \mathbf{Z}_i) = \beta_{Z_1}Z_{i1} + \beta_{Z_2}Z_{i2}$  to estimate survival probabilities at the administrative censoring time t = 10 years,  $(\hat{S}_i(t) = \widehat{S}_0(t)^{exp(\hat{\beta}}Z_1Z_{i1}+\widehat{\beta}Z_2Z_{i2})$ ), which were then averaged by risk group (defined by nearest integer of risk profile,  $Z_1 + Z_2$ ) (straight solid line - predicted survival). Observed survival probabilities (curved lines - observed survival) were obtained from a Cox model comparing patients in each risk group with the reference group. The reference group was the subset of individuals with risk profile  $Z_1 + Z_2 = 0$ . The distribution of subjects across the risk profile is shown as a density curve. Survival probabilities estimated for other time points (t < 10 years) yielded similar results.

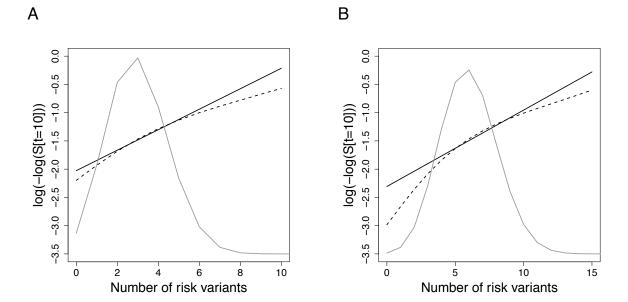


Figure 4. log-log survival probability [t=10 years] versus cumulative number of risk variants under additive effects models. Time to event data were simulated under additive effects, model (Eq. 3), using  $g_{A4}(\beta, G_i) = log(1 + \sum_{j=1}^p \beta_{G_j} G_{ij})$  with (A) p=5 SNPs or (B) p=10 SNPs (MAF=0.3 for each). Data were simulated with a fixed  $\beta_{G_j}$  across all SNPs so that the marginal HR for each SNP in the fitted Cox model was 1.15 and 1.2, for the 10-SNP and 5-SNP models, respectively. Cox models were fit, assuming multiplicative effects on the hazard, using the null model  $g_{01}(\beta, G_i) = \sum_{j=1}^{10} \beta_{G_j} G_{ij}$  to estimate survival probabilities at time t=10 years  $(\hat{S}_i(t) = \widehat{S}_0(t)^{exp(\sum_{j=1}^{10} \widehat{\beta}_{G_j} G_{ij})})$ , which were then averaged by risk group (defined by cumulative number of risk variants) (solid line - predicted survival). Observed survival probabilities (dashed line - observed survival) were obtained from a Cox model comparing patients in each risk group with the reference group. The reference group was the subset of individuals with 3 and 6 risk alleles for the 5-SNP and 10-SNP models respectively. The distribution of subjects per number of risk alleles is shown as a density curve. Survival probabilities estimated for other time points (t < 10 years) yielded similar results.

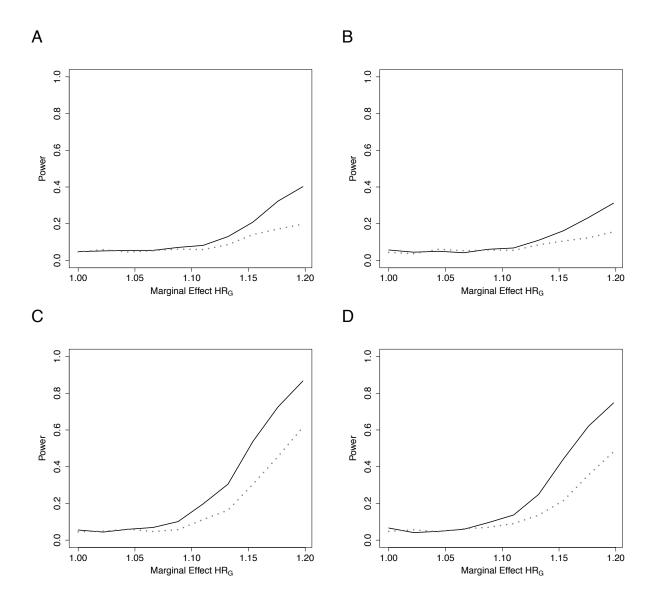


Figure 5. Power of tests under additive effects models with 10 SNPs. Power of the ER (solid line) and GB (dotted line) tests for detecting departures from the multiplicative model with 10 SNPs. Data simulated under additive effects on the HR,  $g_{A4}$ , and fit using the multiplicative effects Cox model (Eq. 3), with  $g_{01}$ . n = 5000 for the 10-SNP model event rate 20% (**A** and **B**) or 50% (**C** and **D**), and either 0% (**A** and **C**) or 50% (**B** and **D**) lost to follow-up censoring.  $G_j \sim Bin(2, 0.3)$ .

### 3. Web Tables

Table 1

Type-1 error of tests for decreasing baseline hazard. Event times simulated under the Weibull hazard model (Eq. 3) with  $g_{01}$  for a decreasing baseline hazard,  $\alpha=0.3$ , with p=5 or 10 covariates (genotypes). Empirical size is presented for the Gronnesby and Borgen test (GB) and the proposed Extreme Risk test (ER). The empirical type 1 error was estimated from 10,000 simulated replicates at the nominal 5% level.

	p = 5				p = 10			
	0% censoring		50% censoring		0% censoring		50% censoring	
Event Rate	GB	$\bar{\mathrm{ER}}$	GB	$\bar{\mathrm{ER}}$	GB	$\overline{\mathrm{ER}}$	$_{\mathrm{GB}}$	$\bar{\mathrm{ER}}$
n = 5000								
0.05	0.050	0.049	0.046	0.050	0.051	0.050	0.051	0.052
0.1	0.052	0.053	0.052	0.049	0.05	0.048	0.050	0.05
0.2	0.051	0.051	0.052	0.052	0.051	0.049	0.050	0.053
n = 1500								
0.05	0.053	0.053	0.049	0.053	0.052	0.048	0.050	0.048
0.1	0.053	0.049	0.052	0.052	0.057	0.052	0.053	0.053
0.2	0.049	0.052	0.049	0.050	0.058	0.051	0.053	0.052

Table 2

Type-1 error of tests for increasing baseline hazard. Event times simulated under the Weibull hazard model (Eq. 3) with  $g_{01}$  for a increasing baseline hazard,  $\alpha=3$ , with p=5 or 10 covariates (genotypes). Empirical size is presented for the Gronnesby and Borgen test (GB) and the proposed Extreme Risk test (ER). The empirical type 1 error was estimated from 10,000 simulated replicates at the nominal 5% level.

	p = 5				p = 10			
	0% censoring		50% censoring		0% censoring		50% censoring	
Event Rate	GB	$\widetilde{\mathrm{ER}}$	GB	$\widetilde{\mathrm{ER}}$	GB	$\widetilde{\mathrm{ER}}$	$_{\mathrm{GB}}$	$\overrightarrow{\mathrm{ER}}$
n = 5000								
0.05	0.052	0.049	0.053	0.052	0.049	0.046	0.051	0.048
0.1	0.050	0.053	0.055	0.054	0.051	0.050	0.053	0.053
0.2	0.051	0.050	0.049	0.050	0.051	0.049	0.054	0.053
n = 1500								
0.05	0.052	0.052	0.049	0.051	0.053	0.049	0.055	0.056
0.1	0.054	0.052	0.052	0.049	0.049	0.054	0.053	0.053
0.2	0.051	0.053	0.056	0.054	0.056	0.053	0.058	0.053

Table 3

Type-1 error of tests applying the 'no less than 5' convention for constant baseline hazard. Event times simulated under the Weibull hazard model (Eq. 3) with  $g_{01}$  for a constant baseline hazard,  $\alpha=1$ , with p=5 or 10 covariates (genotypes). Empirical size is presented for the Gronnesby and Borgen test (GB<sub>adj</sub>) and the proposed Extreme Risk test (ER<sub>adj</sub>). Both tests were applied to the augmented data following collapsing of groups based on the no less than 5 expected events per group convention. The empirical type 1 error was estimated from 10,000 simulated replicates at the nominal 5% level.

	p=5				p = 10			
Event Rate	$0\%$ censoring $GB_{adj}$ $ER_{adj}$		$50\%$ censoring $GB_{adj}$ $ER_{adj}$		$0\%$ censoring $GB_{adj}$ $ER_{adj}$		$50\%$ censoring $GB_{adj}$ $ER_{adj}$	
n = 5000								
0.05	0.049	0.049	0.050	0.049	0.053	0.056	0.054	0.045
0.1	0.050	0.052	0.053	0.049	0.054	0.051	0.052	0.054
0.2	0.052	0.049	0.052	0.05	0.048	0.051	0.052	0.050
n = 1500								
0.05	0.050	0.047	0.052	0.052	0.050	0.048	0.054	0.051
0.1	0.051	0.051	0.048	0.044	0.053	0.049	0.050	0.049
0.2	0.051	0.051	0.052	0.052	0.055	0.053	0.057	0.056