BAYESIAN STATISTICS

Chapter 10

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10. Simultaneous Estimation of Mean and Variance of Normal Population

It is not very common that mean and variance is known *apriori*. Here, we consider the Bayes inference the normal population without the prior information on them.

10.1. Prior distribution for mean and variance without the prior information

Suppose $X \sim N(\mu, \sigma^2) = N(\mu, 1/p_{rec})$ and μ , σ^2 , p_{rec} are unknown. Prior distribution is the joint distribution of (μ, σ^2) or (μ, p_{rec}) , which is specified as the multiplication of the conditional distribution of $\mu|p_{rec}$ and the marginal distribution of p_{rec} .

• conditional prior for $\mu | p_{rec} \colon \mu | p_{rec} \sim {\sf N}\left(\nu, \frac{1}{\pi_{rec} p_{rec}}\right)$

$$g(\mu|
ho_{
m rec}) \propto (\pi_{
m rec}
ho_{
m rec})^{1/2} \exp\left\{-rac{\pi_{
m rec}
ho_{
m rec}(\mu-
u)^2}{2}
ight\}$$

• prior for p_{rec} : $p_{rec} \sim \mathsf{Gamma}(\alpha, \beta)$

$$g(p_{rec}) = rac{1}{\Gamma(lpha)eta^lpha} p_{rec}^{lpha-1} e^{-p_{rec}/eta}$$

• joint prior for (μ, p_{rec}) :

$$\begin{array}{ll} g(\mu, p_{\textit{rec}}) & \propto & g(\mu|p_{\textit{rec}})g(p_{\textit{rec}}) \\ & \propto & (\pi_{\textit{rec}}p_{\textit{rec}})^{1/2}\exp\left\{-\frac{\pi_{\textit{rec}}p_{\textit{rec}}(\mu-\nu)^2}{2}\right\} \times p_{\textit{rec}}^{\alpha-1}\mathrm{e}^{-p_{\textit{rec}}/\beta} \end{array}$$

10.2. Joint posterior for (μ, p_{rec})

The likelihood for a single observation X = x is given as

$$\ell(x|\mu, p_{rec}) \propto p_{rec}^{1/2} \exp\left\{-rac{p_{rec}(x-\mu)^2}{2}
ight\}$$
 .

Thus, the posterior density function becomes

$$\begin{split} h(\mu,\rho_{rec}|x) & \propto \quad \ell(x|\mu,\rho_{rec})g(\mu,\rho_{rec}) \\ & \propto \quad p_{rec}^{1/2} \exp\left\{-\frac{\rho_{rec}(x-\mu)^2}{2}\right\} \\ & \times (\pi_{rec}\rho_{rec})^{1/2} \exp\left\{-\frac{\pi_{rec}\rho_{rec}(\mu-\nu)^2}{2}\right\} \times \rho_{rec}^{\alpha-1} e^{-\rho_{rec}/\beta} \\ & \propto \quad (\pi'_{rec}\rho_{rec})^{1/2} \exp\left\{-\frac{\pi'_{rec}\rho_{rec}(\mu-\nu')^2}{2}\right\} \times \rho_{rec}^{\alpha'-1} e^{-\rho_{rec}/\beta'}, \end{split}$$

where
$$\nu' = (\pi_{rec}\nu + x)/(\pi_{rec} + 1)$$
, $\pi'_{rec} = \pi_{rec} + 1$, $\alpha' = \alpha + 1/2$, and $\beta' = \left\{\frac{1}{\beta} + \frac{\pi_{rec}(x - \nu)^2}{2(\pi_{rec} + 1)}\right\}^{-1}$.

10.3. Marginal posterior distributions

The marginal posterior distribution for p_{rec} is obtained by integrating out the joint posterior with respect to μ :

$$\begin{array}{ll} \textit{h}(\textit{p}_{\textit{rec}}|\textit{x}) & \propto & \int_{-\infty}^{\infty} (\pi'_{\textit{rec}}\textit{p}_{\textit{rec}})^{1/2} \exp\left\{-\frac{\pi'_{\textit{rec}}\textit{p}_{\textit{rec}}(\mu-\nu')^2}{2}\right\} \times \textit{p}_{\textit{rec}}^{\alpha'-1} e^{-\textit{p}_{\textit{rec}}/\beta'} \, d\mu \\ & \propto & \textit{p}_{\textit{rec}}^{\alpha'-1} e^{-\textit{p}_{\textit{rec}}/\beta'} \, \sim \, \mathsf{Gamma}(\alpha',\beta') \end{array}$$

It is little bit complicated to obtain the marginal posterior for μ since integration of the joint posterior with respect to p_{rec} is hard. Instead,

$$h(\mu|x) = \int h(\mu, p_{rec}|x) dp_{rec} = \int h(\mu|p_{rec}, x) h(p_{rec}|x) dp_{rec}.$$

From the joint posterior distribution, we get

$$h(\mu|p_{rec},x) = \frac{h(\mu,p_{rec}|x)}{h(p_{rec}|x)} \propto (\pi_{rec}'p_{rec})^{1/2} \exp\left\{-\frac{\pi_{rec}'p_{rec}(\mu-\nu')^2}{2}\right\}$$

which is $\mu | p_{rec}, x \sim N\left(\nu', \frac{1}{\pi'_{rec}p_{rec}}\right)$.

The marginal posterior for p_{rec} becomes

$$h(\mu|x) = \int_{0}^{\infty} h(\mu|p_{rec}, x) h(p_{rec}|x) dp_{rec}$$

$$= \frac{1}{\sqrt{2\alpha'}B(\alpha', 1/2)} \cdot \frac{\sqrt{\alpha'\beta'\pi'_{rec}}}{\{1 + \beta'\pi'_{rec}(\mu - \nu')^{2}/2\}^{(2\alpha'+1)/2}}$$

with the beta function $B(\alpha,\beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha,\beta)}$. Thus,

$$(\alpha'\beta'\pi'_{rec})^{1/2}(\mu-\nu')|p_{rec},x\sim t(2\alpha').$$

t(r) is t-distribution with a degree of freedom r. Its probability density function is

$$f(t|r) = \frac{1}{\sqrt{r}B\left(\frac{r}{2},\frac{1}{2}\right)} \left(1 + \frac{t^2}{r}\right)^{-(r+1)/2}$$

Suppose the independent samples X_1, \dots, X_n are obtained from $N(\mu, \sigma^2 = 1/p_{rec})$.

prior:

$$g(\mu, p_{\rm rec}) \propto (\pi_{\rm rec} p_{\rm rec})^{1/2} \exp\left\{-\frac{\pi_{\rm rec} p_{\rm rec} (\mu - nu)^2}{2}\right\} \cdot p_{\rm rec}^{\alpha - 1} {\rm e}^{-p_{\rm rec}/\beta}$$

likelihood:

$$\ell(\mathbf{x}|\mu, p_{\text{rec}}) \propto \prod_{i=1}^n p_{\text{rec}}^{1/2} \exp\left\{-\frac{p_{\text{rec}}(\mathbf{x}_i - \mu)^2}{2}\right\} \propto p_{\text{rec}}^{n/2} \exp\left\{-\frac{p_{\text{rec}}\sum_{i=1}^n (\mathbf{x}_i - \mu)^2}{2}\right\}$$

posterior:

$$\begin{split} h(\mu, p_{rec} | \mathbf{x}) &= \ell(\mathbf{x} | \mu, p_{rec}) g(\mu, p_{rec}) \\ &\propto p_{rec}^{n/2} \exp\left\{-\frac{p_{rec} \sum_{i=1}^{n} (x_i - \mu)^2}{2}\right\} \\ &\times (\pi_{rec} p_{rec})^{1/2} \exp\left\{-\frac{\pi_{rec} p_{rec} (\mu - \nu)^2}{2}\right\} \cdot p_{rec}^{\alpha - 1} e^{-p_{rec}/\beta} \\ &\propto (\pi_{rec}^* p_{rec})^{1/2} \exp\left\{-\frac{\pi_{rec}^* p_{rec} (\mu - \nu^*)^2}{2}\right\} \cdot p_{rec}^{\alpha^* - 1} e^{-p_{rec}/\beta^*} \end{split}$$

where
$$\nu^* = \frac{\pi_{rec}\nu + n\bar{\mathbf{x}}}{\pi_{rec} + n}$$
, $\pi^*_{rec} = \pi_{rec} + n$, $\alpha^* = \alpha + n/2$, and
$$\beta^* = \left\{\frac{1}{\beta} + \frac{nS^2}{2} + \frac{\pi_{rec}n(\bar{\mathbf{x}} - \nu)^2}{2(\pi_{rec} + n)}\right\}^{-1} \text{ with } S^2 = \sum_{i=1}^n (x_i - \bar{\mathbf{x}})^2/n.$$

marginal posterior for p_{rec}:

$$\begin{split} \textit{h}(\textit{p}_{\textit{rec}}|\mathbf{x}) & \propto & \int_{-\infty}^{\infty} (\pi_{\textit{rec}}^* \textit{p}_{\textit{rec}})^{1/2} \exp\left\{-\frac{\pi_{\textit{rec}}^* \textit{p}_{\textit{rec}} (\mu - \nu^*)^2}{2}\right\} \cdot \textit{p}_{\textit{rec}}^{\alpha^* - 1} e^{-\textit{p}_{\textit{rec}}/\beta^*} \, d\mu \\ & \propto & \textit{p}_{\textit{rec}}^{\alpha^* - 1} e^{-\textit{p}_{\textit{rec}}/\beta^*} \int_{-\infty}^{\infty} (\pi_{\textit{rec}}^* \textit{p}_{\textit{rec}})^{1/2} \exp\left\{-\frac{\pi_{\textit{rec}}^* \textit{p}_{\textit{rec}} (\mu - \nu^*)^2}{2}\right\} d\mu \\ & \propto & \textit{p}_{\textit{rec}}^{\alpha^* - 1} e^{-\textit{p}_{\textit{rec}}/\beta^*} & \sim \mathsf{Gamma}(\alpha^*, \beta^*). \end{split}$$

ullet marginal posterior for μ : In the similar manner as in the single observation,

$$(lpha^*eta^*\pi^*_{rec})^{1/2}(\mu-
u^*)\sim t(2lpha^*).$$
 (show it!!)

10.5. Bayes estimator

Bayes estimators for μ and p_{rec} are given as the marginal posterior mean.

$$\hat{p}_{rec,B} = \alpha^* \beta^* = \frac{\alpha + n/2}{\frac{1}{\beta} + \frac{nS^2}{2} + \frac{\pi_{rec} n(\bar{x} - \nu)^2}{2(\pi_{rec} + n)}}, \quad \hat{\mu}_B = \nu^* = \frac{\pi_{rec} \nu + n\bar{x}}{\pi_{rec} + n}.$$

Now suppose there is no prior information. To ignore the prior distribution by making it flat or vague, we consider $\alpha \to 0$, $\beta \to \infty$, $\pi_{rec} \rho_{rec} \to 0$ (or $\pi_{rec} \to 0$). This procedure makes the prior as

$$g(\mu, p_{\textit{rec}}) \propto (\pi_{\textit{rec}} p_{\textit{rec}})^{1/2} \exp\left\{-\frac{\pi_{\textit{rec}} p_{\textit{rec}} (\mu - \nu)^2}{2}\right\} \cdot p_{\textit{rec}}^{\alpha - 1} \text{e}^{-p_{\textit{rec}}/\beta} \longrightarrow \text{constant}$$

In this limiting case, the parameters for the posterior distribution become

$$\begin{split} \nu^* &= \frac{\pi_{rec} \nu + n\bar{\mathbf{x}}}{\pi_{rec} + n} \longrightarrow \bar{\mathbf{x}} \\ \pi^*_{rec} &= \pi_{rec} + n \longrightarrow n \\ \alpha^* &= \alpha + n/2 \longrightarrow n/2 \\ \beta^* &= \left\{ \frac{1}{\beta} + \frac{nS^2}{2} + \frac{\pi_{rec} n(\bar{\mathbf{x}} - \nu)^2}{2(\pi_{rec} + n)} \right\}^{-1} \longrightarrow \left\{ \frac{nS^2}{2} \right\}^{-1} \end{split}$$

In the same limiting case, the marginal posteriors are following:

• marginal posterior for μ :

$$\frac{\mu-\bar{x}}{S/\sqrt{n}}\sim t(n)$$

marginal posterior for p_{rec}:

$$p_{rec}|\mathbf{X} \sim \mathsf{Gamma}\left(\frac{n}{2}, \frac{2}{nS^2}\right)$$
 or $nS^2p_{rec}|\mathbf{X} \sim \chi^2(n)$.